

Homework 1 - Stat 37710: Machine Learning

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1. Question 1: Let A be a symmetric $d \times d$ matrix

(a) ANSWER:

Because A is symmetric,

$$\langle v, Av' \rangle = \langle Av, v' \rangle$$

And each is an eigenvector so,

$$\langle v, \lambda' v' \rangle = \langle \lambda v, v' \rangle$$

$$(\lambda' - \lambda) \langle v, v' \rangle = 0$$

But $\lambda' \neq \lambda$ so $v \perp v'$

(b) ANSWER:

By assumption, S spans V of dimension k so there exist k linearly independent vectors in S . By applying Gram-Schmidt we can construct an orthonormal basis v^1, \dots, v^k from this linearly independent set such that $\text{span}\{v^1, \dots, v^k\} = V$. Further, each v^i will be a linear combination of eigenvectors corresponding to λ . Thus, each v^i is in the λ eigenspace and thus is an eigenvector with eigenvalue λ .

W.L.O.G let our spanning linearly independent eigenvectors be $\{w_i, \dots, w_k\}$. Then performing Gram-Schmidt we have:

$$\begin{aligned} v_1 &= \frac{w_1}{\|w_1\|} \\ v_2 &= \frac{w_2 - \langle w_2, v_1 \rangle v_1}{\|w_2 - \langle w_2, v_1 \rangle v_1\|} \\ v_i &= \frac{w_i - \langle w_i, v_1 \rangle v_1 - \dots - \langle w_i, v_{i-1} \rangle v_{i-1}}{\|w_i - \langle w_i, v_1 \rangle v_1 - \dots - \langle w_i, v_{i-1} \rangle v_{i-1}\|} \end{aligned}$$

(c) ANSWER:

Taken together statements (a) and (b) imply that a $d \times d$ symmetric matrix A has d linearly independent eigenvectors (even if eigenvalues are repeated) and there exist eigenvectors v_i, \dots, v_d such that $\|v_i\| = 1$. This further implies that A is orthogonally diagonalizable and therefore we can write A as:

$$A = \begin{bmatrix} v_1 & \dots & v_d \end{bmatrix} \begin{bmatrix} \lambda_1 & \dots & 0 \\ \dots & \dots & \dots \\ 0 & \dots & \lambda_d \end{bmatrix} \begin{bmatrix} v_1 & \dots & v_d \end{bmatrix}^T$$

$$A = \begin{bmatrix} \lambda_1 v_1 & \dots & \lambda_d v_d \end{bmatrix} \begin{bmatrix} v_1 & \dots & v_d \end{bmatrix}^T$$

$$A = \sum_{i=1}^d \lambda_i v_i v_i^T$$

NOTE: $P = \begin{bmatrix} v_1 & \dots & v_d \end{bmatrix}$ has orthonormal columns so $P^{-1} = P^T$

(d) ANSWER:

Below we find constrained extremum of the desired function.

$$f(x) = \frac{w^T A w}{||w||^2} = \frac{w^T}{||w||} A \frac{w}{||w||} = u^T A u$$

Where $||u|| = 1$. So, do Langrange maximization of $f(u)$ such that $u^T u = 1$

$$L(u, \lambda) = u^T A u - \lambda(u^T u - 1)$$

$$\frac{\partial L}{\partial u} = 2A u - 2\lambda u = 0$$

$$\frac{\partial L}{\partial \lambda} = u^T u - 1 = 0$$

Together these equations show we have the following extremum:

$$A u = \lambda u$$

$$||u|| = 1$$

So extremum at eigenvalues with size 1. And because of the ordering of each λ_i we can see that the maximum occurs at v_d and the minimum occurs at v_1 .

2. Question 2: Orthogonal projections minimize distance to a subspace

(a) ANSWER:

To find the point in $V \in R^k$ closest to x we can take each dimension $1, \dots, k$ independently. So, WLOG fix some dimension k .

Let $x_{V_k} = (x \cdot p_k)p_k$ so x_{V_k} is the orthog projection of x to p_k . Let y be some point in the p_k space.

$$\begin{aligned} ||x - x_{V_k}||^2 &\leq ||x - x_{V_k}||^2 + ||x_{V_k} - y||^2 \\ &= ||x - x_{V_k} + x_{V_k} - y||^2 = ||x - y||^2 \end{aligned}$$

Thus, to minimize $||x - y||$ we must set $y = x_{V_k}$. Now, when projecting to the space of p_1, \dots, p_k we simply repeat this for k dimensions and we have our result.

(b) ANSWER:

For notations sake, we show the result for a single observation x_i and then summing over all x_i the result can be repeated.

$$\begin{aligned} ||x_i - \sum_j \langle v_j, x_i \rangle v_j||^2 &= \langle x_i - \sum_j \langle v_j, x_i \rangle v_j, x_i - \sum_j \langle v_j, x_i \rangle v_j \rangle \\ &= \langle x_i, x_i \rangle - 2 \langle x, \sum_j \langle v_j, x_i \rangle v_j \rangle + \langle \sum_j \langle v_j, x_i \rangle v_j, \sum_j \langle v_j, x_i \rangle v_j \rangle \\ &= \langle x_i, x_i \rangle - 2 \left[\sum_j \langle x, v_j \rangle \langle x, v_j \rangle \right] + \sum_j \langle x, v_j \rangle \langle x, v_j \rangle + \sum_{i,j:i \neq j} \langle x, v_j \rangle \langle x, v_j \rangle \langle v_i, v_j \rangle \\ &= ||x||^2 - \sum_j \langle x, v_j \rangle^2 \end{aligned}$$

The first term is clearly not affected by v_j and the left term, being negative, implies that our original minimization problem is equivalent to *maximizing* $\sum_j \langle x, v_j \rangle^2$ which, per slide 12 of 02DimensionalityReduction reduces to the Rayleigh quotient problem $v^T \Sigma v$. And we know we maximize this by choosing the largest so desired Eigenvectors from the var-cov matrix Σ NOTE: slide 12 has the proof of this final result, so rather than regurgitate it I am simply appealing to the result.

3. Question 3: Gram matrix questions

(a) ANSWER:

Let $A = [x_1, x_2, \dots, x_n]$ where $x_i \in R^d$ be the centered data matrix which therefore has dimension $d \times n$.

Further, because $n \geq d$, the matrix A has max rank of d .

The gram matrix, G , can be defined as $G = A^T A$.

We will show that $\text{rank}(A^T A) = \text{rank}(A)$ and thus $\text{rank}(G) \leq d$

To show that $\text{rank}(A^T A) = \text{rank}(A)$ we show that the dimension of each null space is equal thus implying the respective ranks are equal. Let $N(A)$ denote the null space of A .

Let $x \in N(A)$. Thus:

$$\begin{aligned} &\Rightarrow Ax = 0 \\ &\Rightarrow A^T Ax = 0 \\ &\Rightarrow x \in N(A^T A) \end{aligned}$$

Similarly, let $x \in N(A^T A)$. Thus:

$$\begin{aligned} &\Rightarrow A^T Ax = 0 \\ &\Rightarrow x^T A^T Ax = 0 \\ &\Rightarrow (Ax)^T Ax = 0 \\ &\Rightarrow Ax = 0 \\ &\Rightarrow x \in N(A) \end{aligned}$$

Putting this together, the null spaces are equal which by rank-nullity implies their ranks are equal

(b) ANSWER:

The matrix $K \in R^{n \times n}$ is PSD which implies we can compute the Cholesky decomposition such that $K = R^T R$ where R is an upper-triangle matrix. Further, K is clearly the Gram matrix corresponding to R and because K has rank r , we know R also has rank r . Let us extract the n columns from R . We know there are n columns and we know there are r linearly independent columns but each column is in some unknown space $R^?$ which may have dimension less than d . Thus, if needed we can append zeros to each column such that each column, which we can call $x_i \in R^d$, and the resulting Gram matrix of this data is K .

4. Question 4: Centering matrix P questions

(a) ANSWER:

As defined we have P is symmetric and thus $P^T = P$ so $P^2 = PP$. Below I show that $PP = P$

$$\begin{aligned} PP &= \begin{bmatrix} 1 - \frac{1}{n} & -\frac{1}{n} & \dots & -\frac{1}{n} \\ -\frac{1}{n} & 1 - \frac{1}{n} & \dots & -\frac{1}{n} \\ \vdots & \vdots & \ddots & \vdots \\ -\frac{1}{n} & -\frac{1}{n} & \dots & 1 - \frac{1}{n} \end{bmatrix} \begin{bmatrix} 1 - \frac{1}{n} & -\frac{1}{n} & \dots & -\frac{1}{n} \\ -\frac{1}{n} & 1 - \frac{1}{n} & \dots & -\frac{1}{n} \\ \vdots & \vdots & \ddots & \vdots \\ -\frac{1}{n} & -\frac{1}{n} & \dots & 1 - \frac{1}{n} \end{bmatrix} \\ PP_{i=i} &= (1 - \frac{1}{n})^2 + (n-1)\frac{1}{n^2} = 1 - \frac{1}{n} = P_{i=i} \\ PP_{i \neq j} &= 2[-\frac{1}{n}(1 - \frac{1}{n})] + (n-2)\frac{1}{n^2} = -\frac{1}{n} = P_{i \neq j} \end{aligned}$$

Thus, $P^2 = P$

(b) ANSWER:

First, we show going the " \Rightarrow " direction:

So, assume $Pv = 0$

$$Pv = \begin{bmatrix} 1 - \frac{1}{n} & -\frac{1}{n} & \dots & -\frac{1}{n} \\ -\frac{1}{n} & 1 - \frac{1}{n} & \dots & -\frac{1}{n} \\ \vdots & \vdots & \ddots & \vdots \\ -\frac{1}{n} & -\frac{1}{n} & \dots & 1 - \frac{1}{n} \end{bmatrix} \begin{bmatrix} v_1 \\ v_2 \\ \vdots \\ v_n \end{bmatrix} = 0$$

This results in the n system of equations. NOTE: the $\langle v_i \rangle$ denotes the i th element is excluded from the list:

$$(1 - \frac{1}{n})v_1 - \frac{1}{n}(v_2 + \dots + v_n) = 0 \quad (\text{eq. 1})$$

$$(1 - \frac{1}{n})v_i - \frac{1}{n}(v_1 + \dots \langle v_i \rangle \dots + v_n) = 0 \quad (\text{eq. i})$$

Solving this system yields $v_1 = v_2 = \dots = v_n$. To illustrate, we can solve for $n = 2$. If $n = 2$, then we have:

$$\begin{aligned} v_1 - \frac{v_1}{n} - \frac{v_2}{n} &= 0 \\ -\frac{v_1}{n} + v_2 - \frac{v_2}{n} &= 0 \end{aligned}$$

Combining, we have $v_1 - v_2 = 0 \Rightarrow v_1 = v_2$.

So v is the vector of ones times some constant (which could be zero) and we have our result.

Next, we go the " \Leftarrow " direction:

If $v = 0$ the result is obvious. If $v = [1]\lambda$ we can simply do the following:

$$Pv = \begin{bmatrix} 1 - \frac{1}{n} & -\frac{1}{n} & \dots & -\frac{1}{n} \\ -\frac{1}{n} & 1 - \frac{1}{n} & \dots & -\frac{1}{n} \\ \vdots & \vdots & \ddots & \vdots \\ -\frac{1}{n} & -\frac{1}{n} & \dots & 1 - \frac{1}{n} \end{bmatrix} \begin{bmatrix} 1 \\ 1 \\ \vdots \\ 1 \end{bmatrix} \lambda$$

Each n row has equation of the form $1 - \frac{1}{n} - \frac{n-1}{n}$ which clearly equals zero so we have our result.

5. QUESTION 5: Local linear embedding eigenvector problem derivation

ANSWER: Define $\Psi(y_1, \dots, y_n) = \sum_i^n \|y_i - \sum_j w_{i,j} y_j\|^2$. Without loss of generality, to simplify notation let $y_i \in R^1$

Thus, the objective function is

$$\begin{aligned} \Psi(y_1, \dots, y_n) &= \sum_i^n (y_i - \sum_j w_{i,j} y_j)^2 \\ &= \sum_i^n [y_i^2 - y_i (\sum_j w_{i,j} y_j) - (\sum_j w_{i,j} y_j) y_i + (\sum_j w_{i,j} y_j)^2] \\ &= \mathbf{Y}^T \mathbf{Y} - \mathbf{Y}^T (\mathbf{WY}) - (\mathbf{WY})^T \mathbf{Y} + (\mathbf{WY})^T (\mathbf{WY}) \\ &= ((\mathbf{I} - \mathbf{W})\mathbf{Y})^T ((\mathbf{I} - \mathbf{W})\mathbf{Y}) \\ &= \mathbf{Y}^T (\mathbf{I} - \mathbf{W})^T (\mathbf{I} - \mathbf{W}) \mathbf{Y} \end{aligned}$$

Let $\mathbf{M} = \mathbf{I} - \mathbf{W}$. Then $\Psi = \mathbf{Y}^T \mathbf{M} \mathbf{Y}$. Note, \mathbf{M} is the Gram matrix.

Because we assume $y_i \in R^1$ the var-cov I constraint becomes $\mathbf{Y}^T \mathbf{Y} = 1$. Constructing the Lagrangian we have:

$$\mathcal{L}(\mathbf{Y}, \lambda) = \mathbf{Y}^T \mathbf{M} \mathbf{Y} - \lambda(\mathbf{Y}^T \mathbf{Y} - 1)$$

$$\frac{\partial \mathcal{L}}{\partial \mathbf{Y}} = 2\mathbf{M}\mathbf{Y} - 2\lambda\mathbf{Y} = 0$$

$$\mathbf{M}\mathbf{Y} = \lambda\mathbf{Y}$$

Just as with PCA, \mathbf{M} is a symmetric matrix and thus has n orthonormal eigenvectors. Thus we can maximize and minimize (in this case we want to minimize) by selecting the eigenvectors corresponding to the smallest desired eigenvalues.