Problem solution

Junior risk quant

Автор

Черепахин Иван Иванович icherepaxin@bk.ru

1 Task 1

Введем из условия задачи

$$\pi_t = \mathbb{1}\{t \geqslant \tau\} - \sum_{i=1}^t 2^{i-1} \mathbb{1}\{t < \tau\}. \tag{1}$$

1.1 Prove that $\tau < \infty$ a.s.

Хотим показать, что τ является моментом остановки. Измеримость τ вытекает из того, что это inf от измеримых функций. Далее рассмотрим серии эквивалентных событий $\{\tau=\infty\} \Leftrightarrow \{\pi_\infty=1\}$, то есть мы никогда не останавливаем игру, при этом выйгрыш равен 1. На математическом языке

$$P(\tau = \infty) = P(\pi_{\infty} = 1) = P(\sum_{k=1}^{\infty} 2^{k-1} = 1) = 0.$$

1.2 $\mathbb{E}\pi_{\tau} = ?$

Используем (1) и возьмем матожидание π_{τ} . Соответсвенно получаем $\mathbb{E}\pi_{\tau}=1$.

1.3 $\mathbb{E}\pi_t = ?$

Используем (1) и возьмем матожидание π_t

$$\mathbb{E}\pi_t = P(t \ge \tau) - \sum_{i=1}^t 2^{i-1} P(t < \tau) = 1 - 2^t (1 - P(\tau \le t)).$$

Заметим, что распределние au выглядит следующим образом

$$P(\tau = k) = \frac{1}{2^k}, \quad \forall k \in \mathbb{N},$$

так как из условия на каждом шаге равновероятен выйгрыш и проигрыш. Соответственно получаем $\mathbb{E}\pi_t=0.$

1.4 Compare result

Theorem 1.1 (Дуба). Пусть $X = \{X_t, t \in \mathbb{N}\}$ - мартингал и τ - момент остановки со значениями в \mathbb{N} и фильтрация $\{\mathcal{F}_t\}_{t\in\mathbb{N}}$. Пусть $\exists C$ такая, что

$$|X_{min(\tau,n)}| \leqslant C \ \forall n \geqslant 0 \ \text{п.н.}$$

Тогда выполнено $\mathbb{E}X_{\tau} = \mathbb{E}X_1$.

В нашем же случае матожидания различаются. Соответсвенно этого достаточно, что утверждать, что процесс π_t не является мартингалом.

Number 2

Now lets simulate the task. Futhermore I will estimate $ar{X} = rac{1}{N} \sum_{i=1}^N X_i$ and $X_{(1)}$.

```
import numpy as np
import matplotlib.pyplot as plt
import pandas as pd
import seaborn as sns
import numpy.typing as npt
from typing import Callable
import statsmodels.api as sm
```

Parametrs of samples.

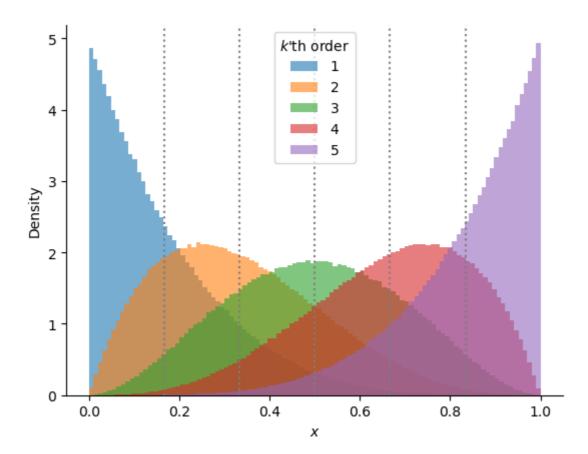
This helpful function is creating plot of sampels.

```
In []:
    def draw(u : np.array):
        __, ax = plt.subplots()
        for k in range(u.shape[0]):
            ax.hist(u[k, :], density=True, alpha=0.6, label=f'${k+1}$', bins=100);
        ax.legend(loc='best', title="$k$'th order")
        sns.despine();
        ax.set_xlabel('$x$')
        ax.set_ylabel('Density');
        for k in range(u.shape[0]):
            ax.axvline(u[k, :].mean(), color='gray', linestyle=':');
        plt.show()
```

2.1

```
In []: R = 1000000 # samples
    N = 5 # amount of variable
    a = 0
    b = 1

In []: x = np.random.uniform(a,b,(N,R))
    x_sort = np.sort(x, 0)
    x_mean_estimate = x.mean(0)
    data = {"x_mean" : list(x_mean_estimate)}
    for i in range(1, N + 1):
        data["x({})".format(i)] = list(x_sort[i - 1])
    data = pd.DataFrame(data)
In []: draw(x_sort)
```



```
In []: estimate = x_sort[0]
  res = sm.OLS(estimate, x_sort.T).fit()
  print(res.summary())
```

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Dep. Variable:		У	R–squa	red (uncent	ered):		1.0		
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Method:	Least Squa	ares	F-stat	istic:	4	.791e+			
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00	wed, of Nov 2	2023	FIUD (I -Statistic		0.			
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07									
Df Model:		5							
Covariance Type:	nonrol 	oust							
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×1 1.00	000 1.29e-18	7.74	 4e+17	0.000	1.000	1.000			

	coef	std err	t	P> t	[0.025	0.975]
x1 x2 x3 x4 x5	1.0000 2.044e-16 3.856e-16 -5.149e-17 -2.739e-17	1.29e-18 1.29e-18 1.29e-18 1.29e-18 8.36e-19	7.74e+17 157.956 298.070 -39.765 -32.764	0.000 0.000 0.000 0.000	1.000 2.02e-16 3.83e-16 -5.4e-17 -2.9e-17	1.000 2.07e-16 3.88e-16 -4.9e-17
Omnibus Prob(Om Skew: Kurtosi	nnibus):	-0.		•	:	0.688 45113.929 0.00 15.7

Notes:

- [1] R^2 is computed without centering (uncentered) since the model does not contain a constant.
- [2] Standard Errors assume that the covariance matrix of the errors is correctly spec ified.

```
In []: estimate = x_mean_estimate
  res = sm.OLS(estimate, x_sort.T).fit()
  print(res.summary())
```

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==								
Dep. Variable:		у	R-squ	uared (unce	entere	:(b		1.0
00								
Model:		0LS	Adj.	R-squared	(uncer	ntered):		1.0
00								
Method:		Least Squares	F-sta	atistic:				4.631e+
36								
Date:	Wed	, 01 Nov 2023	Prob	(F-statist	ic):			0.
00								
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07								
No. Observations:		1000000	AIC:					-7.070e+
07								
Df Residuals:		999995	BIC:					-7.070e+
07								
Df Model:		5						
Covariance Type:		nonrobust						
	coef	std err	t	P> t		 [0 . 025	0.9	75]

	coef	std err	t	P> t	[0.025	0.975]
x1 x2 x3	0.2000 0.2000 0.2000	9.83e-19 9.84e-19 9.84e-19	2.04e+17 2.03e+17 2.03e+17	0.000 0.000 0.000	0.200 0.200 0.200	0.200 0.200 0.200
x4 x5	0.2000 0.2000	9.85e-19 6.36e-19	2.03e+17 3.15e+17	0.000 0.000	0.200 0.200	0.200 0.200
Omnibus: Prob(Omnibus): Skew: Kurtosis:		0.		· · · · ·		1.862 27629.430 0.00 15.7

Notes:

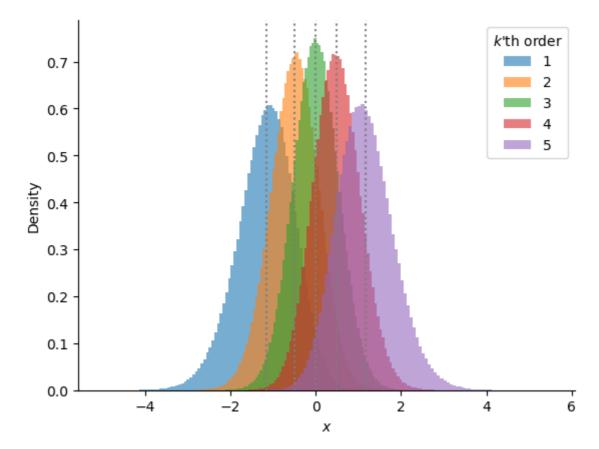
- [1] R^2 is computed without centering (uncentered) since the model does not contain a constant.
- [2] Standard Errors assume that the covariance matrix of the errors is correctly spec ified.

2.2

```
In []: R = 1000000 # samples
    N = 5 # amount of variable
    a = 0
    sigma = 1

In []: x = np.random.normal(a,b,(N,R))
    x_sort = np.sort(x, 0)

In []: draw(x_sort)
```



```
In []: estimate = x_sort[0]
  res = sm.OLS(estimate, x_sort.T).fit()
  print(res.summary())
```

==								
Dep. Variable: 00			У	R-squa	red (uncent	ered):		1.0
Model:			0LS	Adj. F	R-squared (u		1.0	
00 Method:		Least Squa	ares	F-stat	istic:		4	.791e+
35 Date:	We	ed. 01 Nov 2	2023	Prob (F-statistic):		0.
00		·				•	2	
Time: 07		12:41	1:34	Log-L1	kelihood:		3.	5078e+
No. Observations: 07		1000	0000	AIC:			-7	.016e+
Df Residuals:		999	9995	BIC:			-7	.016e+
07 Df Model:			5					
Covariance Type:		nonrob 	oust 					
(oef	std err		t	P> t	[0.025	0.975]	
×1 1.0	000	1.29e-18	7.74	 1e+17	0.000	1.000	1.000	

======	coef	std err	t	P> t	[0.025	0.975]
x1 x2 x3 x4 x5	1.0000 2.044e-16 3.856e-16 -5.149e-17 -2.739e-17	1.29e-18 1.29e-18 1.29e-18 1.29e-18 8.36e-19	7.74e+17 157.956 298.070 -39.765 -32.764	0.000 0.000 0.000 0.000 0.000	1.000 2.02e-16 3.83e-16 -5.4e-17 -2.9e-17	1.000 2.07e-16 3.88e-16 -4.9e-17
Omnibus Prob(Om Skew: Kurtosi	nibus):	-0		•	:	0.688 45113.929 0.00 15.7

Notes:

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```
In []: estimate = x_mean_estimate
  res = sm.OLS(estimate, x_sort.T).fit()
  print(res.summary())
```

```
Dep. Variable:
                                          R-squared (uncentered):
                                                                                      1.0
                                     У
00
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Method:
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36
                     Wed, 01 Nov 2023
                                          Prob (F-statistic):
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Date:
00
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                                                                                 3.5351e+
07
No. Observations:
                                          AIC:
                                                                                 -7.070e+
                               1000000
Df Residuals:
                                999995
                                          BIC:
                                                                                 -7.070e+
07
                                      5
Df Model:
Covariance Type:
                            nonrobust
```

	coef	std err		t	P> t	[0.025	0.975]
 x1	0.2000	9.83e-19	2.04	 e+17	0.000	0.200	0.200
x2	0.2000	9.84e-19	2.03	e+17	0.000	0.200	0.200
x3	0.2000	9.84e-19	2.03	e+17	0.000	0.200	0.200
×4	0.2000	9.85e-19	2.03	e+17	0.000	0.200	0.200
x5	0.2000	6.36e-19	3.15	e+17	0.000	0.200	0.200
Dmnibus:		 22567	 845	 Durbir	 n-Watson:		 1.862
<pre>Prob(Omnibus):</pre>		0	.000	Jarque	e-Bera (JB):		27629.430
Skew:		0	307	Prob(J	B):		0.00
Kurtosis:		3.	534	Cond.	No.		15.7

Notes:

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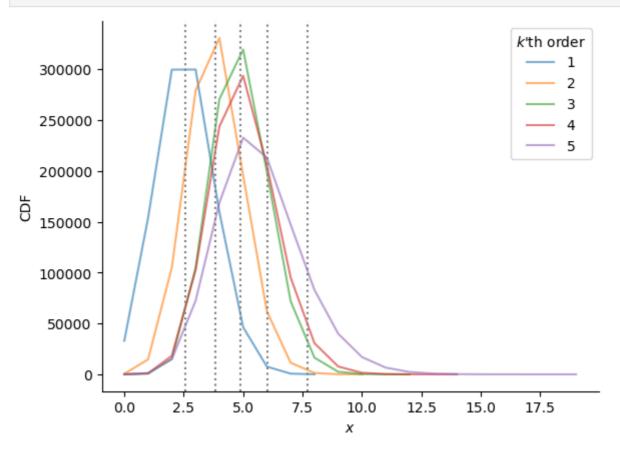
2.3

```
In []:
        def draw_for_pois(u : np.array):
            _, ax = plt.subplots()
            for k in range(u.shape[0]):
                 mass = pd.DataFrame(np.unique(u[k, :], return_counts=True))
                 ax.plot(mass.iloc[1], alpha=0.6, label=f'${k+1}$')
                 # ax.hist(, density=True, alpha=0.6, label=f'${k+1}$', bins=100);
            ax.legend(loc='best', title="$k$'th order")
            sns.despine();
            ax.set_xlabel('$x$')
            ax.set_ylabel('CDF');
            for k in range(u.shape[0]):
                 ax.axvline(u[k, :].mean(), color='gray', linestyle=':');
            plt.show()
In []: R = 1000000 # samples
        N = 5 # amount of variable
         _{\rm lambda} = 5
```

```
In [ ]: x = np.random.poisson(_lambda,(N,R))
x_sort = np.sort(x, 0)
```

```
x_mean_estimate = x.mean(0)
```

```
In [ ]: draw_for_pois(x_sort)
```



```
In []: estimate = x_sort[0]
  res = sm.OLS(estimate, x_sort.T).fit()
  print(res.summary())
```

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== Dep. Variable: 00		y R-sq	R-squared (uncentered):					
Model:	0	LS Adj.	Adj. R-squared (uncentered):					
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Time: 07	12:43:	05 Log-	Likelihood:		3.	5078e+		
No. Observations:	10000	00 AIC:			-7	.016e+		
Df Residuals: 07	9999	95 BIC:			-7	.016e+		
Df Model:		5						
Covariance Type:	nonrobu	st						
CO	ef std err	t	P> t	[0.025	0.975]			
x1 1.00 x2 2.044e-			0.000 0.000	1.000 2.02e-16	1.000 2.07e-16			

	coef	std err	t	P> t	[0.025	0.975]
x1 x2 x3 x4 x5	1.0000 2.044e-16 3.856e-16 -5.149e-17 -2.739e-17	1.29e-18 1.29e-18 1.29e-18 1.29e-18 8.36e-19	7.74e+17 157.956 298.070 -39.765 -32.764	0.000 0.000 0.000 0.000	1.000 2.02e-16 3.83e-16 -5.4e-17 -2.9e-17	1.000 2.07e-16 3.88e-16 -4.9e-17
Omnibus Prob(On Skew: Kurtosi	nnibus):	-0.		•	:	0.688 45113.929 0.00 15.7

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```
In []: estimate = x_mean_estimate
  res = sm.OLS(estimate, x_sort.T).fit()
  print(res.summary())
```

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Method:		Least Squa	ares	F-sta	ntistic:		4	ŀ.631e+
36 Date:	\al.	Wed, 01 Nov 2023			(E statistic	١.		0.
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Df Residuals: 07		999	1995	BIC:			-/	7.070e+
Df Model:			5					
Covariance Typ	e:	nonrob	oust					
=========	======:			:=====				
	coef	std err		t	P> t	[0.025	0.975]	
x1	0.2000	9.83e-19	2.04	le+17	0.000	0.200	0.200	
x2		9.84e-19				0.200	0.200	
x3	0.2000	9.84e-19			0.000	0.200	0.200	
x4			2.03		0.000	0.200	0.200	
x5	0.2000	6.36e-19	3.15	e+17	0.000	0.200 	0.200	
Omnibus:		 22567.	845	Durbi	in-Watson:		1.862	
<pre>Prob(Omnibus):</pre>			000		ue-Bera (JB):		27629.430	
Skew:			307	Prob(•		0.00	
Kurtosis:		3.	534	Cond.	No.		15 . 7	

Notes:

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In []:

Number 3

Derivation

Assume given a filtered probability space $(\Omega, \mathscr{F}, (\mathscr{F}_t)_{t \in [0,T]}, P)$, where the filtration is generated by a Brownian motion W_t .

The market consists of two assets: a *riskless asset* with price B_t (e.g. money market account) and a *risky asset* with price S_t (e.g. FX Spot) given by a geometric Brownian motion:

$$egin{aligned} dB_t &= rB_t dt, & B_0 &= 1, & (B_t = e^{rt}), \ dS_t &= S_t (\mu dt + \sigma dW_t), & S_0 > 0, & \left(S_t &= S_0 e^{\sigma W_t + \left(\mu - rac{\sigma^2}{2}
ight)t}
ight). \end{aligned}$$

Also denote some parametrs:

- 1. r_f foreign risk free rate,
- 2. K the FX strike rate,
- 3. T maturity period.

And the model Garman-Kohlhagen has similar state like Black-Scholes model. The call and put price can be calculated by formula below:

$$V^{
m call} = S e^{-r_f T} \Phi(d_1) - e^{-r T} K \Phi(d_2), \quad V^{
m put} = e^{-r T} K \Phi(-d_2) - S e^{-r_f T} \Phi(-d_1),$$

where $\Phi(x)$ is the standard normal distribution function, and

$$d_1 = rac{1}{\sigma\sqrt{T}}igg(\lnrac{S}{K} + igg(r-r_f + rac{\sigma^2}{2}igg) T igg) \,, d_2 = rac{1}{\sigma\sqrt{T}}igg(\lnrac{S}{K} + igg(r-r_f - rac{\sigma^2}{2}igg) T igg) \,.$$

The greeks

Delta

$$\Delta^{call} = rac{\partial V^{call}}{\partial S} = e^{-r_f T} \Phi(d_1)$$

Gamma

$$\Gamma = rac{\partial^2 V}{\partial S^2} = rac{e^{-r_f T} \Phi'(d_1)}{S \sigma \sqrt{T-t}}$$

Theta

$$heta^{call} = rac{\partial V^{call}}{\partial t} = r_f S e^{-r_f(T-t)} \Phi(d_1) - S e^{-r_f(T-t)} rac{\Phi'(d_1)\sigma}{2\sqrt{T-t}} - rK e^{-r(T-t)} \Phi(d_2).$$

```
In []: from typing import Union
   from dataclasses import dataclass
   import numpy as np
   import numpy.typing as npt
   from scipy import stats
```

```
In []: FloatArray = npt.NDArray[np.float_]
Floats = Union[float, FloatArray]
```

Here we make support class for future application. Names of classes is meaning what it names)

```
In [ ]:
        @dataclass
        class MarketState:
            spot_price: Floats
            domestic_rate: Floats
            time: Floats = 0
        @dataclass
        class StockOption:
            strike_price: Floats
            expiration_time: Floats # in years
            is_call: Union[bool, npt.NDArray[np.bool_]]
            def payoff(self, spot_price: Floats) -> Floats:
                 call_payoff = np.maximum(0, spot_price - self.strike_price)
                put_payoff = np.maximum(0, self.strike_price - spot_price)
                 return np.where(self.is_call, call_payoff, put_payoff)
        @dataclass
        class CallStockOption(StockOption):
            def __init__(self, strike_price, expiration_time):
                 super().__init__(strike_price, expiration_time, True)
        @dataclass
        class PutStockOption(StockOption):
            def __init__(self, strike_price, expiration_time):
                super().__init__(strike_price, expiration_time, False)
        @dataclass
        class GKParams:
            volatility: Floats
            foreign_rate: Floats
```

This block is realisation of d_1 and d_2 by formulas under. And \mbox{dt} function is helpful function for correct pricing becouse we can divide one zero at the formula.

This block is realisation of price of put/call options by formulas under.

This block is realisation of greeks by formulas under.

Now we can launch our model and, at the first, let's set some parametrs.

```
In []: spot = 1.0581
    strikes = np.array((0.9*spot, 1.1*spot))
    matur = 1
    dom_rate=2.7
    for_rate=3
    vol=6

    calls = CallStockOption(strike_price=strikes, expiration_time=matur)

    ms = MarketState(spot_price=spot, domestic_rate=dom_rate)

    params = GKParams(volatility=vol, foreign_rate=for_rate)

    print("Price of options: ", price(calls, ms, params))
    print("Delta : ", delta(calls, ms, params))
    print("Gamma : ", gamma(calls, ms, params))

Price of options: [0.05252301 0.05250671]
```

Now let describe p.2. We need to solve system of equation:

Delta: [0.04971234 0.0497038]

Gamma: [3.82877988e-05 4.22592126e-05]

```
\begin{cases} \Delta_1^{call} N_1^* - \Delta_2^{call} N_2^* - N_3^* = 0 \\ \Gamma_1 N_1^* - \Gamma_2 N_2^* = 0 \\ N_1^* + N_2^* + N_3^* = N \end{cases}
```

Propotions: [0.52336782 0.47418304 0.00244914]

Let's make p.3.

```
In []: budget = 100000
    price_portf = np.append(price(calls, ms, params), spot)
    theta_of_options = np.append(theta(calls, ms, params), 0)

amount_in_portfolio = budget / np.sum(price_portf*propotion_of_portf)
    propotion_of_portf *= amount_in_portfolio
    propotion_of_portf[1] = -propotion_of_portf[1]

theta_portf = np.dot(theta_of_options,propotion_of_portf)

print("Theta of portfolio: ", theta_portf)
```

Theta of portfolio: 14138.894547928823

At the end, we make fix many of paramaters in the model and this way isn't right, althought volatility is changable. In real life we can get implied vol from model and market data. Also for FX option there are another model (SABR, local vol) for more suffitient result.