

## Individual Player Analysis

Now that our data is properly configured, we can begin with our first stage of analysis. We will be investigating if, on the individual player level, the result of the first free throw affects the probability of making the second free throw.

We will do this by performing hypothesis tests with the null hypothesis  $H_0$ : Free throw 2 of 2 is independent of free throw 1 of 2. In other words, knowing the result of the first free throw does not affect the probability of making the second free throw.

```
[33]: import pandas as pd
import scipy.stats as st
import numpy as np

#Create a DataFrame by reading CSV
df = pd.read_csv(r'free_throw_counts.csv')

df.head(5)
```

```
[33]:      player  missed_1st_made_2nd  missed_first  made_1st_made_2nd  \
0    A.J. Price                21             28                62
1    Aaron Brooks              69             83               295
2    Aaron Gordon              27             38                53
3    Aaron Gray               32             65                51
4    Aaron Harrison              2              3                 0

      made_first
0              77
1            345
2              72
3              87
4               3
```

Because we will be performing multiple hypothesis tests for different players, we will need to adjust our significance level to account for the increased likelihood of making a Type I error (false positive) across multiple tests.

When we adjust our significance level, it will reduce the power of our individual hypothesis tests. So, we will only perform tests for the 50 players with the largest total number of free throws.

```
[34]: #Create a new column for the sum of 'missed_first' and 'made_first'
df['total'] = df['missed_first'] + df['made_first']

#Sort the DataFrame by the 'total_shots' column in descending order
df = df.sort_values(by='total', ascending=False)

#Remove the 'total' column
df.drop('total', axis='columns', inplace=True)
```

```
#Only keep the 50 players with the largest total number of free throws
df=df.head(50)

df.head(5)
```

```
[34]:
```

	player	missed_1st_made_2nd	missed_first	made_1st_made_2nd	\
293	Dwight Howard	853	1589	1134	
607	LeBron James	730	951	1871	
563	Kevin Durant	284	322	1854	
295	Dwyane Wade	472	622	1424	
576	Kobe Bryant	333	382	1607	

  

	made_first
293	1887
607	2389
563	2086
295	1784
576	1872

We will assume that pairs of free throws are independent of each other and that our probabilities of success within each group (made/missed first free throw) remain constant for all trials.

Let the random variable  $Y_1$  represent the number of successful 2nd free throw attempts in a sample of  $n_1$  free throw pairs, where the 1st free throw was unsuccessful, drawn from a large population.

Based on our assumptions,  $Y_1$  can be closely modeled by a binomial distribution  $BIN(\theta_1)$ , where  $\theta_1$  represents the probability of making the 2nd free throw after missing the 1st free throw.

Our maximum likelihood estimate for  $\theta_1$  is  $\hat{\theta}_1 = \frac{y_1}{n_1}$ .

As an example, our observed value for LeBron James is  $y_1 = 730$  and  $n_1 = 951$ , giving us  $\hat{\theta}_1 = \frac{730}{951} \approx 0.7676$ .

Defining  $Y_2$  as the number of successful 2nd free throw attempts in a sample of  $n_2$  free throw pairs where the 1st free throw was successful and applying the same process, we get  $\hat{\theta}_2 = \frac{1871}{2389} \approx 0.7832$ , where  $\theta_2$  is defined as the probability of making the 2nd free throw after making the 1st free throw.

We would like to see if the difference between  $\theta_1$  and  $\theta_2$  is statistically significant at the 0.05 significance level.

Our null hypothesis is  $H_0 : \theta_1 = \theta_2$ .

From the National Institute of Standards and Technology ([www.itl.nist.gov/div898/handbook/prc/section3/prc33.htm](http://www.itl.nist.gov/div898/handbook/prc/section3/prc33.htm)), we have that

$$\frac{\tilde{\theta}_1 - \tilde{\theta}_2}{\sqrt{\tilde{\theta}(1-\tilde{\theta})(\frac{1}{n_1} + \frac{1}{n_2})}} \sim G(0, 1), \text{ where } \tilde{\theta} = \frac{n_1\tilde{\theta}_1 + n_2\tilde{\theta}_2}{n_1 + n_2}$$

Our test statistic is  $D = |Z|$ , and our observed value of  $D$  is  $d \approx 0.9777$ .

Our approximate p-value based on the Gaussian approximation is

$$p\text{-value} = P(D \geq d; H_0)$$

$$\begin{aligned}
&= P\left(\frac{|\tilde{\theta}_1 - \tilde{\theta}_2|}{\sqrt{\tilde{\theta}(1-\tilde{\theta})\left(\frac{1}{n_1} + \frac{1}{n_2}\right)}} \geq 0.9777\right) \\
&\approx P(|Z| \geq 0.9777) \\
&= 2[1 - P(Z \leq 0.9777)] \\
&\approx 2(1 - 0.8359) \\
&\approx 0.3282
\end{aligned}$$

Since our p-value is greater than 0.05, we conclude that there is insufficient evidence to reject the null hypothesis for LeBron James.

We can repeat this process for each of our 50 players using python.

```
[35]: df['theta_hat_1'] = df['missed_1st_made_2nd'] / df['missed_first']

df['theta_hat_2'] = df['made_1st_made_2nd'] / df['made_first']

df['theta_tilde'] = ((df['missed_first'] * df['theta_hat_1'] +
                    df['made_first'] * df['theta_hat_2']) /
                    (df['missed_first'] + df['made_first']))

df['d'] = np.abs((df['theta_hat_1'] - df['theta_hat_2']) /
                np.sqrt(df['theta_tilde'] * (1-df['theta_tilde']) *
                        (1/df['missed_first'] + 1/df['made_first'])))

df['p-val'] = 2 * (1 - st.norm.cdf(df['d']))

df.head(5)
```

```
[35]:
```

	player	missed_1st_made_2nd	missed_first	made_1st_made_2nd	\
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	made_first	theta_hat_1	theta_hat_2	theta_tilde	d	p-val
293	1887	0.536816	0.600954	0.571634	3.806789	0.000141
607	2389	0.767613	0.783173	0.778743	0.977652	0.328247
563	2086	0.881988	0.888782	0.887874	0.359669	0.719095
295	1784	0.758842	0.798206	0.788030	2.068395	0.038603
576	1872	0.871728	0.858440	0.860692	0.683505	0.494288

We can already see a few players that seem to provide sufficient evidence to reject the null hypothesis at the 0.05 level. However, we still need to make some adjustments to account for multiple hypothesis tests.

To do this, we will use the Benjamini-Hochberge Procedure (<https://www.statisticshowto.com/benjamini-hochberg-procedure/>).

```
[36]: #Sort the rows from smallest to largest p-value
df = df.sort_values(by='p-val', ascending=True)

#Reset the dataframe indices
df = df.reset_index()
df.drop('index', axis='columns', inplace=True)

#Add the critical value as a column
df['critical val'] = ((df.index + 1) / 50) * 0.05

df.head(5)
```

```
[36]:
```

	player	missed_1st_made_2nd	missed_first	made_1st_made_2nd	\
0	Dwight Howard	853	1589	1134	
1	Josh Smith	380	628	629	
2	Corey Maggette	192	241	881	
3	Andre Iguodala	345	496	752	
4	LaMarcus Aldridge	220	293	982	

  

	made_first	theta_hat_1	theta_hat_2	theta_tilde	d	p-val	\
0	1887	0.536816	0.600954	0.571634	3.806789	0.000141	
1	903	0.605096	0.696567	0.659046	3.713779	0.000204	
2	1016	0.796680	0.867126	0.853620	2.781425	0.005412	
3	990	0.695565	0.759596	0.738223	2.647787	0.008102	
4	1201	0.750853	0.817652	0.804552	2.585265	0.009730	

  

	critical val
0	0.001
1	0.002
2	0.003
3	0.004
4	0.005

```
[37]: #Find the largest p-value less than or equal to their critical value
pval = (df[df['p-val'] <= df['critical val']]).max()['p-val']

#Players with p-values less than or equal to this value are significant
df_stat_sig = df[df['p-val'] <= pval]

print(df_stat_sig)
```

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	made_first	theta_hat_1	theta_hat_2	theta_tilde	d	p-val	\
0	1887	0.536816	0.600954	0.571634	3.806789	0.000141	
1	903	0.605096	0.696567	0.659046	3.713779	0.000204	

	critical val
0	0.001
1	0.002

Therefore, of our 50 players, only Dwight Howard and Josh Smith have statistically significant differences between the probability of them making the second free throw after making and after missing the first free throw.

Both are more likely to make the second free throw if they make the first free throw.

Dwight Howard is approximately 6.4% more likely to make the second free throw if he makes the first.

Josh Smith is approximately 9.1% more likely to make the second free throw if he makes the first.

So, not only is the difference statistically significant for these players, the difference is practically significant.