# Bayesian multimodeling: Variational inference-2

MIPT

2023

# Model selection: coherent Bayesian inference

First level: find optimal parameters:

$$\mathbf{w} = \arg\max \frac{p(\mathfrak{D}|\mathbf{w})p(\mathbf{w}|\mathbf{h})}{p(\mathfrak{D}|\mathbf{h})},$$

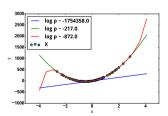
Second level: find optimal model:

Evidence:

$$p(\mathfrak{D}|\mathbf{h}) = \int_{\mathbf{w}} p(\mathfrak{D}|\mathbf{w}) p(\mathbf{w}|\mathbf{h}) d\mathbf{w}.$$



Model selection scheme



Polynomial regression example

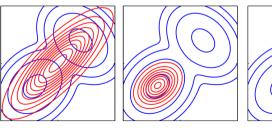
## Evidence lower bound, ELBO

Evidence lower bound is a method of approximation of intractable distribution  $p(\mathbf{w}|\mathfrak{D}, \mathbf{h})$  with a distribution  $q(\mathbf{w}) \in \mathfrak{Q}$ .

Evidence lower bound estimation often reduces to optimization problem

$$\log p(\mathfrak{D}|\mathbf{h}) \ge \mathsf{KL}(q(\mathbf{w})||p(\mathbf{w}|\mathfrak{D})) =$$

$$= -\int_{\mathbf{w}} q(\mathbf{w}) \log \frac{p(\mathbf{w}|\mathfrak{D})}{q(\mathbf{w})} d\mathbf{w} = \mathsf{E}_{\mathbf{w}} \log p(\mathfrak{D}|\mathbf{w}) - \mathsf{KL}(q(\mathbf{w})||p(\mathbf{w}|\mathbf{h})).$$







0.4



Variational inference vs. expectation propogation (Bishop)

## **ELBO** estimation

#### **ELBO** maximization

$$\int_{\mathbf{w}} q(\mathbf{w}) \log \frac{p(\mathbf{y}, \mathbf{w} | \mathbf{X}, \mathbf{h})}{q(\mathbf{w})} d\mathbf{w}$$

is equivalent to KL-divergence minimization between  $q(\mathbf{w}) \in \mathfrak{Q}$  and posteriod distribution  $p(\mathbf{w}|\mathbf{y},\mathbf{X},\mathbf{h})$ :

$$\begin{split} \hat{q} &= \operatorname*{arg\,max}_{q \in \mathfrak{Q}} \int_{\mathbf{w}} q(\mathbf{w}) \log \, \frac{p(\mathbf{y}, \mathbf{w} | \mathbf{X}, \mathbf{h})}{q(\mathbf{w})} d\mathbf{w} \Leftrightarrow \\ \hat{q} &= \operatorname*{arg\,min}_{q \in \mathfrak{Q}} \mathrm{D}_{\mathsf{KL}} \big( q(\mathbf{w}) || p(\mathbf{w} | \mathbf{y}, \mathbf{X}, \mathbf{h}) \big), \\ \mathrm{D}_{\mathsf{KL}} \big( q(\mathbf{w}) || p(\mathbf{w} | \mathbf{y}, \mathbf{X}, \mathbf{h}) \big) &= \int_{\mathbf{w}} q(\mathbf{w}) \log \left( \frac{q(\mathbf{w})}{p(\mathbf{w} | \mathbf{y}, \mathbf{X}, \mathbf{h})} \right) d\mathbf{w}. \end{split}$$

## Outline

- Can we use something except Gaussian distribution?
  - ► Yes, we can
- Does it need to have an analytical form?
  - ► No
- Does it need to have some specific properties except continuity?
  - ► No
- Do we need to optimize KL-divergence w.r.t. distribution parameters for ELBO estimation?
  - ► No
- Do we need to optimize ELBO for posterior approximation?
  - ► In general, no

# Reparametrization trick

Reparamterization idea:

$$arepsilon = S_{m{ heta}}(\mathbf{w}), \quad \mathbf{w} = S_{m{ heta}}^{-1}(arepsilon).$$

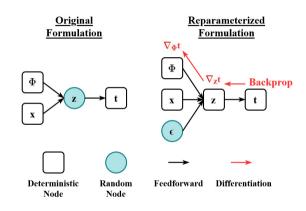
Then:

$$\nabla_{\boldsymbol{\theta}} \mathsf{E}_q f(\mathbf{w}) = \mathsf{E}_q \nabla_{\boldsymbol{\theta}} f(S_{\boldsymbol{\theta}}^{-1}(\varepsilon)).$$

#### Example:

$$w \sim \mathcal{N}(\mu, \sigma^2) \rightarrow S(w) = \frac{w - \mu}{\sigma} \sim \mathcal{N}(0, 1).$$

**Challenge:** calculation of  $S^{-1}$  is an expensive operation.



Source: wikipedia

# **Normalizing Flows**

Given an invertible smooth mapping  $\mathbf{g}$  (flow) and a distribution  $\mathbf{z} \sim q$ . Then  $q(\mathbf{g}(\mathbf{z}))$  is a distribution:

$$\mathbf{g}(g(\mathbf{z})) = q(\mathbf{z}) \left( \det \frac{\partial g}{\partial \mathbf{z}} \right)^{-1}.$$

Example: planar flow:

$$\mathbf{g}(\mathbf{z}) = \mathbf{z} + \mathbf{w}_1 \boldsymbol{\sigma}(\mathbf{w}_2^\mathsf{T} \mathbf{x}).$$

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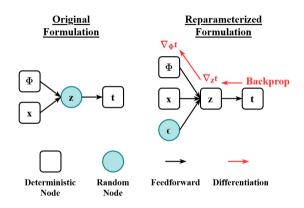
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Source: wikipedia

# Implicit reparametrization trick

$$\nabla_{\theta} E_q f(\mathbf{w}) = \mathsf{E}_q \nabla_{\mathbf{w}} f(\mathbf{w}) \nabla_{\theta} \mathbf{w}.$$

Use a total gradient formula for  $\varepsilon = S_{\theta}(\mathbf{w})$ :

$$abla_{\mathbf{w}}S_{m{ heta}}(\mathbf{w})
abla_{m{ heta}}\mathbf{w} + 
abla_{m{ heta}}S_{m{ heta}}(\mathbf{w}) = 0 
ightarrow$$

$$egin{aligned} igtarrow 
abla_{m{ heta}} \mathbf{w} &= -(
abla_{m{ well}} S_{m{ heta}}(\mathbf{w}))^{-1} 
abla_{m{ heta}} S_{m{ heta}}. \end{aligned}$$

Obtain an expression without inverse function for S.

For 1d samples we can use, for example:

$$S(\mathbf{w}) = F(\mathbf{w}|\boldsymbol{\theta}) \sim \mathcal{U}(0,1).$$

Table 4: Test negative log-likelihood (lower is better) for VAE on MNIST. Mean  $\pm$  standard deviation over 5 runs. The von Mises-Fisher results are from [9].

Prior	Variational posterior	D=2	D=5	D = 10	D = 20	D = 40
$\mathcal{N}(0,1)$	$\mathcal{N}(\mu, \sigma^2)$	$131.1 \pm 0.6$	$107.9 \pm 0.4$	$92.5 \pm 0.2$	$88.1 \pm 0.2$	$88.1 \pm 0.0$
Gamma(0.3, 0.3)	$Gamma(\alpha, \beta)$	$132.4 \pm 0.3$	$108.0 \pm 0.3$	$94.0 \pm 0.3$	$90.3 \pm 0.2$	$90.6 \pm 0.2$
Gamma(10, 10)	$\operatorname{Gamma}(\alpha,\beta)$	$135.0 \pm 0.2$	$107.0 \pm 0.2$	$92.3 \pm 0.2$	$88.3 \pm 0.2$	$88.3 \pm 0.1$
Uniform(0,1)	$\mathrm{Beta}(lpha,eta)$	$128.3 \pm 0.2$	$107.4 \pm 0.2$	$94.1 \pm 0.1$	$88.9 \pm 0.1$	$88.6 \pm 0.1$
Beta(10, 10)	$\mathrm{Beta}(lpha,eta)$	$131.1 \pm 0.4$	$106.7 \pm 0.1$	$92.1 \pm 0.2$	$87.8 \pm 0.1$	$87.7 \pm 0.1$
$\operatorname{Uniform}(-\pi,\pi)$	$vonMises(\mu, \kappa)$	$127.6 \pm 0.4$	$107.5 \pm 0.4$	$94.4 \pm 0.5$	$90.9 \pm 0.1$	$91.5 \pm 0.4$
vonMises(0, 10)	$\mathrm{vonMises}(\mu,\kappa)$	$130.7 \pm 0.8$	$107.5 \pm 0.5$	$92.3 \pm 0.2$	$87.8 \pm 0.2$	$87.9 \pm 0.3$
$\operatorname{Uniform}(S^D)$	von Mises Fisher ( $\pmb{\mu}, \kappa)$	$132.5 \pm 0.7$	$108.4 \pm 0.1$	$93.2 \pm 0.1$	$89.0 \pm 0.3$	$90.9 \pm 0.3$

## MCMC and variational inference

MCMC idea: Sample from the simple distribution and accept them, if the ratio is greater than some threshold:

$$\min\left(1, rac{p(\mathbf{w}^{ au}|\mathbf{y}, \mathbf{X}, \mathbf{h})}{p(\mathbf{w}^{ au-1}|\mathbf{y}, \mathbf{X}, \mathbf{h})}
ight),$$

where  $\mathbf{w}^{\tau}$  is set based on the previous sample:

$$\mathbf{w}^{ au} = T(\mathbf{w}^{ au-1}).$$

**Salimans et al., 2014:** let's interperete the sequence of some operator T application as a variational optimization:

$$\mathcal{T}^1 \circ \dots \mathcal{T}^\eta(\mathsf{w}) o p(\mathsf{w}^ au|\mathsf{y},\mathsf{X},\mathsf{h}).$$

Maclaurin et. al, 2015: use gradient descent as such operator. Do not reject samples at all.

# Optimization operator, Maclaurin et. al, 2015

#### Definition

Let T be an algorithm of changing model parameters  $\mathbf{w}'$  using previous parameter values  $\mathbf{w}$ :

$$\mathbf{w}' = T(\mathbf{w}).$$

#### Definition

Let L be a continuos loss function.

Define a gradient descent operator in the following way:

$$T(\mathbf{w}) = \mathbf{w} - \beta \nabla L(\mathbf{w}, \mathbf{y}, \mathfrak{D}).$$

## Gradient descent for evidence estimation

Consider posterior probability maximization:

$$L = -\log p(\mathfrak{D}, \mathbf{w}|\mathbf{h}) = -\sum_{\mathfrak{D} \in \mathfrak{D}} \log p(\mathfrak{D}|\mathbf{w}, \mathbf{h}) p(\mathbf{w}|\mathbf{h})$$

Optimize neural network in a multi-start regime with r initial parameter values  $\mathbf{w}_1, \dots, \mathbf{w}_r$  using (stochastic) gradient descent:

$$\mathbf{w}' = T(\mathbf{w}).$$

The parameter vectors  $\mathbf{w}_1, \dots, \mathbf{w}_r$  are from some latent distribution  $q(\mathbf{w})$ .

# **Entropy**

We can rewrite variational inference using differential entropy term:

$$egin{aligned} \log p(\mathfrak{D}|\mathbf{f}) &\geq \int_{\mathbf{w}} q(\mathbf{w}) \log rac{p(\mathfrak{D},\mathbf{w}|\mathbf{h})}{q(\mathbf{w})} d\mathbf{w} = \ & \mathsf{E}_{q(\mathbf{w})}[\log p(\mathfrak{D},\mathbf{w}|\mathbf{h})] + \mathsf{S}(q(\mathbf{w})), \end{aligned}$$

where  $S(q(\mathbf{w}))$  is a differential entropy:

$$S(q(\mathbf{w})) = -\int_{\mathbf{w}} q(\mathbf{w}) \log q(\mathbf{w}) d\mathbf{w}.$$

## Gradient descent for evidence estimation

#### Statement

Let L be a Lipschitz function, and optimization operator be a bijection. Then entropy difference for two steps is:

$$S(q'(\mathbf{w})) - S(q(\mathbf{w})) \simeq \frac{1}{r} \sum_{s=1}^{r} (-\beta Tr[\mathbf{H}(\mathbf{w}'^g)] - \beta^2 Tr[\mathbf{H}(\mathbf{w}'^g)]).$$

Final estimation for the  $\tau$  optimization step:

$$\log \hat{p}(\mathbf{Y}|\mathfrak{D},\mathbf{h}) \sim rac{1}{r} \sum_{g=1}^{r} L(\mathbf{w}_{ au}^{g},\mathfrak{D},\mathbf{Y}) + \mathsf{S}(q^{0}(\mathbf{w})) +$$

$$+\frac{1}{r}\sum_{b=1}^{\tau}\sum_{c=1}^{r}\left(-\beta \text{Tr}[\mathbf{H}(\mathbf{w}_{b}^{g})]-\beta^{2}\text{Tr}[\mathbf{H}(\mathbf{w}_{b}^{g})\mathbf{H}(\mathbf{w}_{b}^{g})]\right),$$

 $\mathbf{w}_b^g$  is a parameter vector for optimization g on the step b,  $\mathsf{S}(q^0(\mathbf{w}))$  is an initial entropy.

## How to calculate Hessian trace?

## **Problem**

$$\mathsf{Tr}[\mathbf{H}(\mathbf{w}_b^g)]$$

#### Statement

Let **U** be a symmetric matrix and **v** be the random vector with the following properties:

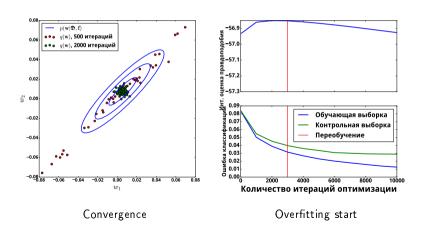
- ①  $Ev_i = 0$ ;
- ②  $Var(v_i) = 1$ .

Then

$$\mathsf{E}\mathbf{v}^\mathsf{T}\mathbf{U}\mathbf{v} = \mathit{Tr}[\mathbf{U}].$$

# Overfitting, Maclaurin et. al, 2015

Gradient descent does not optimize KL-divergence  $\mathrm{KL}(q(\mathbf{w})||p(\mathbf{w}|\mathfrak{D},\mathbf{h}))$ . Evidence estimation gets worse while optimization tends to the optimal parameter values. This can be considered as a overfitting start.



# Stochastic gradient Langevin dynamics

A modification of SGD:

$$T = \mathbf{w} - \beta \nabla L + \epsilon, \quad \epsilon \sim \mathcal{N}(0, \frac{\beta}{2})$$

where  $\beta$  changes with a number of iterations:

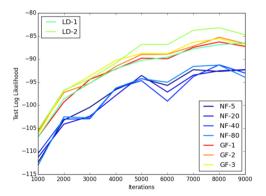
$$\sum_{\tau=1}^{\infty} \beta_{\tau} = \infty, \quad \sum_{\tau=1}^{\infty} \beta_{\tau}^{2} < \infty.$$

**Statement [Welling, 2011].** Distribution  $q^{\tau}(\mathbf{w})$  converges to posterior distribution  $p(\mathbf{w}|\mathbf{X},\mathbf{f})$ . Entropy adjustment:

$$\hat{\mathsf{S}}\big(q^{\tau}(\mathbf{w})\big) \geq \frac{1}{2}|\mathbf{w}|\mathsf{log}\big(\mathsf{exp}\big(\frac{2\mathsf{S}(q^{\tau}(\mathbf{w}))}{|\mathbf{w}|}\big) + \mathsf{exp}\big(\frac{2\mathsf{S}(\epsilon)}{|\mathbf{w}|}\big)\big).$$

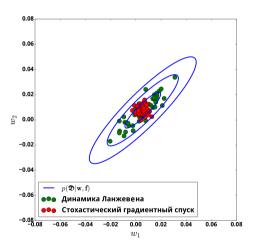
# Stochastic gradient Langevin dynamics for generative models

Altieri et al., 2015: sample latent variable z and use SGLD as a normalizing flow.



## SGLD vs SGD

Parameter distribution after 2000 iterations:



## Stein operator

Given a smooth probability function p and a smooth vector function  $\phi$ . Define a Stein operator as the following:

$$\mathcal{A}_p \phi(\mathbf{x}) = \nabla_{\mathbf{x}} \log p(\mathbf{x}) \phi^{\mathsf{T}} + \nabla_{\phi} \phi(\mathbf{x}).$$

Stein's identity:

$$\mathsf{E}_{\mathsf{x}\sim p}\mathcal{A}_p\phi(\mathsf{x})=0.$$

If we use q instead of p in the  $A_p$  we get a non-zero result, but close to zero as soon as p is close to q.

Let  $T(\mathbf{x}) = \mathbf{x} + \varepsilon \phi(\mathbf{x})$ . Then:

$$abla_{arepsilon} \mathsf{KL}(q||p)|_{arepsilon=0} = \mathsf{E}_{\mathsf{x}\sim q}\mathsf{trace}\mathcal{A}_p \phi.$$

Given a kernel K, the optimal  $\phi$  for minimizing KL is:

$$\phi^*(\mathbf{x}') = \mathsf{E}_{\mathbf{x} \sim q} \nabla_{\mathbf{x}} \log \ p(\mathbf{x}) \mathsf{K}(\mathbf{x}, \mathbf{x}') + \nabla_{\mathbf{x}} \mathsf{K}(\mathbf{x}, \mathbf{x}').$$

# Stein operator: algorithm

#### Algorithm 1 Bayesian Inference via Variational Gradient Descent

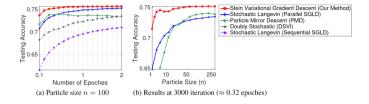
**Input:** A target distribution with density function p(x) and a set of initial particles  $\{x_i^0\}_{i=1}^n$ . **Output:** A set of particles  $\{x_i\}_{i=1}^n$  that approximates the target distribution p(x). **for** iteration  $\ell$  **do** 

$$x_i^{\ell+1} \leftarrow x_i^{\ell} + \epsilon_{\ell} \hat{\boldsymbol{\phi}}^*(x_i^{\ell}) \quad \text{where} \quad \hat{\boldsymbol{\phi}}^*(x) = \frac{1}{n} \sum_{j=1}^n \left[ k(x_j^{\ell}, x) \nabla_{x_j^{\ell}} \log p(x_j^{\ell}) + \nabla_{x_j^{\ell}} k(x_j^{\ell}, x) \right], \tag{8}$$

where  $\epsilon_{\ell}$  is the step size at the  $\ell$ -th iteration.

#### end for

## Stein operator: results



	Avg. Test RMSE		Avg. 7	Avg. Time (Secs)		
Dataset	PBP	Our Method	PBP	Our Method	PBP	Ours
Boston	$2.977 \pm 0.093$	$2.957 \pm 0.099$	$-2.579 \pm 0.052$	$-2.504 \pm 0.029$	18	16
Concrete	$5.506 \pm 0.103$	$5.324 \pm 0.104$	$-3.137 \pm 0.021$	$-3.082 \pm 0.018$	33	24
Energy	$1.734 \pm 0.051$	$1.374 \pm 0.045$	$-1.981 \pm 0.028$	$-1.767 \pm 0.024$	25	21
Kin8nm	$0.098 \pm 0.001$	$0.090 \pm 0.001$	$0.901 \pm 0.010$	$0.984 \pm 0.008$	118	41
Naval	$0.006 \pm 0.000$	$0.004 \pm 0.000$	$3.735 \pm 0.004$	$4.089 \pm 0.012$	173	49
Combined	$4.052 \pm 0.031$	$4.033 \pm 0.033$	$-2.819 \pm 0.008$	$-2.815 \pm 0.008$	136	51
Protein	$4.623 \pm 0.009$	$4.606 \pm 0.013$	$-2.950 \pm 0.002$	$-2.947 \pm 0.003$	682	68
Wine	$0.614 \pm 0.008$	$0.609 \pm 0.010$	$-0.931 \pm 0.014$	$-0.925 \pm 0.014$	26	22
Yacht	$0.778 \pm 0.042$	$0.864 \pm 0.052$	$-1.211 \pm 0.044$	$-1.225 \pm 0.042$	25	25
Year	$8.733 \pm NA$	$8.684 \pm NA$	$-3.586 \pm NA$	$-3.580 \pm NA$	7777	684

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