

Note: No lecture on  
Wed Oct 12.

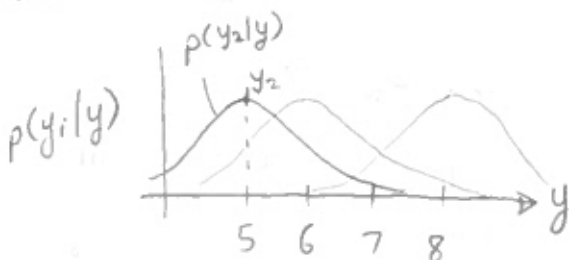
On Fri Oct 14 I will do  
more blackboard+lab, but I am  
going to try preparing a notebook  
ahead of time with the overall  
structure but with places where you  
have to fill in code.

We will do:

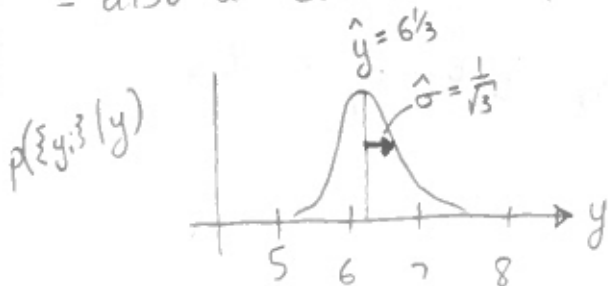
- handling outliers via foreground-background modeling
- intro to Markov Chain Monte Carlo
- using Julia's nonlinear optimizer to find solutions to this model

## Review & Comments from last time

Likelihoods of our observations



Likelihood of getting the 3 observations we got  
= product of 3 individual likelihoods (Gaussian)  
= also a Gaussian ( $\leftarrow$  magical!)



Then we used Bayes' rule to take

$$p(\{y_i\}_3 | y) \text{ to } p(y | \{y_i\}_3)$$

(which is shorthand for  $p(y | \text{observation and all our assumptions})$ )

## Comments

- probabilities : "beliefs", "knowledge",  
"uncertainty"

Bayes' rule: math for updating our statement of our knowledge.

You take a new obs.  $y_{\text{new}}$

$$p(y|y_{\text{new}}) = \frac{p(y_{\text{new}}|y) p(y)}{p(y_{\text{new}})}$$

Implicitly,

Include "everything you already know",  $E$ :

$$p(y|y_{\text{new}}, E) = \frac{p(y_{\text{new}}|y, E) p(y, E)}{p(y_{\text{new}}, E)}$$

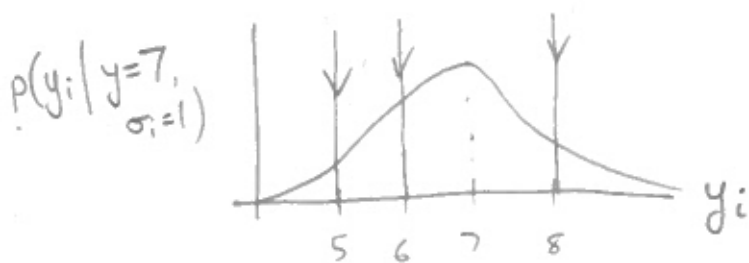
" $E$ " can include previous measurements!

$$p(y|y_{\text{new}}) \propto p(y_{\text{new}}|y) \cdot N(\hat{y}, \hat{\sigma}^2)$$

Updates our statement of what we know/believe about the value of  $y$ .

~~"you don't know"~~

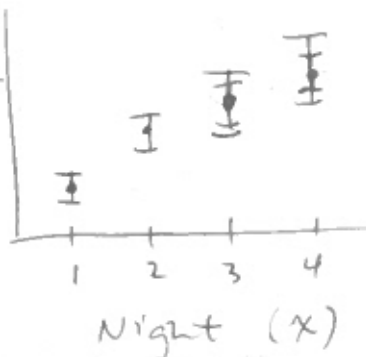
- imagine you know  $y$ . Then can plot:



"you don't know where your observation sits  
on this curve"

## Fitting a line

Night	Pos	$\sigma$	Position y
1	14	1	
2	26	1	
3	37	2	
4	43	2	



Model:

$$y = mx + b$$

Assumptions:

- uncertainties in  $x$  are negligible
- measured  $y_i = m x_i + b + e_i$
- $e_i$  are iid. drawn from  $N(0, \sigma_i^2)$

As before, can write likelihood of the observations given parameters:

$$p(\underbrace{\{y_i\}}_Y | m, b) = \prod_i \frac{1}{\sqrt{2\pi}\sigma_i} \exp\left(-\frac{1}{2} \frac{(mx_i + b - y_i)^2}{\sigma_i^2}\right)$$

$$\log \mathcal{L} = C - \frac{1}{2} \sum \frac{(mx_i + b - y_i)^2}{\sigma_i^2}$$

could take derivatives to find  $m, b$  as before, but that gets messy.

Instead, take the  $\chi^2$  term

$$\chi^2 = \sum_i \frac{(mx_i + b - y_i)^2}{\sigma_i^2}$$

and write as matrix math

$$Y = \begin{bmatrix} y_1 \\ y_2 \\ \vdots \end{bmatrix} \quad A = \begin{bmatrix} 1 & x_1 \\ 1 & x_2 \\ \vdots & \vdots \end{bmatrix} \quad W = \begin{bmatrix} \frac{1}{\sigma_1^2} & & 0 \\ & \frac{1}{\sigma_2^2} & \\ 0 & & \ddots \end{bmatrix}$$

$$X = \begin{bmatrix} b \\ m \end{bmatrix} \quad \leftarrow \text{awkward change of convention on "x"!$$

$$\chi^2 = (AX - Y)^T W (AX - Y)$$

$$\chi^2 = [W^{1/2}(AX - Y)]^T [W^{1/2}(AX - Y)]$$

This is a matrix least-squares problem!

Julia has a built-in

"\ " operator to solve this:

$$r = \hat{X} = \underset{X^T X}{\operatorname{argmin}} \left( r = (W^{1/2} A) X - (W^{1/2} Y) \right)$$

min

$$\hat{X} = (W^{1/2} A) \setminus (W^{1/2} Y)$$

We want to find a "least squares" solution:

$$X =$$

$$\text{cov} = (A^T W A)^{-1}$$

What this solves is

$$\hat{X} = (A^T W A)^{-1} (A^T W Y)$$

## Lab exercises

- set up this matrix equation using our data set, and solve it with and without the weight matrix  $W$ .
- plot the data points (with error bars) and the weighted and unweighted fit lines.
- compute the covariance of the maximum likelihood  $m$  and  $b$  values
- plot, in the  $m, b$  plane, the best-fit (weighted and unweighted) values, and a sampling of values from the covariance matrix.
- plot, in  $x, y$  space, a sampling of lines drawn from the best-fit  $m, b$  plus covariance.
- optional: fit a quadratic



- write the log-likelihood function for our data set given values of  $m, b$ .
- evaluate this function on a grid of  $m, b$  values
- plot this log-likelihood surface (eg, as a heatmap)