# Point of view

# A farewell to P-values?

A recent paper in the medical literature by Sterne and Davey-Smith<sup>1</sup> has focused attention on some of the problems of interpretation of significance/hypothesis tests; in particular, the meaning and status of P-values associated with these tests. What were these concerns: that the division of results into "significant" and "nonsignificant" according to a P-value = 0.05 was arbitrary and not in accordance with the prescriptions of the founders of statistical inference; that the P-value is misinterpreted as the probability that the null hypothesis is true; that, as the absolute value of the P-value indexes the level of evidence against the null hypothesis, measures of effect should attract a P-value of 0.001, in preference to 0.05, where the evidence against the null hypothesis is not strong; Bayesian approaches to reporting of results may have advantage; and "significance" should not be a primary claim of the reporting of results, which should be accompanied by (90%) confidence intervals and interpreted in the context of the type of study and other available evidence.

As the authors observe, these matters are not new and have repeatedly surfaced in the literature of various scientific disciplines since the establishment of the "testing" paradigm in the 1920s and 1930s by RA Fisher and J Neyman & E Pearson.<sup>2</sup> Two volumes, separated by almost 30 years and authored from within the behavioural science disciplines: "The significance test controversy - a reader" (published in 1969) and "What if there were no significance tests" (published in 1997), further attest to these controversies. In a provocatively entitled paper, "Two cheers for P-values", Stephen Senn, in a "limited defence of P-values", noted that "P-values are a practical success, but a critical failure. Scientists the world over use them, but scarcely a statistician can be found to defend them. Bayesians in particular find them ridiculous...".5 In 1996, Nester suggested that "statisticians would be unwise to seek the limelight in any forthcoming 75<sup>th</sup> anniversary, centennial or tricentennial celebrations of hypothesis testing".<sup>6</sup> Rindskopf asked why "Given the many attacks on it. null hypothesis testing...(was not) ...dead", and the demise of the P-value has been rhetorically reported.<sup>8</sup>

Within the medical literature similar sentiments have been expressed, as reflected in the titles of certain lead articles: "Confidence intervals rather than P values: estimation rather than hypothesis testing", "That confounded P-value" and "Are all significant P values created equal?". "I That this was not merely an academic

question, was revealed by a decision of the editor of the journal Epidemiology: 12 "When writing for Epidemilogy, you can also enhance your prospects if you omit tests of statistical significance....we do not publish them at all". 13 In psychology and the social sciences, the tone of discourse (against P-values) has at times been shrill, as noted by Nickerson in a recent exhaustive review. 14 In the biological 15 and econometric literature, 16 others have added to the chorus of complaint. How did this come about or is it all "a tempest in a tea pot"? 14

### History: the paradigm established

From our current medical perspective, "testing", P-values and Type I and II errors appear non-problematic; a "single, unified, uncontroversial means of statistical inference", <sup>17</sup> and the history of the development of statistical inference in the 20<sup>th</sup> century, a remote echo of current concerns. <sup>18</sup> The state of statistics in the first decade of the 20<sup>th</sup> century has been described as "an unexplored archeological site" <sup>19</sup> and the construction of the "testing" paradigm, first by Fisher and then Neyman-Pearson, was self-consciously defined with respect to the 19<sup>th</sup> century Bayesian dominance of "inverse probability". It was paradoxical that Gosset, the inventor of the *t*-test (1909), which initiated hypothesis testing, <sup>20</sup> was a Bayesian. <sup>21</sup>

It is obviously difficult to relive the impact of the Fisherian revolution upon the statistical practice of the first decades of the  $20^{th}$  century, but we may be assured that it was fundamental, although the full significance of the first edition (1925) of the classic "Statistical methods for Research Workers" was , perhaps not surprisingly, "not immediately recognized".  $^{22}$  Fisher and Gosset, in fact, cooperated in calculating tables for the t-distribution presented (along with  $\chi^2$  and z-transformation) in the book.  $^{16}$  The initial "common currency" of significant at 5% and 1% may well have been related to the fact that Fisher's tables (copied subsequently to many text books) were given for P-values of 0.01 and 0.05, partly in deference to the copyright limitations of the journal Biometrika, edited by Karl Pearson, the founder of the  $\chi^2$  test.  $^{23}$ 

R.A. Fisher's position as the "founder" of modern statistics is presumably secure;<sup>24</sup> the comments of Savage (a Bayesian) attest to this: "It would be more economical to list the few statistical topics in which he displayed no interest than those in which he did".<sup>25</sup> The Fisherian *significance* test, deriving from inductive inference, established a null hypothesis (H<sub>0</sub>) and used discrepancies in the data to reject the null hypothesis: that is, H<sub>0</sub> posited a sample coming from a hypothetical (infinite) population with a known sampling distribution and H<sub>0</sub> was rejected if the sample estimate deviated from the mean of the sampling distribution by more than a specified criterion (the level of significance;

formally,  $Pr(x|H_0)$ .<sup>17</sup> The P-value then, was the (tail area) probability of obtaining a result equal to or more extreme than what was actually observed;<sup>21</sup> for, say a P value of 0.05, it was not that the null hypothesis had a probability of (only) 5%. Under an assumption that the null hypothesis was true, it could not then be assumed that the P value was a "direct measure of the probability that the null hypothesis is false". <sup>26</sup> A P value of  $\leq 0.05$ on the null hypothesis indicated, according to Fisher, that: "Either an exceptionally rare chance has occurred or the theory is not true". It did not imply that the investigator accepted being deceived one in twenty occasions, rather it suggested what should be ignored: "all experiments in which significant results are not obtained". Fisher's further advice, oft quoted and ignored, was that "If one in twenty does not seem high enough odds, we may, if we prefer it, draw the line at one in fifty...or one in a hundred". 17 A subtle, but important point, was that the inference from the P value involved only one hypothesis and was partially based on unobserved data in the tail region (of the sampling distribution); thus the "likelihood" of a hypothesis, deriving from the data, was not, at the same time, the "probability of being true". 18,27 That is, P-values were not to be misinterpreted as posterior probabilities, a Bayesian proposition.<sup>28</sup>

The statistical methodology of Neyman & Pearson, articulated primarily in two papers in 1928 and 1933. sought to revise and improve upon Fisher's formulation, but from a deductive position, a paradigmatic difference, which from the Fisherian viewpoint of "rigorous inferences from the particular to the general", was a function of a certain mathematical "bias" of Neyman. Although somewhat distrustful of mathematicians ("Statistical methods for Research Workers" contained no formal mathematical proofs or lemmas), Fisher was in fact a Cambridge trained mathematician.<sup>25</sup> The Fisher-Neyman rivalry (Pearson later "distanced" himself from the Neyman-Pearson paradigm, more so when Neyman relocated to Berkley) was somewhat of a cause-celebre, 30 although the influence of Fisher's "Statistical methods for Research Workers" on the Neyman-Pearson enterprise was acknowledged by the latter, 20 specifically the tabulation of the three distributions mentioned above. Neyman, for his part, charged Fisher with a persistent inability to operate with concepts;<sup>31</sup> Fisher's circumlocutions were also a cause of irritation to those sympathetic to his view-point, such as Kempthorne: "The last sentence, particularly, leads me to the view that Fisher was talking on a plane barely understandable to the rest of humanity". 32

It was Neyman-Pearson methodology which formulated the now familiar two competing hypotheses paradigm, the null  $(H_0)$  and the alternate  $(H_A)$ . This involved the probability of committing two kinds of errors with

respect to the null hypothesis; false rejection (Type I or  $\alpha$  error) and false acceptance (Type II or  $\beta$  error). Power, defined as  $(1-\beta)$  or the P(rejecting  $H_0 \mid a$ particular alternative hypothesis), was introduced as a new and critical concept. Within the Fisherian schema, there was a notion of "sensitivity" in detecting departures from the null, but no formal concept of power;<sup>33</sup> although Barnard has argued otherwise. The  $\alpha$  error was, in theory, prescribed prior to data collection and the focus was on minimizing  $\beta$  errors, subject to the bound upon  $\alpha$ . What in effect was established were rules for making decisions between two hypotheses ("inductive behavior", although this behavioral aspect may only have been heuristic or hypothetical), 35 on the basis that in the "long run of experience, we shall not be too often wrong". 17 Thus for Neyman-Pearson, there was a tension between the control of long term error rates and judgment of the status of the individual experiment. The  $\alpha$  and  $\beta$  error rates defined a rejection region for a test statistic; the significance level,  $\alpha$ , was the "probability of a set of future outcomes", represented by the "tail" area of the null distribution. In principle, Neyman-Pearson theory also avoided an arbitrary element in Fisher's approach, the decision regarding the test statistic.5 The Neyman-Pearson fundamental lemma guaranteed the existence of an optimal (uniformly) Most Powerful- $\alpha$  test;<sup>36</sup> for a simple H<sub>0</sub> tested against a simple alternative H<sub>A</sub>, the optimal test criterion was the likelihood ratio (LR) test. 23,37 The question of Fisher's approach to the alternative hypothesis is one of some difficulty: arguments that a small probability  $p(E|h_0)$  of event E is "not enough per se to discredit the null hypothesis  $h_0$ " have been "forcefully" advanced.33 In particular, Berkson's oft cited paper from 1941 in which the question was posed: "If an event has occurred, the definitive question is not, "Is this an event which would be rare if  $H_0$  is true?" but "Is there an alternative hypothesis under which the event would be relatively frequent?" If there is no plausible alternative at all, the rarity is quite irrelevant to a decision..." <sup>38</sup> Undoubtedly Fisher considered the 'alternative' as obligatory, as revealed in conversation with Kruskal and Savage in the 1950's, where the former recalls that ".. Fisher agreed that, yes, naturally one had to think about distributions for the sample other than that of the hypothesis under test and why were we making such a fuss about an elementary and trivial question".39

That significance testing could provide statistical inference, rather than behavioural decisions, was the gulf that separated Fisher from (early) Neyman-Pearson. Some modification of the Neyman-Pearson position on this did occur, both internally (Pearson's description of a statistical test as a "means of learning" and Neyman's subsequent equivocations on the matter of

inference<sup>41</sup>) and externally, such as the position of Lehmann, in the classic 1959 volume "Testing statistical hypotheses", that in a hypothesis test the "information will be used for guidance...In such cases the emphasis is on the inference...". <sup>42</sup>

A data-based P-value is a random variable with a distribution, under the null hypothesis and for continuous test statistics, uniform over the interval [0,1], regardless of the size of the study.  $^{43,44}$  Under the alternative hypothesis, the distribution of the P-value is skewed and is a function of both sample size ("a natural concept of power") $^{46}$  and the distribution of the test statistic that is used. P-values are therefore not  $\alpha$ -error rates,  $^{5,17,20,26}$  although both are tail-area probabilities under the null hypothesis. As Berger and Delampady note: "P values are not a repetitive error rate, at least in any real sense. A Neyman-Pearson error probability,  $\alpha$ , has the actual frequentist interpretation that a long series of  $\alpha$  level tests will reject no more than  $100\alpha\%$  of true  $H_0$ , but the data dependent P values have no such interpretation".  $^{46}$ 

For a fixed, pre-specified  $\alpha$ , the Neyman-Pearson decision rule "could be defined equivalently in terms of the P-value" but what would be of interest was the fact that  $P < \alpha$ , not the specific value of P. It is ironic that standard applied statistical practice lies easily with an amalgam of the two methods, although from an alternate perspective, this amalgam may be considered as one of statistics "greater triumphs". Tests of significance, as reported in journals, would appear to follow Neyman-Pearson "formally" but Fisher "philosophically" and practically. 33

In the debate about the utility of P-values versus confidence intervals (CI), it is often forgotten that CI were introduced by Neyman in 1937, <sup>48</sup> and were considered by Neyman, and commentators, as integral to the overall theory of hypothesis testing, <sup>49</sup> which embodied the frequentist theory of repeated sampling (an anathema to Fisher<sup>50</sup>). Thus for a 95% CI of a parameter  $\theta$ , the interpretation is that in an infinite number of repetitions of a study, an exact proportion (95%) of all such intervals would enclose  $\theta$ . Once the data has been collected and a single 95% CI has been calculated, the probability that  $\theta$  lies within this CI is now 0 or 1. That is, a 95% CI is not equivalent to a 95% probability interval (which has a Bayesian interpretation). <sup>51-54</sup>

Besides the theory of testing, the second great divide between Neyman and Fisher was that of confidence versus fiducial intervals, the latter being based on fiducial probability,<sup>55</sup> introduced by Fisher as an alternative to Bayesian posterior probabilities.<sup>55</sup> Fiducial probability had, as its basis, certain sufficient statistics (*F* and *t* statistics and correlations) which contained all the information in a sample relevant to population parameters (inference from sample to population).

Although agreement can be demonstrated between CI and fiducial intervals; the classic paper establishing so called exact (or Clopper-Pearson) CI of the binomial was entitled "The use of confidence or fiducial limits illustrated in the case of the binomial", <sup>56</sup> the latter has not stood the test of Neyman's withering attacks, <sup>31,57,58</sup> nor time (Fisher's "biggest blunder"). It is of interest, however, to record that Fisher (as early as 1935) apparently "recognized that 'confidence intervals' are 'only another way of saying that, by a certain test of significance, some kinds of hypothetical possibilities are to be rejected, while others are not' ".<sup>59</sup>

## History: the paradigm revised

The literature in response to the Fisher and Neyman-Pearson paradigm is, not surprisingly, enormous in breadth and detail, both from within the statistical, 35,60-64 and applied scientific disciplines. One of the more systematic and useful developments is that associated with DR Cox, in which P-values are treated as "rough tools for inspecting data". 65 The perspective, developed fully in the 1974 volume of Cox and Hinkley, 66 is one of eclecticism, the central theme being that "it is fruitful to contemplate problems formulated in different depths of detail and to use different methods accordingly...the most primitive formulation is that for a pure significance test, where only the null hypothesis under test need be explicitly formulated, and the richest formulation is that for Bayesian decision analysis...". Null hypotheses are not viewed as undifferentiated species, rather, are divided into plausible (close to the truth) and dividing (divide the range of possibilities into qualitatively different types) hypotheses, which may also be further sub-divided and specified. Statistical strategies, the use of significance tests and the actual Pvalue level, are determined contingent upon these hypothesis specifications.<sup>68,69</sup>

A significance test (measuring the consistency of the data with a null hypothesis) has the following form: a function t = t(v) of the observations exists, such that, the larger t(y), the greater the inconsistency of y (the observed vector of responses) with  $H_0$ . T = t(y) is the test statistic (a random variable). If the distribution of T is known (when Ho is true), then the level of significance  $p_{\text{obs}} = \text{pr} \ (T \ge t_{\text{obs}}: H_0)$ . The result of such a test is a significance level (not a decision);  $p_{obs}$  is a "guide, and no more, to interpretation" (albeit the mathematical connection between  $p_{\rm obs}$  and "critical regions of pre-assigned size", that is, Neyman-Pearson testing). 66,68,70 A similar approach was also recommended by Kempthorne. 71 Of interest, Cox, in a somewhat sympathetic response to the paper by Sterne and Davey-Smith, suggested that "To distinguish several types of hypotheses that might be tested helps to understand the issues". 72

#### The problem revisited

Where then do we stand? In a response to the debate over the paper by Sterne and Davey-Smith, 1 Berger, 73 outlined potential problems with hypothesis testing and it is useful to consider some of these:

- 1. *P-values are misunderstood*. The frequent misrepresentation of P-values and hypothesis testing, especially in textbooks, has been repeatedly documented. Such, as with other statistical misrepresentations, is not an argument for their abolition.
- 2. P-values as a measure of support. Sterne and Davey-Smith, suggest a graded level of evidence against the null hypothesis, indexed by the P-value; such scales date back to the 1970s. 70 A corollary to this is the so called α-postulate of Cornfield (rejected by him in favour of likelihood ratios), that "All hypotheses rejected at the same critical level have equal amounts of evidence against them".76 However, the question of sample size for equal Pvalues needs to be considered; a number of commentators have argued that for, say a P-value of 0.05, there is stronger evidence against  $H_0$  for a small sample than a large one. <sup>27,38,70,77-79</sup> Schervish also demonstrated that "the interpretation of a particular value on the scale of support, such as the popular .05, must vary with the hypothesis" and was "unable to construct a consistent interpretation of the P-value as anything similar to a measure of support for its hypothesis".
- 3. *P-values are associated with rigid cut-off values*. A flexible (eclectic) attitude to tests and P-values, associated with the approach of Cox and Kempthorne, has been outlined above. As opposed to Berger, it would indeed appear reasonable that there "should be no sharp distinction made between cases having a P-value of say 4.9% and those having a P-value of 5.1%- a distinction forced by the language of confidence interval testing". <sup>80</sup> In the presence of a bewildering array of possible statistics, the advice of Kempthorne to "Look at it" seems apposite; similar to the admonitions of the adherents of the likelihood principle <sup>81</sup> and the Bayesians.
- 4. P-values are the wrong measure of evidence. From the Bayesian perspective, P-values overstate the evidence against the null hypothesis and other methods to adduce evidence (likelihood ratios) may be of more utility. 82

In a frequently cited paper by J.O. Berger and Sellke, it was shown (two sided testing a normal mean) that with a P-value of 0.05, the posterior probability of the null was at least 0.30 for any objective prior distribution.  $^{83}$  However, in the one-sided setting, where the different geometry of  $H_0$  and  $H_A$  was not operative, the discrepancy, Bayesian

posterior probability versus P-value, was no longer evident. 84 Technically, this 'geometry' relates to the fixing of a (prior) probability mass on the null and varying it on the alternative; Casella and R. L. Berger suggest that the discrepancy between Pvalues and  $P(H_0|x)$  is a function of the large (50%) prior probability mass placed on  $H_0$  by J.O. Berger and co-workers,  $^{46,83}$  and conclude that "there is agreement between P-values and Bayesian interval null calculations in the more typical situation in which small prior probability is assigned to  $H_0$ ". As Casella and R. L. Berger note, "We would be surprised if most researchers would place even a 10% prior probabilty on  $H_0$  ", in accord with the sentiments of Meehl, who maintained that the point null hypothesis is "...[quasi-] always false in biological and social science". 86 P-values and posterior probabilities are not necessarily in competition and any difference in conclusions reached does not "....by itself invalidate either measure". 74 Of interest, J.O. Berger and co-workers recently proposed calibrating P-values such that they may be interpreted in either a Bayesian fashion (B(p))=  $-e.p \log(p)$ , when p < 1/e) or a frequentist way  $(\alpha(p) = (1+[-e.p \log(p)]^{-1})^{-1}$ .<sup>87</sup>

Goodman, again from a Bayesian perspective, calculated the replication probability (using an uninformative prior) of trials at a P-value of 0.05; this was found to be 50% and lower than "expected" (by non-statisticians).88 Senn has subjected the import of this finding to close scrutiny. 5,21 Firstly, from a Bayesian perspective, P-values are not unreasonable given an uninformative prior. However, the problem is that "...the 'uninformative' prior is rarely appropriate...it is not possible to survive as a Bayesian on uninformative priors...".21 Secondly, the requirement that a single significant Pvalue should entail near certainty that a second will follow, is deemed by Senn to be an undesirable property: "Anticipated evidence is not evidence, nor do we want it to be. To expect that it is, is to make exactly the same mistake that physicians make in saying, 'the result was not significant, p = 0.09, because the trial was too small". This being said of the general problem of replicability, empirical studies have suggested that the P-value does provide "a continuous measure that has an orderly and monotonic mapping onto confidence in the replicabiltiy of a null hypothesis rejection"89 and statistically significant exact replication (SSER) may be a useful interpretative measure. 90

#### P-values and CI

There would appear to be considerable virtue in reporting both P-values and CI, on the basis that

singular statements such as P < 0.05, or P = nil sig, convey little useful information, although for a 100(1α)% CI, it must be remembered that any violation of the assumptions that effect the true vale of  $\alpha$  (obviously) effect CI precision.<sup>91</sup> From the Bayesian perspective, Lindley has summarized the position thus: "significance tests, as inference procedures, are better replaced by estimation methods...it is better to quote a confidence or credible interval....estimation procedures provide more information....Nevertheless there remain cases where significance test have an advantage...". Numerous papers within the medical literature have attested to the utility of CI, 93-95 in the epidemiological literature, the P-value has been condemned as confounded, in that the information "mixed" in the Pvalue should be separately reported: the size of the effect (estimated by, say, the risk ratio) and the precision of the estimate (described by the SE or CI).<sup>10</sup> Poole has suggested that for epidemiological measures such as relative risk, the estimates least influenced by chance are those with narrow confidence intervals, not low P-values. 96 However, as pointed out by Feinstein, Pvalues and CI methods are essentially reciprocal and do not provide "an evaluation of substantive importance for the 'big' or 'small' magnitude of the observed distinction...they offer no guidance for the basic quantitative scientific appraisals that depend on purely descriptive rather than inferential boundaries.". Similarly, Poole suggested that CI "are usually taken as nothing more than tests of significance" and proposed that the complete P-value (that is, the graph of all possible P-values or CI) or likelihood function be used for the "main result of an epidemiological study". 98

An interesting case study of the interpretation of "CI without P-values" and a focal point for a lively exchange in the American Journal of Public Health in the mid 1980's on the role of P-values, was the response of Fleiss to a relatively small case-control study<sup>99</sup> in which all 12 reported CI for summary odds ratios included 1 (extending from 0.4 to 17); yet the associations were variously described as 'strong' 'negative' and 'positive'. Fleiss asked the not unreasonable question "There is no gainsaying that tests of significance have been abused, but at least they have the virtue of providing explicit, pre-specifiable criteria for inferring that an association is real. This is not the case with confidence intervals, at least as far as the paper in question is concerned....I would appreciate learning just what criteria were employed to conclude that the ....associations were 'strong', 'negative' and 'positive'". 100 The authors acknowledged the small study size, and suggested that the reader should be "cautious in generalizing our results", but even when not "statistically significant" point estimates may "significantly add to our understanding". 101 The latter

position would appear to distance itself considerably from the Fisherian requirement of "rigorous uncertainty". <sup>30</sup>

The question of the "propriety' of associations and/or claims of efficacy also resonates with the reporting of drug trials; 102 the case for confidence intervals as estimates of effect has been well argued 103 and indeed, would appear to be non-controversial. In the presence of competing claims and professional enthusiasts, one can, with V.W. Berger, question the likelihood of being mislead 13 and find P-values eminently applicable to control this probability, 104 not withstanding the utility of other (for example, Bayesian) approaches.

#### Overview

Efron summarized the major reasons why, as opposed to the Bayesian 19<sup>th</sup> century, Fisherian and Neyman-Pearson ideas have held sway in the 20<sup>th</sup>: ease of use, model building, division of labour (parts of a complicated problem may be addressed separately) and objectivity. Thus, despite the belief that P-values are dead and buried (by some journals), we would agree with Fleiss that significance tests are "alive and well". 107

#### Post-script

As this is published in an Australian journal, we provide the following further information: Ronald Aylmer Fisher was born February 17 1890 in East Finchley, London and died July 29, 1962 in Adelaide, South Australia. His ashes are interred in St Peter's Cathedral in the latter city.

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