

Exponential Model

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1 Introduction

The exponential model is one where each arm i is assumed to follow an exponential distribution with hazard λ_i . Depending on the design procedure (the test statistic) there are numerous choices of grid spaces and parametrizations of the natural parameter space. This document will focus on the log-rank statistic with two arms (control and a treatment).

2 Model Assumptions

Assume that there are n patients in each of the $d = 2$ arms with independent $X_{ci} \sim E(\lambda_c)$, $X_{ti} \sim E(\lambda_t)$, $i = 1, \dots, n$. X_c are the samples for the control arm and X_t are for the treatment arm. Then, the distribution of $X \in \mathbb{R}^{n \times 2}$ forms an exponential family with sufficient statistic $T(x) = \left(\sum_{i=1}^n x_{ci}, \sum_{i=1}^n x_{ti} \right)$ natural parameter $\eta = (-\lambda_c, -\lambda_t)$, and log-partition function $A(\eta) := -n \log(\eta_c \eta_t)$.

3 Grid Space

Since the log-rank statistic only depends on the hazard rate $h = \lambda_t/\lambda_c$, it is convenient to parametrize the natural parameter space as a function of (λ_c, h) . Moreover, we will see in Section 4.4 that we get major computation benefits of parametrizing in the log-space $\theta = (\log(\lambda_c), \log(h))$.

This parametrization defines a mapping $\eta(\theta) = (-e^{\theta_1}, -e^{\theta_1+\theta_2})$ from the grid space to the natural parameter space. We conclude this section with the Jacobian and hessian computations needed in the later sections.

$$D_\theta \eta(\theta) = \begin{bmatrix} -e^{\theta_1} & 0 \\ -e^{\theta_1+\theta_2} & -e^{\theta_1+\theta_2} \end{bmatrix} \quad (1)$$

$$\nabla_\theta^2 \eta_1(\theta) = -e^{\theta_1} e_1 e_1^\top \quad (2)$$

$$\nabla_\theta^2 \eta_2(\theta) = -e^{\theta_1+\theta_2} \vec{1} \vec{1}^\top \quad (3)$$

where e_i is the i th standard basis vector and $\vec{1}$ is a vector of ones.

4 Upper Bound

For any model, we must be able to compute the upper bound estimate. The generalized upper bound estimate requires model-specific quantities, which are given by

$$\begin{aligned}
& \text{Gradient Term} : T(x) - \nabla_{\eta} A(\eta) \\
& \eta \text{ transform} : D_{\theta} \eta(\theta) v \\
& \text{Covariance quadratic form} : u^{\top} \text{Var}(T)_{\eta} u \\
& \text{Max covariance quadratic form} : \sup_{\theta \in R} \left[v^{\top} (D_{\eta}(\theta))^{\top} \text{Var}(T)_{\eta(\theta)} (D_{\eta}(\theta)) v \right] \\
& \text{Max covariance and } \eta \text{ hessian} : \|v\|^2 \sum_{k=1}^d \sup_{\theta \in R} \left[\|\nabla^2 \eta_k(\theta)\|_{op} \sqrt{\text{Var}(T_k)_{\eta(\theta)}} \right]
\end{aligned}$$

for any $v, u \in \mathbb{R}^d$ and a bounded subset $R \subseteq \mathbb{R}^d$.

The next few subsections will derive the formulas for each of the quantities above.

4.1 Gradient Term

As shown in Section 2, we have the form for $T(x)$ and $A(\eta)$.

$$\nabla_{\eta} A(\eta) = -n (\eta_c^{-1}, \eta_t^{-1}) = n (\lambda_c^{-1}, \lambda_t^{-1})$$

This gives us

$$T(x) - \nabla_{\eta} A(\eta) = \left(\sum_{i=1}^n x_{ci} - n\lambda_c^{-1}, \sum_{i=1}^n x_{ti} - n\lambda_t^{-1} \right)$$

4.2 η Transform

Using Eq. 1, for any $v \in \mathbb{R}^d$,

$$D_{\theta} \eta(\theta) v = - \begin{bmatrix} e^{\theta_1} & 0 \\ e^{\theta_1+\theta_2} & e^{\theta_1+\theta_2} \end{bmatrix} v = - \begin{bmatrix} e^{\theta_1} v_1 \\ e^{\theta_1+\theta_2}(v_1 + v_2) \end{bmatrix} = - \begin{bmatrix} \lambda_c v_1 \\ \lambda_t(v_1 + v_2) \end{bmatrix}$$

4.3 Covariance Quadratic Form

The covariance of T is given by

$$\text{Var}(T)_{\eta} = n \begin{bmatrix} \eta_c^{-2} & 0 \\ 0 & \eta_t^{-2} \end{bmatrix} = n \begin{bmatrix} \lambda_c^{-2} & 0 \\ 0 & \lambda_t^{-2} \end{bmatrix} \quad (4)$$

and so,

$$u^{\top} \text{Var}(T)_{\eta} u = n(u_1^2 \lambda_c^{-2} + u_2^2 \lambda_t^{-2})$$

4.4 Max Covariance Quadratic Form

Using Eq. 1, 4,

$$\begin{aligned}
D_{\theta} \eta(\theta)^{\top} \text{Var}(T)_{\eta} D_{\theta} \eta(\theta) &= n D_{\theta} \eta(\theta)^{\top} \begin{bmatrix} \eta_c^{-2} & 0 \\ 0 & \eta_t^{-2} \end{bmatrix} \begin{bmatrix} e^{\theta_1} & 0 \\ e^{\theta_1+\theta_2} & e^{\theta_1+\theta_2} \end{bmatrix} \\
&= n D_{\theta} \eta(\theta)^{\top} \begin{bmatrix} e^{-2\theta_1} & 0 \\ 0 & e^{-2(\theta_1+\theta_2)} \end{bmatrix} \begin{bmatrix} e^{\theta_1} & 0 \\ e^{\theta_1+\theta_2} & e^{\theta_1+\theta_2} \end{bmatrix} \\
&= n \begin{bmatrix} e^{\theta_1} & e^{\theta_1+\theta_2} \\ 0 & e^{\theta_1+\theta_2} \end{bmatrix} \begin{bmatrix} e^{-\theta_1} & 0 \\ e^{-(\theta_1+\theta_2)} & e^{-(\theta_1+\theta_2)} \end{bmatrix} \\
&= n \begin{bmatrix} 2 & 1 \\ 1 & 1 \end{bmatrix}
\end{aligned}$$

Note the incredible simplification due to our choice of the η transformation. This gives us

$$\sup_{\theta \in R} \left[v^\top D_\theta \eta(\theta)^\top \text{Var}(T)_\eta D_\theta \eta(\theta) v \right] = n v^\top \begin{bmatrix} 2 & 1 \\ 1 & 1 \end{bmatrix} v$$

4.5 Max Covariance and η Hessian

From Eq. 2, 3,

$$\begin{aligned} \|\nabla^2 \eta_1(\theta)\|_{op} &= e^{\theta_1} \|e_1 e_1^\top\|_{op} = e^{\theta_1} \\ \|\nabla^2 \eta_2(\theta)\|_{op} &= e^{\theta_1 + \theta_2} \|\vec{1} \vec{1}^\top\|_{op} = e^{\theta_1 + \theta_2} d \end{aligned}$$

This gives us

$$\|v\|^2 \sum_{k=1}^d \sup_{\theta \in R} \left[\|\nabla^2 \eta_k(\theta)\|_{op} \sqrt{\text{Var}(T_k)_{\eta(\theta)}} \right] = \|v\|^2 \sqrt{n} (1 + d)$$