# Homework R Markdown Skeleton

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#### Problem 1



## Problem 2

2(a)

For MA(1) model:

$$X_t = Z_t + \theta Z_{t-1}, \quad where \quad Z_t \stackrel{i.i.d}{\sim} N(0, \sigma^2).$$

For autocovariance:

2)=11.48\ \  $\gamma(1)\&=COV(X\_t,X\_t)$ 

$$\label{eq:logical_problem} \left. \left. \left. \left. \left. \right. \right\} \right\} \gamma(0) \right\} = Var(X_t) = Var(Z_t) + \theta Z_t \left\{ t - 1 \right\} \right) = Var(Z_t) + \theta^{2 \operatorname{Var}(Z_t) + \theta^{2 \operatorname{V$$

$$\operatorname{Var}(\overline{X}) = \frac{\tau_n^2}{n}, \quad where$$

$$\tau_n^2 = \gamma(0) + 2\sum_{h=1}^{n-1} (1 - \frac{h}{n})\gamma(h) = \gamma(0) + 2(1 - \frac{1}{n}\gamma(1)) = (1 + \theta^2)\sigma^2 + 2 \cdot \frac{n-1}{n}\theta\sigma^2 = 0.3978947$$

$$Var(\overline{X}) = \frac{\tau_n^2}{n} = 0.004188366$$

2(c)

When  $\theta = 0.8$ , we get  $Var(\overline{X}) = 0.2374958$ .

2(d)

95% confidence interval for  $\mu = E(X_t)$  is:

$$(\overline{X} - 1.96 \frac{\tau_n}{\sqrt{n}}, \overline{X} + 1.96 \frac{\tau_n}{\sqrt{n}})$$

That is, (37.11824, 41.28176).

#### Problem 3

3(a)

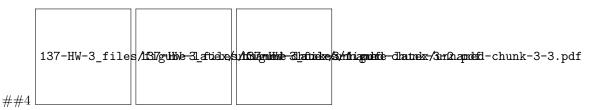
In the case of AR(1) model, i.e.

$$X_t - \mu = \phi(X_{t-1} - \mu) + \epsilon_t$$

For  $\gamma(0)$  and  $\rho(h)$ :

 $\label{eq:local_property} $$ \left( aligned \right) \& \gamma(0) = Var(X_t) = (1 - \phi^2) \{-1\} \sigma^{2 = 19.44444} \\ \& \rho(h) = \phi h \\ \& \tau_n^{2 \approx \gamma(0)[1 + 2 \sum_{h=1}^{2} (\infty) \rho(h)] = \sigma^2/(1 - \phi) \\ \& \tau_n^{2 \approx \gamma(0)[1 + 2 \sum_{h=1}^{2} (\infty) \rho(h)] = \sigma^2/(1 - \phi) \\ \& \tau_n^{2 \approx \gamma(0)[1 + 2 \sum_{h=1}^{2} (\infty) \rho(h)] = \sigma^2/(1 - \phi) \\ \& \tau_n^{2 \approx \gamma(0)[1 + 2 \sum_{h=1}^{2} (\infty) \rho(h)] = \sigma^2/(1 - \phi) \\ \& \tau_n^{2 \approx \gamma(0)[1 + 2 \sum_{h=1}^{2} (\infty) \rho(h)] = \sigma^2/(1 - \phi) \\ \& \tau_n^{2 \approx \gamma(0)[1 + 2 \sum_{h=1}^{2} (\infty) \rho(h)] = \sigma^2/(1 - \phi) \\ \& \tau_n^{2 \approx \gamma(0)[1 + 2 \sum_{h=1}^{2} (\infty) \rho(h)] = \sigma^2/(1 - \phi) \\ \& \tau_n^{2 \approx \gamma(0)[1 + 2 \sum_{h=1}^{2} (\infty) \rho(h)] = \sigma^2/(1 - \phi) \\ \& \tau_n^{2 \approx \gamma(0)[1 + 2 \sum_{h=1}^{2} (\infty) \rho(h)] = \sigma^2/(1 - \phi) \\ \& \tau_n^{2 \approx \gamma(0)[1 + 2 \sum_{h=1}^{2} (\infty) \rho(h)] = \sigma^2/(1 - \phi) \\ \& \tau_n^{2 \approx \gamma(0)[1 + 2 \sum_{h=1}^{2} (\infty) \rho(h)] = \sigma^2/(1 - \phi) \\ \& \tau_n^{2 \approx \gamma(0)[1 + 2 \sum_{h=1}^{2} (\infty) \rho(h)] = \sigma^2/(1 - \phi) \\ \& \tau_n^{2 \approx \gamma(0)[1 + 2 \sum_{h=1}^{2} (\infty) \rho(h)] = \sigma^2/(1 - \phi) \\ \& \tau_n^{2 \approx \gamma(0)[1 + 2 \sum_{h=1}^{2} (\infty) \rho(h)] = \sigma^2/(1 - \phi) \\ \& \tau_n^{2 \approx \gamma(0)[1 + 2 \sum_{h=1}^{2} (\infty) \rho(h)] = \sigma^2/(1 - \phi) \\ \& \tau_n^{2 \approx \gamma(0)[1 + 2 \sum_{h=1}^{2} (\infty) \rho(h)] = \sigma^2/(1 - \phi) \\ \& \tau_n^{2 \approx \gamma(0)[1 + 2 \sum_{h=1}^{2} (\infty) \rho(h)] = \sigma^2/(1 - \phi) \\ \& \tau_n^{2 \approx \gamma(0)[1 + 2 \sum_{h=1}^{2} (\infty) \rho(h)] = \sigma^2/(1 - \phi) \\ \& \tau_n^{2 \approx \gamma(0)[1 + 2 \sum_{h=1}^{2} (\infty) \rho(h)] = \sigma^2/(1 - \phi) \\ \& \tau_n^{2 \approx \gamma(0)[1 + 2 \sum_{h=1}^{2} (\infty) \rho(h)] = \sigma^2/(1 - \phi) \\ \& \tau_n^{2 \approx \gamma(0)[1 + 2 \sum_{h=1}^{2} (\infty) \rho(h)] = \sigma^2/(1 - \phi) \\ \& \tau_n^{2 \approx \gamma(0)[1 + 2 \sum_{h=1}^{2} (\infty) \rho(h)] = \sigma^2/(1 - \phi) \\ \& \tau_n^{2 \approx \gamma(0)[1 + 2 \sum_{h=1}^{2} (\infty) \rho(h)] = \sigma^2/(1 - \phi) \\ \& \tau_n^{2 \approx \gamma(0)[1 + 2 \sum_{h=1}^{2} (\infty) \rho(h)] = \sigma^2/(1 - \phi) \\ \& \tau_n^{2 \approx \gamma(0)[1 + 2 \sum_{h=1}^{2} (\infty) \rho(h)] = \sigma^2/(1 - \phi) \\ \& \tau_n^{2 \approx \gamma(0)[1 + 2 \sum_{h=1}^{2} (\infty) \rho(h)] = \sigma^2/(1 - \phi) \\ \& \tau_n^{2 \approx \gamma(0)[1 + 2 \sum_{h=1}^{2} (\infty) \rho(h)] = \sigma^2/(1 - \phi) \\ \& \tau_n^{2 \approx \gamma(0)[1 + 2 \sum_{h=1}^{2} (\infty) \rho(h)] = \sigma^2/(1 - \phi) \\ \& \tau_n^{2 \approx \gamma(0)[1 + 2 \sum_{h=1}^{2} (\infty) \rho(h)] = \sigma^2/(1 - \phi) \\ \& \tau_n^{2 \approx \gamma(0)[1 + 2 \sum_{h=1}^{2} (\infty) \rho(h)] = \sigma^2/(1 - \phi) \\ \& \tau_n^{2 \approx \gamma(0)[1 + 2 \sum_{h=1}^{2} (\infty) \rho(h)] = \sigma^2/(1 - \phi) \\ \& \tau_n^{2 \approx \gamma(0)[1 + 2 \sum_{h=1}^{2} (\infty) \rho(h)] = \sigma^2/(1 - \phi) \\ \& \tau_n^{2 \approx \gamma$ 

When  $\phi = 0.8$ , we get  $Var(\overline{X}) = 1.842105$ .



## AR1.aic AR2.aic AR3.aic AR4.aic AR5.aic AR6.aic
## 1 234.5559 235.2847 237.2512 239.1987 240.9296 242.8016

```
##
## Call:
## arima(x = data, order = c(1, 0, 0))
##
## Coefficients:
## ar1 intercept
## 0.7113 5.6856
## s.e. 0.0807 0.4407
##
```

##  $sigma^2$  estimated as 1.273: log likelihood = -114.28, aic = 234.56

## ar1 intercept ## 0.7112991 5.6855723

```
137-HW-3_files/MSignible-3Lafticles/Miniagnence-dlattreks/Junthapmeti-chunk-3-5.pdf
##
## Box-Ljung test
##
## data: AR1$residuals
## X-squared = 2.4072, df = 10, p-value = 0.9922
## [1] 3.750000 6.050000 5.208333 3.283333 3.025000 2.925000 5.591667 4.366667
## [9] 4.125000 4.300000 6.841667 5.450000 5.541667 6.691667 5.566667 5.641667
## [17] 5.158333 4.508333 3.791667 3.841667 3.558333 3.491667 4.983333 5.950000
## [25] 5.600000 4.858333 5.641667 8.475000 7.700000 7.050000 6.066667 5.850000
## [33] 7.175000 7.616667 9.708333 9.600000 7.508333 7.191667 7.000000 6.175000
## [41] 5.491667 5.258333 5.616667 6.850000 7.491667 6.908333 6.100000 5.591667
## [49] 5.408333 4.941667 4.500000 4.216667 3.966667 4.741667 5.783333 5.991667
## [57] 5.541667 5.083333 4.608333 4.616667 5.800000 9.283333 9.608333 8.933333
## [65] 8.075000 7.358333 6.158333 5.275000 4.875000 4.358333 3.891667 3.675000
     ARn1.aic ARn2.aic ARn3.aic ARn4.aic ARn5.aic ARn6.aic
##
## 1 213.3339 210.2611 211.9065 213.8531 215.4704 217.1695
##
         ar1
                    ar2 intercept
   0.9934626 -0.2690977 5.6128301
##
## Please cite as:
   Hlavac, Marek (2018). stargazer: Well-Formatted Regression and Summary Statistics Tables.
   R package version 5.2.2. https://CRAN.R-project.org/package=stargazer
##
  _____
##
                        Dependent variable:
##
##
                         data
                                      newdata
##
                          (1)
                                        (2)
##
## ar1
                       0.711***
                                     0.993***
##
                        (0.081)
                                      (0.116)
##
## ar2
                                      -0.269**
##
                                      (0.117)
##
                       5.686***
                                     5.613***
## intercept
```

(0.412)

(0.441)

##

##

```
## Observations
                     74
                                    72
                                 -101.131
## Log Likelihood
                    -114.278
## sigma2
                     1.273
                                   0.957
## Akaike Inf. Crit. 234.556
                                   210.261
## Note:
                   *p<0.1; **p<0.05; ***p<0.01
## [1] 4.185842
## [1] 4.75126
## $pred
## Time Series:
## Start = 73
## End = 73
## Frequency = 1
## [1] 4.15083
##
## $se
## Time Series:
## Start = 73
## End = 73
## Frequency = 1
## [1] 0.9782596
## $pred
## Time Series:
## Start = 73
## End = 74
## Frequency = 1
## [1] 4.150830 4.681853
##
## $se
## Time Series:
## Start = 73
## End = 74
## Frequency = 1
## [1] 0.9782596 1.3789533
```

### Code Appendix

```
knitr::opts_chunk$set(echo = FALSE,warning=FALSE)
library(tidyverse)
library(ggplot2)
#Problem 1(a)
set.seed(123)
simMA1.1a<-arima.sim(n=275,model=list(c(ma=0.7)),sd=1)

plot.ts(simMA1.1a)
acf(simMA1.1a,lag.max = 10)</pre>
```

```
#Problem 1(b)
set.seed(123)
simMA2.1b < -arima.sim(n=275, model=list(ma=c(1.1,0.7)), sd=1)
plot.ts(simMA2.1b)
acf(simMA2.1b,lag=10)
library(readxl)
Unemp1948.2021 <- read_excel("Unemp1948-2021.xls", skip = 10)</pre>
tm < -1:74
plot1 = plot.ts(Unemp1948.2021$UNRATE)
acf(Unemp1948.2021$UNRATE,main = "acf plot")
pacf(Unemp1948.2021$UNRATE,main = " PACF")
data = Unemp1948.2021$UNRATE
AR1 = arima(data, order = c(1,0,0))
AR2 = arima(data, order = c(2,0,0))
AR3 = arima(data, order = c(3,0,0))
AR4 = arima(data, order = c(4,0,0))
AR5 = arima(data, order = c(5,0,0))
AR6 = arima(data, order = c(6,0,0))
aic_table = data.frame(AR1$aic,AR2$aic,AR3$aic,AR4$aic,AR5$aic,AR6$aic)
aic_table
# choose AR 1 because smallest AIC value
AR1
AR1$coef
residual_plot = plot(AR1$residuals) # residuals
acf(AR1$residuals, main = " acf plot of residuals")
test <- Box.test(AR1$residuals, lag = 10, type = "Ljung-Box")</pre>
# We do not rehect the null hypothesis that residuals are i.i.d. or all the aurocorrelations are zero.
newdata = Unemp1948.2021$UNRATE[1:72]
newdata
ARn1 = arima(newdata, order = c(1,0,0))
ARn2 = arima(newdata, order = c(2,0,0))
ARn3 = arima(newdata, order = c(3,0,0))
ARn4 = arima(newdata, order = c(4,0,0))
ARn5 = arima(newdata, order = c(5,0,0))
ARn6 = arima(newdata, order = c(6,0,0))
aic_table2 = data.frame(ARn1$aic,ARn2$aic,ARn3$aic,ARn4$aic,ARn5$aic,ARn6$aic)
aic table2
# we choose AR2 now because smallest aic value
ARn2$coef
u = mean(newdata)
library(stargazer)
stargazer(AR1,ARn2,type = "text")
forecast_2020 = 0.9927 * (newdata[72]-u) - 0.2691 * (newdata[71] - u) + u
forecast 2020
forecast_2021 = 0.9927 * (forecast_2020-u) -0.2691 * (newdata[72] - u) + u
forecast_2021
predict(ARn2,n.ahead = 1)
predict(ARn2,n.ahead = 2)
```