## SYSTEM ESTIMATION METHODS III: SUBSPACE IDENTIFICATION

# **ADVANCED TOPICS**

Today we will see a more detailed derivation of the subspace methods

## **REMINDER: PRELIMINARIES**

System (MIMO):

$$x_{t+1} = Ax_t + Bu_t + w_t$$
$$y_t = Cx_t + Du_t + v_t$$

Input:  $u_t \in \mathbb{R}^m$  Process noise:  $w_t \in \mathbb{R}^n$ 

State:  $x_t \in \mathbb{R}^n$  Measurement noise:  $v_t \in \mathbb{R}^p$ 

Output:  $y_t \in \mathbb{R}^p$ 

Assumptions:  $\{w_t\}$  and  $\{v_t\}$  are white noise sequences

There are no constraints on A, B, C, D

#### **REMINDER: PREDICTORS**

$$Y_t^r = O^r x_t + S^r U_t^r + V_t$$
 Fundamental Equation

where

$$Y^{r} \coloneqq \begin{bmatrix} y_{t} \\ y_{t+1} \\ \vdots \\ y_{t+r-1} \end{bmatrix}, \quad U^{r} \coloneqq \begin{bmatrix} u_{t} \\ u_{t+1} \\ \vdots \\ u_{t+r-1} \end{bmatrix}, \quad V_{t} \coloneqq \begin{bmatrix} v_{t} \\ Cw_{t} + v_{t+1} \\ \vdots \\ CA^{r-2}w_{t} + CA^{r-3}w_{t+1} + \dots + Cw_{t+r-2} + v_{t+r-1} \end{bmatrix}$$

$$O^r \coloneqq egin{bmatrix} C \ CA \ dots \ CA^{r-1} \end{bmatrix}, \qquad S^r \coloneqq egin{bmatrix} D & 0 & \cdots & 0 & 0 \ CB & D & \cdots & 0 & 0 \ dots & dots & \ddots & dots & dots \ CA^{r-2}B & CA^{r-3}B & \cdots & CB & D \end{bmatrix}$$

# ESTIMATION OF THE (EXTENDED) OBSERVABILITY MATRIX

This equation can be written as

$$\mathbf{Y} = O^r \mathbf{X} + S^r \mathbf{U} + \mathbf{V}$$

where

$$\mathbf{Y} \coloneqq [Y_1^r \quad \cdots \quad Y_N^r]$$

$$\mathbf{X} \coloneqq [x_1 \quad \cdots \quad x_N]$$

$$U \coloneqq [U_1^r \quad \cdots \quad U_N^r]$$

 $\mathbf{V} := [V_1 \quad \cdots \quad V_N]$ 

Objective: To estimate  $O^r \mathbf{X}$ , given data  $\mathbf{U}$  and  $\mathbf{Y}$ .

**Remark:** A state transformation  $A \to T^{-1}AT$ ,  $C \to CT$  changes  $O^r$  to  $O^rT$ . Thus, postmultiplying  $O^r$  by an invertible T simply changes the resulting realization

# ESTIMATION OF THE (EXTENDED) OBSERVABILITY MATRIX (CONT.)

To eliminate **U**, post-multiply by a *projector*  $\Pi_{\mathbf{U}^{\perp}}^{\perp} := I - \mathbf{U}^{T} (\mathbf{U}\mathbf{U}^{T})^{-1}\mathbf{U}$ , giving

$$\mathbf{Y}\Pi_{\mathbf{U}^{\perp}}^{\perp} = O^{r}\mathbf{X}\Pi_{\mathbf{U}^{\perp}}^{\perp} + S^{r}\mathbf{U}\Pi_{\mathbf{U}^{\perp}}^{\perp} + \mathbf{V}\Pi_{\mathbf{U}^{\perp}}^{\perp} = O^{r}\mathbf{X}\Pi_{\mathbf{U}^{\perp}}^{\perp} + \mathbf{V}\Pi_{\mathbf{U}^{\perp}}^{\perp}$$

The noise term,  $\mathbf{V}\Pi_{\mathbf{U}^{\perp}}^{\perp}$ , can be eliminated using *instrumental variables*, i.e., post-multiplying by a matrix  $\Phi^T := [\varphi_1^s \cdots \varphi_N^s]^T \in \mathbb{R}^{N \times s}$ , so that

$$\underbrace{\frac{1}{N}\mathbf{Y}\Pi_{\mathbf{U}^{\perp}}^{\perp}\Phi^{T}}_{G} = O^{r}\underbrace{\frac{1}{N}\mathbf{X}\Pi_{\mathbf{U}^{\perp}}^{\perp}\Phi^{T}}_{\tilde{T}_{N}} + \underbrace{\frac{1}{N}\mathbf{V}\Pi_{\mathbf{U}^{\perp}}^{\perp}\Phi^{T}}_{V_{N}} \xrightarrow{N \to \infty} O^{r}\tilde{T}$$

# ESTIMATION OF THE (EXTENDED) OBSERVABILITY MATRIX (CONT.)

Therefore, we need

$$\begin{split} 0 &= \lim_{N \to \infty} V_{N} \\ &= \lim_{N \to \infty} \frac{1}{N} \mathbf{V} \Pi_{\mathbf{U}^{\perp}}^{\perp} \Phi^{T} \\ &= \lim_{N \to \infty} \frac{1}{N} \sum_{t=1}^{N} V_{t} (\varphi_{t}^{s})^{T} - \lim_{N \to \infty} \frac{1}{N} \sum_{t=1}^{N} V_{t} (U_{t}^{r})^{T} \left[ \frac{1}{N} \sum_{t=1}^{N} U_{t}^{r} (U_{t}^{r})^{T} \right]^{-1} \frac{1}{N} \sum_{t=1}^{N} U_{t}^{r} (\varphi_{t}^{s})^{T} \\ &= \overline{E} \{ V_{t} (\varphi_{t}^{s})^{T} \} - \overline{E} \{ V_{t} (U_{t}^{r})^{T} \} \left[ \overline{E} \{ U_{t}^{r} (U_{t}^{r})^{T} \} \right]^{-1} \overline{E} \{ U_{t}^{r} (\varphi_{t}^{s})^{T} \} \end{split}$$

In open loop,  $\{U_t^r\}$  is independent of  $\{V_t\}$ , so  $\overline{E}\{V_t(U_t^r)^T\} = 0$ . The first term is 0 if we build  $\varphi_t^s$  from past data, e.g.,

$$\boldsymbol{\varphi}_{t}^{s} = [y_{t-1} \quad \cdots \quad y_{t-s_{1}} \quad \boldsymbol{u}_{t-1} \quad \cdots \quad \boldsymbol{u}_{t-s_{2}}]^{T}$$

# ESTIMATION OF THE (EXTENDED) OBSERVABILITY MATRIX (CONT.)

We also need  $\tilde{T} = \lim_{N \to \infty} N^{-1} \mathbf{X} \Pi_{\mathbf{U}^{\perp}}^{\perp} \Phi^{T}$  nonsingular. For the previous choice of  $\varphi_{t}^{s}$  this holds under some conditions (see Problem 10G.6 of Ljung)

#### **ESTIMATION OF THE ORDER**

If we have a noisy estimation  $G = O^r T + E_N$ , where  $E_N$  is small, then rank $\{O^r\}$  can be estimated via SVD:

$$G = USV^{T} = U \begin{bmatrix} \sigma_{1} & 0 & \cdots & 0 \\ 0 & \sigma_{2} & \cdots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \cdots & \sigma_{n^{*}} \\ 0 & 0 & \cdots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \cdots & 0 \end{bmatrix} V^{T}$$

The smaller  $\sigma_i$ 's ( $<\varepsilon$  for some predetermined  $\varepsilon>0$ ) can be replaced by 0, thus replacing G by a lower rank matrix  $G_1=U_1S_1V_1^T$ . From a previous remark, only  $U_1$  is important for estimating A,B,C,D

## **ESTIMATION OF THE ORDER (CONT.)**

Many methods using weighting matrices for the SVD step, i.e.,

$$W_1 G W_2 = U S V^T \approx U_1 S_1 V_1^T$$

and then consider:

$$\hat{O}^r = W_1^{-1} U_1 R$$

where R is an arbitrary matrix (to determine a particular state realization)  $W_2$  corresponds to a state transformation

 $W_1$  only affects  $\hat{O}^r$  when there is noise, so it affects the quality of  $\hat{A}, \hat{C}$ 

| <b>Typical choices:</b> | $R = I$ , $R = S_1$ or $R = S_1^{1/2}$ |
|-------------------------|--|
| LODGD                   |  |

| MOESP | $W_1 = I, W_2 = (N^{-1}\Phi\Pi_{U^T}^{\perp}\Phi^T)^{-1}\Phi\Pi_{U^T}^{\perp}$               |
|-------|--|
| N4SID | $W_1 = I, W_2 = (N^{-1}\Phi\Pi_{U^T}^{\perp}\Phi^T)^{-1}\Phi$                                |
| IVM   | $W_1 = (N^{-1}Y\Pi_{U^T}^{\perp}Y)^{-1/2}, W_2 = (N^{-1}\Phi\Phi^T)^{-1/2}$                  |
| CVA   | $W_1 = (N^{-1}Y\Pi_{U^T}^{\perp}Y)^{-1/2}, W_2 = (N^{-1}\Phi\Pi_{U^T}^{\perp}\Phi^T)^{-1/2}$ |

### ESTIMATION OF A AND C

A and C can be estimated from the (extended) observability matrix

$$O^r \coloneqq \begin{bmatrix} C \\ CA \\ \vdots \\ CA^{r-1} \end{bmatrix}$$

by solving the equations:

$$\hat{C} = O^{r}(1: p, 1: n)$$

$$O^{r}(p+1: pr, 1: n) = O^{r}(1: p(r-1), 1: n)\hat{A}$$

#### ESTIMATION OF B AND D

Given  $\hat{A}$  and  $\hat{C}$ , we can estimate B and D (and the initial state  $x_0$ ) via LS from:

$$y_{t} = \hat{C}(qI - \hat{A})^{-1} \mathbf{x}_{0} \delta_{t} + \hat{C}(qI - \hat{A})^{-1} \mathbf{B} u_{t} + \mathbf{D} u_{t} + \varepsilon_{t}$$

where  $\varepsilon_t := \hat{C}(qI - \hat{A})^{-1}w_t + v_t$ .

*Remark:* It is possible to find the state  $x_t$ , and from this to estimate the noise statistics. For more details, see Ljung, pp. 348-349.

### **SUMMARY OF SUBSPACE METHODS**

- 1. From data, form  $G = N^{-1} \mathbf{Y} \Pi_{\mathbf{U}^T}^{\perp} \Phi^T$
- 2. Choose  $W_1$ ,  $W_2$  and perform SVD:  $\hat{G} = W_1 G W_2 = U S V^T \approx U_1 S_1 V_1^T$
- 3. Select R and define  $\hat{O}^r = W_1^{-1}U_1R$ , from which estimate  $\hat{A}$ ,  $\hat{C}$  via

$$\hat{C} = O^{r}(1:p,1:n)$$

$$O^{r}(p+1:pr,1:n) = O^{r}(1:p(r-1),1:n)\hat{A}$$

4. Estimate  $\hat{B}$ ,  $\hat{D}$  via LS from

$$y_{t} = \hat{C}(qI - \hat{A})^{-1} \mathbf{x}_{0} \delta_{t} + \hat{C}(qI - \hat{A})^{-1} \mathbf{B} u_{t} + \mathbf{D} u_{t} + \varepsilon_{t}$$