

Machine Learning Notebook

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Chapter 1

Introduction

1.1 About this Notebook

1.2 Policy of Use

Chapter 2

Mathematics Basics

2.1 Probability

2.1.1 Basic Rules

Three Axioms of Probability Let Ω be a sample space. A probability assigns a real number $P(X)$ to each event $X \subseteq \Omega$ in such a way that

1. $P(X) \geq 0, \forall X$
2. If X_1, X_2, \dots are pairwise disjoint events ($X_1 \cap X_2 = \emptyset, i \neq j, i, j = 1, 2, \dots$), then $P(\bigcup_{i=1}^{\infty} X_i) = \sum_{i=1}^{\infty} P(X_i)$. (This property is called countable additivity.)
3. $P(\Omega) = 1$

Joint Probability The probability both event A and B occur. $P(X, Y) = P(X \cap Y)$.

Marginalization The probability distribution of any variable in a joint distribution can be recovered by integrating (or summing) over the other variables.

1. For continuous r.v. $P(x) = \int P(x, y) dy ; P(y) = \int P(x, y) dx$.
2. For discrete r.v. $P(x) = \sum_y P(x, y) ; P(y) = \sum_x P(x, y)$.
3. For mixed r.v. $P(x, y) = \sum_w \int P(w, x, y, z) dz$, where w is discrete and z is continuous.

Conditional Probability $P(X = x|Y = y)$ is the probability $X = x$ occurs given the knowledge $Y = y$ occurs. Conditional probability can be extracted from joint probability that

$$P(x|y = y^*) = \frac{P(x, y = y^*)}{\int P(x, y = y^*) dx} = \frac{P(x, y = y^*)}{P(y = y^*)}$$

Usually, the formula is written as $P(x|y) = \frac{P(x, y)}{P(y)}$.

Product Rule The formula can be rearranged as $P(x, y) = P(x|y) P(y) = P(y|x) P(x)$.
In case of multiple variables

$$\begin{aligned} P(w, x, y, z) &= P(w, x, y|z) P(z) \\ &= P(w, x|y, z) P(y|z) P(z) \\ &= P(w|x, y, z) P(x|y, z) P(y|z) P(z) \end{aligned}$$

Independence If two variables x and y are independent, then r.v. x tells nothing about r.v. y (and vice-versa)

$$\begin{aligned} P(x|y) &= P(x) \\ P(y|x) &= P(y) \\ P(x, y) &= P(x) P(y) \end{aligned}$$

Baye's Rule By rearranging formula in Product Rule, we have

$$\begin{aligned} P(y|x) &= \frac{P(x|y) P(y)}{P(x)} \\ &= \frac{P(x|y) P(y)}{\int P(x, y) dy} \\ &= \frac{P(x|y) P(y)}{\int P(x|y) P(y) dy} \end{aligned}$$

Expectation Expectation tells us the expected or average value of some function $f(x)$, taking into account the distribution of x .

$$\begin{aligned} \mathbf{E}[f(x)] &= \sum_x f(x) P(x) \\ \mathbf{E}[f(x)] &= \int f(x) P(x) dx \end{aligned}$$

Definition in two dimensions: $\mathbf{E}[f(x, y)] = \iint f(x, y) P(x, y) dx dy$

Function $f(\bullet)$	Expectation
x	mean, μ_x
$(x - \mu_x)^2$	variance
$(x - \mu_x)^3$	skew
$(x - \mu_x)^4$	kurtosis
$(x - \mu_x)(x - \mu_y)$	covariance of x and y

Besides, Expectation has the following four rules

1. Expected value of a constant is the constant $\mathbf{E}[\kappa] = \kappa$.
2. Expected value of constant times function is constant times expected value of function $\mathbf{E}[kf(x)] = k\mathbf{E}[f(x)]$.
3. Expectation of sum of functions is sum of expectation of functions $\mathbf{E}[f(x) + g(y)] = \mathbf{E}[f(x)] + \mathbf{E}[g(y)]$.
4. Expectation of product of functions in variables x and y is product of expectations of functions if x and y are independent $\mathbf{E}[f(x)g(y)] = \mathbf{E}[f(x)]\mathbf{E}[g(y)]$, $x \perp\!\!\!\perp y$.

2.1.2 Common Probability Distributions

Bernoulli Bernoulli distribution describes situation where only two possible outcomes $y = 0/y = 1$ or failure/success.

1. $P(x) = \mathbf{Bern}_x[\lambda] = \lambda^x(1 - \lambda)^{1-x}$
2. univariate, discrete, binary
3. $x \in \{0, 1\}$; $\lambda \in [0, 1]$
4. $\mathbf{E}[x] = \lambda$, $\mathbf{Var}[x] = \lambda(1 - \lambda)$

Beta Beta distribution is the conjugate distribution to Bernoulli distribution.

1. $P(\lambda) = \mathbf{Beta}_\lambda[\alpha, \beta] = \frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)}\lambda^{\alpha-1}(1 - \lambda)^{\beta-1}$
2. univariate, continuous, unbounded
3. $\lambda \in \mathbb{R}$; $\alpha \in \mathbb{R}_+$, $\beta \in \mathbb{R}_+$
4. $\mathbf{E}[\lambda] = \frac{\alpha}{\alpha+\beta}$, $\mathbf{Var}[\lambda] = \frac{\alpha\beta}{(\alpha+\beta)^2(\alpha+\beta+1)}$

Categorical Categorical distribution describes situation with K possible outcomes.

1. $P(x) = \mathbf{Cat}_x[\boldsymbol{\lambda}]$, $P(x = k) = \lambda_k$, $P(\mathbf{x} = \mathbf{e}_k) = \prod_{j=1}^K \lambda_j^{x_j} = \lambda_k$
2. univariate, discrete, multi-valued
3. $x \in \{1, 2, \dots, K\}$; $\lambda_k \in [0, 1]$ where $\sum_k \lambda_k = 1$
4. $\mathbf{E}[x_i] = \lambda_i$, $\mathbf{Var}[x_i] = \lambda_i(1 - \lambda_i)$, $\mathbf{Cov}[x_i, x_j] = -\lambda_i\lambda_j$ ($i \neq j$)

Dirichlet Dirichlet distribution is the conjugate distribution to categorical distribution.

1. $P(\boldsymbol{\lambda}) = \mathbf{Dir}_\lambda[\boldsymbol{\alpha}] = \frac{\Gamma(\sum_{k=1}^K \alpha_k)}{\prod_{k=1}^K \Gamma(\alpha_k)} \prod_{k=1}^K \lambda_k^{\alpha_k-1}$
2. multivariate, continuous, bounded, sums to one
3. $\boldsymbol{\lambda} = [\lambda_1, \lambda_2, \dots, \lambda_K]^\top$, $\lambda_k \in [0, 1]$, $\sum_{k=1}^K \lambda_k = 1$; $\boldsymbol{\alpha} = [\alpha_1, \alpha_2, \dots, \alpha_K]$, $\alpha_k \in \mathbb{R}_+$
4. $\mathbf{E}[\lambda_i] = \frac{\alpha_i}{\sum_k \alpha_k}$, $\mathbf{Var}[\lambda_i] = \frac{\alpha_i(\sum_k \alpha_k - \alpha_i)}{(\sum_k \alpha_k)^2(\sum_k \alpha_k + 1)}$, $\mathbf{Cov}[\lambda_i, \lambda_j] = \frac{-\alpha_i\alpha_j}{(\sum_k \alpha_k)^2(\sum_k \alpha_k + 1)}$ ($i \neq j$)

Univariate Normal Univariate normal distribution describes single continuous variable.

1. $P(x) = \mathbf{Norm}_x[\mu, \sigma^2] = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(x-\mu)^2}{2\sigma^2}\right)$
2. univariate, continuous, unbounded
3. $x \in \mathbb{R}; \mu \in \mathbb{R}, \sigma^2 \in \mathbb{R}_+$
4. $\mathbf{E}[x] = \mu, \mathbf{Var}[x] = \sigma^2$

Normal Inverse Gamma Normal inverse gamma distribution is a conjugate distribution to univariate normal distribution.

1. $P(\mu, \sigma^2) = \mathbf{NormInvGam}_{\mu, \sigma^2}[\alpha, \beta, \gamma, \delta] = \frac{\sqrt{\gamma}}{\sqrt{2\pi\sigma^2}} \frac{\beta^\alpha}{\Gamma(\alpha)} \left(\frac{1}{\sigma^2}\right)^{\alpha+1} \exp\left(-\frac{2\beta+\gamma(\delta-\mu)^2}{2\sigma^2}\right)$
2. bivariate, continuous, μ unbounded, σ^2 bounded below
3. $\mu \in \mathbb{R}, \sigma^2 \in \mathbb{R}_+; \alpha \in \mathbb{R}_+, \beta \in \mathbb{R}_+, \gamma \in \mathbb{R}_+, \delta \in \mathbb{R}$
4. $\mathbf{E}[\mu] = \delta, \mathbf{E}[\sigma^2] = \frac{\beta}{\alpha-1} \ (\alpha > 1), \mathbf{Var}[\mu] = \frac{\beta}{(\alpha-1)\gamma} \ (\alpha > 1), \mathbf{Var}[\sigma^2] = \frac{\beta^2}{(\alpha-1)^2(\alpha-2)} \ (\alpha > 2), \mathbf{Cov}[\mu, \sigma^2] = 0 \ (\alpha > 1)$

Multivariate Normal Multivariate normal distribution describes multiple continuous variables. It takes two parameters: a vector containing mean position $\boldsymbol{\mu}$, and a symmetric positive definite covariance matrix $\boldsymbol{\Sigma}$.

1. $P(\mathbf{x}) = \mathbf{Norm}_x[\boldsymbol{\mu}, \boldsymbol{\Sigma}] = \frac{1}{\sqrt{\det(2\pi\boldsymbol{\Sigma})}} \exp\left(-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu})^\top \boldsymbol{\Sigma}^{-1}(\mathbf{x} - \boldsymbol{\mu})\right)$
2. multivariate, continuous, unbounded
3. $\mathbf{x} \in \mathbb{R}^K; \boldsymbol{\mu} \in \mathbb{R}^K, \boldsymbol{\Sigma} \in \mathbb{R}^{K \times K}$ (positive semi-definite matrix)
4. $\mathbf{E}[\mathbf{x}] = \boldsymbol{\mu}, \mathbf{Var}[\mathbf{x}] = \boldsymbol{\Sigma}$

Normal Inverse Wishart Normal inverse wishart distribution is a conjugate distribution to multivariate normal distribution.

1. $P(\boldsymbol{\mu}, \boldsymbol{\Sigma}) = \mathbf{NormInvWis}_{\boldsymbol{\mu}, \boldsymbol{\Sigma}}[\alpha, \boldsymbol{\Psi}, \gamma, \boldsymbol{\delta}]$
 $= \frac{\gamma^{D/2} |\boldsymbol{\Psi}|^{\alpha/2} |\boldsymbol{\Sigma}|^{-\frac{\alpha+D+2}{2}}}{(2\pi)^{D/2} 2^{(\alpha\boldsymbol{\Sigma})/2} \Gamma_D(\alpha/2)} \exp\left(-\frac{1}{2}(\text{Tr}(\boldsymbol{\Psi}\boldsymbol{\Sigma}^{-1}) + \gamma(\boldsymbol{\mu} - \boldsymbol{\delta})^\top \boldsymbol{\Sigma}^{-1}(\boldsymbol{\mu} - \boldsymbol{\delta}))\right)$
2. multivariate, $\boldsymbol{\mu}$ unbounded, $\boldsymbol{\Sigma}$ square, positive definite
3. $\boldsymbol{\mu} \in \mathbb{R}^K, \boldsymbol{\Sigma} \in \mathbb{R}^{K \times K}; \alpha \in \mathbb{R}_{>D-1}, \boldsymbol{\Psi} \in \mathbb{R}^{K \times K}, \gamma \in \mathbb{R}_+, \boldsymbol{\delta} \in \mathbb{R}^K$

2.2 Linear Algebra

2.3 Calculus

2.4 Informatics

2.5 Optimization

Chapter 3

Machine Learning Basics

3.1 Regularization

3.1.1 Under-fitting & Over-fitting

Under-fitting If $N > D$ (e.g. 30 data points, 2 dimensions) we have more equations than unknowns: over-determined system. Input-output relations can only hold approximately.

Over-fitting If $N < D$ (e.g. 30 points, 15265 dimensions) we have more unknowns than equations: under-determined system. Input-output equations hold exactly, but we are simply memorizing data.

3.1.2 Bias & Variance

High Bias & Low Variance A rigid model's (low complexity) performance is more predictable in the test set but the model may not be good even on the training set.

Low Bias & High Variance A flexible model (high complexity) approximates the target function well in the training set but can "overtrain" and have poor performance on the test set.

3.1.3 Vector Norm

L1, ("Manhattan") norm $\|\mathbf{w}\|_1 = \sum_{d=1}^D |w_d|$

L2, ("Euclidean") norm $\|\mathbf{w}\|_2 = \sqrt{\sum_{d=1}^D w_d^2} = \sqrt{\langle \mathbf{w}, \mathbf{w} \rangle} = \sqrt{\mathbf{w}^\top \mathbf{w}}$

Lp norm, $p > 1$ $\|\mathbf{w}\|_p = \left(\sum_{d=1}^D w_d^p \right)^{1/p}$

3.1.4 Penalize Complexity

In linear regression, the residual vector is $\epsilon = \mathbf{y} - \Psi\mathbf{w}$. The loss function is $L(\mathbf{w}) = \epsilon^\top \epsilon$. We add a complexity term $R(\mathbf{w}) = \|\mathbf{w}\|_2^2 = \mathbf{w}^\top \mathbf{w}$ to the loss function. Hence, the original loss function becomes $L(\mathbf{w}) = \epsilon^\top \epsilon + \lambda \mathbf{w}^\top \mathbf{w}$.

Without regularization, the loss function is $L(\mathbf{w}) = \epsilon^\top \epsilon$. Let $\nabla L(\mathbf{w}^*) = 0$, we have $\mathbf{w}^* = (\mathbf{X}^\top \mathbf{X})^{-1} \mathbf{X}^\top \mathbf{y}$.

With L2-regularization, the loss function is $L(\mathbf{w}) = \epsilon^\top \epsilon + \lambda \mathbf{w}^\top \mathbf{w}$. Let $\nabla L(\mathbf{w}^*) = 0$, we have $\mathbf{w}^* = (\mathbf{X}^\top \mathbf{X} + \lambda \mathbf{I})^{-1} \mathbf{X}^\top \mathbf{y}$. The additional $\lambda \mathbf{I}$ makes the data matrix more robust to calculate inversion.

3.2 Cross-Validation

We can select hyperparameters with (cross-)validation. Cross-validation excludes part of the training data from parameter estimation, and use them only to predict the test error.

K-fold cross validation: split data set into K folds and each time train on (K-1) folds and valid on the remaining fold until all folds have been used as validation fold. The cross-validation error is the average of K validation errors. We pick hyperparameters that minimize cross-validation error.

3.3 Bayesian Learning

3.3.1 Bayes' Rule Terminology

Bayes' Rule:

$$P(y|x) = \frac{P(x|y) P(y)}{\int P(x|y) P(y) dy}$$

Prior $P(y)$ what we know about y before seeing x . In parameters learning we choose prior that is conjugate to likelihood.

Likelihood $P(x|y)$ propensity for observing a certain value of x given a certain value of y .

Posterior $P(y|x)$ what we know about y after seeing x . Posterior must have same form as conjugate prior distribution.

Evidence $\int P(x|y) P(y) dy$ a constant to ensure that the LHS is a valid distribution. Posterior must be a distribution which implies that evidence equals to a constant κ from conjugate relation.

3.3.2 Maximum Likelihood

Fitting As the name suggests we find the parameters under which the data $\mathbf{x}_{1...I}$ are most likely. Here, we have assumed that data was independent.

$$\begin{aligned}\hat{\theta} &= \operatorname{argmax}_{\theta} P(\mathbf{x}_{1...I}|\theta) \\ &= \operatorname{argmax}_{\theta} \prod_{i=1}^I P(\mathbf{x}_i|\theta)\end{aligned}$$

Predictive Density Evaluate new data point \mathbf{x}^* under probability distribution $P(\mathbf{x}^*|\hat{\theta})$ with best parameters.

3.3.3 Maximum a Posterior (MAP)

Fitting As the name suggests we find the parameters which maximize the posterior probability $P(\theta|\mathbf{x}_{1...I})$. Again we have assumed that data was independent.

$$\begin{aligned}\hat{\theta} &= \operatorname{argmax}_{\theta} P(\theta|\mathbf{x}_{1...I}) \\ &= \operatorname{argmax}_{\theta} \frac{P(\mathbf{x}_{1...I}|\theta) P(\theta)}{P(\mathbf{x}_{1...I})} \\ &= \operatorname{argmax}_{\theta} \frac{\prod_{i=1}^I P(\mathbf{x}_i|\theta) P(\theta)}{P(\mathbf{x}_{1...I})}\end{aligned}$$

Since the denominator does not depend on the parameters we can instead maximize

$$\hat{\theta} = \operatorname{argmax}_{\theta} \prod_{i=1}^I P(\mathbf{x}_i|\theta) P(\theta)$$

Predictive Density Evaluate new data point \mathbf{x}^* under probability distribution with MAP parameters $P(\mathbf{x}^*|\hat{\theta})$

3.3.4 Bayesian Approach

Fitting Compute the posterior distribution over possible parameter values using Bayes' rule. Principle: There are many values that could have explained the data. Instead of picking one set of parameters, try to capture all of the possibilities.

$$P(\theta|\mathbf{x}_{1...I}) = \frac{\prod_{i=1}^I P(\mathbf{x}_i|\theta) P(\theta)}{P(\mathbf{x}_{1...I})}$$

Predictive Density (a) Each possible parameter value makes a prediction. (b) Some parameters more probable than others.

$$P(\mathbf{x}^*|\mathbf{x}_{1...I}) = \int P(\mathbf{x}^*|\boldsymbol{\theta}) P(\boldsymbol{\theta}|\mathbf{x}_{1...I}) d\boldsymbol{\theta}$$

Make a prediction that is an infinite weighted sum (integral) of the predictions for each parameter value ($P(\mathbf{x}^*|\boldsymbol{\theta})$), where weights are the probabilities ($P(\boldsymbol{\theta}|\mathbf{x}_{1...I})$).

3.3.5 Example: Univariate Normal Distribution

Maximum Likelihood

Likelihood given by normal distribution pdf:

$$P(x|\mu, \sigma^2) = \mathbf{Norm}_x[\mu, \sigma^2] = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(x - \mu)^2}{2\sigma^2}\right)$$

Apply maximum likelihood:

$$\begin{aligned} \hat{\mu}, \hat{\sigma}^2 &= \operatorname{argmax}_{\mu, \sigma^2} P(\mathbf{x}_{1...I}|\mu, \sigma^2) \\ &= \operatorname{argmax}_{\mu, \sigma^2} \prod_{i=1}^I P(x_i|\mu, \sigma^2) \\ &= \operatorname{argmax}_{\mu, \sigma^2} \prod_{i=1}^I \mathbf{Norm}_{x_i}[\mu, \sigma^2] \\ &= \operatorname{argmax}_{\mu, \sigma^2} \sum_{i=1}^I \log \mathbf{Norm}_{x_i}[\mu, \sigma^2] \\ &= \operatorname{argmax}_{\mu, \sigma^2} \left(-\frac{I}{2} \log 2\pi - \frac{I}{2} \log \sigma^2 - \frac{1}{2} \sum_{i=1}^I \frac{(x_i - \mu)^2}{\sigma^2} \right) \end{aligned}$$

Let $\nabla L(\hat{\mu}, \hat{\sigma}^2) = 0$, we have the solution:

$$\begin{aligned} \hat{\mu} &= \frac{\sum_{i=1}^I x_i}{I} \\ \hat{\sigma}^2 &= \sum_{i=1}^I \frac{(x_i - \hat{\mu})^2}{I} \end{aligned}$$

Maximum a Posterior

Likelihood given by normal distribution pdf:

$$P(x|\mu, \sigma^2) = \mathbf{Norm}_x[\mu, \sigma^2] = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(x - \mu)^2}{2\sigma^2}\right)$$

Prior given by normal inverse gamma distribution pdf:

$$P(\mu, \sigma^2) = \mathbf{NormInvGam}_{\mu, \sigma^2}[\alpha, \beta, \gamma, \delta] = \frac{\sqrt{\gamma}}{\sqrt{2\pi\sigma^2}} \frac{\beta^\alpha}{\Gamma(\alpha)} \left(\frac{1}{\sigma^2}\right)^{\alpha+1} \exp\left(-\frac{2\beta + \gamma(\delta - \mu)^2}{2\sigma^2}\right)$$

Apply maximum a posterior:

$$\begin{aligned} \hat{\mu}, \hat{\sigma}^2 &= \underset{\mu, \sigma^2}{\operatorname{argmax}} \prod_{i=1}^I P(x_i | \mu, \sigma^2) P(\mu, \sigma^2) \\ &= \underset{\mu, \sigma^2}{\operatorname{argmax}} \prod_{i=1}^I \mathbf{Norm}_{x_i}[\mu, \sigma^2] \mathbf{NormInvGam}_{\mu, \sigma^2}[\alpha, \beta, \gamma, \delta] \\ &= \underset{\mu, \sigma^2}{\operatorname{argmax}} \left(\sum_{i=1}^I \log \mathbf{Norm}_{x_i}[\mu, \sigma^2] + \log \mathbf{NormInvGam}_{\mu, \sigma^2}[\alpha, \beta, \gamma, \delta] \right) \end{aligned}$$

Let $\nabla L(\hat{\mu}, \hat{\sigma}^2) = 0$, we have the solution:

$$\begin{aligned} \hat{\mu} &= \frac{\sum_{i=1}^I x_i + \gamma\delta}{I + \gamma} \\ \hat{\sigma}^2 &= \frac{\sum_{i=1}^I (x_i - \mu)^2 + 2\beta + \gamma(\delta - \mu)^2}{I + 3 + 2\alpha} \end{aligned}$$

Bayesian Approach

Compute the posterior distribution using Bayes' rule:

$$\begin{aligned} P(\mu, \sigma^2 | x_{1...I}) &= \frac{\prod_{i=1}^I P(x_i | \mu, \sigma^2) P(\mu, \sigma^2)}{P(x_{1...I})} \\ &= \frac{\prod_{i=1}^I \mathbf{Norm}_{x_i}[\mu, \sigma^2] \mathbf{NormInvGam}_{\mu, \sigma^2}[\alpha, \beta, \gamma, \delta]}{P(x_{1...I})} \\ &= \frac{\kappa(\alpha, \beta, \gamma, \delta, x_{1...I}) \mathbf{NormInvGam}_{\mu, \sigma^2}[\tilde{\alpha}, \tilde{\beta}, \tilde{\gamma}, \tilde{\delta}]}{P(x_{1...I})} \\ &= \mathbf{NormInvGam}_{\mu, \sigma^2}[\tilde{\alpha}, \tilde{\beta}, \tilde{\gamma}, \tilde{\delta}] \end{aligned}$$

where

$$\begin{aligned} \tilde{\alpha} &= \alpha + \frac{I}{2} \\ \tilde{\beta} &= \frac{\sum_i x_i^2}{2} + \beta + \frac{\gamma\delta^2}{2} - \frac{(\gamma\delta + \sum_i x_i)^2}{2(\gamma + I)} \\ \tilde{\gamma} &= \gamma + I \\ \tilde{\delta} &= \frac{\gamma\delta + \sum_i x_i}{\gamma + I} \end{aligned}$$

Take weighted sum of predictions from different parameter values:

$$\begin{aligned}
P(x^*|x_{1...I}) &= \iint P(x^*|\mu, \sigma^2) P(\mu, \sigma^2|x_{1...I}) d\mu d\sigma \\
&= \iint \mathbf{Norm}_{x^*}[\mu, \sigma^2] \mathbf{NormInvGam}_{\mu, \sigma^2}[\tilde{\alpha}, \tilde{\beta}, \tilde{\gamma}, \tilde{\delta}] d\mu d\sigma \\
&= \iint \kappa(\alpha, \beta, \gamma, \delta, x_{1...I}) \mathbf{NormInvGam}_{\mu, \sigma^2}[\check{\alpha}, \check{\beta}, \check{\gamma}, \check{\delta}] d\mu d\sigma \\
&= \kappa(\alpha, \beta, \gamma, \delta, x_{1...I}) \iint \mathbf{NormInvGam}_{\mu, \sigma^2}[\check{\alpha}, \check{\beta}, \check{\gamma}, \check{\delta}] d\mu d\sigma \\
&= \kappa(\alpha, \beta, \gamma, \delta, x_{1...I}) \\
&= \frac{1}{\sqrt{2\pi}} \frac{\sqrt{\tilde{\gamma}} \tilde{\beta}^{\tilde{\alpha}} \Gamma(\check{\alpha})}{\sqrt{\check{\gamma}} \check{\beta}^{\check{\alpha}} \Gamma(\check{\alpha})}
\end{aligned}$$

where

$$\begin{aligned}
\check{\alpha} &= \tilde{\alpha} + \frac{1}{2} \\
\check{\beta} &= \frac{x^{*2}}{2} + \tilde{\beta} + \frac{\tilde{\gamma}^{\tilde{\delta}^2}}{2} - \frac{(\tilde{\gamma}^{\tilde{\delta}} + x^*)^2}{2(\tilde{\gamma} + 1)} \\
\check{\gamma} &= \tilde{\gamma} + 1
\end{aligned}$$

3.3.6 Example: Categorical Distribution

Maximum Likelihood

Likelihood given by categorical distribution pdf:

$$P(x|\boldsymbol{\lambda}) = \mathbf{Cat}_x[\boldsymbol{\lambda}] = \prod_{j=1}^K \lambda_j^{x_j} = \lambda_k$$

Apply maximum likelihood:

$$\begin{aligned}
\hat{\boldsymbol{\lambda}} &= \underset{\boldsymbol{\lambda}}{\operatorname{argmax}} \prod_{i=1}^I P(x_i|\boldsymbol{\lambda}) & s.t. \sum_k \lambda_k &= 1 \\
&= \underset{\boldsymbol{\lambda}}{\operatorname{argmax}} \prod_{i=1}^I \mathbf{Cat}_{x_i}[\boldsymbol{\lambda}] & s.t. \sum_k \lambda_k &= 1 \\
&= \underset{\boldsymbol{\lambda}}{\operatorname{argmax}} \prod_{k=1}^K \lambda_k^{N_k} & s.t. \sum_k \lambda_k &= 1 \\
&= \underset{\boldsymbol{\lambda}}{\operatorname{argmax}} \sum_{k=1}^K N_k \log \lambda_k & s.t. \sum_k \lambda_k &= 1
\end{aligned}$$

Here, N_k represents the number of times the data is classified in class k . As before, we will instead optimize log probability. Since there is a constraint *s.t.* $\sum_k \lambda_k = 1$, we use Lagrange multiplier to reconstruct the loss function.

$$L(\boldsymbol{\lambda}) = \sum_{k=1}^K N_k \log \lambda_k + v \left(\sum_{k=1}^K \lambda_k - 1 \right)$$

Let $\nabla L(\boldsymbol{\lambda}, v) = 0$, we have the solution:

$$\hat{\lambda}_k = \frac{N_k}{\sum_{m=1}^K N_m}$$

Maximum a Posterior

Bayesian Approach

3.4 Machine Learning Models

Chapter 4

Regression

4.1 Linear Regression

4.2 Non-linear Regression

4.3 Logistic Regression

Chapter 5

Support Vector Machines

Bibliography

Appendix A

Test

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