

Stabilizing niche differences are required to maintain species-rich communities in temporally variable environments

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Abstract

Explaining coexistence in species-rich communities of primary producers remains a challenge for ecologists because of the likely competition for shared resources. Following Hutchinson's seminal suggestion, many theoreticians have tried to create diversity through a fluctuating environment, which impairs or slows down competitive exclusion. There are now several fluctuating-environment models allowing coexistence, but they often produce only a dozen of coexisting species at best. Here, we investigate how to create even richer communities in fluctuating environments, using an empirically parameterized model. Building on the forced Lotka-Volterra model of Scranton and Vasseur (2016) inspired by phytoplankton communities, we have investigated the effect of two coexistence mechanisms, namely the storage effect and stabilizing niche differences (SNDs). SNDs occur when intraspecific competition is stronger than interspecific; we tuned the competition ratio based on empirical data. Although SNDs maintained more species (50%) than the storage effect (25%), we show that none of the two coexistence mechanisms considered could, by itself, ensure the coexistence of all species present at the beginning of our simulations. Realistic seasonal environments only aggravated that picture, as they decreased persistence relative to white noise. Our results suggest that combining different mechanisms for biodiversity maintenance into community models might be more fruitful than trying to find which mechanism explains best the observed diversity levels. Our results on community composition (formation of clumps, biomass distribution) in the different scenarios give further hints about likely coexistence mechanisms at work.

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1 Introduction

There has been a rich debate in theoretical ecology on how to reconcile niche and neutral perspectives on species coexistence (Gravel et al, 2006; Carmel et al, 2017). Under the neutral perspective, all species have equal birth and death rates and compete equally (since space is limited) whilst under the niche perspective, birth and death rates can vary between species and various mechanisms contribute to increasing intraspecific over interspecific competition (Hubbell, 2001). However, as it has been pointed out repeatedly, niche and neutral processes are not mutually exclusive: they may actually act together to produce observed species coexistence (Gravel et al, 2006; Mutshinda et al, 2009; Götzenberger et al, 2012).

For instance, Scheffer and van Nes (2006) put forward the concept of ‘clumpy coexistence’, whereby a simultaneous influence of both niche and neutral processes create several clumps of similar species along a single trait axis. Classical stabilizing niche differences (SNDs) enable coexistence of multiple clumps through stronger net intraspecific competition (Chesson, 2000), while within-clump coexistence occurs through neutral processes (Hubbell, 2001; Munoz and Huneman, 2016), as species that differ too little in their fitnesses cannot exclude themselves out on reasonable timeframes. Indeed, clumps on the trait axis eventually thin out in absence of immigration, but transient coexistence can last for extended periods of time. This ‘emergent neutrality’ within groups (Holt, 2006) has been proposed as a unifying concept for the niche and neutral theories. The findings of Scheffer and van Nes (2006) have been disputed due to hidden niches in the original model (Barabás et al, 2013). Hidden niches emerge through stronger intraspecific competition mediated by an additional intraspecific predation-like term (Barabás et al, 2013). This makes coexistence in the Scheffer and van Nes model more similar to that of the classical Lotka-Volterra model (Barabás et al, 2016), so that coexistence within clumps is not exactly neutral. Still, the idea that niche and neutral assembly can mould communities stays potent (Vernon et al, 2013). Since then, several studies have suggested that ‘clumpy coexistence’ can occur in theoretical models, most notably models incorporating temporal variation (Scranton and Vasseur, 2016; Sakavara et al, 2018). In these temporal-variation models, equal competitive strengths are combined with other mechanisms like the storage effect (or temporal niche partitioning, that is an equivalent

concept for forced Lotka-Volterra models, Barabás et al, 2012; Scranton and Vasseur, 2016). The storage effect increases the possibility of coexistence by making the interaction strength covary positively with a fluctuating environment (see also Barabás et al, 2012).

Here, we build on the work of Scranton and Vasseur (2016) to investigate the possibility of coexistence through species response to fluctuating environments. Our enthusiasm for the Scranton and Vasseur (2016) model stems from our interest in phytoplankton communities, that inspired their thermal preference curves modeling intrinsic growth rates. However, Scranton and Vasseur (2016) described temperature as a white noise, i.e., independent and identically distributed Gaussian random variates over time. This appeared to us a key assumption to relax. Under most latitudes, temperature is indeed a seasonal signal, and seasonal forcing can strongly affect the dynamics of the community considered (Vesipa and Ridolfi, 2017). Hence, over short timescales, random temporal variations often only add noise to a largely deterministic seasonal trend. Our present work can therefore be seen as an attempt to blend Scranton and Vasseur (2016)’s framework with the periodic environments of Barabás et al (2012), to better represent the mixture of stochastic and deterministic environmental forces affecting phytoplankton community dynamics.

Because many phytoplankton species or genera respond in similar ways to temperature despite having different optimas (Moisan et al, 2002), we hypothesized that a large seasonal variation might not necessarily foster species coexistence. In fact, similar responses to seasonal fluctuations should lead to an increased synchrony of species abundances which, in turn, should theoretically mitigate the expected temporal partitioning. How seasonality affects coexistence (as opposed to a purely randomly fluctuating environment) is therefore a key feature of this paper. In particular, we contrast cases where the storage effect is present vs. absent, which conveniently maps to two different parameterizations of the forced Lotka-Volterra model. Moreover, the overall diversity obtained at the end of the simulations with Scranton and Vasseur (2016)’s model was relatively low compared to what we usually observe in phytoplankton communities (several dozens to hundreds of species). We have therefore sought out which mechanisms would foster a truly species-rich community for extended periods of time.

In an empirical study combining phytoplankton community-level time series and multivariate

autoregressive models (Barraquand et al, 2018)¹, we found that despite a large influence of the environment (including temperature, irradiance, and other factors), intraspecific (or intragenus) competition was most likely the key driver of species coexistence. In other words, stabilizing niche differences had a large role to play in maintaining species diversity in coastal phytoplankton (Barraquand et al, 2018). These SNDs mirror those found in a number of terrestrial plant communities (Adler et al, 2018) and in animal communities (Mutshinda et al, 2009).

Here, we therefore try to establish what are the relative contributions of the storage effect vs SNDs to coexistence in a phytoplankton-like theoretical community model. This led us to cross different combinations of seasonality in the forcing signal, presence of the storage effect or not, and intra- vs interspecific competition, in order to disentangle the contributions of all these factors to biodiversity maintenance.

2 Methods

Models description

The model described in Scranton and Vasseur (2016) is based on a variant of the Lotka-Volterra competition model. Fluctuations in the environment are introduced in the model by temperature-dependent intrinsic growth rates (see Eq. 1-2, all coefficients are defined in Table 1) so that species growth rates can be expressed as:

$$\frac{dN_i}{dt} = r_i(\tau)N_i \left(1 - \sum_{j=1}^S \alpha_{ij}N_j\right) - mN_i \quad (1)$$

$$r_i(\tau) = a_r(\tau_0)e^{E_r \frac{(\tau - \tau_0)}{k\tau\tau_0}} f_i(\tau) \quad (2)$$

$$\text{where } f_i(\tau) = \begin{cases} e^{-|\tau - \tau_i^{opt}|^3/b_i}, & \tau \leq \tau_i^{opt} \\ e^{-5|\tau - \tau_i^{opt}|^3/b_i}, & \tau > \tau_i^{opt} \end{cases} \quad (3)$$

$$\text{and } b_i \text{ defined by numerically solving } \int r_i(\tau)d\tau = A \quad (4)$$

¹Preprint version available: see Barraquand et al (2017) in the reference list

Model parameters are detailed in Table 1, and we set their values to match the features of phytoplankton communities as in Scranton and Vasseur’s work (2016). The niche of each species is defined by its thermal optimum τ_i^{opt} . Thermal performance curves defined in Eq. 3 are parameterized so that all species share the same niche area (Eq. 4), which sets a trade-off between maximum growth rates and niche width.

Table 1: Parameter definition and values of the model described in Eq. 1-4

Name	Definition	Value (unit)
S	Number of species	60
N_i	Biomass density of the i th species	(kg/area)
τ	Temperature	K
$r_i(\tau)$	Growth rate of species i as a function of temperature	$\frac{\text{kg}}{\text{kg}^*\text{year}}$
α_{ij}	Strength of competition of species j on species i	0.001 area/kg
b_i	Normalization of the thermal decay rate	
m	Mortality rate	$15 \frac{\text{kg}}{\text{kg}^*\text{year}}$
τ_0	Reference temperature	293 K
$a_r(\tau_0)$	Growth rate at reference temperature	$386 \frac{\text{kg}}{\text{kg}^*\text{year}}$
E_r	Activation energy	0.467 eV
k	Boltzmann’s constant	$8.6173324 \cdot 10^{-5} \text{eV} \cdot \text{K}^{-1}$
$f_i(\tau)$	Fraction of the maximum rate achieved for the i th species	
μ_τ	Mean temperature	293 K
σ_τ	Standard deviation for temperature	5 K
τ_{\min}	Minimum thermal optimum	288K
τ_{\max}	Maximum thermal optimum	298 K
A	Niche breadth	$10^{3.1} \frac{\text{kg}}{\text{kg}^*\text{year}}$
τ_i^{opt}	Thermal optimum for growth of the i th species	K
θ	Scaling between white noise and seasonal signal	$[0, \sqrt{(2)}]$
ρ	Ratio of intra-to-intergroup competition strengths	(1;10)

We kept the mean and standard deviation of the forcing signal but included a lower-frequency component using a sinusoidal function with a period of 365 days (1 time unit being one day, Eq. 5). We tune the ratio of low-to-high frequency with the variable θ so as to keep the same energy content - i.e., equal total variance - in the forcing signal.

$$\tau(t) = \mu_\tau + \theta \sigma_\tau \sin\left(\frac{2\pi t}{365}\right) + \epsilon_t, \text{ where } \epsilon_t \sim \mathcal{N}(0, \sigma_\tau \sqrt{1 - \frac{\theta^2}{2}}) \quad (5)$$

Note that the upper limit for θ , $\sqrt{2}$, corresponds to a completely deterministic model which

we do not explore here (but see Zhao (1991) proving bounded coexistence). We chose to keep the stochasticity in the signal and to model a plausible temperature signal with $\theta = 1.3$ (see Fig. 1b) when considering a seasonal forcing of the dynamics.

The formulation of the forced Lotka-Volterra model of Scranton and Vasseur (2016) implies a storage effect, as the competition strengths covary positively with the growth rate values $r_i(\tau)$ (Chesson, 1994; Ellner et al, 2016). To test for the effect of an explicit storage effect in the model, we formulated a new version of this model, where we removed this assumption by using the mean value of a species' growth rate (\bar{r}_i) to weight the interaction coefficients (see Table 2).

$$\frac{dN_i}{dt} = N_i \left(r_i(\tau) - \sum_{j=1}^S \bar{r}_i \alpha_{ij} N_j \right) - m N_i \quad (6)$$

In this way, competition strengths remain unaffected by the environmental conditions, in contrast to growth rates (Eq. 6), while preserving the same average magnitude as in Eq. 1.

Stabilizing niche differences are ensured by the addition of the coefficient ρ , which is the ratio of intra-to-intergroup competition strength. We can therefore re-write the interaction coefficients α_{ij} in Eq. 7

$$\alpha_{ij} = \alpha (1 + (\rho - 1) \delta_{ij}) \quad (7)$$

where δ_{ij} is the Kronecker symbol, equal to 1 if $i = j$ and to 0 otherwise. The value of the parameter $\rho = 10$ was chosen from analyses of phytoplanktonic data (Barraquand et al, 2018) ².

In addition to two types of environmental forcings (white noise with $\theta = 0$, and seasonal forcing with $\theta = 1.3$), we therefore compare the results for four formulations of the model: with and without an explicit storage effect (Eq. 1 and Eq. 6, respectively) ; with and without stabilizing niche differences ($\rho = 10$ or $\rho = 1$, respectively). These are summed up in Table 2

²The values of non-zero interspecific competition coefficients reported in Fig. 5 of Barraquand et al (2018) are somewhat higher than here (and ρ lower, closer to 4) because the best-fitting model actually set to zero some intergroup competition coefficients; using a model where all α_{ij} are non-zero leads in contrast to $\rho = 10$.

$\frac{dN_i}{dt} + mN_i$	Storage effect	No storage effect
Stabilizing niche differences	$r_i(\tau)N_i \left(1 - \sum_{j=1}^S \alpha (1 + 9\delta_{ij}) N_j\right)$	$N_i \left(r_i(\tau) - \sum_{j=1}^S \bar{r}_i \alpha (1 + 9\delta_{ij}) N_j\right)$
No stabilizing niche differences	$r_i(\tau)N_i \left(1 - \sum_{j=1}^S \alpha N_j\right)$	$N_i \left(r_i(\tau) - \sum_{j=1}^S \bar{r}_i \alpha N_j\right)$

Table 2: Growth rate of species i in the four formulations of the model we present

126

127 Set-up

128 We replicate the ‘Species sorting’ experiment of Scranton and Vasseur (2016) so as to investigate
129 how the structure of synthetic phytoplankton communities varies under the different scenarios we
130 described above. We focused on the dynamics of a community initialized with 60 species with
131 thermal optima uniformly spaced along the interval $[15^\circ\text{C}, 25^\circ\text{C}]$, and with the same initial density
132 $\left(\frac{1}{\alpha S}\right)$. Each simulation was run for 5000 years in 1-day intervals. When the density of a species
133 dropped below 10^{-6} , it was considered extinct. For each combination of parameters (type of
134 environmental signal, storage effect and stabilizing niche differences), we ran 100 simulations.

135 All simulations were ran with Matlab’s ode45 algorithm, an explicit Runge-Kutta (4,5) algo-
136 rithm with an absolute error tolerance of 10^{-8} , and relative error tolerance 10^{-3} . The code is
137 available in a GitHub repository ³.

138 3 Results

139 Typical dynamics of the community following Eq. 1 (the model of Scranton and Vasseur, 2016),
140 with both the environmental signals described in Eq. 1 (original choice of Scranton and Vasseur,
141 2016; Fig. 1a) and Eq. 5 (our variant, Fig. 1b), are shown in Fig. 1c and d. A sinusoidal
142 forcing produces the strongly seasonally structured dynamics that are typical of phytoplankton.
143 Even though only 5 species can be seen in Fig. 1c, there were 14 species still present at the end
144 of the simulation forced by a white noise, with large disparities in the range of variation of their
145 biomasses. A third of the species kept a biomass above 10 kg/area (setting area = 1 ha, with a

³<https://github.com/CoraliePicoche/Seasonality>, will be made public upon acceptance or at the reviewer’s request

146 depth of a few meters, produce a realistic standing biomasses; Reynolds, 2006) while 6 out of the
147 14 remained below the unit. All persisting species in the white noise simulations were clustered
148 within a 3.2°C-range of thermal optima (see the biomass distribution as a function of the thermal
149 optimum in Fig. A1 in the Supplementary Material). No obvious temporal patterns (e.g., cycles)
150 could be seen in the communities forced by white noise. On the contrary, seasonal cycles were clear
151 in the seasonally-forced case of Fig. 1d. Only 4 species coexisted at the end of the simulation with
152 seasonal forcing, gathered in two groups with large thermal optimum differences (5.7°C between
153 the maximum thermal optimum of the first group and the minimum thermal optimum of the
154 second group). When temperatures are high, the group with higher thermal optima reaches its
155 maximum biomass, then as temperature decreases through the season, these species leave room
156 for the growth of the low-temperature group.

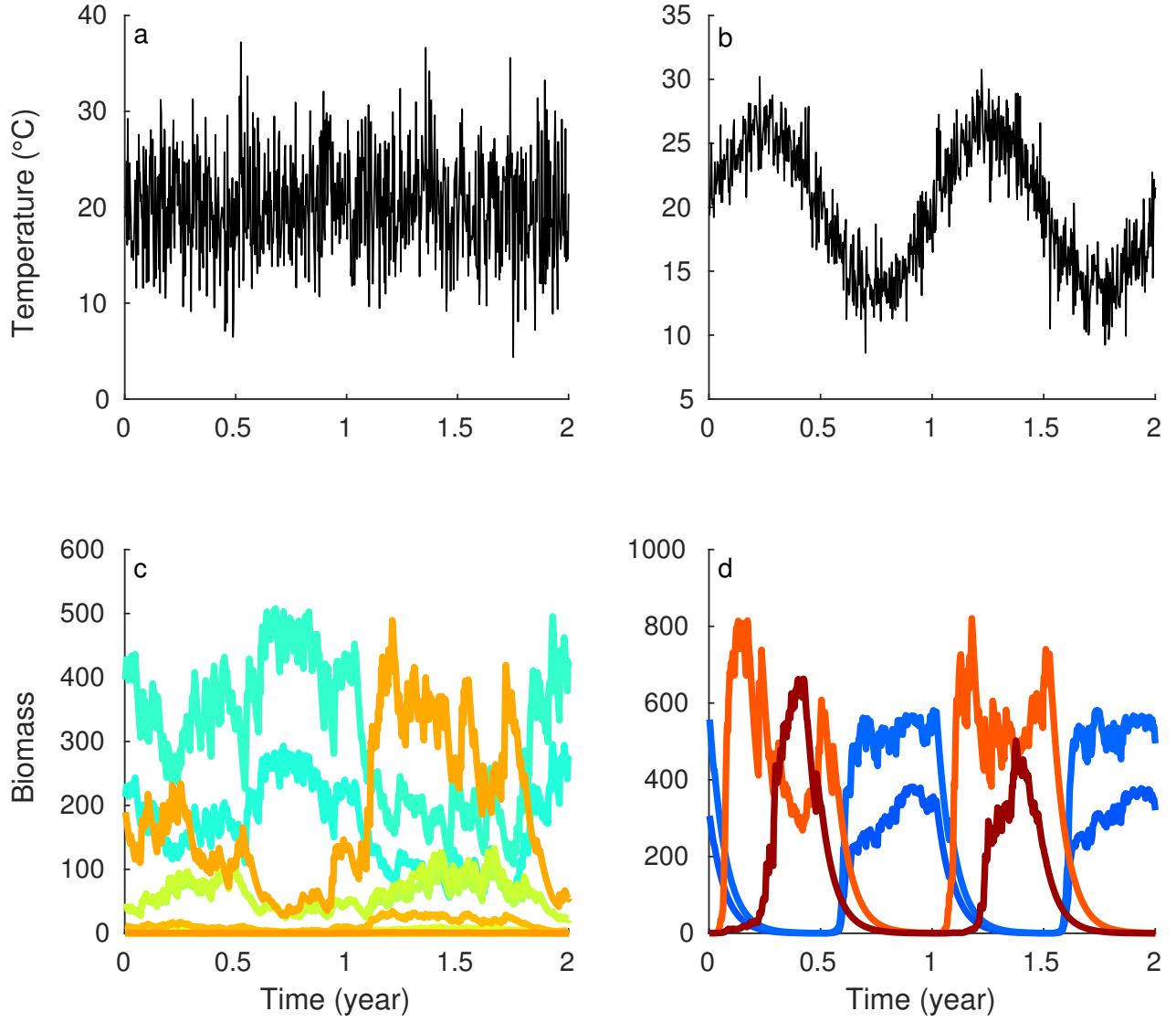


Figure 1: Time series of temperature (top; a-b) and extant species (bottom; c-d) for the 2 last years of a 5000 years simulation, with storage effect and no stabilizing niche differences. The forcing temperature is either a white noise (a) or a noisy seasonal signal (b), leading to community dynamics with more erratic fluctuations (c) vs seasonally structured fluctuations (d). Line colors of species biomasses correspond to their thermal optimum (from blue, corresponding to low thermal optimum, to red, corresponding to high thermal optimum).

157 The decrease in persistence due to seasonality observed in Fig. 1 was confirmed in all our
 158 simulations (Fig. 2). In cases where final species richness varied from one simulation to another
 159 (namely, the two middle cases in Fig. 2: with storage effect but without stabilizing niche differences,
 160 or without storage effect but with stabilizing niche differences), seasonality reduced the number
 161 of extant species to, on average, 27% and 48% of their original values, respectively (Fig. 2). A
 162 seasonal signal therefore led to a much smaller average persistence. There was also less variance

in persistence between seasonally forced simulations when compared to white noise simulations.

Both the stabilizing niche differences and the storage effect markedly increased persistence. Without any of these coexistence mechanisms, only one species persisted at the end of the simulations. When only the storage effect was present, the number of extant species varied between 8 and 20 (14.8 ± 2.4) with a white noise, or 2 and 6 (4.1 ± 0.7) with a seasonal signal. On the other hand, when only stabilizing niche differences were present, the number of extant species nearly doubled, varying between 20 and 32 (27.5 ± 2.4), or 12 and 15 (13.3 ± 0.6), with a white noise or a seasonal signal, respectively. Remarkably, when the storage effect and SNDs both affected the community dynamics, all species persisted in the community, while neither of these mechanisms was able to produce that result alone, for either white noise and seasonal forcing.

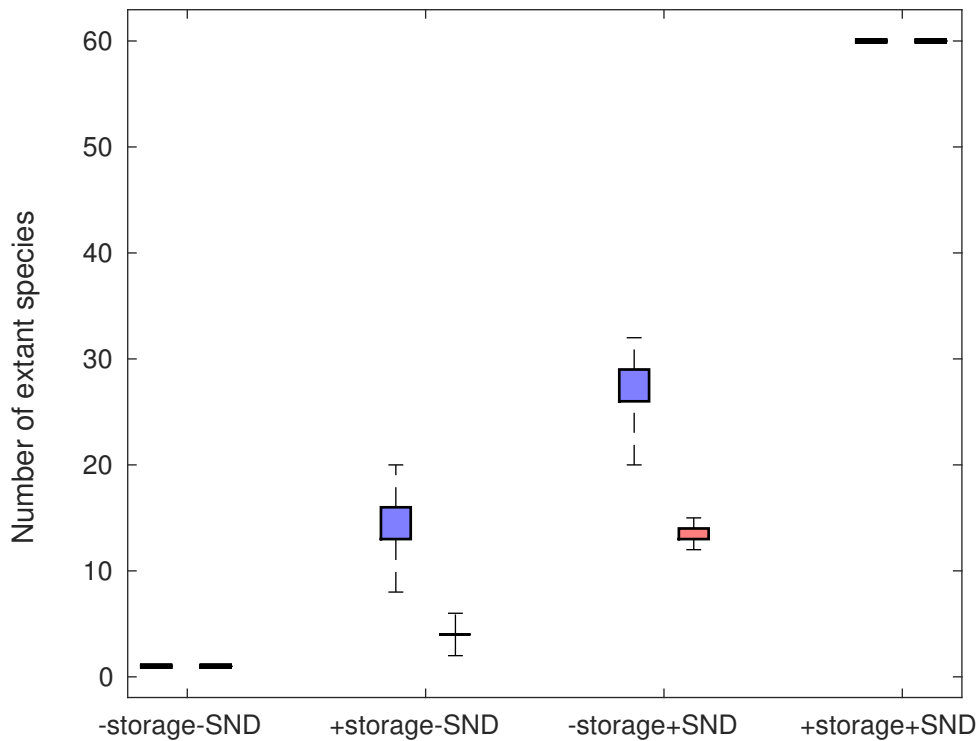


Figure 2: Number of species still present at the end of 100 simulations (5000 years each), initialized with 60 species, with a white noise forcing signal (blue) or a noisy seasonal signal (red). The signs + or - storage refer to presence or absence of a storage effect, respectively; + / - SND, presence or absence of Stabilizing Niche Differences, respectively. Community compositions are stable in the cases -storage-SND and +storage+SND, for which 1 or 60 species are still present at the end of all simulations, respectively. Due to low variance, the whiskers here represent min and max rather than 1.5 interquartile range.

173 When the richness of the community was stable (either 1 or 60 species at the end of the
 174 simulation), there were still large differences in the structure of the community with respect to
 175 temperature, due to both stochasticity and the type of forcing (Fig. 3). Without storage effect
 176 nor SNDs, a white noise forcing favored species with intermediate thermal optima, with two thirds
 177 of the simulations ending with a species with a thermal optimum between 18.9°C and 21.4°C
 178 (corresponding to only one fourth of the range of thermal optima present at the beginning of the
 179 simulation) and reaching a maximum average biomass in this range (Fig. 3a). This distribution
 180 can be related to a selection for the highest long-term growth rates, averaged over time (see scaled
 181 growth rates in Fig. 3). On the contrary, seasonality tended to favor species with larger maximum
 182 growth rates (thermal optima above 22°C). Species with a higher thermal optimum are more likely
 183 to persist and to reach a higher biomass at the end of the simulation. 38% of the simulations
 184 therefore ended with the species having the highest temperature optimum, 25°C.

185 When both coexistence mechanisms were present, the 60 initial species coexisted with small
 186 variations in biomasses for each species over the 100 simulations (mean CV=0.008 across simu-
 187 lations with either a white noise or a seasonal forcing signal, Fig. 3b and d). The forcing signal
 188 modified only the distribution of biomasses resulting in contrasted community structures despite
 189 equal richness in both simulation types. With a white noise, the distribution was unimodal with
 190 a maximum biomass reached for the second best long-term average growth rate (corresponding to
 191 a thermal optimum of 20.2°C). On the contrary, a seasonal signal led to a bi-modal distribution
 192 (centered on 17.0°C and 24.4°C), each corresponding to one season, with highest biomasses for
 193 higher thermal optima (Fig. 3d). The minimum biomass was reached for the best long-term av-
 194 erage growth rate at an intermediate temperature (20.4°C), one species apart from the maximum
 195 biomass in the white noise case.

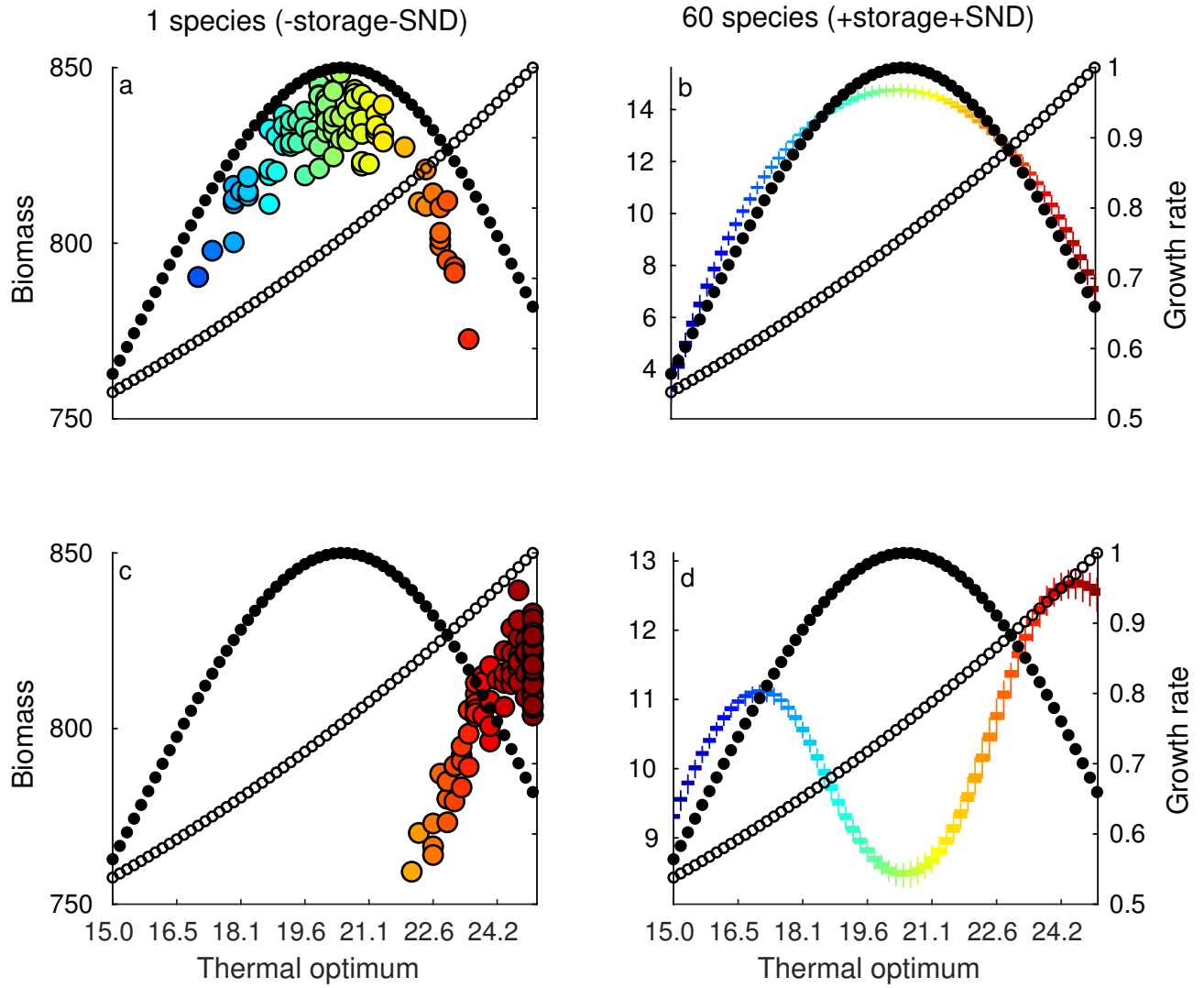


Figure 3: Mean biomass distribution over the last 200 years for 100 simulations, as a function of the species thermal optima. Here we consider the two stable-composition cases and two types of forcing signal. On the left column, simulations without storage effect nor stabilizing niche differences are presented. Only one species is present at the end of the simulations and its mean value is represented by one large colored circle per simulation. There can be several circles for the same species, corresponding to multiple simulations ending with this species alone. On the right column, simulations with storage effect and stabilizing niche differences are represented. All species are present at the end of the simulations and small boxplots present the variation in the temporal average of biomass with a given trait, for 100 simulations. The forcing signal is either a white noise (top) or a seasonal signal (bottom). Each species is identified by its thermal optimum through its color code. Scaled (divided by maximum) average and maximum growth rates are shown as small filled and open circles, respectively, and are indexed on the right y-axis.

196 In cases where the richness of the community varied, the overall shape (multimodal vs. uni-
 197 modal) of the marginal distribution of extant species with respect to the trait axis were similar
 198 for both types of environmental forcings (Fig. 4). By contrast, the type of coexistence mechanism

199 generated different shapes. Indeed, the storage effect led to a multimodal biomass distribution with
200 respect to thermal optima. We always observed 3 modes with a white noise and 3 modes in 95% of
201 the seasonal simulations (Fig. 4a). With a white noise, extant species are grouped in rather similar
202 clumps regarding species thermal optima (between 18.8°C and 22.2°C) whereas clumps tended to
203 be further apart in the seasonal case, covering a total range of 7.7°C, with species grouping in the
204 higher part of the thermal range, above 22°C. On the other hand, stabilizing niche differences led
205 to a quasi-uniform biomass distribution (Fig. 4 b). Species characterising communities forced by
206 a white noise stayed in the lower range of temperatures (in 96% of the simulations, the highest
207 thermal optimum was 22.4°C, see Fig. A2 in the Supplementary Material) while they were filtered
208 out in communities subjected to a seasonal fluctuation of their environment, for which species with
209 thermal optima above 20.5°C persisted.

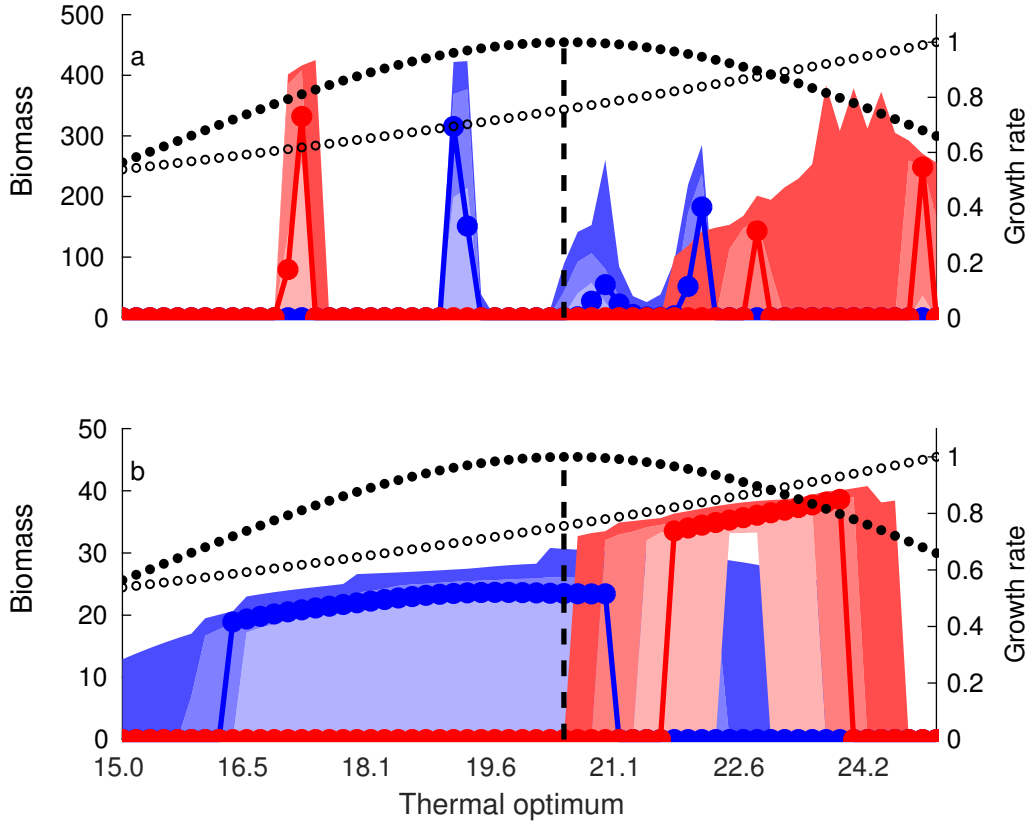


Figure 4: Mean biomass distribution over the last 200 years for 100 simulations, as a function of thermal optima, with storage effect and no stabilizing niche differences (a) and without storage effect, with stabilizing niche differences (b). The forcing signal is either a white noise (in blue) or a seasonal signal (in red). Shades of the same color correspond to the 50th, 90th and 100th percentiles of the distributions while colored lines correspond to one representative simulation. Scaled (divided by maximum) average (whose maximum is indicated by the dashed line) and maximum growth rates are shown as filled and open circles, respectively, and indexed on the right y-axis.

4 Discussion

We have simulated competitive Lotka-Volterra dynamics forced by a fluctuating environment (e.g., temperature fluctuations) under a range of scenarios allowing more or less coexistence. Two coexistence mechanisms, the storage effect and stabilizing niches differences, could be either present or absent, which led to four scenarios. These four scenarios were crossed with two possibilities for the forcing signal, a white noise and a stochastic yet seasonal signal, both with equal temporal variance. Our investigation therefore built on the model of Scranton and Vasseur (2016),

217 which included a white noise forcing and a storage effect, but considered seven additional com-
 218 binations of mechanisms. This was motivated by our wish to include two observed features of
 219 phytoplankton dynamics: seasonal cycles (Winder and Cloern, 2010) and stabilizing niche differ-
 220 ences (Chesson, 2000; Adler et al, 2010; Barraquand et al, 2018). Stabilizing niche differences,
 221 that occur whenever intraspecific competition is stronger than interspecific competition, can arise
 222 from many mechanisms: nonlinearities in the functional forms of competition or mutualism that
 223 contribute to increasing self-regulation (Kawatsu and Kondoh, 2018), predation or parasitism (see
 224 e.g., the generalist predators in Haydon, 1994), etc. They seem nonetheless an ubiquitous feature
 225 in primary producers (Adler et al, 2018).

226 We have found that first, temporally forced Lotka-Volterra dynamics cannot sustain any di-
 227 versity with our phytoplankton-based set of parameters, unless the structure is geared to include
 228 either a storage effect or SNDs. Although this absence of diversity-enhancing effect of “pure” en-
 229 vironmental variation has already been stated by other authors (Barabás et al, 2012; Fox, 2013;
 230 Scranton and Vasseur, 2016), this is not always intuitive (Fox, 2013), so we feel compelled to stress
 231 it once more: temporal variation in growth rate alone cannot help coexistence within competitive
 232 communities. A nice point made by Scranton and Vasseur (2016) was that a built-in storage ef-
 233 fect in a forced Lotka-Volterra model, parameterized for phytoplankton communities, could lead
 234 to some degree of coexistence. Our investigation reproduced these results, using the white noise
 235 forcing considered by Scranton and Vasseur (2016). However, an arguably more lifelike noisy and
 236 seasonal temperature forcing considerably lessened the richness of the community after 5000 years,
 237 decreasing from 15 to 4 species on average. Even imagining that groups represented here are genera
 238 or classes rather than species, this is a fairly low diversity for a phytoplankton-like community (see
 239 Chapter 1 in Reynolds, 2006). This suggests that the storage effect may not, on its own, be suffi-
 240 cient to maintain species-rich communities (e.g., dozens to hundreds of species). We have therefore
 241 sought out whether stabilizing niche differences could maintain a higher diversity, using field-based
 242 intra- vs intergroup (species or genera) competition strength ratio (Barraquand et al, 2018), where
 243 the intragroup density-dependence was chosen 10 times stronger. On their own, SNDs produced
 244 a higher level of diversity than the storage effect (almost double for white noise), which not only

aligns with our results on phytoplankton but also with results on perennial plants (Adler et al, 2010).

However, the seasonal forcing still considerably reduced diversity when only SNDs were considered, especially the “neutral” kind of diversity, i.e., diversity within clumps of similar traits. This diversity reduction occurs because within a season, the signal autocorrelation gives long, contiguous time intervals to the best competitor to exclude its less adapted heterospecifics. This makes the results likely to hold not only for seasonal environments, but more generally for autocorrelated ones, i.e., “red” noise. This could be relevant for species whose population dynamics occurs at timescales largely above one year. In contrast, a white noise generates large temperature shifts more frequently, and thereby forbid such competitive exclusion. In a seasonal setting, a species with the highest long-term averaged growth rate may not be the best competitor, and can disappear as a result of a strong competition from both low- and high-temperature tolerant species. This holds with or without a storage effect.

Our results may appear at odds with recent proposals that seasonal forcing in itself would help maintain diversity (Sakavara et al, 2018). However, we compared the effect of seasonal forcing to that of other forcing signals while controlling for total variance. Thus, the contrast between our results and those of Sakavara et al (2018) may be due to the role of forcing variance over time (we compare scenarios under a constant total variance). Overall, while seasonality may be slightly better than no forcing at all in maintaining diversity, seasonal forcing of parameters does not improve coexistence: diversity within clumps is lower when seasonality is included.

In addition to community diversity, the biomass-trait relationship also varied from one simulation to another. Some regularities did emerge across simulations though. The storage effect begot several clumps along the trait space (as observed by Scranton and Vasseur, 2016). The seasonality that we added to the temperature signal led to more distant clumps on the trait axis (as said above, less species per clump). Conversely, SNDs alone led to relatively uniform biomass distributions, with species forming a single large cluster, which covers a fraction of the initial trait space. The identification of multiple modes in biomass-trait relationships and SADs is relatively recent (Dornelas and Connolly, 2008; Matthews et al, 2014) and is a rare pattern in theoretical models

(McGill et al, 2007). Barabás et al (2013) convincingly argued that multimodality could arise from the demographic stochasticity of a single model run (with either SNDs or neutrality, but without the clumpy coexistence emerging from a storage effect). However, our results are based on many model runs, for which either the storage effect alone or a storage effect + SNDs in a seasonal context consistently produced multimodal distributions, while simulations without the storage effect always led to a single cluster along the trait axis. Our suggestion for empirical studies is as follows: if only one spatial location is observed, caution in interpreting multiple clumps on the trait axis is of course required, as Barabás et al (2013) highlighted. However, with several locations - or in a theoretical context - one could average across locations to reproduce similar graphs to the ones produced here. Clumps in the trait axis when averaged across model runs/locations is therefore a signature of the storage effect for the cases that we considered in the article. Of course, other mechanisms that we did not include in our models may produce similar patterns (Rael et al, 2018). Still, clustering on the trait axis, in scenarios where the environment fluctuates strongly in time, suggests that storage effects could be at work.

In our previous empirical study of coastal phytoplankton dynamics (Barraquand et al, 2018), we did not find any storage effect (which does not mean that it could not be observed in other systems). Given the results on species richness and composition presented here, we are skeptical that the storage effect alone could help explaining phytoplankton diversity. However, our results suggest that in phytoplankton-like seasonal environments, even though empirically-based SNDs produce much more diversity than the storage effect when considered in isolation, the storage effect can help diversity maintenance when combined to other mechanisms. Indeed, the combination storage effect + SNDs is non-additive: the cases where both SNDs and the storage effect were present showed more diversity than generated by any mechanism on its own.

The above results suggest the very exciting idea that multiple coexistence mechanisms might combine superadditively, thus helping us to better understand the astounding diversity of primary producers. This logic could, in principle, be extended to mechanisms that we have not considered here (e.g., spatial structure, specialized natural enemies, that could be as important here for plankton as they are for tropical trees, see Bagchi et al, 2014; Comita et al, 2014; Stump and Chesson,

2015; Barraquand et al, 2017). Previous research has however demonstrated that generalist seed predation could weaken the storage effect (Kuang and Chesson, 2009, 2010) thus different mechanisms might not always combine superadditively as we found here. That said, superadditivity has been found in some cases, i.e., pathogens could enhance the storage effect (Mordecai, 2015). Better explaining plant or microbial diversity would then not be about selecting the best unique mechanism susceptible to explain the observed diversity, but rather better combining those mechanisms together. This may obviously be an annoyance for those who like to sharpen Occam’s razor, but it clearly holds opportunities for theoreticians wishing to investigate synergies between coexistence mechanisms in highly diverse communities. Aside from the synergies between predator diversity-enhancing effects and SNDs or storage effects (on the temporal axis), one obvious follow-up of this research would be interactions with spatial structure. Spatial structure occurs both endogeneously, through spatially restricted movements and interactions, and exogeneously, through spatial variation in environmental covariates (Bolker, 2003). Numerous studies (e.g., Bolker and Pacala, 1999; Murrell and Law, 2002) have shown that spatially restricted movements and interactions - very small-scale spatial structure - can help coexistence, which we believe would be especially important for phytoplankton since many species form colonies (Reynolds, 2006; see discussion in Barraquand et al, 2018). Moreover, although temperature is usually relatively spatially homogeneous over space, other drivers (e.g., rainfall, pH in terrestrial ecosystems; salinity in aquatic ones) may exhibit spatial variation which is a main factor for coexistence (Snyder, 2008). The odds that different (resource) niches, natural enemies, spatial limits to competition and temporal niche partitioning all interact to promote the very high-dimensional coexistence observed in the field seem much higher than for any of those mechanisms alone. Whether the diversity-enhancing effects of these mechanisms combine sub-additively (as in Kuang and Chesson, 2010) or super-additively like here is therefore worthy of further research.

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