## Strong self-regulation and widespread facilitative interactions between genera of phytoplankton

Coexistence and facilitation in phytoplankton

Coralie Picoche<sup>1</sup>, Frédéric Barraquand<sup>1,2\*</sup>

- 1 University of Bordeaux, Integrative and Theoretical Ecology, LabEx COTE, Bât. B2 Allée Geoffroy St-Hilaire, 33615 Pessac, France; coralie.picoche@u-bordeaux.fr
- ${\bf 2}$  CNRS, Institute of Mathematics of Bordeaux, 351 Cours de la Libération, 33405 Talence, France; frederic.barraquand@u-bordeaux.fr

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Data accessibility: The REPHY dataset has already been published (REPHY, 2017) and all scripts for MAR and subsequent network analyses are available online in a GitHub repository (https://github.com/CoraliePicoche/REPHY-littoral). This repository will be made public upon acceptance and codes can be shared with referees should they wish to access them.

<sup>\*</sup> corresponding author

## Abstract

The persistence of phytoplanktonic diversity in spite of potentially strong competition is a long-standing puzzle of ecology. Here, we analyse a large dataset of phytoplankton abundance time series, counted every two weeks over twenty years, at ten sites along the French coastline. We estimate biotic interaction strengths using dynamic, multispecies autoregressive models. We show that a strong self-regulation, with competition strength within a taxon (genus) an order of magnitude higher than between taxa, was present in all interaction networks. This much stronger intra-taxon competition suggests that niche differentiation - rather than neutrality - was commonplace. Facilitation was also widespread, being even more frequent that inter-taxa competition. While network stability was unrelated to complexity measures, we unveiled links between self-regulation, inter-taxa interaction strengths and abundance. Strong self-regulation, widespread facilitation and stabilizing covariances between interaction strengths seem to be common features of coexisting phytoplanktonic communities in the field.

#### 1 Introduction

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How species or close genera can coexist together in spite of competition is one of the main puzzles of
   community ecology, especially for primary producers that seemingly share the same basic resources
   (Hutchinson, 1961). Many theoretical studies of competition models have shown that competitive exclusion
   is likely in those circumstances (Armstrong & McGehee, 1980; Chesson, 2018), unless mechanisms
   involving spatial or temporal variation are at play (Armstrong & McGehee, 1976; Chesson & Huntly,
   1997; Huisman & Weissing, 2001; Li & Chesson, 2016). Neutral theory, that assumes a non-equilibrium
   coexistence maintained by dispersal and equal competitive abilities for all species (Hubbell 2001, though
   there are exceptions, see Volkov et al. 2003, 2007) has been proposed as a solution to explain highly
   diverse communities (Hubbell, 2001; Rosindell et al., 2011).
      However, the evidence gathered from terrestrial plant communities starts to suggest that, in fact, niche
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   rather than neutral processes may be paramount to explain coexistence, with intraspecific competition
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   dwarfing interspecific competition in most cases (Adler et al., 2010, 2018b). Whether these conclusions
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   drawn from studies of annual plants and forest trees apply to other ecosystems and taxa is currently
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   little known (but see Mutshinda et al. 2009).
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      Moreover, even within a single trophic level, competition may not be the rule: the meta-analysis
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   by Adler et al. (2018b) reported a large number of facilitative interactions (30%) and several reviews
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   (Brooker et al., 2008; McIntire & Fajardo, 2014) have highlighted that facilitation may be much more
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   widespread than ecologists usually tend to think. Although some theoretical studies suggest that
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   facilitative interactions can be destabilizing (sensu resilience) and therefore undermine coexistence in
   Lotka-Volterra models (Coyte et al., 2015), multiple other modelling (Gross, 2008) and empirical (Brooker
   et al., 2008; Cavieres & Badano, 2009) studies have suggested that facilitative interactions can to a large
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   degree benefit coexistence, especially when multiple interaction types are considered simultaneously
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   (Mougi & Kondoh, 2012; García-Callejas et al., 2018).
      Here, we study a large multi-species dataset consisting of several multivariate long-term time series
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   of phytoplankton dynamics along the French coastline, which we then analyse using multivariate
   autoregressive (MAR) time series models, allowing for interactions between genera. Although many
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   ecological studies focus on interactions between species, competition has been shown experimentally to
   occur between different genera of phytoplankton (Titman, 1976; Descamps-Julien & Gonzalez, 2005).
   The genus level is also a rather fine taxonomic scale for phytoplankton interaction studies, as most studies
   are restricted to interaction between different classes or even phyla (Ives et al., 2003; Hampton et al.,
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2008; Griffiths et al., 2015). Studying interactions between different genera of phytoplankton therefore
both makes empirical sense in light of competition experiments and allows to estimate better-resolved
networks. We focus here on genera that belong mostly to diatoms and dinoflagellates. To put our
results into a more general context, we then compare our interaction strength estimates to previously
published interaction networks produced under the same statistical framework, both in plankton and
other empirical systems.

#### 33 Material and methods

#### 39 Sampling methods

All phytoplankton counts were collected by Ifremer coastal laboratories as part of the National Phytoplankton and Phycotoxin Monitoring Network (REPHY, 2017). Since 1987, this monitoring program has required the sampling of 26 sites along the French coastline every 2 weeks within 2 hours of high tide to document both biotic (phytoplankton counts) and abiotic (water temperature, salinity) variables. We 43 focused on sites which had the longest time series. We also excluded time series which had missing data for over 6 months or an average delay between sampling dates above 20 days. This reduced the number of study sites to 10 sites nested within 4 regions (Brittany, Oléron, Arcachon and the Mediterranean Sea; Fig. S1). 47 Abiotic variables (temperature, salinity) were measured directly from the boat during the sampling process while water samples for biotic analyses were fixed with a Lugol's solution and examined later. Phytoplankton cells above 20  $\mu$ m were identified at the lowest possible taxonomic level and counted with the Utermöhl method using an optical microscope (Utermöhl, 1958). Throughout the years and 51 sites, more than 600 taxa were identified at different taxonomic levels. We aggregated them at the 52 genus (or group of genera when not possible) level based on previous work (Hernández Fariñas et al., 2015; Barraquand et al., 2018), except for cryptophytes and euglenophytes in Arcachon, which could not be identified below the family level. Although the taxonomic resolution used here may seem coarse in comparison to land plants, it is in fact more refined than 86% of the MAR(1) studies of phytoplankton listed in Table S3. For each region, the MAR(1) analysis focused on the most abundant and most frequently observed genera to avoid most of the gaps in the time series. When gaps did not exceed a month, missing values were linearly interpolated; remaining missing values were replaced by a random number between 0 and half of the lowest observed abundance (Hampton et al., 2006). We tested extensively this and other

methods to deal with missing data in a previous publication on a subset of this dataset (Barraquand et al., 2018). All time series were scaled and centered before MAR analyses.

#### $_{64}$ MAR(1) model

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Multivariate autoregressive (MAR) models are used to determine the interspecific interactions and abiotic effects shaping a community's dynamics (Ives et al., 2003). MAR(1) models are based on a stochastic, discrete-time Gompertz equation which relates log-abundance of S species at time t+1 to log-abundances of the whole community at time t, with possible interactions between species, and effects of V abiotic variables at time t+1. These assumptions are encapsulated in eq. 1:

where  $\mathbf{n}_t$  is the  $1 \times S$  vector of log-abundance of phytoplankton groups,  $\mathbf{B}$  is the  $S \times S$  community

$$\mathbf{n}_{t+1} = \mathbf{B}\mathbf{n}_t + \mathbf{C}\mathbf{u}_{t+1} + \mathbf{e}_t, \mathbf{e}_t \sim \mathcal{N}_{\mathcal{S}}(0, \mathbf{Q})$$
(1)

(interaction) matrix,  $\mathbf{C}$  is the  $S \times V$  environment matrix describing the effects of V variables (stacked in vector  $\mathbf{u}_{t+1}$ ) on species growth rates, and  $\mathbf{e}_t$  is a  $1 \times S$  noise vector which covers both process and 72 observation error, following a multivariate normal distribution with a variance-covariance matrix Q. Q is diagonal and we have previously showed that this parsimonious choice did not affect qualitatively the 74 results (Barraquand et al., 2018). 75 We used the MARSS package (Holmes et al., 2014) v3.9, in R v3.3.2 (Venables & Smith, 2013), to estimate parameters with a maximum likelihood procedure. 77 We have previously published a detailed analysis of one of the dataset (Arcachon) for which more covariables were available (Barraquand et al., 2018), including nutrients and hydrodynamics variables. We 79 found that hydrodynamics variables were more influential than nutrients; nutrient dynamics contributed little to phytoplankton dynamics on the two-week timescale. Because temperature and salinity sum up seasonal changes in light as well as hydrology (salinity is inversely related to freshwater inflow), these represent the two key drivers needed to account for abiotic influences (Scheef et al., 2013). The analysis 83 of real data in Barraquand et al. (2018) was complemented by that of phytoplankton-like simulated data, 84 which confirmed the ability of MAR(1) models to infer biotic interactions and abiotic forcings (e.g., no need for extra non-linearities to model the storage effect, which was found to be nearly non-existent, as in previous analyses of plant data for which strong-self regulation was observed (Adler et al., 2010; Ellner et al., 2016; Adler et al., 2018b). Furthermore, using two abiotic variables (temperature and salinity) in this study rather than the full set used in Barraquand et al. (2018) led to almost identical estimates for the Arcachon study site. We are therefore confident that the MAR(1) models presented here are robust to small changes in model specification. In general, MAR(1) models tend to be fairly robust to small deviations of the underlying (non-linear) data-generating model, provided that one asks mainly order of magnitude of coefficients values (rather than precise point estimates) and sign of interaction coefficients (Certain et al., 2018), which is how these models are used here. For ease of interpretation of those coefficients, we also prove the correspondance between the magnitude of intra/inter interaction strength in a MAR(1) model and a multispecies Beverton-Holt model, i.e., a discrete-time Lotka-Volterra model (Cushing et al., 2004), in the Supplementary Information.

In this study, the number of phytoplankton groups, S, varies between regions but we keep the same 2 covariates, i.e. water temperature and salinity, that were measured at all study sites. Therefore, the 99 dimension of the dynamical system only depends on the (square of the) number of phytoplankton groups 100 we study, which ranges between 7 (Mediterranean Sea) and 14 (Brittany). The smallest system still 101 requires 70 parameters to be estimated if we consider all possible interactions between species. To reduce this dimensionality and remove unnecessary parameters, we compared different 'interaction scenarios' 103 based on BIC (Fig S3), which proved to be satisfactory in our previous analyses of both real data and similar simulated datasets (Barraquand et al., 2018). The null interaction scenario assumed no 105 interaction between groups of species (diagonal interaction matrix) and was compared to four other interaction scenarios. The first interaction scenario assumed that interactions could only occur between 107 phylogenetically close organisms, i.e., within a class (groups were then diatoms, dinoflagellates, and other 108 phytoplanktonic organisms) while the second interaction scenario further differentiated pennate and centric diatoms. The third interaction scenario considered the reverse hypothesis, that only unrelated 110 organisms could interact (i.e., a diatom could only interact with a dinoflagellate or a cryptophyte, but 111 not with another diatom), and the last interaction scenario did not constrain the interactions at all (full 112 interaction matrix). The second interaction scenario, hereafter called the pennate-centric scenario, had the lowest BIC for all sites and was therefore the most parsimonious, and was chosen as the basis for 114 further investigations of network structure.

#### Analysis of interaction strengths

The interaction matrix obtained from MAR(1) analyses can be used to determine the stability of a discrete-time dynamical system (Ives *et al.*, 2003). We compared the maximum modulus of the eigenvalues of the pennate/centric matrices in each site, as a proxy of stability, to network metrics which could be related to complexity, such as weighted connectance and linkage density (Breier *et al.*, 2018). Weighted

connectance is a measure of the proportion of realized links, taking into account the shape of the flux 121 distribution, while link density measures the average proportion and strength of interactions for a given species. These metrics are adapted to weighted interaction matrices but cannot accommodate for both 123 positive and negative coefficients: we therefore chose to focus on the absolute values of these coefficients, which can be linked to their strength, irrespective of interaction sign. 125

In addition to these network-level metrics, we also computed the average vulnerability (average effect of other taxa on a focal taxon, Eq S5) and impact (average effect of a focal taxon on other taxa, Eq S6) on both raw and absolute values of the coefficients, and compared these to the regulation a focal species exerted on itself. Raw values indicate the average effect (i.e., is the effect mostly positive or negative?) that can be expected on a species' growth rate from other planktonic species while absolute effects 130 characterise the strength of all types of interactions on a species (i.e., is a species strongly affected by the others?).

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Finally, we compared our results on the ratio between mean self-regulation/intraspecific interaction strength and mean interspecific interaction strength to other published studies based on a MAR(1) model. 134 A list of references is given in Table S3. Authors usually reported only coefficients that were significant at the 95% significance threshold, thus ignoring potentially many weak effects. For mean intergroup 136 interactions, we therefore computed both the mean value of all coefficients outside of the matrix diagonal, including zeroes (which decreases the mean intergroup interaction strength), and the mean value of 138 statistically significant intergroup coefficients only (which increases the mean intergroup interaction 139 strength). We should mention two potential biases associated with this comparison across the published literature: low-dimensional matrices tended to be more complete (less sparse) than high-dimensional 141 matrices, as these small interaction matrices were used to study known interaction phenomena (observed predation between organisms, for instance). Conversely, the number of parameters to estimate increases 143 as the square of the number of interacting groups, leading most authors to reduce this set before the estimation process for large interaction matrices. There is therefore a positive correlation between 145 sparsity and dimensionality (Fig S9). A second caveat is that while we informed our model selection by phylogeny (see above), several authors have instead reduced the number of estimated parameters by an 147 automated procedure, usually based on the comparison of 100 randomly chosen interaction matrices by 148 BIC (Ives et al., 2003). The latter choice may bias high non-zero interactions in the literature. This is why we decided to present in the main text intra/inter ratios including interspecific (or intergroup) 150 coefficients set to zero, which should be less sensitive to the model selection method and therefore make 151 comparisons across studies possible. 152

#### $_{\scriptscriptstyle{153}}$ Results

#### 54 Interaction estimates

Using MAR(1) autoregressive models, we have produced interaction matrices (Ives et al., 2003; Hampton 155 et al., 2013) – i.e., Jacobian community matrices on the logarithmic scale (Ives et al., 2003). Best-156 fitting models corresponded to a phylogenetically-structured interaction scenario, where interaction only 157 occurred betwen closely related genera (Fig S3). This led to sparse, modular matrices that have two main 158 features. First, we observe a strong self-regulation for all sites (Fig. 1, diagonal elements of all matrices), 159 a feature that we have previously highlighted in a more detailed analysis on one of the considered study 160 regions (Barraquand et al., 2018). The ratio of mean intragenus to intergenus interaction coefficients 161 varies between 6 and 10, not counting coefficients set to 0 in the estimation process. If we include the 162 zeroes in the interaction matrix in the computation of the intra/inter mean interaction strength, the 163 ratio rises to 21-43. Therefore, intragenus interactions are on average much stronger than intergenus 164 interactions, approximately 10 to 20 times stronger. 165

Second, although the percentage of facilitative interactions seemed to vary among sites (between 40% and 71% of interactions in the selected models), facilitation remained predominant in 9 sites out of 10 (only Lazaret, in the Mediterranean Sea, has 60% negative interactions). Our observational setup being nested, with sites within regions, we can examine whether locally positive interactions remain positive in a regional context: the percentage of consistently positive interactions at the regional level varies between 30% and 53%, higher than the percentage of similarly defined negative interactions (between 15% and 40%), except for sites in the Mediterranean Sea.

We have found that the percentage of true mutualism (+/+) is substantial: averaged over all sites, 32% of all interactions are (+/+) while only 12% of them are (-/-), see also Fig S5. The sign correspondence is not always maintained between French regions: the only interaction that is non-zero in the 10 sites (CHA/SKE) is mutualistic in Men er Roue only (Brittany) and mixed (+/-) at all other sites. Within the same region, however, interactions measured at different sites tend to keep the same sign. In the 3 sites of Oléron, for instance, there were 4 interactions which remained positive on both sides (CHA/GUI, DIT/GUI, LEP/THP, SKE/THP), 3 of them being also mutualistic in some of the Brittany sites. This contradicts previous observations that mutualistic interactions tend to be more context-dependent than competitive interactions (Chamberlain et al., 2014).

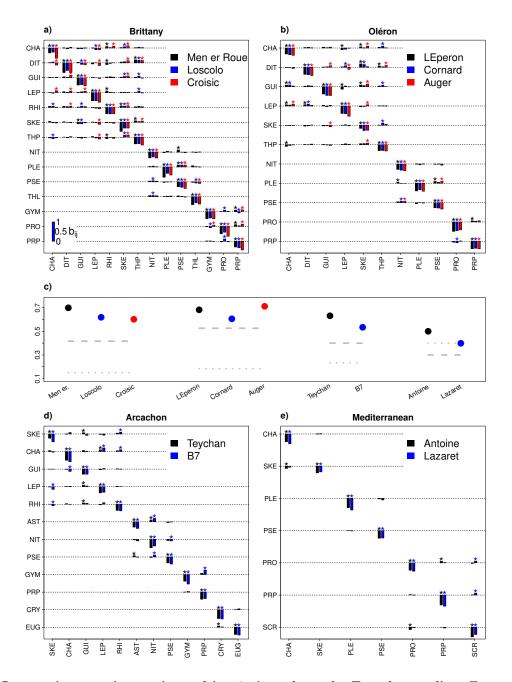


Fig 1. Interaction matrices estimated in 10 sites along the French coastline. Four regions are distinguished: Brittany (a), Oléron (b), Arcachon (d) and the Mediterranean Sea (e). Only interactions between clades (pennate and centric diatoms, dinoflagellates, other planktonic taxa) are allowed, as this is the best fitting interaction scenario (Fig S3). Taxon j (in columns) has an effect illustrated by the bar height on taxon i's growth rate (in rows). We present the interaction matrix minus the identity matrix ( $\mathbf{B} - \mathbf{I}$ ) because this compares unambigously intra- and intergenera interactions. The scale for the coefficient values is given at the bottom left of panel a). 95% significance of coefficients is marked by asterisks (\*). The community composition is given in Table S2. The fraction of positive interactions in each matrix is given by points in c) while the dashed (resp., dotted) line represents the ratio of interactions remaining positive (resp., negative) for all sites of a given region.

#### Interaction network analysis

The stability (sensu resilience, Ives & Carpenter 2007) of all interaction matrices was not strongly affected by the percentage of positive interactions or their connectivity properties (Fig. 2). The maximum modulus of the interaction matrix eigenvalues remained between 0.65 and 0.80. There was also a slight increase in stability with weighted connectance, with a drop in eigenvalue modulus for weighted connectances between 0.09 and 0.1.

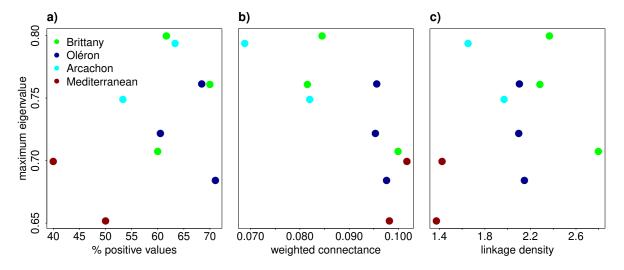


Fig 2. Relation between stability and complexity of the interaction networks. The maximum modulus of the interaction matrix **B** eigenvalues indicates stability *sensu* resilience. Each point color corresponds to a given region. Metrics formula for weighted connectance and linkage density are given in the Supplementary Information.

Given that a direct complexity-stability link was not obvious, we investigated whether the matrix coefficients had some particular structure that could help theoretical ecology to make better null models 189 of joint community dynamics and interactions (James et al., 2015). We defined two scores, vulnerability 190 (summed effect of others on the focal species growth rate, Eq S5) and impact (summed effect of the focal 191 species onto other species growth rates, Eq S6). Relations between inter- and intra-genus interactions emerged (Fig. 3): genera that were more self-regulating also had also a higher vulnerability score and a 193 lower impact score. Those two influences are likely to trade-off: a high degree of self-regulation somehow 194 buffers outside influences. Taxa that were less self-regulating were also more likely to have a broad 195 range of effects onto other taxa. As these genera tended to be more abundant (Fig S7), they were more 196 likely to affect other genera. It is important to note, however, that these trends are weak and there is 197 therefore a considerable amount of randomness dominating the interaction matrix: many scenarios of 198 self-regulation vs limitation by others are therefore possible.

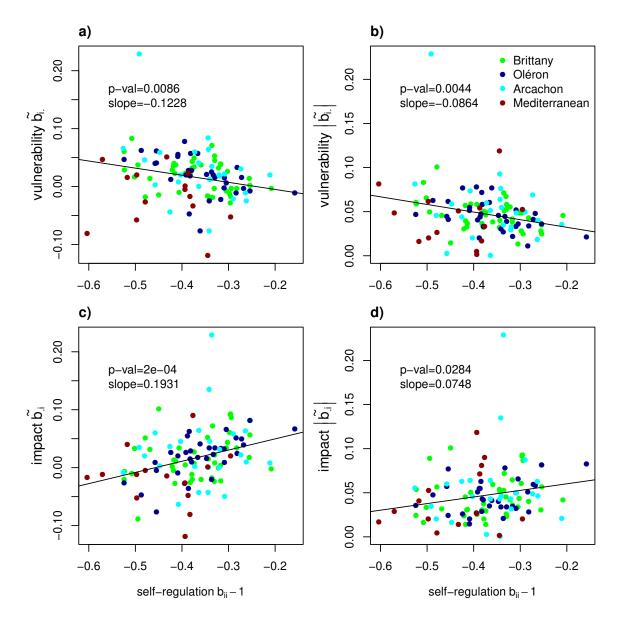


Fig 3. Relation between vulnerability/impact and self-regulation. Average vulnerability (effects of others on the focal taxon growth rate, a-b) and impact (effects of the focal taxon on others' growth rates, c-d), as well as self-regulation, are computed for untransformed (a-c) or absolute (b-d) values of the coefficients of the interaction matrix  $(\mathbf{B} - \mathbf{I})$  for the 10 study sites. Each color corresponds to a given region (Fig S1). The p-value of the Pearson correlation between vulnerability (respectively impact) and self-regulation is given in the top left of each panel and the slope of the linear regression is given by the black line across scatter plots.

Aside from these trade-offs, we found no remarkable patterns of covariation between matrix elements (other than a mean-variance scaling of interaction coefficients, Fig S6).

#### 202 Literature comparison

Finally, we sought to put these results in a broader context by compiling the intra vs inter group estimates 203 of previous MAR(1) studies of long-term observational count data (listed in Table S3). We found that 204 the order of magnitude of intra/inter interaction strengths considered here is not particularly above 205 those found for most planktonic systems to which MAR(1) models have been fitted, considering that our systems are relatively high-dimensional and that the higher the number of taxa, the larger the 207 intraspecific regulation (Barabás et al., 2017). We included in Fig. 4 not only plankton studies but also a couple of vertebrate or insect studies on less diverse communities, where interactions are stronger. 209 The conclusion from this comparison seems to be that, unlike small communities that can be tight-knit, 210 any diverse system of competitors and facilitators has evolved large niche differences making intragroup 211 competition much larger in magnitude than intergroup interactions. 212

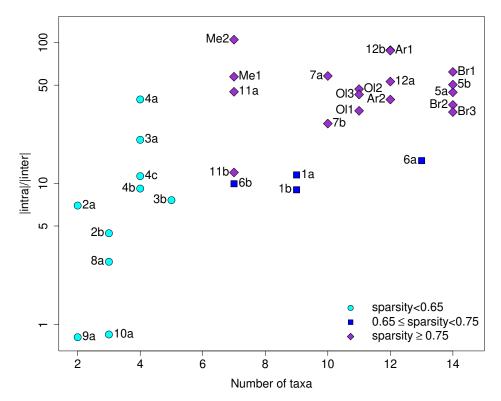


Fig 4. Ratio of intra- to inter-group interaction strength in Multivariate AutoRegressive (MAR) models in the ecological literature as a function of the number of species they include. The name of each studies, corresponding to each code, is given in Table S3. Codes beginning with letters correspond to the present study. The symbol color and shape correspond to the sparsity of the interaction matrix (e.g., the proportion of null interactions in the matrix). Intergroup interactions were set to 0 when they were not specified in the articles (in most cases, articles removed non-significant interactions at the 95% threshold; Fig S8 is the same figure taking into account only significant interactions)

#### Discussion

#### 214 Strong self-regulation and facilitation

We found very large niche differences between genera, translating into much higher intragenus than intergenus effects on growth rates (i.e., strong self-regulation), together with a high degree of facilitative net interactions.

The rather high intra/intertaxon interaction strength ratio (Levine & HilleRisLambers, 2009) that we 218 found, from 5 to 20, depending on how one counts the interactions set to zero in the estimation process, 219 could appear extremely high in light of previous intra/interspecific competition strength estimates of 4 220 to 5 by Adler et al. (2018b). Even though their model is a different one, i.e., Lotka-Volterra competition, 221 we prove in the Supplementary Information that the intra/inter ratio should remain commensurate. The 222 difference in the intra/inter ratio that we found should therefore lie elsewhere, which requires some 223 explanation. First, one could argue that such high intra/inter ratio arises because we consider the genus 224 as our baseline taxonomic unit, rather than the species. Although it is logical that niche differentiation 225 increases as one gets up the phylogenetic tree, and that getting down to the species level could slightly decrease that ratio, there are two arguments suggesting that the niche differences found here extend to 227 the species level. First, phytoplankton species belonging to different genera are often found to compete in experiments (Titman, 1976; Tilman et al., 1982; Descamps-Julien & Gonzalez, 2005). There is therefore 229 a massive difference between niche differentiation in the lab (weak) and in the field (strong, Barraquand 230 et al. 2018). Second, the only other study that managed to provide MAR(1) estimates down to the 231 species level for phytoplankton, that of Huber & Gaedke (2006), provides an intra/interspecific strength 232 ratio similar to ours (point 7a in Fig. 4). Strong self-regulation seems therefore a genuine feature of field 233 phytoplanktonic communities. 234

Another main finding of our study is the large frequency of positive interactions, with 30% truly mutualistic (+/+) interactions and between 40 and 70% facilitative effects. This again can be compared to the meta-analysis by Adler et al. (2018b) who also found facilitative interactions, but a little less than here (≈30%). However, Adler et al. (2018b)'s review contains many experiments while the plant literature is replete with field examples of facilitation (Brooker et al., 2008; McIntire & Fajardo, 2014), so that plant facilitation could be higher in the field. At the moment, it is therefore unknown how the predominance of facilitative interactions that we found in phytoplankton compares to facilitation in terrestrial plants. We note that several authors using MAR(1) models previously forbade positive interactions within the same trophic level, so that the fraction of facilitative interactions in plankton

<sup>244</sup> cannot be computed from literature-derived MAR(1) estimates.

The large niche differences and facilitative interactions that arise when considering a single trophic 245 level are an emergent property, arising from hidden effects of resources or predator partitioning/sharing 246 (Chesson, 2018). In our previous publication investigating in detail the Arcachon study sites (Barraquand et al., 2018), we have argued that for phytoplankton, the strong intragroup density-dependence could 248 arise from effects of natural enemies (Haydon, 1994; Barraquand et al., 2018). Natural enemies could also very well create apparent mutualism between prey species (Abrams et al., 1998; Barraquand et al., 250 2015; de Ruiter & Gaedke, 2017). We believe this to be likely true for the present study as well, given 251 that the new study regions (Oléron, Brittany, Mediterranean) have similar predators to the Arcachon site 252 (zooplankton, Jamet et al. 2001; Modéran et al. 2010; Tortajada et al. 2012, e.g.) and parasites (viruses, 253 Ory et al. 2010, e.g., fungi). Though natural enemies are good candidates to explain the observed niche 254 differences and emerging facilitation, one must bear in mind that other known drivers of phytoplankton 255 dynamics such as allelopathy (Felpeto et al., 2018), auxotrophy (Tang et al., 2010) or hydrodynamics (Lévy et al., 2018) can all, in theory, help create different niches and an emerging facilitation (see last 257 subsection of the Discussion). Finally, resources that are usually considered limiting for all species might in fact not always be: the changes in phytoplankton absorption spectrum documented by Burson et al. 259 (2018) constitute an example of fine-scale resource partitioning of one resource, light, that is usually believed to be limiting for all species and genera. 261

# No complexity-stability relationship but connections between self-regulation and intergroup interactions

There was no relation between the complexity of the communities (measured as either the weighted 264 connectance or linkage density of the interaction matrices) and their stability, as measured by the dominant eigenvalue of the interaction matrix, which measures the return time to a point equilibrium 266 (i.e., resilience). This result is conditional upon our model being a good approximate description of the 267 system (i.e., no multiyear limit cycles or chaotic attractors as the mapping between eigenvalues and 268 actual stability is distorted in that case (Certain et al., 2018) but we showed on a subset of this data 269 that a fixed point in a MAR(1) model, perturbed by seasonality and abiotic variables, is an accurate 270 description of the system (Barraquand et al., 2018). Therefore, we are confident that the absence of 271 complexity-resilience found here is genuine. This absence of direct link between complexity and stability could be an actual feature of empirical systems, as shown previously by Jacquet et al. (2016) using a 273 different technique, even though it does contradict previous results on random matrices, especially for

competitive and/or mutualistic networks (Allesina & Tang, 2012). We also found that the percentage of 275 mutualistic interactions, that is thought to affect the stability of the network (Mougi & Kondoh, 2012; Coyte et al., 2015; García-Callejas et al., 2018), does not have a major impact on the network's resilience. 277 In addition to weighted connectance, indices at the network-level (e.g., linkage density) and at the species or genus level (vulnerability and impact) approximate the average effects exerted and sustained 279 by any given taxa in the different study sites. While, at the network level, network structure (either complexity measures or the percentage of mutualistic interactions) did not affect stability, a relation 281 emerged between self-regulation, necessary for coexistence, and genus-level indices. We found that the 282 more a genus is self-regulated, the more it tends to be vulnerable to other genera impacts and the less it 283 impacts other genera. We examined whether vulnerability and impact could be affected by phylogenetic 284 correlations; they were not as on Fig. 3 points were not clustered according to genus, family or phylum. 285 High self-regulation usually indicates large niche differences with the rest of the community, and it makes 286 therefore sense that a species/genus whose needs strongly differ from the others only marginally impacts the resources of the other coexisting species. Furthermore, a low self-regulation was correlated with high 288 average abundance, which echoes findings by Yenni et al. (2017) who found that rare species usually show stronger self-regulation. This correlation between rarity and self-regulation could explain the lesser 290 impact effect of high self-regulated species/genus: a species which dominates the community composition can have a major effect on the others, especially as they usually cover more space, while it is harder for 292 rare, localised species to have large impacts. However, it was more difficult to explain the relationship 293 between self-regulation and vulnerability: a genus that is more self-regulated and rarer was found here to be on average more vulnerable to other genera's increases in densities. Such relation implies greater 295 stability (sensu resilience, Ives et al. 2003, and also invariability, Arnoldi et al. 2018) for the network as a whole, because the taxa that are the most vulnerable to other taxa's impacts are also those whose 297 dynamics are intrinsically more buffered. By which mechanisms this could happen is so far unclear and open to speculation. We caution, however, that the relationships between vulnerability, impact 299 and self-regulation that we evidenced are all relatively weak: considerable stochasticity dominates the distribution of interaction matrix coefficients. 301

#### Ghosts of competition past and present

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Overall, the dominance of niche differentiation in observational plankton studies – based on our analysis of the REPHY dataset and re-analysis of the MAR(1) literature – is similar to what has been recently found in plant community studies (Volkov *et al.*, 2007; Adler *et al.*, 2018b) or empirical food webs

including horizontal diversity (Barabás et al., 2017). Large niche differences might be due to the ghost
of the competition past, i.e., competition has occurred in the past, leading to strong selection and
subsequent evolution leading to progressive niche separation. In this scenario, species have evolved niches
that allow them not to compete or to interact only weakly (very strong facilitative effects might be
likewise destabilizing (Coyte et al., 2015). The likely predator effects that we highlighted above could be
comprised within such niche differentiation sensu largo: specialized predators can make strong conspecific
density-dependence emerge (Bagchi et al., 2014; Comita et al., 2014), while switching generalists can also
promote diversity (Vallina et al., 2014). Both predators and resources have often symmetrical effects
and can therefore contribute almost equally to such past niche differentiation (Chesson, 2018).

An intriguing new possibility, dubbed the "ghost of competition present" (Tuck et al., 2018), suggests 315 by contrast that spatial distributions in relation to abiotic factors might have a large impact on the 316 interaction strengths inferred from temporal interaction models such as ours. Recent combinations of 317 model fitting and removal experiments have shown that the model fitting usually underestimates the effect of competitors that are uncovered by removal experiments (Tuck et al., 2018; Adler et al., 2018a). 319 This could occur for instance if species are spatially segregated (at a small scale) because each species only exists within a domain where it is relatively competitive (Pacala's spatial segregation hypothesis 321 Pacala & Levin 1997), while a focal species could spread out if competitors were removed. This means that a species can be limited by competitors, but act so as to minimize competition (a little like avoidance 323 behaviour in animals), which implies that competition is in effect hard to detect when all species are 324 present. This would require spatial segregation between phytoplankton species at the scale of interactions, i.e., at the microscale. At the moment, it is known that turbulence generates inhomogeneities at the 326 microscale (Barton et al., 2014; Breier et al., 2018) but it is quite unclear how this affects multivariate 327 spatial patterns of species distributions (sensu Bolker & Pacala 1999, or Murrell & Law 2003). Moreover, 328 even if turbulence generates spatial structure with segregation between species, it is not quite clear that the "ghost of competition present" mechanism could work for plankton, because turbulence rather than 330 organism movement dictates where the phytoplankton patches can or cannot appear.

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## 337 Supporting Information

This article contains supporting information.

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