Universität Konstanz

III. Nonresponse

Session 6: Empirical strategies for missing data

Overview

- Introduction
- Assumptions about missingness
- Conventional strategies
- Information available about those missing
- Reweighting
- Hot decking
- Single imputation
- Multiple imputation
- Final remarks

Introduction

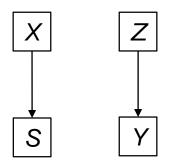
- Nonresponse introduces statistical inefficiency
- Nonresponse will also lead to biased estimates if the propensity to respond to the survey (or an item) is related to the attribute of interest
- Ex ante attempts to reduce unit and item nonreponse seldom are fully successful (e.g. increase and dispersion of contact attempts, incentives, conversion by specially trained interviewers, reduction of respondent burden / interview length via matrix sampling etc.)
- Nonresponse rate and nonresponse bias seem hardly related
- Importance of ex post strategies for missing data
- Different strategies are based on different assumptions about missingness and different kinds of information available on those missing

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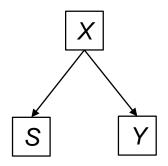
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Assumptions about missingness

Missing completely at random (MCAR)



Missing at random (MCAR)



Missing not at random (MNAR)



Conventional strategies

Listwise deletion, complete case analysis

- Default in many statistical software packages
- Loss of ~1/3 of the cases in typical regression models based on survey data
- Decrease in statistical efficiency
- "Listwise deletion is evil" (King 1998)
- Potential bias if MCAR does not hold
- If MAR holds, analysis including the relevant Xs yield unbiased results

Pairwise deletion

- Uses different sets of sample units for different params
- Nonconstant n, therefore difficult to obtain estimates of standard error / CIs of analytic statistics

Conventional strategies

Mean imputation

- Descriptive statistics: leaves \overline{y} unchanged, but decreases variances
- Overestimation of confidence
- Analytical statistics: decrease in correlations

Logical / best guess imputation

- E.g. due to filtering
- Eg. knowledge questions

Information about those missing

- Population information about the distribution of variables x presumedly responsible for missingness from censuses, other sources
- Individual measurements on nonrespondents from sampling frames, record linkages, screening interviews, interviewer observations, previous panels in a panel study, other items in the same survey (in instances of item nonresponses)
- Individual measurements on "unlikely" respondents using contact histories, reports of intentions to respond to a later survey, multi-phase sampling strategies

Poststratification

- Ex post stratification of the sample along one or several categorical variables x
 ...
 - ... which are suspected to co-determine missingness and the attribute of interest *y* ("common causes" according to Groves, 2006)
 - ... whose (joint) population distribution is known
- Conceive of the categories / cross-classifications of x as groups or strata h, use the stratified estimator
- Corrects for disproportionalities between $\frac{N_h}{N}$ and $\frac{n_h}{N}$ that may be either due to sampling or due to nonresponse
- Assumes MAR

Poststratification

Poststratified estimator of \(\overline{Y} \):

$$\bar{y}_{PS} = \sum_{h=1}^{H} \frac{N_h}{N} \bar{y}_h'$$

The sampling variance of of the poststratified estimator is estimated by

$$V(\bar{y}_{PS}) = \sum_{h=1}^{H} \left(\frac{N_h}{N}\right)^2 \frac{\sum_{i=1}^{n_h} (y_i' - \bar{y}_h')^2}{n_h' - 1} \frac{N_h - n_h'}{N_h}$$

Poststratification:

- Inclusion probability: $\pi_i = \frac{n_h'}{N_h} = \frac{n_h}{N_h} \frac{n_h'}{n_h}$
- Combined sampling and response weight: $\omega_i = \frac{1}{\pi_i}$
- Poststratified estimator:

$$\bar{y}_{PS} = \sum_{h=1}^{H} \frac{N_h}{N} \bar{y}'_h = \frac{1}{N} \sum_{h=1}^{H} \sum_{i=1}^{n_h} \frac{y'_i}{\pi_i}$$

Raking

- Poststratification requires that N_h is known and that $n_h^{'} \neq 0$ for all h
- This is rarely met for joint distributions of several X
- Situation resembles an ecological inference problem where the margins of a contingency table are known and the inner cells are unknown:

	$X_2 = 1$	$X_2 = 0$	
$X_1 = 1$?	?	P_1
$X_1 = 0$?	?	$1-P_1$
	P_2	$1 - P_2$	1

	$x_2 = 1$	$x_2 = 0$	
$x_1 = 1$	✓	✓	p_1
$x_1 = 0$	✓	✓	$1 - p_1$
	p_2	$1 - p_2$	1

• Iterative proportional fitting of the cell weights ω_h so that the marginal distributions in the sample approximate those in the population

Hot decking

- Requires individual measurements of x on nonrespondents
- Match individuals with missing y-values to one or more individuals for which y
 is observed based on x; impute (average) y-value
- Nonparametric; considers response patterns

Potential problems

- Assumes MAR
- What if there are no identical individuals in the data set?
- How to incorporate matching variance?
- How to incorporate imputation uncertainty?
- What if x also contains missing values?

Hot decking

Potential solutions

- Helpful if there are no identical individuals in the data set
- Model the response indicator using x as covariates (e.g. logit, probit)
- Matching based on predicted response probabilities
- Alternatively, use inverse predicted response probabilities as weights

Single imputation

Deterministic regression-based (conditional mean) imputation

(Linear) regression of y on x using complete cases:

$$y_{i}^{'} = \alpha + \beta x_{i}^{'} + e_{i}$$
$$e_{i} \sim Normal(0, \sigma_{e}^{2})$$

 Impute prediced y-values for nonrespondents using known x-values and estimated regression parameters:

$$\hat{y}_i = \hat{\alpha} + \widehat{\beta} x_i$$

Single imputation

Potential problems

- Assumes MAR
- Regression model needs to be specified correctly, potentially different distributions of x among respondents and nonrespondents
- Decreases variance in y, increases correlations between imputed y and x
- How to incorporate prediction uncertainty?
- How to incorporate imputation uncertainty?

Single imputation

Stochastic regression-based imputation

- Easy way to incorporate prediction uncertainty
- As previously, linear regression of y on x using complete cases
- Random draws of \hat{e}_i from Normal(0, $\hat{\sigma}_e^2$)
- Impute predicted y-values for nonrespondents using known x-values and estimated regression parameters plus ê_i: ŷ_i=α̂+β̂x_i+ê_i

Potential problems

- Assumes MAR
- Regression model needs to be specified correctly, potentially different distributions of x among respondents and nonrespondents
- How to incorporate imputation uncertainty?

Survey Methodology

Multiple imputation

- Way to also incorporate imputation uncertainty
- As previously, but multiple stochastic imputations m=1,2,..., M
- Combination of estimates:

$$\hat{y}_i = \frac{1}{M} \sum_{m=1}^{M} \hat{y}_{im}$$

- Sample mean of observed and (average) imputed *y*-values, \bar{y}_{MI} , is an unbiased and consistent estimate of \bar{Y} (assuming SRS)
- The combined sampling and imputation variance of this estimate is given by

$$V(\bar{y}_{MI}) = \frac{1}{M} \sum_{m=1}^{M} \frac{\sum_{i=1}^{n} (y_i - \bar{y}_m)^2}{n-1} \frac{N-n}{N} + \left(1 + \frac{1}{M}\right) \frac{1}{M-1} \sum_{m=1}^{M} (\bar{y}_m - \bar{y})^2$$

Final remarks

- Before applying a strategy for missing data, we ought to ask whether an underlying "true" value exists and, if so, whether the missing value is truly unknown
- If individual measurements for nonrespondents are available, multiple imputation (MI) is the current gold standard for handling nonresponse, as it incorporates both prediction and imputation uncertainty
- MI makes data analysis more complex (we have to handle m data sets)
- Model specification is still a major concern, also MI assumes MAR
- If the missingness process is suspected to be non-ignorable (NMAR), one has to refrain to selection models or pattern-mixture models
- Some NMAR problems could perhaps be converted into MAR problems by collecting additional data on response propensities

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