Simple linear regression



Overview

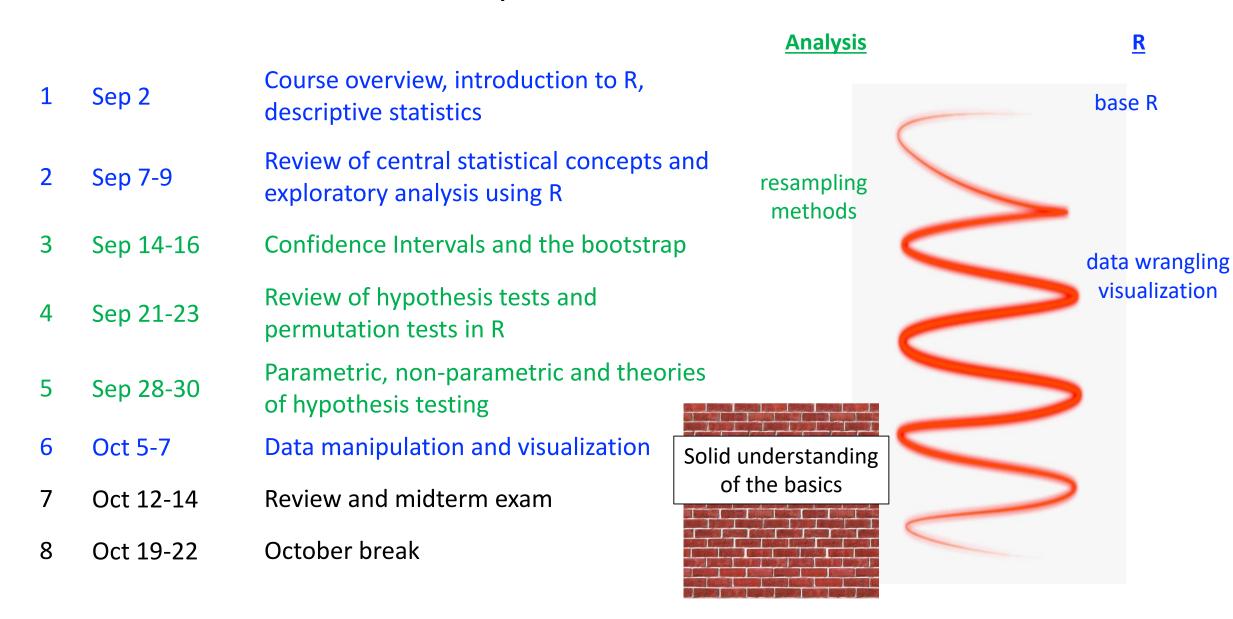
Simple linear regression

Simple linear regression in R

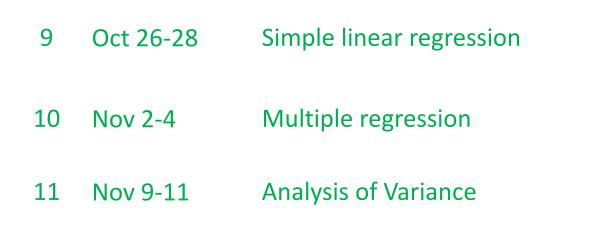
Inference for simple linear regression

- Hypothesis tests on regression coefficients
 - Hypothesis tests on regression coefficients in R
- If there is time: confidence and prediction intervals

Where we are: completed



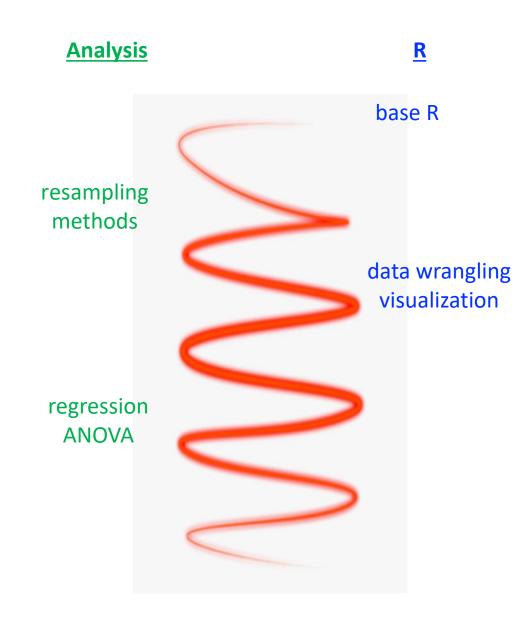
Where we are: up next



Next: building statistical models to predict the mean of a response variable y, based on explanatory variables x_i's

We will use these models to:

- 1. Make predictions for new values y
- 2. Understand which of the explanatory variables x_i 's are related to the response variable y



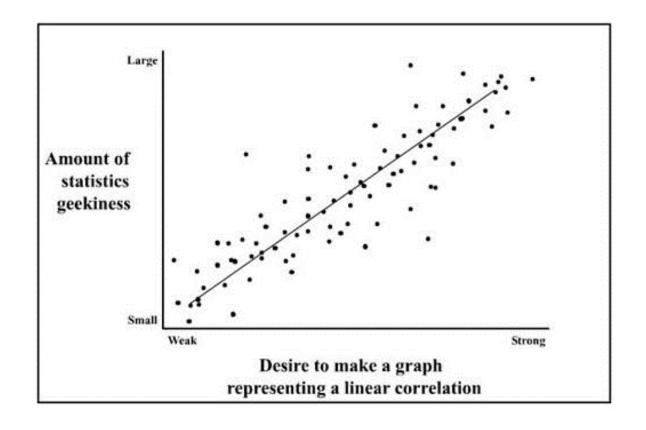
Linear regression

Regression is method of using one variable **x** to predict the value of a second variable **y**

$$\hat{y} = f(x)$$

In **linear regression** we fit a <u>line</u> to the data, called the **regression line**

• In *simple* linear regression, we use a single variable x, to predict y



Motivation: Predicting the 2020 election



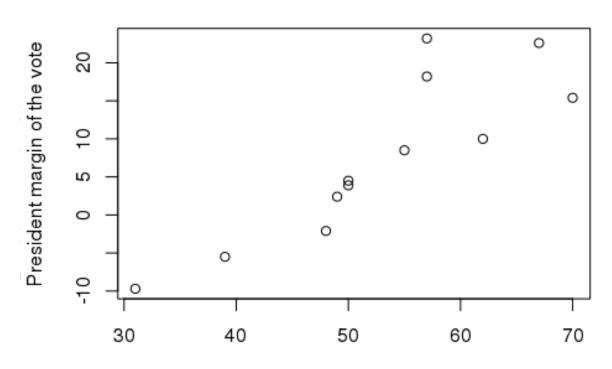
Predict the margin of the popular vote based on the president's approval rating

Data from an article on the 2012 election on the Five Thirty Eight website



Approval rating vote margin regression line

From previous 12 US president's running for reelection



President approval rating on election day

$$\hat{y} = b_0 + b_1 \cdot x$$

$$\hat{y} = \hat{\beta_0} + \hat{\beta_1} x$$

R:
$$lm(y \sim x)$$

$$\hat{\beta}_0 = -36.76$$

$$\hat{\beta}_1 = 0.84$$

$$\hat{y} = -36.76 + .84 \cdot x$$

Approval rating vote margin survey questions

1. If a president had a 0% approval rating, what percent of the vote margin does this model predict the president would get?

$$\hat{y} = b_0 + b_1 \cdot x$$

R: $lm(y \sim x)$

2. If a president's approval rating increased by 1%, how much of would the president's margin of the vote increase by?

$$b_0 = -36.76$$

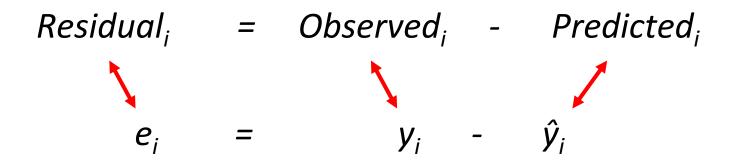
$$b_1 = 0.84$$

3. At what presidential approval level would there be an exactly even split of the vote?

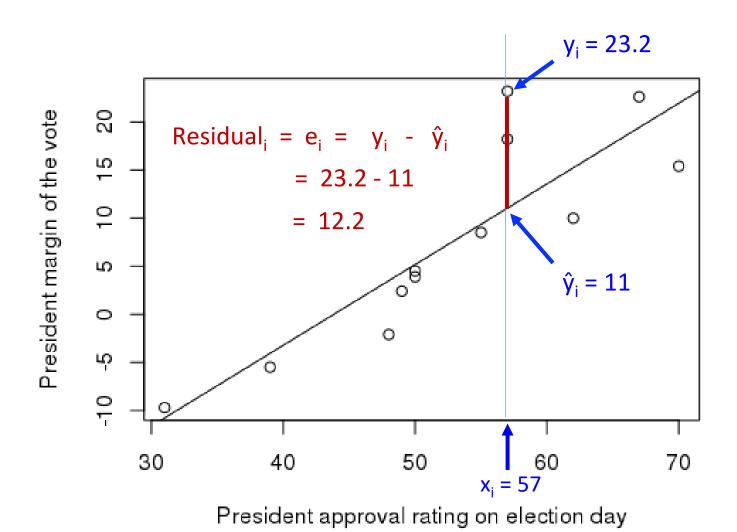
$$\hat{y} = -36.76 + .84 \cdot x$$

Residuals

The **residual** at a data value is the difference between the observed (y) and predicted value of the response variable



Approval rating vote margin regression line



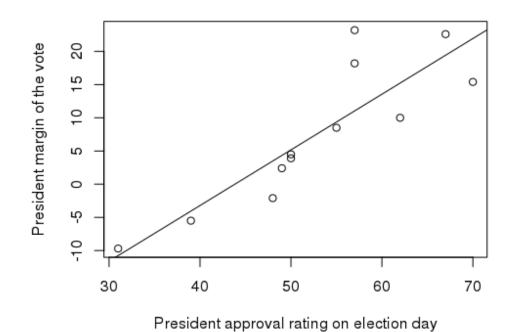
Approval rating vote margin regression line

$$\hat{y} = \hat{\beta_0} + \hat{\beta_1} x$$

Approval	Margin obs	Margin pred	Residuals
X	у	ŷ	e = y - ŷ
62	10	15.23	-5.23
50	4.5	5.17	-0.67
70	15.4	21.94	-6.54
67	22.6	19.43	3.17
57	23.2	11.04	12.16
48	-2.1	3.49	-5.59
31	-9.7	-10.76	1.06
57	18.2	11.04	7.16

Line of 'best fit'

The **least squares line**, also called 'the line of best fit', is the line which minimizes the sum of squared residuals



Try to find the line of best fit

Approval rating vote margin regression line

$$\hat{y} = \hat{\beta_0} + \hat{\beta_1} x$$

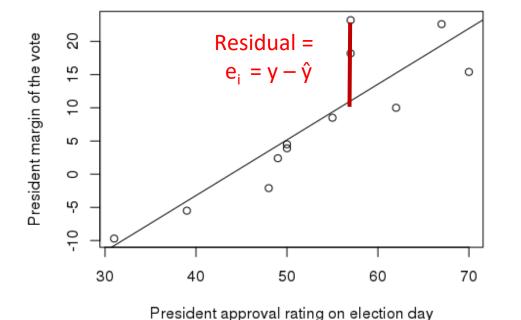
Approval	Margin obs	Margin pred	Residuals	Residuals ²
X	у	ŷ	e = y - ŷ	$e^2 = (y - \hat{y})^2$
62	10	15.23	-5.23	27.40
50	4.5	5.17	-0.67	0.45
70	15.4	21.94	-6.54	42.81
67	22.6	19.43	3.17	
57	23.2	11.04	12.16	
48	-2.1	3.49	-5.59	
31	-9.7	-10.76	1.06	
57	18.2	11.04	7.16	

Q: Why do we minimize the sum of *squared* residuals rather than just the sum of residuals?

Minimizing the sum of the squared residuals to find the regression coefficients

To find the regression coefficients \hat{eta}_0 and \hat{eta}_0 we minimize the sum of squared residuals

- We will use the notation **SSResidual (SSRes)** to denote the some of squared residuals
 - (The residual sum of squares is also called the error sum of squares (SSE))



$$residual = e_{i} = y_{i} - \hat{y}_{i}$$

$$SSRes = \sum_{i=1}^{n} e_{i}^{2} = \sum_{i=1}^{n} (y_{i} - \hat{y}_{i})^{2}$$

$$= \sum_{i=1}^{n} (y_{i} - \hat{f}(x_{i}))^{2} = \sum_{i=1}^{n} (y_{i} - (\hat{\beta}_{0} + \hat{\beta}_{1}x_{i}))^{2}$$

R: $lm(y \sim x)$

How do we minimize the SSE?

$$SSRes = \sum_{i=1}^{n} (y_i - (\hat{\beta}_0 + \hat{\beta}_1 x))^2$$

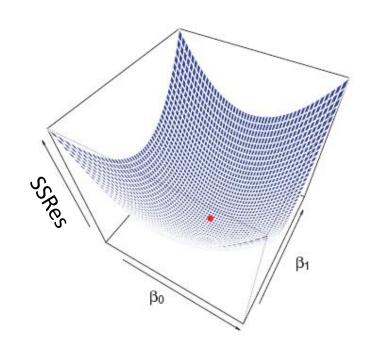
How do we find $\hat{\beta_0}, \hat{\beta_1}$?



- Take the derivative, set to 0 and solve
- This mathematical convenience is why the squared loss is so commonly used

$$\hat{\beta}_0 = \bar{y} - \hat{\beta}_1 \bar{x}$$

$$\hat{\beta}_1 = \frac{\sum_{i=1}^n (x_i - \bar{x})(y_i - \bar{y})}{\sum_{i=1}^n (x_i - \bar{x})^2}$$



Basic regression caution # 1

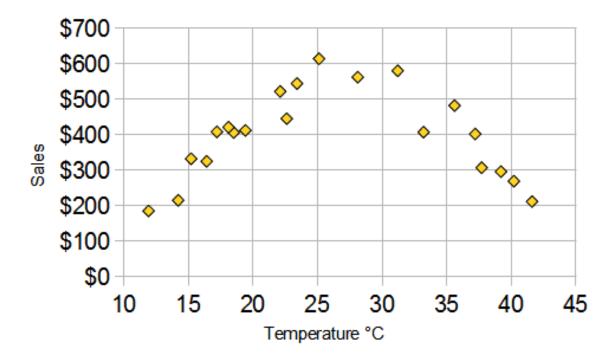
Avoid trying to apply the regression line to predict values far from those that were used to create the line.

• i.e., do not extrapolate too far

Basic regression caution # 2

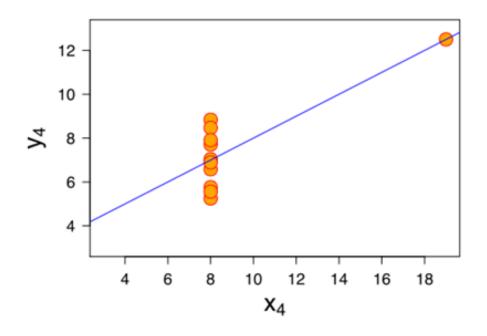
Plot the data! Linear regression is only appropriate when there is a linear trend in the data.

• We will discuss a set of checks on the appropriateness of using linear models soon



Basic regression caution #3

Be aware of outliers and high leverage points. They can have a large effect on the regression line.



Outlier: big $| y - \overline{y} |$

Leverage: big $|x - \overline{x}|$

Influential point: big outlier and leverage

There are statistics that quantify/describe influential points

We will discuss these soon as well

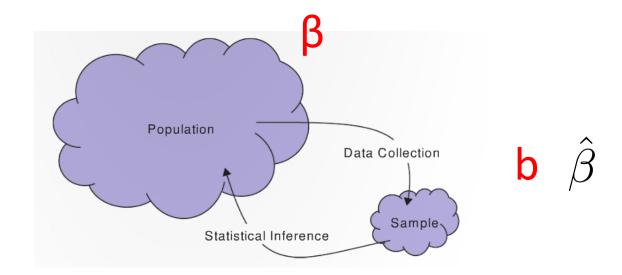
Let's try simple linear regression in R...

Inference for simple linear regression

Inference for simple linear regression

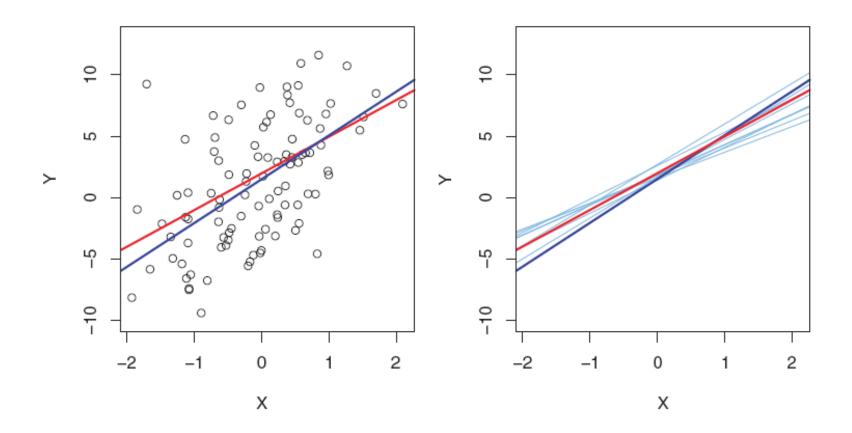
The letter **b** or $\hat{\beta}$ is typically used to denote the slope **of the sample**

The Greek letter β is used to denote the slope of the population



Population: β

Sample estimates: b \hat{eta}



Linear regression underlying model

Intercept Slope } Parameters

Error

True regression line: $\mu_Y(x) = \beta_0 + \beta_1 x$

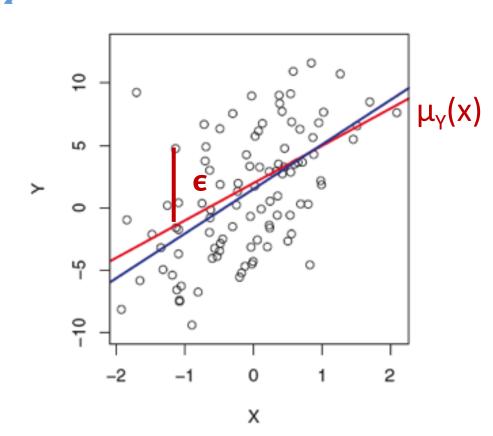
Observed data point:

$$Y = \beta_0 + \beta_1 x + \epsilon$$

$$=\mu_Y(x)+\epsilon$$

Errors ϵ_i are the difference between the **true** regression line $\mu_Y(x_i)$ and observed data points Y_i

•
$$\epsilon_i = Y_i - \mu_V(x_i)$$



Linear regression underlying model

Intercept Slope } Parameters

Error

True regression line: $\mu_Y(x) = \beta_0 + \beta_1 x$

Observed data point: $Y = \beta_0 + \beta_1 x + \epsilon$

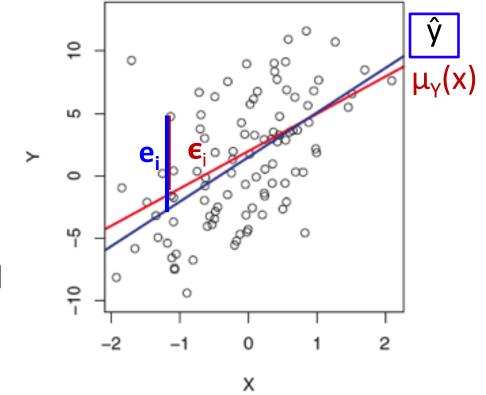
Estimated regression line: $\hat{y} = \hat{\beta_0} + \hat{\beta_1} x$

Errors ϵ_i are the difference between the **true** regression line $\mu_{Y}(x_i)$ and observed data points Y_i

•
$$\epsilon_i = Y_i - \mu_Y(x_i)$$

Residuals e_i are the difference between the **estimated** regression line \hat{y}_i and observed data points Y_i

•
$$\mathbf{e}_{i} = Y_{i} - \hat{y}_{i}$$



Linear regression underlying model

Intercept Slope } **Parameters** True regression line: $\mu_Y(x) = \beta_0 + \beta_1 x$ **Error**

Observed data point:
$$Y = \beta_0 + \beta_1 x + \epsilon$$

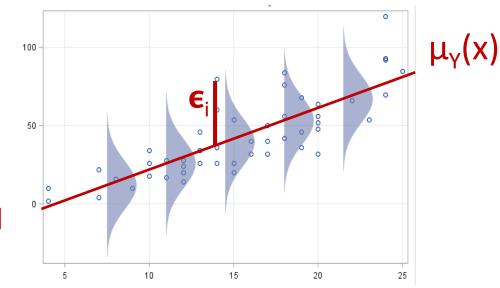
 $\epsilon \sim N(0, \sigma_{\epsilon})$

Errors ϵ_i are the difference between the **true** regression line $\mu_{v}(x_{i})$ and observed data points Y_{i}

•
$$\epsilon_i = Y_i - \mu_Y(x_i)$$

We will assume that the errors ϵ_i are normally distributed

- This is needed for inference using parametric methods
 - e.g., to use t-distributions and F-distributions



Recap: Errors vs. residuals

The data:

$$Y_i = \beta_0 + \beta_1 x_i + \epsilon_i$$

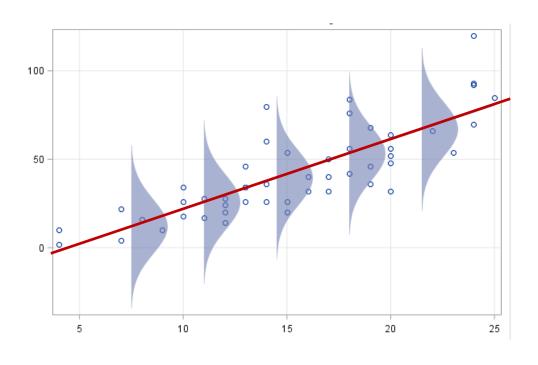
$$\epsilon_i \sim N(0, \sigma_\epsilon)$$

"True" model: $\mu_Y(x_i) = \beta_0 + \beta_1 x_i$

• Errors: $\epsilon_i = Y_i - \mu_Y(x_i)$

Estimated model: $\hat{y}_i = \hat{\beta}_0 + \hat{\beta}_1 x_i$

• Residuals: $e_i = Y_i - \hat{y}_i$



Standard deviation of the errors: σ_{ε}

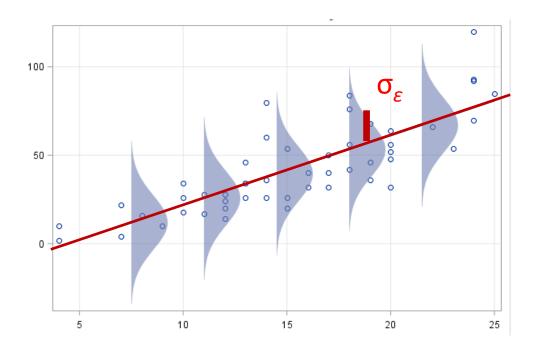
The standard deviation of the errors is denoted $\sigma_{arepsilon}$

We can use the **standard deviation of residuals** as an estimate standard deviation of the errors σ_{ε} .

• σ_{ε} called the "regression standard error"

$$\hat{\sigma}_{\epsilon} = \sqrt{\frac{1}{n-2} SSRes}$$

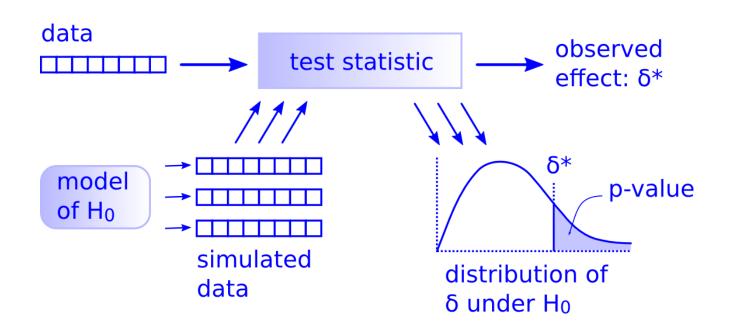
$$= \sqrt{\frac{1}{n-2} \sum_{i=1}^{n} (y_i - \hat{y}_i)^2}$$

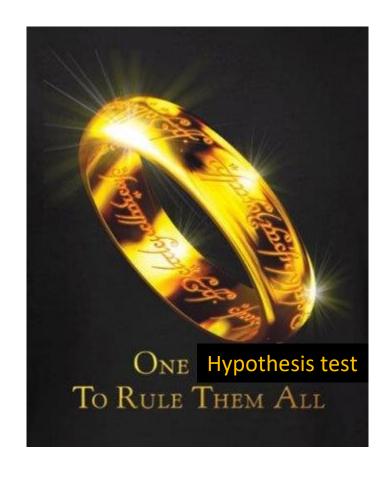


Inference for linear regression: hypothesis tests

Hypothesis test for regression coefficients

There is only one <u>hypothesis test!</u>





Hypothesis test for regression coefficients

We can run hypothesis tests to assess whether there is a relationship between y and x, and calculate p-values

- H_0 : $\beta_1 = 0$ (slope is 0, so no relationship between x and y
- H_A : $\beta_1 \neq 0$

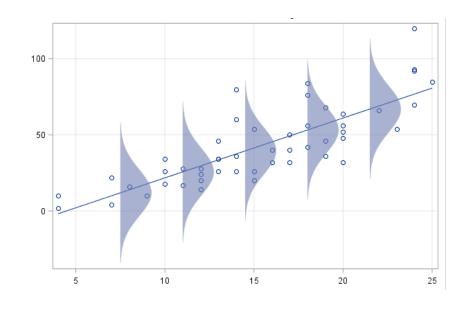
One type of hypothesis test we can run is based on a t-statistic: $t=\frac{\beta_1-0}{\hat{SE}_{\hat{\beta_1}}}$ • The t-statistic comes from a t-distribution with n - 2 degrees of freedom

$$\hat{SE}_{\hat{\beta}_{1}} = \frac{\hat{\sigma}_{\epsilon}}{\sqrt{\sum_{i=1}^{n} (x_{i} - \bar{x})^{2}}} \qquad \hat{SE}_{\hat{\beta}_{0}} = \hat{\sigma}_{\epsilon} \sqrt{\frac{1}{n} + \frac{\bar{x}^{2}}{\sum_{i=1}^{n} (x_{i} - \bar{x})^{2}}}$$

Inference using parametric methods

When using parametric methods, we make the following (LINE) assumptions:

- Linearity: A line can describe the relationship between x and y
- Independence: each data point is independent from the other points
- Normality: errors are normally distributed around the true regression line $\mu_v(x)$
- Equal variance (Homoscedasticity): constant variance of errors over the whole range of x values



These assumptions are usually checked after the models are fit using 'regression diagnostic' plots.

Let's look at inference for simple linear regression in R

Back to faculty salaries...



Inference for linear regression: confidence intervals

We can estimate three types of intervals for a regression:

- 1. Confidence intervals for the regression coefficients: eta_0 and eta_1
- 2. Confidence intervals for the full line μ_v

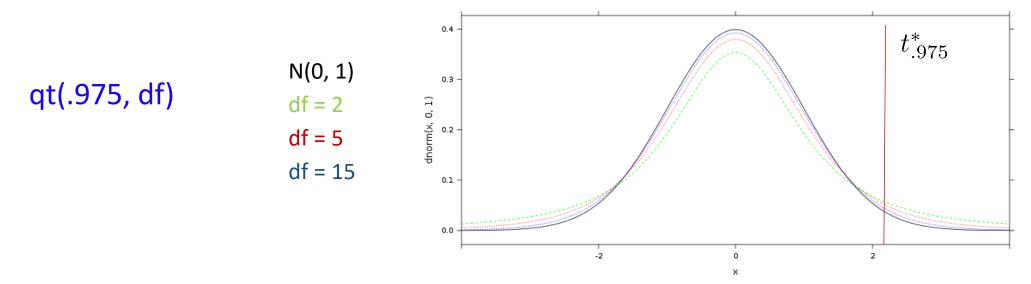
3. Prediction intervals where most of the data is expected

Confidence intervals for regression coefficients

For the slope coefficient , the confidence interval is: $\hat{eta}_1 \pm t^* \cdot SE_{\hat{eta}_1}$

Where:
$$SE_{\hat{\beta_1}} = \frac{\sigma_{\epsilon}}{\sqrt{\sum_{i=1}^{n}(x_i - \bar{x})^2}}$$

 t^* is the critical value for the t_{n-2} density curve needed to obtain a desired confidence level



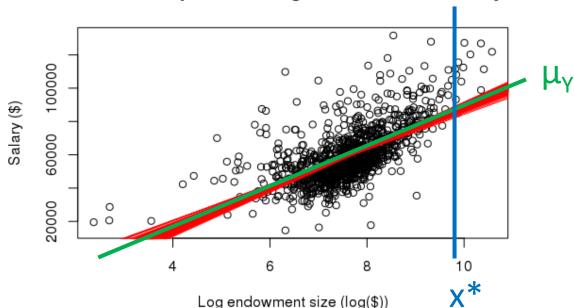
Confidence intervals for the regression line μ_{γ}

A confidence interval for the mean response for the **true regression line** μ_{γ} when $X = x^*$ is:

$$\hat{y} \pm t^* \cdot SE_{\hat{\mu}}$$
 where

$$SE_{\hat{\mu}} = \sigma_{\epsilon} \sqrt{\frac{1}{n} + \frac{(x^* - \bar{x})^2}{\sum_{i=1}^{n} (x_i - \bar{x})^2}}$$

Relationship between log endowment and salary



Note:

- There is more uncertainty at the ends of the regression line
- The confidence interval for the regression line μ_{γ} is different than the confidence interval for slope β_1

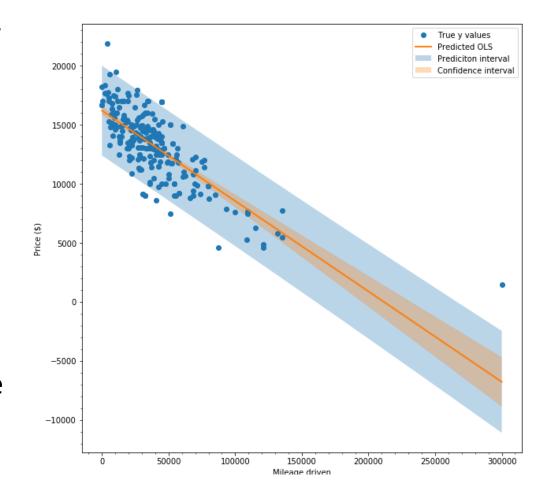
Prediction intervals

Confidence intervals give us a measure of uncertain about our the true relationship between x and y for:

- The true regression slope β_1
- The true regression line μ_{Y}

Prediction intervals give us a range of plausible values for y

• i.e., 95% of our y's with be within this range



Prediction intervals

A **prediction intervals** for the y can be calculated using:

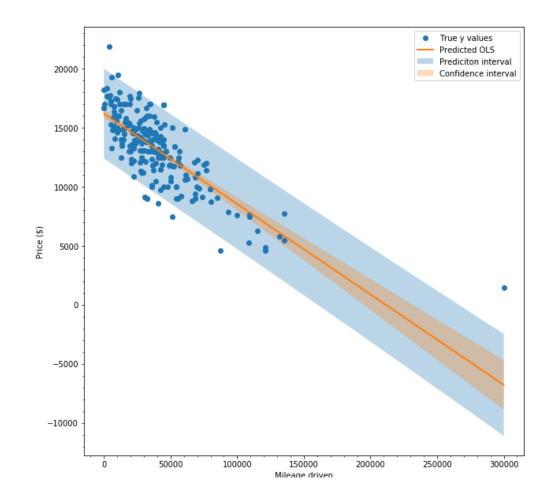
$$\hat{y} \pm t^* \cdot SE_{\hat{y}}$$

where

$$SE_{\hat{y}} = \sigma_{\epsilon} \sqrt{1 + \frac{1}{n} + \frac{(x^* - \bar{x})^2}{\sum_{i=1}^{n} (x_i - \bar{x})^2}}$$

Due to y's scattering around the true regression line

Due to uncertainty in where the true regression line is



Summary of confidence and prediction intervals

1. CI for Slope β
$$\hat{\beta}_1 \pm t^* \cdot SE_{\hat{\beta}_1}$$
 $SE_{\hat{\beta}_1} = \sigma_\epsilon \sqrt{\frac{1}{\sum_{i=1}^n (x_i - \bar{x})^2}}$

2. CI for regression line μ_v at point x^*

$$\hat{y} \pm t^* \cdot SE_{\hat{\mu}}$$
 $SE_{\hat{\mu}} = \sigma_{\epsilon} \sqrt{\frac{1}{n} + \frac{(x^* - \bar{x})^2}{\sum_{i=1}^n (x_i - \bar{x})^2}}$

3. Prediction interval y

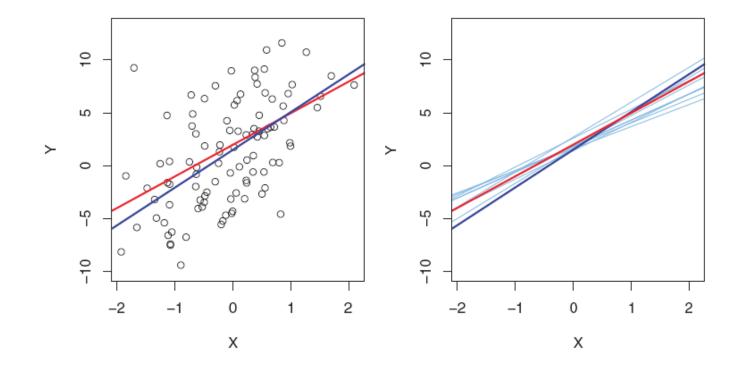
$$\hat{y} \pm t^* \cdot SE_{\hat{y}}$$
 $SE_{\hat{y}} = \sigma_{\epsilon} \sqrt{1 + \frac{1}{n} + \frac{(x^* - \bar{x})^2}{\sum_{i=1}^n (x_i - \bar{x})^2}}$

Resampling methods for inference in regression

We can also use resampling methods to estimate run hypothesis tests and create confidence intervals for the regression coefficients

Bootstrap

Permutation test



Let's look at inference for simple linear regression in R

More faculty salary data

