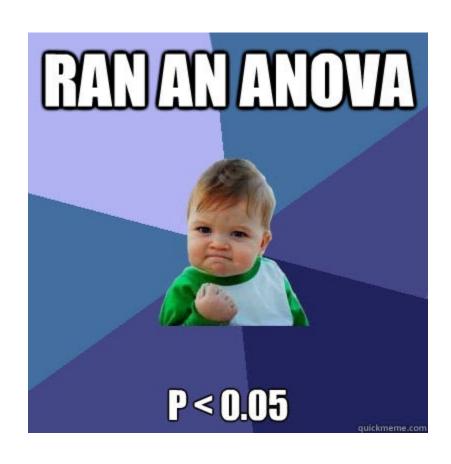
Analysis of Variance



Overview

One-way analysis of variance (ANOVA) concepts and R

Connections between ANOVA and linear regression

Paired tests after running an ANOVA

If there is time: explanation of experimental data used for homework

Any questions about the final project?

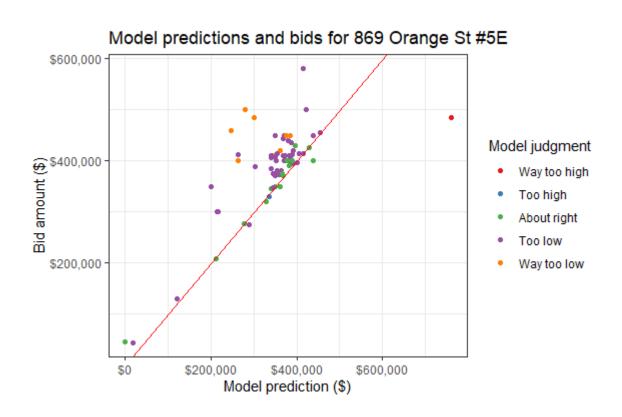
Note: you can submit a collaborative project with one other person

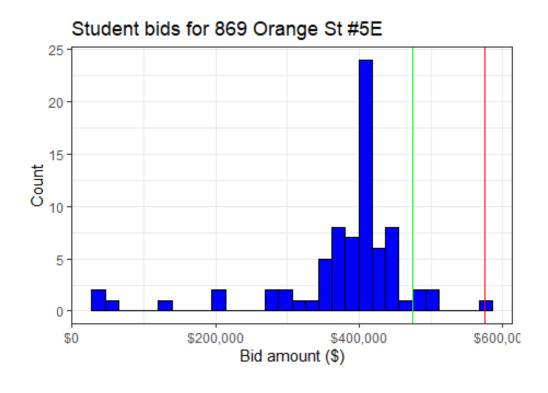
Project should be twice as long and twice as impressive!

• i.e., 10-16 pages, more in depth analyses, etc.



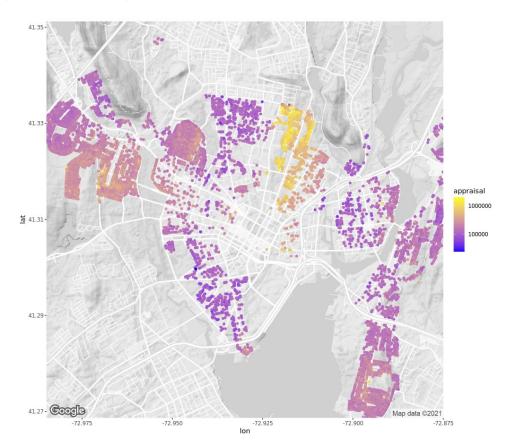
How did predicting condo process go?





How did predicting condo process go?

Map of single-family house prices by location in New Haven





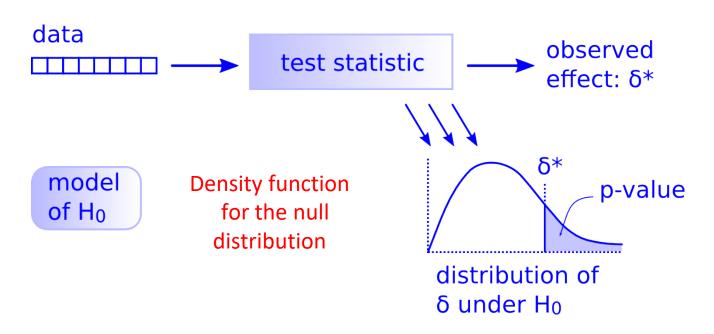
The company shocked the market in November by <u>announcing it was closing</u> <u>Zillow Offers</u> because the <u>tech-powered platform failed</u> to accurately predict movements in home prices.

One-way analysis of variance (ANOVA)

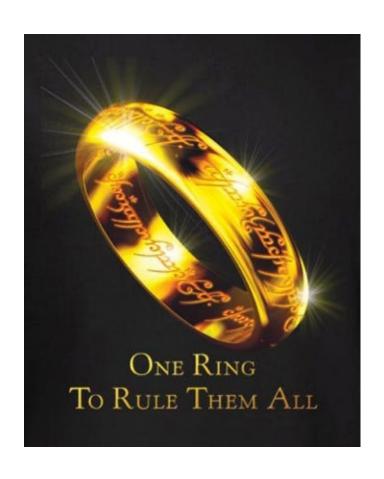
One-way ANOVA

A **one-way analysis of variance (ANOVA)** is a parametric hypothesis test that can be used to examine if a set of means are all the same

There is only one <u>hypothesis test!</u>



Just follow the 5 hypothesis tests steps!



Faculty salaries again...

Did you find it useful that used the same data sets for several weeks (i.e., the faculty salaries and car transactions data sets), or would you have preferred that we had used different data set each week?

I found it useful to use the same data set for several weeks	54 respondents	69 %	\ <u></u>
I have no preference	13 respondents	17 %	
I would have preferred using new data sets each week	11 respondents	14 %	

Silly question: Do Assistant, Associate and Fully Professors get paid the same on average?



One-way ANOVA

A **one-way analysis of variance (ANOVA)** is a parametric hypothesis test that can be used to examine if a set of means are all the same.

$$H_0$$
: $\mu_1 = \mu_2 = ... = \mu_k$

 H_A : $\mu_i \neq \mu_j$ for some i, j

The statistic we use for a one-way ANOVA is the F-statistic

$$F = \frac{\text{between-group variability}}{\text{within-group variability}} = \frac{\frac{1}{K-1} \sum_{i=1}^{K} n_i (\bar{x}_i - \bar{x}_{tot})^2}{\frac{1}{N-K} \sum_{i=1}^{K} \sum_{j=1}^{n_i} (x_{ij} - \bar{x}_i)^2}$$

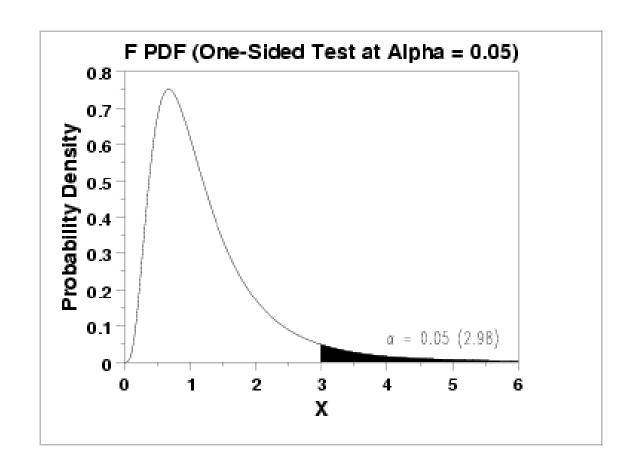
One-way ANOVA — the central idea

If H₀ is true, the F-statistic will come from an F distribution with parameters

- $df_1 = K 1$
- $df_2 = N K$

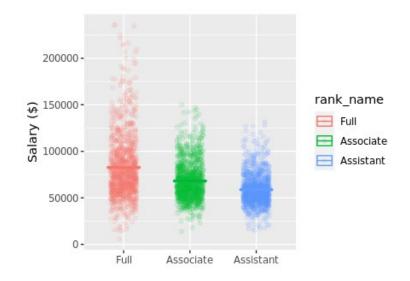
The F-distribution is valid if these conditions are met:

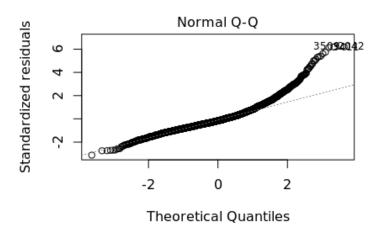
- The data in each group should follow a normal distribution
- The variances in each group should be approximately equal



Checking ANOVA conditions ('assumptions')

- 1. We can check if the data in each group is relatively normal by visually examining the residuals between each point and its group mean:
 - Residuals as a function of the group mean
 - Q-Q plots
 - Histograms of residuals
- 2. We can check the equal variance condition by seeing if the ratio of the largest to smallest standard deviation is greater than 2
 - $s_{max}/s_{min} < 2$





Note: the one-way ANOVA is fairly robust to violations of these conditions

Calculating the observed F-statistic

$$F = \frac{\text{between-group variability}}{\text{within-group variability}} = \frac{\frac{1}{K-1} \sum_{i=1}^{K} n_i (\bar{x}_i - \bar{x}_{tot})^2}{\frac{1}{N-K} \sum_{i=1}^{K} \sum_{j=1}^{n_i} (x_{ij} - \bar{x}_i)^2}$$

K: the number of groups

N: total number of points

 \overline{X}_{tot} : the mean across all the data

 \overline{x}_i : the mean of group i

n_i: the number of points in group i

x_{ij}: the jth data point from group i

K = 3 different ranks

N = 690 total salaries

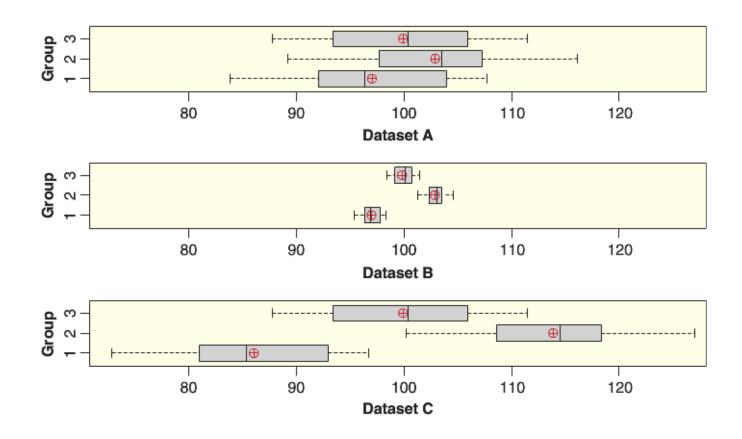
 \overline{x}_{tot} : the mean salary over all professors

 \overline{x}_i : the mean salary for the ith professor rank

 n_i = num professors of the i^{th} rank (330)

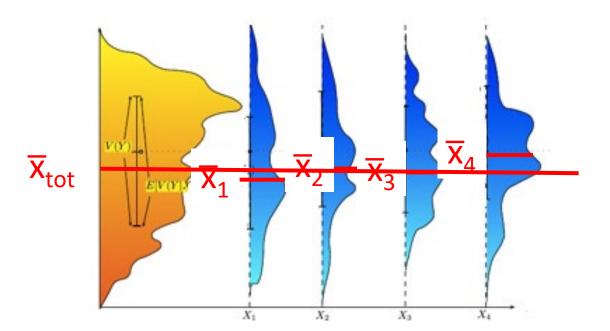
x_{ii}: the jth professor salary in the ith rank

Why use the F-Statistic?



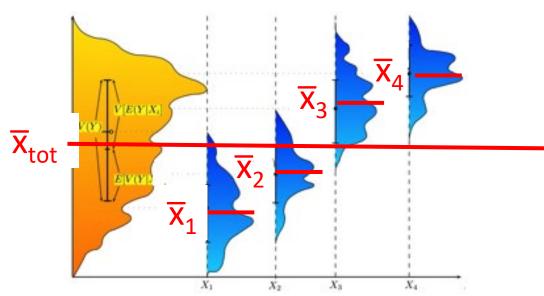
Which dataset gives the strongest evidence that there is a difference in population means?

If H₀ is true, the data from all groups have the same means



- Similar means \bar{x}_i
- Similar spread s_i

If H₀ is not true, the data from all groups do not have the same mean

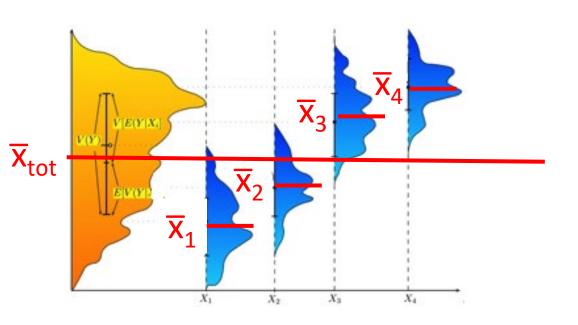


- Different means \overline{x}_i
- Smaller spreads s_i

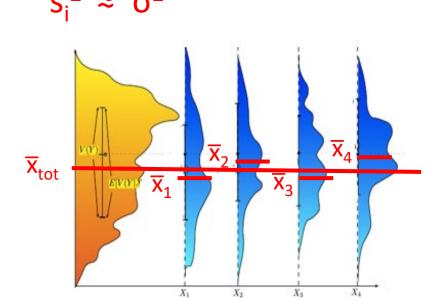
$$F = \frac{\text{between-group variability}}{\text{within-group variability}} = \frac{\frac{1}{K-1} \sum_{i=1}^{K} n_i (\bar{x}_i - \bar{x}_{tot})^2}{\frac{1}{N-K} \sum_{i=1}^{K} \sum_{j=1}^{n_i} (x_{ij} - \bar{x}_i)^2}$$

The F statistic measures a fraction of:

variability between group means variability within each group



$$F = \frac{\text{between-group variability}}{\text{within-group variability}} = \frac{\frac{1}{K-1} \sum_{i=1}^{K} n_i (\bar{x}_i - \bar{x}_{tot})^2}{\frac{1}{N-K} \sum_{i=1}^{K} \sum_{j=1}^{n_i} (x_{ij} - \bar{x}_i)^2}$$

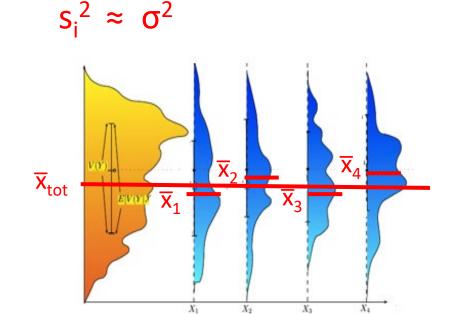


$$F = \frac{\text{between-group variability}}{\text{within-group variability}} = \frac{\frac{1}{K-1} \sum_{i=1}^{K} n_i (\bar{x}_i - \bar{x}_{tot})^2}{\frac{1}{N-K} \sum_{i=1}^{K} \sum_{j=1}^{n_i} (x_{ij} - \bar{x}_i)^2}$$

The F statistic measures a fraction of:

variability between group means

$$= \frac{\text{variability between group filed}}{\approx \sigma^2}$$

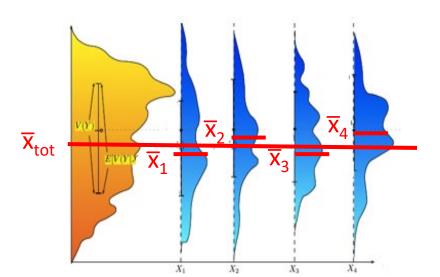


$$F = \frac{\text{between-group variability}}{\text{within-group variability}} = \frac{\frac{1}{K-1} \sum_{i=1}^{K} n_i (\bar{x}_i - \bar{x}_{tot})^2}{\frac{1}{N-K} \sum_{i=1}^{K} \sum_{j=1}^{n_i} (x_{ij} - \bar{x}_i)^2}$$

The F statistic measures a fraction of:

variability between group means

$$F = \frac{\sqrt{\alpha \pi a \sin t y} \sec w \cot g \cos p \cot p}{\approx \sigma^2}$$



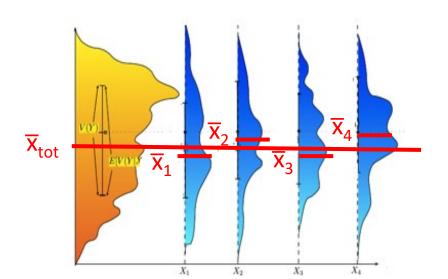
SE²

$$F = \frac{\text{between-group variability}}{\text{within-group variability}}$$

$$\frac{\frac{1}{K-1} \sum_{i=1}^{K} n_i (\bar{x}_i - \bar{x}_{tot})^2}{\frac{1}{N-K} \sum_{i=1}^{K} \sum_{j=1}^{n_i} (x_{ij} - \bar{x}_i)^2}$$

$$SE^2 \approx \sigma^2/n$$

$$= \frac{\approx \sigma^2}{\approx \sigma^2} \approx 1$$



Sum of Squares Group (SSG)

$$F = \frac{\text{between-group variability}}{\text{within-group variability}} = \frac{\frac{1}{K-1} \sum_{i=1}^{K} n_i (\bar{x}_i - \bar{x}_{tot})^2}{\frac{1}{N-K} \sum_{i=1}^{K} \sum_{j=1}^{n_i} (x_{ij} - \bar{x}_i)^2}$$

$$F = \frac{\approx \sigma^2}{\approx \sigma^2} \approx 1$$

Mean Squares Group (MSG)

$$F = \frac{\text{between-group variability}}{\text{within-group variability}} = \frac{\frac{1}{K-1} \sum_{i=1}^{K} n_i (\bar{x}_i - \bar{x}_{tot})^2}{\frac{1}{N-K} \sum_{i=1}^{K} \sum_{j=1}^{n_i} (x_{ij} - \bar{x}_i)^2}$$

$$F = \frac{MSG}{\approx \sigma^2} \approx 1$$

Sum of Squares Error (SSE)

$$F = \frac{\text{between-group variability}}{\text{within-group variability}} = \frac{\frac{1}{K-1} \sum_{i=1}^{K} n_i (\bar{x}_i - \bar{x}_{tot})^2}{\frac{1}{N-K} \sum_{i=1}^{K} \sum_{j=1}^{n_i} (x_{ij} - \bar{x}_i)^2}$$

$$F = \frac{MSG}{\approx \sigma^2} \approx 1$$

Mean of Squares Error (MSE)

$$F = \frac{\text{between-group variability}}{\text{within-group variability}} = \frac{\frac{1}{K-1} \sum_{i=1}^{K} n_i (\bar{x}_i - \bar{x}_{tot})^2}{\frac{1}{N-K} \sum_{i=1}^{K} \sum_{j=1}^{n_i} (x_{ij} - \bar{x}_i)^2}$$

The F statistic measures a fraction of:

Same as what we saw in regression: SSTotal = SSG + SSE

ANOVA table

Source	df	Sum of Sq.	Mean Square	F-statistic	p-value
Groups	k – 1	SSG	$MSG = rac{SSG}{k-1}$	$F=rac{MSG}{MSE}$	Upper tail $F_{k-1,n-k}$
Error	n – k	SSE	$MSE = rac{SSE}{n-k}$		
Total	n – 1	SSTotal			

Where:

$$SST = \sum_{i=1}^{k} \sum_{j=1}^{n_i} (x_{ij} - \bar{x}_{tot})^2$$

$$SSG = \sum_{i=1}^{k} n_i (\bar{x}_i - \bar{x}_{tot})^2$$

$$SSE = \sum_{i=1}^{k} \sum_{j=1}^{n_i} (x_{ij} - \bar{x}_i)^2$$

ANOVA table

Just as we saw for linear regression, we have the relationship:

$$SST = SSG + SSE$$

Where:

$$SST = \sum_{i=1}^{k} \sum_{j=1}^{n_i} (x_{ij} - \bar{x}_{tot})^2$$

$$SSG = \sum_{i=1}^{k} n_i (\bar{x}_i - \bar{x}_{tot})^2$$

$$SSE = \sum_{i=1}^{k} \sum_{j=1}^{n_i} (x_{ij} - \bar{x}_i)^2$$

Running a one-way ANOVA

Step 1: State the null and alternative hypothesis

Step 2: Calculate the F-statistic on using actual data

Step 3: Create the appropriate F-distribution

Step 4: Calculate the p-value

Step 5: Make a decision



Check our underlying assumptions were met

Let's try it out in R!



Connections between regression and ANOVAs

ANOVA as regression with only categorical predictors

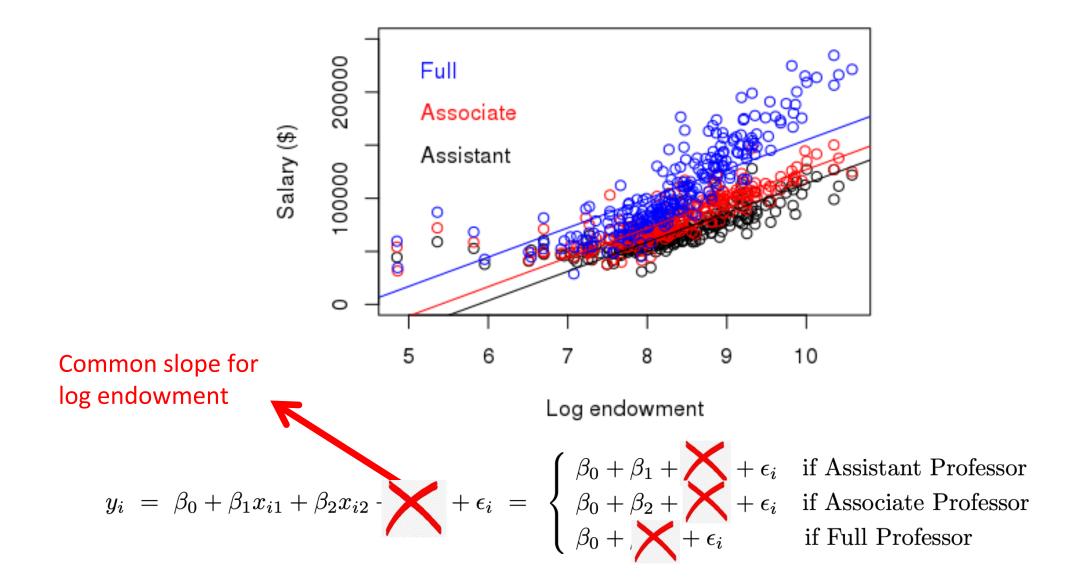
Recall we can have categorical predictors with k levels in a regression model by using k -1 dummy variables:

 e.g., we would need two dummy variables to have different intercepts for Assistant, Associate and Full Professors

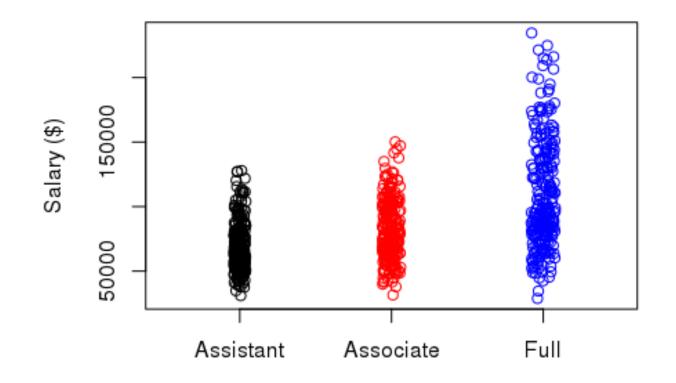
$$x_{i1} = \begin{cases} 1 & \text{if Assistant Professor} \\ 0 & \text{if Full Professor} \end{cases} \qquad x_{i2} = \begin{cases} 1 & \text{if Associate Professor} \\ 0 & \text{if Full Professor} \end{cases}$$

$$y_{i} = \beta_{0} + \beta_{1}x_{i1} + \beta_{2}x_{i2} + \beta_{3}x_{i} + \epsilon_{i} = \begin{cases} \beta_{0} + \beta_{1} + \beta_{3}x_{i} + \epsilon_{i} & \text{if Assistant Professor} \\ \beta_{0} + \beta_{2} + \beta_{3}x_{i} + \epsilon_{i} & \text{if Full Professor} \end{cases}$$

ANOVA as regression with only categorical predictors



ANOVA as regression with only categorical predictors



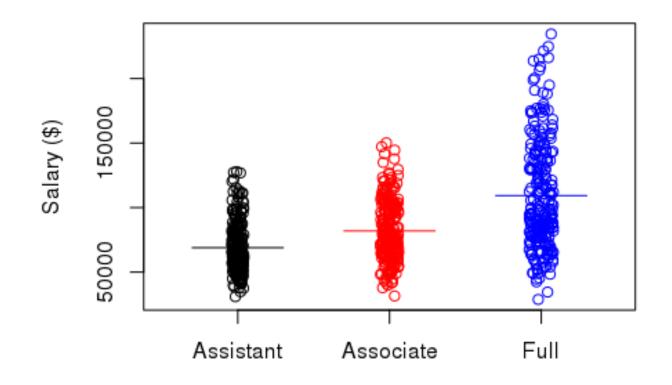
$$y_i = \beta_0 + \beta_1 x_{i1} + \beta_2 x_{i2} + \epsilon_i = \begin{cases} \beta_0 + \beta_1 + \epsilon_i & \text{if Assistant Professor} \\ \beta_0 + \beta_2 + \epsilon_i & \text{if Associate Professor} \end{cases}$$
 $\mu_1 = \beta_0 + \beta_1$
 $\beta_0 + \beta_2 + \epsilon_i & \text{if Full Professor} \end{cases}$ $\mu_2 = \beta_0 + \beta_2$

$$\mu_1 = \beta_0 + \beta_2$$

$$\mu_2 = \beta_0 + \beta_2$$

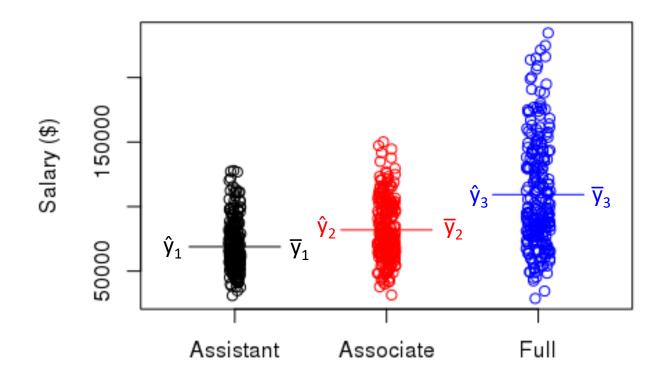
$$\mu_3 = \beta_0$$

Least squares prediction for \hat{y}_i is \overline{y}_k



$$y_i = \mu_k + \epsilon_i = \begin{cases} \mu_1 + \epsilon_i & \text{if Assistant Professor} \\ \mu_2 + \epsilon_i & \text{if Associate Professor} \\ \mu_3 + \epsilon_i & \text{if Full Professor} \end{cases}$$

Least squares prediction for \hat{y}_i is \overline{y}_k



$$\hat{y}_i = \bar{y}_k = \begin{cases} \bar{y}_1 & \text{if Assistant professor} \\ \bar{y}_2 & \text{if Associate professor} \\ \bar{y}_3 & \text{if Full} \end{cases}$$

ANOVA decomposition

$$F = \frac{\frac{1}{K-1} \sum_{i=1}^{K} n_i (\bar{y}_i - \bar{y}_{tot})^2}{\frac{1}{N-K} \sum_{i=1}^{K} \sum_{j=1}^{n_i} (y_{ij} - \bar{y}_i)^2}$$

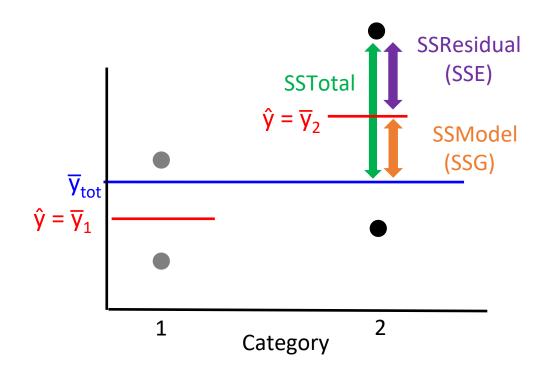
The ANOVA decomposes the variance as:

SSTotal = SSModel (SSC) + SSResidual (SSE)

$$y_{ij} - \bar{y}_{tot} = (\hat{y}_{ij} - \bar{y}_{tot}) + (y_{ij} - \hat{y}_{ij})$$

$$(y_{ij} - \bar{y}_{tot})^2 = (\hat{y}_{ij} - \bar{y}_{tot})^2 + (y_{ij} - \hat{y}_{ij})^2$$

$$(y_{ij} - \bar{y}_{tot})^2 = (\bar{y}_i - \bar{y}_{tot})^2 + (y_{ij} - \bar{y}_i)^2$$



$$\hat{y}_{ji} = \overline{y}_{i}$$

(the prediction for each class is the group mean)

Let's examine these relationships in R...

Planned comparisons/posthoc tests

Suppose we run a one-way ANOVA and we are able to reject the null hypothesis.

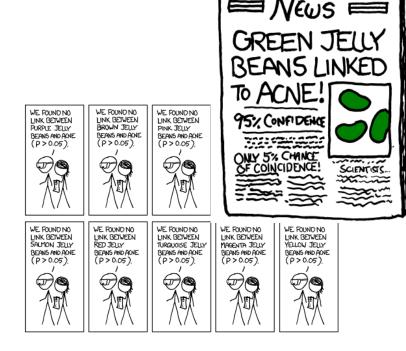
$$H_0$$
: $\mu_1 = \mu_2 = ... = \mu_k$
 H_A : $\mu_i \neq \mu_i$ for some i, j

Q: What else would we like to know?

A: We would like to know which groups actually differed!

Q: What would be a problem if we ran two sample tests on all pairs?

A: The problem of multiplicity



Pairwise comparisons

There are several tests that can be used to examine which pairs of means differed; i.e., to test:

- H_0 : $\mu_i = \mu_i$
- H_A : $\mu_i \neq \mu_j$

These tests include:

- Fisher's Least Significant Difference
- Bonferroni procedure/correction
- Tukeys Honest significantly different

Fisher's Least Significant Difference (LSD)

- 1. Perform the ANOVA
- 2. If the ANOVA F-test is not significant, stop
- 3. If the ANOVA F-test is significant, then you can test H_0 for a pairwise comparisons using:

$$t = \frac{\bar{x}_i - \bar{x}_j}{\sqrt{MSE \cdot (\frac{1}{n_i} + \frac{1}{n_j})}}$$
 Uses the MSE as a pooled estimate of the σ^2

Very 'liberal' tests

- Likely to make Type I errors (lots of false rejections of H₀)
- Less likely to make Type II errors (highest chance of detecting effects)

Bonferroni correction

Controls for the *family-wise error rate*

- i.e., $\alpha = 0.05$ for making **any** Type I error **over all pairs of comparisons**
- 1. Choose an α -level for the family-wise error rate α
- 2. Decide how many comparisons you will make. Call this m.
- 3. Reject any hypothesis tests that have p-values less than α/m
 - Pairwise tests typically done using a t-statistic, where the MSE is used in the estimate of the SE

$$t = \frac{\bar{x}_i - \bar{x}_j}{\sqrt{MSE \cdot (\frac{1}{n_i} + \frac{1}{n_j})}} \qquad \text{Use a t-distribution with n-k degrees of freedom}$$

Very 'conservative' tests

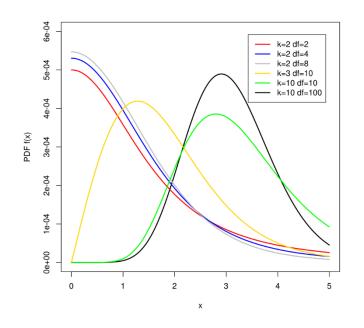
- Unlikely to make Type I errors (few false rejections of H₀)
- Likely to make Type II errors (insensitive at detecting real effects)

Tukey's Honest Significantly Different Test

Controls for the family-wise error rate

$$q = \frac{\sqrt{2}(\bar{x}_{max} - \bar{x}_{min})}{\sqrt{MSE \cdot (\frac{1}{n_{max}} + \frac{1}{n_{min}})}}$$

Where q comes from a **studentized range distribution**



The test is based on the distribution of $|\overline{x}_{max} - \overline{x}_{min}|$ that would be expected under the null hypothesis that none of the pairs of means are different

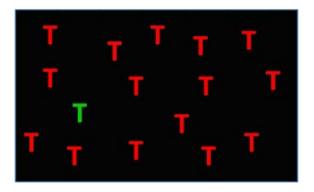
- Controls for the familywise error rate but less conservative than the Bonferroni correction
- Still based on assumptions that the data in each group is normal with equal variance

Let's try pairwise comparisons in R...

You will analyze data from a psychophysics experiment that explored popout attention

 Study done at Hampshire College by Jacob Prescott, Tapujit Debnath Tapu, Julian Oks, Kirsten Lydic

Background:

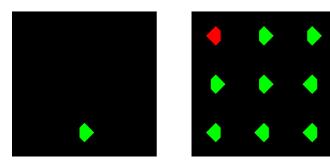


Exogenous attention



Endogenous attention

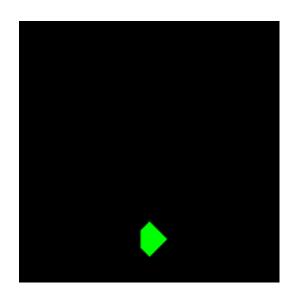
Single item Multiple items

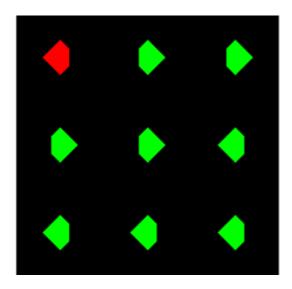


Do reaction times differ for:

- 1. Position of target stimulus
- 2. Single vs. multiple item displays

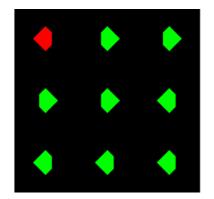
Participants engaged in a reaction time task where they needed to respond as quickly and as accurately as possible





Press "z" because left side is cut off

Press "/" because right side is cut off

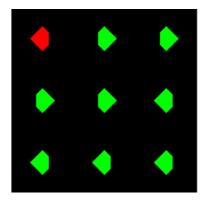


The experiment had a 9 x 2 x 2 x 2 factorial design:

- 1. Position (9 levels): 9 locations where the target stimulus could appear
- 2. Isolated/distractor condition (2 levels): isolated or cluttered display
- 3. Target color (2 levels): red or green target
 - For cluttered displays, the distractors always had the opposite color of the target
- 4. Cut direction (2 levels): left or right side of the target diamond was cut off
 - Corresponds to pressing the "z" or "/" key

The experiment had 10 blocks where all 72 (9 x 2 x 2 x 2) stimuli were shown

8 volunteer participants participated in the experiment



On homework 10 you will run:

- A one-way ANOVA to see if the mean reaction time is the same at all target positions
- A two-way ANOVA to look at how both position and isolated/cluttered displays affect mean reaction times.
- Explore another question using this data

Questions?