

Evaluating Informative Hypotheses With
Equality and Inequality Constraint: The Bayes
Factor Encompassing Prior Approach
Supplemental Material

Claudio Zandonella Callegher, Tatiana Marci, Pietro De Carli, and Gianmarco Altoè

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Introduction

This is the Supplemental Material of the article “*Evaluating informative hypotheses with equality and inequality constraints: A tutorial using the Bayes factor via the encompassing prior approach*” [TODO: add cit]. The paper aims to provide a clear and detailed description of the Bayes Factor with the encompassing prior approach considering an applied example.

Study Summary

When conducting a study, researchers usually have expectations based on hypotheses or theoretical perspectives they want to evaluate. Equality and inequality constraints on the model parameters are used to formalize researchers’ expectations or theoretical perspectives into the so-called informative hypotheses. However, traditional statistical approaches, such as the Null Hypothesis Significance Testing (NHST) or the model comparison using information criteria (e.g., AIC and BIC), are unsuitable for testing complex informative hypotheses. An alternative approach is to use the Bayes factor. In particular, the Bayes factor based on the encompassing prior approach allows researchers to easily evaluate complex informative hypotheses in a wide range of statistical models (e.g., generalized linear). This paper provides a detailed introduction to the Bayes factor with encompassing prior. First, all steps and elements involved in the formalization of informative hypotheses and the computation of the Bayes factor with encompassing prior are described. Next, we apply this method to a real case scenario, considering the attachment theory. Specifically, we analyzed the relative influence of maternal and paternal attachment on children’s social-emotional development by comparing the various theoretical perspectives debated in the literature.

Supplemental Material Structure

In the supplemental material, we provide the complete statistical analyses of the case study regarding attachment. For more information regarding the Bayes factor with encompassing prior approach or the formalization of theoretical per-

spectives regarding the role of mother attachment and father attachment on children socio-emotional development, see the main article [TODO: add link article]. The supplemental material is structured into three parts:

1) Presentation

In this section, we describe the aim of the study and the sample included in the study.

- **Chapter 1 - Case Study.** We describe the sample included in the study and present the results of the cluster analysis to classify the different attachment patterns.

2) Externalizing Problems

In this section, we present the analyses evaluating the different roles of mother attachment and father attachment on children's externalizing problems. We follow three different approaches.

- **Chapter 2 - Models Family Choice.** We discuss the appropriate models' family to take into account data characteristics.
- **Chapter 3 - NHST.** Analysis is conducted following the traditional Null Hypothesis Significance Testing (NHST).
- **Chapter 4 - Model Comparison.** Analysis is conducted according to the Model Comparison approach using the AIC and BIC criteria.
- **Chapter 5 - Bayes Factor.** Analysis is conducted following the Bayes Factor with the encompassing prior approach.
- **Chapter 6 - Conclusions.** Results of the three approaches are discussed.

3) Internalizing Problems

In this section, we present the analyses evaluating the different roles of mother attachment and father attachment on children's internalizing problems. We follow the three different approaches like in the analysis of the externalizing problems. This section is shorter than the previous one as we focus only on the results without repeating all the descriptions of the methods.

- **Chapter 7 - Models Family Choice.** We discuss the appropriate models' family to take into account data characteristics.
- **Chapter 8 - NHST.** Analysis is conducted following the traditional Null Hypothesis Significance Testing (NHST).
- **Chapter 9 - Model Comparison.** Analysis is conducted according to the Model Comparison approach using the AIC and BIC criteria.
- **Chapter 10 - Bayes Factor.** Analysis is conducted following the Bayes Factor with the encompassing prior approach.
- **Chapter 11 - Conclusions.** Results of the three approaches are discussed.

GitHub

All the materials are available at the GitHub repository <https://github.com/ClaudioZandonella/Attachment>.

Refer to the **README** file for further information.

Presentation

Chapter 1

Case Study

In this chapter, first, we describe the aim of the study. Subsequently, we describe the sample included in the study. We present general information about the sample and the results of the cluster analysis to classify the different attachment patterns. Finally, we consider children’s social-emotional problems.

1.1 Aim of the Study

The study aims to compare the different theoretical perspectives regarding the role of mother attachment and father attachment on children’s social-emotional development. In the literature, four different main theoretical perspectives have been identified [Bretherton]:

- **Monotropy Theory** - only the mother has an impact on children’s development.
- **Hierarchy Theory** - the mother has a greater impact on children’s development than the father.
- **Independence Theory** - all attachment figures are equally important but they affect the children’s development differently.
- **Integration Theory** - to understand the impact on children’s development it is necessary to consider attachment relationships taken together.

Contrasting results have been reported by studies investigating which is the “correct” theory. No study, however, has tried to properly evaluate the different theoretical perspectives by directly comparing the different hypotheses.

For more information regarding Attachment theory and the formalization of theoretical perspectives regarding the role of mother attachment and father attachment on children’s socio-emotional development, see the main article [TODO: add link article].

1.2 The Study Sample

To evaluate the different roles of father attachment and mother attachment, we consider 854 Italian children from third to sixth school grade. Only middle-childhood children are included in the analysis (5 participants are excluded as older than 12.3 years and 2 participants are excluded as the age is missing). The total sample size included in the analysis is 847 (50.65 % Females). Participants' frequencies by grade and gender are reported in Table~1.1.

Table 1.1: Participants frequencies by grade and gender ($n_{subj} = 847$).

Grade	Gender		Total
	Females	Males	
3rd	133	127	260
4th	81	75	156
5th	36	32	68
6th	179	184	363
Total	429	418	847

Participants' age mean is 10.28 (SD = 1.34). The participants' age summary statistics are reported below and the distribution is presented in Figure~1.1.

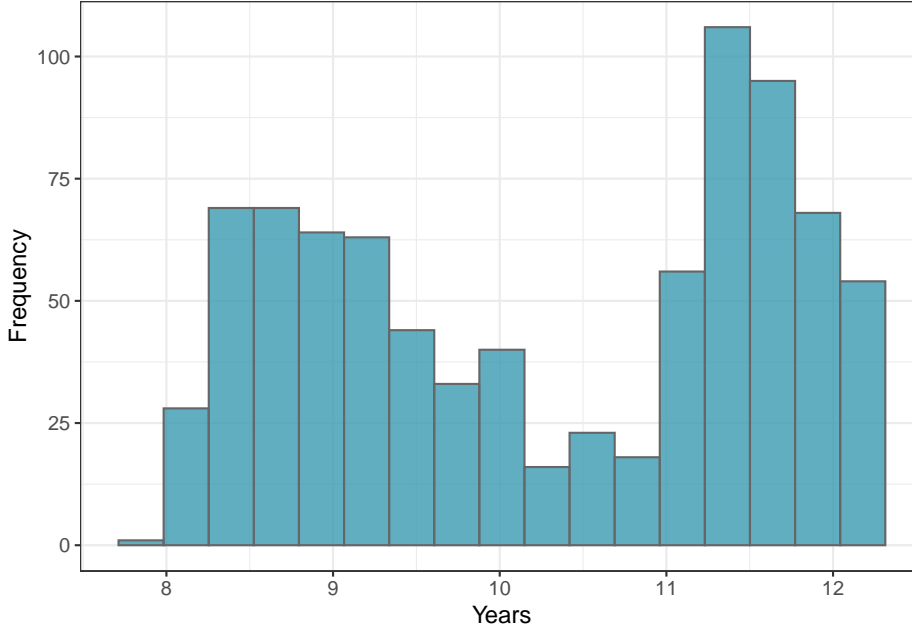
##	# Age summary statistics					
##	Min.	1st Qu.	Median	Mean	3rd Qu.	Max.
##	7.92	9.00	10.25	10.28	11.58	12.25

1.3 Attachment Styles

Attachment towards the mother and the father was measured separately using the brief Experiences in Close Relationships Scale - Revised Child Version [ECR-RC, Brenning et al.; marciBriefExperiencesClose2019] completed by the children. Four main attachment styles have been recognized in the literature according to three levels of *anxiety* and *avoidance*:

- **Secure Attachment** - children with low levels of anxiety and avoidance.
- **Anxious Attachment** - children characterized by high levels of anxiety.
- **Avoidant Attachment** - children characterized by high levels of avoidance.
- **Fearful Attachment** - children with high levels of anxiety and avoidance.

To identify children's attachment styles towards the mother and towards the father, we conduct two separate cluster analyses. Clusters are obtained using the function `hclust()` (with argument `method="ward.D"`) considering *Euclidean* distances between participants responses to the ECR items.

Figure 1.1: Participants age distribution ($n_{subj} = 847$)

1.3.1 Mother Cluster Analysis

Regarding mother attachment, 4 groups are selected from the cluster analysis results (see Figure~1.2).

In Figure~1.3, Anxiety (“*Anx*”) and Avoidance (“*Av*”) scores are presented according to mother attachment styles. The frequencies of mother attachment styles according to gender are reported in Table~1.2.

Table 1.2: Mother attachment styles by gender ($n_{subj} = 847$).

Gender	Attachment Style			
	Secure	Anxious	Avoidant	Fearful
Females	132 (16%)	131 (15%)	111 (13%)	55 (6%)
Males	99 (12%)	155 (18%)	119 (14%)	45 (5%)
Total	231 (27%)	286 (34%)	230 (27%)	100 (12%)

Compared to the overall average values of anxiety and avoidance, we can observe that:

- *Secure* children have lower levels of anxiety and avoidance
- *Anxious* children have higher levels of anxiety and about the average levels

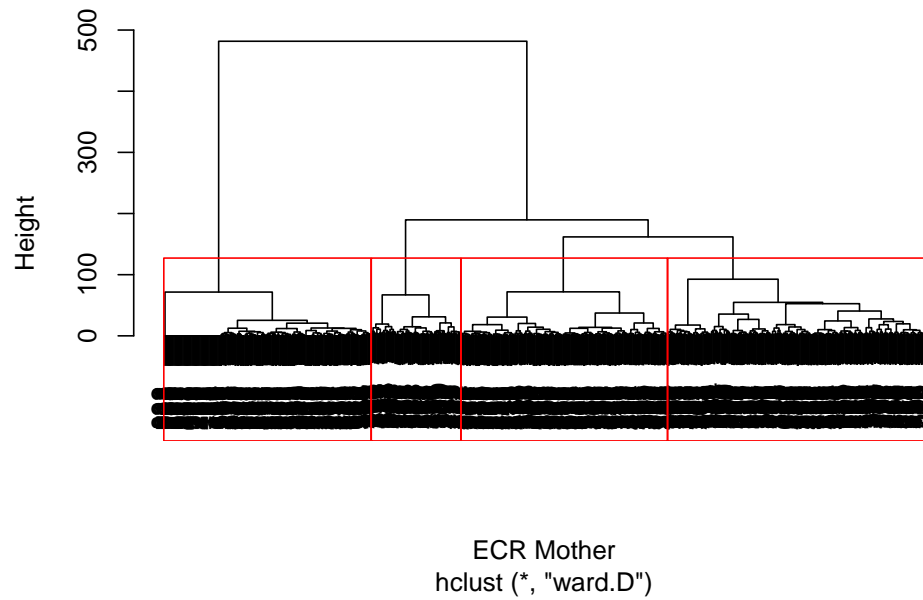


Figure 1.2: Mother attachment cluster dendrogram ($n_{subj} = 847$).

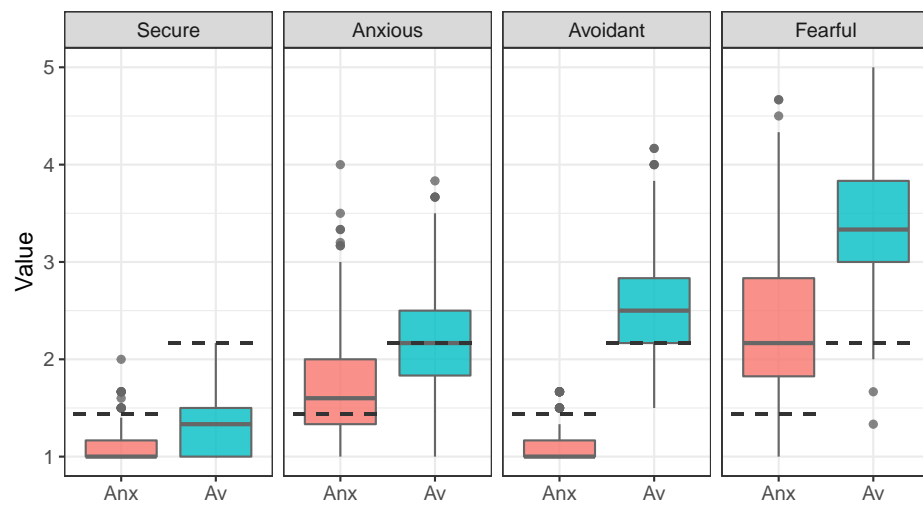


Figure 1.3: Anxiety (“Anx”) and Avoidance (“Av”) scores according to mother attachment styles. Dashed lines represents overall average values ($n_{subj} = 847$).

of avoidance

- *Avoidant* children have lower levels of anxiety and higher levels of avoidance
- *Fearful* children have higher levels of anxiety and avoidance

These results reflect the traditional definition of the four attachment styles.

Mclust Check

To evaluate if 4 clusters is an appropriate choice, we compare the BIC values of different model-based clustering options. To do that we use the function `mclustBIC()` from the `mclust` R-package [Scrucca et al.]. BIC values of the best three model-based clusterings are reported below and overall results are presented in Figure~1.4. See `?mclustBIC()` help page for further information.

```
## # Best model-based clustering for mother attachment
## Best BIC values:
##          EEV,6      EEV,7      EEV,8
## BIC      -23978.93 -24842.2437 -25039.952
## BIC diff      0.00   -863.3168  -1061.025
```

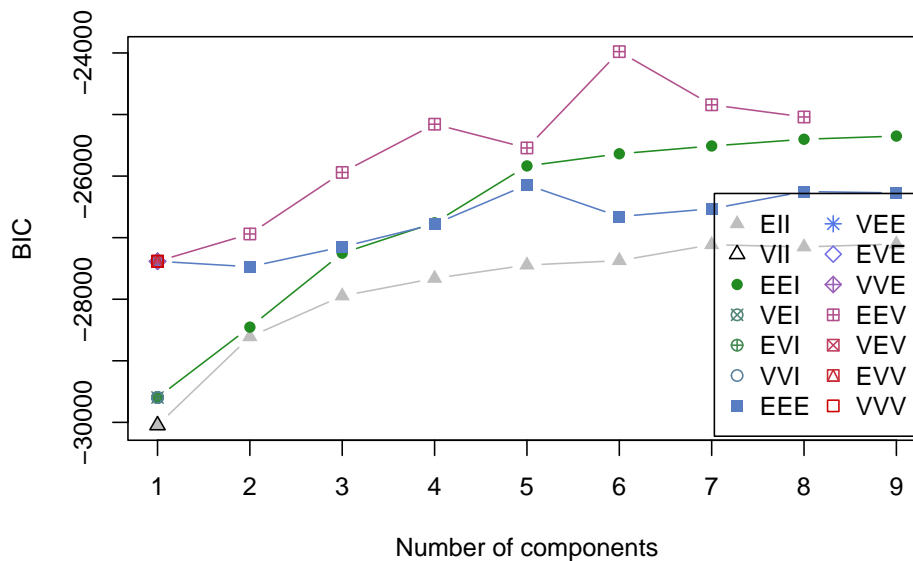


Figure 1.4: BIC values of different model-based clustering options. See `?mclustBIC()` help page for further information.

Results indicate the possible presence of a larger number of clusters. However, considering only 4 clusters seems to us the most reasonable choice. This is in line with attachment theory and general results in the literature. The important thing is that no smaller number of clusters provided better results than the division in 4 clusters.

1.3.2 Father Cluster Analysis

Regarding father attachment, 4 groups are selected from the cluster analysis results (see Figure~1.5).

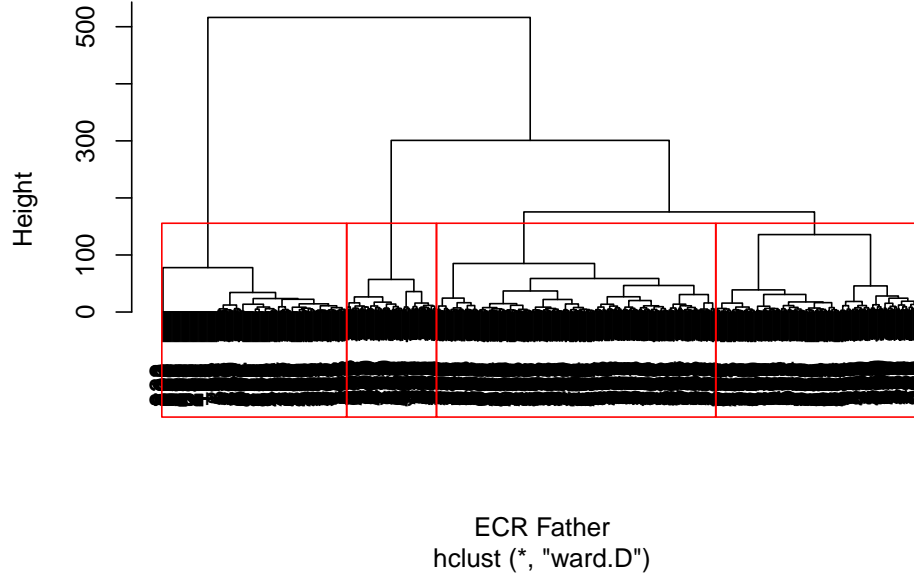


Figure 1.5: Father attachment cluster dendrogram ($n_{subj} = 847$).

In Figure~1.6, Anxiety (“*Anx*”) and Avoidance (“*Av*”) scores are presented according to father attachment styles. The frequencies of father attachment styles according to gender are reported in Table~1.3.

Table 1.3: Father attachment styles by gender ($n_{subj} = 847$).

Gender	Attachment Style			
	Secure	Anxious	Avoidant	Fearful
Females	99 (12%)	104 (12%)	163 (19%)	63 (7%)
Males	107 (13%)	126 (15%)	148 (17%)	37 (4%)
Total	206 (24%)	230 (27%)	311 (37%)	100 (12%)

Compared to the overall average values of anxiety and avoidance, we can observe that:

- *Secure* children have lower levels of anxiety and avoidance
- *Anxious* children have lower levels of anxiety and avoidance
- *Avoidant* children have lower levels of anxiety and higher levels of avoidance

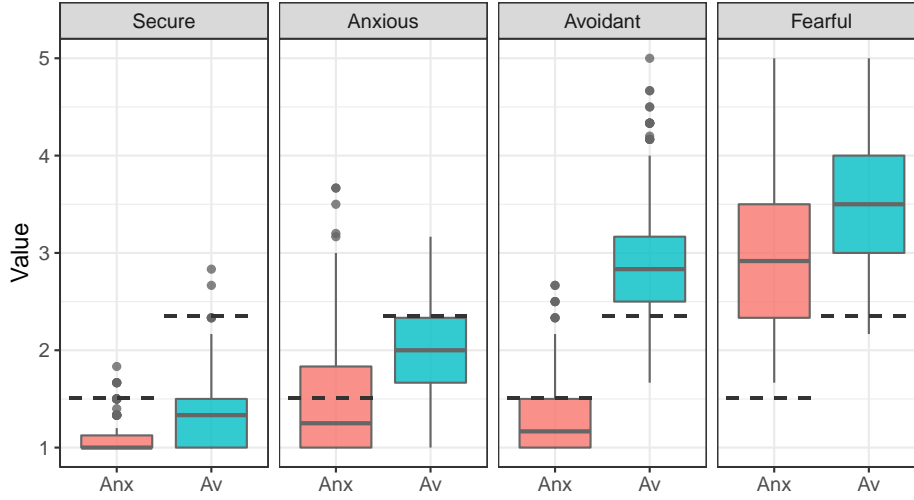


Figure 1.6: Anxiety (“Anx”) and Avoidance (“Av”) scores according to father attachment styles. Dashed lines represents overall average values ($n_{subj} = 847$).

- *Fearful* children have higher levels of anxiety and avoidance

These results are not as good as in the mother attachment classification, but they are still reasonable.

Mclust Check

To evaluate if 4 clusters is an appropriate choice, we compare the BIC values of different model-based clustering options as before. BIC values of the best three model-based clusterings are reported below and overall results are presented in Figure~1.7.

```
## # Best model-based clustering for father attachment
## Best BIC values:
##          EEV,9      EEV,8      EEV,5
## BIC      -25351.43 -25790.9130 -25905.197
## BIC diff         0.00   -439.4807   -553.765
```

As for the mother, results indicate the possible presence of a larger number of clusters. However, considering only 4 clusters seems to us the most reasonable choice for the same reasons as before.

1.3.3 Mother and Father

The frequencies of mother attachment and father attachment styles are reported in Table~1.4.

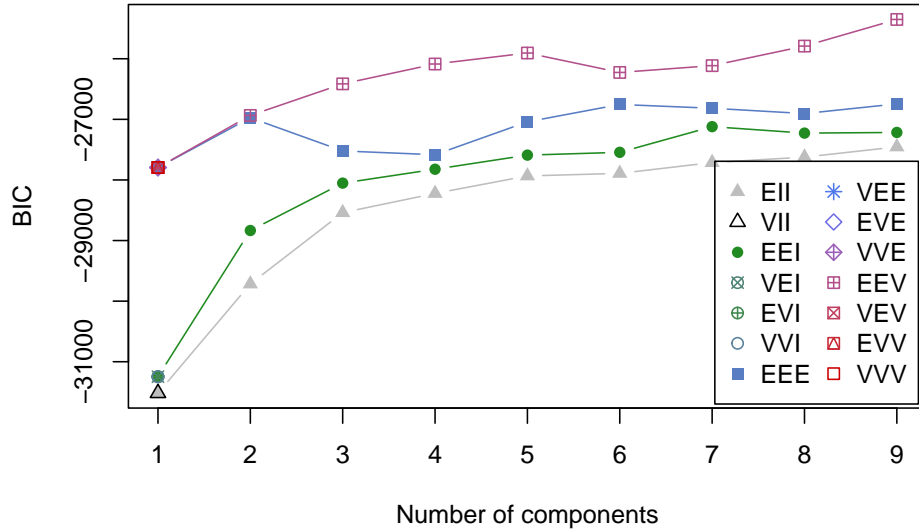


Figure 1.7: BIC values of different model-based clustering options. See `?mclust-BIC()` help page for further information.

Table 1.4: Mother attachment and father attachment styles ($n_{subj} = 847$).

Mother Attachment	Father Attachment			
	Secure	Anxious	Avoidant	Fearful
Secure	125 (15%)	49 (6%)	49 (6%)	8 (1%)
Anxious	51 (6%)	100 (12%)	98 (12%)	37 (4%)
Avoidant	25 (3%)	67 (8%)	126 (15%)	12 (1%)
Fearful	5 (1%)	14 (2%)	38 (4%)	43 (5%)

We can observe how values on the diagonal (same attachment styles towards both parents) tend to be larger than others.

1.4 Children Outcomes

Children's social-emotional development was measured using the Strength and Difficulties Questionnaire [SDQ, Goodman et al.] completed by the teachers. Separate scores for externalizing and internalizing problems were obtained as the sum of the questionnaire items.

1.4.1 Externalizing Problems

Participants externalizing problems mean is 3.35 (SD = 3.91). The participants' externalizing problems summary statistics are reported below and the distribution is presented in Figure~1.8.

```
## # Externalizing problems summary statistics
##   Min. 1st Qu.  Median    Mean 3rd Qu.    Max.
##  0.000  0.000   2.000   3.349  6.000  18.000
```

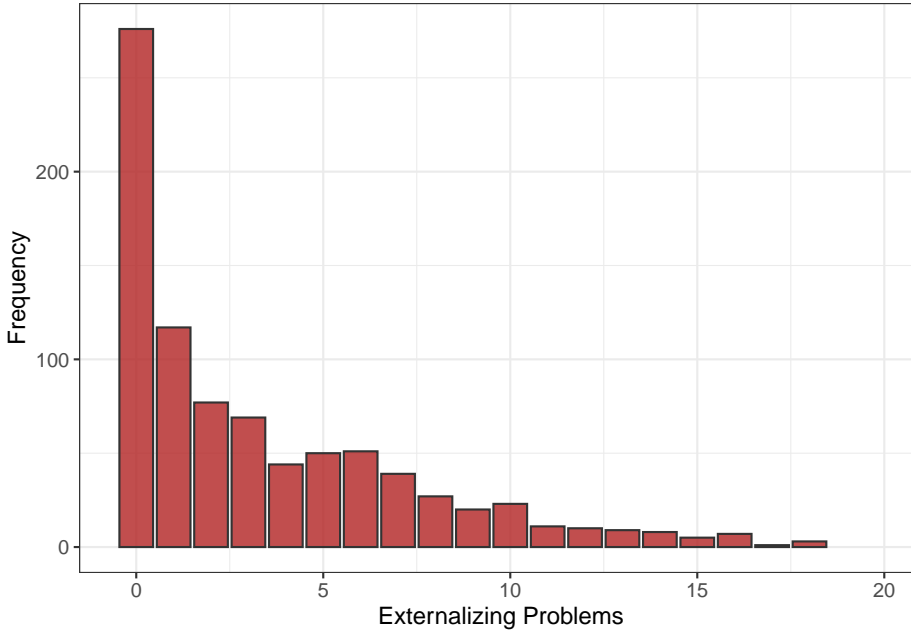


Figure 1.8: Participants externalizing problems distribution ($n_{subj} = 847$)

Overall, externalizing problems are low. This is expected as the sample is not clinical. Externalizing problems according to attachment styles are reported in Table~1.5.

Table 1.5: Externalizing problems according to attachment styles ($n_{subj} = 847$).

Mother Attachment	Father Attachment							
	Secure		Anxious		Avoidant		Fearful	
	Mean (SD)	Median	Mean (SD)	Median	Mean (SD)	Median	Mean (SD)	Median
Secure	2.63 (3.57)	1.0	3.45 (4.48)	2.0	1.61 (2.13)	1.0	2.88 (3.44)	2.0
Anxious	3.69 (4.07)	2.0	3.01 (3.61)	2.0	3.32 (4.12)	2.0	4.05 (3.61)	3.0
Avoidant	2.84 (3.34)	1.0	3.31 (3.65)	2.0	3.71 (4.19)	2.0	3.75 (4.81)	1.0
Fearful	7.60 (4.04)	8.0	4.64 (3.84)	4.5	4.76 (4.67)	3.0	4.26 (4.07)	4.0

1.4.2 Internalizing Problems

Participants internalizing problems mean is 2.96 (SD = 3.14). The participants internalizing problems summary statistics are reported below and the distribution is presented in Figure~1.9.

```
## # Internalizing problems summary statistics
## Min. 1st Qu. Median Mean 3rd Qu. Max.
## 0.000 0.000 2.000 2.955 4.000 18.000
```

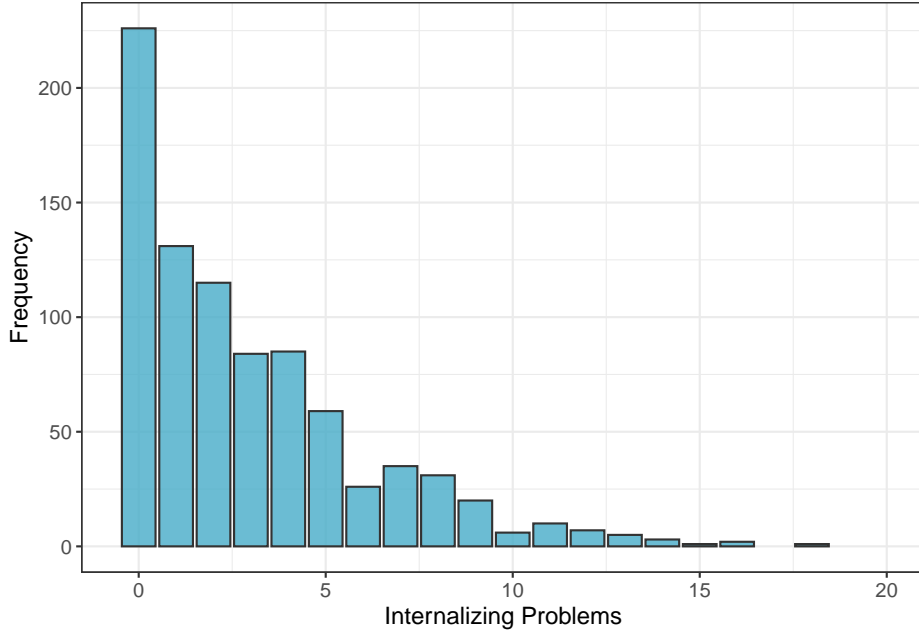


Figure 1.9: Participants internalizing problems distribution ($n_{subj} = 847$)

Overall, internalizing problems are low. Again, this is expected as the sample is not clinical. Internalizing problems according to attachment styles are reported in Table~??tab:table-cluster-ext).

Table 1.6: Internalizing problems according to attachment styles ($n_{subj} = 847$).

Mother Attachment	Father Attachment							
	Secure		Anxious		Avoidant		Fearful	
	Mean (SD)	Median	Mean (SD)	Median	Mean (SD)	Median	Mean (SD)	Median
Secure	2.29 (2.64)	1.0	2.63 (2.75)	2.0	2.67 (2.75)	3.0	3.62 (3.46)	3.0
Anxious	3.29 (3.48)	2.0	3.15 (3.09)	2.0	3.28 (3.72)	2.0	3.11 (3.81)	2.0
Avoidant	1.88 (2.28)	1.0	2.99 (3.22)	2.0	2.65 (2.86)	2.0	4.08 (3.78)	3.5
Fearful	4.40 (3.65)	4.0	2.07 (1.69)	2.0	3.68 (3.68)	3.0	4.37 (3.12)	4.0

Externalizing Problems

Chapter 2

Models Family Choice

In this chapter, we discuss the appropriate models' family to take into account data characteristics.

2.1 Negative Binomial

Externalizing problems are computed as the sum of 10 items of the SDQ, obtaining discrete scores that range from 0 to 20. Thus, we should use appropriate discrete distribution such as the *Poisson* distribution or the *Negative Binomial*. In the Poisson distribution mean and variance are defined according to the same parameter λ . On the contrary, Negative Binomial has an extra parameter to adjust the variance allowing more flexibility. Considering data distribution (see Figure~1.8), we can observe that data have high dispersion with a long right tail. In this case, the Poisson distribution would be a poor choice and we prefer Negative Binomial instead.

Again, considering data distribution (see Figure~1.8), we can observe a high peak of values at zero. Remember that this is not a clinical sample, thus it is expected that the majority of children have no problems or really few problems. We could question ourselves, however, whether a *Zero-Inflated* model may be appropriate

2.2 Zero Inflated Negative Binomial

To evaluate the presence of zero inflation in our data, we compare the number of observed zeros and expected zeros in a Negative Binomial mixed-effects model. We consider in the model the role of gender and the interaction between mother attachment and father attachment. Moreover, we consider the children's classroom ID as a random effect to account for teachers' different ability to evaluate children's problems. Using R formula syntax, we have

```
# model formula
externalizing_sum ~ gender + mother * father + (1|ID_class)
```

The model is fitted using the function `glmmTMB()` from the `glmmTMB` R-package [Brooks et al.]. Next, we compare the number of observed zero and expected zeros using an adapted version of the function `check_zeroinflation()` from the R-package `performance` [Lüdtke et al.] that solves a small bug (see issue <https://github.com/easystats/performance/issues/367>).

```
my_check_zeroinflation(fit_ext_nb)
## # Check for zero-inflation
##
##      Observed zeros: 276
##      Predicted zeros: 245
##              Ratio: 0.89
## Model is underfitting zeros (probable zero-inflation).
```

Results indicate that the model is slightly under-fitting the number of zeros. Now, we can try to fit a *Zero Inflated Negative Binomial* (ZINB) model and compare the performance of the two models. ZINB models are defined as

$$y_{ij} \sim ZINB(p_{ij}, \mu_{ij}, \phi),$$

where p_{ij} is the probability of an observation y_{ij} being an extra zero (i.e., a zero not coming from the Negative Binomial distribution) and $1 - p_{ij}$ indicates the probability of a given observation y_{ij} being generated from a Negative Binomial distribution with mean μ_{ij} and variance $\sigma_{ij}^2 = \mu_{ij} + \frac{\mu_{ij}^2}{\phi}$. Moreover, we have that

$$p_{ij} = \text{logit}^{-1}(X_i^T \beta_p + Z_j^T u_p), \mu_{ij} = \exp(X_i^T \beta_\mu + Z_j^T u_\mu).$$

That is, both p and μ are modelled separately according to (possibly) different variables. In our case, we consider only the role of gender for p (i.e., the probability of having externalizing problems depends on gender), whereas for μ we also consider the interaction between mother attachment and father attachment. In both cases, we consider the children's classroom ID as a random effect (teachers may differ in the ability to detect children's problems and quantify them). Using R formula syntax, we have

```
# formula for p
p ~ gender + (1|ID_class)

# formula for mu
mu ~ gender + mother * father + (1|ID_class)
```

The ZINB model is fitted using the function `glmmTMB()`. To compare the ZINB model and the Negative Binomial model we conduct an analysis of *Deviance*. Note that, in the case of generalized linear models (GLM), the deviance is the

corresponding of the residual variance used in the traditional ANOVA in the case of linear models.

```
anova(fit_ext_nb, fit_ext_zinb)
## Data: data_cluster
## Models:
## fit_ext_nb: externalizing_sum ~ gender + mother * father + (1 | ID_class), zi=~0, disp=~1
## fit_ext_zinb: externalizing_sum ~ gender + mother * father + (1 | ID_class), zi=~gender + (1 |
##           Df      AIC      BIC logLik deviance  Chisq Chi Df Pr(>Chisq)
## fit_ext_nb   19 3910.1 4000.2 -1936.1   3872.1
## fit_ext_zinb 22 3867.6 3971.9 -1911.8   3823.6 48.556      3 1.622e-10 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

Overall, results indicate that the ZINB model performs better than the Negative Binomial model. Thus, in the following analyses, we decide to use ZINB models.

Chapter 3

NHST

Following the traditional NHST approach, we consider the model previously defined that includes all effects of interest. That is the gender effect and the interaction between mother attachment and father attachment. Subsequently, we can run an analysis of deviance to evaluate the significance of the predictors using the function `Anova()` from the R-package `car` [Fox and Weisberg].

```
car::Anova(fit_ext_zinb)
## Analysis of Deviance Table (Type II Wald chisquare tests)
##
## Response: externalizing_sum
##           Chisq Df Pr(>Chisq)
## gender      15.2947  1 9.197e-05 ***
## mother      15.6195  3  0.001357 **
## father       1.1006  3  0.776938
## mother:father  8.9096  9  0.445657
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

Results indicate a statistically significant effect of gender and mother attachment. On the contrary, the interaction and father attachment are not significant. The model summary is reported below.

```
summary(fit_ext_zinb)
## Family: nbinom2 ( log )
## Formula:      externalizing_sum ~ gender + mother * father + (1 | ID_class)
## Zero inflation: ~gender + (1 | ID_class)
## Data: data_cluster
##
##      AIC      BIC   logLik deviance df.resid
##  3867.6   3971.9 -1911.8   3823.6      825
```

```
##
## Random effects:
##
## Conditional model:
## Groups   Name          Variance Std.Dev.
## ID_class (Intercept) 0.08132  0.2852
## Number of obs: 847, groups: ID_class, 50
##
## Zero-inflation model:
## Groups   Name          Variance Std.Dev.
## ID_class (Intercept) 0.7669   0.8757
## Number of obs: 847, groups: ID_class, 50
##
## Dispersion parameter for nbinom2 family (): 1.86
##
## Conditional model:
##                                     Estimate Std. Error z value Pr(>|z|)
## (Intercept)                      1.01996    0.12814   7.960 1.72e-15 ***
## genderM                          0.31556    0.08069   3.911 9.20e-05 ***
## motherAnxious                    0.30735    0.18664   1.647  0.0996 .
## motherAvoidant                   0.14496    0.25509   0.568  0.5699
## motherFearful                    0.62950    0.39550   1.592  0.1115
## fatherAnxious                    0.21543    0.19156   1.125  0.2607
## fatherAvoidant                   -0.46555    0.21138  -2.202  0.0276 *
## fatherFearful                    0.07138    0.41737   0.171  0.8642
## motherAnxious:fatherAnxious      -0.38346    0.26981  -1.421  0.1553
## motherAvoidant:fatherAnxious     -0.09675    0.33041  -0.293  0.7697
## motherFearful:fatherAnxious      -0.40811    0.49348  -0.827  0.4082
## motherAnxious:fatherAvoidant      0.34281    0.28290   1.212  0.2256
## motherAvoidant:fatherAvoidant     0.63111    0.32765   1.926  0.0541 .
## motherFearful:fatherAvoidant      0.34025    0.46490   0.732  0.4642
## motherAnxious:fatherFearful      -0.04630    0.47448  -0.098  0.9223
## motherAvoidant:fatherFearful      0.11794    0.56413   0.209  0.8344
## motherFearful:fatherFearful      -0.20316    0.58602  -0.347  0.7288
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Zero-inflation model:
##                                     Estimate Std. Error z value Pr(>|z|)
## (Intercept)                    -1.1239     0.2417   -4.65 3.32e-06 ***
## genderM                       -0.7145     0.2516   -2.84 0.00451 **
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

To evaluate the effect of gender and mother attachment, the marginal predicted

values according to gender and mother attachment are presented separately in Figure~3.1. Not that the marginal predicted values for gender are averaged over mother and father attachment effects. Whereas, the marginal predicted values for mother attachment are averaged over father attachment and gender effect.

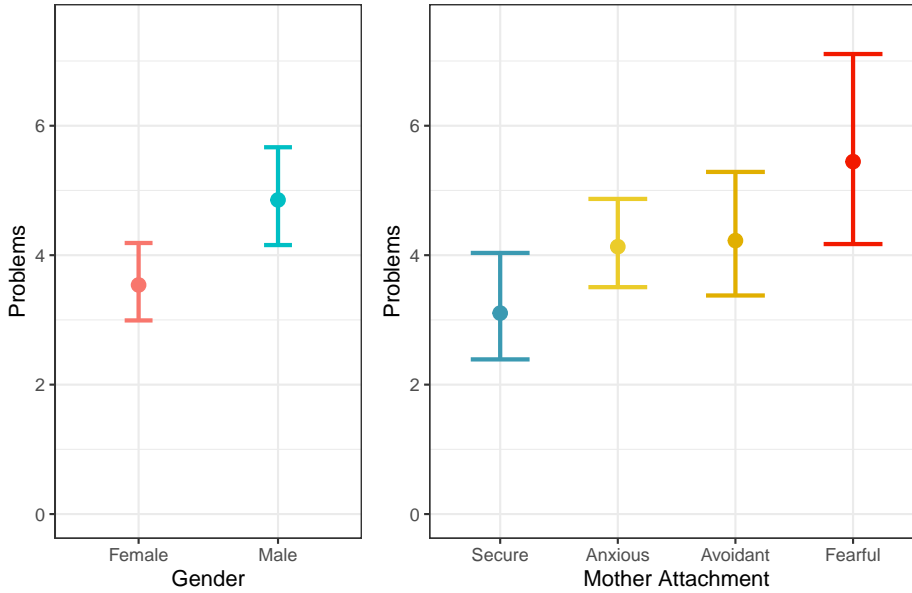


Figure 3.1: Marginal predicted values according to gender and mother attachment. Values are averaged over the other effects ($n_{subj} = 847$).

Post-hoc tests are run to evaluate differences between mother attachment styles. To do that we use the `contrast()` function from the `emmeans` R-package, considering pairwise comparisons and adjusting p -values according to multivariate t -distribution. This approach is less restrictive than the traditional “Bonferroni” method, as it determines the adjustment according to a multivariate t -distribution with the same covariance structure as the estimates. Results are reported below,

```
emmeans::contrast(emmeans::emmeans(fit_ext_zinb, specs = ~ mother ),
  "pairwise", adjust = "mvt")
## NOTE: Results may be misleading due to involvement in interactions
## contrast      estimate    SE df t.ratio p.value
## Secure - Anxious   -0.2856 0.140 825  -2.034  0.1720
## Secure - Avoidant  -0.3080 0.162 825  -1.902  0.2228
## Secure - Fearful   -0.5617 0.178 825  -3.149  0.0089
## Anxious - Avoidant -0.0224 0.124 825  -0.181  0.9978
## Anxious - Fearful  -0.2761 0.146 825  -1.885  0.2300
## Avoidant - Fearful -0.2537 0.167 825  -1.522  0.4180
##
```

```
## Results are averaged over the levels of: gender, father
## Results are given on the log (not the response) scale.
## P value adjustment: mvt method for 6 tests
```

Overall, results indicate that Males have more externalizing problems than Females and, regarding mother attachment, Fearful children have more problems than Secure children.

To evaluate the fit of the model to the data, we used R^2 . In the case of generalized mixed-effects models, however, there are several definitions of R^2 . We computed the *Marginal* R^2 and the *Conditional* R^2 as suggested by Nakagawa et al.. *Marginal* R^2 is concerned with the variance explained by fixed factors of the model, and *Conditional* R^2 is concerned with the variance explained by both fixed and random factors of the model. To do that we use the function `performance::r2()`.

```
performance::r2(fit_ext_zinb)
## Warning: mu of 4.2 is too close to zero, estimate of random effect variances may be
##   unreliable.
## # R2 for Mixed Models
##
##   Conditional R2: 0.191
##   Marginal R2: 0.091
```

We can see that the actual variance explained by fixed effects is almost 10%, not bad for psychology.

Conclusions

Considering attachment theoretical perspectives, results indicate only the role of mother attachment. Note, however, that traditional NHST does not allow us to evaluate evidence in favour of a hypothesis. Moreover, we actually have not tested our hypotheses but only the catch-all null hypothesis that “*nothing is going on*”.

Chapter 4

Model Comparison

Model comparison allows us to compare multiple hypotheses and identify which is the most supported by the data [McElreath]. First, we need to formalize models according to our hypotheses. Subsequently, we can evaluate which is the most supported model among those considered according to the data using the AIC and BIC [Wagenmakers and Farrell, Akaike, Schwarz].

4.1 Formalize Models

Following the same reasons as before (see Section~2), we consider Zero Inflated Negative Binomial Mixed-Effects models. Again, we consider only the role of gender as a fixed effect and children's classroom ID as a random effect for p . Whereas, considering μ , we define four different models to take into account the different theoretical perspectives:

- `fit_ext_zero`: we consider only the effect of gender. This model assumes that attachment plays no role.
- `fit_ext_mother`: we consider the additive effects of gender and mother attachment. This model supports the idea that only mother attachment is important (**Monotropy Theory**).
- `fit_ext_additive`: we consider the additive effects of gender, mother attachment, and father attachment. This model supports the idea that both mother attachment and father attachment are important, but not their interaction (**Hierarchy Theory** or **Independence Theory**).
- `fit_ext_inter`: we consider the additive effects of gender and the interaction between mother attachment and father attachment. This model supports the idea that the interaction between mother attachment and father attachment is important (**Integration Theory**).

Moreover, in all models, we include children's classroom ID as a random effect to take into account teachers' different ability to evaluate children's problems.

Using R formula syntax, we have

```
# formula for p (same for all models)
p ~ gender + (1|ID_class)

# formula for mu

# fit_ext_zero
mu ~ gender + (1|ID_class)

# fit_ext_mother
mu ~ gender + mother + (1|ID_class)

# fit_ext_additive
mu ~ gender + mother + father + (1|ID_class)

# fit_ext_inter
mu ~ gender + mother * father + (1|ID_class)
```

4.2 AIC and BIC Results

After estimating the models, the AIC and BIC values together with their relative weights are computed. Results are reported in Table~4.1.

Table 4.1: Model comparison externalizing problems ($n_{subj} = 847$).

Model	Df	AIC	AIC _{weights}		BIC	BIC _{weights}	
fit-ext-zero	6	3865.1	0.00	◦	3898.3	0.78	●
fit-ext-mother	9	3853.4	0.92	●	3900.8	0.22	◐
fit-ext-additive	12	3858.4	0.08	◐	3920.0	0.00	◦
fit-ext-inter	21	3867.6	0.00	◦	3971.9	0.00	◦

According to AIC, the most likely model is **fit_ext_mother** (92%) and the second most likely model is **fit_ext_additive** (8%) given the data and the set of models considered. According to BIC, instead, the most likely model is **fit_ext_zero** (78%) and the second most likely model is **fit_ext_mother** (22%) given the data and the set of models considered.

To interpret these results, note that, AIC tends to select more complex models that can better explain the data, on the contrary, BIC penalizes complex models to a greater extent. As pointed out by Kuha, using the two criteria together is

always advocated as agreement provides reassurance on the robustness of the results and disagreement still provides useful information for the discussion. We can say that there is evidence in favour of the role of mother attachment but probably this effect is small.

4.3 Selected Model

Considering the model `fit_ext_mother`, we can run an analysis of deviance to evaluate the significance of the predictors.

```
car::Anova(fit_ext_mother)
## Analysis of Deviance Table (Type II Wald chisquare tests)
##
## Response: externalizing_sum
##           Chisq Df Pr(>Chisq)
## gender 17.732  1  2.544e-05 ***
## mother 17.398  3  0.0005851 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

Results confirm a statistically significant effect of gender and mother attachment. The model summary is reported below.

```
summary(fit_ext_mother)
## Family: nbinom2 ( log )
## Formula:          externalizing_sum ~ gender + mother + (1 | ID_class)
## Zero inflation:    ~gender + (1 | ID_class)
## Data: data_cluster
##
##      AIC      BIC   logLik deviance df.resid
##  3853.4   3900.8 -1916.7   3833.4      837
##
## Random effects:
##
## Conditional model:
## Groups   Name      Variance Std.Dev.
## ID_class (Intercept) 0.07747  0.2783
## Number of obs: 847, groups: ID_class, 50
##
## Zero-inflation model:
## Groups   Name      Variance Std.Dev.
## ID_class (Intercept) 0.766    0.8752
## Number of obs: 847, groups: ID_class, 50
##
## Dispersion parameter for nbinom2 family (): 1.81
##
```

```
## Conditional model:
##               Estimate Std. Error z value Pr(>|z|)
## (Intercept)    0.98649    0.10770   9.160 < 2e-16 ***
## genderM        0.33533    0.07964   4.211 2.54e-05 ***
## motherAnxious  0.24321    0.10338   2.353 0.01864 *
## motherAvoidant 0.30803    0.10910   2.823 0.00475 **
## motherFearful  0.52680    0.13135   4.011 6.05e-05 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Zero-inflation model:
##               Estimate Std. Error z value Pr(>|z|)
## (Intercept)  -1.1211     0.2432  -4.609 4.05e-06 ***
## genderM      -0.6990     0.2491  -2.807  0.005 **
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

To evaluate the effect of gender and mother attachment, the marginal predicted values according to gender and mother attachment are presented separately in Figure~4.1. Not that the marginal predicted values for gender are averaged over mother attachment. Whereas, the marginal predicted values for mother attachment are averaged over gender.

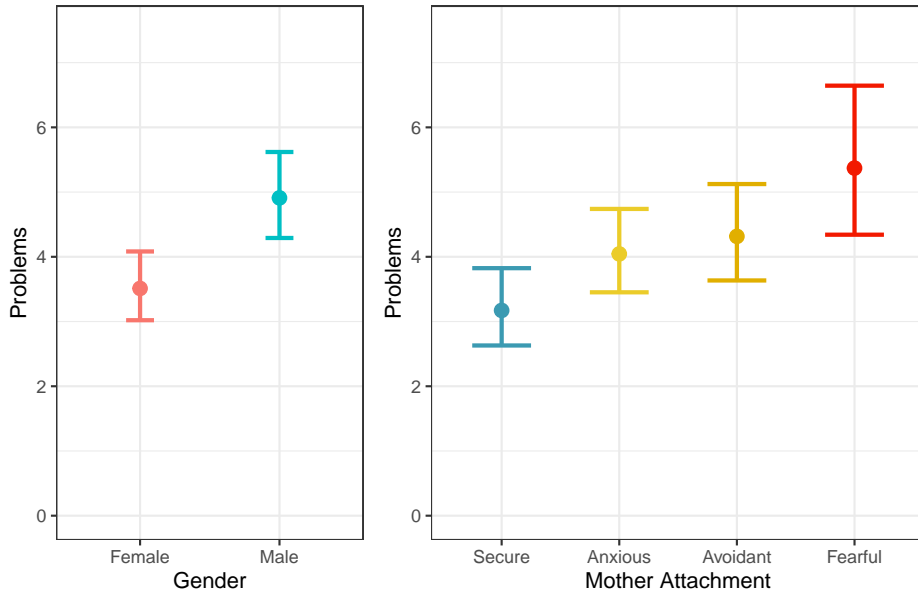


Figure 4.1: Marginal predicted values according to gender and mother attachment ($n_{subj} = 847$).

Post-hoc tests are run to evaluate differences between mother attachment styles, considering pairwise comparisons and adjusting p -values according to multivariate t -distribution. Results are reported below,

```
emmeans::contrast(emmeans::emmeans(fit_ext_mother, specs = ~ mother ),
                  "pairwise", adjust = "mvt")
```

## contrast	estimate	SE	df	t.ratio	p.value
## Secure - Anxious	-0.2432	0.1034	837	-2.353	0.0859
## Secure - Avoidant	-0.3080	0.1091	837	-2.823	0.0247
## Secure - Fearful	-0.5268	0.1313	837	-4.011	0.0003
## Anxious - Avoidant	-0.0648	0.0981	837	-0.661	0.9107
## Anxious - Fearful	-0.2836	0.1216	837	-2.333	0.0898
## Avoidant - Fearful	-0.2188	0.1276	837	-1.714	0.3135

```
##
## Results are averaged over the levels of: gender
## Results are given on the log (not the response) scale.
## P value adjustment: mvt method for 6 tests
```

Overall, results indicate that Males have more externalizing problems than Females. Regarding mother attachment, Fearful and Avoidant children have more problems than Secure children. Moreover, also the difference between Anxious and Secure children and the difference between Anxious and Fearful children have a low (but not statistically significant) p -value.

To evaluate the fit of the model to the data, we computed the *Marginal* R^2 and the *Conditional* R^2 .

```
performance::r2(fit_ext_mother)
```

```
## Warning: mu of 4.2 is too close to zero, estimate of random effect variances may be
##   unreliable.
## # R2 for Mixed Models
##
##   Conditional R2: 0.168
##   Marginal R2: 0.071
```

We can see that the actual variance explained by fixed effects is around 7%, not bad for psychology.

Conclusions

Considering attachment theoretical perspectives, results indicate only the role of mother attachment so we can support the **Monotropy Theory**. Note, however, that the compared models contain no information regarding the expected direction of the effects but we only include/exclude predictors.

Chapter 5

Bayes Factor

To properly evaluate hypotheses with information regarding the expected direction of the effects, we use the Bayes factor with the encompassing prior approach. See the main article for a detailed introduction to this approach [TODO: add link article].

First, we define the encompassing model. Subsequently, we obtain the hypotheses matrices according to the informative hypotheses. Next, we compute the Bayes factor and, finally, we describe the selected model.

5.1 Encompassing Model

We define a Zero-Inflated Negative Binomial (ZINB) mixed-effects model to take into account the characteristics of the dependent variable and its distribution (see Section~2). Again, we consider only the role of gender as a fixed effect and children's classroom ID as a random effect for p . Whereas, regarding μ , we consider the interaction between mother and father attachment together with gender as fixed effects and children's classroom ID as a random effect. In the R formula syntax, we have

```
# formula for p
p ~ gender + (1|ID_class)

# formula for mu
mu ~ gender + mother * father + (1|ID_class)
```

5.1.1 Prior Choice

The prior choice is important for the parameters involved in the equality and inequality constraints. In our case, the parameters of interest (i.e., those related to mother and father attachment interaction) are unbounded. Thus, we can

simply specify as prior a normal distribution with mean 0 and a given standard deviation. Considering the standard deviation, however, we have to choose a value so that the resulting prior is non-informative but without being excessively diffuse.

We can evaluate the consequences of different values' choice considering the resulting prior predictions. To facilitate this step, we compute prior prediction considering only the intercept and a single parameter of interest. Remembering that the inverse link function (i.e., function that in a GLM transform the model linear prediction into the value on the original response scale) is the exponential function, we consider as intercept the value 1 because $\exp(1) \approx 2.7$ that is close to the externalizing problems sample mean 3.35. In Table~5.1, summary information about prior predictions for different standard deviation values is reported.

Table 5.1: Prior prediction according to different prior settings assuming $\exp(1)$ as intercept value.

Prior	Predicted Problems				
	-1 SD	-.5 SD	+0 SD	+.5 SD	+1 SD
$\mathcal{N}(0, 0.5)$	1.6	2.1	2.7	3.5	4.5
$\mathcal{N}(0, 1)$	1.0	1.6	2.7	4.5	7.4
$\mathcal{N}(0, 3)$	0.1	0.6	2.7	12.2	54.6
$\mathcal{N}(0, 5)$	0.0	0.2	2.7	33.1	403.4
$\mathcal{N}(0, 10)$	0.0	0.0	2.7	403.4	59874.1

Considering that externalizing problems are bounded between 0 and 20, a reasonable prior is $\mathcal{N}(0, 3)$. With these settings, prior predicted values cover all possible values without including excessively large values. More diffuse priors would result in values with a higher order of magnitude and tighter priors would exclude plausible values. The influence of prior specification will be subsequently evaluated in a prior sensitivity analysis.

Regarding the other nuisance parameters (i.e., intercepts, random effects and shapes parameters) **brms** default priors are maintained. The resulting prior settings are

```
##           prior      class      coef      group resp dpar
##      normal(0, 3)      b
##      normal(0, 3)      b      fatherAnxious
##      normal(0, 3)      b      fatherAvoidant
##      normal(0, 3)      b      fatherFearful
##      normal(0, 3)      b      genderM
##      normal(0, 3)      b      motherAnxious
##      normal(0, 3)      b      motherAnxious:fatherAnxious
```

```

##          normal(0, 3)          b  motherAnxious:fatherAvoidant
##          normal(0, 3)          b  motherAnxious:fatherFearful
##          normal(0, 3)          b           motherAvoidant
##          normal(0, 3)          b  motherAvoidant:fatherAnxious
##          normal(0, 3)          b  motherAvoidant:fatherAvoidant
##          normal(0, 3)          b  motherAvoidant:fatherFearful
##          normal(0, 3)          b           motherFearful
##          normal(0, 3)          b  motherFearful:fatherAnxious
##          normal(0, 3)          b  motherFearful:fatherAvoidant
##          normal(0, 3)          b  motherFearful:fatherFearful
##          normal(0, 3)          b                                     zi
##          normal(0, 3)          b           genderM                    zi
## student_t(3, 0.7, 2.5) Intercept
##          logistic(0, 1) Intercept                                     zi
##          student_t(3, 0, 2.5) sd
##          student_t(3, 0, 2.5) sd                                     zi
##          student_t(3, 0, 2.5) sd                                     ID_class
##          student_t(3, 0, 2.5) sd           Intercept ID_class
##          student_t(3, 0, 2.5) sd           ID_class                zi
##          student_t(3, 0, 2.5) sd           Intercept ID_class                zi
##          gamma(0.01, 0.01)      shape

```

5.1.2 Posterior

The encompassing model is estimated using 6 independent chains with 10,000 iterations (warm-up 2,000). To do that we use the `brm()` function from the `brms` R-package [Bürkner, b,a], which is based on STAN [Stan Development Team]. Summary of the encompassing model is presented below.

```

## Family: zero_inflated_negbinomial
## Links: mu = log; shape = identity; zi = logit
## Formula: externalizing_sum ~ gender + mother * father + (1 | ID_class)
##          zi ~ gender + (1 | ID_class)
## Data: data (Number of observations: 847)
## Samples: 6 chains, each with iter = 10000; warmup = 2000; thin = 1;
##          total post-warmup samples = 48000
##
## Group-Level Effects:
## ~ID_class (Number of levels: 50)
##          Estimate Est.Error 1-95% CI u-95% CI Rhat Bulk_ESS Tail_ESS
## sd(Intercept)      0.30      0.07   0.17   0.45 1.00   14161   20827
## sd(zi_Intercept)    1.14      0.29   0.68   1.82 1.00   22361   31673
##
## Population-Level Effects:
##          Estimate Est.Error 1-95% CI u-95% CI Rhat Bulk_ESS Tail_ESS
## Intercept      0.98      0.13   0.72   1.24 1.00   35521   38291

```

```

## zi_Intercept          -1.33      0.33    -2.09    -0.78 1.00    36797
## genderM               0.32      0.08     0.16     0.49 1.00    81905
## motherAnxious         0.32      0.19    -0.05     0.70 1.00    39619
## motherAvoidant        0.17      0.26    -0.33     0.69 1.00    37098
## motherFearful         0.69      0.41    -0.07     1.54 1.00    34182
## fatherAnxious         0.22      0.20    -0.16     0.60 1.00    40404
## fatherAvoidant       -0.45      0.22    -0.87    -0.02 1.00    40412
## fatherFearful         0.11      0.42    -0.69     0.96 1.00    35432
## motherAnxious:fatherAnxious -0.39      0.28    -0.94     0.15 1.00    36701
## motherAvoidant:fatherAnxious -0.12      0.34    -0.80     0.53 1.00    36329
## motherFearful:fatherAnxious -0.43      0.51    -1.46     0.55 1.00    35084
## motherAnxious:fatherAvoidant 0.33      0.29    -0.25     0.89 1.00    36157
## motherAvoidant:fatherAvoidant 0.61      0.34    -0.06     1.26 1.00    34169
## motherFearful:fatherAvoidant 0.30      0.48    -0.68     1.20 1.00    32533
## motherAnxious:fatherFearful -0.06      0.48    -1.02     0.86 1.00    34008
## motherAvoidant:fatherFearful 0.09      0.57    -1.04     1.19 1.00    34795
## motherFearful:fatherFearful -0.26      0.60    -1.48     0.86 1.00    30725
## zi_genderM           -0.78      0.29    -1.38    -0.23 1.00    70555
##
## Family Specific Parameters:
##      Estimate Est.Error 1-95% CI u-95% CI Rhat Bulk_ESS Tail_ESS
## shape      1.63      0.24      1.22      2.14 1.00    34643    32767
##
## Samples were drawn using sampling(NUTS). For each parameter, Bulk_ESS
## and Tail_ESS are effective sample size measures, and Rhat is the potential
## scale reduction factor on split chains (at convergence, Rhat = 1).

```

5.2 Hypothesis Matrices

For each informative hypothesis, we obtain a hypothesis matrix that translates equality and inequality constraints according to the encompassing model parametrization. Formalization of informative hypotheses and the procedure to derive hypothesis matrices are described in the main paper ([TODO: add link]).

Here we present the obtained hypothesis matrices where on the columns we have the parameters of the encompassing model (excluding the intercept) and each row expresses an equality constraint or an inequality constraint. Matrices' row names follow this notation: names without square brackets indicate a constraint directly on the model parameter; names within square brackets indicate a group condition (that could be the resulting composition of more parameters). For example, "M_Avoidant:F_Anxious" indicates the actual interaction term of the model, whereas "[M_Avoidant_F_Anxious]" indicates the group condition. This is done because when assuming no interaction or no father attachment effect we set constraints directly on the model parameters, instead, when constraints involve group comparisons, we need to obtain the resulting conditions.

Equality and inequality constraints are presented separately.

Null Hypothesis

- **Equality matrix**, $R_{iE} = 0$ (each row is set to zero).

##	M_Anxious	M_Avoidant	M_Fearful	F_Anxious	F_Avoidant	F_Fearful	M_Anxious:
## M_Anxious	1	0	0	0	0	0	
## M_Avoidant	0	1	0	0	0	0	
## M_Fearful	0	0	1	0	0	0	
## F_Anxious	0	0	0	1	0	0	
## F_Avoidant	0	0	0	0	1	0	
## F_Fearful	0	0	0	0	0	1	
## M_Anxious:F_Anxious	0	0	0	0	0	0	
## M_Avoidant:F_Anxious	0	0	0	0	0	0	
## M_Fearful:F_Anxious	0	0	0	0	0	0	
## M_Anxious:F_Avoidant	0	0	0	0	0	0	
## M_Avoidant:F_Avoidant	0	0	0	0	0	0	
## M_Fearful:F_Avoidant	0	0	0	0	0	0	
## M_Anxious:F_Fearful	0	0	0	0	0	0	
## M_Avoidant:F_Fearful	0	0	0	0	0	0	
## M_Fearful:F_Fearful	0	0	0	0	0	0	

- **Inequality matrix**, $R_{iI} > 0$ (each row is set greater than zero). There are no inequality constraints.

Monotropy Hypothesis

- **Equality matrix**, $R_{iE} = 0$ (each row is set to zero).

##	M_Anxious	M_Avoidant	M_Fearful	F_Anxious	F_Avoidant	F_Fearful	M_Anxious:
## [M_AnxF_Sec - M_AvF_Sec]	1	-1	0	0	0	0	
## F_Anxious	0	0	0	1	0	0	
## F_Avoidant	0	0	0	0	1	0	
## F_Fearful	0	0	0	0	0	1	
## M_Anxious:F_Anxious	0	0	0	0	0	0	
## M_Avoidant:F_Anxious	0	0	0	0	0	0	
## M_Fearful:F_Anxious	0	0	0	0	0	0	
## M_Anxious:F_Avoidant	0	0	0	0	0	0	
## M_Avoidant:F_Avoidant	0	0	0	0	0	0	
## M_Fearful:F_Avoidant	0	0	0	0	0	0	
## M_Anxious:F_Fearful	0	0	0	0	0	0	
## M_Avoidant:F_Fearful	0	0	0	0	0	0	
## M_Fearful:F_Fearful	0	0	0	0	0	0	

- **Inequality matrix**, $R_{iI} > 0$ (each row is set greater than zero).

##	M_Anxious	M_Avoidant	M_Fearful	F_Anxious	F_Avoidant	F_Fearful	M_Anxious:
## [M_AnxF_Sec]	1	0	0	0	0	0	

```
## [M_Fear_F_Sec - M_Av_F_Sec]          0          -1          1          0          0
```

Hierarchy Hypothesis

- **Equality matrix**, $R_{iE} = 0$ (each row is set to zero).

```
##                                     M_Anxious M_Avoidant M_Fearful F_Anxious F_Avoidant F_Fearful
## [M_Anx_F_Sec - M_Av_F_Sec]          1          -1          0          0          0
## [M_Sec_F_An timer - M_Sec_F_Av]      0          0          0          1          -1
## M_Anxious:F_An timer                0          0          0          0          0
## M_Avoidant:F_An timer                0          0          0          0          0
## M_Fearful:F_An timer                 0          0          0          0          0
## M_Anxious:F_Avoidant                0          0          0          0          0
## M_Avoidant:F_Avoidant               0          0          0          0          0
## M_Fearful:F_Avoidant                0          0          0          0          0
## M_Anxious:F_Fearful                 0          0          0          0          0
## M_Avoidant:F_Fearful                0          0          0          0          0
## M_Fearful:F_Fearful                 0          0          0          0          0
```

- **Inequality matrix**, $R_{iI} > 0$ (each row is set greater than zero).

```
##                                     M_Anxious M_Avoidant M_Fearful F_Anxious F_Avoidant F_Fearful
## [M_An timer_F_Sec]                  1          0          0          0          0
## [M_Fear_F_Sec - M_Av_F_Sec]          0          -1          1          0          0
## [M_Sec_F_An timer]                  0          0          0          1          0
## [M_Sec_F_Fear - M_Sec_F_Av]           0          0          0          0          -1
## [M_An timer_F_Sec - M_Sec_F_An timer] 1          0          0          -1          0
## [M_Av_F_Sec - M_Sec_F_Av]            0          1          0          0          -1
## [M_Fear_F_Sec - M_Sec_F_Fear]         0          0          1          0          0
```

Independence Hypothesis

- **Equality matrix**, $R_{iE} = 0$ (each row is set to zero).

```
##                                     M_Anxious M_Avoidant M_Fearful F_Anxious F_Avoidant F_Fearful
## [M_An timer_F_Sec - M_Av_F_Sec]      1          -1          0          0          0
## M_An timer:F_An timer                0          0          0          0          0
## M_Avoidant:F_An timer                0          0          0          0          0
## M_Fearful:F_An timer                 0          0          0          0          0
## M_An timer:F_Avoidant               0          0          0          0          0
## M_Avoidant:F_Avoidant               0          0          0          0          0
## M_Fearful:F_Avoidant                0          0          0          0          0
## M_An timer:F_Fearful                 0          0          0          0          0
## M_Avoidant:F_Fearful                0          0          0          0          0
## M_Fearful:F_Fearful                 0          0          0          0          0
```

- **Inequality matrix**, $R_{iI} > 0$ (each row is set greater than zero).

```
##                                     M_Anxious M_Avoidant M_Fearful F_Anxious F_Avoidant F_Fearful
```

## [M_AnxF_Sec]	1	0	0	0	0	0
## [M_Fear_F_Sec - M_Av_F_Sec]	0	-1	1	0	0	0
## [M_Sec_F_AnxF]	0	0	0	1	0	0
## [M_Sec_F_Av - M_Sec_F_AnxF]	0	0	0	-1	1	0
## [M_Sec_F_Fear - M_Sec_F_Av]	0	0	0	0	-1	1

Integration Hypothesis

- **Equality matrix**, $R_{iE} = 0$ (each row is set to zero).

##	M_Anxious	M_Avoidant	M_Fearful	F_Anxious	F_Avoidant	F_Fearful	M_
## [M_AnxF_Sec - M_Av_F_Sec]	1	-1	0	0	0	0	
## [M_AnxF_Sec - M_Sec_F_AnxF]	1	0	0	-1	0	0	
## [M_AnxF_Sec - M_Sec_F_Av]	1	0	0	0	-1	0	
## [M_AnxF_AnxF - M_AnxF_Av]	0	0	0	1	-1	0	
## [M_AnxF_AnxF - M_Av_F_AnxF]	1	-1	0	0	0	0	
## [M_AnxF_AnxF - M_Av_F_Av]	1	-1	0	1	-1	0	
## [M_Fear_F_AnxF - M_Fear_F_Av]	0	0	0	1	-1	0	
## [M_Fear_F_AnxF - M_AnxF_Fear]	-1	0	1	1	0	-1	
## [M_Fear_F_AnxF - M_Av_F_Fear]	0	-1	1	1	0	-1	

- **Inequality matrix**, $R_{iI} > 0$ (each row is set greater than zero).

##	M_Anxious	M_Avoidant	M_Fearful	F_Anxious	F_Avoidant	F_Fearful	M_
## [M_AnxF_Sec]	1	0	0	0	0	0	
## [M_AnxF_AnxF - M_AnxF_Sec]	0	0	0	1	0	0	
## [M_Fear_F_AnxF - M_AnxF_AnxF]	-1	0	1	0	0	0	
## [M_Fear_F_Fear - M_Fear_F_AnxF]	0	0	0	-1	0	1	

5.3 Centering and Adjusting

So far we have specified the encompassing prior, obtained the model posterior distribution, and defined the hypotheses matrices. Now, we need to transform our parameters of interest and center the distribution on the constraints focal points of interest. We apply the following transformation

$$\beta = R\theta - r$$

but we can ignore r as in all our constraints it is always a vector of zeros.

Next, we get the adjusted prior and the posterior of the transformed parameters vector β (i.e., the parameters that identify the constraints) for each hypothesis. The adjusted prior is given by

$$\pi_{adj}(\beta) \sim \mathcal{N}(0, \Sigma_\beta) = \mathcal{N}(0, R\Sigma_\theta R^T).$$

Note that we set the mean vector to zero. The posterior is given by the same transformation

$$Pr(\beta|Y) \sim \mathcal{N}(\hat{\beta}, \hat{\Sigma}_\beta) = \mathcal{N}(R\hat{\theta} - r, R\hat{\Sigma}_\theta R^T).$$

See the main article for more details [TODO: add link].

This adjustment, however, requires the hypothesis matrix R to be *full-row-rank* (i.e., all constraints are linearly independent). However, this is not the case with the Hierarchy Hypothesis. To overcome this issue, we follow the solution presented in the main article. First, define R^* selecting the maximum number of independent rows. In this case, 15 contrast are independent

##	M_Anxious	M_Avoidant	M_Fearful	F_Anxious	F_Avoidant	F_Fearful
## [M_Anx_F_Sec - M_Av_F_Sec]	1	-1	0	0	0	0
## [M_Sec_F_An timer - M_Sec_F_Av]	0	0	0	1	-1	0
## M_Anxious:F_An timer	0	0	0	0	0	0
## M_Avoidant:F_An timer	0	0	0	0	0	0
## M_Fearful:F_An timer	0	0	0	0	0	0
## M_Anxious:F_Avoidant	0	0	0	0	0	0
## M_Avoidant:F_Avoidant	0	0	0	0	0	0
## M_Fearful:F_Avoidant	0	0	0	0	0	0
## M_Anxious:F_Fearful	0	0	0	0	0	0
## M_Avoidant:F_Fearful	0	0	0	0	0	0
## M_Fearful:F_Fearful	0	0	0	0	0	0
## [M_An timer_F_Sec]	1	0	0	0	0	0
## [M_Fear_F_Sec - M_Av_F_Sec]	0	-1	1	0	0	0
## [M_Sec_F_An timer]	0	0	0	1	0	0
## [M_Sec_F_Fear - M_Sec_F_Av]	0	0	0	0	0	-1

The remaining contrasts, instead, are obtained as linear combinations of the other constraints. In particular,

```
# Constraint 16: [M_An timer_F_Sec - M_Sec_F_An timer]
all(R["[M_An timer_F_Sec - M_Sec_F_An timer]",] == R["[M_An timer_F_Sec]", ] - R["[M_Sec_F_An timer]", ])
## [1] TRUE

# Constraint 17: [M_Av_F_Sec - M_Sec_F_Av]
all(R["[M_Av_F_Sec - M_Sec_F_Av]",] == - R["[M_An timer_F_Sec - M_Av_F_Sec]", ] + R["[M_Sec_F_Av]", ])
## [1] TRUE

# Constraint 18: [M_Fear_F_Sec - M_Sec_F_Fear]
all(R["[M_Fear_F_Sec - M_Sec_F_Fear]",] == R["[M_Av_F_Sec - M_Sec_F_Av]", ] + R["[M_Fear_F_Sec - M_Sec_F_Fear]", ])
## [1] TRUE
```

Before computing the Bayes factor, note that we have a set of comparable hypotheses as it exists a common solution to the set of linear equations obtained by setting all hypothesis constraints equal to zero. The solution is the trivial solution of simply considering all parameters equal to zero. Finally, we do not

need to standardize our parameters as they represent mean groups' differences. See the main article for a detailed explanation [TODO: add link article].

5.4 Results and Sensitivity

To compute the Bayes factor we evaluate marginal densities and conditional probabilities as described in detail in the main article [TODO: add link article]. Bayes factor and posterior probability of each hypothesis are reported in Table~5.2.

Table 5.2: Bayes factor encompassing model and hypothesis posterior probabilities ($n_{subj} = 847$).

Hypothesis	Bayes Factor	Posterior Probability	
Null	2.9e+11	0.01	○
Monotropy	2.6e+13	0.98	●
Hierarchy	2.7e+11	0.01	○
Independence	3.9e+09	0.00	○
Integration	3.2e+09	0.00	○

Remember, however, that prior specification affects the Bayes factor results. Therefore, we also evaluate the results considering different prior settings. In particular, we consider as possible priors for the parameters of interest:

- $\mathcal{N}(0, .5)$ - unreasonable tight prior
- $\mathcal{N}(0, 1)$ - tighter prior
- $\mathcal{N}(0, 3)$ - original prior
- $\mathcal{N}(0, 5)$ - more diffuse prior
- $\mathcal{N}(0, 10)$ - unreasonably diffuse prior

The results of the prior sensitivity analysis are reported in Table~5.3.

Overall results consistently indicate the Monotropy Hypothesis as the most supported by the data. However, we can observe two distinct patterns. As the prior gets more diffuse, the order of magnitude of the Bayes factor comparing each hypothesis with the encompassing model increases. Moreover, the probability of the Null Hypothesis increases with more diffuse prior, whereas the probabilities of the Hierarchy, Independence and Integration Hypothesis increases with tighter priors.

Table 5.3: Bayes factor encompassing model and hypothesis posterior probabilities (PP) under different prior settings ($n_{subj} = 847$).

Hypothesis	$\mathcal{N}(0, .5)$		$\mathcal{N}(0, 1)$		$\mathcal{N}(0, 3)$		$\mathcal{N}(0, 5)$		$\mathcal{N}(0, 10)$	
	BF	PP	BF	PP	BF	PP	BF	PP	BF	PP
Null	8.2e+01	0.00	9.4e+04	0.00	2.9e+11	0.01	4.7e+14	0.03	1.5e+19	0.11
Monotropy	1.2e+05	0.67	6.3e+07	0.90	2.6e+13	0.98	1.6e+16	0.97	1.2e+20	0.89
Hierarchy	4.9e+04	0.28	6.1e+06	0.09	2.7e+11	0.01	6.7e+13	0.00	1.2e+17	0.00
Independence	4.4e+03	0.02	2.6e+05	0.00	3.9e+09	0.00	5.8e+11	0.00	5.1e+14	0.00
Integration	4.6e+03	0.03	3.3e+05	0.00	3.2e+09	0.00	3.3e+11	0.00	1.5e+14	0.00

To interpret these patterns, remember that order constraints are insensitive to the distribution specification as long as the distribution is symmetric and centred on the constraint focal point. On the contrary, equality constraints are highly affected by the prior definition (see the main article for more details [TODO: add link article]).

All the defined hypotheses include equality constraints. Thus, for more diffuse prior we observe that the order of magnitude of the Bayes factor comparing each hypothesis with the encompassing model increases. Moreover, the hypothesis with a higher number of equality constraints (e.g., Null Hypothesis) will be favoured over hypotheses with a smaller number of equality constraints (e.g., Hierarchy, Independence and Integration Hypothesis).

5.5 Selected Model

One of the limits of the Bayes factor with the encompassing prior approach is that we only get the selected hypothesis but we do not obtain the actual estimates of the parameters posterior. To overcome this limit we rely on Bayesian inference that allows us to effectively estimate the model parameter posteriors.

This time in the model we consider only the role of gender and mother attachment as fixed effects of μ . In the R formula syntax, we have

```
# formula for p
p ~ gender + (1|ID_class)

# formula for mu
mu ~ gender + mother + (1|ID_class)
```

Again, we specify a normal distribution with mean 0 and standard deviation of 3, $\mathcal{N}(0, 3)$, as prior for the beta parameters (i.e., those related to gender and mother attachment). Whereas, for the other nuisance parameters (i.e., intercepts, random effects and shapes parameters) **brms** default priors are maintained. The resulting prior settings are

```
##                prior      class      coef      group resp dpar nlpar bound
```

```

##          normal(0, 3)          b          user
##          normal(0, 3)          b          genderM          (vectorized)
##          normal(0, 3)          b motherAnxious          (vectorized)
##          normal(0, 3)          b motherAvoidant          (vectorized)
##          normal(0, 3)          b motherFearful          (vectorized)
##          normal(0, 3)          b          zi          (vectorized)
##          normal(0, 3)          b          genderM          zi          (vectorized)
## student_t(3, 0.7, 2.5) Intercept          default
##          logistic(0, 1) Intercept          zi          default
##          student_t(3, 0, 2.5)          sd          default
##          student_t(3, 0, 2.5)          sd          zi          default
##          student_t(3, 0, 2.5)          sd          ID_class          (vectorized)
##          student_t(3, 0, 2.5)          sd          Intercept ID_class          (vectorized)
##          student_t(3, 0, 2.5)          sd          ID_class          zi          (vectorized)
##          student_t(3, 0, 2.5)          sd          Intercept ID_class          zi          (vectorized)
##          gamma(0.01, 0.01)          shape          default

```

The model is estimated using 6 independent chains with 6,000 iterations (warm-up 2,000). The model summary is presented below.

```

## Family: zero_inflated_negbinomial
## Links: mu = log; shape = identity; zi = logit
## Formula: externalizing_sum ~ gender + mother + (1 | ID_class)
##          zi ~ gender + (1 | ID_class)
## Data: data (Number of observations: 847)
## Samples: 6 chains, each with iter = 6000; warmup = 2000; thin = 1;
##          total post-warmup samples = 24000
##
## Group-Level Effects:
## ~ID_class (Number of levels: 50)
##          Estimate Est.Error 1-95% CI u-95% CI Rhat Bulk_ESS Tail_ESS
## sd(Intercept)      0.29      0.07    0.17    0.43 1.00    7372    10149
## sd(zi_Intercept)   1.10      0.27    0.65    1.74 1.00    9371    12428
##
## Population-Level Effects:
##          Estimate Est.Error 1-95% CI u-95% CI Rhat Bulk_ESS Tail_ESS
## Intercept      0.96      0.11    0.74    1.18 1.00    14818    16592
## zi_Intercept   -1.27      0.31   -1.97   -0.75 1.00    14384    12961
## genderM        0.34      0.08    0.18    0.50 1.00    30197    18984
## motherAnxious   0.25      0.10    0.05    0.46 1.00    21564    18963
## motherAvoidant   0.32      0.11    0.10    0.53 1.00    21623    19168
## motherFearful    0.54      0.14    0.28    0.81 1.00    22199    18755
## zi_genderM     -0.74      0.27   -1.30   -0.22 1.00    27915    17381
##
## Family Specific Parameters:
##          Estimate Est.Error 1-95% CI u-95% CI Rhat Bulk_ESS Tail_ESS
## shape      1.70      0.25    1.27    2.23 1.00    15644    16568

```

```
##
## Samples were drawn using sampling(NUTS). For each parameter, Bulk_ESS
## and Tail_ESS are effective sample size measures, and Rhat is the potential
## scale reduction factor on split chains (at convergence, Rhat = 1).
```

Marginal effects are presented in Figure~5.1 and differences between mother attachment patterns are reported in Figure~5.2.

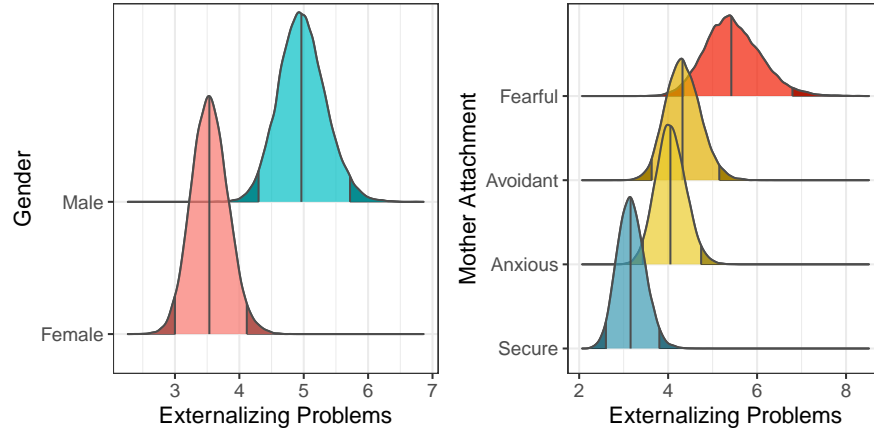


Figure 5.1: Marginal predicted values according to gender and mother attachment ($n_{subj} = 847$).

Overall, results indicate that Males have more externalizing problems than Females. Regarding mother attachment, Fearful, Avoidant, and Anxious children have more problems than Secure children. Moreover, Fearful children have more problems than Anxious children.

To evaluate the fit of the model to the data, we computed the *Bayesian* R^2 , using the function `brms::bayes_R2()`, and we present Posterior Predictions in Figure~5.3.

```
r2_ext
##      Estimate Est.Error      Q2.5      Q97.5
## R2 0.1482686 0.02972484 0.09435914 0.2106949
```

We can see that the actual variance explained by fixed effects and random effects is around 15%. Moreover, the posterior predictive check indicates a good fit to the data.

Conclusions

Considering attachment theoretical perspectives, results indicate only the role of mother attachment so we can support the Monotropy Theory.

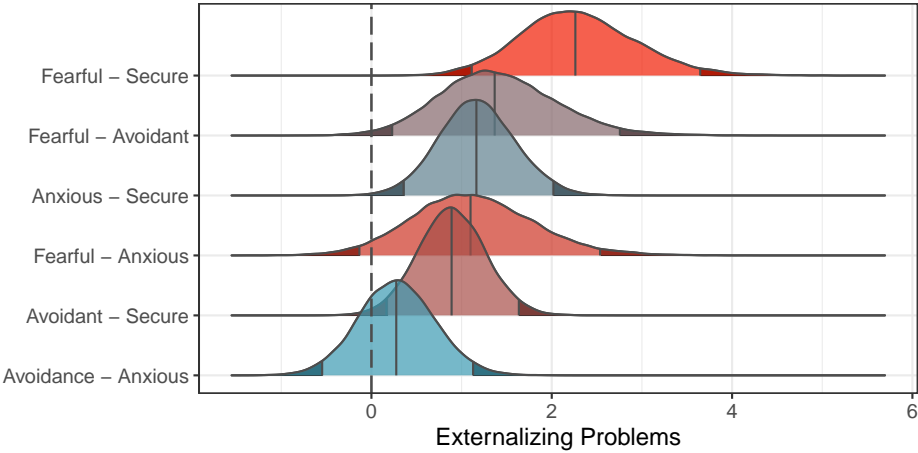


Figure 5.2: Predicted differences between mother attachment patterns ($n_{subj} = 847$).

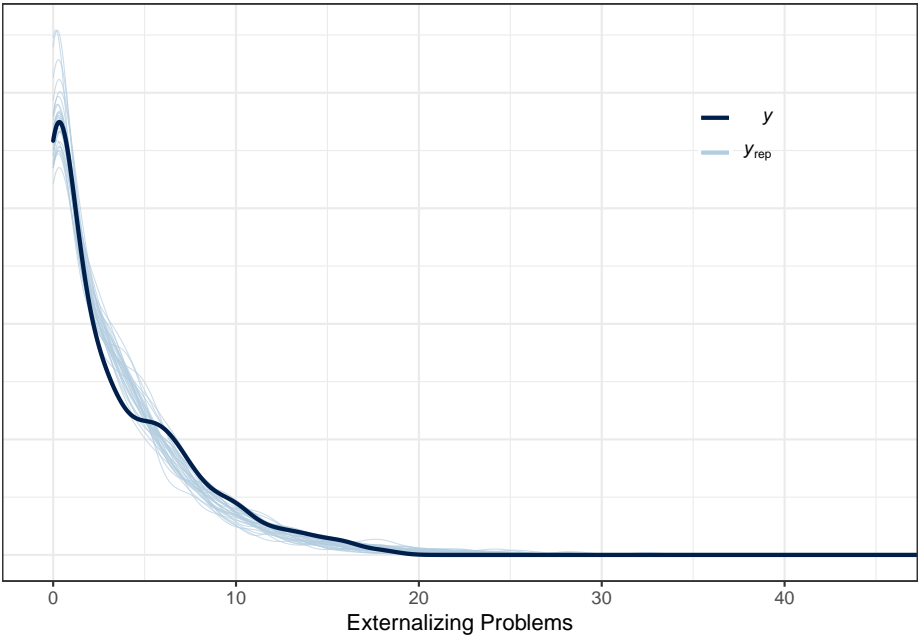


Figure 5.3: Posterior predictive check ($n_{subj} = 847$).

Chapter 6

Conclusions

Overall, we obtained consistent results from the three different approaches. Males have more problems than Females and, regarding attachment, only mother attachment influences children's externalizing problems. In particular, we observed the following pattern: Secure children have the lowest level of problems, Anxious and Avoidant children have a similar, intermediate, level of problems, and Fearful children have the highest level of problems. Taken together, these results support the Monotropy Theory.

Although if the different approaches lead to apparently similar results, the rigorous interpretation of the results is very different.

- Considering the NHST approach, we actually only found that it is unlikely that mother attachment has no effect. Thus, we reject the null hypothesis but we can not quantify the evidence in favour of any of our hypotheses.
- Only model comparison allows us to quantify the relative evidence of our models. Model comparison results clearly indicated evidence in favour of an effect of mother attachment but not father attachment. However, using information criteria we could not directly evaluate our informative hypotheses regarding the expected effects, but only the presence of any effect.
- Bayes factor with encompassing prior approach allowed us to directly test our informative hypotheses regarding the expected effects. Results clearly selected the Monotropy theory as the most likely theory among those considered.

To summarize, the apparently identical results actually have a completely different meaning and what we learn from the data is very different. Hopefully, now it is clear that statistical inference is a complex process that requires careful thinking. In particular, to answer the questions we are actually interested in, we need to apply the appropriate statistical techniques.

Internalizing Problems

Chapter 7

Models Family Choice

In this chapter, we discuss the appropriate models' family to take into account data characteristics.

Internalizing problems are computed as the sum of 10 items of the SDQ, obtaining discrete scores that range from 0 to 20. Considering data distribution (see Figure~1.9), we choose a **Negative Binomial** distribution to model the data.

7.1 Zero Inflated Negative Binomial

As in the case of externalizing problems, we evaluate whether a *Zero-Inflated* model may be appropriate. We compare the number of observed zeros and expected zeros in a Negative Binomial mixed-effects model considering the same predictors as in the case of externalizing problems. Using R formula syntax, we have

```
# model formula
internalizing_sum ~ gender + mother * father + (1|ID_class)
```

Comparing the number of observed zero and expected zeros, we get

```
my_check_zeroinflation(fit_int_nb)
## # Check for zero-inflation
##
##      Observed zeros: 226
##      Predicted zeros: 195
##      Ratio: 0.86
## Model is underfitting zeros (probable zero-inflation).
```

Results indicate that the model is slightly under-fitting the number of zeros. Thus, we can fit a *Zero Inflated Negative Binomial* (ZINB) model and compare the performance of the two models. Using R formula syntax, we have

```
# formula for p
p ~ gender + (1|ID_class)

# formula for mu
mu ~ gender + mother * father + (1|ID_class)
```

Below we report results of the analysis of deviance.

```
anova(fit_int_nb, fit_int_zinb)
## Data: data_cluster
## Models:
## fit_int_nb: internalizing_sum ~ gender + mother * father + (1 | ID_class), zi=~0, d
## fit_int_zinb: internalizing_sum ~ gender + mother * father + (1 | ID_class), zi=~ge
##           Df      AIC      BIC logLik deviance Chisq Chi Df Pr(>Chisq)
## fit_int_nb  19 3680.7 3770.8 -1821.3   3642.7
## fit_int_zinb 22 3669.0 3773.3 -1812.5   3625.0 17.677      3 0.0005127 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

Overall, results indicate that the ZINB model performs better than the Negative Binomial model. Note, however, that BIC actually prefers the model without zero inflation. Nevertheless, in the following analyses, we decide to use ZINB models to be consistent with the analysis of the externalizing problems.

Chapter 8

NHST

Following the traditional NHST approach, we consider the model previously defined that includes all effects of interest (gender effect and the interaction between mother attachment and father attachment). Results of the analysis of deviance are reported below.

```
car::Anova(fit_int_zinb)
## Analysis of Deviance Table (Type II Wald chisquare tests)
##
## Response: internalizing_sum
##           Chisq Df Pr(>Chisq)
## gender      0.9449  1  0.331015
## mother     12.4302  3  0.006046 **
## father      2.6855  3  0.442698
## mother:father 9.6213  9  0.382006
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

Results indicate only a statistically significant effect of mother attachment. On the contrary, the interaction and father attachment are not significant. The model summary is reported below.

```
summary(fit_int_zinb)
## Family: nbinom2 ( log )
## Formula:      internalizing_sum ~ gender + mother * father + (1 | ID_class)
## Zero inflation: ~gender + (1 | ID_class)
## Data: data_cluster
##
##      AIC      BIC   logLik deviance df.resid
##  3669.0   3773.3 -1812.5   3625.0      825
##
```

```
## Random effects:
##
## Conditional model:
##   Groups   Name          Variance Std.Dev.
##   ID_class (Intercept) 0.1364    0.3693
## Number of obs: 847, groups: ID_class, 50
##
## Zero-inflation model:
##   Groups   Name          Variance Std.Dev.
##   ID_class (Intercept) 3.225     1.796
## Number of obs: 847, groups: ID_class, 50
##
## Dispersion parameter for nbinom2 family (): 2.6
##
## Conditional model:
##                                     Estimate Std. Error z value Pr(>|z|)
## (Intercept)                      0.85592    0.12115   7.065 1.61e-12 ***
## genderM                          0.06854    0.07051   0.972 0.33101
## motherAnxious                    0.47960    0.16672   2.877 0.00402 **
## motherAvoidant                   -0.08057    0.24671  -0.327 0.74398
## motherFearful                     0.66801    0.41101   1.625 0.10410
## fatherAnxious                     0.07410    0.16668   0.445 0.65661
## fatherAvoidant                    0.09901    0.17349   0.571 0.56822
## fatherFearful                     0.38182    0.33578   1.137 0.25550
## motherAnxious:fatherAnxious      -0.26373    0.23441  -1.125 0.26055
## motherAvoidant:fatherAnxious      0.25023    0.30643   0.817 0.41416
## motherFearful:fatherAnxious      -0.80565    0.50634  -1.591 0.11158
## motherAnxious:fatherAvoidant     -0.25807    0.23926  -1.079 0.28076
## motherAvoidant:fatherAvoidant     0.10204    0.30109   0.339 0.73467
## motherFearful:fatherAvoidant     -0.25422    0.46295  -0.549 0.58292
## motherAnxious:fatherFearful      -0.53674    0.39927  -1.344 0.17885
## motherAvoidant:fatherFearful      0.23131    0.47623   0.486 0.62718
## motherFearful:fatherFearful      -0.41834    0.54125  -0.773 0.43958
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Zero-inflation model:
##                                     Estimate Std. Error z value Pr(>|z|)
## (Intercept)                    -2.6112     0.6183  -4.223 2.41e-05 ***
## genderM                        -0.2890     0.3531  -0.818 0.413
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

To evaluate the effect of mother attachment, the marginal predicted values are presented in Figure~8.1. Note that the marginal predicted values are averaged

over father attachment and gender effect.

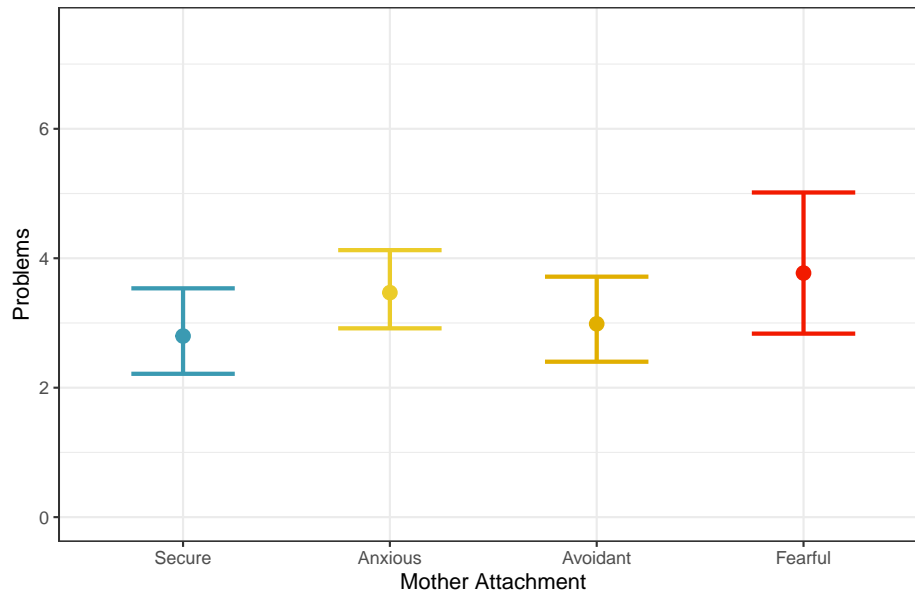


Figure 8.1: Marginal predicted values according to mother attachment. Values are averaged over the other effects ($n_{subj} = 847$).

Post-hoc tests are run to evaluate differences between mother attachment styles, considering pairwise comparisons and adjusting p -values according to multivariate t -distribution. Results are reported below,

```
emmeans::contrast(emmeans::emmeans(fit_int_zinb, specs = ~ mother ),
  "pairwise", adjust = "mvt")
## NOTE: Results may be misleading due to involvement in interactions
## contrast      estimate    SE  df t.ratio p.value
## Secure - Anxious   -0.2150 0.117 825  -1.836  0.2508
## Secure - Avoidant  -0.0653 0.136 825  -0.481  0.9622
## Secure - Fearful   -0.2985 0.166 825  -1.793  0.2707
## Anxious - Avoidant  0.1496 0.112 825   1.336  0.5321
## Anxious - Fearful  -0.0835 0.147 825  -0.566  0.9403
## Avoidant - Fearful -0.2331 0.164 825  -1.423  0.4774
##
## Results are averaged over the levels of: gender, father
## Results are given on the log (not the response) scale.
## P value adjustment: mvt method for 6 tests
```

After adjusting p -values we actually get that there are no statistically significant differences. Without adjusting, the difference between Anxious and Secure children and the difference between Fearful and Secure children have a low (but

not statistically significant) p -value.

```
emmeans::contrast(emmeans::emmeans(fit_int_zinb, specs = ~ mother ),
  "pairwise", adjust = NULL)
## NOTE: Results may be misleading due to involvement in interactions
## contrast      estimate      SE df t.ratio p.value
## Secure - Anxious   -0.2150 0.117 825  -1.836  0.0668
## Secure - Avoidant  -0.0653 0.136 825  -0.481  0.6307
## Secure - Fearful   -0.2985 0.166 825  -1.793  0.0734
## Anxious - Avoidant   0.1496 0.112 825   1.336  0.1818
## Anxious - Fearful   -0.0835 0.147 825  -0.566  0.5713
## Avoidant - Fearful  -0.2331 0.164 825  -1.423  0.1551
##
## Results are averaged over the levels of: gender, father
## Results are given on the log (not the response) scale.
```

To evaluate the fit of the model to the data, we computed the *Marginal R^2* and the *Conditional R^2* .

```
performance::r2(fit_int_zinb)
## Warning: mu of 3.2 is too close to zero, estimate of random effect variances may be
## # R2 for Mixed Models
##
## Conditional R2: 0.217
## Marginal R2: 0.047
```

We can see that the actual variance explained by fixed effects is less than 5%, not that much.

Conclusions

Results are difficult to interpret because we have a statistically significant effect of mother attachment, but, considering pot-hoc tests, we get no statistically significant difference. Overall, we could say that, in some way, results indicate a role of mother attachment but this probably is small.

Chapter 9

Model Comparison

9.1 Formalize Models

We consider Zero Inflated Negative Binomial Mixed-Effects models with only the role of gender as a fixed effect and children's classroom ID as a random effect for p . Whereas, considering μ , we define the same models as in the analysis of externalizing problems. Using R formula syntax, we have

```
# formula for p (same for all models)
p ~ gender + (1|ID_class)

# formula for mu

# fit_int_zero
mu ~ gender + (1|ID_class)

# fit_int_mother
mu ~ gender + mother + (1|ID_class)

# fit_int_additive
mu ~ gender + mother + father + (1|ID_class)

# fit_int_inter
mu ~ gender + mother * father + (1|ID_class)
```

9.2 AIC and BIC Results

AIC and BIC values together with their relative weights are computed and reported in Table-9.1.

Table 9.1: Model comparison internalizing problems ($n_{subj} = 847$).

Model	Df	AIC	AIC _{weights}		BIC	BIC _{weights}	
fit-int-zero	6	3671.4	0.00	◦	3704.6	0.52	●
fit-int-mother	9	3657.4	0.83	●	3704.8	0.48	●
fit-int-additive	12	3660.6	0.16	●	3722.2	0.00	◦
fit-int-inter	21	3669.0	0.00	◦	3773.3	0.00	◦

According to AIC, the most likely model is `fit_int_mother` (83%) and the second most likely model is `fit_int_additive` (16%) given the data and the set of models considered. According to BIC, instead, `fit_int_zero` and `fit_int_mother` models have almost the same probability, 52% and 48% respectively.

We can say that there is evidence in favour of the role of mother attachment but probably this effect is small.

9.3 Selected Model

Results of analysis of deviance for model `fit_int_mother` are reported below.

```
car::Anova(fit_int_mother)
## Analysis of Deviance Table (Type II Wald chisquare tests)
##
## Response: internalizing_sum
##           Chisq Df Pr(>Chisq)
## gender   0.754  1  0.3852261
## mother  19.894  3  0.0001786 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

Results confirm a statistically significant effect of mother attachment. The model summary is reported below.

```
summary(fit_int_mother)
## Family: nbinom2 ( log )
## Formula:      internalizing_sum ~ gender + mother + (1 | ID_class)
## Zero inflation: ~gender + (1 | ID_class)
## Data: data_cluster
##
##           AIC           BIC    logLik deviance df.resid
```

```
##    3657.4    3704.8   -1818.7    3637.4        837
##
## Random effects:
##
## Conditional model:
##   Groups   Name          Variance Std.Dev.
## ID_class (Intercept) 0.1423    0.3772
## Number of obs: 847, groups: ID_class, 50
##
## Zero-inflation model:
##   Groups   Name          Variance Std.Dev.
## ID_class (Intercept) 2.966     1.722
## Number of obs: 847, groups: ID_class, 50
##
## Dispersion parameter for nbinom2 family (): 2.56
##
## Conditional model:
##               Estimate Std. Error z value Pr(>|z|)
## (Intercept)    0.91263    0.10407   8.770 < 2e-16 ***
## genderM         0.06096    0.07021   0.868  0.38523
## motherAnxious   0.28519    0.08921   3.197  0.00139 **
## motherAvoidant  0.11979    0.09462   1.266  0.20550
## motherFearful   0.46860    0.11682   4.011 6.04e-05 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Zero-inflation model:
##               Estimate Std. Error z value Pr(>|z|)
## (Intercept)  -2.5021     0.5425  -4.612 3.99e-06 ***
## genderM       -0.2774     0.3493  -0.794  0.427
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

To evaluate the effect of mother attachment, the marginal predicted values are presented in Figure~9.1. Note that the marginal predicted values for gender are averaged over mother attachment. Whereas, the marginal predicted values for mother attachment are averaged over gender.

Post-hoc tests are run, considering pairwise comparisons and adjusting p -values according to multivariate t -distribution. Results are reported below,

```
emmeans::contrast(emmeans::emmeans(fit_int_mother, specs = ~ mother ),
                   "pairwise", adjust = "mvt")
##   contrast      estimate      SE df t.ratio p.value
## Secure - Anxious   -0.285 0.0892 837  -3.197  0.0075
## Secure - Avoidant  -0.120 0.0946 837  -1.266  0.5809
```

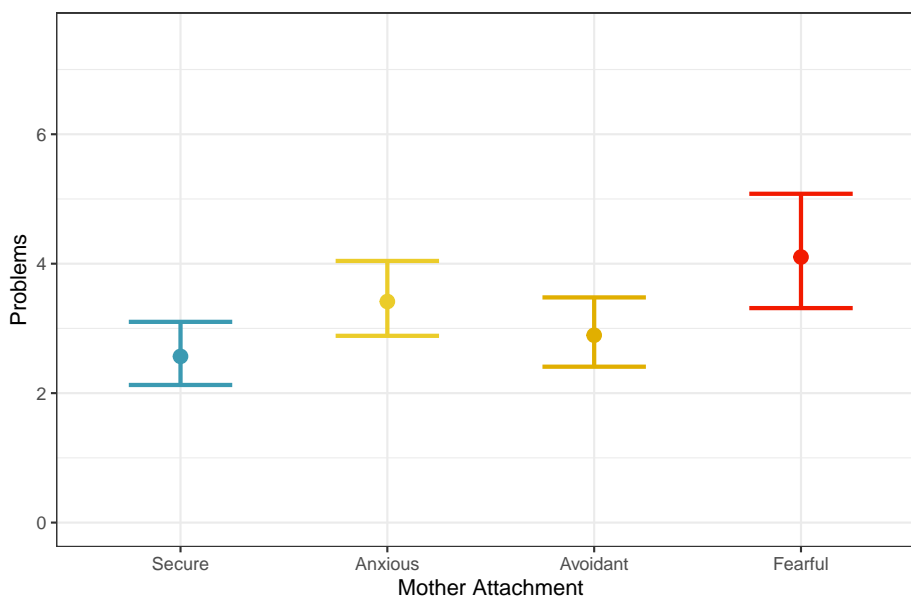


Figure 9.1: Marginal predicted values according to gender and mother attachment ($n_{subj} = 847$).

```
## Secure - Fearful      -0.469 0.1168 837  -4.011  0.0003
## Anxious - Avoidant    0.165 0.0869 837   1.903  0.2243
## Anxious - Fearful    -0.183 0.1096 837  -1.673  0.3347
## Avoidant - Fearful    -0.349 0.1167 837  -2.989  0.0148
##
## Results are averaged over the levels of: gender
## Results are given on the log (not the response) scale.
## P value adjustment: mvt method for 6 tests
```

Results indicate that Fearful and Anxious children have more problems than Secure children. Moreover, also the difference between Avoidant and Fearful children is significant.

To evaluate the fit of the model to the data, we computed the *Marginal R^2* and the *Conditional R^2* .

```
performance::r2(fit_int_mother)
## Warning: mu of 3.2 is too close to zero, estimate of random effect variances may be
## # R2 for Mixed Models
##
## Conditional R2: 0.209
## Marginal R2: 0.031
```


We can see that the actual variance explained by fixed effects is 3%, not that much.

Conclusions

Considering attachment theoretical perspectives, results indicate only the role of mother attachment so we can support the **Monotropy Theory**. Note, however, that the compared models contain no information regarding the expected direction of the effects but we only include/exclude predictors.

Chapter 10

Bayes Factor

10.1 Encompassing Model

We define a Zero-Inflated Negative Binomial (ZINB) mixed-effects considering only the role of gender as a fixed effect and children's classroom ID as a random effect for p . Whereas, regarding μ , we consider the interaction between mother and father attachment together with gender as fixed effects and children's classroom ID as a random effect. In the R formula syntax, we have

```
# formula for p
p ~ gender + (1|ID_class)

# formula for mu
mu ~ gender + mother * father + (1|ID_class)
```

10.1.1 Prior Choice

We set the same prior as for the externalizing problems analysis. The resulting prior settings are

##	prior	class	coef	group	resp	dpar	nlpar	bound
##	normal(0, 3)	b						
##	normal(0, 3)	b	fatherAnxious					
##	normal(0, 3)	b	fatherAvoidant					
##	normal(0, 3)	b	fatherFearful					
##	normal(0, 3)	b	genderM					
##	normal(0, 3)	b	motherAnxious					
##	normal(0, 3)	b	motherAnxious:fatherAnxious					
##	normal(0, 3)	b	motherAnxious:fatherAvoidant					
##	normal(0, 3)	b	motherAnxious:fatherFearful					
##	normal(0, 3)	b	motherAvoidant					

```

##          normal(0, 3)          b  motherAvoidant:fatherAnxious
##          normal(0, 3)          b motherAvoidant:fatherAvoidant
##          normal(0, 3)          b  motherAvoidant:fatherFearful
##          normal(0, 3)          b          motherFearful
##          normal(0, 3)          b  motherFearful:fatherAnxious
##          normal(0, 3)          b motherFearful:fatherAvoidant
##          normal(0, 3)          b  motherFearful:fatherFearful
##          normal(0, 3)          b
##          normal(0, 3)          b          genderM          zi
## student_t(3, 0.7, 2.5) Intercept
##          logistic(0, 1) Intercept          zi
##          student_t(3, 0, 2.5)      sd
##          student_t(3, 0, 2.5)      sd          zi
##          student_t(3, 0, 2.5)      sd          ID_class
##          student_t(3, 0, 2.5)      sd          Intercept ID_class
##          student_t(3, 0, 2.5)      sd          ID_class          zi
##          student_t(3, 0, 2.5)      sd          Intercept ID_class          zi
##          gamma(0.01, 0.01)      shape

```

10.1.2 Posterior

The encompassing model is estimated using 6 independent chains with 10,000 iterations (warm-up 2,000). Summary of the encompassing model is presented below.

```

## Family: zero_inflated_negbinomial
## Links: mu = log; shape = identity; zi = logit
## Formula: internalizing_sum ~ gender + mother * father + (1 | ID_class)
##          zi ~ gender + (1 | ID_class)
## Data: data (Number of observations: 847)
## Samples: 6 chains, each with iter = 10000; warmup = 2000; thin = 1;
##          total post-warmup samples = 48000
##
## Group-Level Effects:
## ~ID_class (Number of levels: 50)
##          Estimate Est.Error 1-95% CI u-95% CI Rhat Bulk_ESS Tail_ESS
## sd(Intercept)      0.44      0.08    0.29    0.62 1.00    9467    20478
## sd(zi_Intercept)   2.24      0.56    1.37    3.52 1.00   22914    28211
##
## Population-Level Effects:
##          Estimate Est.Error 1-95% CI u-95% CI Rhat Bulk_ESS Ta
## Intercept          0.83      0.13    0.57    1.07 1.00    21805
## zi_Intercept       -2.94      0.70   -4.58   -1.84 1.00    23476
## genderM            0.08      0.07   -0.06    0.22 1.00    69575
## motherAnxious      0.46      0.17    0.13    0.79 1.00    33349
## motherAvoidant     -0.08      0.25   -0.56    0.41 1.00    28640

```

```

## motherFearful           0.63      0.40     -0.14      1.45 1.00      28081      31710
## fatherAnxious           0.07      0.17     -0.27      0.40 1.00      36862      37511
## fatherAvoidant          0.08      0.18     -0.26      0.43 1.00      36867      35693
## fatherFearful           0.36      0.34     -0.28      1.04 1.00      30611      33358
## motherAnxious:fatherAnxious -0.24      0.24     -0.70      0.22 1.00      31718      36313
## motherAvoidant:fatherAnxious 0.24      0.31     -0.38      0.85 1.00      28308      33805
## motherFearful:fatherAnxious -0.75      0.50     -1.75      0.22 1.00      30436      34570
## motherAnxious:fatherAvoidant -0.23      0.24     -0.70      0.24 1.00      32215      36746
## motherAvoidant:fatherAvoidant 0.11      0.30     -0.49      0.71 1.00      27005      33164
## motherFearful:fatherAvoidant -0.19      0.45     -1.09      0.69 1.00      27741      32785
## motherAnxious:fatherFearful -0.47      0.40     -1.25      0.31 1.00      27639      33504
## motherAvoidant:fatherFearful 0.28      0.48     -0.67      1.21 1.00      26850      35151
## motherFearful:fatherFearful -0.35      0.53     -1.41      0.68 1.00      24657      31248
## zi_genderM              -0.30      0.40     -1.10      0.48 1.00      49145      34266
##
## Family Specific Parameters:
##      Estimate Est.Error 1-95% CI u-95% CI Rhat Bulk_ESS Tail_ESS
## shape      2.46      0.32      1.90      3.17 1.00      38857      36697
##
## Samples were drawn using sampling(NUTS). For each parameter, Bulk_ESS
## and Tail_ESS are effective sample size measures, and Rhat is the potential
## scale reduction factor on split chains (at convergence, Rhat = 1).

```

10.2 Hypothesis Matrices

Hypothesis matrices are the same as in the analysis of the externalizing problems.

10.3 Centering and Adjusting

Centering and adjusting procedures are the same as in the analysis of the externalizing problems.

10.4 Results and Sensitivity

Bayes factor and posterior probability of each hypothesis are reported in Table~10.1.

Prior sensitivity analysis is conducted considering the same prior as in the analysis of the externalizing problems. The results of the prior sensitivity analysis are reported in Table~10.2.

Overall results consistently indicate the Monotropy Hypothesis as the most supported by the data. However, we can observe the same patterns as in the analysis of the externalizing problems. More diffuse prior, favour hypothesis

Table 10.1: Bayes factor encompassing model and hypothesis posterior probabilities ($n_{subj} = 847$).

Hypothesis	Bayes Factor	Posterior Probability	
Null	3.5e+11	0.03	○
Monotropy	9.2e+12	0.92	●
Hierarchy	4.7e+11	0.05	○
Independence	1.1e+10	0.00	○
Integration	1.7e+10	0.00	○

Table 10.2: Bayes factor encompassing model and hypothesis posterior probabilities (PP) under different prior settings ($n_{subj} = 847$).

Hypothesis	$\mathcal{N}(0, .5)$		$\mathcal{N}(0, 1)$		$\mathcal{N}(0, 3)$		$\mathcal{N}(0, 5)$		$\mathcal{N}(0, 10)$	
	BF	PP	BF	PP	BF	PP	BF	PP	BF	PP
Null	3.8e+01	0.00	8.9e+04	0.00	3.5e+11	0.03	5.5e+14	0.09	1.8e+19	0.27
Monotropy	2.3e+04	0.29	1.8e+07	0.63	9.2e+12	0.92	5.4e+15	0.89	4.9e+19	0.72
Hierarchy	4.0e+04	0.51	8.8e+06	0.30	4.7e+11	0.05	1.0e+14	0.02	2.2e+17	0.00
Independence	5.7e+03	0.07	6.4e+05	0.02	1.1e+10	0.00	1.5e+12	0.00	1.6e+15	0.00
Integration	9.8e+03	0.13	1.4e+06	0.05	1.7e+10	0.00	1.5e+12	0.00	8.4e+14	0.00

with more equality constraints. Whereas, tighter prior penalizes hypotheses with more equality constraints.

10.5 Selected Model

In the model, we consider only the role of gender and mother attachment as fixed effects of μ . In the R formula syntax, we have

```
# formula for p
p ~ gender + (1|ID_class)

# formula for mu
mu ~ gender + mother + (1|ID_class)
```

Again, we specify the same prior distributions as before. The resulting prior settings are

```
##           prior      class      coef      group resp dpar nlpar bound      source
##           normal(0, 3)      b      genderM      (vectorized)
##           normal(0, 3)      b      motherAnxious      (vectorized)
##           normal(0, 3)      b      motherAvoidant      (vectorized)
##           normal(0, 3)      b      motherFearful      (vectorized)
##           normal(0, 3)      b      genderM      zi      (vectorized)
##           normal(0, 3)      b      genderM      zi      (vectorized)
## student_t(3, 0.7, 2.5) Intercept      default
##           logistic(0, 1) Intercept      zi      default
## student_t(3, 0, 2.5)      sd      default
## student_t(3, 0, 2.5)      sd      zi      default
## student_t(3, 0, 2.5)      sd      ID_class      (vectorized)
## student_t(3, 0, 2.5)      sd      Intercept ID_class      (vectorized)
## student_t(3, 0, 2.5)      sd      ID_class      zi      (vectorized)
## student_t(3, 0, 2.5)      sd      Intercept ID_class      zi      (vectorized)
##           gamma(0.01, 0.01)      shape      default
```

The model is estimated using 6 independent chains with 6,000 iterations (warm-up 2,000). The model summary is presented below.

```
## Family: zero_inflated_negbinomial
## Links: mu = log; shape = identity; zi = logit
## Formula: internalizing_sum ~ gender + mother + (1 | ID_class)
##           zi ~ gender + (1 | ID_class)
## Data: data (Number of observations: 847)
## Samples: 6 chains, each with iter = 6000; warmup = 2000; thin = 1;
##           total post-warmup samples = 24000
##
## Group-Level Effects:
## ~ID_class (Number of levels: 50)
##           Estimate Est.Error 1-95% CI u-95% CI Rhat Bulk_ESS Tail_ESS
## sd(Intercept)      0.43      0.08      0.29      0.61 1.00      5203      10756
## sd(zi_Intercept)    2.18      0.53      1.35      3.43 1.00      11849      14075
##
## Population-Level Effects:
##           Estimate Est.Error 1-95% CI u-95% CI Rhat Bulk_ESS Tail_ESS
## Intercept          0.88      0.11      0.66      1.10 1.00      9744      13174
## zi_Intercept       -2.82      0.65     -4.33     -1.78 1.00      11781      12633
## genderM            0.07      0.07     -0.07      0.21 1.00      36050      19093
## motherAnxious      0.29      0.09      0.11      0.46 1.00      26032      19760
## motherAvoidant     0.12      0.10     -0.07      0.31 1.00      29724      20101
## motherFearful      0.48      0.12      0.25      0.71 1.00      26736      20112
## zi_genderM        -0.29      0.39     -1.06      0.46 1.00      28148      18297
##
## Family Specific Parameters:
##           Estimate Est.Error 1-95% CI u-95% CI Rhat Bulk_ESS Tail_ESS
```

```
## shape      2.53      0.33      1.95      3.25 1.00      21936      18166
##
## Samples were drawn using sampling(NUTS). For each parameter, Bulk_ESS
## and Tail_ESS are effective sample size measures, and Rhat is the potential
## scale reduction factor on split chains (at convergence, Rhat = 1).
```

Marginal effects are presented in Figure~10.1 and differences between mother attachment patterns are reported in Figure~10.2.

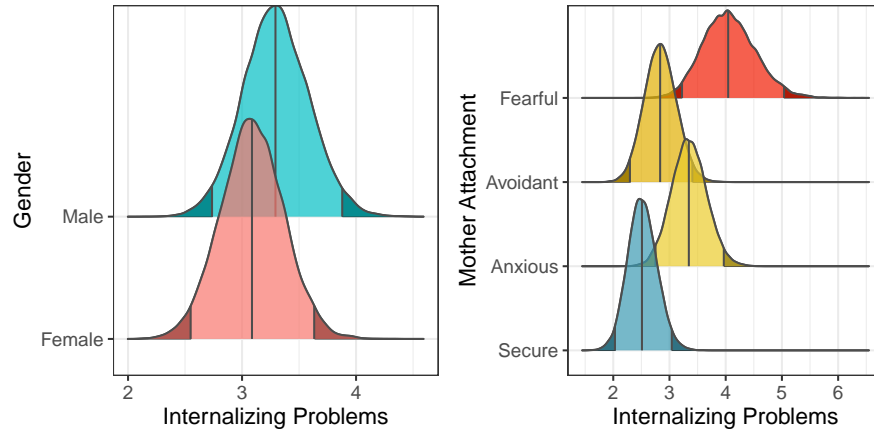


Figure 10.1: Marginal predicted values according to gender and mother attachment ($n_{subj} = 847$).

Overall, results indicate that Fearful, and Anxious children have more problems than Secure children. Moreover, Fearful children have more problems than Avoidant children. Finally, also the difference between Avoidant and Anxious children is close to the threshold.

To evaluate the fit of the model to the data, we computed the *Bayesian* R^2 using the function `brms::bayes_R2()`, and we present Posterior Predictions in Figure~10.3.

```
r2_int
##      Estimate Est.Error      Q2.5      Q97.5
## R2 0.2224337 0.02977471 0.1651088 0.281678
```

We can see that the actual variance explained by fixed effects and random effects is around 22%. Moreover, the posterior predictive check indicates a good fit to the data.

Conclusions

Considering attachment theoretical perspectives, results indicate only the role of mother attachment so we can support the Monotropy Theory.

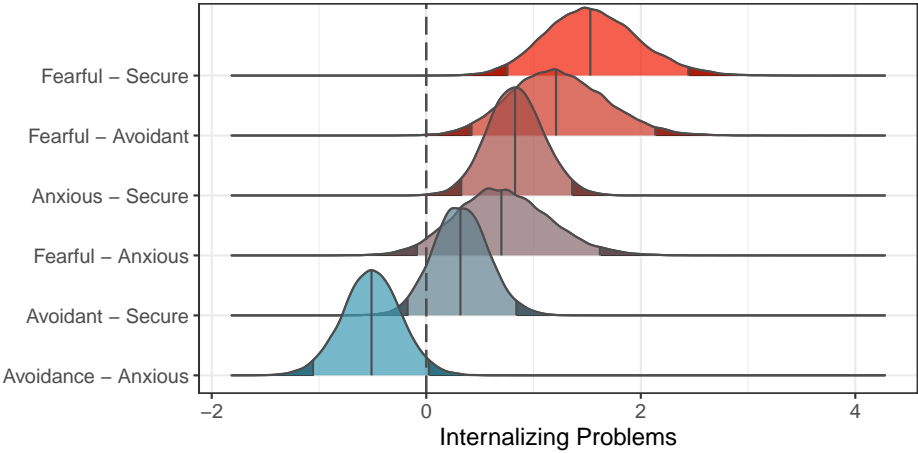


Figure 10.2: Predicted differences between mother attachment patterns ($n_{subj} = 847$).

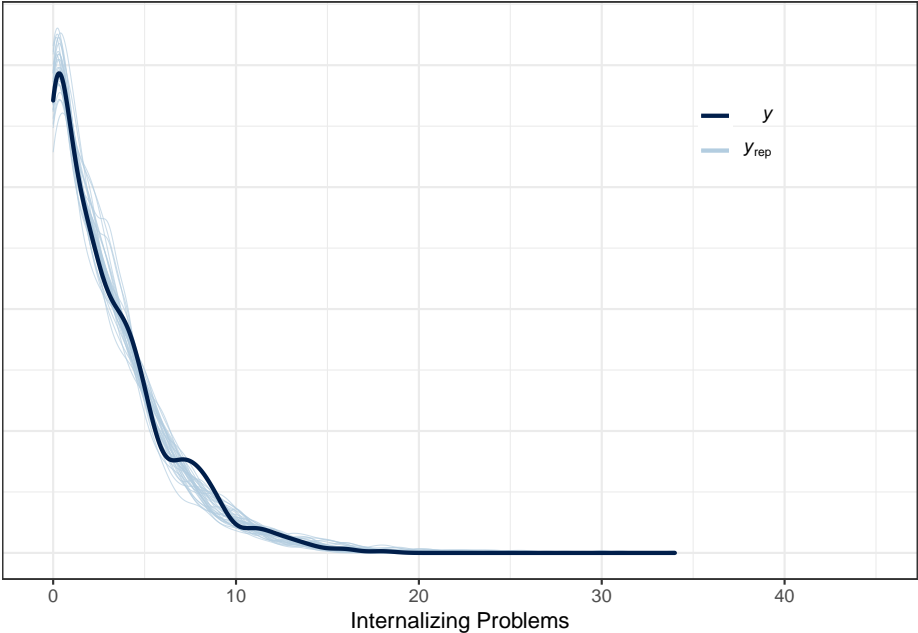


Figure 10.3: Posterior predictive check ($n_{subj} = 847$).

Chapter 11

Conclusions

Overall, we obtained consistent results from the three different approaches: only mother attachment influences children's externalizing problems. In particular, we observed the following pattern: Secure children have the lowest level of problems, Anxious and Avoidant children have a similar, intermediate, level of problems, and Fearful children have the highest level of problems. Compared to externalizing problems, however, differences between attachment groups are less pronounced. Taken together, these results support the Monotropy Theory.

Bibliography

- H. Akaike. Information theory and an extension of the maximum likelihood principle. In B.N. Petrov and F. Csaki, editors, *Proceedings of the Second International Symposium on Information Theory*, pages 267–281. Budapest Akademiai Kiado.
- Katrijn Brenning, Stijn Van Petegem, Janne Vanhalst, and Bart Soenens. The psychometric qualities of a short version of the Experiences in Close Relationships Scale – Revised Child version. 68:118–123. doi: 10.1016/j.paid.2014.04.005.
- Inge Bretherton. Fathers in attachment theory and research: A review. 180 (1-2):9–23. ISSN 0300-4430, 1476-8275. doi: 10.1080/03004430903414661. URL <http://www.tandfonline.com/doi/abs/10.1080/03004430903414661>.
- Mollie E. Brooks, Kasper Kristensen, Koen J. van Benthem, Arni Magnusson, Casper W. Berg, Anders Nielsen, Hans J. Skaug, Martin Maechler, and Benjamin M. Bolker. glmmTMB balances speed and flexibility among packages for zero-inflated generalized linear mixed modeling. 9(2):378–400. URL <https://journal.r-project.org/archive/2017/RJ-2017-066/index.html>.
- Paul-Christian Bürkner. Advanced Bayesian multilevel modeling with the R package brms. 10(1):395, a. ISSN 2073-4859. doi: 10.32614/RJ-2018-017. URL <https://journal.r-project.org/archive/2018/RJ-2018-017/index.html>.
- Paul-Christian Bürkner. Brms: An R package for Bayesian multilevel models using Stan. 80(1):1–28, b. ISSN 1548-7660. doi: 10.18637/jss.v080.i01. URL <http://www.jstatsoft.org/v80/i01/>.
- John Fox and Sanford Weisberg. *An R Companion to Applied Regression*. Sage, third edition. URL <https://socialsciences.mcmaster.ca/jfox/Books/Companion/>.
- Anna Goodman, Donna L. Lamping, and George B. Ploubidis. When to use broader internalising and externalising subscales instead of the hypothesised five subscales on the Strengths and Difficulties Questionnaire (SDQ): Data from British parents, teachers and children. 38(8):1179–1191. ISSN 1573-

2835. doi: 10.1007/s10802-010-9434-x. URL <https://doi.org/10.1007/s10802-010-9434-x>.
- Jouni Kuha. AIC and BIC: Comparisons of Assumptions and Performance. 33 (2):188–229. ISSN 0049-1241, 1552-8294. doi: 10.1177/0049124103262065. URL <http://journals.sagepub.com/doi/10.1177/0049124103262065>.
- Daniel Lüdtke, Mattan S. Ben-Shachar, Indrajeet Patil, Philip Waggoner, and Dominique Makowski. performance: An R package for assessment, comparison and testing of statistical models. 6(60):3139. doi: 10.21105/joss.03139.
- Richard McElreath. *Statistical Rethinking: A Bayesian Course with Examples in R and Stan*. CRC Texts in Statistical Science. Taylor and Francis, 2 edition. ISBN 978-0-367-13991-9.
- Shinichi Nakagawa, Paul C D Johnson, and Holger Schielzeth. The coefficient of determination R^2 and intra-class correlation coefficient from generalized linear mixed-effects models revisited and expanded. (14):20170213. doi: <http://dx.doi.org/10.1098/rsif.2017.0213>.
- Gideon Schwarz. Estimating the dimension of a model. 6(2):461–464. ISSN 0090-5364. doi: 10.1214/aos/1176344136.
- Luca Scrucca, Michael Fop, T. Brendan Murphy, and Adrian E. Raftery. mclust 5: Clustering, classification and density estimation using Gaussian finite mixture models. 8(1):289–317. URL <https://doi.org/10.32614/RJ-2016-021>.
- Stan Development Team. RStan: The R interface to Stan. URL <http://mc-stan.org/>.
- Eric-Jan Wagenmakers and Simon Farrell. AIC model selection using Akaike weights. 11(1):192–196. ISSN 1069-9384, 1531-5320. doi: 10.3758/BF03206482. URL <http://www.springerlink.com/index/10.3758/BF03206482>.