# intro math

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## 1 Introduction

We use the X below represents data samples.

$$X_{N \times p} = (x_1, x_2, \cdots, x_N)^T, x_i = (x_{i1}, x_{i2}, \cdots, x_{ip})^T$$

As we know, there are two point of views: Frequentists and Bayesians

For Frequentists, they assumed  $\theta$  in  $p(x|\theta)$  is a constant value. The probability of the data samples is  $p(X|\theta) = \prod_{i \neq i}^{N} p(x_i|\theta)$ . Thus, the  $\theta$  should give the maximum probability. In order to get that  $\theta$ , we can compute **MLE**:

$$\theta_{MLE} = \underset{\theta}{argmax} \log p(X|\theta) \underset{iid}{=} \underset{\theta}{argmax} \sum_{i=1}^{N} \log p(x_i|\theta)$$
 (1)

For bayesians, they assumed  $\theta$  subject to a prior distribution. According to the Bayesian Theorem, the posterior probability could be written as:

$$p(\theta|X) = \frac{p(X|\theta) \cdot p(\theta)}{p(X)} = \frac{p(X|\theta) \cdot p(\theta)}{\int_{0}^{\infty} p(X|\theta) \cdot p(\theta) d\theta}$$

In order to find  $\theta$ , we should maximize the posterior. However, there's no need to actually get the value of posterior probability, since p(X) is unrelated with  $\theta$ . So here come to the **MAP**:

$$\theta_{MAP} = \underset{\theta}{\operatorname{argmax}} p(\theta|X) = \underset{\theta}{\operatorname{argmax}} p(X|\theta) \cdot p(\theta)$$
 (2)

### 2 Basic Math

#### 2.1 Gaussian Distribution

#### 2.1.1 MLE in One Dimension

Basically, the probability density function of Gaussian Distribution could be written as:

$$p(x|\mu, \Sigma) = \frac{1}{(2\pi)^{p/2} |\Sigma|^{1/2}} e^{-\frac{1}{2}(x-\mu)^T \Sigma^{-1}(x-\mu)}$$
(3)

Next, we compute its MLE only consider in one dimension situation:

$$\log p(X|\theta) = \sum_{i=1}^{N} \log p(x_i|\theta) = \sum_{i=1}^{N} \log \frac{1}{\sqrt{2\pi}\sigma} \exp(-(x_i - \mu)^2/2\sigma^2)$$

We could see that the parameter has two parts: mean and variance, so we need to compute them respectively. For  $\mu$  we have:

$$\mu_{MLE} = \underset{\mu}{argmax} \log p(X|\theta) = \underset{\mu}{argmax} \sum_{i=1}^{N} (x_i - \mu)^2$$

Let its partial derivatives equals 0, we could get the value of  $\mu$ :

$$\frac{\partial}{\partial \mu} \sum_{i=1}^{N} (x_i - \mu)^2 = 0 \longrightarrow \mu_{MLE} = \frac{1}{N} \sum_{i=1}^{N} x_i \tag{4}$$

And same process for  $\sigma$ :

$$\sigma_{MLE} = \mathop{argmax}_{\sigma} \log p(X|\theta) = \mathop{argmax}_{\sigma} \sum_{i=1}^{N} [-\log \sigma - \frac{1}{2\sigma^2} (x_i - \mu)^2] = \mathop{argmin}_{\sigma} \sum_{i=1}^{N} [\log \sigma + \frac{1}{2\sigma^2} (x_i - \mu)^2]$$

$$\frac{\partial}{\partial \sigma} \sum_{i=1}^{N} [\log \sigma + \frac{1}{2\sigma^2} (x_i - \mu)^2] = 0 \longrightarrow \sigma_{MLE}^2 = \frac{1}{N} \sum_{i=1}^{N} (x_i - \mu)^2$$
 (5)

The  $\mu$  and  $\sigma$  above are estimated value based on input data. We could compute whether there is a bias on them. For  $\mu$ :

$$\mathbb{E}_{\mathcal{D}}[\mu_{MLE}] = \mathbb{E}_{\mathcal{D}}\left[\frac{1}{N}\sum_{i=1}^{N}x_i\right] = \frac{1}{N}\sum_{i=1}^{N}\mathbb{E}_{\mathcal{D}}[x_i] = \mu$$

$$\mathbb{E}_{\mathcal{D}}[\sigma_{MLE}^{2}] = \mathbb{E}_{\mathcal{D}}[\frac{1}{N}\sum_{i=1}^{N}(x_{i}-\mu_{MLE})^{2}] = \mathbb{E}_{\mathcal{D}}[\frac{1}{N}\sum_{i=1}^{N}(x_{i}^{2}-2x_{i}\mu_{MLE}+\mu_{MLE}^{2})]$$

$$= \mathbb{E}_{\mathcal{D}}[\frac{1}{N}\sum_{i=1}^{N}x_{i}^{2}-\mu_{MLE}^{2}] = \mathbb{E}_{\mathcal{D}}[\frac{1}{N}\sum_{i=1}^{N}x_{i}^{2}-\mu^{2}+\mu^{2}-\mu_{MLE}^{2}]$$
For  $\sigma$ :
$$= \mathbb{E}_{\mathcal{D}}[\frac{1}{N}\sum_{i=1}^{N}x_{i}^{2}-\mu^{2}] - \mathbb{E}_{\mathcal{D}}[\mu_{MLE}^{2}-\mu^{2}] = \sigma^{2} - (\mathbb{E}_{\mathcal{D}}[\mu_{MLE}^{2}]-\mu^{2})$$

$$= \sigma^{2} - (\mathbb{E}_{\mathcal{D}}[\mu_{MLE}^{2}] - \mathbb{E}_{\mathcal{D}}^{2}[\mu_{MLE}]) = \sigma^{2} - Var[\mu_{MLE}]$$

$$= \sigma^{2} - Var[\frac{1}{N}\sum_{i=1}^{N}x_{i}] = \sigma^{2} - \frac{1}{N^{2}}\sum_{i=1}^{N}Var[x_{i}] = \frac{N-1}{N}\sigma^{2}$$

We could see that  $\sigma$  has a bias comes from using  $\mu_{MLE}$ 

#### 2.1.2 MLE in Multi-Dimension

Multi-Dimension Gaussian Distribution could be written as:

$$p(x|\mu, \Sigma) = \frac{1}{(2\pi)^{p/2}|\Sigma|^{1/2}} e^{-\frac{1}{2}(x-\mu)^T \Sigma^{-1}(x-\mu)}$$
(6)