Foundation of Analytics: Lecture 3

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Content

- Random Variables: Dependent, Independent, Correlation
- Linear Regression of One Variable
- Linear Regression of Multiple Variables
- Logistic Regression

Let's look at a few pairs of data points?

•
$$\vec{x} = [0.5, 0.6, 0.1, -0.3, 2.3], \vec{y} = [0.5, 0.6, 0.1, -0.3, 2.3]$$

•
$$\vec{x} = [0.5, 0.6, 0.1, -0.3, 2.3], \vec{y} = [0.6, 0.6, 0.12, -0.3, 2.3]$$

•
$$\vec{x} = [0.5, 0.6, 0.1, -0.3, 2.3], \vec{y} = [0.02, -0.2, 0.2, 2.1, -0.5]$$

What can you tell about the relationship between \vec{x} and \vec{x} ?

Given two random variables X and Y, denote the mean and variance of the two variables as $E[X] = \mu_X$, $E[Y] = \mu_Y$, $Var[X] = \sigma_X^2$, $Var[Y] = \sigma_Y^2$.

The covariance of X and Y is the number defined by

$$Cov(X,Y) = E[(X - \mu_X)(Y - \mu_Y)]$$

= $E[XY] - \mu_X \mu_Y$

Empricial Estimation of Covariance

$$Cov(X,Y) = \frac{(x - \mu_x)(y^T - \mu_y)}{N}$$
 (empirical)
$$Cov(X,Y) = \frac{(x - \mu_x)(y^T - \mu_y)}{N - 1}$$
 (unbiased)

The correlation of the two random variables is the number defined by

$$\rho_{XY} = \frac{Cov(X,Y)}{\sigma_X \sigma_Y}$$

Calculate the covariance/correlation of

• Example 1:

$$\vec{x} = [2, -2, -2, 2], \vec{y} = [2, -2, -2, 2]$$

We have $\mu_x=0$, $\mu_y=0$, $\sigma_x^2=4$, $\sigma_y^2=4$, E[XY]=4 Therefore Cov(X,Y)=4-0=4 and $\rho_{xy}=4/(2*2)=1$

Example 2:

$$\vec{x} = [2, -2, -2, 2], \vec{y} = [2, 0, -2, 0]$$

We have $\mu_x=0$, $\mu_y=0$, $\sigma_x^2=4$, $\sigma_y^2=2$, E[XY]=2 Therefore Cov(X,Y)=2-0=2 and $\rho_{xy}=2/(2*\sqrt{2})=1/\sqrt{2}$

Linear Regression with One Variable

Data set:

$$y = \begin{bmatrix} y^1 \\ y^2 \\ \vdots \\ y^n \end{bmatrix}, X = \begin{bmatrix} x^1 \\ x^2 \\ \vdots \\ x^n \end{bmatrix}$$

Linear Regression with One Variable

Assume y is linearly depending on x i.e.

$$\hat{y} = \beta_0 + \beta_1 x$$

Find $\hat{\beta}$ that minimize the estimation error

$$\epsilon = \sum_{i=1}^{n} (y^i - \hat{y}^i)^2 = \sum_{i=1}^{n} (y^i - \beta_0 - \beta_1 x^i)^2$$

i.e.

$$\frac{\partial \epsilon}{\partial \beta_1} = 0 \to \sum_{i=1}^n (y^i - \beta_0 - \beta_1 x^i) x^i = 0$$

$$\frac{\partial \epsilon}{\partial \beta_0} = 0 \rightarrow \sum_{i=1}^n (y^i - \beta_0 - \beta_1 x^i) = 0$$

$$\beta_0 \sum_{i=1}^n x^i = \sum_{i=1}^n y^i x^i - \beta_1 \sum_{i=1}^n x^i x^i$$

$$\beta_0 = \frac{1}{n} \sum_{i=1}^{n} (y^i - \beta_1 x^i) = \bar{y} - \beta_1 \bar{x}$$

Insert the second equation to the first, we have

$$n\bar{x}\bar{y} - \beta_1 n\bar{x}\bar{x} = \sum_{i=1}^{n} y^i x^i - \beta_1 \sum_{i=1}^{n} x^i x^i$$

Therefore,

$$\beta_1 = \frac{\frac{1}{n} \sum_{i=1}^{n} x^i y^i - \bar{x}\bar{y}}{\frac{1}{n} \sum_{i=1}^{n} x^i x^i - \bar{x}^2} = \frac{Cov(X,Y)}{Var(X)} = \rho_{XY} \frac{\sigma_Y}{\sigma_X}$$

Data set:

$$\begin{bmatrix} y, X \end{bmatrix} = \begin{bmatrix} y^1 & x_0^1 & x_1^1 & x_2^1 & \dots & x_m^1 \\ y^2 & x_0^2 & x_1^2 & x_2^2 & \dots & x_m^2 \\ \vdots & \vdots & \vdots & \vdots & \ddots & \vdots \\ y^n & x_0^n & x_1^n & x_2^n & \dots & x_m^n \end{bmatrix}$$

Assume y is a linear superposition of multiple x's

$$\hat{y} = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \ldots + \beta_m x_m$$

or simply

$$\hat{y} = \sum_{j=1}^{m} \beta_j x_j$$

Estimate β 's that best fits the data, we need to minimize the error

$$\epsilon = \sum_{i=1}^{n} (y^i - \hat{y}^i)^2$$
$$= (y - \hat{y})^T (y - \hat{y})$$

Use basic calculus we know, we want to have the β s satisfy the following equation set:

$$\frac{\partial \epsilon}{\partial \beta_i} = 0, j = 1, 2, 3, 4...m$$

i.e.

$$\sum_{i=1}^{n} \frac{\partial (y^{i} - \hat{y}^{i})^{2}}{\partial \beta_{j}} = 0$$

$$\sum_{i=1}^{n} (y^{i} - \hat{y}^{i}) \frac{\partial \hat{y}^{i}}{\partial \beta_{j}} = 0$$

$$\sum_{i=1}^{n} (y^{i} - \hat{y}^{i}) x_{j}^{i} = 0$$

Written in matrix formula we require

$$(y - X\beta)^T X = 0$$

or after transposing

$$X^T y - X^T X \beta = \mathbf{0}$$

Therefore

$$\beta = (X^T X)^{-1} X^T y$$

Logistic Regression: Likelihood Function

Assuming two possible outcomes 1 and 0, the probability of being 1 is modeled as

$$p_i = \frac{1}{1 + \exp(-\vec{\beta} \cdot \vec{x}^i)}$$

The likelihood function is defined as

$$Likelihood = \prod_{i=1}^{n} p_i^{y^i} (1 - p_i)^{1-y^i}$$

The log-likelihood function is the defined as the log transformation of the likelihood function

$$\ell = \log(Likelihood) = \sum_{i=1}^{n} y^{i} \log(p_{i}) + (1 - y^{i}) \log(1 - p_{i})$$



Logistic Regression: Optimization Attempt

It follows that

$$\ell = \sum_{i=1}^{n} y^{i} \log \frac{p_{i}}{1 - p_{i}} + \log(1 - p_{i})$$

$$= \sum_{i=1}^{n} y^{i} (\vec{\beta} \cdot \vec{x}^{i}) - \log(1 + \exp(\vec{\beta} \cdot \vec{x}^{i})))$$

Take the gradient against β s, we have

$$\frac{\partial \ell}{\partial \beta_j} = \sum_{i=1}^n \left(y^i - \frac{1}{1 + \exp(-\vec{\beta} \cdot \vec{x}^i)} \right) x_j^i, j = 1, 2, 3, ..., m$$

 β s can NOT be solved by setting $\nabla \ell = 0$ because of the nonlinear term of x^i , which is $\frac{1}{1+\exp(\vec{x}^i\cdot\vec{\beta})}$.



Newton-Raphson Method for Optimizing Non-linear Functions

Consider a function of one parameter $\ell(\beta)$ and assume β_0 is close to the point that minimizes $\ell(\beta)$. We can therefore use Talyor expansion for approximation

$$\ell(\beta) = \ell(\beta_0) + \ell'(\beta_0)(\beta - \beta_0) + \frac{1}{2}\ell''(\beta_0)(\beta - \beta_0)^2$$

The β^* that minimize the function have derivative at the point 0 i.e. $\ell'(\beta)|_{\beta=\beta^*}=0$, by setting $\ell'(\beta)=0$, we get an iterative evaluation methods for β^*

$$\ell'(\beta_{0}) + \frac{1}{2} 2\ell''(\beta_{0})(\beta - \beta_{0}) = 0 \to \beta = \beta_{0} - \frac{\ell'(\beta_{0})}{\ell''(\beta_{0})}$$
i.e.
$$\beta^{(k+1)} = \beta^{(k)} - \frac{\ell'(\beta^{(k)})}{\ell''(\beta^{(k)})}$$

Multivariate Newton-Raphson Method

For multivarite function, the iteration formula becomes

$$\beta^{(k+1)} = \beta^{(k)} - H^{-1}(\beta^{(k)}) \nabla \ell(\beta^{(k)})$$

here $H(\beta^{(k)})$ is the Hessian matrix of $\ell(\beta)$ evaluated at $\beta = \beta^{(k)}$, defined as

$$H_{ab} = \frac{\partial^2 \ell}{\partial \beta_a \partial \beta_b} |_{\beta = \beta^{(k)}}$$

and $H^{-1}(\beta^{(k)})$ is the inverse of $H(\beta^{(k)})$



Logistic Regression

Apply Newton-Raphson methods to optimize the logistic regression, we calculate the Hessian of the log-likelihood function

$$\frac{\partial^2 \ell}{\partial \beta_a \partial \beta_b} = -\sum_{i=1}^n x_b^i \frac{\exp(-\vec{\beta} \cdot \vec{x}^i)}{(1 + \exp(-\vec{\beta} \cdot \vec{x}^i))^2} x_a^i$$
$$= -\sum_{i=1}^n x_b^i p_i (1 - p_i) x_a^i$$

written in matrix formula, the Hessian of the loglikelihood function is

$$H = -X^T W X$$
, $W = \begin{bmatrix} p_1(1-p_1) & & & \\ & \ddots & & \\ & & p_n(1-p_n) \end{bmatrix}$

Logistic Regression: Optimization Algorithm

Use Newton Raphson Methods, we have

$$\vec{\beta}^{(k+1)} \leftarrow \vec{\beta}^{(k)} - H^{-1} \nabla \ell$$
$$\vec{\beta}^{(k+1)} \leftarrow \vec{\beta}^{(k)} + (X^T W X)^{-1} X^T (y - p)$$

Recall in linear regression case

$$\beta = (X^T X)^{-1} X^T y$$

