SYSTEM ESTIMATION METHODS II: STRUCTURED ESTIMATION

Mathematical Models can be derived from

- Physical Modeling (Analytic approach)
- Identification (Experimental approach)

CLASSIFICATION OF MODELS

- SISO v/s MIMO
- Linear v/s Nonlinear
- Parametric v/s Nonparametric
- Time invariant v/s Time variant
- Time domain v/s Frequency domain
- Discrete time v/s Continuous time
- Deterministic v/s Stochastic

GENERAL LTI-SISO MODEL STRUCTURE (BOX-JENKINS)

$$y_{t} = G(q;\theta)u_{t} + H(q;\theta)w_{t}$$

$$G(q;\theta) = \frac{B(q)}{A(q)} = \frac{b_{1}q^{-nk} + \dots + b_{nb}q^{-nk-nb+1}}{1 + a_{1}q^{-1} + \dots + a_{na}q^{-na}}$$

$$H(q;\theta) = \frac{C(q)}{D(q)} = \frac{1 + c_{1}q^{-1} + \dots + c_{nc}q^{-nc}}{1 + d_{1}q^{-1} + \dots + a_{nd}q^{-nd}}$$

where $\{w_t\}$ is white noise of variance λ^2 , $\{u_t\}$ is the input, and

$$\theta = [a_1 \quad \cdots \quad a_{na} \quad b_1 \quad \cdots \quad b_{nb} \quad c_1 \quad \cdots \quad c_{nc} \quad d_1 \quad \cdots \quad d_{nd}]^T$$

- Time delay $n_k \ge 1 \implies G(\infty; \theta) = 0$ (also, $H(\infty; \theta) = 1$)
- $H^{-1}(q;\theta)$ and $H^{-1}(q;\theta)G(q;\theta)$ are asymptotically stable

Often, $H(q;\theta)$ is also required to be asymptotically stable

TIME SERIES MODELS (WITHOUT CONTROLLABLE INPUT)

AR
$$A(q)y_t = w_t$$

MA $y_t = C(q)w_t$

ARMA $A(q)y_t = C(q)w_t$

BASIC MODELS FOR CONTROL

ARX
$$A(q)y_{t} = B(q)u_{t} + w_{t}$$
ARMAX
$$A(q)y_{t} = B(q)u_{t} + C(q)w_{t}$$
FIR
$$y_{t} = B(q)u_{t} + w_{t}$$
OE
$$y_{t} = \frac{B(q)}{A(q)}u_{t} + w_{t}$$

- Selection of model structure depends on prior knowledge and estimation method
- Preferable: model structure contains true system (no undermodeling)
- Preferable: different models predict different results (identifiability)

STATE SPACE MODELS

$$x_{t+1} = A(\theta)x_t + B(\theta)u_t + w_t$$
$$y_t = C(\theta)x_t + D(\theta)u_t + v_t$$

Advantages:

- Physical Insight (physical parameters)
- Natural extension to MIMO systems

Problem: Over-parameterization

$$\overline{x}_{t+1} = TA(\theta)T^{-1}\overline{x}_t + TB(\theta)u_t + Tw_t$$
$$y_t = C(\theta)T^{-1}\overline{x}_t + D(\theta)u_t + v_t$$

gives the same input-output relation!

PREDICTION ERROR FORMULATION

$$y_{t} = f(Y_{t-1}, U_{t}, t; \theta) + \varepsilon_{t}$$

where $Y_{t-1} := \{y_{t-1}, \dots, y_1\}, U_t := \{u_t, \dots, u_1\}$ and $\{\varepsilon_t\}$ is an *innovations sequence*, i.e.

$$E\{\varepsilon_t \mid Y_{t-1}, U_t\} = 0$$

Important Case: LTI models with wide-sense stationary disturbances

$$\begin{aligned} y_t &= G(q; \theta) u_t + H(q; \theta) w_t \\ &= \underbrace{[I - H^{-1}(q; \theta)] y_t + H^{-1}(q; \theta) G(q; \theta) u_t}_{\hat{y}_{t|t-1}(\theta) := E\{y_t | Y_{t-1}, U_{t-1}\}} + \underbrace{w_t}_{\mathcal{E}_t} \end{aligned}$$

ML Estimation for the Prediction Error Formulation

If $\{\varepsilon_t\}$ are IID Gaussian with (known) covariance Σ , the ML estimator is

$$\hat{\theta}_{ML} = \arg \max_{\theta \in \Theta} l(\theta) = \arg \min_{\theta \in \Theta} \sum_{t=1}^{N} \varepsilon_{t}^{T}(\theta) \Sigma^{-1} \varepsilon_{t}(\theta)$$

Interpretation:

 $\hat{ heta}_{ ext{ iny ML}}$ tries to make the prediction errors "small"

⇒ *Prediction Error Methods* (PEM):

$$\hat{\theta}_{PEM} = \arg\min_{\theta \in \Theta} f(\{\varepsilon_t(\theta)\}; \theta)$$

Typical cost functions: $\frac{1}{2N} \sum_{t=1}^{N} \varepsilon_{t}^{2}(\theta) \ (SISO) \quad \text{or} \quad \det \left[\frac{1}{N} \sum_{t=1}^{N} \varepsilon_{t}(\theta) \varepsilon_{t}^{T}(\theta) \right] (MIMO)$

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LINEAR REGRESSION

For some model structures (e.g. ARX, FIR), the predictor is linear in θ :

$$\hat{\mathbf{y}}_{t|t-1} = \boldsymbol{\varphi}_t^T \boldsymbol{\theta}$$

Then, the PEM cost function is a *least-squares (LS) criterion*:

$$V_N(\theta) := \frac{1}{2N} \sum_{t=1}^N [y_t - \varphi_t^T \theta]^2$$

The minimizer of $V_N(\theta)$ has an explicit expression:

$$\hat{\theta}_{LS} := \arg\min_{\theta} V_N(\theta) = \left[\frac{1}{N} \sum_{t=1}^N \varphi_t \varphi_t^T \right]^{-1} \frac{1}{N} \sum_{t=1}^N \varphi_t y_t$$

LINEAR REGRESSION (CONT.)

Properties:

If $y_t = \varphi_t^T \theta_0 + w_t$, and the signals are quasi-stationary, then

$$\hat{\theta}_{LS} - \theta_0 = \left[\frac{1}{N} \sum_{t=1}^{N} \varphi_t \varphi_t^T\right]^{-1} \frac{1}{N} \sum_{t=1}^{N} \varphi_t w_t \xrightarrow{a.s.} [R^*]^{-1} f^*$$

where

$$R^* = \overline{E}\{\varphi_t \varphi_t^T\}, \qquad f^* = \overline{E}\{\varphi_t w_t\}$$

Therefore, LS is consistent if:

- R^* is nonsingular (persistence of excitation and identifiability)
- $f^* = 0$, i.e., $\{w_t\}$ is uncorrelated with $\{\varphi_t\}$ (e.g. if $\{w_t\}$ is white noise, or for FIR)

LINEAR REGRESSION (CONT.)

Example: ARX

$$A(q)y_t = B(q)u_t + w_t \qquad \Rightarrow \qquad \hat{y}_{t|t-1} = [1 - A(q)]y_t + B(q)u_t$$

Then, $\varphi_t = \begin{bmatrix} -y_{t-1} & \cdots & -y_{t-na} & u_{t-1} & \cdots & u_{t-nb} \end{bmatrix}^T$

and $\theta = [a_1 \quad \cdots \quad a_{na} \quad b_1 \quad \cdots \quad b_{nb}]^T$

What happens if the noise is colored, e.g. $A(q)y_t = B(q)u_t + \frac{1}{D(q)}w_t$?

One possibility: High-order ARX model

$$A(q)D(q)y_t = B(q)D(q)u_t + w_t$$

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THE CORRELATION APPROACH (INSTRUMENTAL VARIABLES)

In LS: $\frac{1}{N} \sum_{t=1}^{N} \varphi_t [y_t - \varphi_t^T \hat{\theta}_{LS}] = 0$

There is a problem when $\{\varphi_t\}$ is correlated with $\{w_t\}$

Idea: Instrumental Variables

$$\frac{1}{N} \sum_{t=1}^{N} \zeta_t [y_t - \varphi_t^T \hat{\theta}_{IV}] = 0 \qquad \Rightarrow \qquad \hat{\theta}_{IV} = \left[\frac{1}{N} \sum_{t=1}^{N} \zeta_t \varphi_t^T \right]^{-1} \frac{1}{N} \sum_{t=1}^{N} \zeta_t y_t$$

where for consistency:

• $\{\zeta_t\}$ highly correlated with $\{u_t\}$ $(\overline{E}\{\zeta_t\varphi_t^T\}$ nonsingular)

• $\{\zeta_t\}$ uncorrelated with $\{w_t\}$ $(\overline{E}\{\zeta_t w_t\} = 0)$

Good choices: past inputs, noiseless outputs (from LS estimation)