In lecture 16, we looked at precipitation amounts in Madison County (at Morrisville station). We found that the Weibull distribution had a good fit to the monthly precipitation amounts.

We found that the MLEs for the Weibull distribution were

$$\hat{a} = 2.1871$$

 $\hat{\sigma} = 3.9683$

and

$$-\mathcal{L}(\{\hat{a}, \hat{\sigma}\}|\mathbf{x}) = 2166.496$$

is the realized negative log-likelihood. Note this means that the log-likelihood is

$$\mathcal{L}(\{\hat{a}, \hat{\sigma}\}|\mathbf{x}) = -2166.496,$$

and the usual likelihood is

$$L(\{\hat{a}, \hat{\sigma}\}|\mathbf{x}) = e^{[\mathcal{L}(\{\hat{a}, \hat{\sigma}\}|\mathbf{x})]} \approx = e^{-2166.496},$$

which R cannot differentiate from 0.

1. Someone asked "why Weibull?" in class. That is, why wouldn't we use another right-skewed distribution like the Gamma (see Lecture 15), or the Log-Normal (see Lecture 17).

(a) Compute the MLEs for these data using a Gamma distribution.

```
#Function to compute Maximum Likelihood
llgamma <- function(par, data, neg = F) {</pre>
 alpha <- par[1] #get alpha and beta
 beta <- par[2]
  #compute log likelihood
 loglik <- sum(log(dgamma(x = data, shape = alpha, rate = beta)), na.rm = T)</pre>
 return(ifelse(neg, - loglik, loglik))
#Compute MLE
MLE.gamma <- optim(par = c(1,1),
               fn = llgamma.
               data=dat.precip.long$Precipitation,
#extract alpha and beta
alpha.MLE <- MLE.gamma$par[1]
beta.MLE <- MLE.gamma$par[2]
#print the values
alpha.MLE
## [1] 4.174581
beta MLE
## [1] 1.189099
```

We computed the MLEs for these data using a Gamma distribution to obtain $\alpha = 4.1745814$ and $\beta = 1.1890993$.

(b) Compute the MLEs for these data using the Log-Normal distribution.

```
#Function to compute Maximum Likelihood
lllognorm <- function(par, data, neg = F)\{
 mu <- par[1] #get mu and sigma
  sigma <- par[2]
 #compute log likelihood
loglik <- sum(log(dlnorm(x = data, meanlog = mu, sdlog = sigma)), na.rm = T)</pre>
  return(ifelse(neg, - loglik, loglik))
#Compute MLE
MLE.lognorm <- optim(par = c(1,1),</pre>
                fn = lllognorm,
                data=dat.precip.long$Precipitation,
                neg=T)
#extract alpha and beta
mu.MLE <- MLE.lognorm$par[1]</pre>
sigma.MLE <- MLE.lognorm$par[2]
#print the values
mu.MLE
## [1] 1.131261
sigma.MLE
## [1] 0.5333417
```

We computed the MLEs for these data using a Log-Normal distribution to obtain $\mu = 1.1312609$ and $\sigma = 0.5333417$.

(c) Compute the likelihood ratio to compare the Weibull and the Gamma distribution. Which has a better fit according to the likelihood ratio?

$$Q = \frac{L(\{\hat{a}, \hat{\sigma}\} | \mathbf{x})}{L(\{\hat{a}, \hat{\beta}\} | \mathbf{x})} = e^{\left[\mathcal{L}(\{\hat{a}, \hat{\sigma}\} | \mathbf{x}) - \mathcal{L}(\{\hat{a}, \hat{\beta}\} | \mathbf{x})\right]}$$

```
#compute log-likelihood for Gamma distribution
Gamma.loglik <- llgamma(par = c(alpha.MLE, beta.MLE), data = dat.precip.long$Precipitation, neg = T)
#compute the likelihood ratio
Weibull.loglik <- 2166.496
q.gamma.weibull <- exp(Weibull.loglik-Gamma.loglik)
#print Q
q.gamma.weibull
## [1] 4626807</pre>
```

The likelihood ratio of the Weibull and the Gamma distribution is 4.6268066×10^6 . Because Q > 1 (Q is significantly larger than 1), the numerator provides a better fit for the data than the denominator. Weibull distribution is in the numerator, so it is a better fit according to the likelihood ratio.

(d) Compute the likelihood ratio to compare the Weibull and the Log-Normal distribution. Which has a better fit according to the likelihood ratio?

$$Q = \frac{L(\{\hat{a}, \hat{\sigma}\} | \mathbf{x})}{L(\{\hat{\mu}, \hat{\sigma}\} | \mathbf{x})} = e^{[\mathcal{L}(\{\hat{a}, \hat{\sigma}\} | \mathbf{x}) - \mathcal{L}(\{\hat{\mu}, \hat{\sigma}\} | \mathbf{x})]}$$

```
#compute log-likelihood for Log-Normal distribution
Lognorm.loglik <- lllognorm(c(mu.MLE, sigma.MLE), data = dat.precip.long$Precipitation, neg = T)
#compute the likelihood ratio
q.lognorm.weibull <- exp(Weibull.loglik-Lognorm.loglik)
#print Q
q.lognorm.weibull
## [1] 4.218221e-17</pre>
```

The likelihood ratio of the Weibull and the Log-Normal distribution is $4.2182211 \times 10^{-17}$. Because Q < 1, the denominator provides a better fit for the data than the numerator. Log-Normal distribution is in the denominator, so it is a better fit according to the likelihood ratio.

(e) Compute the likelihood ratio to compare the Gamma and the Log-Normal distribution. Which has a better fit according to the likelihood ratio?

$$Q = \frac{L(\{\hat{\alpha}, \hat{\beta}\} | \mathbf{x})}{L(\{\hat{\mu}, \hat{\sigma}\} | \mathbf{x})} = e^{\left[\mathcal{L}(\{\hat{\alpha}, \hat{\beta}\} | \mathbf{x}) - \mathcal{L}(\{\hat{\mu}, \hat{\sigma}\} | \mathbf{x})\right]}$$

```
#compute the likelihood ratio
q.gamma.lognorm <- exp(Gamma.loglik-Lognorm.loglik)
#print Q
q.gamma.lognorm
## [1] 9.116917e-24</pre>
```

The likelihood ratio of the Gamma and the Log-Normal distribution is $9.1169169 \times 10^{-24}$. Because Q < 1, the denominator provides a better fit for the data than the numerator. Log-Normal distribution is in the denominator, so it is a better fit according to the likelihood ratio.

- 2. Optional Coding Challenge. Choose the "best" distribution and refit the model by season.
 - (a) Fit the Distribution for Winter (December-February).
 - (b) Fit the Distribution for Spring (March-May).
 - (c) Fit the Distribution for Summer (June-August).
 - (d) Fit the Distribution for Fall (September-November).
 - (e) Plot the four distributions in one plot using cyan3 for Winter, chartreuse3 for Spring, red3 for Summer, and chocolate3 for Fall. Note any similarities/differences you observe across the seasons.