Machine Learning Notebook

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Introduction

- 1.1 About this Notebook
- 1.2 Policy of Use

Mathematics Basics

2.1 Probability

2.1.1 Basic Rules

Three Axioms of Probability Let Ω be a sample space. A probability assigns a real number P(X) to each event $X \subseteq \Omega$ in such a way that

- 1. $P(X) \ge 0, \forall X$
- 2. If $X_1, X_2, ...$ are pairwise disjoint events $(X_1 \cap X_2 = \emptyset, i \neq j, i, j = 1, 2, ...)$, then $P(\bigcup_{i=1}^{\infty} X_i) = \sum_{i=1}^{\infty} P(X_i)$. (This property is called countable additivity.)
- 3. $P(\Omega) = 1$

Joint Probability The probability both event A and B occur. $P(X,Y) = P(X \cap Y)$.

Marginalization The probability distribution of any variable in a joint distribution can be recovered by integrating (or summing) over the other variables.

- 1. For continuous r.v. $P(x) = \int P(x,y) dy$; $P(y) = \int P(x,y) dx$.
- 2. For discrete r.v. $P(x) = \sum_{y} P(x, y)$; $P(y) = \sum_{x} P(x, y)$.
- 3. For mixed r.v. $P(x,y) = \sum_{w} \int P(w,x,y,z) dz$, where w is discrete and z is continuous.

Conditional Probability P(X = x | Y = y) is the probability X = x occurs given the knowledge Y = y occurs. Conditional probability can be extracted from joint probability that

$$P(x|y = y^*) = \frac{P(x, y = y^*)}{\int P(x, y = y^*) dx} = \frac{P(x, y = y^*)}{P(y = y^*)}$$

Usually, the formula is written as $P(x|y) = \frac{P(x,y)}{P(y)}$.

Product Rule The formula can be rearranged as P(x, y) = P(x|y) P(y) = P(y|x) P(x). In case of multiple variables

$$P(w, x, y, z) = P(w, x, y|z) P(z)$$

$$= P(w, x|y, z) P(y|z) P(z)$$

$$= P(w|x, y, z) P(x|y, z) P(y|z) P(z)$$

Independence If two variables x and y are independent, then r.v. x tells nothing about r.v. y (and vice-versa)

$$P(x|y) = P(x)$$

$$P(y|x) = P(y)$$

$$P(x,y) = P(x) P(y)$$

Baye's Rule By rearranging formula in Product Rule, we have

$$P(y|x) = \frac{P(x|y) P(y)}{P(x)}$$

$$= \frac{P(x|y) P(y)}{\int P(x,y) dy}$$

$$= \frac{P(x|y) P(y)}{\int P(x|y) P(y) dy}$$

Expectation Expectation tells us the excepted or average value of some function f(x), taking into account the distribution of x.

$$\mathbf{E}[f(x)] = \sum_{x} f(x)P(x)$$
$$\mathbf{E}[f(x)] = \int f(x)P(x) dx$$

Definition in two dimensions: $\mathbf{E}\left[f(x,y)\right] = \iint f(x,y)P\left(x,y\right)dx\ dy$

Function $f(\bullet)$	Expectation
x^k	k^{th} moment about zero
$(x-\mu_x)^k$	k^{th} moment about the mean

Function $f(\bullet)$	Expectation
x	mean, μ_x
$(x-\mu_x)^2$	variance
$(x-\mu_x)^3$	skew
$(x-\mu_x)^4$	kurtosis
$(x-\mu_x)(x-\mu_y)$	covariance of x and y

Besides, Expectation has the following four rules

- 1. Expected value of a constant is the constant $\mathbf{E}\left[\kappa\right] = \kappa$.
- 2. Expected value of constant times function is constant times excepted value of function $\mathbf{E}[kf(x)] = k\mathbf{E}[f(x)].$
- 3. Expectation of sum of functions is sum of expectation of functions $\mathbf{E}[f(x) + g(y)] = \mathbf{E}[f(x)] + \mathbf{E}[g(x)].$

2.1.2 Common Probability Distributions

Bernoulli Bernoulli distribution describes situation where only two possible outcomes y = 0/y = 1 or failure/success.

- 1. $P(x) = \mathbf{Bern}_x[\lambda] = \lambda^x (1 \lambda)^{1-x}$
- 2. univariate, discrete, binary
- 3. $x \in \{0, 1\}; \lambda \in [0, 1]$
- 4. $\mathbf{E}[x] = \lambda$, $\mathbf{Var}[x] = \lambda(1 \lambda)$

Beta Beta distribution is the conjugate distribution to Bernoulli distribution.

- 1. $P(\lambda) = \mathbf{Beta}_{\lambda}[\alpha, \beta] = \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} \lambda^{\alpha 1} (1 \lambda)^{\beta 1}$
- 2. univariate, continuous, unbounded
- 3. $\lambda \in \mathbb{R}$; $\alpha \in \mathbb{R}_+$, $\beta \in \mathbb{R}_+$
- 4. $\mathbf{E}[\lambda] = \frac{\alpha}{\alpha + \beta}$, $\mathbf{Var}[\lambda] = \frac{\alpha\beta}{(\alpha + \beta)^2(\alpha + \beta + 1)}$

Categorical Categorical distribution describes situation with K possible outcomes.

1.
$$P(x) = \mathbf{Cat}_x[\lambda], P(x = k) = \lambda_k, P(\mathbf{x} = \mathbf{e}_k) = \prod_{j=1}^K \lambda_j^{x_j} = \lambda_k$$

- 2. univariate, discrete, multi-valued
- 3. $x \in \{1, 2, \dots, K\}; \lambda_k \in [0, 1] \text{ where } \sum_k \lambda_k = 1$
- 4. $\mathbf{E}[x_i] = \lambda_i$, $\mathbf{Var}[x_i] = \lambda_i(1 \lambda_i)$, $\mathbf{Cov}[x_i, x_j] = -\lambda_i\lambda_j \ (i \neq j)$

Dirichlet Dirichlet distribution is the conjugate distribution to categorical distribution.

1.
$$P(\lambda) = \mathbf{Dir}_{\lambda}[\alpha] = \frac{\Gamma(\sum_{k=1}^{K} \alpha_k)}{\prod_{k=1}^{K} \Gamma(\alpha_k)} \prod_{k=1}^{K} \lambda_k^{\alpha_k - 1}$$

- 2. multivariate, continuous, bounded, sums to one
- 3. $\lambda = [\lambda_1, \lambda_2, \dots, \lambda_K]^{\top}, \ \lambda_k \in [0, 1], \ \sum_{k=1}^K \lambda_k = 1; \ \alpha = [\alpha_1, \alpha_2, \dots, \alpha_K], \ \alpha_k \in \mathbb{R}_+$

4.
$$\mathbf{E}\left[\lambda_{i}\right] = \frac{\alpha_{i}}{\sum_{k} \alpha_{k}}, \, \mathbf{Var}\left[\lambda_{i}\right] = \frac{\alpha_{i}\left(\sum_{k} \alpha_{k} - \alpha_{i}\right)}{\left(\sum_{k} \alpha_{k}\right)^{2}\left(\sum_{k} \alpha_{k} + 1\right)}, \, \mathbf{Cov}\left[\lambda_{i}, \lambda_{j}\right] = \frac{-\alpha_{i}\alpha_{j}}{\left(\sum_{k} \alpha_{k}\right)^{2}\left(\sum_{k} \alpha_{k} + 1\right)} \, \left(i \neq j\right)$$

Univariate Normal Univariate normal distribution describes single continuous variable.

1.
$$P(x) = \mathbf{Norm}_x[\mu, \sigma^2] = \frac{1}{\sqrt{2\pi\sigma^2}} exp\left(-\frac{(x-\mu)^2}{2\sigma^2}\right)$$

- 2. univariate, continuous, unbounded
- 3. $x \in \mathbb{R}$; $\mu \in \mathbb{R}$, $\sigma^2 \in \mathbb{R}_+$
- 4. $\mathbf{E}[x] = \mu, \, \mathbf{Var}[x] = \sigma^2$

Normal Inverse Gamma Normal inverse gamma distribution is a conjugate distribution to univariate normal distribution.

1.
$$P(\mu, \sigma^2) = \mathbf{NormInvGam}_{\mu, \sigma^2}[\alpha, \beta, \gamma, \delta] = \frac{\sqrt{\gamma}}{\sqrt{2\pi\sigma^2}} \frac{\beta^{\alpha}}{\Gamma(\alpha)} \left(\frac{1}{\sigma^2}\right)^{\alpha+1} exp\left(-\frac{2\beta+\gamma(\delta-\mu)^2}{2\sigma^2}\right)$$

- 2. bivariate, continuous, μ unbounded, σ^2 bounded below
- 3. $\mu \in \mathbb{R}, \ \sigma^2 \in \mathbb{R}_+; \ \alpha \in \mathbb{R}_+, \ \beta \in \mathbb{R}_+, \ \gamma \in \mathbb{R}_+, \ \delta \in \mathbb{R}_+$

4.
$$\mathbf{E}[\mu] = \delta$$
, $\mathbf{E}[\sigma^2] = \frac{\beta}{\alpha - 1} (\alpha > 1)$, $\mathbf{Var}[\mu] = \frac{\beta}{(\alpha - 1)\gamma} (\alpha > 1)$, $\mathbf{Var}[\sigma^2] = \frac{\beta^2}{(\alpha - 1)^2(\alpha - 2)} (\alpha > 2)$, $\mathbf{Cov}[\mu, \sigma^2] = 0 (\alpha > 1)$

Multivariate Normal Multivariate normal distribution describes multiple continuous variables. It takes two parameters: a vector containing mean position μ , and a symmetric positive definite covariance matrix Σ .

1.
$$P(\boldsymbol{x}) = \mathbf{Norm}_{\boldsymbol{x}}[\boldsymbol{\mu}, \boldsymbol{\Sigma}] = \frac{1}{\sqrt{\det(2\pi\boldsymbol{\Sigma})}} exp\left(-\frac{1}{2}(\boldsymbol{x} - \boldsymbol{\mu})^{\top} \boldsymbol{\Sigma}^{-1}(\boldsymbol{x} - \boldsymbol{\mu})\right)$$

- 2. multivariate, continuous, unbounded
- 3. $\boldsymbol{x} \in \mathbb{R}^K$; $\boldsymbol{\mu} \in \mathbb{R}^K$, $\boldsymbol{\Sigma} \in \mathbb{R}^{K \times K}$ (positive semi-definite matrix)
- 4. $\mathbf{E}[x] = \mu$, $\mathrm{Var}[x] = \Sigma$

Normal Inverse Wishart Normal inverse wishart distribution is a conjugate distribution to multivariate normal distribution.

1.
$$P(\boldsymbol{\mu}, \boldsymbol{\Sigma}) = \mathbf{NormInvWis}_{\boldsymbol{\mu}, \boldsymbol{\Sigma}}[\alpha, \boldsymbol{\Psi}, \gamma, \boldsymbol{\delta}]$$

= $\frac{\gamma^{D/2} |\boldsymbol{\Psi}|^{\alpha/2} |\boldsymbol{\Sigma}|^{-\frac{\alpha+D+2}{2}}}{(2\pi)^{D/2} 2^{(\alpha\boldsymbol{\Sigma})/2} \Gamma_D(\alpha/2)} exp\left(-\frac{1}{2}(\text{Tr}(\boldsymbol{\Psi}\boldsymbol{\Sigma}^{-1}) + \gamma(\boldsymbol{\mu} - \boldsymbol{\delta})^{\top} \boldsymbol{\Sigma}^{-1}(\boldsymbol{\mu} - \boldsymbol{\delta})\right)$

- 2. multivariate, μ unbounded, Σ square, positive definite
- 3. $\boldsymbol{\mu} \in \mathbb{R}^K$, $\boldsymbol{\Sigma} \in \mathbb{R}^{K \times K}$; $\alpha \in \mathbb{R}_{>D-1}$, $\boldsymbol{\Psi} \in \mathbb{R}^{K \times K}$, $\gamma \in \mathbb{R}_+$, $\boldsymbol{\delta} \in \mathbb{R}^K$

- 2.2 Linear Algebra
- 2.3 Calculus
- 2.4 Informatics
- 2.5 Optimization

Machine Learning Basics

3.1 Regularization

3.1.1 Under-fitting & Over-fitting

Under-fitting If N > D (e.g. 30 data points, 2 dimensions) we have more equations than unknowns: over-determined system. Input-output relations can only hold approximately.

Over-fitting If N < D (e.g. 30points, 15265 dimensions) we have more unknowns than equations: under-determined system. Input-output equations hold exactly, but we are simply memorizing data.

3.1.2 Bias & Variance

High Bias & Low Variance A rigid model's (low complexity) performance is more predictable in the test set but the model may not be good even on the training set.

Low Bias & High Variance A flexible model (high complexity) approximates the target function well in the training set but can "overtrain" and have poor performance on the test set.

3.1.3 Vector Norm

L1, ("Manhattan") norm
$$||\boldsymbol{w}||_1 = \sum_{d=1}^D |w_d|$$

L2, ("Euclidean") norm
$$||\boldsymbol{w}||_2 = \sqrt{\sum_{d=1}^D w_d^2} = \sqrt{\langle \boldsymbol{w}, \boldsymbol{w} \rangle} = \sqrt{\boldsymbol{w}^{\top} \boldsymbol{w}}$$

Lp norm, p>1
$$||\boldsymbol{w}||_p = \left(\sum_{d=1}^D w_d^p\right)^{1/p}$$

3.1.4 Penalize Complexity

In linear regression, the residual vector is $\boldsymbol{\epsilon} = \boldsymbol{y} - \boldsymbol{\Psi} \boldsymbol{w}$. The loss function is $L(\boldsymbol{w}) = \boldsymbol{\epsilon}^{\top} \boldsymbol{\epsilon}$. We add a complexity term $R(\boldsymbol{w}) = ||\boldsymbol{w}||_2 = \boldsymbol{w}^{\top} \boldsymbol{w}$ to the loss function. Hence, the original loss function becomes $L(\boldsymbol{w}) = \boldsymbol{\epsilon}^{\top} \boldsymbol{\epsilon} + \lambda \boldsymbol{w}^{\top} \boldsymbol{w}$.

Without regularization, the loss function is $L(\boldsymbol{w}) = \boldsymbol{\epsilon}^{\top} \boldsymbol{\epsilon}$. Let $\nabla L(\boldsymbol{w}^*) = 0$, we have $\boldsymbol{w}^* = (\mathbf{X}^{\top} \mathbf{X})^{-1} \mathbf{X}^{\top} \boldsymbol{y}$.

With L2-regularization, the loss function is $L(\boldsymbol{w}) = \boldsymbol{\epsilon}^{\top} \boldsymbol{\epsilon} + \lambda \boldsymbol{w}^{\top} \boldsymbol{w}$. Let $\nabla L(\boldsymbol{w}^*) = 0$, we have $\boldsymbol{w}^* = (\mathbf{X}^{\top} \mathbf{X} + \lambda \mathbf{I})^{-1} \mathbf{X}^{\top} \boldsymbol{y}$. The additional $\lambda \mathbf{I}$ makes the data matrix more robust to calculate inversion.

3.2 Cross-Validation

We can select hyperparameters with (cross-)validation. Cross-validation excludes part of the training data from parameter estimation, and use them only to predict the test error.

K-fold cross validation: split data set into K folds and each time train on (K-1) folds and valid on the remaining fold until all folds have been used as validation fold. The cross-validation error is the average of K validation errors. We pick hyperparameters that minimize cross-validation error.

3.3 Bayesian Learning

3.3.1 Bayes' Rule Terminology

Bayes' Rule:

$$P(y|x) = \frac{P(x|y) P(y)}{\int P(x|y) P(y) dy}$$

Prior P(y) what we know about y before seeing x. In parameters learning we choose prior that is conjugate to likelihood.

Likelihood P(x|y) propensity for observing a certain value of x given a certain value of y.

Posterior P(y|x) what we know about y after seeing x. Posterior must have same form as conjugate prior distribution.

Evidence $\int P(x|y) P(y) dy$ a constant to ensure that the LHS is a valid distribution. Posterior must be a distribution which implies that evidence equals to a constant κ from conjugate relation.

3.3.2 Maximum Likelihood

Fitting As the name suggests we find the parameters under which the data $x_{1...I}$ are most likely. Here, we have assumed that data was independent.

$$\hat{\boldsymbol{\theta}} = \underset{\boldsymbol{\theta}}{\operatorname{argmax}} P(\boldsymbol{x}_{1...I} | \boldsymbol{\theta})$$
$$= \underset{\boldsymbol{\theta}}{\operatorname{argmax}} \prod_{i=1}^{I} P(\boldsymbol{x}_{i} | \boldsymbol{\theta})$$

Predictive Density Evaluate new data point x^* under probability distribution $P\left(x^*|\hat{\theta}\right)$ with best parameters.

3.3.3 Maximum a Posterior (MAP)

Fitting As the name suggests we find the parameters which maximize the posterior probability $P(\theta|\mathbf{x}_{1...I})$. Again we have assumed that data was independent.

$$\hat{\boldsymbol{\theta}} = \underset{\boldsymbol{\theta}}{\operatorname{argmax}} P\left(\boldsymbol{\theta}|\boldsymbol{x}_{1...I}\right)$$

$$= \underset{\boldsymbol{\theta}}{\operatorname{argmax}} \frac{P\left(\boldsymbol{x}_{1...I}|\boldsymbol{\theta}\right) P\left(\boldsymbol{\theta}\right)}{P\left(\boldsymbol{x}_{1...I}\right)}$$

$$= \underset{\boldsymbol{\theta}}{\operatorname{argmax}} \frac{\prod_{i=1}^{I} P\left(\boldsymbol{x}_{i}|\boldsymbol{\theta}\right) P\left(\boldsymbol{\theta}\right)}{P\left(\boldsymbol{x}_{1...I}\right)}$$

Since the denominator does not depend on the parameters we can instead maximize

$$\hat{\boldsymbol{\theta}} = \operatorname*{argmax}_{\boldsymbol{\theta}} \prod_{i=1}^{I} P\left(\boldsymbol{x}_{i} | \boldsymbol{\theta}\right) \ P\left(\boldsymbol{\theta}\right)$$

Predictive Density Evaluate new data point x^* under probability distribution with MAP parameters $P\left(x^*|\hat{\theta}\right)$

3.3.4 Bayesian Approach

Fitting Compute the posterior distribution over possible parameter values using Bayes' rule. Principle: There are many values that could have explained the data. Instead of picking one set of parameters, try to capture all of the possibilities.

$$P\left(\boldsymbol{\theta}|\boldsymbol{x}_{1...I}\right) = \frac{\prod_{i=1}^{I} P\left(\boldsymbol{x}_{i}|\boldsymbol{\theta}\right) \ P\left(\boldsymbol{\theta}\right)}{P\left(\boldsymbol{x}_{1...I}\right)}$$

Predictive Density (a) Each possible parameter value makes a prediction. (b) Some parameters more probable than others.

$$P\left(\boldsymbol{x}^{*}|\boldsymbol{x}_{1...I}\right) = \int P\left(\boldsymbol{x}^{*}|\boldsymbol{\theta}\right) P\left(\boldsymbol{\theta}|\boldsymbol{x}_{1...I}\right) d\boldsymbol{\theta}$$

Make a prediction that is an infinite weighted sum (integral) of the predictions for each parameter value $(P(\boldsymbol{x}^*|\boldsymbol{\theta}))$, where weights are the probabilities $(P(\boldsymbol{\theta}|\boldsymbol{x}_{1...I}))$.

3.3.5 Example: Univariate Normal Distribution

Maximum Likelihood

Likelihood given by normal distribution pdf:

$$P(x|\mu,\sigma^2) = \mathbf{Norm}_x[\mu,\sigma^2] = \frac{1}{\sqrt{2\pi\sigma^2}} exp\left(-\frac{(x-\mu)^2}{2\sigma^2}\right)$$

Apply maximum likelihood:

$$\begin{split} \hat{\mu}, \hat{\sigma}^2 &= \operatorname*{argmax}_{\mu, \sigma^2} P\left(x_{1...I} | \mu, \sigma^2\right) \\ &= \operatorname*{argmax}_{\mu, \sigma^2} \prod_{i=1}^{I} P\left(x_i | \mu, \sigma^2\right) \\ &= \operatorname*{argmax}_{\mu, \sigma^2} \prod_{i=1}^{I} \mathbf{Norm}_{x_i} [\mu, \sigma^2] \\ &= \operatorname*{argmax}_{\mu, \sigma^2} \sum_{i=1}^{I} \log \mathbf{Norm}_{x_i} [\mu, \sigma^2] \\ &= \operatorname*{argmax}_{\mu, \sigma^2} \left(-\frac{I}{2} \log 2\pi - \frac{I}{2} \log \sigma^2 - \frac{1}{2} \sum_{i=1}^{I} \frac{(x_i - \mu)^2}{\sigma^2} \right) \end{split}$$

Let $\nabla L(\hat{\mu}, \hat{\sigma}^2) = 0$, we have the solution:

$$\hat{\mu} = \frac{\sum_{i=1}^{I} x_i}{I}$$

$$\hat{\sigma}^2 = \sum_{i=1}^{I} \frac{(x_i - \hat{\mu})^2}{I}$$

Maximum a Posterior

Likelihood given by normal distribution pdf:

$$P(x|\mu, \sigma^2) = \mathbf{Norm}_x[\mu, \sigma^2] = \frac{1}{\sqrt{2\pi\sigma^2}} exp\left(-\frac{(x-\mu)^2}{2\sigma^2}\right)$$

Prior given by normal inverse gamma distribution pdf:

$$P\left(\mu,\sigma^{2}\right) = \mathbf{NormInvGam}_{\mu,\sigma^{2}}[\alpha,\beta,\gamma,\delta] = \frac{\sqrt{\gamma}}{\sqrt{2\pi\sigma^{2}}} \frac{\beta^{\alpha}}{\Gamma(\alpha)} \left(\frac{1}{\sigma^{2}}\right)^{\alpha+1} exp\left(-\frac{2\beta + \gamma(\delta - \mu)^{2}}{2\sigma^{2}}\right)$$

Apply maximum a posterior:

$$\begin{split} \hat{\mu}, \hat{\sigma}^2 &= \operatorname*{argmax}_{\mu, \sigma^2} \prod_{i=1}^{I} P\left(x_i | \mu, \sigma^2\right) \ P\left(\mu, \sigma^2\right) \\ &= \operatorname*{argmax}_{\mu, \sigma^2} \prod_{i=1}^{I} \mathbf{Norm}_{x_i} [\mu, \sigma^2] \ \mathbf{NormInvGam}_{\mu, \sigma^2} [\alpha, \beta, \gamma, \delta] \\ &= \operatorname*{argmax}_{\mu, \sigma^2} \left(\sum_{i=1}^{I} \log \mathbf{Norm}_{x_i} [\mu, \sigma^2] + \log \mathbf{NormInvGam}_{\mu, \sigma^2} [\alpha, \beta, \gamma, \delta] \right) \end{split}$$

Let $\nabla L(\hat{\mu}, \hat{\sigma}^2) = 0$, we have the solution:

$$\hat{\mu} = \frac{\sum_{i=1}^{I} x_i + \gamma \delta}{I + \gamma}$$

$$\hat{\sigma}^2 = \frac{\sum_{i=1}^{I} (x_i - \mu)^2 + 2\beta + \gamma (\delta - \mu)^2}{I + 3 + 2\alpha}$$

Bayesian Approach

Compute the posterior distribution using Bayes' rule:

$$\begin{split} P\left(\mu,\sigma^{2}|x_{1...I}\right) &= \frac{\prod_{i=1}^{I} P\left(x_{i}|\mu,\sigma^{2}\right) \ P\left(\mu,\sigma^{2}\right)}{P\left(x_{1...I}\right)} \\ &= \frac{\prod_{i=1}^{I} \mathbf{Norm}_{x_{i}}[\mu,\sigma^{2}] \ \mathbf{NormInvGam}_{\mu,\sigma^{2}}[\alpha,\beta,\gamma,\delta]}{P\left(x_{1...I}\right)} \\ &= \frac{\kappa(\alpha,\beta,\gamma,\delta,x_{1...I}) \mathbf{NormInvGam}_{\mu,\sigma^{2}}[\tilde{\alpha},\tilde{\beta},\tilde{\gamma},\tilde{\delta}]}{P\left(x_{1...I}\right)} \\ &= \mathbf{NormInvGam}_{\mu,\sigma^{2}}[\tilde{\alpha},\tilde{\beta},\tilde{\gamma},\tilde{\delta}] \end{split}$$

where

$$\begin{split} \tilde{\alpha} &= \alpha + \frac{I}{2} \\ \tilde{\beta} &= \frac{\sum_{i} x_{i}^{2}}{2} + \beta + \frac{\gamma \delta^{2}}{2} - \frac{(\gamma \delta + \sum_{i} x_{i})^{2}}{2(\gamma + I)} \\ \tilde{\gamma} &= \gamma + I \\ \tilde{\delta} &= \frac{\gamma \delta + \sum_{i} x_{i}}{\gamma + I} \end{split}$$

Take weighted sum of predictions from different parameter values:

$$\begin{split} P\left(\boldsymbol{x}^{*}|\boldsymbol{x}_{1...I}\right) &= \int\!\!\!\int P\left(\boldsymbol{x}^{*}|\boldsymbol{\mu},\sigma^{2}\right) P\left(\boldsymbol{\mu},\sigma^{2}|\boldsymbol{x}_{1...I}\right) d\boldsymbol{\mu} d\sigma \\ &= \int\!\!\!\!\int \mathbf{Norm}_{\boldsymbol{x}^{*}}[\boldsymbol{\mu},\sigma^{2}] \mathbf{NormInvGam}_{\boldsymbol{\mu},\sigma^{2}}[\tilde{\alpha},\tilde{\beta},\tilde{\gamma},\tilde{\delta}] d\boldsymbol{\mu} d\sigma \\ &= \int\!\!\!\!\int \kappa(\alpha,\beta,\gamma,\delta,\boldsymbol{x}_{1...I}) \mathbf{NormInvGam}_{\boldsymbol{\mu},\sigma^{2}}[\boldsymbol{\alpha},\boldsymbol{\beta},\tilde{\gamma},\tilde{\delta}] d\boldsymbol{\mu} d\sigma \\ &= \kappa(\alpha,\beta,\gamma,\delta,\boldsymbol{x}_{1...I}) \int\!\!\!\!\!\int \mathbf{NormInvGam}_{\boldsymbol{\mu},\sigma^{2}}[\boldsymbol{\alpha},\boldsymbol{\beta},\tilde{\gamma},\tilde{\delta}] d\boldsymbol{\mu} d\sigma \\ &= \kappa(\alpha,\beta,\gamma,\delta,\boldsymbol{x}_{1...I}) \\ &= \frac{1}{\sqrt{2\pi}} \frac{\sqrt{\tilde{\gamma}}\tilde{\beta}^{\tilde{\alpha}}}{\sqrt{\tilde{\gamma}}\tilde{\beta}^{\tilde{\alpha}}} \frac{\Gamma(\boldsymbol{\alpha})}{\Gamma(\tilde{\alpha})} \end{split}$$

where

$$\ddot{\alpha} = \tilde{\alpha} + \frac{1}{2}$$

$$\ddot{\beta} = \frac{x^{*2}}{2} + \tilde{\beta} + \frac{\tilde{\gamma}^{\tilde{\delta}^2}}{2} - \frac{(\tilde{\gamma}^{\tilde{\delta}} + x^*)^2}{2(\tilde{\gamma} + 1)}$$

$$\ddot{\gamma} = \tilde{\gamma} + 1$$

3.3.6 Example: Categorical Distribution

Maximum Likelihood

Likelihood given by categorical distribution pdf:

$$P(x|\boldsymbol{\lambda}) = \mathbf{Cat}_x[\boldsymbol{\lambda}] = \prod_{j=1}^K \lambda_j^{x_j} = \lambda_k$$

Apply maximum likelihood:

$$\hat{\lambda} = \underset{\lambda}{\operatorname{argmax}} \prod_{i=1}^{I} P(x_{i}|\lambda) \qquad s.t. \sum_{k} \lambda_{k} = 1$$

$$= \underset{\lambda}{\operatorname{argmax}} \prod_{i=1}^{I} \mathbf{Cat}_{x_{i}}[\lambda] \qquad s.t. \sum_{k} \lambda_{k} = 1$$

$$= \underset{\lambda}{\operatorname{argmax}} \prod_{k=1}^{K} \lambda_{k}^{N_{k}} \qquad s.t. \sum_{k} \lambda_{k} = 1$$

$$= \underset{\lambda}{\operatorname{argmax}} \sum_{k=1}^{K} N_{k} \log \lambda_{k} \qquad s.t. \sum_{k} \lambda_{k} = 1$$

$$= \underset{\lambda}{\operatorname{argmax}} \sum_{k=1}^{K} N_{k} \log \lambda_{k} \qquad s.t. \sum_{k} \lambda_{k} = 1$$

Here, N_k represents the number of times the data is classified in class k. As before, we will instead optimize log probability. Since there is a constraint $s.t. \sum_k \lambda_k = 1$, we use Lagrange multiplier to reconstruct the loss function.

$$L(\lambda) = \sum_{k=1}^{K} N_k \log \lambda_k + v \left(\sum_{k=1}^{K} \lambda_k - 1 \right)$$

Let $\nabla L(\lambda, v) = 0$, we have the solution:

$$\hat{\lambda}_k = \frac{N_k}{\sum_{m=1}^K N_m}$$

Maximum a Posterior

Likelihood given by categorical distribution pdf:

$$P(x|\lambda) = \mathbf{Cat}_x[\lambda] = \prod_{j=1}^K \lambda_j^{x_j} = \lambda_k$$

Prior given by Dirichlet distribution pdf:

$$P(\boldsymbol{\lambda}) = \mathbf{Dir}_{\boldsymbol{\lambda}}[\boldsymbol{\alpha}] = \frac{\Gamma(\sum_{k=1}^{K} \alpha_k)}{\prod_{k=1}^{K} \Gamma(\alpha_k)} \prod_{k=1}^{K} \lambda_k^{\alpha_k - 1}$$

Apply maximum a posterior:

$$\hat{\lambda} = \underset{\lambda}{\operatorname{argmax}} \prod_{i=1}^{I} P(x_{i}|\lambda) P(\lambda) \qquad s.t. \sum_{k} \lambda_{k} = 1$$

$$= \underset{\lambda}{\operatorname{argmax}} \operatorname{Cat}_{x_{i}}[\lambda] \operatorname{Dir}_{\lambda}[\alpha] \qquad s.t. \sum_{k} \lambda_{k} = 1$$

$$= \underset{\lambda}{\operatorname{argmax}} \prod_{k=1}^{K} \lambda_{k}^{N_{k}} \prod_{k=1}^{K} \lambda_{k}^{\alpha_{k}-1} \qquad s.t. \sum_{k} \lambda_{k} = 1$$

$$= \underset{\lambda}{\operatorname{argmax}} \prod_{k=1}^{K} \lambda_{k}^{N_{k}+\alpha_{k}-1} \qquad s.t. \sum_{k} \lambda_{k} = 1$$

$$= \underset{\lambda}{\operatorname{argmax}} \sum_{k=1}^{K} (N_{k} + \alpha_{k} - 1) \log \lambda_{k} \qquad s.t. \sum_{k} \lambda_{k} = 1$$

The loss function is very similar to maximum likelihood (same when the prior is uniform, i.e. $\alpha_{1...k} = 1$). Take derivative with Lagrange multiplier, we have the solution:

$$\hat{\lambda}_k = \frac{N_k + \alpha_k - 1}{\sum_{m=1}^K (N_m + \alpha_m - 1)}$$

Bayesian Approach

Compute the posterior distribution using Bayes' rule:

$$\begin{split} P\left(\boldsymbol{\lambda}|\boldsymbol{x}_{1...I}\right) &= \frac{\prod_{i=1}^{I} P\left(\boldsymbol{x}_{i}|\boldsymbol{\lambda}\right) \ P\left(\boldsymbol{\lambda}\right)}{P\left(\boldsymbol{x}_{1...I}\right)} \\ &= \frac{\prod_{i=1}^{I} \mathbf{Cat}_{\boldsymbol{x}_{i}}[\boldsymbol{\lambda}] \ \mathbf{Dir}_{\boldsymbol{\lambda}}[\boldsymbol{\alpha}]}{P\left(\boldsymbol{x}_{1...I}\right)} \\ &= \frac{\kappa(\boldsymbol{\alpha}, \boldsymbol{x}_{1...I}) \mathbf{Dir}_{\boldsymbol{\lambda}}[\tilde{\boldsymbol{\alpha}}]}{P\left(\boldsymbol{x}_{1...I}\right)} \\ &= \mathbf{Dir}_{\boldsymbol{\lambda}}[\tilde{\boldsymbol{\alpha}}] \end{split}$$

Compute predictive distribution:

$$P(x^*|x_{1...I}) = \int P(x^*|\lambda) P(\lambda|x_{1...I}) d\lambda$$
$$= \int \mathbf{Cat}_{x^*}[\lambda] \mathbf{Dir}_{\lambda}[\tilde{\alpha}] d\lambda$$
$$= \int \kappa(x^*, \tilde{\alpha}) \mathbf{Dir}_{\lambda}[\check{\alpha}] d\lambda$$
$$= \kappa(x^*, \alpha)$$

3.4 Machine Learning Models

3.4.1 Learning and Inference

In real world problems, we usually have two tasks:

- 1. Observe measured data, \mathbf{x}
- 2. Draw inferences from it about world, w

and

- 1. When the world state \mathbf{w} is *continuous*, we'll call this *regression*.
- 2. When the world state \mathbf{w} is discrete, we'll call this classification.

We want take observations \mathbf{x} , and return probability distribution $P(\mathbf{w}|\mathbf{x})$ over possible worlds compatible with data. To solve this, we need

- 1. A model that mathematically relates the visual data \mathbf{x} to the world state \mathbf{w} . Model specifies family of relationships, particular relationship depends on parameter θ .
- 2. A learning algorithm fits parameters θ from paired training examples $\mathbf{x_i}$, $\mathbf{w_i}$.
- 3. An inference algorithm uses model to return $P(\mathbf{w}|\mathbf{x})$ given new observation data \mathbf{x} .

3.4.2 Three Types of Model

We have three types of model:

- 1. Model contingency of the world on the data P(w|x). (Discriminative Model)
- 2. Model joint occurrence of world and data P(x, w). (Generative Model)
- 3. Model contingency of data on world P(x|w). (Generative Model)

Within the three models, type 1 is called *Discriminative Model*. Type 2 and 3 are called *Generative Model*.

Model P(w|x) - Discriminative

- 1. $P(w|x,\theta) = \mathbf{Distrib}_w[f(x,\theta)]$
- 2. How to model: (a) Choose an appropriate form for P(w). (b) Make parameters a function of x. (c) Function takes parameters θ that define its shape.
- 3. Learning algorithm: Learn parameters θ from training data x, w.
- 4. Inference algorithm: Just evaluate P(w|x)

Model P(x, w) - Generative

- 1. $P(z|\theta) = \mathbf{Distrib}_z[\theta]$
- 2. How to model: (a) Concatenate x and w to make $z = [x^{\top}, w^{\top}]^{\top}$. (b) Model the pdf of z. (c) pdf takes parameters θ that define its shape.
- 3. Learning algorithm: Learn parameters θ from training data x, w.
- 4. Inference algorithm: Compute P(w|x) using Bayes' rule $P(w|x) = \frac{P(x,w)}{P(x)} = \frac{P(x,w)}{\int P(x,w)dw}$.

Model P(x|w) - Generative

- 1. $P(x|w,\theta) = \mathbf{Distrib}_x[f(w,\theta)]$
- 2. How to model: (a) Choose an appropriate form for P(x). (b) Make parameters a function of w. (c) Function takes parameters θ that define its shape.
- 3. Learning algorithm: Learn parameters θ from training data x, w.
- 4. Define prior $P\left(w\right)$ and then compute $P\left(w|x\right)$ using Bayes' rule $P\left(w|x\right) = \frac{P\left(x|w\right)P\left(w\right)}{\int P\left(x|w\right)P\left(w\right)dw}$

3.4.3 Example: Regression

Consider a simple case:

- 1. We make a univariate continuous measurement x.
- 2. Use this to predict a univariate continuous state w.

Model P(w|x) - Discriminative

- 1. $P(w|x, \theta) = \mathbf{Norm}_w[\phi_0 + \phi_1 x, \sigma^2], \theta = \{\phi_0, \phi_1, \sigma^2\}$
- 2. How to model: (a) Choose normal distribution over w. (b) Make mean μ linear function of x (variance constant). (c) Parameters are ϕ_0 (y-offset), ϕ_1 (slope), σ^2 (variance). This model is called *linear regression*.
- 3. Learning algorithm: Learn $\boldsymbol{\theta}$ from training data x, w. e.g. MAP:

$$\hat{\boldsymbol{\theta}} = \underset{\boldsymbol{\theta}}{\operatorname{argmax}} P\left(\boldsymbol{\theta}|w_{1...I}, x_{1...I}\right)$$

$$= \underset{\boldsymbol{\theta}}{\operatorname{argmax}} P\left(w_{1...I}|x_{1...I}, \boldsymbol{\theta}\right) P\left(\boldsymbol{\theta}\right)$$

$$= \underset{\boldsymbol{\theta}}{\operatorname{argmax}} \prod_{i=1}^{I} P\left(w_{i}|x_{i}, \boldsymbol{\theta}\right) P\left(\boldsymbol{\theta}\right)$$

4. Inference algorithm: Just evaluate P(w|x) for new data x.

Model P(x, w) - Generative

- 1. $P(x, w|\boldsymbol{\theta}) = \mathbf{Norm}_{x,w}[\boldsymbol{\mu}, \boldsymbol{\Sigma}], \boldsymbol{\theta} = \{\boldsymbol{\mu}, \boldsymbol{\Sigma}\}$
- 2. How to model: (a) Concatenate x and w to make $z = [x^{\top}, w^{\top}]^{\top}$. (b) Model the pdf of z as normal distribution. (c) pdf makes parameters μ and Σ that define its shape.
- 3. Learning algorithm: Learn parameters θ from training data x, w.
- 4. Inference algorithm: Compute P(w|x) using Bayes' rule $P(w|x) = \frac{P(x,w)}{P(x)} = \frac{P(x,w)}{\int P(x,w)dw}$.

Model P(x|w) - Generative

- 1. $P(x|w, \theta) = \mathbf{Norm}_x[\phi_0 + \phi_1 w, \sigma^2], \theta = \{\phi_0, \phi_1, \sigma^2\}$
- 2. How to model: (a) Choose normal distribution over x. (b) Make mean μ linear function of w (variance constant). (c) Parameters are ϕ_0 , ϕ_1 , σ^2 .

- 3. Learning algorithm: Learn $\boldsymbol{\theta}$ from training data x, w. e.g. MAP
- 4. Inference algorithm: Compute P(w|x) using Bayes' rule $P(w|x) = \frac{P(x,w)}{P(x)} = \frac{P(x,w)}{\int P(x,w)dw}$.

3.4.4 Example: Classification

Consider a simple case:

- 1. We make a univariate continuous measurement x.
- 2. Use this to predict a discrete binary world $w \in \{0, 1\}$.

Model P(w|x) - Discriminative

- 1. $P(w|x, \boldsymbol{\theta}) = \mathbf{Bern}_w[\sigma(\phi_0 + \phi_1 x)], \boldsymbol{\theta} = \phi_0, \phi_1$
- 2. How to model: (a) Choose Bernoulli distribution for P(w). (b) Make parameters a sigmoid-activated function of x. (c) Function takes parameters ϕ_0 and ϕ_1 . This model is called *logistic regression*.
- 3. Learning algorithm: Learning by standard methods, e.g. ML, MAP, Bayesian Approach.
- 4. Inference algorithm: Just evaluate P(w|x).

Model P(x, w) - Generative

Can't build this mode very easily:

- 1. Concatenate continuous vector x and discrete w to make z.
- 2. No obvious probability distribution to model joint probability of discrete and continuous.

Model P(x|w) - Generative

- 1. $P(x|w, \theta) = \mathbf{Norm}_x[\mu_w, \sigma_w^2], \theta = \{\mu_0, \mu_1, \sigma_0^2, \sigma_1^2\}$
- 2. How to model: (a) Choose a Normal distribution for P(x). (b) Make parameters a function of discrete binary w. (c) Function takes parameters μ_0 , μ_1 , σ_0^2 , σ_1^2 that define its shape.
- 3. Learning algorithm: Learning by standard methods, e.g. ML, MAP, Bayesian Approach.
- 4. Define prior P(w) and then compute P(w|x) using Bayes' rule $P(w|x) = \frac{P(x|w)P(w)}{\int P(x|w)P(w)dw}$

Regression

- 4.1 Linear Regression
- 4.2 Non-linear Regression
- 4.3 Kernel Trick & Gaussian Processes
- 4.4 Sparse Linear Regression
- 4.5 Dual Linear Regression
- 4.6 Relevance Vector Regression

Classification

- 5.1 Logistic Regression
- 5.2 Non-linear Logistic Regression
- 5.3 Kernel Trick & Gaussian Process Classification
- 5.4 Multi-class Classification

Support Vector Machines

- 6.1 Geometric Margins
- 6.2 Primal & Dual Problems
- 6.3 Support Vectors
- 6.4 Slack Variables
- 6.5 Hinge Loss
- 6.6 Non-linear SVMs
- 6.7 Kernel Trick

EM Algorithm

- 7.1 Expectation Maximization
- 7.2 Example: Mixture of Gaussians
- 7.3 Example: t-distributions
- 7.4 Example: Factor Analysis

Boosting

- 8.1 Ensemble Methods
- 8.2 Bagging
- 8.3 Boosting
- 8.4 Adaboost

Decision Tree & Random Forest

- 9.1 CART
- 9.2 ID3
- 9.3 C4.5
- 9.4 Random Forest

Graphical Models & Markov Network

- 10.1 Graph Definitions
- 10.2 Belief Networks
- 10.3 Markov Networks
- 10.4 Markov Chains
- 10.5 Hidden Markov Models

Bibliography

Appendix A

Statistical Assessment

- A.1 Hypothesis Testing
- A.2 Confidence Intervals