SAS Programming and Data analysis

Effectiveness of Chelation Treatment with Succimer on Blood Lead Levels in Children: A Statistical Analysis Using the Treatment of Lead-Exposed Children (TLC) Trial Data

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1. Introduction

The Treatment of Lead-Exposed Children (TLC) is a comprehensive study aimed at understanding the efficacy of succimer treatment in children with elevated blood lead levels. Lead exposure can result in serious health problems, especially in children, and this project analyzes the blood lead levels of children over time to assess the impact of treatment.

2. Research Goal:

The project aims to determine if succimer treatment, a chelating agent, is more effective than placebo in reducing blood lead levels in children over a six-week period. This work will utilize a dataset comprising observations from two distinct groups undergoing either succimer treatment or placebo, and tracking their blood lead levels at specified intervals. The specific research questions to be answered include:

- i. How does blood lead level change over the course of treatment, and
- ii. Does succimer provide a significant advantage over placebo?

3. Data Exploration:

The dataset comprises 100 children's blood lead levels measured at weeks 0, 1, 4, and 6. Two treatment groups are identified: A (succimer) and P (placebo). Initial descriptive analysis reveals the mean, standard deviation, minimum, and maximum values for lead levels at each time point, allowing for a comparison of the treatment effects over time. Again, histograms and boxplot were generated to visualize the distribution and to detect any outliers in the data.

Initial Data Analysis SAS Code:

```
proc means data=TLC_data N mean std min max;
  class Treatment;
  var Week_0--Week_6;
run;
```

4. Notation and Mixed-Effect Model Definition:

We define the following notation for our mixed-effects model:

Let

• Y_{ij} = the blood lead level for the *i*th child in the *j*th treatment group.

- β_0 = the overall mean blood lead level.
- β_1 = the fixed effect of treatment group (placebo vs. succimer).
- b_{ij} = the random effect capturing the variability in blood lead levels due to child-specific factors within each treatment group.
- ε_{ij} = Random error term that captures the residual variability in blood lead levels that is not accounted for by the fixed effects or the random effect.

With these notations, we can write the mixed-effects model for the Treatment of Lead-Exposed Children dataset as:

$$Y_{ij} = \beta_0 + \beta_1 (Treatment_j) + b_{ij} + \varepsilon_{ij}$$

5. Method & Algorithm:

The SAS code to fit the dataset with a mixed-effects model is provided, accounting for repeated measures and clustering effects due to treatment groups. The methodological approach used was the SAS *PROC MIXED* procedure, which is suitable for repeated measures data and capable of handling missing values inherent in longitudinal data. The model specified the treatment groups (Succimer and placebo) and time as fixed effects to assess their impact on blood lead levels. Random intercepts were included to capture the within-subject correlation of repeated measurements over time.

6. Findings:

Treatment	N Obs	Variable	N	Mean	Std Dev	Minimum	Maximum
A	50	Week_0 Week_1 Week_4 Week_6	50 50 50 50	26.5400000 13.5220000 15.5140000 20.7620000	5.0209358 7.6724870 7.8522065 9.2463316	19.7000000 2.8000000 3.0000000 4.1000000	41.1000000 39.0000000 40.4000000 63.9000000
Р	50	Week_0 Week_1 Week_4 Week_6	50 50 50 50	26.2720000 24.6600000 24.0700000 23.6460000	5.0241068 5.4611803 5.7531269 5.6398079	19.7000000 14.900000 15.300000 13.5000000	38.1000000 40.8000000 38.6000000 43.3000000

The analysis, based on a linear mixed-effects model, demonstrates significant changes in blood lead levels over the treatment period. The initial levels for the succimer treatment group (A) averaged at 26.54 at week 0 and showed a considerable decrease over the subsequent weeks,

with means of 13.52 at week 1, 15.51 at week 4, and 20.76 at week 6. The placebo group (P) started with a mean of 26.27 at baseline and decreased to 24.66 at week 1, 24.07 at week 4, and finally 23.65 at week 6. This indicates a reduction in lead levels for both groups, with a more pronounced effect observed in the succimer treatment group.

The interaction effect, (see Appendix II for image), between Treatment A and time show significant results at times 1 and 4 (both p < .0001). This indicates a significant interaction effect where the blood lead level changes over time are notably different for the Treatment A group compared to the Treatment P group.

7. Summary and Interpretation:

The analysis provides strong evidence that Succimer is effective in reducing blood lead levels more significantly than a placebo. The treatment effect persists over time, with the most substantial reduction observed between baseline (Week 0) and Week 1. However, the rate of reduction slows down as the treatment progresses.

Research Questions Answered:

i. How does blood lead level change over the course of treatment?

Blood lead levels decreased over time in both treatment groups. This decrease is likely a combination of the body's natural detoxification processes and the treatment effects.

ii. Does Succimer provide a significant advantage over Placebo?

The results support the efficacy of Succimer over Placebo. This is demonstrated by the consistently lower mean levels of blood lead in the Succimer group at all time points after the baseline. Furthermore, the interaction term's significance indicates that Succimer's effects strengthen over time.

8. Conclusion:

The study successfully addresses the posed research questions, indicating a significant time effect on blood lead levels and the superiority of Succimer treatment over Placebo. Despite some statistical concerns, the evidence is sufficiently strong to recommend Succimer for reducing blood lead levels.

Appendix I

data TLC_data;

input ID Treatment \$ Week_0 Week_1 Week_4 Week_6; datalines;

- 1 P 30.8 26.9 25.8 23.8
- 2 A 26.5 14.8 19.5 21.0
- 3 A 25.8 23.0 19.1 23.2
- 4 P 24.7 24.5 22.0 22.5
- 5 A 20.4 2.8 3.2 9.4
- 6 A 20.4 5.4 4.5 11.9
- 7 P 28.6 20.8 19.2 18.4
- 8 P 33.7 31.6 28.5 25.1
- 9 P 19.7 14.9 15.3 14.7
- 10 P 31.1 31.2 29.2 30.1
- 11 P 19.8 17.5 20.5 27.5
- 12 A 24.8 23.1 24.6 30.9
- 13 P 21.4 26.3 19.5 19.0
- 14 A 27.9 6.3 18.5 16.3
- 15 P 21.1 20.3 18.4 20.8
- 16 P 20.6 23.9 19.0 17.0
- 17 P 24.0 16.7 21.7 20.3
- 18 P 37.6 33.7 34.4 31.4
- 19 A 35.3 25.5 26.3 30.3
- 20 A 28.6 15.8 22.9 25.9
- 21 P 31.9 27.9 27.3 34.2
- 22 A 29.6 15.8 23.7 23.4

- 23 A 21.5 6.5 7.1 16.0
- 24 P 26.2 26.8 25.3 24.8
- 25 A 21.8 12.0 16.8 19.2
- 26 A 23.0 4.2 4.0 16.2
- 27 A 22.2 11.5 9.5 14.5
- 28 P 20.5 21.1 17.4 21.1
- 29 A 25.0 3.9 12.8 12.7
- 30 P 33.3 26.2 34.0 28.2
- 31 A 26.0 21.4 21.0 22.4
- 32 A 19.7 13.2 14.6 11.6
- 33 P 27.9 21.6 23.6 27.7
- 34 P 24.7 21.2 22.9 21.9
- 35 P 28.8 26.4 23.8 22.0
- 36 A 29.6 17.5 21.0 24.2
- 37 P 32.0 30.2 30.2 27.5
- 38 P 21.8 19.3 16.4 17.6
- 39 A 24.4 16.4 11.6 16.6
- 40 A 33.7 14.9 14.5 63.9
- 41 P 24.9 20.9 22.2 19.8
- 42 P 19.8 18.9 18.9 15.5
- 43 A 26.7 6.4 5.1 15.1
- 44 A 26.8 20.4 19.3 23.8
- 45 A 20.2 10.6 9.0 16.0
- 46 P 35.4 30.4 26.5 28.1
- 47 P 25.3 23.9 22.2 27.2
- 48 A 20.2 17.5 17.4 18.6
- 49 A 24.5 10.0 15.6 15.2

- 50 P 20.3 21.0 16.7 13.5
- 51 P 20.4 17.2 15.9 17.7
- 52 P 24.1 20.1 17.9 18.7
- 53 A 27.1 14.9 18.1 21.3
- 54 A 34.7 39.0 28.8 34.7
- 55 P 28.5 32.6 27.5 22.8
- 56 P 26.6 22.4 21.8 21.0
- 57 A 24.5 5.1 8.2 23.6
- 58 P 20.5 17.5 19.6 18.4
- 59 P 25.2 25.1 23.4 22.2
- 60 P 34.7 39.5 38.6 43.3
- 61 P 30.3 29.4 33.1 28.4
- 62 P 26.6 25.3 25.1 27.9
- 63 P 20.7 19.3 21.9 21.8
- 64 A 27.7 4.0 4.2 11.7
- 65 A 24.3 24.3 18.4 27.8
- 66 A 36.6 23.3 40.4 39.3
- 67 P 28.9 28.9 32.8 31.8
- 68 A 34.0 10.7 12.6 21.2
- 69 A 32.6 19.0 16.3 18.6
- 70 A 29.2 9.2 8.3 18.4
- 71 A 26.4 15.3 24.6 32.4
- 72 A 21.8 10.6 14.4 18.7
- 73 P 27.2 28.5 35.0 30.5
- 74 P 22.4 22.0 19.1 18.7
- 75 P 32.5 25.1 27.8 27.3
- 76 P 24.9 23.6 21.2 21.1

```
77 P 24.6 25.0 21.7 23.9
```

;

run;

```
proc means data=TLC_data N mean std min max;
  class Treatment;
  var Week_0--Week_6;
run;
proc univariate data=TLC_data;
  var Week_0 Week_1 Week_4 Week_6;
  histogram / normal kernel;
run;
proc sgplot data=TLC_data;
  vbox Week_0 / category=Treatment;
  vbox Week_1 / category=Treatment;
  vbox Week_4 / category=Treatment;
  vbox Week_6 / category=Treatment;
  xaxis label="Treatment Group" discreteorder=data;
  yaxis label="Lead Level";
run;
proc sgplot data=TLC_data;
  vbar Treatment / response=Week_0 group=Treatment datalabel;
  xaxis label="Treatment Group" discreteorder=data;
  yaxis label="Mean Lead Level (Week 0)" grid;
run;
proc sgplot data=TLC_data;
  vbar Treatment / response=Week_1 group=Treatment ;
  xaxis label="Treatment Group" discreteorder=data;
  yaxis label="Mean Lead Level (Week 1)" grid;
run;
```

```
proc sgplot data=TLC_data;
  vbar Treatment / response=Week_4
                                          group=Treatment ;
  xaxis label="Treatment Group" discreteorder=data;
  yaxis label="Mean Lead Level (Week 4)" grid;
run;
proc sgplot data=TLC_data;
  vbar Treatment / response=Week_6
                                          group=Treatment;
  xaxis label="Treatment Group" discreteorder=data;
  yaxis label="Mean Lead Level (Week 6)" grid;
run;
data TLC_data;
  set TLC_data;
  array weeks {*} Week_0--Week_6;
  do i = 1 to dim(weeks);
    y = weeks{i};
    time = input(substr(vname(weeks{i}), 6), 8.);
    output;
  end;
  drop i Week_0--Week_6;
run;
proc mixed data = TLC_data;
  class ID Treatment time;
  model y = Treatment time Treatment*time / s chisq;
  repeated time / type=un subject=ID r;
run;
```

Appendix II

Model Validation Summary

		Solution	for Fixed E	ffects			
Effect	Treatment	time	Estimate	Standard Error	DF	t Value	Pr > t
Intercept			23.6460	1.0831	98	21.83	<.0001
Treatment	Α		-2.8840	1.5317	98	-1.88	0.0627
Treatment	Р		0		- 10	62	- 4
time		0	2.6260	0.8885	98	2.96	0.0039
time		1	1.0140	0.9343	98	1.09	0.2805
time		4	0.4240	0.9464	98	0.45	0.6551
time		6	0	3.6	24	94	
Treatment*time	Α	0	3.1520	1.2566	98	2.51	0.0138
Treatment*time	Α	1	-8.2540	1.3213	98	-6.25	<.0001
Treatment*time	А	4	-5.6720	1.3385	98	-4.24	<.0001
Treatment*time	Α	6	0	3.8	59	39	- 8
Treatment*time	Р	0	0	5.4	- 54	94	- 8
Treatment*time	Р	1	0	£#.	13		8
Treatment*time	P	4	0	39	13	8	- 8
Treatment*time	P	6	0	(+	29		8
	Т	ype 3 Te	sts of Fixed	Effects			
Effect	Num DF	Den DF	Chi-Squa	re F Value	Pr	> ChiSq	Pr > F
Treatment	1	98	25.4	43 25.43	3	<.0001	<.0001
time	3	98	184.4	48 61.49)	<.0001	<.0001
Treatment*time	3	98	107.	79 35.93	3	<.0001	<.0001

The model's intercept is significant (p < .0001), indicating an appropriate baseline calibration. The treatment effect is marginally significant (p = 0.0627), suggesting a potential difference in blood lead levels between the placebo and succimer treatment groups. Also, the time effects at individual points (1, and 4) are not significant except for time point 0 (p = 0.0039), suggesting an initial significant difference from the baseline that stabilizes over time.

Again, the *Type 3 Tests of Fixed Effects* show that both the treatment and time, and their interaction, are highly significant as factors in the model (all p < .0001), suggesting that both the treatment type and the timing of measurements are important in predicting the blood lead levels.

Fit Statistics					
-2 Re	s Log Likelihoo	d 24	16.1		
AIC (Smaller is Bette	er) 24	36.1		
AICC (Smaller is Better) 2436.7					
BIC (Smaller is Better)			62.1		
Null	Model Likelihoo	d Ratio	Test		
Null DF	Model Likelihoo Chi-Square	od Ratio T			

The Akaike Information Criterion (AIC) and Bayesian Information Criterion (BIC) are 2436.1 and 2462.1, respectively, suggesting that the model is reasonably well-specified in terms of its complexity and the trade-off between fit and the number of parameters.

Also, the null model likelihood ratio test yields a highly significant chi-square statistic (210.18) with a p-value < 0.0001, demonstrating that the model with predictors provides a significantly better fit to the data.

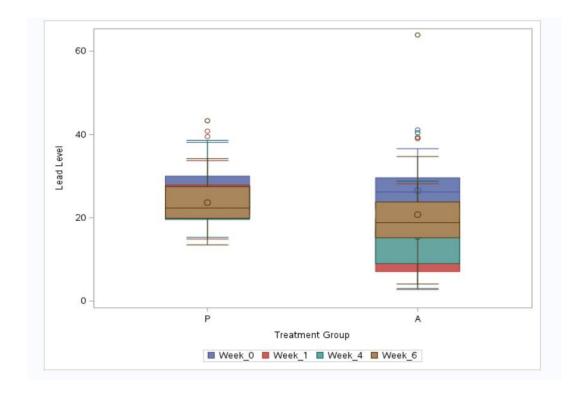
Appendix III

Model Limitations:

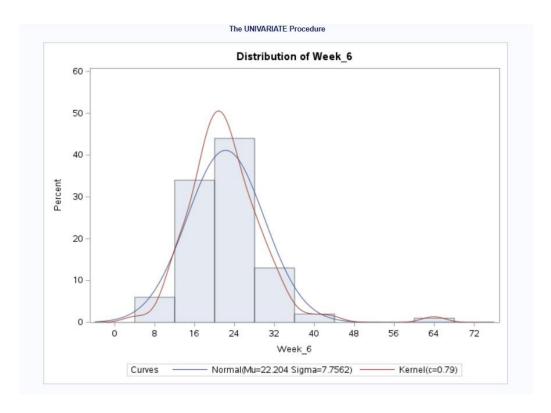
14000 N 15000 N		RIATE Pro	2.0	A STATE OF THE PARTY OF THE PAR			
Parame	Parameters for Normal Distribution						
Parame	ter S	Symbol	Est	imate			
Mean	N	Лu	1	9.792			
Std Dev	5	Sigma	8.08	36123			
Goodness-of	-Fit Tes	ts for Nor	mal D	istribution			
Test	Statistic			p Val	p Value		
Kolmogorov-Smirnov	D	0.05292	2487	Pr > D	>0.150		
Cramer-von Mises	W-Sq	0.06678	3003	Pr > W-Sq	>0.250		
Anderson-Darling	A-Sq	0.44310	1126	Pr > A-Sq	>0.250		

The statistical tests for normality, including Kolmogorov-Smirnov, Cramer-von Mises, and Anderson-Darling, indicate that while some weeks align with the assumption of normally distributed residuals (p > 0.05 for Week 4), others suggest deviations from normality (p < 0.05 for Week 6). This raises concerns about the validity of parametric methods used and their assumptions.

				n for V	Veek_6	
	Parame	ters for	Normal [Distrib	oution	
	Parame	ter S	ymbol Iu	Est	imate	
	Mean	N		2	2.204	
	Std Dev	S	igma	7.7	56222	
G	oodness-of	-Fit Test	ts for No	rmal [Distribution	
		Statistic			p Val	ue
Test						0.047
Test Kolmogoro	v-Smirnov	D	0.0989	2606	Pr > D	0.017
		D W-Sq	0.0989		Pr > D Pr > W-Sq	<0.005



The presence of outliers, as indicated in the boxplots, can have a significant effect on the mean and standard deviation. These outliers may represent extreme but valid variations in response to treatment or may be due to data entry errors or other issues.



The histogram of Week 6's lead levels, along with the superimposed normal distribution curve, indicates a slight deviation from the assumed normality. The normal distribution curve, defined by a mean of 22.204 and a standard deviation of 7.7562, appears to be a reasonable fit at first glance. However, the tails of the distribution suggest potential skewness or kurtosis that is not captured by a normal distribution.

Improvement Potential:

For normality, non-parametric methods that do not rely on the assumption of normality could be considered for analysis. These statistical techniques are less sensitive to deviations from normality.

Also, using median and interquartile range as primary descriptive statistics can provide a better understanding of the central tendency and dispersion in the presence of outliers.