

Notes on Importance Sampling and Policy Gradient

Nan Jiang

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张楚珩

清华大学 交叉信息院博士在读

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这是UIUC姜楠老师开设的CS598统计强化学习（理论）课程的第六讲，这一讲的主要内容是Fitted Q-iteration。

原文传送门

CS598 Note6

nanjiang.cs.illinois.edu



一、Importance sampling

1. 估计期望

Consider the problem of estimating $\mathbb{E}_{x \sim p}[f(x)]$ for distribution $p \in \Delta(\mathcal{X})$ and function $f: \mathcal{X} \rightarrow \mathbb{R}$. If we can sample $x \sim p$, the standard Monte-Carlo estimate is $f(x)$, and averaging such estimates over multiple i.i.d. samples of x will give us an accurate estimate of $\mathbb{E}_{x \sim p}[f(x)]$. This is particularly useful if it is easy to sample from p but difficult to calculate the integral in $\mathbb{E}_{x \sim p}[f(x)]$.

Now what if we cannot sample from p , but have access to $x \sim q$ for some other distribution $q \in \Delta(\mathcal{X})$? It turns out that, if p is fully supported on q , that is, for all $x \in \mathcal{X}$ where $p(x) > 0$ we have $q(x) > 0$, then the following *importance weighted* estimator also gives an unbiased estimate of $\mathbb{E}_{x \sim p}[f(x)]$:

$$\frac{p(x)}{q(x)} f(x). \quad (1)$$

To verify unbiasedness:

$$\mathbb{E}_{x \sim q} \left[\frac{p(x)}{q(x)} f(x) \right] = \sum_{x \in \mathcal{X}} q(x) \frac{p(x)}{q(x)} f(x) = \sum_{x \in \mathcal{X}} p(x) f(x) = \mathbb{E}_{x \sim p}[f(x)].$$

$p(x)/q(x)$ has many names: importance weight, importance ratio, or inverse propensity score (IPS). A useful property of importance ratio to keep in mind is that

$$\mathbb{E}_{x \sim q} \left[\frac{p(x)}{q(x)} \right] = 1.$$

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当要估计 $\mathbb{E}_{\mathbf{a} \sim p} [f(\mathbf{a})]$ 的时候，如果可以从分布 p 中采样，那么可以直接把样本上的平均函数值作为该期望的无偏估计；如果只能从另外一个分布 q 中采样，那么可以使用一个调整系数 $\frac{p(\mathbf{a})}{q(\mathbf{a})}$ 来使其仍然是一个无偏估计。这样做的要求是 p is fully supported on q （具体定义见截图）。这种方法叫做 importance sampling（IS）。

2. 考虑单步情形

构造一个无偏估计

考虑单步的情形，即contextual bandit问题。同时假设状态分布为 μ ，奖励范围在 $[0, 1]$ 之间。假设有使用behavior policy $\mathbf{a} \sim \mu, \mathbf{a} \sim \pi_b(\mathbf{x})$ 采样到的数据集 $\{(\mathbf{x}, \mathbf{a}, r)\}$ ，目标是使用该数据集估计target policy下的性能 $v^* := \mathbb{E}[r | \mathbf{a} \sim \pi]$ 。

方式很简单，就是选用如下estimate

$$\rho r, \text{ where } \rho = \frac{\pi(\mathbf{a} | \mathbf{x})}{\pi_b(\mathbf{a} | \mathbf{x})}. \quad (3)$$

它是unbiased的，根据之前的结论，使用importance weight是unbiased，即

$$\mathbb{E}[r | \mathbf{a} \sim \pi] = \mathbb{E}_{(\mathbf{x}, \mathbf{a}, r) \sim p}[r] = \mathbb{E}_{(\mathbf{x}, \mathbf{a}, r) \sim q} \left[\frac{p(\mathbf{x}, \mathbf{a}, r)}{q(\mathbf{x}, \mathbf{a}, r)} r \right].$$

而该importance weight可以仅使用behavior policy和target policy来计算出来

$$\frac{p(\mathbf{x}, \mathbf{a}, r)}{q(\mathbf{x}, \mathbf{a}, r)} = \frac{\mu(\mathbf{x})\pi(\mathbf{a} | \mathbf{x})R(r | \mathbf{x}, \mathbf{a})}{\mu(\mathbf{x})\pi_b(\mathbf{a} | \mathbf{x})R(r | \mathbf{x}, \mathbf{a})} = \frac{\pi(\mathbf{a} | \mathbf{x})}{\pi_b(\mathbf{a} | \mathbf{x})} = \rho.$$

方差分析

下面来分析这种方法的方差，为了便于分析，假设behavior policy是对于各个action均匀采样，而target policy是一个确定性策略，同时奖励是一个确定性的常数。那么其方差可以较为容易地推导出来

$$\begin{aligned}
\mathbb{V}[\rho r | a \sim U] &= r^2 \mathbb{V}[\rho | a \sim U] \\
&= r^2 (\mathbb{E}[\rho^2 | a \sim U] - (\mathbb{E}[\rho | a \sim U])^2) \\
&= r^2 (\mathbb{E}[\rho^2 | a \sim U] - 1) \quad (\text{the mean of } \rho \text{ is always } 1) \\
&= r^2 \left(\mathbb{E} \left[\frac{\mathbb{I}[a = \pi(x)]}{1/K^2} \mid a \sim U \right] - 1 \right) \\
&= r^2 (K - 1).
\end{aligned}$$

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其中 $K := |\mathcal{A}|$ 是可能动作的个数。观察到，当 behavior policy = target policy 时，方差应该为零，而加上了 IS 之后产生了更大的方差。可以想象成只有采样到的样本选择的动作和 target policy 选择的动作相同的时候，该样本才能用上，因此会有了一个产生与 K 有关的方差。

当奖励不是确定性的常数而是在 $[0, 1]$ 之间的随机变量时，方差上界为

$$\mathbb{V}[\rho r | a \sim U] \leq \mathbb{E}[\rho^2 r^2 | a \sim U] \leq \mathbb{E}[\rho^2 | a \sim U] = K.$$

有限样本分析

下面考虑使用有限数目的样本能够使用 ρ 估计 r 到何种精度。使用 Bernstein 不等式相比于 Hoeffding 不等式能够得到更紧的上界。考虑 $r \in [0, 1], \rho \in [0, K], \mathbb{V}[\rho] \leq K$ ，可以得到上界为

$$\sqrt{\frac{2K}{n} \ln \frac{2}{\delta}} + \frac{2K}{3n} \ln \frac{2}{\delta}.$$

其中使用到的 Bernstein 不等式可以表述为

Theorem Let x_1, x_2, \dots, x_n be independent bounded random variables such that $\mathbb{E}x_i = 0$ and $|x_i| \leq \varsigma$ with probability 1 and let $\sigma^2 = \frac{1}{n} \sum_{i=1}^n \text{Var}\{x_i\}$. Then for any $a > 0$ we have

$$P\left(\frac{1}{n} \sum_{i=1}^n x_i \geq \epsilon\right) \leq e^{-\frac{n\epsilon^2}{2\sigma^2 + 2\varsigma\epsilon/3}}$$

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Weighted importance sampling

回到奖励是确定性常数以及 target policy 为确定性策略的情形，在这种情况下，如果只使用和 target policy 一致的样本 (subsamples)，那么应该该估计的方差应该为零，但是前面介绍的 IS 方法得到的方差却不为零。IS 方法的估计为

$$\frac{1}{n} \sum_{i=1}^n \frac{\mathbb{I}[a_i = \pi(x_i)]}{1/K} r_i = \frac{1}{n/K} \sum_{i: a_i = \pi(x_i)} r_i.$$

如果分母部分为 $|\{i: a_i = \pi(x_i)\}|$ ，那么就相当于在 subsample 上做估计，得到的估计方差为零；但是现

在这里分母却不是 $|\{i: a_i = \pi(a_i)\}|$ ，而是它的期望 n/K ，这导致了估计的方差大于零。weighted importance sampling (WIS) 则另外构造了下面这样的估计方法，可以在此情况下方差为零，起到了减小方差的目的。

$$\frac{1}{\sum_{i=1}^n \rho_i} \sum_{i=1}^n \rho_i r_i. \quad (4)$$

WIS的缺点是它是biased的，即样本有限时，期望不等于真实数值；不过它是consistent的，即当样本数目趋向于无穷的时候，其分布会趋向于真实分布。

进一步减小方差

根据前面的结果，方差和奖励的平方有关，由此想到可以先把奖励减去某个常数 c ，然后求导相应的估计之后，再加上这个常数，由此能够得到更小的方差 $(r-c)^2(K-1)$ 。

进一步推广可以得到 doubly robust (DR) estimate，即使用一个估计来的 $\hat{Q}(x, a)$ 来替代前面根据先验知识设置的 c ，并且期望 $\hat{Q}(x, a) \approx \mathbb{E}_{\pi'}[r|x, a]$ 。

$$\mathbb{E}_{a' \sim \pi} [\hat{Q}(x, a')] + \rho (r - \hat{Q}(x, a)). \quad (5)$$

3. 考虑多步情形

无偏估计

类似地，对于多步情形（即标准的RL设定），可以得到无偏的估计

$$\begin{aligned} v^\pi &= \mathbb{E} \left[\sum_{h=1}^H \gamma^{h-1} r_h \mid a_{1:H} \sim \pi \right] = \mathbb{E}_{\tau \sim p} \left[\sum_{h=1}^H \gamma^{h-1} r_h \right] = \mathbb{E}_{\tau \sim q} \left[\frac{p(\tau)}{q(\tau)} \sum_{h=1}^H r_h \right] \\ &= \mathbb{E}_{\tau \sim q} \left[\frac{\mu(s_1) \pi(a_1|s_1) R(r_1|s_1, a_1) P(s_2|s_1, a_1) \cdots \pi(a_H|s_H) R(r_H|s_H, a_H)}{\mu(s_1) \pi_b(a_1|s_1) R(r_1|s_1, a_1) P(s_2|s_1, a_1) \cdots \pi_b(a_H|s_H) R(r_H|s_H, a_H)} \sum_{h=1}^H \gamma^{h-1} r_h \right] \\ &= \mathbb{E}_{\tau \sim q} \left[\frac{\pi(a_1|s_1) \cdots \pi(a_H|s_H)}{\pi_b(a_1|s_1) \cdots \pi_b(a_H|s_H)} \sum_{h=1}^H \gamma^{h-1} r_h \right] = \mathbb{E} \left[\frac{\pi(a_1|s_1) \cdots \pi(a_H|s_H)}{\pi_b(a_1|s_1) \cdots \pi_b(a_H|s_H)} \sum_{h=1}^H \gamma^{h-1} r_h \right] \end{aligned}$$

由此可以得到 **per-trajectory IS estimator**

So the expression in the bracket is an unbiased estimate of v^π . Let $\rho_h := \pi(a_h|s_h)/\pi_b(a_h|s_h)$ and $\rho_{1:h}$ be a shorthand for $\prod_{h'=1}^h \rho_{h'}$, the **per-trajectory IS estimator** is [2, 3]:

$$\rho_{1:H} \sum_{h=1}^H \gamma^{h-1} r_h. \quad (6)$$

观察到第h步之后的样本不会影响到第h步的奖励，因此，对于第h步来说，计算IS的时候，可以把h步之后都去掉。得到 **per-step IS estimator**

$$\sum_{h=1}^H \gamma^{h-1} \rho_{1:h} r_h. \quad (7)$$

它还可以写成递归的形式，即

$$v_{H-h+1} := \rho_h(r_h + \gamma v_{H-h}). \quad (8)$$

类似地，可以得到**DR estimator**

$$v_{H-h+1}^{DR} := \mathbb{E}_{a \sim \pi}[\hat{Q}^\pi(s_h, a)] + \rho_h(r_h + \gamma v_{H-h}^{DR} - \hat{Q}^\pi(s_h, a_h)) \quad (9)$$

通过递归的形式，可以顺着递归地证明上述estimator是都是无偏的。

方差分析

下面分析DR estimator的方差

Variance of per-step IS The variance of Eq.(7) also satisfies an interesting recursion, which has important implications outside off-policy evaluation. Let $\mathbb{V}_h[\cdot]$ and $\mathbb{E}_h[\cdot]$ denote conditional variance and expectation, respectively, conditioned on $s_1, a_1, r_1, \dots, s_{h-1}, a_{h-1}, r_{h-1}$. For simplicity assume reward is a deterministic function of state and action, then

$$\begin{aligned} \mathbb{V} V_h[v_{H-h+1}] &= \mathbb{E}_h[v_{H-h+1}^2] - (\mathbb{E}_h[v_{H-h+1}])^2 \\ &= \mathbb{E}_h[v_{H-h+1}^2] - (\mathbb{E}_h[V^\pi(s_h)])^2 & (V^\pi(s_h) = \mathbb{E}_h[v_{H-h+1} | s_h]) \\ &= \mathbb{E}_h[(\rho_h Q^\pi(s_h, a_h) + \rho_h(r_h + \gamma v_{H-h} - Q^\pi(s_h, a_h)))^2] - (\mathbb{E}_h[V^\pi(s_h)])^2 \\ &= \mathbb{E}_h[(\rho_h Q^\pi(s_h, a_h))^2] + \mathbb{E}_h[\rho_h^2(r_h + \gamma v_{H-h} - Q^\pi(s_h, a_h))^2] - (\mathbb{E}_h[V^\pi(s_h)])^2 \\ &= \mathbb{E}_h[(V^\pi(s_h) + \rho_h Q^\pi(s_h, a_h) - V^\pi(s_h))^2] + \gamma^2 \mathbb{E}_h[\rho_h^2(v_{H-h} - V^\pi(s_{h+1}))^2] - (\mathbb{E}_h[V^\pi(s_h)])^2 \\ &= \mathbb{E}_h[V^\pi(s_h)^2] + \mathbb{E}_h[\mathbb{V}_h[\rho_h Q^\pi(s_h, a_h) | s_h]] + \gamma^2 \mathbb{E}_h[\rho_h^2 \mathbb{V}_{h+1}[v_{H-h}]] - (\mathbb{E}_h[V^\pi(s_h)])^2 \\ &= \mathbb{V}_h[V^\pi(s_h)] + \mathbb{E}_h[\mathbb{V}_h[\rho_h Q^\pi(s_h, a_h) | s_h]] + \gamma^2 \mathbb{E}_h[\rho_h^2 \mathbb{V}_{h+1}[v_{H-h}]]. \end{aligned}$$

考虑on-policy并且测量是确定性的情形，上式可以化简为如下形式，即Bellman equation for variance。

$$\mathbb{V}_h[v_{H-h+1}] = \mathbb{V}_h[V^\pi(s_h)] + \gamma^2 \mathbb{E}_h[\mathbb{V}_{h+1}[v_{H-h}]].$$

二、策略梯度

略。主要讲了策略梯度的推导和使用baseline来减小variance，本专栏的入门系列有讲，同时贴一下个人之前总结的笔记。

Define value functions and optimization objective as usual.

Definition 1.1 (State value function).

$$V^\pi(s) = \mathbb{E}[\sum_{t=0}^{\infty} \sum_{s'} p(s_t = s' | s_0 = s, \pi_\theta) \gamma^t r_t] \quad (1)$$

Definition 1.2 (Action value function).

$$Q^\pi(s, a) = \mathbb{E}[\sum_{t=0}^{\infty} \sum_{s'} p(s_t = s' | s_0 = s, a_0 = a, \pi_\theta) \gamma^t r_t] \quad (2)$$

Definition 1.3 (Optimization objective). *The ultimate goal of model-free reinforcement learning is to maximize*

$$\eta(\pi) := \mathbb{E}_{s_0}[V^\pi(s_0)] \quad (3)$$

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1.1 Policy gradient

Theorem 1.1 (Policy gradient).

$$\nabla_\theta \eta(\pi_\theta) = \mathbb{E}_{s \sim \rho_\pi, a \sim \pi}[\nabla_\theta \log \pi_\theta(a|s) Q^\pi(s, a)]$$

where $\rho_\pi(s) = \mathbb{E}_{s_0}[\sum_{t=0}^{\infty} \gamma^t p(s_t = s | s_0, \pi)]$ is the state visitation frequency.

Proof. Take derivative and iteratively unroll.

$$\begin{aligned} \nabla_\theta V^\pi(s_0) &= \nabla_\theta [\sum_{a_0} \pi_\theta(a_0 | s_0) Q^\pi(s_0, a_0)] \\ &= \sum_{a_0} [\nabla_\theta \pi_\theta(a_0 | s_0) Q^\pi(s_0, a_0) + \pi_\theta(a_0 | s_0) \nabla_\theta Q^\pi(s_0, a_0)] \\ &= \sum_{a_0} [\nabla_\theta \pi_\theta(a_0 | s_0) Q^\pi(s_0, a_0) + \pi_\theta(a_0 | s_0) \nabla_\theta \sum_{s_1, r_0} p(s_1, r_0 | s_0, a_0) (r + \gamma V^\pi(s_1))] \\ &= \sum_{a_0} [\nabla_\theta \pi_\theta(a_0 | s_0) Q^\pi(s_0, a_0) + \pi_\theta(a_0 | s_0) \sum_{s_1} \gamma p(s_1 | s_0, a_0) \nabla_\theta V^\pi(s_1)] \\ &= \sum_{s \in \mathcal{S}} \sum_{t=0}^{\infty} \gamma^t p(s_t = s | s_0, \pi_\theta) \mathbb{E}_{a \sim \pi}[\nabla_\theta \log \pi_\theta(a | s) Q^\pi(s, a)] \end{aligned}$$

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$$\begin{aligned}\nabla_{\theta}\eta(\pi_{\theta}) &= \sum_{s_0} \sum_{s \in \mathcal{S}} \sum_{t=0}^{\infty} p(s_0) \gamma^t p(s_t = s | s_0, \pi_{\theta}) \mathbb{E}_{a \sim \pi} [\nabla_{\theta} \log \pi_{\theta}(a|s) Q^{\pi}(s, a)] \\ &= \mathbb{E}_{s \sim \rho_{\pi}, a \sim \pi} [\nabla_{\theta} \log \pi_{\theta}(a|s) Q^{\pi}(s, a)]\end{aligned}$$

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1.2 Policy gradient with baseline

Theorem 1.2 (Policy gradient with baseline). *Baseline $b(s)$ dose not change expected value of policy gradient.*

$$\nabla_{\theta}\eta(\theta) = \mathbb{E}_{s \sim \rho_{\pi}, a \sim \pi} [\nabla_{\theta} \log \pi_{\theta}(a|s) (Q^{\pi}(s, a) - b(s))] \quad (4)$$

Moreover, when the baseline is chosen to be

$$b^*(s) = \frac{\mathbb{E}_{a \sim \pi} [\log \pi_{\theta}(a|s)^T \log \pi_{\theta}(a|s) Q^{\pi}(s, a)]}{\mathbb{E}_{a \sim \pi} [\log \pi_{\theta}(a|s)^T \log \pi_{\theta}(a|s)]} \quad (5)$$

the variance of policy gradient is minimized.

Proof. In terms of expected value, we only need to notice that

$$\mathbb{E}_a [\nabla_{\theta} \log \pi_{\theta}(a|s) b(s)] = \sum_a [\nabla_{\theta} \pi_{\theta}(a|s)] b(s) = \nabla_{\theta} (\sum_a \pi_{\theta}(a|s)) b(s) = 0$$

In terms of the variance, let $g = \nabla_{\theta} \log \pi_{\theta}(a|s) (Q^{\pi}(s, a) - b(s))$ and $g_0 = \nabla_{\theta} \log \pi_{\theta}(a|s) Q^{\pi}(s, a)$.

$$\begin{aligned}Var(g) &= \mathbb{E}_{s \sim \rho_{\pi}, a \sim \pi} [(g - \mathbb{E}_{s \sim \rho_{\pi}, a \sim \pi} [g])^T (g - \mathbb{E}_{s \sim \rho_{\pi}, a \sim \pi} [g])] \\ &= Var(g_0) - 2 \mathbb{E}_{s \sim \rho_{\pi}, a \sim \pi} [\log \pi_{\theta}(a|s)^T \log \pi_{\theta}(a|s) Q^{\pi}(s, a)] b(s) \\ &\quad + \mathbb{E}_{s \sim \rho_{\pi}, a \sim \pi} [\log \pi_{\theta}(a|s)^T \log \pi_{\theta}(a|s)] b^2(s)\end{aligned}$$

Optimal baseline can be found to minimize the variance

$$b^*(s) = \frac{\mathbb{E}_{a \sim \pi} [\log \pi_{\theta}(a|s)^T \log \pi_{\theta}(a|s) Q^{\pi}(s, a)]}{\mathbb{E}_{a \sim \pi} [\log \pi_{\theta}(a|s)^T \log \pi_{\theta}(a|s)]}$$

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 \square

For convenience, a popular baseline is chosen to be

$$b(s) = \mathbb{E}_{a \sim \pi} Q^\pi(s, a) = V^\pi(s) \quad (7)$$

or function approximated state value function

$$b(s) = V_\phi^\pi(s) \quad (8)$$

This is an unbiased estimate of policy gradient, though function approximated state value function is used for baseline.

$$\nabla_\theta \eta(\theta) = \mathbb{E}_{s \sim \rho_\pi, a \sim \pi} [\nabla_\theta \log \pi_\theta(a|s) (Q^\pi(s, a) - V_\phi^\pi(s))] \quad (9)$$

When the state value function is unbiased (e.g. learn from Monte Carlo return samples), the following policy gradient estimate is also unbiased, where s' is the next state following s and a .

$$\nabla_\theta \eta(\theta) = \mathbb{E}_{s \sim \rho_\pi, a \sim \pi} [\nabla_\theta \log \pi_\theta(a|s) (r + \gamma V_\phi^\pi(s') - V_\phi^\pi(s))] \quad (10)$$

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