

Last week: Chapters 15,16,17  
Plus ...

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# This week

- Complete all your assigned work
- Monday:
  - Chapter 15 requirements
  - $S^2$  (one sample t-test) proofs (notice that  $S^2$  has a different meaning in Regression)
- Wednesday, Friday:
  - Chapters 16,17
  - Quality Control
  - Product and System Reliability

# The Exam

- Make sure you review Chapter 9 – we did a lot on this
- Exam will major on chapters 11-14
  - Proof
  - Examples
  - R
  - Book worked examples
- Chapter 15 – see next slides
- Chapters 16,17 will not be examined
- In addition, you learned in MATH 4753  $E(S^2) = \sigma^2$ , now you have the tools to prove  $\frac{(n-1)S^2}{\sigma^2} \sim \chi_{n-1}^2$

# Nonparametric Statistics

## OBJECTIVE

*To present some statistical tests that require fewer, or less stringent, assumptions than the methods of Chapters 8 and 10–14*

## CONTENTS

- 15.1 Introduction: Distribution-Free Tests
- 15.2 Testing for Location of a Single Population
- 15.3 Comparing Two Populations: Independent Random Samples
- 15.4 Comparing Two Populations: Matched-Pairs Design
- 15.5 Comparing Three or More Populations: Completely Randomized Design
- 15.6 Comparing Three or More Populations: Randomized Block Design
- 15.7 Nonparametric Regression

### Sign Test for a Population Median, $\tau$

#### One-Tailed Test

$$H_0: \tau = \tau_0$$

$$H_a: \tau > \tau_0 \text{ [or, } H_a: \tau < \tau_0]$$

Test statistic:

$S$  = Number of sample

observations greater than  $\tau_0$

[or,  $S$  = Number of sample

observations less than  $\tau_0$ ]

#### Two-Tailed Test

$$H_0: \tau = \tau_0$$

$$H_a: \tau \neq \tau_0$$

Test statistic:

$S$  = larger of  $S_1$  and  $S_2$

where

$S_1$  = Number of sample

observations greater than  $\tau_0$

$S_2$  = Number of sample

observations less than  $\tau_0$

[Note: Eliminate observations from the analysis that are exactly equal to the hypothesized median,  $\tau_0$ , and reduce the sample size accordingly.]

Observed significance level:

$$p\text{-value} = P(x \geq S)$$

Observed significance level:

$$p\text{-value} = 2P(x \geq S)$$

where  $x$  has a binomial distribution with parameters  $n$  and  $p = .5$ .

Rejection region: Reject  $H_0$  if  $\alpha > p\text{-value}$ .

Assumption: The sample is randomly selected from a continuous probability distribution. (Note: No assumptions have to be made about the shape of the probability distribution.)

### Wilcoxon Rank Sum Test for a Shift in Population Locations: Independent Random Samples\*

Let  $D_1$  and  $D_2$  represent the relative frequency distributions for populations 1 and 2, respectively.

*One-Tailed Test*

$H_0$ :  $D_1$  and  $D_2$  are identical

$H_a$ :  $D_1$  is shifted to the right of  $D_2$   
(or  $H_a$ :  $D_1$  is shifted to the left of  $D_2$ )

*Two-Tailed Test*

$H_0$ :  $D_1$  and  $D_2$  are identical

$H_a$ :  $D_1$  is shifted either to the left  
or to the right of  $D_2$

Rank the  $n_1 + n_2$  observations in the two samples from the smallest (rank 1) to the largest (rank  $n_1 + n_2$ ). Calculate  $T_1$  and  $T_2$ , the rank sums associated with sample 1 and sample 2, respectively. Then calculate the test statistic.

*Test statistic:*

$T_1$ , if  $n_1 < n_2$ ;  $T_2$ , if  $n_2 < n_1$   
(Either rank sum can be used  
if  $n_1 = n_2$ .)

*Test statistic:*

$T_1$ , if  $n_1 < n_2$ ;  $T_2$ , if  $n_2 < n_1$   
(Either rank sum can be used if  
 $n_1 = n_2$ .) We will denote this rank  
sum as  $T$ .

*Rejection region:*

$T_1$ :  $T_1 \geq T_U$  [or  $T_1 \leq T_L$ ]

$T_2$ :  $T_2 \leq T_L$  [or  $T_2 \geq T_U$ ]

*Rejection region:*

$T \leq T_L$  or  $T \geq T_U$

where  $T_L$  and  $T_U$  are obtained from Table 15, Appendix B.

*Note:* Tied observations are assigned ranks equal to the average of the ranks that would have been assigned to the observations had they not been tied.

### The Wilcoxon Signed Ranks Test: Matched Pairs

Let  $D_1$  and  $D_2$  represent the relative frequency distributions for populations 1 and 2, respectively.

*One-Tailed Test*

$H_0$ :  $D_1$  and  $D_2$  are identical

$H_a$ :  $D_1$  is shifted to the right of  $D_2$   
(or  $H_a$ :  $D_1$  is shifted to the left of  $D_2$ )

*Two-Tailed Test*

$H_0$ :  $D_1$  and  $D_2$  are identical

$H_a$ :  $D_1$  is shifted either to the left  
or to the right of  $D_2$

Calculate the difference within each of the  $n$  matched pairs of observations. Then rank the absolute values of the  $n$  differences from the smallest (rank 1) to the highest (rank  $n$ ) and calculate the rank sum  $T_-$  of the negative differences and the rank sum  $T_+$  of the positive differences.

*Test statistic:*

$T_-$ , the rank sum of the negative differences  
(or  $T_+$ , the rank sum of the positive  
differences)

*Test statistic:*

$T$ , the smaller of  $T_-$  or  $T_+$

*Rejection region:*

$T_- \leq T_0$  (or  $T_+ \leq T_0$ )

*Rejection region:*

$T \leq T_0$

where  $T_0$  is given in Table 16 of Appendix B

(*Note:* Differences equal to 0 are eliminated and the number  $n$  of differences is reduced accordingly. Tied absolute differences receive ranks equal to the average of the ranks they would have received had they not been tied.)

### The Wilcoxon Signed Ranks Test for the Median, $\tau$ , of a Single Population

#### One-Tailed Test

$$H_0: \tau = \tau_0$$

$$H_a: \tau > \tau_0 \text{ [or, } H_a: \tau < \tau_0]$$

Test statistic:

$T_-$ , the negative rank sum

[or,  $T_+$ , the positive rank sum]

[Note: The sample differences are computed as  $(y_i - \tau_0)$ .]

Rejection region:

$$T_- \leq T_0 \text{ [or, } T_+ \leq T_0]$$

where  $T_0$  is found in Table 16 of Appendix B.

#### Two-Tailed Test

$$H_0: \tau = \tau_0$$

$$H_a: \tau \neq \tau_0$$

Test statistic:

$T$ , the smaller of the positive and negative rank sums,  $T_+$  and  $T_-$

Rejection region:

$$T \leq T_0$$

- Assumptions:**
1. A random sample of observations has been selected from the population.
  2. The absolute differences  $y_i - \tau_0$  can be ranked. [No assumptions must be made about the form of the population probability distribution.]
  3. Differences equal to 0 are eliminated and  $n$  is reduced accordingly. Tied differences are assigned ranks equal to the average of the ranks of the tied observations.



## Kruskal–Wallis $H$ Test for Comparing $k$ Population Probability Distributions: Completely Randomized Design

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$H_0$ : The  $k$  population probability distributions are identical

$H_a$ : At least two of the  $k$  population probability distributions differ in location

$$\text{Test statistic: } H = \frac{12}{n(n+1)} \sum_{i=1}^k \frac{T_i^2}{n_i} - 3(n+1)$$

where

$n_i$  = Number of measurements in sample  $i$

$T_i$  = Rank sum for sample  $i$ , where the rank of each measurement is computed according to its relative magnitude in the totality of data for the  $k$  samples

$n$  = Total sample size =  $n_1 + n_2 + \cdots + n_k$

*Rejection region:*  $H > \chi_{\alpha}^2$  with  $(k - 1)$  degrees of freedom

*p-value:*  $P(\chi^2 > H_c)$

*Assumptions:*

1. The  $k$  samples are random and independent.
2. There are 5 or more measurements in each sample.
3. The observations can be ranked.

(Note: No assumptions have to be made about the shape of the population probability distributions.)

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### The Friedman $F_r$ Test for a Randomized Block Design

$H_0$ : The relative frequency distributions for the  $k$  populations are identical

$H_a$ : At least two of the  $k$  populations differ in location (shifted either to the left or to the right of one another)

*Test statistic*: Rank each of the  $k$  observations within each block from the smallest (rank 1) to the largest (rank  $k$ ). Calculate the treatment rank sums,  $T_1, T_2, \dots, T_k$ . Then the test statistic is

$$F_r = \frac{12}{bk(b+1)} \sum T_i^2 - 3b(k+1)$$

where

$b$  = Number of blocks employed in the experiment

$k$  = Number of treatments

$T_i$  = Sum of the ranks for the  $i$ th treatment

*Rejection region*:  $F_r > \chi_\alpha^2$

*p-value*:  $P(\chi^2 > F_r)$

where  $\chi_\alpha^2$  is based on  $(k-1)$  degrees of freedom

*Assumptions*:

1. The  $k$  treatments were randomly assigned to the  $k$  experimental units within each block.
2. For the chi-square approximation to be adequate, either the number  $b$  of blocks or the number  $k$  of treatments should exceed 5.
3. Tied observations are assigned ranks equal to the average of the ranks that would have been assigned to the observations had they not been tied.

### Theil's Test for Zero Slope in the Straight-Line Model $y = \beta_0 + \beta_1 X + \varepsilon$

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*One-Tailed Test*

$$H_0: \beta_1 = 0$$

$$H_a: \beta_1 > 0 \quad (\text{or } H_a: \beta_1 < 0)$$

*Two-Tailed Test*

$$H_0: \beta_1 = 0$$

$$H_a: \beta_1 \neq 0$$

*Test statistic:*  $C = (-1)(\text{Number of negative } y_j - y_i \text{ differences})$   
 $+ (1)(\text{Number of positive } y_j - y_i \text{ differences})$

where  $y_i$  and  $y_j$  are the  $i$ th and  $j$ th observations ranked in increasing order of the  $x$  values,  $i < j$ .

*Observed significance level:*

$$p\text{-value} = \begin{cases} P(x \geq C) & \text{for } H_a: \beta_1 > 0 \\ P(x \leq C) & \text{for } H_a: \beta_1 < 0 \end{cases}$$

*Observed significance level:*

$$p\text{-value} = 2 \min(p_1, p_2)$$

where

$$p_1 = P(x \geq C)$$

$$p_2 = P(x \leq C)$$

where the values of  $P(x \geq C) = P(x \leq -C)$  are given in Table 18 of Appendix B.

*Assumptions:* The random error  $\varepsilon$  is independent.

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## Key Formulas

Test	Test Statistic	Large-Sample Approximation
Sign	$S$ = number of sample measurements greater than (or less than) hypothesized median, $\tau_0$	$Z = \frac{S - .5n}{.5\sqrt{n}}$ 842
Wilcoxon rank sum	$T_1$ = rank sum of sample 1 or $T_2$ = rank sum of sample 2	$Z = \frac{T_1 - \frac{n_1(n_1 + n_2 + 1)}{2}}{\sqrt{\frac{n_1 n_2 (n_1 + n_2 + 1)}{12}}}$ 845, 850
Wilcoxon signed ranks	$T_-$ = negative rank sum or $T_+$ = positive rank sum	$Z = \frac{T_+ - \frac{n(n + 1)}{4}}{\sqrt{\frac{n(n + 1)(2n + 1)}{24}}}$ 854, 855
Kruskal-Wallis	$H = \frac{12}{n(n + 1)} \sum \frac{T_j^2}{n_j} - 3(n + 1)$	860
Friedman	$F_r = \frac{12}{bk(k + 1)} \sum T_j^2 - 3b(k + 1)$	865
Spearman rank correlation (shortcut formula)	$r_s = 1 - \frac{6 \sum d_i^2}{n(n^2 - 1)}$ where $d_i$ = difference in ranks of $i$ th observations for samples 1 and 2	871
Theil's Zero slope	$C = (-1)(\text{number of negative } y_i - y_j \text{ differences})$ $+ (1)(\text{Number of positive } y_i - y_j \text{ differences})$	873

# LANGUAGE LAB

Symbol	Description
$\tau$ (tau)	Population median
$S$	Test statistic for sign test (see Key Formulas)
$T_i$	Sum of ranks of observations in sample $i$
$T_L$	Critical lower Wilcoxon rank sum value
$T_U$	Critical upper Wilcoxon rank sum value
$T_+$	Sum of ranks of positive differences of paired observations
$T_-$	Sum of ranks of negative differences of paired observations
$T_0$	Critical value of Wilcoxon signed ranks test
$H$	Test statistic for Kruskal–Wallis test (see Key Formulas)
$F_r$	Test statistic for Friedman test (see Key Formulas)
$r_s$	Spearman’s rank correlation coefficient (see Key Formulas)
$\rho$ (rho)	Population correlation coefficient
$C$	Test statistic for Theil’s zero slope test (see Key Formulas)

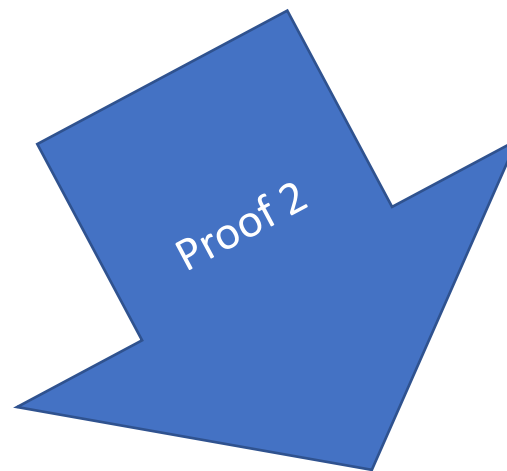
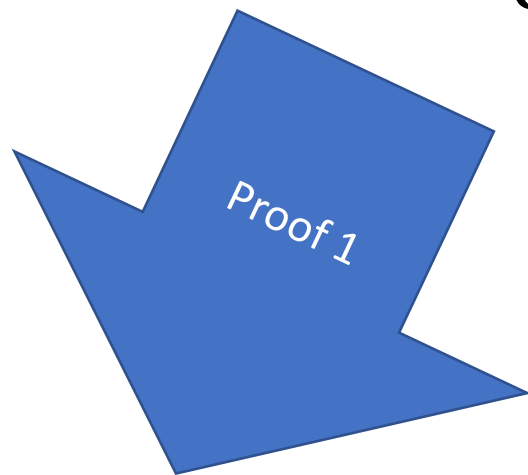
## Chapter Summary Notes

- **Distribution-free tests**—do not rely on assumptions about the probability distribution of the sampled population
- **Nonparametrics**—distribution-free tests that are based on **rank statistics**
- *One-sample* nonparametric test for the population median—**sign test**
- Nonparametric test for *matched pairs*—**Wilcoxon rank test**
- Nonparametric test for a *completely randomized design*—**Kruskal–Wallis test**
- Nonparametric test for a *randomized block design*—**Friedman test**
- Nonparametric test for *rank correlation*—**Spearman's test**
- Nonparametric test for *zero slope*—**Theil's C test**

# To review chapter 15 for the exam

- Go through the above mentioned tests
  - Know what a non-parametric test is
  - Know the name of the test and what parametric test it replaces.
  - Know the NULL for each test.
  - One caveat – You will need to know how the sign test works from a theoretical and more complete perspective

$$\frac{(n-1)S^2}{\sigma^2} \sim \chi_{n-1}^2$$



Identity:

$$\sum_i (x - x_i)^2 = \sum_i x_i^2 - n\bar{x}^2$$

Projection  
theory