

# Take home 2

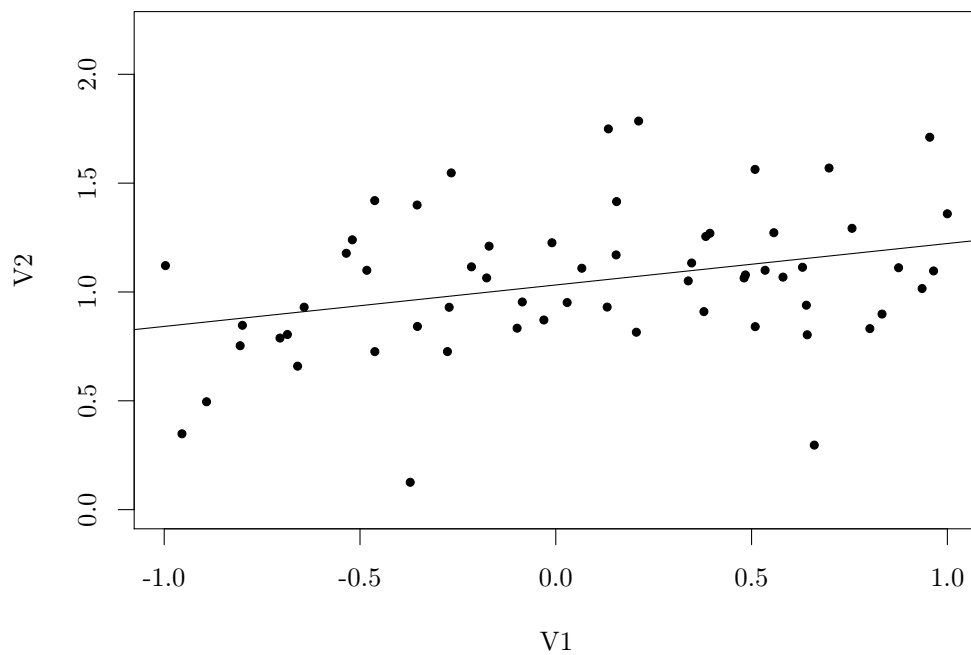
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## Question 1

(a)

The value for the intercept of the fit is 1.0321764 and for the slope of the fit is 0.1904323



Figuur 1: line fit through the data in Ex1.txt

(b)

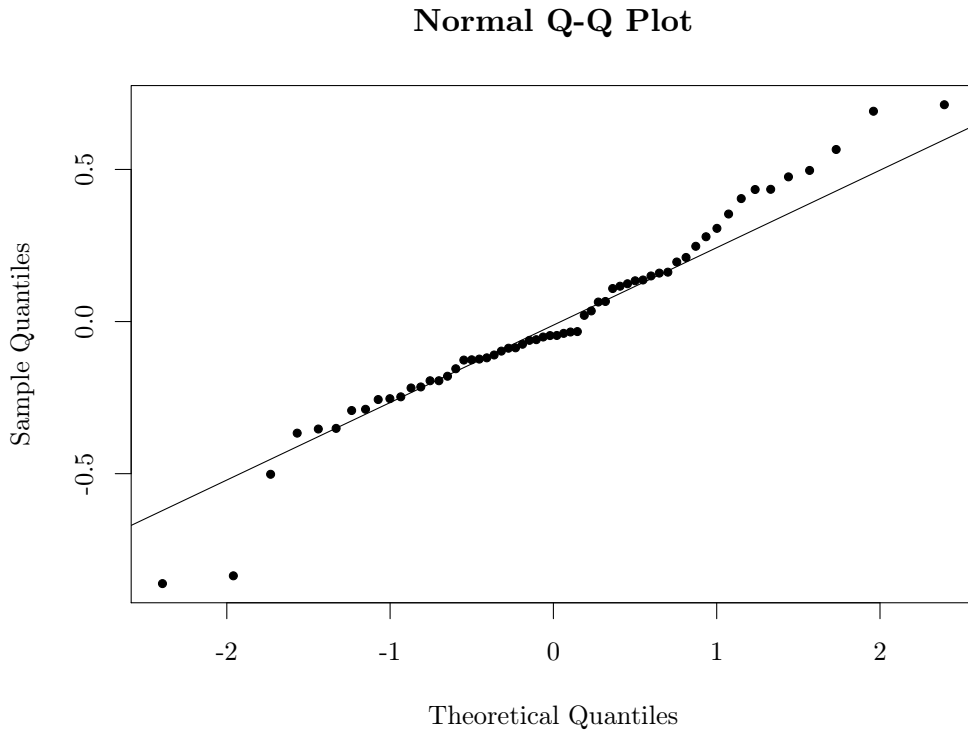
Using the result from section 7.4.1 in <REFERENCE cursus> we know that the random variable

$$T = \frac{\beta_1}{\sqrt{\frac{S^2}{\sum_{i=1}^n (x_i - \bar{x})^2}}} \quad (1)$$

has a Student distribution with  $n - 2$  degrees of freedom. The test value is 2.603. Using a student-t distribution with  $60 - 2 = 58$  degrees of freedom we find a p-value of 0.0117. At the confidence level  $\alpha = 0.01$ , the null hypothesis that  $\beta_1 = 0$  holds. The 99% confidence region for the test value is  $[-2.663, 2.663]$ .

(c)

A q-q plot is a plot where the quantiles of the assumed distribution are plotted against the quantiles from the sample. So in a sample of  $n$  points where we label the observation  $x_i, i \in \{1, 2, \dots, n\}$  from lowest to highest, the  $i$ th point will be plotted at the  $(Q(i/n), x_i)$ . Here  $Q(x)$  is the quantile function of the assumed distribution. This is a function such that  $P(X < Q(p)) = p$ . If the assumed distribution is a good model for the observed sample, the points will well fitted by a linear function.



Figuur 2: qq plot for the errors on the fit

(e)

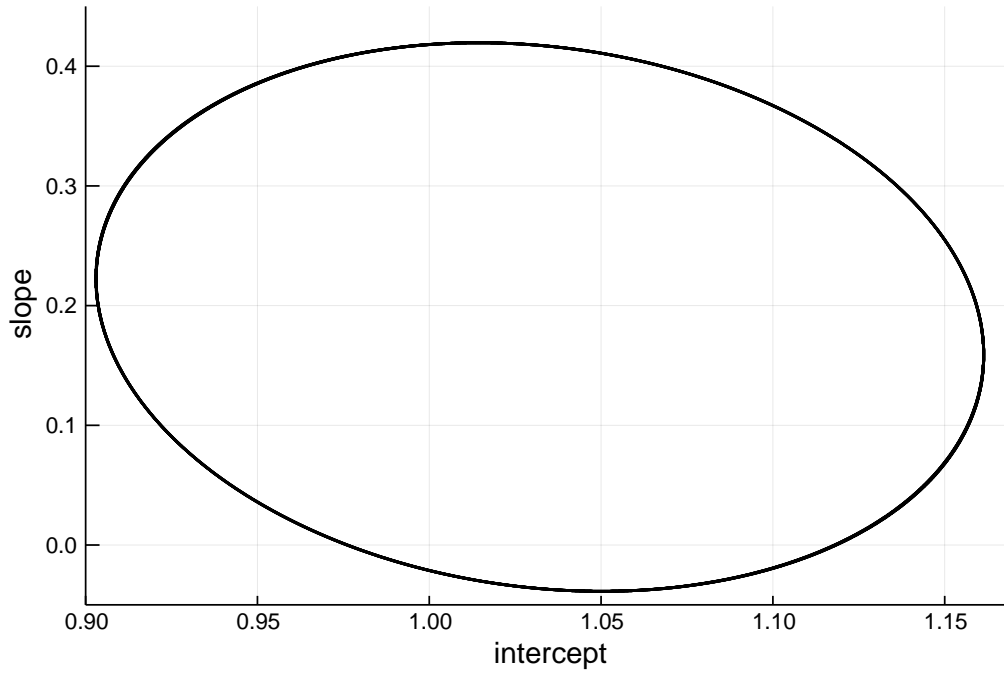
We use the formula from section 7.5.4 in the book to find a confidence region for  $\beta$ . This region is of the form:

$$\left\{ \beta \left| \frac{(\hat{\beta} - \beta)^T (X^T X) (\hat{\beta} - \beta)}{p S^2} \leq F_{p, n-p}(1 - \alpha) \right. \right\} \quad (2)$$

Because the matrix  $X^T X$  is positive definite, this is an ellips with center  $\beta$  and axes along the eigenvectors of  $X^T X$ . The upper bound for this expression is in this case:  $F_{2, 58}(0.99) = 4.9910$ . Filling in the values for  $X, p = 2, n = 60$  and  $S^2$  we get the following result (obtained using the cAS system sympy):

$$\{\beta | 5.2596\alpha^2 + 0.8222\alpha\beta - 0.4950\alpha + 1.6711\beta^2 - 0.6629\beta + 0.0711 \leq 4.9910\}$$

The resulting ellips is plotted in figure 3.



Figuur 3: A plot of the 99% confidence region for  $\beta$ .

## Question 2

(a)

The value for  $\hat{\beta}_{OLS}$  obtained using ordinary least squares with the correlated errors is:

$$\hat{\beta}_{OLS} = (3.923161, 1.214355)^T \quad (3)$$

(b)

we know that the variance is the same for all  $\epsilon_i$  and that they obey the recursive relation:

$$\epsilon_i = \rho\epsilon_{i-1} + \eta_i$$

with  $\eta_i$  standard normaly distributed. Using  $Var(\epsilon_i) = Var(\epsilon_j)$  for all  $i, j \in \{1, 2, \dots, 80\}$  we find:

$$\begin{aligned} Var(\epsilon_i) &= \rho^2 Var(\epsilon_{i-1}) + Var(\eta_i) \\ \Leftrightarrow Var(\epsilon_i) &= \rho^2 Var(\epsilon_i) + 1 \\ \Leftrightarrow Var(\epsilon_i) &= \frac{1}{1 - \rho^2} \\ &= \frac{25}{9} \\ &= 2.777\dots \end{aligned}$$

We conclude that the variance of  $\epsilon$  is  $\frac{25}{9}$ .

(c)

To transform the errors to a basis where they are no longer correlated we will use the variance-covariance matrix  $V$  of the errors. The off-diagonal elements, so the covariances, are (suppose  $i < j$ :

$$\begin{aligned}
Cov(\epsilon_i, \epsilon_j) &= E[(\epsilon_i - E(\epsilon_i))(\epsilon_j - E(\epsilon_j))] \\
&= E[\epsilon_i \epsilon_j] \\
&= E[\epsilon_i (\rho \epsilon_{j-1} \eta_j)] \\
&\vdots \\
&= E[\epsilon_i (\rho^{j-i} \epsilon_i + \rho^{j-i-1} \eta_i + \rho^{j-i-2} \eta_{i-2} + \dots + \eta_j)] \\
&= E[\epsilon_i \rho^{i-j} \epsilon_i] \\
&= \rho^{i-j} Var(\epsilon) \\
&= \rho^{i-j} \frac{25}{9}
\end{aligned}$$

So the matrix  $V$  has  $\frac{25}{9}$  on the diagonal,  $\frac{25}{9}\rho$  on the two sub-diagonals next to the diagonal,  $\frac{25}{9}\rho^k$  on the  $k$ th sub-diagonal. the matrix  $V$  can be diagonalized with an orthogonal (or unitary) transformation  $U$ :  $\Lambda = MV M^T$ . When we use the definition of the variance-covariance matrix (where  $\epsilon$  denotes the column vector with the errors) we find:

$$\begin{aligned}
V &= E((\epsilon - E(\epsilon))(\epsilon - E(\epsilon))^T) \\
&\Leftrightarrow \Lambda = UE(\epsilon\epsilon^T)U^T \\
&\Leftrightarrow \Lambda = E((U\epsilon)(U\epsilon)^T)
\end{aligned}$$

Where we used that  $E(\epsilon = \mathbf{0})$ . The vector  $U\epsilon$  has variance-covariance matrix  $\Lambda$  so we conclude that it is a vector of uncorrelated stochastic variables.

The value for  $\hat{\beta}_{OLS}$  that we find using these uncorrelated errors is exactly the same as with the uncorrelated errors:

$$\hat{\beta}_{OLS} = (3.923161, 1.214355)^T$$

This is expected because we used a unitary transformation to get uncorrelated errors and by filling this in the expression for the Sum of squares of the remainders we find:

$$\begin{aligned}
RSS(\hat{\beta}_{Uncorrelated}) &= (UY - UX\beta)^T (UY - UX\beta) \\
&= (Y - X\beta)^T U^T U (Y - X\beta) \\
&= (Y - X\beta)^T I (Y - X\beta) \\
&= RSS(\hat{\beta}_{Correlated})
\end{aligned}$$

which means the RSS for the correlated and the uncorrelated errors will have the same optimal  $\hat{\beta}$ .

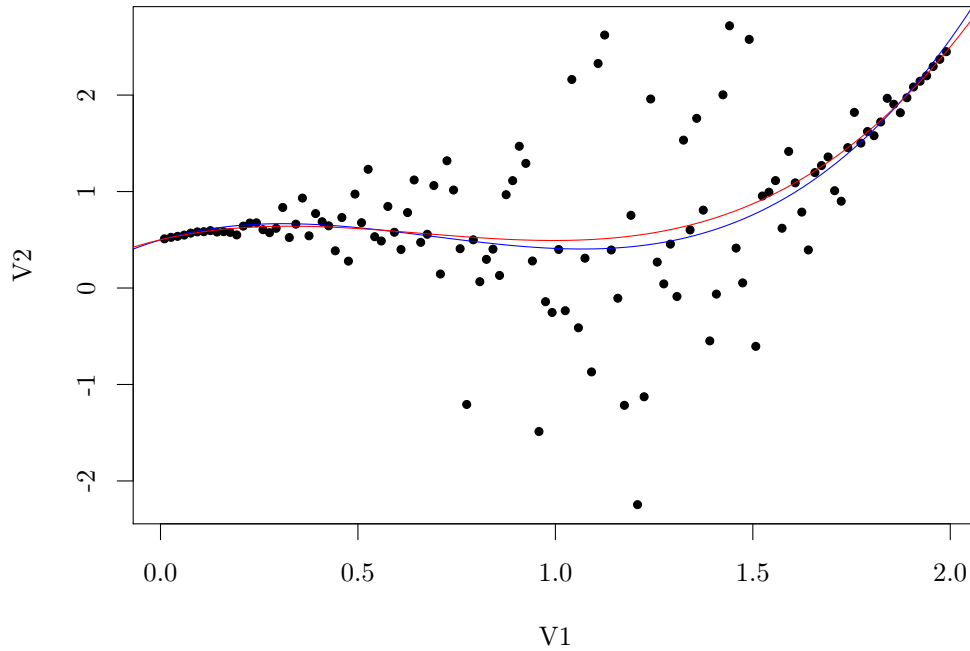


Figure 4: fits of a cubic function using ordinary least square (blue) and weighted least squares (red)

### Question 3

(a)

The values for the fitted coefficients with ordinary least squares are:  $[0.4994, 1.1985, -2.4959, 1.2082]$ . In figure 4 the fit is plotted along with the data.

(b)

The coefficients fitted with the weighted least square algorithm are  $[0.5003, 0.9669, -1.9655, 0.9910]$ .

(c)

The coefficients that are calculated using the least squares method are closer to the real values than the ones calculated using the ordinary least squares. To really say something about the difference and how well they match the given real values of the function we need more information about the distribution of the parameters. In figure 4 the two fits are plotted together with the data to fit through.

(d)

To derive the distribution of  $\hat{\beta}_{WLS}$  we start from the distribution for the errors, which is known:

$$\epsilon \sim N_n(0, W)$$

Where  $W$  is a diagonal matrix with at position  $W_{ii}$  the variance of  $\epsilon_i$ . This is the same matrix  $W$  used in the formula to calculate the estimate for  $\beta$  with the weighted least squares method. This formula is:

$$\hat{\beta}_{WLS} = (X^T W^{-1} X)^{-1} X^T W^{-1} Y \quad (4)$$

here  $Y$  is the vector with samples used to calculate the least squares estimator.  $X$  is the matrix as defined on page 193 in the course notes, this matrix is deterministic.  $Y$  can be rewritten as follows:

$$Y = X\beta + \epsilon$$

substituting this into equation 4 yields:

$$\begin{aligned} \hat{\beta}_{WLS} &= (X^T W^{-1} X)^{-1} X^T W^{-1} (X\beta + \epsilon) \\ &= (X^T W^{-1} X)^{-1} (X^T W^{-1} X)\beta + (X^T W^{-1} X)^{-1} X^T W^{-1} \epsilon \\ &= \beta + X^{-1} W X^{-T} X^T W^{-1} \epsilon \\ &= \beta + X^{-1} \epsilon \end{aligned}$$

We see this is a linear transformation of a multivariate normal distribution with mean  $\mathbf{0}$  and as variance-covariance matrix  $\Sigma_\epsilon = W$ . Again using the results from section 6.2, page 172 in the course notes we find that the distribution of  $\hat{\beta}_{WLS}$  is a multivariate normal distribution with mean  $\beta$  and variance-covariance matrix  $X^{-1} W X^{-T} = (X^T W^{-1} X)^{-1}$ . We conclude that:

$$\hat{\beta}_{WLS} \sim N_p(\beta, (X^T W^{-1} X)^{-1})$$

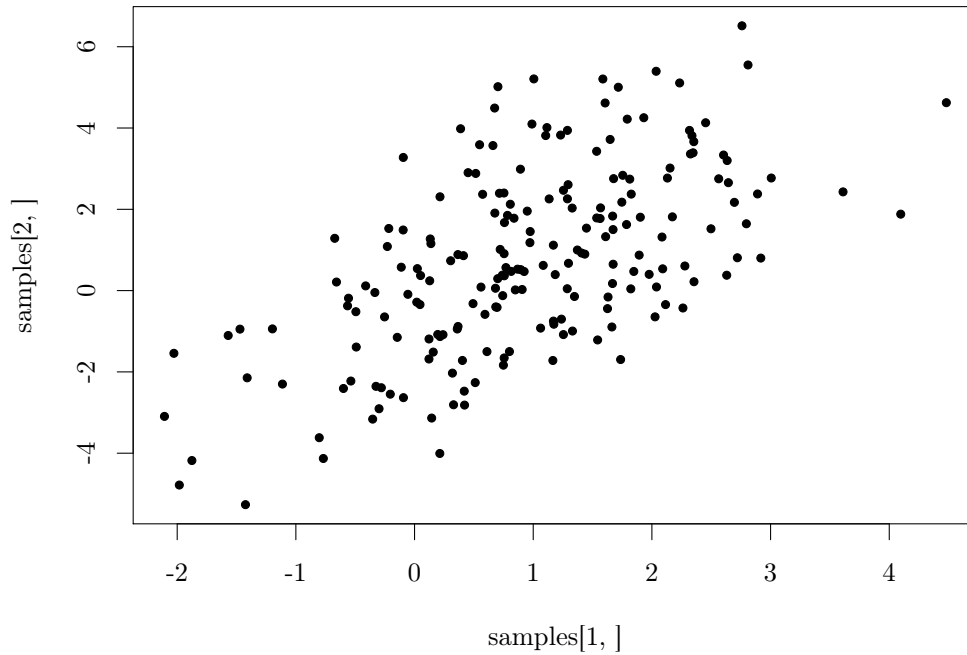


Figure 5: Projection on the  $xy$ -plane of the sample from the multivariate normal distribution

## Question 4

(a)

To generate a sample from a multi-normal distribution we work analogues to the derivation in section 6.2 in the course notes, but backwards. We are given the variance-covariance matrix  $\Sigma_X$  of a multivariate normal model. This matrix must be positive definite, so the square root of this matrix exists. Let  $A$  be a matrix such that  $AA^T = \Sigma_X$ . This matrix can be found by diagonalizing  $\Sigma_X$ :  $\Sigma_X = U\Lambda U^T$  and then taking  $A = U\sqrt{\Lambda}$ . We can do this because all eigenvalues of  $\Sigma_X$  are positive. The result from section 6.2 now tells us that  $X = AY + \mu$  where  $Y$  is a vector of standard normal distributions. This observation can be used to draw a sample from  $X$  by drawing  $n$  independent samples from a standard normal distribution, filling them in in a vector  $y$  and then calculating the sample from  $X$  as  $x = Ay + \mu$ . This was implemented in *R*

(b)

To compute these conditional chances we use proposition 6.2 from the course notes with the following vectors and matrices:

$$\begin{aligned}\Sigma_{11} &= (1.25) & \Sigma_{12} &= (1.50, 0.5) \\ \Sigma_{21} &= \begin{pmatrix} 1.50 \\ 0.50 \end{pmatrix} & \Sigma_{22} &= \begin{pmatrix} 5.25 & 3.5 \\ 3.5 & 3.0 \end{pmatrix} \\ \mathbf{X}_1 &= (\mathbf{X}_1) & \mathbf{x}_2 &= \begin{pmatrix} 1 \\ -2 \end{pmatrix}\end{aligned}$$

When we fill in these values in the the formula for the distribution of the conditional distribution  $f_{\mathbf{X}_1|\mathbf{X}_2}(\mathbf{x}_1, \mathbf{X}_2 = \mathbf{x}_2)$  we find the following distribution:

$$f_{\mathbf{X}_1|\mathbf{X}_2}(\mathbf{x}_1, \mathbf{X}_2 = \mathbf{x}_2) = N_1(1, 0.446) \quad (5)$$

The probability  $P(X_1 > 1 | X_2 = 1, X_3 = -2)$  is equal to 0.5 because of the symetry of the conditional probability distribution around 1.

To calculate the probability  $P(X_1 > 1 | X_2 + X_3 = 1)$  we first transform the random vector  $\mathbf{X}$  to the vector  $\mathbf{Y} = (X_1, X_2 + X_3)^T = \mathbf{A}\mathbf{X}$  where

$$\mathbf{A} = \begin{pmatrix} 1 & 0 & 0 \\ 0 & 1 & 1 \end{pmatrix}$$

Using part (b) of proposition 6.2 we find that the distribution of  $\mathbf{Y}$  is given by:

$$\mathbf{Y} \sim N(A\mu, \mathbf{A}\Sigma\mathbf{A}^T) = N(\dots)$$

Observe that  $P(X_1 > 1 | X_2 + X_3 = -1) = P(Y_1 > 1 | Y_2 = -1)$ . We can again use part (c) of proposition 6.2 to calculate this probability. This yields the following conditional probability distribution:

$$f_{\mathbf{Y}_1|\mathbf{Y}_2}(\mathbf{y}_1 | \mathbf{y}_2 = -1) = N(1, 0.9877)$$

Therefore  $P(X_1 > 1 | X_2 + X_3 = -1) = 0.5$ .

(c)

To determine the density function of  $\mathbf{Y} = (X_1, X_2)^T$  We make again use of proposition 6.2 (c) with the following matrix:

$$\mathbf{A} = \begin{pmatrix} 1 & 0 & 0 \\ 0 & 1 & 0 \end{pmatrix}$$

From which we find the probability distribution of  $\mathbf{Y}$

$$\mathbf{Y} \sim N_2 \left( \begin{pmatrix} 1 \\ 1 \end{pmatrix}, \begin{pmatrix} 1.25 & 1.50 \\ 1.50 & 5.25 \end{pmatrix} \right)$$

We can now use proposition 6.1 to find the contours of of this density function. The contour that comprises 95% of the probability mass is given by the following equation:

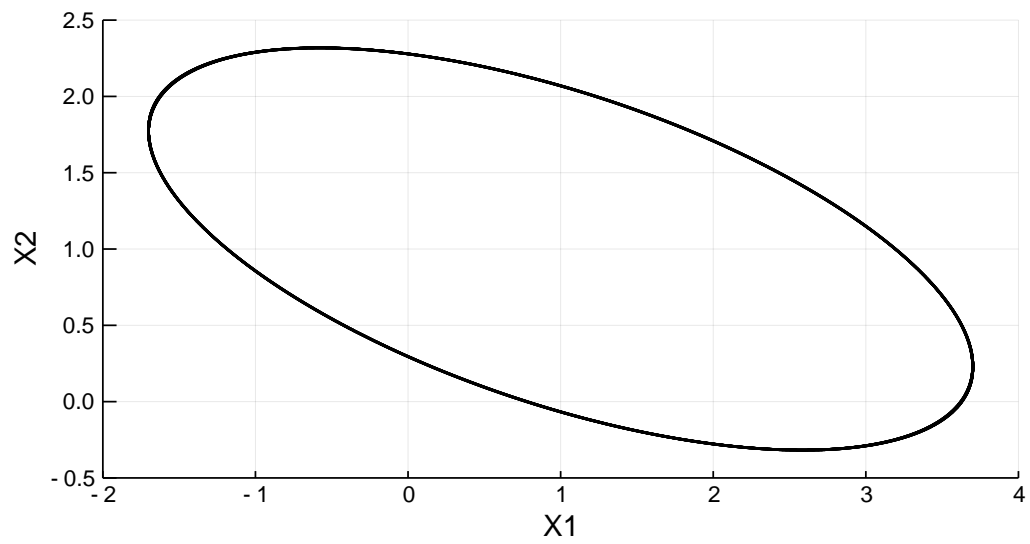
$$Q_{\chi^2_2}(0.95) = (\mathbf{Y} - \mu)^T \Sigma^{-1} (\mathbf{Y} - \mu)$$

This region is plotted in figure

## Acknowledgements

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Figuur 6: The contour that comprises 95% of the density of the probability density function of the random vector  $\mathbf{Y}$ .