

1. $\{X_t \text{ for } t = \dots, -2, -1, 0, 1, 2, \dots\}$

a7. When ①. $E[X_t] = \mu$ is constant

②. $Var(X_t) = \sigma^2$ is constant

$$\textcircled{3}. \gamma(t_1, t_2) = \text{Cov}(t_1, t_2) = E\{[Y_{t_1} - E(Y_{t_1})][Y_{t_2} - E(Y_{t_2})]\} = \gamma(t_1 - t_2)$$

we could say that X_t is weakly stationary

b7. $X_t = \sum_{j=0}^4 \alpha_j \varepsilon_{t-j}$, $\varepsilon_t' \sim \text{iid}(0, \sigma_\varepsilon^2)$ & t'

$$E[X_t] = E\left[\sum_{j=0}^4 \alpha_j \varepsilon_{t-j}\right] = \sum_{j=0}^4 \alpha_j E[\varepsilon_{t-j}] = 0 \quad \text{is constant}$$

$$Var(X_t) = Var\left(\sum_{j=0}^4 \alpha_j \varepsilon_{t-j}\right) = \sum_{j=0}^4 \alpha_j^2 Var(\varepsilon_{t-j}) = \sigma_\varepsilon^2 \cdot \sum_{j=0}^4 \alpha_j^2 \quad \text{is constant}$$

$$\gamma(k) = \text{Cov}(X_t, X_{t-k}) = \text{Cov}\left(\sum_{j=0}^4 \alpha_j \varepsilon_{t-j}, X_{t-k}\right)$$

$$= \sum_{j=0}^4 \alpha_j \text{Cov}(\varepsilon_{t-j}, X_{t-k})$$

$$= \sum_{j=0}^4 \alpha_j \text{Cov}(\varepsilon_{t-j}, \sum_{i=0}^4 \alpha_i \varepsilon_{t-k-i})$$

$$= \sum_{j=0}^4 \sum_{i=0}^4 \alpha_j \alpha_i \text{Cov}(\varepsilon_{t-j}, \varepsilon_{t-k-i})$$

$$\text{Cov}(\varepsilon_{t-j}, \varepsilon_{t-k-i}) = \begin{cases} 0 & \text{if } t-j \neq t-k-i \Rightarrow i \neq k+j \\ \sigma^2 & \text{if } t-j = t-k-i \Rightarrow i = k+j \end{cases}$$

$$\gamma(k) = \sum_{j=0}^4 \alpha_j \alpha_{k+j} \sigma^2 = \sigma^2 \sum_{j=0}^4 \alpha_j \alpha_{j+k}$$

$\Rightarrow X_t$ is weakly stationary

c7. $X_t = \sum_{j=0}^n \alpha_j \varepsilon_{t-j}$, $\varepsilon_t' \sim \text{iid}(0, \sigma_\varepsilon^2)$

$$E[X_t] = E\left[\sum_{j=0}^n \alpha_j \varepsilon_{t-j}\right] = \sum_{j=0}^n \alpha_j E[\varepsilon_{t-j}] = 0 \quad \text{is constant}$$

$$Var(X_t) = Var\left(\sum_{j=0}^n \alpha_j \varepsilon_{t-j}\right) = \sum_{j=0}^n \alpha_j^2 Var(\varepsilon_{t-j}) = \sigma_\varepsilon^2 \sum_{j=0}^n \alpha_j^2 \quad \text{is constant}$$

$$\gamma(k) = \text{Cov}(X_t, X_{t-k}) = \text{Cov}\left(\sum_{j=0}^n \alpha_j \varepsilon_{t-j}, X_{t-k}\right)$$

$$= \sum_{j=0}^n \sum_{i=0}^n \alpha_j \alpha_i \text{Cov}(\varepsilon_{t-j}, \varepsilon_{t-k-i})$$

$$\text{Cov}(\varepsilon_{t-j}, \varepsilon_{t-k-i}) = \begin{cases} 0 & \text{if } t-j = t-k-i \Rightarrow i = k+j \\ \sigma^2 & \text{if } t-j \neq t-k-i \Rightarrow i \neq k+j \end{cases}$$

$$\gamma(k) = \sum_{j=0}^n \alpha_j \alpha_{k+j} \sigma^2 = \sigma^2 \sum_{j=0}^n \alpha_j \alpha_{j+k}$$

$\Rightarrow X_t$ is weakly stationary

d7. $X_t = \sum_{j=0}^{\infty} \alpha_j \varepsilon_{t-j}$, $\varepsilon_t \stackrel{iid}{\sim} (0, \sigma_\varepsilon^2) \quad \forall t$

$$E[X_t] = E\left[\sum_{j=0}^{\infty} \alpha_j \varepsilon_{t-j}\right] = \sum_{j=0}^{\infty} \alpha_j E[\varepsilon_{t-j}] = 0 \quad \text{is constant}$$

$$\text{Var}(X_t) = \text{Var}\left(\sum_{j=0}^{\infty} \alpha_j \varepsilon_{t-j}\right) = \sum_{j=0}^{\infty} \alpha_j^2 \text{Var}(\varepsilon_{t-j}) = \sigma_\varepsilon^2 \sum_{j=0}^{\infty} \alpha_j^2 < \sigma_\varepsilon^2 \cdot C < \infty \text{ is finite and constant}$$

$$\gamma(k) = \text{Cov}(X_t, X_{t-k}) = \text{Cov}\left(\sum_{j=0}^{\infty} \alpha_j \varepsilon_{t-j}, X_{t-k}\right)$$

$$= \sum_{j=0}^{\infty} \sum_{i=0}^{\infty} \alpha_j \alpha_i \text{Cov}(\varepsilon_{t-j}, \varepsilon_{t-k-i})$$

$$\text{Cov}(\varepsilon_{t-j}, \varepsilon_{t-k-i}) = \begin{cases} 0 & \text{if } t-j = t-k-i \Rightarrow i = k+j \\ \sigma^2 & \text{if } t-j \neq t-k-i \Rightarrow i \neq k+j \end{cases}$$

$$\gamma(k) = \sum_{j=0}^{\infty} \alpha_j \alpha_{k+j} \sigma^2 = \sigma^2 \sum_{j=0}^{\infty} \alpha_j \alpha_{j+k}$$

$\Rightarrow X_t$ is weakly stationary if $\sum_{j=0}^{\infty} \alpha_j^2 < C < \infty$

2. AR(1) $y_t = 0.5 y_{t-1} + \varepsilon_t$, $\varepsilon_t \stackrel{iid}{\sim} (0, 1)$

a) $y_t = \sum_{j=0}^{\infty} \alpha_j \varepsilon_{t-j}$

$$y_t = 0.5 y_{t-1} + \varepsilon_t$$

$$= 0.5 (0.5 y_{t-2} + \varepsilon_{t-1}) + \varepsilon_t$$

⋮

$$= \sum_{j=0}^{\infty} 0.5^j \varepsilon_{t-j}$$

$$\Rightarrow \alpha_j = 0.5^j$$

b) $\mu = E[y_t] = E[0.5 y_{t-1} + \varepsilon_t] = 0.5 E[y_{t-1}] = 0.5 \mu \Rightarrow \mu = 0$

c) $\text{Var}(y_t) = \gamma(0) = \text{Var}(0.5 y_{t-1} + \varepsilon_t) = \frac{1^2}{1-0.5^2} = \frac{4}{3}$

d) $\text{cov}(y_t, y_{t-k}) = \gamma(k) = \text{cov}(0.5 y_{t-1} + \varepsilon_t, y_{t-k})$

$$= 0.5 \text{cov}(y_{t-1}, y_{t-k})$$

$$= 0.5 \gamma(k-1)$$

$$= 0.5^2 \gamma(k-2)$$

⋮

$$= 0.5^k \gamma(0)$$

$$= 0.5^k \cdot \frac{4}{3}$$

PS1Q3

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```
library("quantmod")
```

```
##(a) Compile quarterly data for the U.S. real gross private domestic investment (DI) from 1947Q1 to 2023Q4.
```

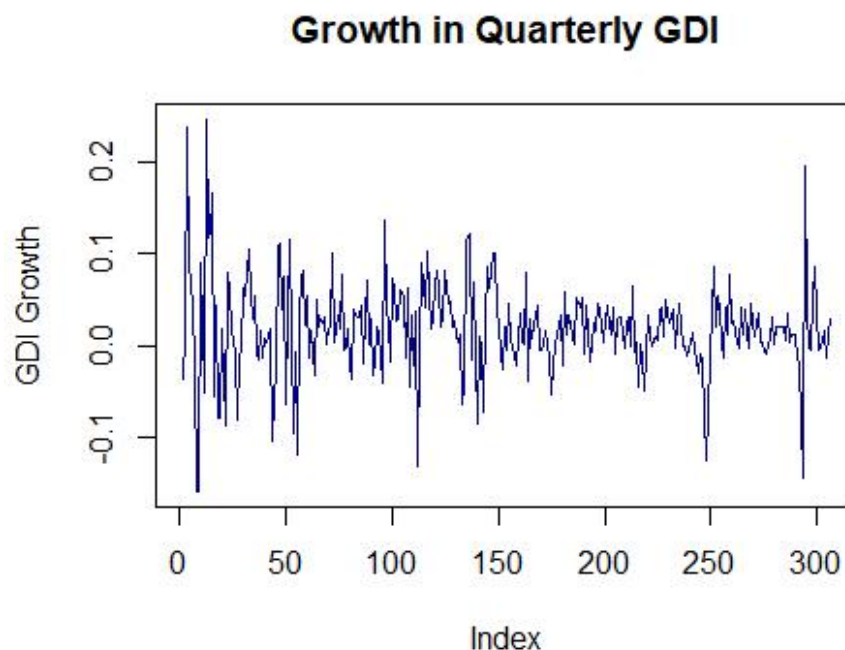
```
getSymbols(Symbols = "GPDI", src = "FRED", from = '1947/01/01')
```

```
## [1] "GPDI"
```

```
qua_di <- as.matrix(GPDI[,1])  
qua_di_date = as.Date(row.names(qua_di))  
n_obs_qua_di = length(qua_di_date)
```

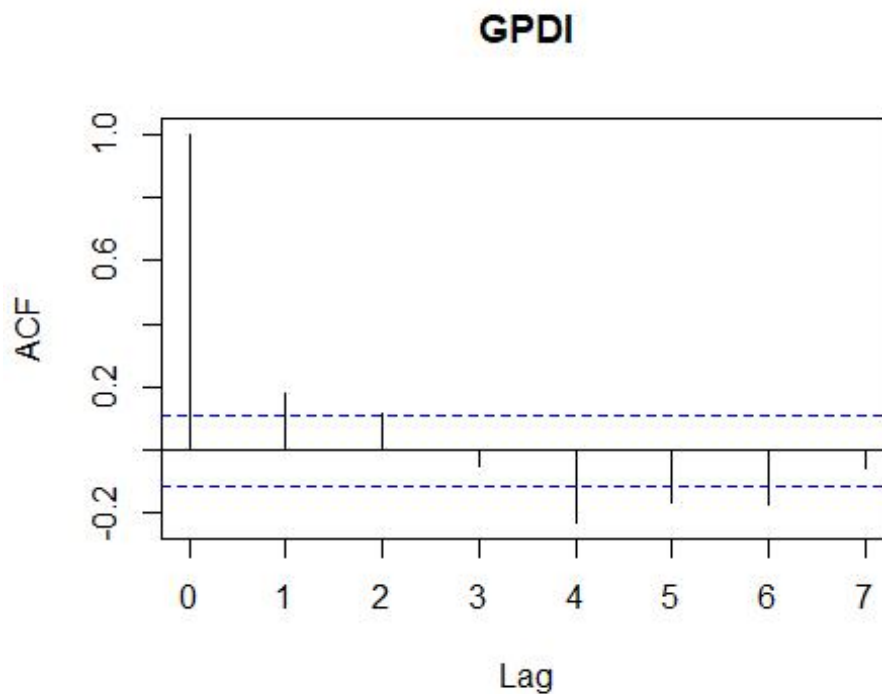
```
##(b) Compute growth in quarterly DI (GDI), provide its summary statistics and plot the data.
```

```
di_return <- diff(qua_di)/qua_di[1:n_obs_qua_di-1,1]  
di_return_date = qua_di_date[2:n_obs_qua_di]  
n_obs_qua_return = length(di_return_date)  
plot(di_return, type = "l", col = "darkblue", main = "Growth in Quarterly GDI", ylab = "GDI Growth")
```



##(c) Compute and plot empirical autocorrelation function. Given the plot, do you expect any time-series correlation among the observations? Explain why?

```
acf(di_return, lag=round(n_obs_qua_return^(1/3)))
```



```
ACF=acf(di_return, lag=round(n_obs_qua_return^(1/3)), plot = FALSE)
ACF$acf
```

```
## , , 1
##
##      [,1]
## [1,] 1.00000000
## [2,] 0.18230712
## [3,] 0.11675470
## [4,] -0.04992628
## [5,] -0.22861728
## [6,] -0.16254377
## [7,] -0.17190977
## [8,] -0.05641716
```

Yes I do expect some time-series correlation among the observations. Because there are some lags where the bars extend beyond the dotted lines.

##(d) Set the maximum number of lags to the integer closest to the number of observations to the power one-third. Perform a test for joint autocorrelation in GDI and report your result. Does your finding consistent with that of Part 3c? Explain why?

```
t_ratio <- ACF$acf[2]*sqrt(n_obs_qua_return)
t_ratio

## [1] 3.189072

Box.test(di_return, lag = round(n_obs_qua_return^(1/3)), type = "Ljung-Box")

##
## Box-Ljung test
##
## data: di_return
## X-squared = 50.143, df = 7, p-value = 1.354e-08

Box.test(di_return, lag = round(n_obs_qua_return^(1/3)), type = "Box-Pierce")

##
## Box-Pierce test
##
## data: di_return
## X-squared = 49.199, df = 7, p-value = 2.074e-08
```

##(e) Consider an AR(1) model and compute the theoretical autocorrelation function. Compare your findings with that of Part 3c.

```
lag_di_return = rbind(NA, as.matrix(di_return[1:(n_obs_qua_return-1),1]))
intercept = matrix(1,n_obs_qua_return)
X = cbind(intercept,lag_di_return)
y = di_return
reg_result = ols(X[2:n_obs_qua_return,],as.matrix(y[2:n_obs_qua_return,1]))
1 - sum(reg_result$u_hat^2)/sum(y^2)

## [1] 0.1381177

beta_hat = reg_result$beta_hat
beta_hat

##           [,1]
## [1,] 0.01449087
## [2,] 0.18233574

var_beta_hat = reg_result$var_beta_hat
test_result = t_test(beta_hat,var_beta_hat)
test_result$t_stat

##           [,1]
## [1,] 4.769901
## [2,] 3.244819
```

```

test_result$p_value

##           [,1]
## [1,] 1.843163e-06
## [2,] 1.175253e-03

ar_coeff <- as.numeric(beta_hat[2])
ma_coeff <- 0
TACF <- ARMAacf(ar_coeff, ma_coeff, lag.max = round(n_obs_qua_return^(1/3)))
plot(c(0:round(n_obs_qua_return^(1/3))),ACF$acf,type='l',xlab='Lag',ylab='ACF',
ylim=c(-0.1,1))
lines(0:round(n_obs_qua_return^(1/3)),TACF,lty=2)
grid(nx = 4, ny = 4)

```

