

# Lecture Notes - 03: Linear Regression

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## 1 Single Variate Linear Regression

A simple linear model can be formulated by assuming the target variable is dependent only on one predictor i.e.

$$\hat{y} = \beta_0 + \beta_1 x$$

In order to have our estimate as close as to the actual value of  $y$ , we want to find the  $\beta$ s that minimize the sum squared error function

$$\epsilon = \sum_{i=1}^n (y^i - \hat{y}^i)^2 = \sum_{i=1}^n (y^i - \beta_0 - \beta_1 x^i)^2 \quad (1)$$

i.e.

$$\begin{cases} \frac{\partial \epsilon}{\partial \beta_1} = 0 \\ \frac{\partial \epsilon}{\partial \beta_0} = 0 \end{cases} \quad (2)$$

$$\Rightarrow \begin{cases} \sum_{i=1}^n (y^i - \beta_0 - \beta_1 x^i) x^i = 0 \\ \sum_{i=1}^n (y^i - \beta_0 - \beta_1 x^i) = 0 \end{cases} \quad (3)$$

Sorting the equations to get the solution for  $\beta_0$  and  $\beta_1$

$$\Rightarrow \begin{cases} \beta_0 \sum_{i=1}^n x^i = \sum_{i=1}^n y^i x^i - \beta_1 \sum_{i=1}^n x^i x^i \\ \sum_{i=1}^n \beta_0 = \sum_{i=1}^n y^i - \beta_1 \sum_{i=1}^n x^i \end{cases}$$

$$\Rightarrow \begin{cases} \beta_1 = \frac{\sum_{i=1}^n y^i x^i - \beta_0 \sum_{i=1}^n x^i}{\sum_{i=1}^n x^i x^i} = \frac{\frac{1}{n} \sum_{i=1}^n y^i x^i - \beta_0 \bar{x}}{\frac{1}{n} \sum_{i=1}^n x^i x^i} \\ \beta_0 = \frac{1}{n} \left( \sum_{i=1}^n y^i - \beta_1 \sum_{i=1}^n x^i \right) = (\bar{y} - \beta_1 \bar{x}) \end{cases} \quad (4)$$

Substitute  $\beta_0$  in to the first equation in equation set (4). We have

$$\beta_1 = \frac{\frac{1}{n} \sum_{i=1}^n y^i x^i - (\bar{y} \bar{x} - \beta_1 \bar{x} \bar{x})}{\frac{1}{n} \sum_{i=1}^n x^i x^i}$$

Solving for  $\beta_1$  we have

$$\begin{aligned} \beta_1 \frac{1}{n} \sum_{i=1}^n x^i x^i &= \frac{1}{n} \sum_{i=1}^n y^i x^i - \bar{y} \bar{x} + \beta_1 \bar{x} \bar{x} \\ \Rightarrow \beta_1 \frac{1}{n} \sum_{i=1}^n x^i x^i &= \frac{1}{n} \sum_{i=1}^n y^i x^i - \bar{y} \bar{x} + \beta_1 \bar{x} \bar{x} \\ \Rightarrow \beta_1 \left( \frac{1}{n} \sum_{i=1}^n x^i x^i - \bar{x} \bar{x} \right) &= \frac{1}{n} \sum_{i=1}^n y^i x^i - \bar{y} \bar{x} \end{aligned}$$

Thus we get the solution of  $\beta_1$

$$\beta_1 = \frac{\frac{1}{n} \sum_{i=1}^n y^i x^i - \bar{y} \bar{x}}{\frac{1}{n} \sum_{i=1}^n x^i x^i - \bar{x} \bar{x}} \quad (5)$$

Take a close look it is not hard to find that the numerator is the co-variance of  $X$  and  $Y$ . The denominator is the variance of  $X$ .

Therefore  $\beta_1$  can also be written as

$$\beta_1 = \frac{Cov(X, Y)}{Var(X)} = \rho_{XY} \frac{\sigma_X}{\sigma_Y} \quad (6)$$

## 2 Multivariate Linear Regression

Assume  $y$  is a linear superposition of multiple  $x$ s, the model for  $y$  is then formulated as

$$\hat{y} = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \dots + \beta_m x_m$$

or simply

$$\hat{y} = \sum_{j=1}^m \beta_j x_j = \vec{\beta} \cdot \vec{x}$$

To estimate  $\beta$ s that best fit the data, we need to minimize the error

$$\epsilon = \sum_{i=1}^n (y^i - \hat{y}^i)^2 \quad (7)$$

$$= \sum_{i=1}^n (y^i - \vec{x}^i \cdot \vec{\beta})^2 \quad (8)$$

Writing in matrix notation, we have

$$\begin{aligned} \epsilon &= (y - \hat{y})^T (y - \hat{y}) \\ &= (y - X\beta)^T (y - X\beta) \end{aligned}$$

here  $\beta$  is a  $m \times 1$  matrix defined as

$$\beta = \begin{bmatrix} \beta_1 \\ \beta_2 \\ \vdots \\ \beta_m \end{bmatrix}$$

To minimize the error  $\epsilon$  we want to the  $\beta$ s satisfy the following equation set:

$$\frac{\partial \epsilon}{\partial \beta_j} = 0, j = 1, 2, 3, 4 \dots m$$

Using equation (7), we have

$$\sum_{i=1}^n \frac{\partial}{\partial \beta_j} (y^i - \hat{y}^i) = 0$$

$$\sum_{i=1}^n (y^i - \hat{y}^i) \frac{\partial \hat{y}^i}{\partial \beta_j} = 0 \quad (9)$$

$$\Rightarrow \sum_{i=1}^n (y^i - \hat{y}^i) x_j^i = 0 \quad (10)$$

Going from equation (9) to equation (10), we use the fact that  $\hat{y} = \vec{x} \cdot \vec{\beta} = \sum_{l=1}^n x_l^i \beta_l$  and

$$\frac{\partial \hat{y}^i}{\partial \beta_j} = \frac{\partial}{\partial \beta_j} \sum_{l=1}^n x_l^i \beta_l = x_j^i \quad (11)$$

Write equation (10) in matrix format we have

$$(y - X\beta)^T X = 0$$

or after transposing

$$X^T y - X^T X \beta = 0$$

Therefore the  $\beta$  that minimize  $\epsilon$  has to satisfy the following equation

$$\beta = (X^T X)^{-1} X^T y \quad (12)$$

Loosely speaking, we can interpreting equation (12) as composed two components the covariance related term  $X^T y$  and a term that is related to the variance of  $X$  i.e.  $X^T X$