Linear Regression (part 1)

Predictive Modeling & Statistical Learning

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Linear Regression

Advertising Data from ISL

```
# file in folder data/ of github repo
Advertising <- read.csv("data/Advertising.csv", row.names = 1)</pre>
```

```
TV
          Radio
                 Newspaper
                            Sales
   230.1
           37.8
                     69.2
                           22.1
    44.5 39.3
                     45.1 10.4
3
    17.2
           45.9
                     69.3
                            9.3
4
   151.5
         41.3
                     58.5
                            18.5
5
   180.8
         10.8
                     58.4 12.9
6
                            7.2
   8.7
           48.9
                     75.0
    57.5
           32.8
                     23.5
                            11.8
   120.2
           19.6
                     11.6
                             13.2
```

(first 8 rows)

Advertising Data from ISL

Advertising consists of:

- ▶ the Sales of a product in 200 different markets
- the advertising budgets for three different media:
 - TV
 - Radio
 - Newspaper
- It is not possible to directly increase the sales of the product
- On the other hand, it is possible to control the advertising expenditure in each of the 3 media

Introduction

- Suppose we observe a quantitative response Y and p different predictors, X_1, X_2, \ldots, X_p
- We assume there is some relationship between Y and $[X_1, \ldots, X_p]$. that can be written in a general form as

$$Y = f(X_1, X_2, \dots, X_p) + \epsilon$$

- lacktriangleright f represents the systematic information that the predictors provide about Y
- ϵ represents an ϵ rror term that is a catch-all for what we miss with the model

Data set Advertising

Response:

▶ Y: Sales

Predictors:

► X₁: TV

 $ightharpoonup X_2$: Radio

▶ X₃: Newspaper

Relationship:

Sales =
$$f(TV, Radio, Newspaper) + \epsilon$$

Introduction

One possibility for f() is a linear relationship of the form:

$$Y = \beta_0 + \beta_1 X_1 + \dots + \beta_p X_p + \varepsilon$$

Introduction

One possibility for f() is a linear relationship of the form:

$$Y = \beta_0 + \beta_1 X_1 + \dots + \beta_p X_p + \varepsilon$$

- ▶ It assumes a linear dependence of Y on the predictors
- $\triangleright \beta_0, \beta_1, \dots, \beta_p$ are unknown constants also known as the model *coefficients* or *parameters*
- ▶ The linearity is in the parameters (i.e. coefficients)

Linear relationship

$$Y = \beta_0 + \beta_1 X_1 + \dots + \beta_p X_p + \epsilon$$

Sales =
$$\beta_0 + \beta_1 \text{ TV} + \beta_2 \text{ Radio} + \beta_3 \text{ Newspaper} + \epsilon$$

Examples of linear models

A couple of examples of other possible linear models

$$Y = \beta_0 + \beta_1 X_1 + \beta_2 log(X_2) + \beta_3 X_1 X_2 + \varepsilon$$

$$Y = \beta_0 + \beta_1 X_1 + \beta_2 X_1^2 + \beta_3 (X_1^{X_2}) + \varepsilon$$

Non-linear models

Some models are **not linear** in the parameters:

$$Y = \beta_0 + \beta_1 X_1^{\beta_2} + \varepsilon$$

Some relationships can be transformed to linearity, for example:

$$Y = \beta_0 X_1^{\beta_1} \varepsilon$$

can be linearized by taking logs (and reexpressing some of the parameters)

$$log(Y) = log(\beta_0) + \beta_1 log(X_1) + log(\varepsilon)$$

Introduction

The challenge involves finding parameter estimates denoted by

$$\hat{\beta}_0, \hat{\beta}_1, \dots, \hat{\beta}_p$$

that provide the "best" approximation for Y:

$$Y \approx \hat{\beta_0} + \hat{\beta_1} X_1 + \dots + \hat{\beta_p} X_p$$

or more commonly

$$\hat{Y} = \hat{\beta}_0 + \hat{\beta}_1 X_1 + \dots + \hat{\beta}_p X_p$$

Introduction

- ▶ Linearity is a BIG assumption.
- ► True regression functions are rarely linear.
- ► Although it may seem overly simplistic, linear regression is extremely useful both conceptually and practically.

Simple Linear Regression

- ► Simple Linear Regression = Univariate regression
- ightharpoonup One predictor variable X and one response variable Y

$$Y = \beta_0 + \beta_1 X + \varepsilon$$

We assume a linear model

$$Y = \beta_0 + \beta_1 X + \varepsilon$$

where:

- ▶ β_0 and β_1 are two unknown constants also known as coefficients or parameters
- \triangleright β_0 represents the *intercept*
- β_1 represents the *slope*
- \triangleright ε is a vector of error terms

In vector notation:

$$\mathbf{y} = \beta_0 + \beta_1 \mathbf{x} + \boldsymbol{\varepsilon}$$

where:

- ▶ y is the vector representing the response variable
- ▶ x is the vector representing the predictor variable
- \triangleright ε is the vector representing the error term

Some vector-matrix notation

In matrix notation:

$$\mathbf{y}_{n \times 1} = \mathbf{X}_{n \times 2} \times \boldsymbol{\beta}_{2 \times 1} + \boldsymbol{\varepsilon}_{n \times 1}$$

which can be represented by:

$$\begin{bmatrix} y_1 \\ y_2 \\ y_3 \\ \vdots \\ y_n \end{bmatrix} = \begin{bmatrix} 1 & x_1 \\ 1 & x_2 \\ 1 & x_3 \\ \vdots & \vdots \\ 1 & x_n \end{bmatrix} \begin{bmatrix} \beta_0 \\ \beta_1 \end{bmatrix} + \begin{bmatrix} \varepsilon_1 \\ \varepsilon_2 \\ \varepsilon_3 \\ \vdots \\ \varepsilon_n \end{bmatrix}$$

Note that if the data is center (mean = 0)

$$\mathbf{y} = \beta_1 \mathbf{x} + \boldsymbol{\varepsilon}$$

then there is no intercept term β_0

Some vector-matrix notation

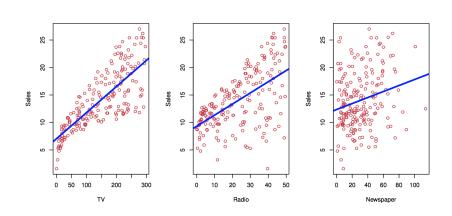
With centered data we have:

$$\mathbf{y}_{n\times 1} = \mathbf{x}_{n\times 1} \times \beta_1 + \mathbf{\varepsilon}_{n\times 1}$$

which can be represented by:

$$\begin{bmatrix} y_1 \\ y_2 \\ y_3 \\ \vdots \\ y_n \end{bmatrix} = \begin{bmatrix} x_1 \\ x_2 \\ x_3 \\ \vdots \\ x_n \end{bmatrix} \begin{bmatrix} \beta_1 \end{bmatrix} + \begin{bmatrix} \varepsilon_1 \\ \varepsilon_2 \\ \varepsilon_3 \\ \vdots \\ \varepsilon_n \end{bmatrix}$$

Various simple regressions



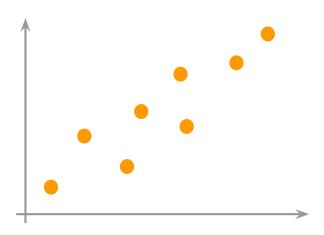
Assuming the model

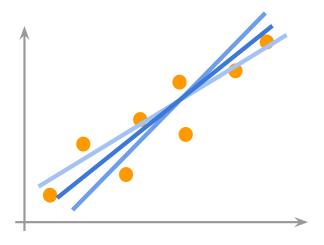
$$\mathbf{y} = \beta_0 + \beta_1 \mathbf{x} + \boldsymbol{\epsilon}$$

and given some estimates $\hat{\beta}_0$ and $\hat{\beta}_1$ for the model coefficients, we predict future sales using

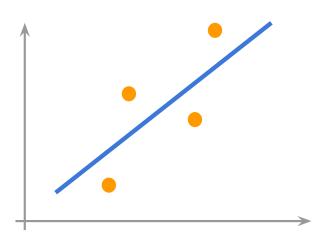
$$\mathbf{\hat{y}} = \hat{\beta_0} + \hat{\beta_1} \mathbf{x}$$

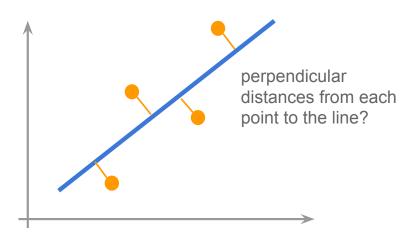
where $\hat{\mathbf{y}}$ indicates the predicted \mathbf{y}

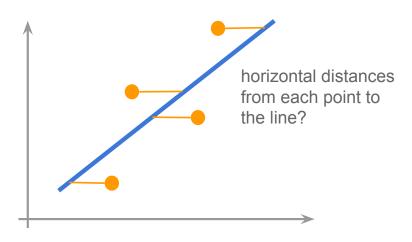


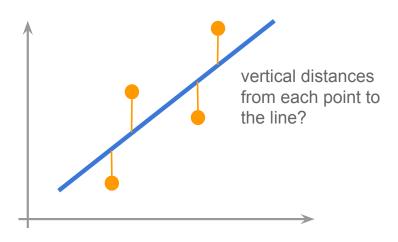


How to find the "best" fitting line?

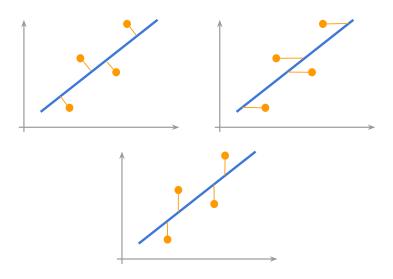


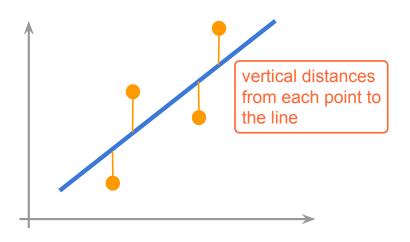






Which Criterion?





Estimation of Parameters

Estimation of the parameters

▶ Let $\hat{y}_i = \hat{\beta}_0 + \hat{\beta}_1 x_i$ be the prediction for \mathbf{y} based on the ith value of \mathbf{x}

Estimation of the parameters

- Let $\hat{y}_i = \hat{\beta}_0 + \hat{\beta}_1 x_i$ be the prediction for y based on the ith value of \mathbf{x}
- ▶ Then $e_i = y_i \hat{y}_i$ represents the *i*th residual

Estimation of the parameters

- ▶ Let $\hat{y}_i = \hat{\beta}_0 + \hat{\beta}_1 x_i$ be the prediction for y based on the *i*th value of x
- ▶ Then $e_i = y_i \hat{y}_i$ represents the *i*th residual
- ▶ We define the Residual Sum of Squares (RSS) as

$$RSS = e_1^2 + e_2^2 + \dots + e_n^2$$

► The **Least Squares** approach chooses $\hat{\beta}_0$ and $\hat{\beta}_1$ to minimize the RSS

Estimation of the parameters

The starting point is to write the model as:

$$\mathbf{e} = \mathbf{y} - (\beta_0 + \beta_1 \mathbf{x})$$

For convenience we define a quadratic loss function

$$L = \sum_{i=1}^{n} e_i^2 = \sum_{i=1}^{n} (y_i - \beta_0 - \beta_1 x_i)^2$$

To minimize L we take partial derivatives with respect to each of the two parameters

Estimation of the parameters

Taking partial derivatives w.r.t each of the two parameters:

$$\frac{\partial L}{\partial \beta_0} = 2 \sum_{i=1}^{n} (y_i - \beta_0 - \beta_1 x_i)(-1) = 0$$

and

$$\frac{\partial L}{\partial \beta_1} = 2\sum_{i=1}^n (y_i - \beta_0 - \beta_1 x_i)(-x_i) = 0$$

Estimation of the parameters

The solutions for β_0 and β_1 would be ontained by solving the so-called *normal equations*

$$\sum_{i=1}^{n} (y_i - \beta_0 - \beta_1 x_i) = 0$$

and

$$\sum_{i=1}^{n} (x_i y_i - x_i \beta_0 - \beta_1 x_i^2) = 0$$

Estimation of the parameters by OLS

The **Least Squares** coefficients are:

$$\hat{\beta}_1 = \frac{\sum_{i=1}^n (x_i - \bar{x})(y_i - \bar{y})}{\sum_{i=1}^n (x_i - \bar{x})^2}$$

$$\hat{\beta_0} = \bar{y} - \hat{\beta_1}\bar{x}$$

where:

$$\bar{y} = \frac{1}{n} \sum_{i=1}^{n} y_i$$
 and $\bar{x} = \frac{1}{n} \sum_{i=1}^{n} x_i$

Estimation of the parameters by OLS

Notice that:

$$\hat{\beta}_1 = \frac{\sum_{i=1}^n (x_i - \bar{x})(y_i - \bar{y})}{\sum_{i=1}^n (x_i - \bar{x})^2}$$

is equivalent to:

$$\hat{\beta}_1 = \frac{cov(\mathbf{x}, \mathbf{y})}{var(\mathbf{x})}$$

Example: Advertising Data

```
# number of observations
n <- nrow(Advertising)

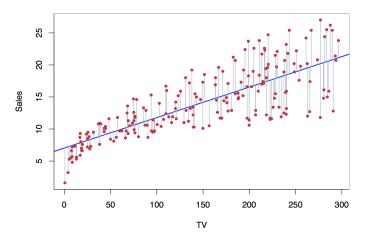
# model matrix
x <- Advertising$TV

# reponse variable
y <- Advertising$Sales</pre>
```

Example: Advertising Data

```
# slope
b1 \leftarrow cov(x, y) / var(x)
b1
## [1] 0.04753664
# intercept
b0 \leftarrow mean(y) - b1 * mean(x)
b0
## [1] 7.032594
```

Example: Advertising Data



The least squares fit for the regression of Sales on TV. In this case a linear fit captures the essence of the relationship, although it is somewhat deficient in the left of the plot.

Another perspective

Projection

Notice that:

$$\hat{\beta}_1 = \frac{\sum_{i=1}^n (x_i - \bar{x})(y_i - \bar{y})}{\sum_{i=1}^n (x_i - \bar{x})^2}$$

Can be expressed in vector notation as:

$$\hat{\beta_1} = \frac{\mathbf{y}^\mathsf{T} \mathbf{x}}{\mathbf{x}^\mathsf{T} \mathbf{x}}$$

with x and y mean-centered.

Projection

Thus, with centered variables x and y, the fitted values \hat{y} are given by:

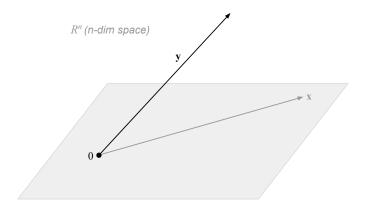
$$\hat{\mathbf{y}} = \hat{\beta}_1 \mathbf{x}$$

$$= \left(\frac{\mathbf{y}^\mathsf{T} \mathbf{x}}{\mathbf{x}^\mathsf{T} \mathbf{x}}\right) \mathbf{x}$$

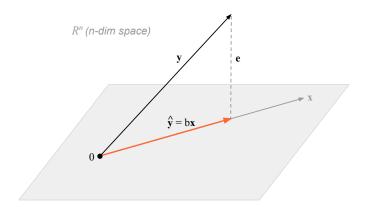
$$= \mathbf{x} \left(\frac{\mathbf{y}^\mathsf{T} \mathbf{x}}{\mathbf{x}^\mathsf{T} \mathbf{x}}\right)$$

$$= \mathbf{x} (\mathbf{x}^\mathsf{T} \mathbf{x})^{-1} \mathbf{x}^\mathsf{T} \mathbf{y}$$

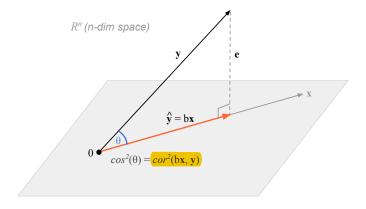
From variables perspective



From variables perspective



From variables perspective



Some Remarks

- There is nothing in the Least Squares method that requires statistical inference: formal tests of null hypotheses or confidence intervals.
- ► In its <u>simplest</u> form, regression analysis can be performed without statistical inference.
- ► The inferential part can sometimes be very useful but goes beyond the definition of a regression analysis.

Some Comments

- ► Linear Regression is a "simple" approach to supervised learning.
- ▶ Don't get fooled by the word "simple".
- "simple" \neq easy / boring / uninteresting.
- ▶ I will use the terms *Regression Analysis* and *Regression Model* interchangeably.

References

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- ► Modern Multivariate Statistical Techniques by Julian Izenman (2008). Springer.
- ▶ Data Mining and Statistics for Decision Making by Stephane Tuffery (2011). Chapter 11: Classification and prediction methods. Wiley.

References (French Literature)

- ▶ Probabilites, analyse des donnees et statistique by Gilbert Saporta (2011). Chapter 17: La regression multiple et le modele lineaire general. Editions Technip, Paris.
- Statistique: Methodes pour decrire, expliquer et prevoir by Michel Tenenhaus (2008). Chapter 5: La Regression Multiple. Dunod, Paris.
- Regression avec R by Cornillon and Matzner-Lober (2011).
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- ➤ Statistique Exploratoire Multidimensionnelle by Lebart et al (2004). Chapter 3, section 3.2: Regression multiple, modele lineaire. Dunod, Paris.
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