STAT 151A HW3 Solutions for #2 and #3 $_{Billy\ Fang}$

 $\mathbf{2}$

Recall the definitions

$$r_i := \frac{\hat{e}_i}{\sqrt{\text{RSS}/(n-p-1)}\sqrt{1-h_i}}, \qquad t_i := \frac{\hat{e}_i}{\sqrt{\text{RSS}_{[i]}/(n-p-2)}\sqrt{1-h_i}}$$

Then,

$$\left(\frac{r_i}{t_i}\right)^2 = \frac{\text{RSS}_{[i]}/(n-p-2)}{\text{RSS}/(n-p-1)}
= \frac{n-p-1}{n-p-2} \left(1 - \frac{\hat{e}_i^2/(1-h_i)}{\text{RSS}}\right)
= \frac{1}{n-p-2} \left(n-p-1 - \frac{\hat{e}_i^2}{(\text{RSS}/(n-p-1)) \cdot (1-h_i)}\right)
= \frac{n-p-1-r_i^2}{n-p-2}.$$
RSS_[i] = RSS - $\frac{\hat{e}_i^2}{1-h_i}$

Some rearranging yields the desired equality.

3

```
bodyfat <- read.csv("bodyfat.csv")
n <- dim(bodyfat)[1]
p <- 4
mod <- lm(bodyfat ~ Age + Weight + Height + Thigh, data=bodyfat)
res <- resid(mod)
fit <- fitted(mod)
sigmahat <- summary(mod)$sigma
X <- as.matrix(cbind(1, bodyfat[,c("Age", "Weight", "Height", "Thigh")]))
lev <- hat(X)</pre>
```

Although you were not asked to do this, let us look at a plot of the leverage values.

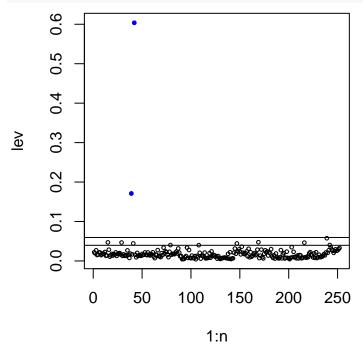
```
plot(1:n, lev, cex=0.5)
hbar <- (p+1)/n
abline(h = 2 * hbar)
abline(h = 3 * hbar)
lev.sorted <- sort(lev, decreasing=T, index.return=T)
n.big <- length(which(lev > 3 * hbar))
idx <- lev.sorted$ix[1:n.big]
idx</pre>
```

[1] 42 39

lev[idx]

```
## [1] 0.6037373 0.1710321
```

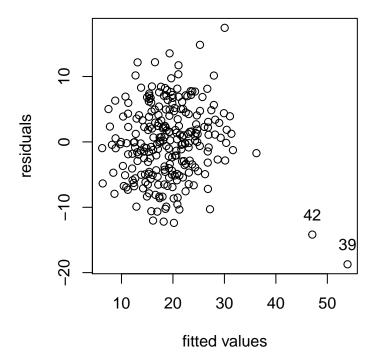
```
points(idx, lev[idx], pch=19, cex=0.5, col='blue')
```



We have two high-leverage points: subjects 42 and 39.

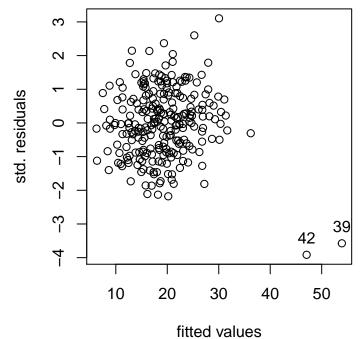
a)

```
plot(fit, res, xlab="fitted values", ylab="residuals")
text(fit[42], res[42]+3, "42")
text(fit[39], res[39]+3, "39")
```



b)

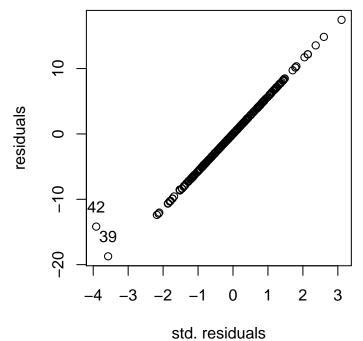
```
res.std <- res / (sigmahat * sqrt(1 - lev))
plot(fit, res.std, xlab="fitted values", ylab="std. residuals")
text(fit[42], res.std[42]+0.5, "42")
text(fit[39], res.std[39]+0.5, "39")</pre>
```



plot(fit, rstandard(mod))

c)

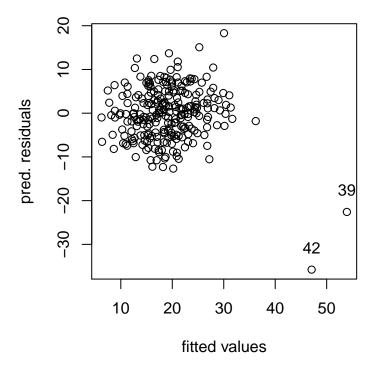
```
plot(res.std, res, xlab="std. residuals", ylab="residuals")
text(res.std[42], res[42]+3, "42")
text(res.std[39], res[39]+3, "39")
```



plot(rstandard(mod), res)

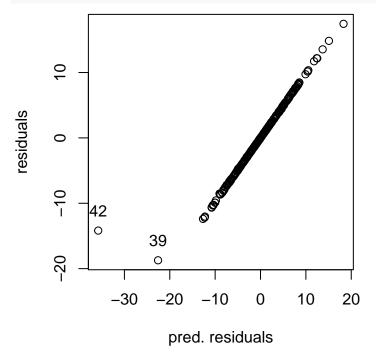
d)

```
res.pred <- res / (1 - lev)
plot(fit, res.pred, xlab="fitted values", ylab="pred. residuals")
text(fit[42], res.pred[42]+5, "42")
text(fit[39], res.pred[39]+5, "39")</pre>
```



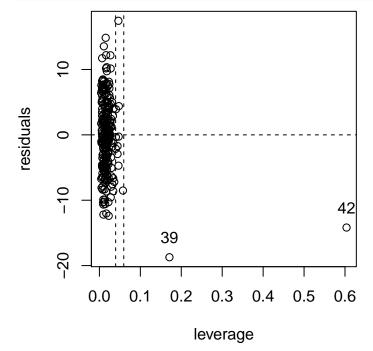
e)

```
plot(res.pred, res, xlab="pred. residuals", ylab="residuals")
text(res.pred[42], res[42]+3, "42")
text(res.pred[39], res[39]+3, "39")
```



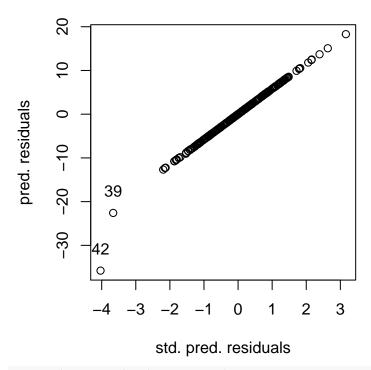
f)

```
plot(lev, res, xlab="leverage", ylab="residuals")
abline(v = 2 * hbar, lty=2)
abline(v = 3 * hbar, lty=2)
text(lev[42], res[42]+3, "42")
text(lev[39], res[39]+3, "39")
abline(h = 0, lty=2)
```



 $\mathbf{g})$

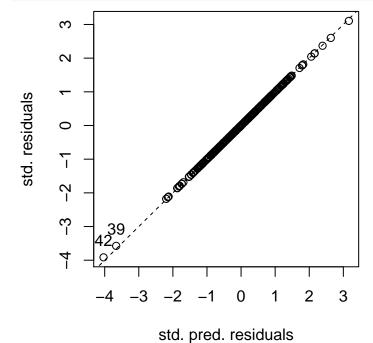
```
res.stdpred <- res.std * sqrt((n - p - 2) / (n - p - 1 - res.std^2))
plot(res.stdpred, res.pred, xlab="std. pred. residuals", ylab="pred. residuals")
text(res.stdpred[42], res.pred[42]+5, "42")
text(res.stdpred[39], res.pred[39]+5, "39")</pre>
```



plot(rstudent(mod), res.pred)

h)

```
plot(res.stdpred, res.std, xlab="std. pred. residuals", ylab="std. residuals")
text(res.stdpred[42], res.std[42]+0.5, "42")
text(res.stdpred[39], res.std[39]+0.5, "39")
abline(a=0, b=1, lty=2)
```

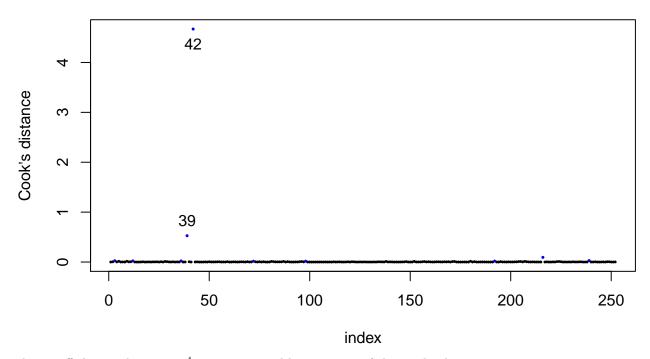


```
# plot(rstudent(mod), rstandard(mod))
```

i)

```
cook <- res.std^2 / (p + 1) * lev / (1 - lev)</pre>
plot(1:n, cook, pch=16, cex=0.4, xlab="index", ylab="Cook's distance",
     main="Cook's distance vs. index (large values in red)")
cook.sorted <- sort(cook, decreasing=T, index.return=T)</pre>
cook.cutoff <- 4 / (n - p - 1) # see page 282
num.big <- length(which(cook > cook.cutoff))
idx <- cook.sorted$ix[1:num.big]</pre>
cook[idx]
##
                       39
## 4.66969177 0.52711418 0.09388593 0.02839649 0.02547257 0.02201873
## 0.02113771 0.01807339 0.01623765 0.01620523
num.big
## [1] 10
points(idx, cook[idx], pch=16, cex=0.4, col='blue')
text(42, cook[42]-0.3, "42")
text(39, cook[39]+0.3, "39")
```

Cook's distance vs. index (large values in red)



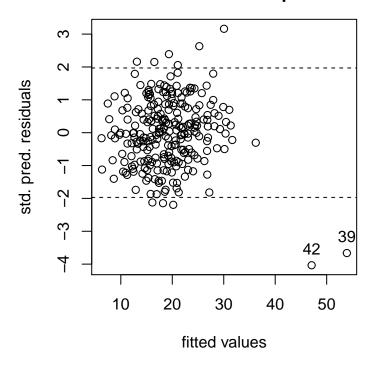
The cutoff chosen above is $\frac{4}{n-p-1}$ as suggested by page 282 of the textbook.

j)

The above plots (a)-(g) suggest that the two high-leverage points 42 and 39 are outliers. They already stand out in the residual plot (a), but standardization (b) makes this even more obvious. The various other plots support this reasoning as well.

We include the following plot to further support this claim.

Std. pred. res. vs. fitted; dashed lines are 5% quantiles



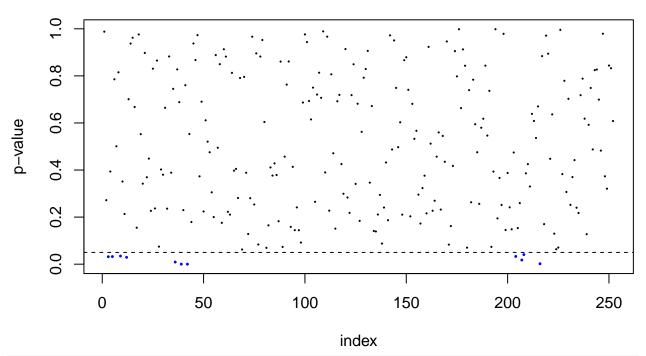
We see that after studentization, these two points have much larger residuals than the other datapoints. The two lines are the quantiles for a 5% t-test, and the eleven points beyond these lines are the eleven blue points in the p-value plot in the next part.

From our work in part i), we see that there are a few influential points. The most influential are the two high-leverage points 42 and 39, followed by 8 other points above the recommended cutoff, but these 8 other points have much smaller influence than the two high-leverage points.

k)

```
abline(h=0.05, lty=2)
#abline(h=0, lty=3)
#abline(h=1, lty=3)
pval.sorted <- sort(pval, index.return=T)
num.small <- length(which(pval < 0.05))
idx <- pval.sorted$ix[1:num.small]
points(idx, pval[idx], pch=16, cex=0.4, col='blue')</pre>
```

P-values vs. index; 0.05 marked by dashed line



pval[idx]

```
## 42 39 216 36 207
## 7.325343e-05 3.054138e-04 1.764147e-03 9.029060e-03 1.760332e-02
## 12 3 5 204 9
## 2.909586e-02 3.169823e-02 3.237542e-02 3.285382e-02 3.461962e-02
## 208
## 4.062107e-02
num.small
```

[1] 11

It does not make sense to rule all of these high p-values as outliers at a 5% level since these are separate t-tests, and some of them may be large purely by chance. We would need to do a Bonferroni correction to properly reject the largest residual at a 5% level.

```
num.small <- length(which(pval < 0.05 / n))
idx <- pval.sorted$ix[1:num.small]
pval[idx]</pre>
```

```
## 42
## 7.325343e-05
```

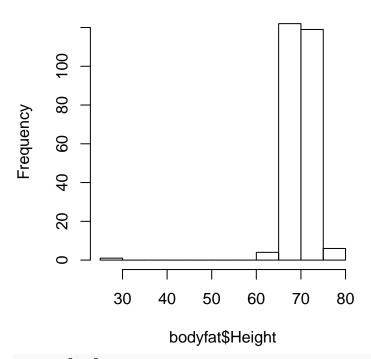
From here, we see that the test only rejects subject 42.

1)

We should be careful when removing data. The analysis above suggests that subjects 42 and 39 have something interesting, but we should look more closely at the data to understand what exactly is going on.

hist(bodyfat\$Height)

Histogram of bodyfat\$Height



bodyfat[42,]

```
##
      Density bodyfat Age Weight Height Neck Chest Abdomen
                                                               Hip Thigh Knee
                              205
                                    29.5 36.6
                                                 106
                                                       104.3 115.5 70.6 42.5
## 42
        1.025
                 32.9
                       44
##
      Ankle Biceps Forearm Wrist
## 42 23.7
              33.6
                       28.7
```

A quick glance at the data shows that subject 42 has an extremely low height of 29.5 inches (2 ft 5.5 in, or 74.93 cm), which is probably the main source of this data point's high leverage. This is likely a recording error, since this subject weights 205 pounds and is 44 years old. Since we do not have a way to correct this error, we should probably remove this point as the above analysis suggests.

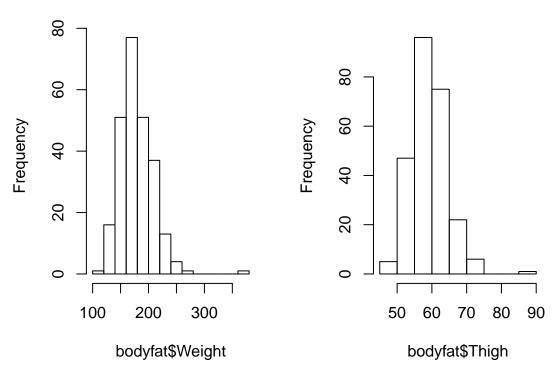
What about subject 39?

```
hist(bodyfat$Weight)
hist(bodyfat$Thigh)
bodyfat[39,]
```

```
## Density bodyfat Age Weight Height Neck Chest Abdomen Hip Thigh Knee
## 39 1.0202 35.2 46 363.15 72.25 51.2 136.2 148.1 147.7 87.3 49.1
## Ankle Biceps Forearm Wrist
## 39 29.6 45 29 21.4
```

Histogram of bodyfat\$Weight

Histogram of bodyfat\$Thigh



We can see that subject 39 has high leverage due to his relatively high weight and thigh measurement that is beyond the range of everyone else in the study; he is over 100 pounds heavier than the next heaviest person in the study. Regarding whether we should remove this datapoint from the dataset or not, one can argue either way: one one hand one might feel that this subject is unusual and our model should only account for males with measurements closer to the rest of the data, while on the other hand one might feel that we have no reason to exclude this subject from the analysis, since it may give some insight into improving our model to account for males with larger measurements. We do not explore this further.

We list the summaries for the original model, the model fitted after removing subject 42, and the model fitted after removing both subjects 42 and 39 below.

summary(mod)

```
##
## Call:
## lm(formula = bodyfat ~ Age + Weight + Height + Thigh, data = bodyfat)
##
## Residuals:
##
       Min
                 1Q
                     Median
                                 3Q
                                         Max
##
  -18.722
            -4.283
                     -0.055
                              4.061
                                     17.449
##
## Coefficients:
##
               Estimate Std. Error t value Pr(>|t|)
   (Intercept) -2.27488
                           11.12642
                                      -0.204
                                               0.8382
##
## Age
                 0.20517
                            0.03274
                                       6.267 1.63e-09 ***
## Weight
                 0.13417
                            0.02952
                                       4.545 8.59e-06 ***
## Height
                -0.49810
                            0.11313
                                      -4.403 1.59e-05 ***
## Thigh
                 0.38970
                            0.16142
                                       2.414
                                               0.0165 *
                   0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
## Signif. codes:
##
```

```
## Residual standard error: 5.753 on 247 degrees of freedom
## Multiple R-squared: 0.5349, Adjusted R-squared: 0.5274
## F-statistic: 71.03 on 4 and 247 DF, p-value: < 2.2e-16
mod2 <- lm(bodyfat ~ Age + Weight + Height + Thigh, data=bodyfat[-42,])</pre>
summary(mod2)
##
## Call:
## lm(formula = bodyfat ~ Age + Weight + Height + Thigh, data = bodyfat[-42,
##
##
## Residuals:
##
       Min
                 1Q
                      Median
                                    3Q
                                            Max
## -22.2729 -3.7828 -0.0947
                               3.9254
                                       13.0096
##
## Coefficients:
##
              Estimate Std. Error t value Pr(>|t|)
                         14.18928
## (Intercept) 34.86048
                                    2.457
                                            0.0147 *
                          0.03284
                                    5.228 3.66e-07 ***
## Age
               0.17168
## Weight
               0.17257
                           0.03019
                                    5.717 3.13e-08 ***
                           0.17072 -6.007 6.77e-09 ***
## Height
              -1.02550
## Thigh
               0.29942
                           0.15824
                                    1.892 0.0596.
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 5.583 on 246 degrees of freedom
## Multiple R-squared: 0.559, Adjusted R-squared: 0.5519
## F-statistic: 77.96 on 4 and 246 DF, p-value: < 2.2e-16
mod3 <- lm(bodyfat ~ Age + Weight + Height + Thigh, data=bodyfat[-c(39,42),])
summary(mod3)
##
## Call:
## lm(formula = bodyfat ~ Age + Weight + Height + Thigh, data = bodyfat[-c(39,
##
       42), ])
##
## Residuals:
       Min
                 1Q
                      Median
                                   ЗQ
## -11.4982 -3.7381 -0.0034
                               3.7581 12.0943
##
## Coefficients:
##
              Estimate Std. Error t value Pr(>|t|)
## (Intercept) 42.82844
                         13.74245
                                    3.117 0.00205 **
## Age
               0.16101
                          0.03164
                                    5.089 7.18e-07 ***
                          0.03020
                                    7.003 2.39e-11 ***
## Weight
               0.21150
              -1.18281
                          0.16753 -7.060 1.70e-11 ***
## Height
## Thigh
               0.24418
                          0.15252
                                    1.601 0.11068
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 5.365 on 245 degrees of freedom
## Multiple R-squared: 0.5883, Adjusted R-squared: 0.5816
## F-statistic: 87.54 on 4 and 245 DF, p-value: < 2.2e-16
```