

S1 Characteristics of standard spatial point processes

In order to familiarize the reader with spatial process metrics, we present here well-known formula for the pair correlation function, Ripley's function and dominance index for standard point processes (a uniform distribution, i.e. the Poisson point process, and a clustered distribution, the Thomas process).

S1.1 Pair correlation function

In the case of a Poisson point process,

$$\forall r \geq 0, g_{ii}(r) = 1 \quad (1)$$

For a Thomas point process with parent intensity C_p and child-point position following a normal distribution centered on the parent point with standard deviation σ , the expected value of the pcf is

$$g_{ii}(r) = 1 + \frac{1}{C_p} \frac{1}{(4\pi\sigma^2)^{3/2}} e^{-\left(\frac{r^2}{4\sigma^2}\right)} \quad (2)$$

More generally, for a random superposition stationary point processes with marks (species, in our model) i and j , $\forall i, j, r \geq 0$, $g_{ij}(r) = 1$.

S1.2 Ripley's K-function

In the case of a Poisson point process,

$$\forall r \geq 0, K_{ii}(r) = \frac{4}{3}\pi r^3 \quad (3)$$

For a Thomas point process,

$$K_{ii}(r) = \frac{4}{3}\pi r^3 + \frac{1}{\lambda_p \sigma \sqrt{\pi}} \left(\sigma \sqrt{\pi} \operatorname{erf} \left(\frac{r}{2\sigma} \right) - r e^{-\left(\frac{r}{2\sigma}\right)^2} \right) \quad (4)$$

More generally, for a random superposition of stationary point processes, $K_{ij}(r) = \frac{4}{3}\pi r^3$.

S1.3 Dominance index

In the Poisson point process, $K_{ii}(r) = K_{ij}(r)$, which means the dominance index can be reduced to ratios of concentrations:

$$D_i(r) = \frac{C_i}{\sum_{j=1}^S C_j} \quad (5)$$

In the Thomas process, using eq. 4,

$$D_i(r) = \frac{C_i \left(\frac{4}{3}\pi r^3 + \frac{F(r)}{\lambda_{p,i}} \right)}{C_i \frac{F(r)}{\lambda_{p,i}} + \sum_j C_j \frac{4}{3}\pi r^3} \quad (6)$$

with $F(r) = \frac{1}{\sigma \sqrt{\pi}} \left(\sigma \sqrt{\pi} \operatorname{erf} \left(\frac{r}{2\sigma} \right) - r e^{-\left(\frac{r}{2\sigma}\right)^2} \right)$.

[[There will be here a 2x2 panels with a)spatial distribution b)pcf c)K function d)dominance index computed only from theoretical formula, just to give an idea of what we obtain with standard distributions in a more visual way.]]

S2 Theoretical behaviour of the dominance index in the BBM

With the theoretical formula for the dominance index in the Brownian Bug Model, we can show the progressive clustering of particles with time when advection is absent. Even after a short period of time, the dominance index without advection is larger than with advection, the difference being much more pronounced for larger cells.

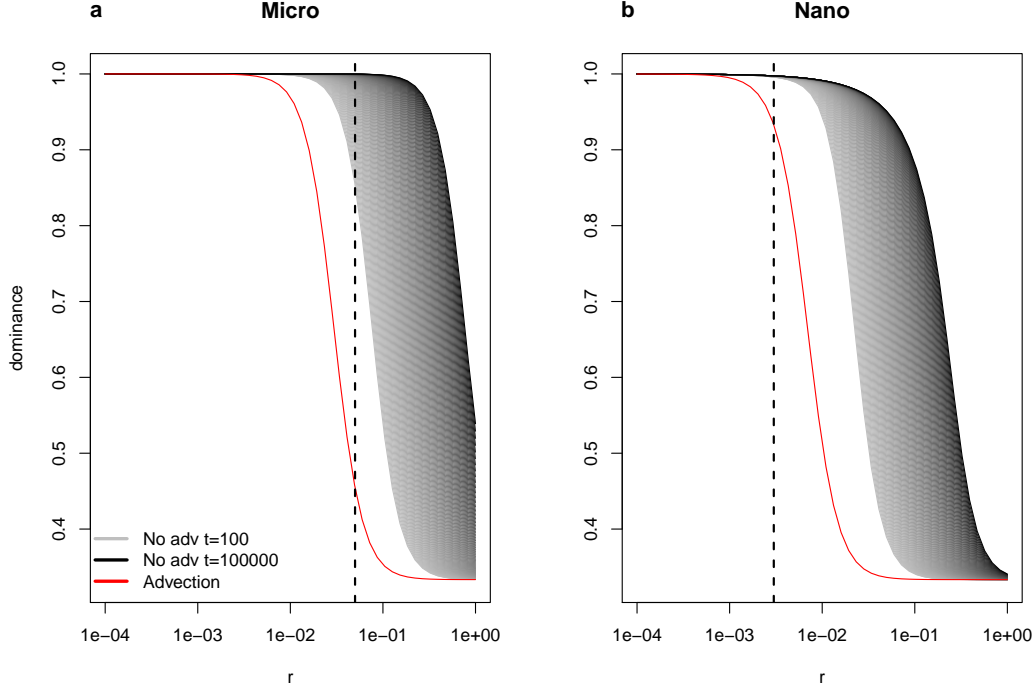


Figure S1: Theoretical dominance indices as a function of the distance from a particle of a given species, for a microphytoplankton (a) and nanophytoplankton (b) 3-species community with an even abundance distribution, with (red line) and without (grey to black lines, with darker lines for longer duration) advection. The vertical dashed line corresponds to the 10-diameter threshold.

S3 Computation of the pair correlation function and Ripley's function

The algorithm for pcf computation was mostly taken from the function `pcf3est` in `spatstat` 2.2-0 (Baddeley *et al.*, 2015) and adapted to interspecific pcf.

Schematically, $g(r)$ is estimated via the use of the Epanechnikov kernel κ_E with bandwidth δ .

$$\hat{g}(r) = \frac{1}{\hat{C}^2} \frac{1}{4\pi r^2} \sum_i \sum_j \kappa_E(r - ||x_i - x_j||) w(x_i, x_j) \quad (7)$$

where $w(x_i, x_j)$ is the Ohser translation correction estimator (Ohser, 1983) and the kernel is defined as follow.

$$\kappa_E(x) = \begin{cases} \frac{3}{4\delta} \left(1 - \frac{x^2}{\delta^2}\right) & \text{for } -\delta \leq x \leq \delta \\ 0 & \text{otherwise} \end{cases} \quad (8)$$

The estimate $\hat{g}(r)$ is therefore very sensitive to the bandwidth: if it is too small, the estimate is noisy and may even be missing several pairs of points; if it is too large, the smoothing might be so important that values are strongly underestimated. In `spatstat` 2.2-0 (Baddeley *et al.*, 2015), the bandwidth default value is $\delta = 0.26C^{-1/3}$. The pcf computation function was first tested on standard distributions (with the default bandwidth), then on the Brownian Bug Model (with different bandwidths, see below).

Estimates of the Ripley's function were also performed with the Ohser translation correction estimator but did not require any kernel smoothing.

S3.1 Standard point processes

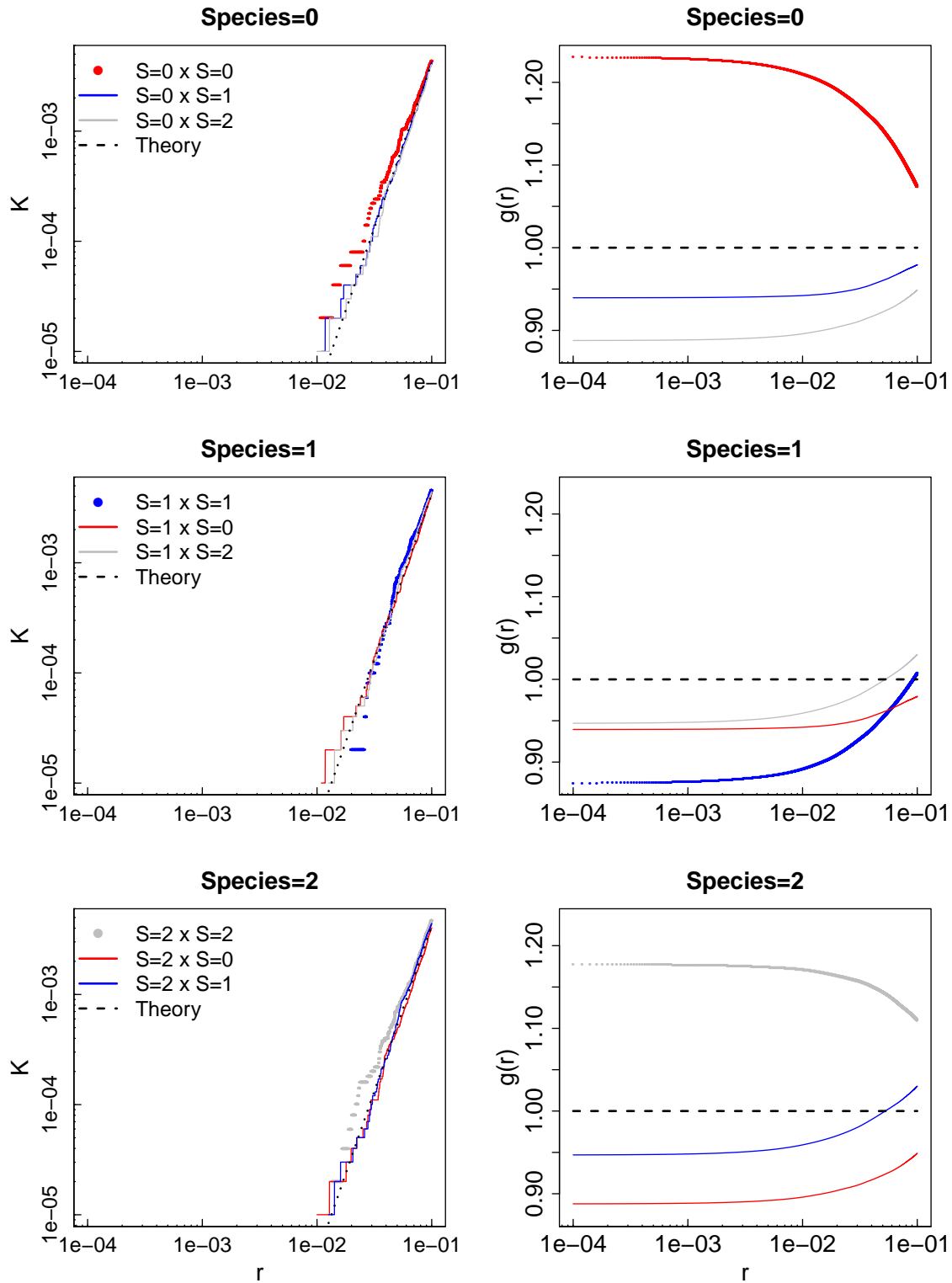


Figure S2: Intra- and inter-specific Ripley's K function and PCF values for 3 species following a Poisson process with intensity 10 cm^{-3} , in a volume of 1000 cm^3 . Values computed from our simulations (circles and solid lines for intra- and interspecific values, respectively) are compared with theoretical formula (dotted lines). Note that theoretical values are the same for intra and interspecific moments for the Poisson distribution. Colors correspond to the different species (red for species 0, blue for species 1, black for species 2).

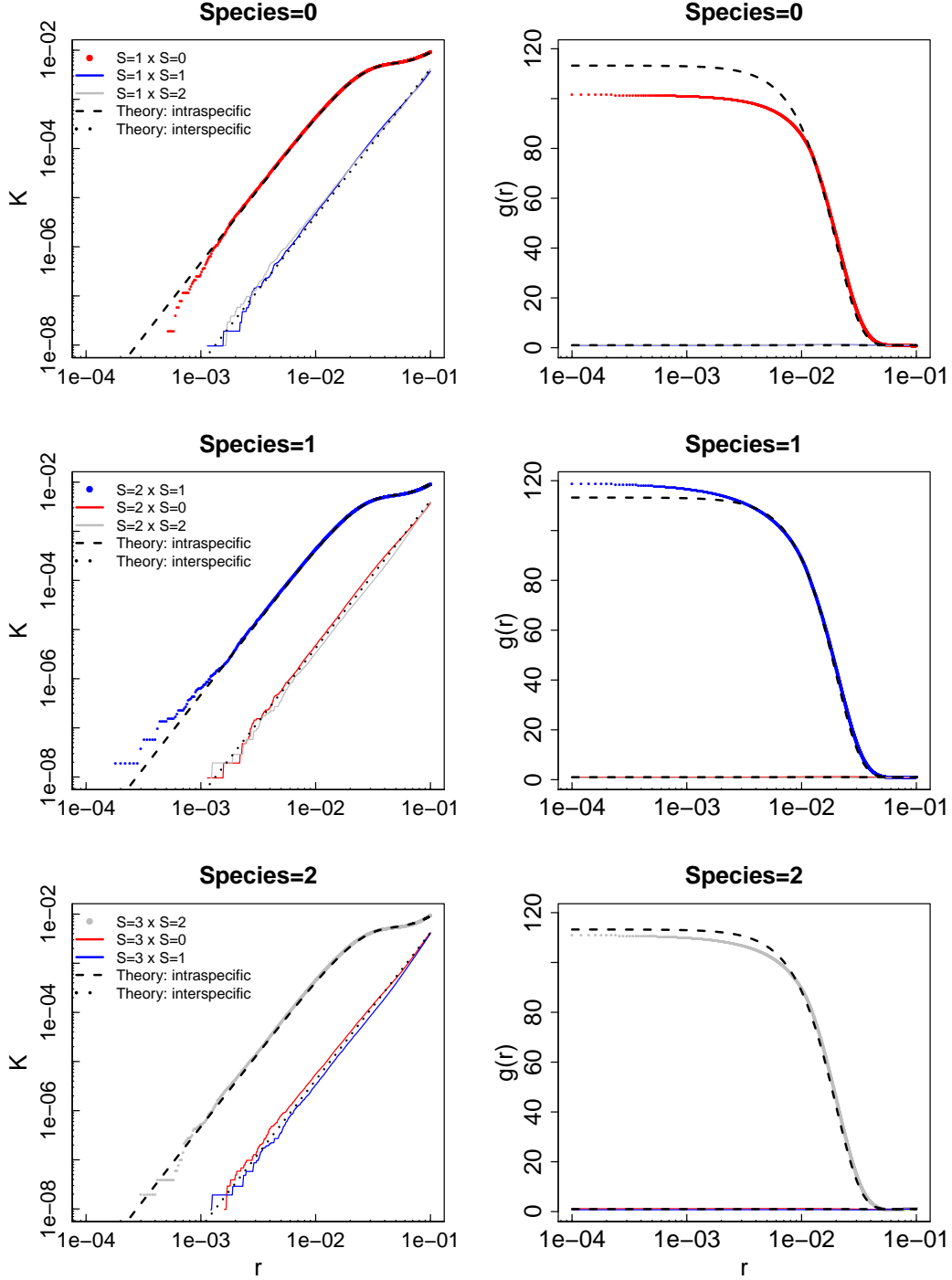


Figure S3: Intra- and inter-specific Ripley's K function and PCF values for 3 species following a Thomas process with parent intensity $\lambda_p = 200 \text{ cm}^{-3}$, number of children per parent $N_c = 50$, in a volume of 1 cm^3 , $\sigma = 0.01$ and $\delta \approx 0.012$. Values computed from our simulations (circles and solid lines for intra- and interspecific values, respectively) are compared with theoretical formula (dashed and dotted lines for intra- and interspecific values, respectively). Colors correspond to the different species (red for species 0, blue for species 1, black for species 2).

S3.2 Brownian Bug Model

While the pcf was one of the first indices we intended to use, we quickly realized that the combination of the large range of radii we wanted to explore (from 10^{-4} to 1 cm) and the low density of particles, at least for microphytoplankton, made the estimation difficult as the choice of the bandwidth was critical. We give an example

of the sensitivity of the pcf computation to the bandwidth below (Fig. S4).

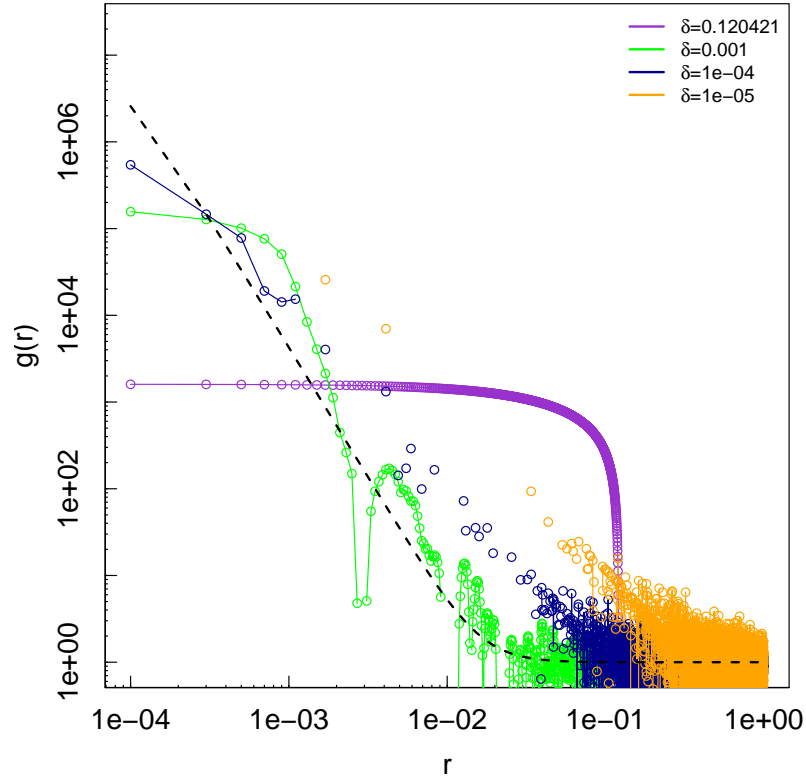


Figure S4: Intraspecific pcf computed for the Brownian bug model with microphytoplankton particles, after 1000 time steps, with different values of the bandwidth δ . The dashed line indicates the theoretical pcf.

We decided, from these results, to focus on Ripley's function, which enabled us to compute the dominance index without having to calibrate a bandwidth beforehand.

S4 Spatial distributions

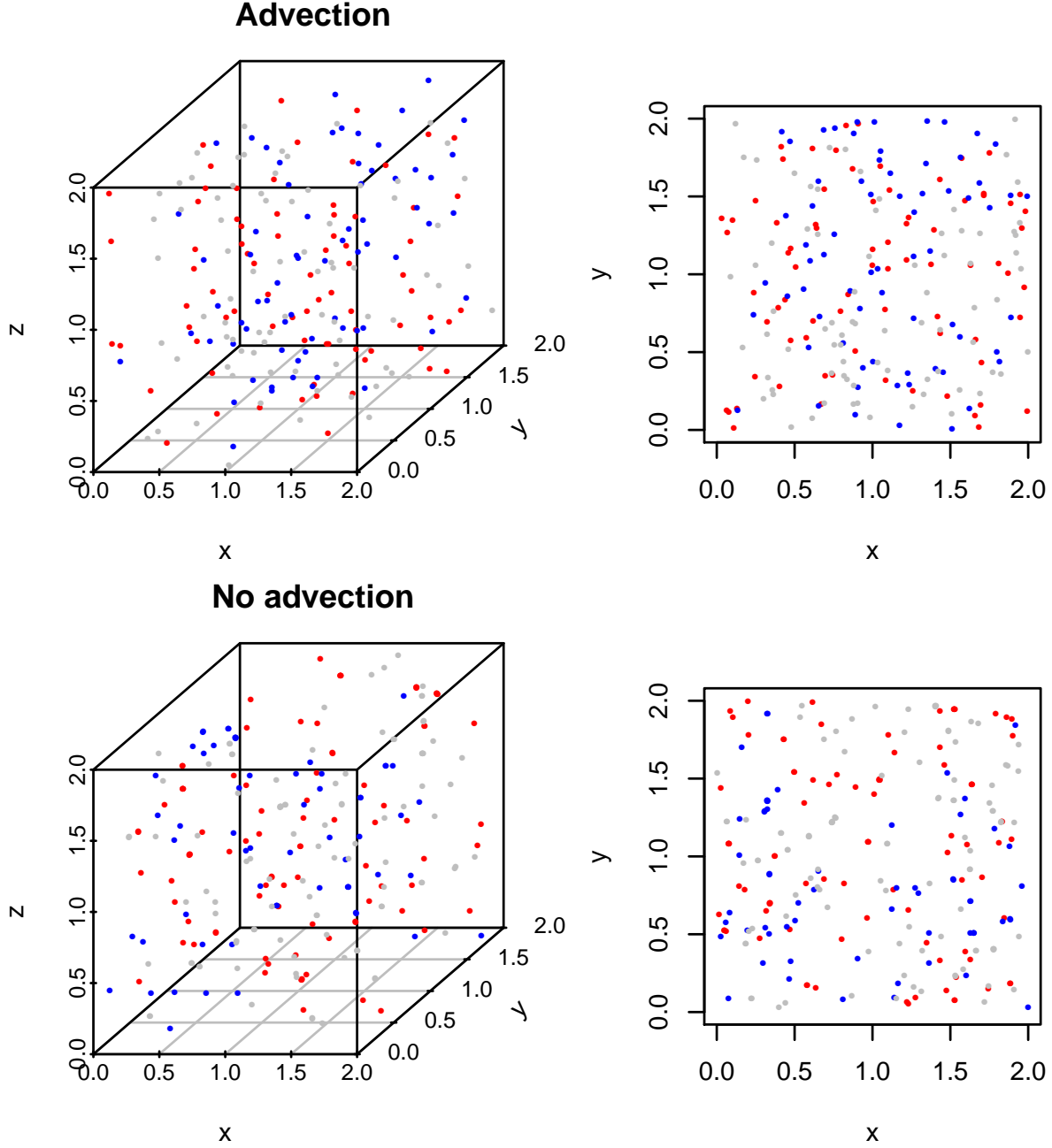


Figure S5: Spatial distributions of a 3-species community of microphytoplankton with and without advection with density $\lambda = 10 \text{ cm}^{-3}$ after 1000 time steps. Each color corresponds to a different species. On the left-hand side, only a zoom on a $2 \times 2 \times 2 \text{ cm}^3$ cube is shown, and its projection on the x-y plane is shown on the right hand-side.

S5 Minimum distances between points

Theory

One of the reason why estimating K and $g(r)$ is difficult is that for small radii (below 10^{-2}), we can find very few observed distances between pairs of points. As a first proxy, we want to estimate the minimum expected distance between points (distance to the nearest neighbour, DNN) when they are uniformly distributed.

In d dimensions, the probability distribution of the distance r to the nearest-neighbour follows $f(r) = db_d \lambda r^{d-1} \exp(-b_d r^d \lambda)$. If we want to find the distribution of the minimum DNN between n realized points of a Poisson process with intensity λ , we can write:

$$\begin{aligned}
Pr(\min(R_1, \dots, R_n) > r) &= Pr(R_1 > r, \dots, R_n > r) \\
&= \prod_i Pr(R_i > r) \\
&= \prod_i \exp(-b_d r^d \lambda) \\
&= \exp(-b_d r^d \sum_i \lambda)
\end{aligned} \tag{9}$$

We can then conclude that the distribution of the minimum distance follows the same distribution as the DNN, but with intensity $n\lambda$.

Clark & Evans (1979) show that a variable with probability distribution (with notations changed to fit our own) $f(r) = \frac{d\lambda\pi^{d/2}r^{d-1}}{\Gamma(\frac{d}{2}+1)} \exp(-\frac{\lambda\pi^{d/2}r^d}{\Gamma(\frac{d}{2}+1)}) = d\lambda b_d r^{d-1} \exp(-\lambda b_d r^d)$ has an expected value of $\mu_d = \frac{(\Gamma(\frac{d}{2}+1))^{1/d} \Gamma(\frac{1}{d}+1)}{\lambda^{1/d} \pi^{1/2}}$.

With intensity $n\lambda$, we can write $\frac{(\Gamma(\frac{d}{2}+1))^{1/d} \Gamma(\frac{1}{d}+1)}{(n\lambda)^{1/d} \pi^{1/2}}$.

In 3D,

$$\begin{aligned}
\mu_d &= (n\lambda)^{-1/3} \frac{(\Gamma(\frac{3}{2}+1))^{1/3} \Gamma(\frac{1}{3}+1)}{\pi^{1/2}} \\
&= (n\lambda)^{-1/3} \left(\frac{3}{2}\Gamma(3/2)\right)^{1/3} \frac{1}{3}\Gamma(1/3) \frac{1}{\pi^{1/2}} \\
&\approx 0.554 \frac{1}{(n\lambda)^{1/3}}
\end{aligned} \tag{10}$$

This needs to be taken into account when defining λ . For microphytoplankton, using $\lambda = 10 \text{ cm}^{-3}$ and $n \approx 10^4$, the smallest expected distance for a uniform distribution is $1.2 \times 10^{-2} \text{ cm}$. For nanophytoplankton, using $\lambda = 10^3 \text{ cm}^{-3}$ and $n \approx 10^4$, it is reduced to $2.6 \times 10^{-3} \text{ cm}$.

Simulations

We can compute the simulated distance to the nearest neighbour compared to what we should obtain with a uniform distribution (Fig. S6 and S7): the simulated average DNN are close to the expected value for a uniform distribution, but the minimum distance to a conspecific is much lower than expected.

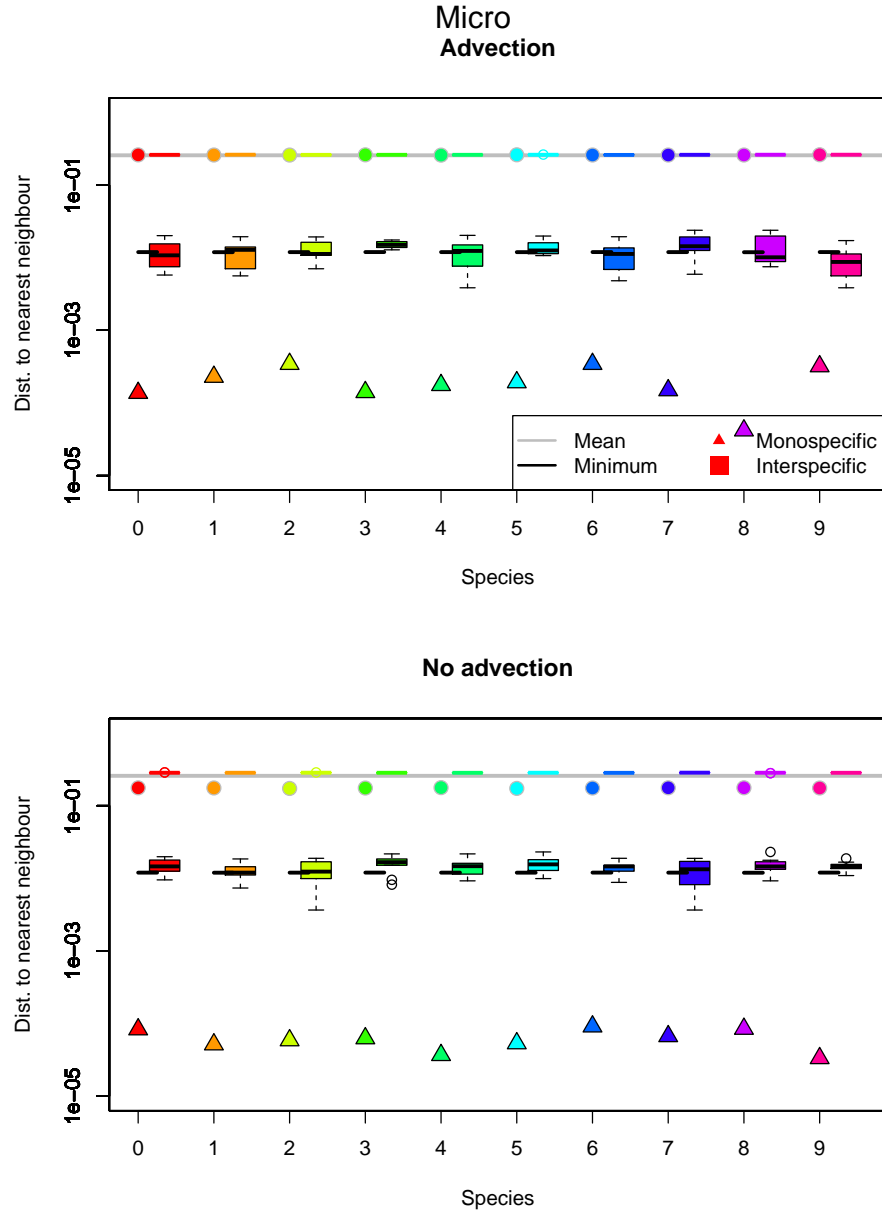


Figure S6: Mean and minimum distance to the nearest neighbour for 10 microphytoplankton species with density $\lambda = 10 \text{ cm}^{-3}$, with and without advection, after 1000 time steps, compared to predictions for a uniform distribution. Horizontal lines show the average distance to the nearest neighbour (grey line) and the expected minimum distance to the nearest neighbour with the actual number of realizations (black line). Circles and triangles represent mean and minimum distance to a conspecific, respectively. Boxplot corresponds to the distribution of mean (grey outlines) and minimum (black outlines) distances to a heterospecific. Colors correspond to different species.

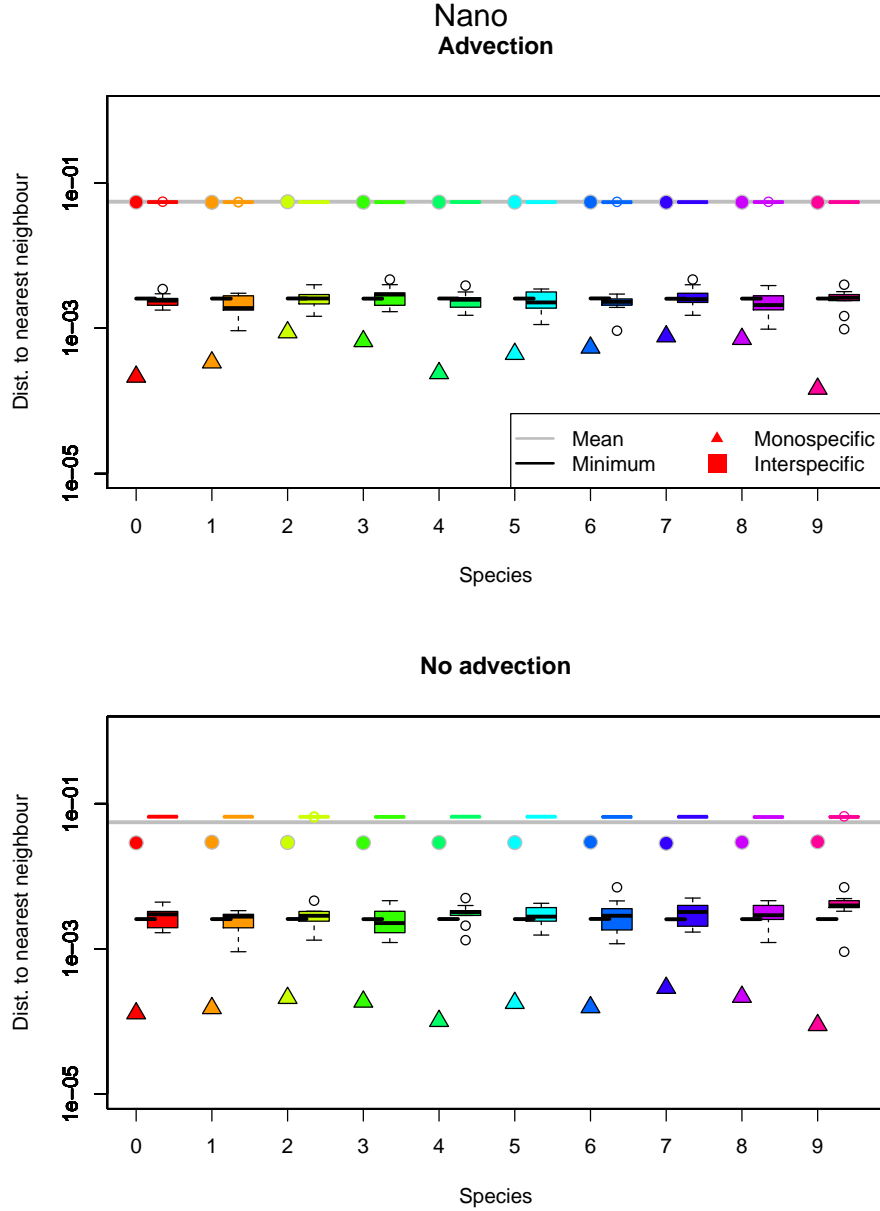


Figure S7: Mean and minimum distance to the nearest neighbour for 10 nanophytoplankton species with density $\lambda = 10^3 \text{ cm}^{-3}$, with and without advection, after 1000 time steps, compared to predictions for a uniform distribution. Horizontal lines show the average distance to the nearest neighbour (grey line) and the expected minimum distance to the nearest neighbour with the actual number of realizations (black line). Circles and triangles represent mean and minimum distance to a conspecific, respectively. Boxplot corresponds to the distribution of mean (grey outlines) and minimum (black outlines) distances to a heterospecific. Colors correspond to different species.

S6 Effect of the initial distribution

Particles are uniformly distributed in the cube at the start of all simulations shown in the manuscript. However, we have no reason to believe that such spatial distribution is more appropriate than a more clustered one to begin with. In Fig. S8, we show the final dominances obtained with and without advection, starting from a superposition of Thomas processes, i.e. each species was distributed with its own Thomas process.

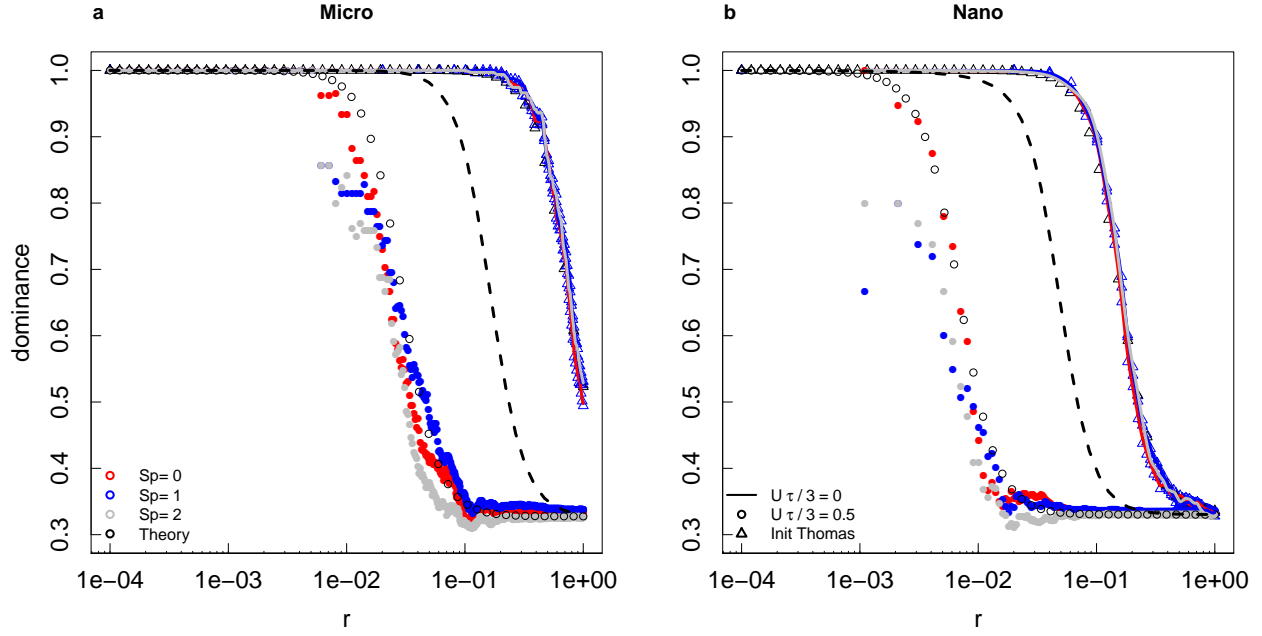


Figure S8: Dominance indices for microphytoplankton (a) and nanophytoplankton (b) in a 3-species community with even distributions after 1000 timesteps starting from a superposition of species-specific Thomas point processes, with (circles) and without (lines) advection. The triangles correspond to the dominance index at the start of the simulation, i.e. when points are only distributed according to a Thomas point process with parent intensity $\lambda_p = 200 \text{ cm}^{-3}$, number of children per parent $N_c = 50$, and $\sigma = 0.001$. Each color represents a different species. The black points and line corresponds to the theoretical values of the dominance index.

While there was a good fit between the expected dominance index and the simulated dominance index when advection was present, this was not the case when only diffusion and growth were taken into account. This might be due to the time needed to shift to the expected distribution from a non-uniform distribution (the nanophytoplankton dominance index finally reaches the expected value for a longer simulation duration), but even very long simulations, up to 10^6 timesteps did not lead to a match for microphytoplankton particles (Fig. S9).

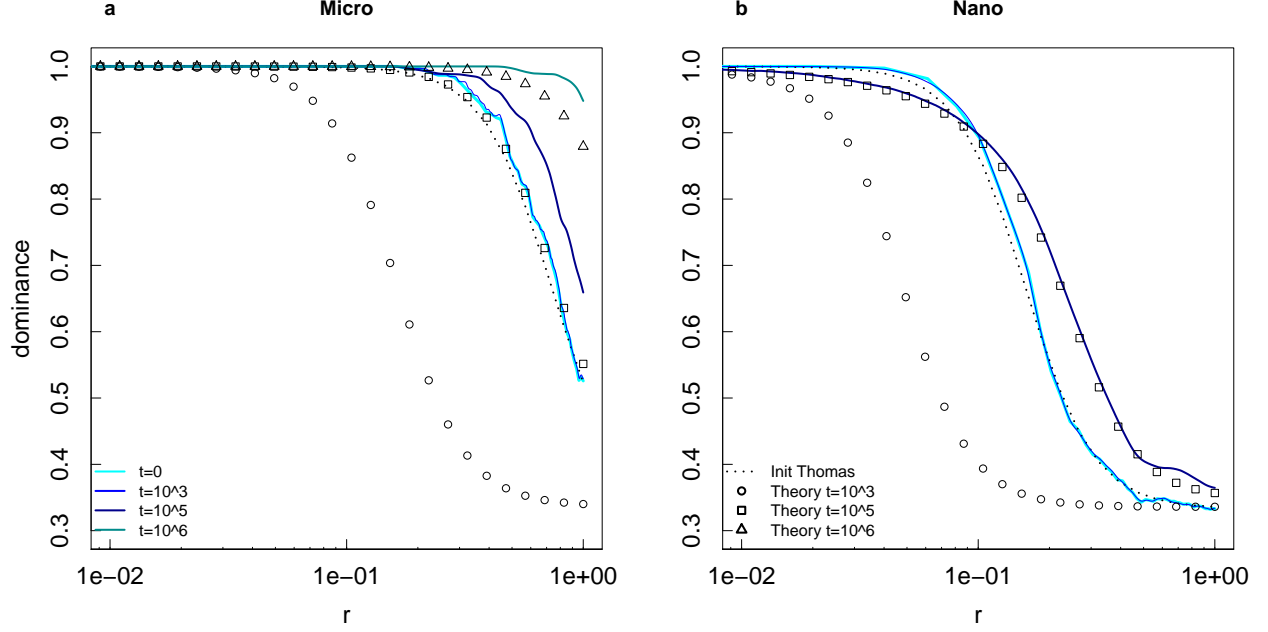


Figure S9: Dominance indices for one species in a microphytoplankton (a) and nanophytoplankton (b) 3-species community with even distributions after 1000 timesteps starting from a superposition of species-specific Thomas point processes in the absence of advection. The dotted line corresponds to the dominance index at the start of the simulation, i.e. when points are only distributed according to a Thomas point process with parent intensity $\lambda_p = 200 \text{ cm}^{-3}$, number of children per parent $N_c = 50$, and $\sigma = 0.001$. Each colored line represents a different simulation duration, and black symbols correspond to the theoretical values of the dominance index.

S7 Sensitivity to the computation of the advection parameter

If we define a length L_v corresponding to the equivalent sphere diameter, i.e.:

$$\begin{aligned}
 \frac{4}{3}\pi \left(\frac{L_v}{2}\right)^3 &= L_c^3 \\
 \Leftrightarrow L_v &= 2L_c \left(\frac{3}{4\pi}\right)^{1/3} \\
 \Leftrightarrow L_v &= 1.24 \text{ cm}
 \end{aligned} \tag{11}$$

If we use $U \approx \nu/L_v$, $U \approx 8.1 \times 10^{-5} \text{ m.s}^{-1}$. Using $U\tau/3 = 0.5 \text{ cm}$, we have $\tau = 185 \text{ s} = 2.1 \times 10^{-3} \text{ d}$. This means that $\gamma = 164$.

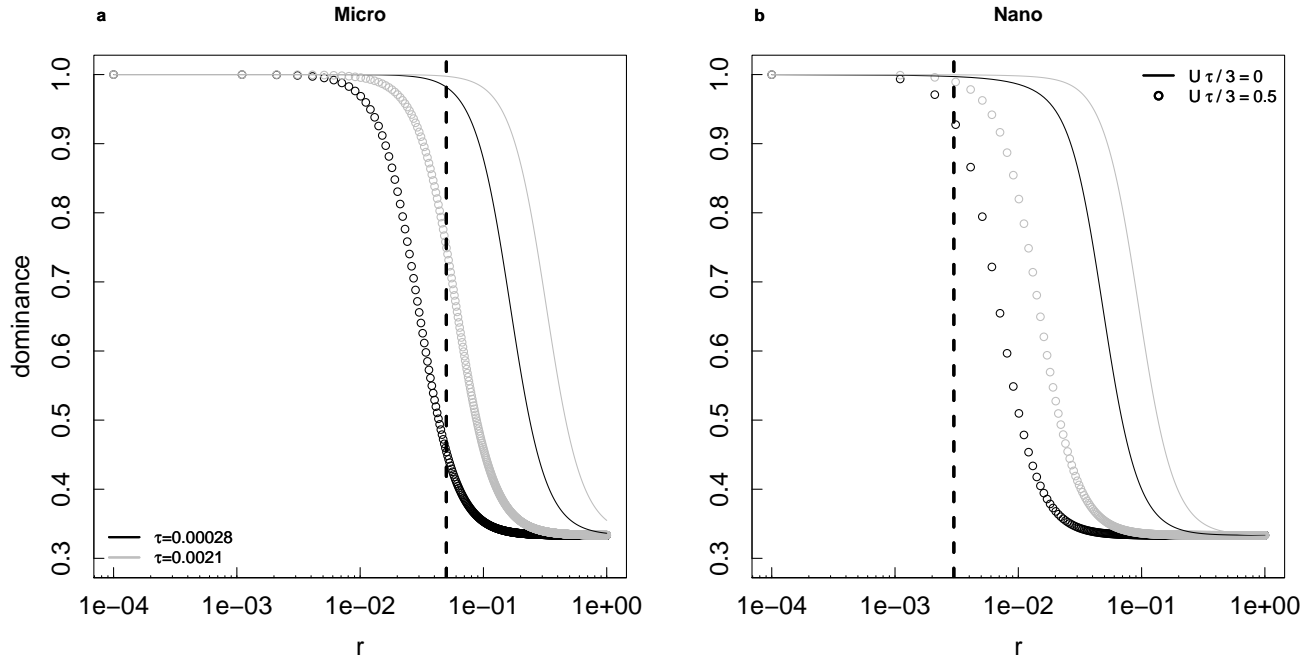


Figure S10: Dominance indices for one species in a microphytoplankton (a) and nanophytoplankton (b) 3-species community with even distributions after 1000 timesteps with (circles) and without (lines) advection for different duration of the timesteps.

References

- Baddeley, A., Rubak, E. & Turner, R. (2015). *Spatial Point Patterns: Methodology and Applications with R*. Chapman and Hall/CRC Press, London.
- Clark, P.J. & Evans, F.C. (1979). Generalization of a nearest neighbor measure of dispersion for use in k dimensions. *Ecology*, 60, 316–317.
- Ohser, J. (1983). On estimators for the reduced second moment measure of point processes. *Series Statistics*, 14, 63–71.