

①

① 3 Doors D_1, D_2, D_3 $P(D_i) = 1/3$ given $i=1,2,3$

* If my friend choose D_1 then the gameshow host will open D_2 or D_3 . In this case prob of the gameshow host opening D_2 or $D_3 \rightarrow$
 $P(H_{D_2 \text{ or } D_3}) = 1/2$

$$P(H_{D_2 \text{ or } D_3} | D_1) = 1/2$$

$$P(H_{D_2 \text{ or } D_3} | D_3) = 0$$

$$P(H_{D_2 \text{ or } D_3} | D_2) = 1$$

$$P(D_1 | H) = \frac{P(H | D_1) P(D_1)}{P(H)} = \frac{1/2 (1/3)}{1/2} = 1/3$$

$$P(D_2 | H) = \frac{P(H | D_2) P(D_2)}{P(H)} = \frac{1 (1/3)}{1/2} = 2/3$$

The posterior prob of switching doors is higher than that if the friend does not switch.

$P(D_2 | H) > P(D_1 | H)$ therefore, the friend should switch doors for better odds!

②

② Bayes Rule: $P(\pi|X) = \frac{p(X|\pi)p(\pi)}{p(X)}$

$$P(\pi|X) \propto p(X|\pi)p(\pi)$$

$$\begin{aligned} \text{likelihood: } p(X|\pi) &= \prod_{n=1}^N p(x_n|\pi) \\ &= \prod_{n=1}^N \prod_{j=1}^K \pi_j^{x_{n,j}} \\ &= \prod_{j=1}^K \pi_j^{\sum_n x_{n,j}} \end{aligned}$$

$$\text{Prior: } p(\pi) = \frac{1}{B(\alpha)} \prod_{j=1}^K \pi_j^{\alpha_j - 1}$$

$$\begin{aligned} P(\pi|X, \alpha) &= p(X|\pi) p(\pi|\alpha) \\ &= \left(\prod_{j=1}^K \pi_j^{\sum_i x_{i,j}} \right) \frac{1}{B(\alpha)} \prod_{j=1}^K \pi_j^{\alpha_j - 1} \\ &= \frac{1}{B(\alpha)} \prod_{j=1}^K \pi_j^{\alpha_j - 1 + \sum_i x_{i,j}} \end{aligned}$$

$$\alpha_j^N = \alpha_j^0 + \sum_i x_{i,j}$$

$$P(\pi|X, \alpha) = \text{Dirichlet}(\alpha_j + \sum_i x_{i,j})$$

[Conjugate prior is Dirichlet]

The most obvious features about the posterior is that it is the prior plus the new data it has seen. Thus it is shifted by the new data. Also the parameters are independent.

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③ a) $X \sim N(\mu, \tau^{-1})$
 $\mu | \tau \sim N(0, a\tau^{-1})$
 $\tau \sim \text{Gamma}(b, c)$

PRIOR = $p(\mu, \tau) \rightarrow p(\mu | \tau) p(\tau)$

likelihood = $p(x | \mu, \tau)$

$p(\mu, \tau | x) \propto p(x | \mu, \tau) p(\mu, \tau)$
 $\propto p(x | \mu, \tau) p(\mu | \tau) p(\tau)$

$N(0, a\tau^{-1}) \text{Ga}(b, c)$
 $\frac{1}{\sqrt{2\pi a\tau}} e^{-\tau/2a\mu^2} \cdot \frac{c^b}{\Gamma(b)} \tau^{b-1} e^{-c\tau}$
 $\sim \text{NG}(0, a^{-1}, b, c)$

$p(\mu, \tau | x) = \prod_{i=1}^N \frac{1}{\sqrt{2\pi\tau}} e^{-\tau/2(x_i - \mu)^2} * \text{NG}(0, a^{-1}, b, c)$

* we know $(x_i - \bar{x})^2 + (\mu - \bar{x})^2 = (x_i - \mu)^2$

$p(x | \mu, \tau) = \left(\frac{2\pi}{\tau}\right)^{-N/2} e^{-\tau/2((x_i - \bar{x})^2 + (\bar{x} - \mu)^2)}$
 $= \left(\frac{2\pi}{\tau}\right)^{-N/2} e^{-\tau/2(Ns + N(\bar{x} - \mu)^2)}$

$p(\mu, \tau | x) = \left(\frac{2\pi}{\tau}\right)^{-N/2} e^{-\tau/2(Ns + N(\bar{x} - \mu)^2)} \frac{1}{\sqrt{2\pi a\tau}} e^{-\tau/2a\mu^2}$

$\cdot \frac{c^b}{\Gamma(b)} \tau^{b-1} e^{-c\tau}$



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$$\begin{aligned}
&\propto \gamma^{b+n/2-1/2} e^{-c\gamma} e^{\gamma/2 \sum_{i=1}^n (x_i - \mu)^2 + a^{-1} \mu^2} \\
&\propto \gamma^{b+n/2-1/2} e^{-c\gamma} e^{\gamma/2 \sum x_i^2 - 2\mu \sum x_i + n\mu^2 + a^{-1} \mu} \\
&\propto \gamma^{b+n/2-1/2} e^{-c\gamma} e^{\gamma/2 \sum x_i^2} e^{\gamma/2 (n+1/a) \mu^2 - 2 \sum x_i \mu} \\
&\propto \gamma^{b+n/2-1/2} e^{-c\gamma} e^{\gamma/2 \sum x_i^2} e^{\gamma/2 (n+1/a) (\mu^2 - \frac{2 \sum x_i}{n+1/a} \mu)} \\
&\propto \gamma^{b+n/2-1/2} e^{-c\gamma} e^{\gamma/2 \sum x_i^2} e^{\gamma/2 (n+1/a) (\mu - \frac{\sum x_i}{n+1/a})^2 + \frac{\sum x_i^2}{n+1/a}} \\
&\propto \gamma^{b+n/2-1/2} e^{-c\gamma} e^{\gamma/2 \sum x_i^2 - \frac{\sum x_i^2}{n+1/a}} e^{\gamma \frac{(n+1/a)}{2} (\mu - \frac{\sum x_i}{n+1/a})^2} \\
&\propto \gamma^{b+n/2-1/2} e^{-\gamma (c + 1/2 \sum (x_i^2) - \frac{\sum x_i^2}{n+1/a})} e^{\gamma \frac{(n+1/a)}{2} (\mu - \frac{\sum x_i}{n+1/a})^2} \\
&= NG\left(\frac{\sum x_i}{n+a^{-1}}, n+a^{-1}, b+n/2 + c + 1/2 \sum (x_i^2) - \frac{\sum x_i^2}{n+a^{-1}}\right)
\end{aligned}$$

Therefore we can see the posterior is also a NG distribution of the form:

$$p(\mu, \gamma | x) = NG(\mu_n, a_n, b_n, c_n)$$

$$\mu_n = \frac{a_0 \mu_0 + N \bar{x}}{a_0 + N}$$

$$a_n = a_0 + N$$

$$b_n = b_0 + N/2$$

$$c_n = c_0 + 1/2 \sum_{i=1}^n (x_i - \bar{x})^2 + \frac{a_0 N (\bar{x} - \mu_0)^2}{2(a_0 + N)}$$

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$$\begin{aligned}
 \textcircled{3} \textcircled{6} \quad p(x^* | x_1, \dots, x_n) &= \int_0^\infty \int_{-\infty}^\infty p(x^* | u, \tau) p(u, \tau | x_1, \dots, x_n) du d\tau \\
 &= \int_0^\infty \int_{-\infty}^\infty p(x^* | u, \tau) p(u | x, \tau) p(\tau | x) du d\tau \\
 &\quad \text{does not depend on } u \\
 &= \int_0^\infty p(\tau | x) \int_{-\infty}^\infty p(x^* | u, \tau) p(u | \tau) du d\tau
 \end{aligned}$$

First part $\int_{-\infty}^\infty p(x^* | u, \tau) p(u | \tau) du d\tau$

$$\begin{aligned}
 &= \frac{1}{2\pi\tau^{-1}} e^{-\tau/2(x^* - u)^2} \frac{1}{2\pi\tau^{-1} \sqrt{N+1/2}} e^{\tau N + \alpha - 1/2} u - \frac{N x}{N+1/2} \tau \\
 &\propto e^{1/2 \tau x^{*2} - 2\tau u x^* + \tau u^2 + \tau(N+\alpha-1)u - 2N\bar{x}\tau u + \frac{N\bar{x}^2}{N+1/2} \tau} \\
 &\propto e^{(1/2 \tau x^{*2} + N\bar{x}^2 / (N+1/2))} e^{\text{other terms}}
 \end{aligned}$$

Not dependent

all other terms

$$\propto e^{1/2 \tau (1+N+\alpha-1) u^2 - 2\tau (x^* + N\bar{x}) u} \times e^{1/2 \tau (1+N+\alpha-1) u - \frac{\tau (x^* + N\bar{x})^2}{\tau(1+N+\alpha-1)}} \cdot \frac{(x^* + N\bar{x})^2}{1+N+\alpha-1}$$

Now integrating $= \int_{-\infty}^\infty 1 = \frac{\sqrt{2\pi}}{\sqrt{\tau(1+N+\alpha-1)}}$

combine w/
other constants

$$= \frac{\tau \sqrt{N+\alpha-1}}{2\pi} \cdot \frac{\sqrt{2\pi}}{\sqrt{\tau(1+N+\alpha-1)}}$$

$$\cdot e^{1/2 \tau x^{*2} + \frac{\tau N\bar{x}^2}{N+1/2} - \frac{\tau (x^* + N\bar{x})^2}{1+N+\alpha-1}}$$

Reassembling $\int_{-\infty}^\infty p(x^* | u, \tau) \cdot p(u | \tau) du \sim N\left(\frac{N\bar{x}}{N+1/2}, \frac{1+N+1/2}{(N+1/2)^2}\right)$



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Now solve integral of NG given definitions
form (3a) \rightarrow

$$\begin{aligned}
 &= \int \text{Gamma}(b, c) \cdot N(\mu, \tau/a) d\tau \\
 &= \int \frac{c^b}{\Gamma(b)} \tau^{b-1} e^{-\tau c} \left(\frac{\tau}{2\pi b}\right)^{1/2} e^{-\tau/2b (x^* - \mu_n)^2} d\tau \\
 &= \frac{c^b}{\Gamma(b) \sqrt{2\pi b}} \int \tau^{b+1/2-1} e^{-\tau(c + 1/2a (x^* - \mu_n)^2)} d\tau
 \end{aligned}$$

This looks like a t -distribution:

$$= \frac{\Gamma\left(\frac{2b+1}{2}\right)}{\Gamma\left(\frac{2b}{2}\right)} (2\pi a c)^{-1/2} \left(1 + \frac{1}{2ac} (x^* - \mu_n)^2\right)^{-2b+1/2}$$

Thus we can the final form is a t -dist
w/ 3 parameters:

$$p(x^* | x) = \pi^{-1/2} \frac{\Gamma(2b+1/2)}{\Gamma(2b/2)} \left(\frac{1}{2bn}\right)^{1/2} \left(1 + \frac{1}{2bn} (x - \mu_n)^2\right)^{-2b+1/2}$$

$\mu_n = \text{center}$

$1/b = \text{precision} = \frac{bn}{c_n(n+1)}$

$df = 2bn$