#### Course 02402 Introduction to Statistics Lecture 10:

# Oneway Analysis of Variance, ANOVA

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Intro: Small example and TV-data from B&O

#### Oversigt

- Intro: Small example and TV-data from B&O
- Model and hypothesis
- Omputation decomposition and the ANOVA table
- A Hypothesis test (F-test)
- Within-Group variability and the relation to 2-Group t-test
- 6 Post hoc analysis
- Model control
- A complete example from the book

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Intro: Small example and TV-data from B&O

# Oneway ANOVA - example

Group A	Group B	Group C
2.8	5.5	5.8
3.6	6.3	8.3
3.4	6.1	6.9
2.3	5.7	6.1

Is there a difference (in means) between the groups A, B and C?

Analysis of variance (ANOVA) can be used for the analysis if the observations in eacg group can be assumed to be normally distributed.

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# TV set development at Bang & Olufsen

Sound and image quality is measured by th human perceptual instrument:



We developed a tool that is used by B&O to ANOVA (among other things) PanelCheck (Show Panelcheck programme with TV data)

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Intro: Small example and TV-data from B&O

# Oneway ANOVA - example

```
## Input data and plot
## Observations
y \leftarrow c(2.8, 3.6, 3.4, 2.3,
      5.5, 6.3, 6.1, 5.7,
      5.8, 8.3, 6.9, 6.1)
## Groups (treatments)
treatm <- factor(c(1, 1, 1, 1,
                 2, 2, 2, 2,
                 3, 3, 3, 3))
## Plot
par(mfrow=c(1,2))
plot(as.numeric(treatm), y, xlab="Treatment", ylab="y")
plot(treatm, y, xlab="Treatment", ylab="y")
```

#### Bang & Olufsen data in R:

```
# Getting the Bang and Olufsen data from the lmerTest-package:
library(lmerTest) # (Developed by us)
data(TVbo)
head(TVbo)
# Defining the factor identifying the 12 TVset and Picture combs:
TVbo$TVPic <- factor(TVbo$TVset:TVbo$Picture)</pre>
# Each of 8 assessors scored each of 12 combinations 2 times
# Averaging the two replicates for each Assessor and TVpic:
library(doBy)
TVbonoise <- summaryBy(Noise ~ Assessor + TVPic, data = TVbo,
                       keep.names = T)
# One-way ANOVA of the Noise: (Not the correct analysis!!)
anova(lm(Noise ~ TVPic, data = TVbonoise))
# Two-way ANOVA of the Noise: (Much better analysis - next week)
anova(lm(Noise ~ Assessor + TVPic, data = TVbonoise))
```

Model and hypothesis

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## Oneway ANOVA, model

Express the model

$$Y_{ij} = \mu + \alpha_i + \varepsilon_{ij}$$

where it is assumed that

$$\varepsilon_{ij} \sim N(0, \sigma^2)$$

- ullet  $\mu$  is the overall mean
- $\alpha_i$  is the effect of Group (treatment) i
- j indicates the measurements in the groups, from 1 to  $n_i$  in each Group

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Computation - decomposition and the ANOVA table

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# Oneway ANOVA, hypothesis

• We want to compare (more than 2) means  $\mu + \alpha_i$  in the model

$$Y_{ij} = \mu + \alpha_i + \varepsilon_{ij}, \quad \varepsilon_{ij} \sim N(0, \sigma^2)$$

• So we can express the hypothesis:

 $H_0: \quad \alpha_i = 0 \quad \text{for all } i$ 

 $H_1: \quad \alpha_i \neq 0 \quad \text{for at least one } i$ 

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Computation - decomposition and the ANOVA tal

# Oneway ANOVA, decomposition and the ANOVA table

• With the model

$$Y_{ij} = \mu + \alpha_i + \varepsilon_{ii}, \quad \varepsilon_{ii} \sim N(0, \sigma^2)$$

• the total variation in the data can be decomposed:

$$SST = SS(Tr) + SSE$$

- 'Oneway' refers to the fact that there is only one factor in the experiment on k levels
- The method is called <u>analysis</u> of variance, because the testing is carried out by comparing certain variances.

#### Computation - decomposition and the ANOVA table

# Formulas for sums of squares

• Total sum of squares ("the total variance")

$$SST = \sum_{i=1}^{k} \sum_{j=1}^{n_i} (y_{ij} - \bar{y})^2$$

 The sum of squares for the residuals ("residual variance after model fit")

$$SSE = \sum_{i=1}^{k} \sum_{j=1}^{n_i} (y_{ij} - \bar{y}_i)^2$$

• Sum of squares of treatment ("variance explained by the model")

$$SS(Tr) = \sum_{i=1}^{k} n_i (\bar{y}_i - \bar{y})^2$$

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#### Computation - decomposition and the ANOVA table

#### Example

```
## Number of Groups
k <- 3
## Number in each Group
ni <- 10
## Simulate data from model with 3 means
yModel1 \leftarrow rep(c(4, 5, \rightarrow3), each=ni) + rnorm(ni*k, sd=1)
## Simulate data from model with 3 other means
yModel2 \leftarrow rep(c(1, 3, 1), each=ni) + rnorm(ni*k, sd=1)
## 3 Groups
group <- rep(1:k, each=ni)</pre>
## Plot them
par(mfrow=c(1,2))
plot(group, yModel1, ylim=range(yModel1,yModel2))
plot(group, yModel2, ylim=range(yModel1,yModel2))
## Compute SST: total variance, which is highest?
(SST1 <- sum( (yModel1 - mean(yModel1))^2 ))
(SST2 <- sum( (yModel2 - mean(yModel2))^2 ))
## Compute SSE: total residual variation, which is highest?
(SSE1 <- sum(tapply(yModel1, group, function(x) { sum((x - mean(x))^2) })))
(SSE2 <- sum(tapply(yModel2, group, function(x){ sum((x - mean(x))^2) })))
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```

#### The ANOVA table

Source of	Deg. of	Sums of	Mean sum of	Test-	<i>p</i> -
variation	freedom	squares	squares	statistic $F$	value
treatment	k-1	SS(Tr)	$MS(Tr) = \frac{SS(Tr)}{k-1}$	$F_{\rm obs} = \frac{MS(Tr)}{MSE}$	$P(F > F_{\rm obs})$
Residual	n-k	SSE	$MSE = \frac{SSE}{n-k}$		
Total	n-1	SST			

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Hypothesis test (F-test)

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# Oneway ANOVA, F-test

• We have: (Theorem 8.2)

$$SST = SS(Tr) + SSE$$

• and can find the test statistic:

$$F = \frac{SS(Tr)/(k-1)}{SSE/(n-k)}$$

where

- k is the number of levels of the factor
- *n* is the total number of observations
- ullet The significance level lpha is chosen and the test statistic F is computed
- ullet The test statistic is compared with a quantile in the F distribution

$$F \sim F_{\alpha}(k-1,n-k)$$
 (Theorem 8.6)

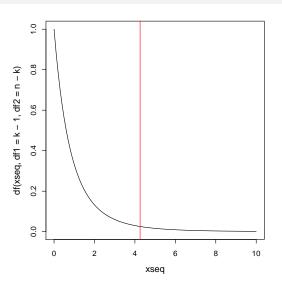
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Hypothesis test (F-test)

#### The F-distribution



#### The F-distribution

```
## Plot the F distribution and see the critical value
## Remember, this is "under HO" (that is we compute as if HO is true):
## Number of Groups
k <- 3
## number of observations
## Sequence for plot
xseq \leftarrow seq(0, 10, by=0.1)
## Plot the density of the F distribution
plot(xseq, df(xseq, df1=k-1, df2=n-k), type="1")
##The critical value for significance level 5 %
cr \leftarrow qf(0.95, df1=k-1, df2=n-k)
## Mark it in the plot
abline(v=cr, col="red")
## The value of the test statistic
(F \leftarrow (SSTr/(k-1)) / (SSE/(n-k)))
## The p-value hence is:
(1 - pf(F, df1=k-1, df2=n-k))
```

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Hypothesis test (F-test)

# The ANOVA table

Source of	Deg. of	Sums of	Mean sum of	Test-	<i>p</i> -
variation	freedom	squares	squares	statistic $F$	value
treatment	k-1	SS(Tr)	$MS(Tr) = \frac{SS(Tr)}{k-1}$	$F_{\rm obs} = \frac{MS(Tr)}{MSE}$	$P(F > F_{\text{obs}})$
Residual	n-k	SSE	$MSE = \frac{SSE}{n-k}$		
Total	n-1	SST			

Within-Group variability and the relation to 2-Group t-test

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Post hoc analysi

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# Within-Group variability and the relation to 2-Group t-test (Theorem 8.4)

The residual sum of squares SSE divided by n-k, also called Residual mean square MSE = SSE/(n-k) is the average within group variability:

$$MSE = \frac{SSE}{n-k} = \frac{(n_1 - 1)s_1^2 + \dots + (n_k - 1)s_k^2}{n-k}$$
 (1)

$$s_i^2 = \frac{1}{n_i - 1} \sum_{i=1}^{n_i} (y_{ij} - \bar{y}_i)^2$$

IF k = 2:(cf. Method 3.51)

For 
$$k = 2$$
:  $MSE = s_p^2 = \frac{(n_1 - 1)s_1^2 + (n_2 - 1)s_2^2}{n - 2}$ 

For 
$$k = 2$$
:  $F_{obs} = t_{obs}^2$ 

where  $t_{\rm obs}$  is the pooled version coming from Methods 3.51 and 3.52.

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Post hoc analys

#### Post hoc confidence interval - Method 8.9

• A single pre-planned confidence interval for the difference between treatment *i* and *j* is found as:

$$\bar{y}_i - \bar{y}_j \pm t_{1-\alpha/2} \sqrt{\frac{SSE}{n-k} \left(\frac{1}{n_i} + \frac{1}{n_j}\right)} \tag{2}$$

where  $t_{1-\alpha/2}$  is based on the t-distribution with n-k degrees of freedom.

- Note the fewer degrees of freedom as more unknowns are estimated in the computationen of  $MSE=SSE/(n-k)=s_p^2$  (i.e. pooled variance estimate)
- If all M = k(k-1)/2 combinations of pairwise confidence intervals are found use the formula M times but each time with  $\alpha_{\mathsf{Bonferroni}} = \alpha/M$ .

# Post hoc pairwise hypothesis test- Method 8.10

• A single pre-planned level  $\alpha$  hypothesis tests:

$$H_0: \mu_i = \mu_j, \ H_1: \mu_i \neq \mu_j$$

is carried out as:

$$t_{\text{obs}} = \frac{\bar{y}_i - \bar{y}_j}{\sqrt{MSE\left(\frac{1}{n_i} + \frac{1}{n_j}\right)}} \tag{3}$$

and:

$$p$$
 – value =  $2P(t > |t_{obs}|)$ 

where the *t*-distribution with n-k degrees of freedom is used.

• HIf all M = k(k-1)/2 combinations of pairwise hypothesis tests are carried out use the approach M times but each time with test level  $\alpha_{\text{Bonferroni}} = \alpha/M$ .

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# Variance homogeneity

Look at box-plot to check whether the variability seems different for the groups

```
## Check assumption of homogeneous variance
## Box plot
plot(treatm,y)
```

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Model control

#### Normal assumption

#### Look at qq-normal plot

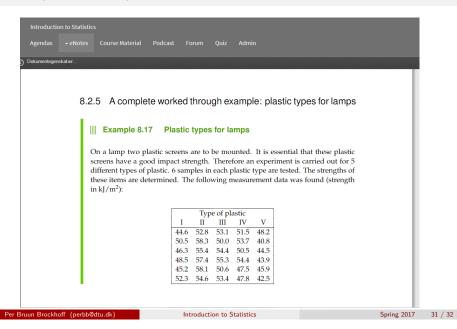
```
#####################################
## Check the assumption of normality of residuals
## qq-normal plot of residuals
fit1 <- lm(y ~ treatm)
qqnorm(fit1$residuals)
qqline(fit1$residuals)
## Or with a Wally plot
require(MESS)
qqwrap <- function(x, y, ...) {qqnorm(y, main="",...);</pre>
qqline(y)}
## Can we see a deviating gg-norm plot?
wallyplot(fit1$residuals, FUN = qqwrap)
```

# Next week: Two-way ANOVA

```
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```

A complete example - from the book

## A complete example - from the book



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