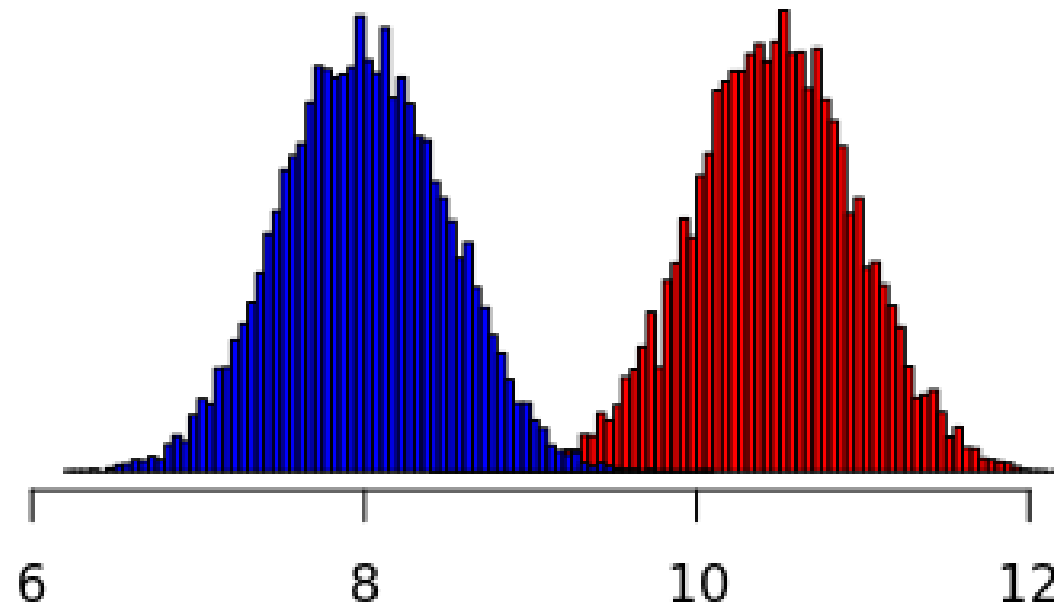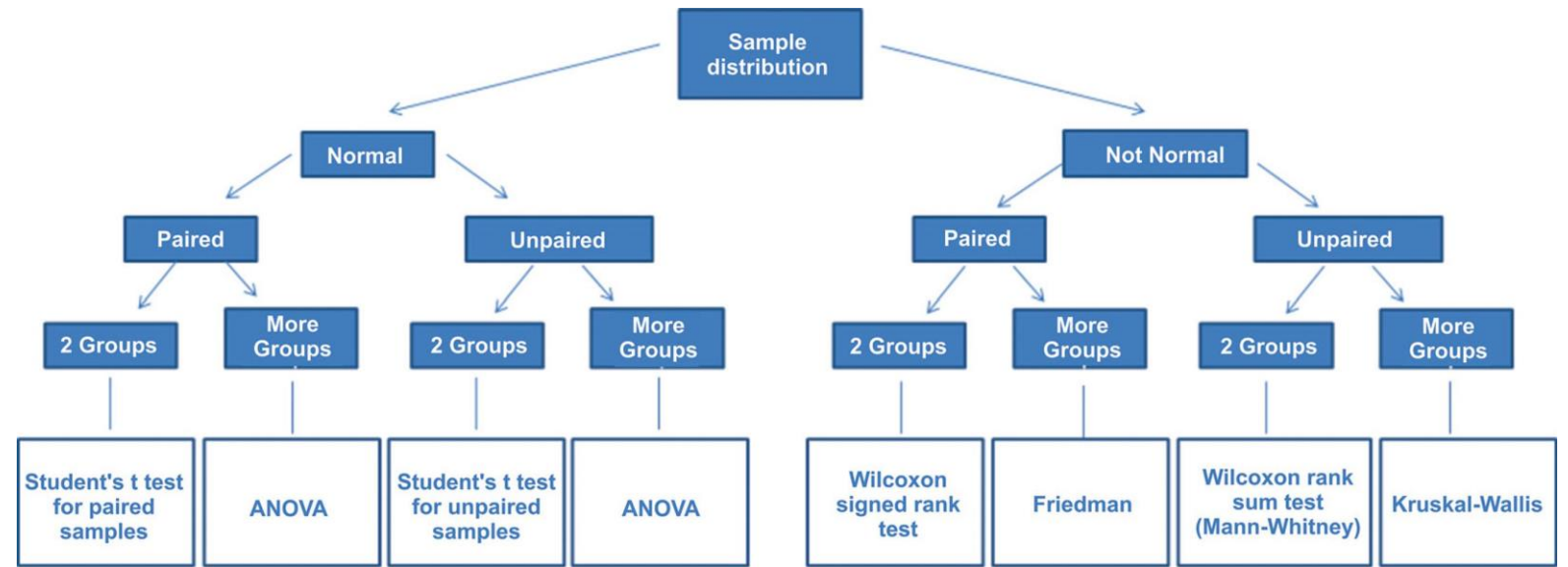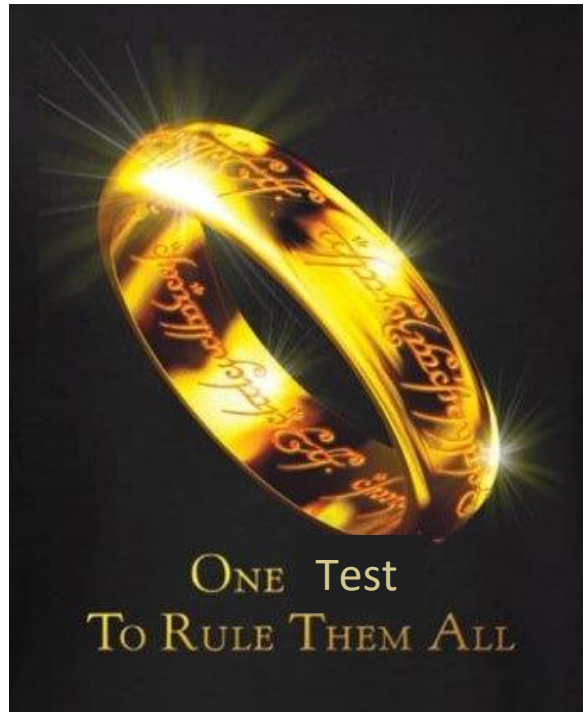# Parametric hypothesis tests continued

Overview

Quick review and extensions of parametric and nonparametric tests
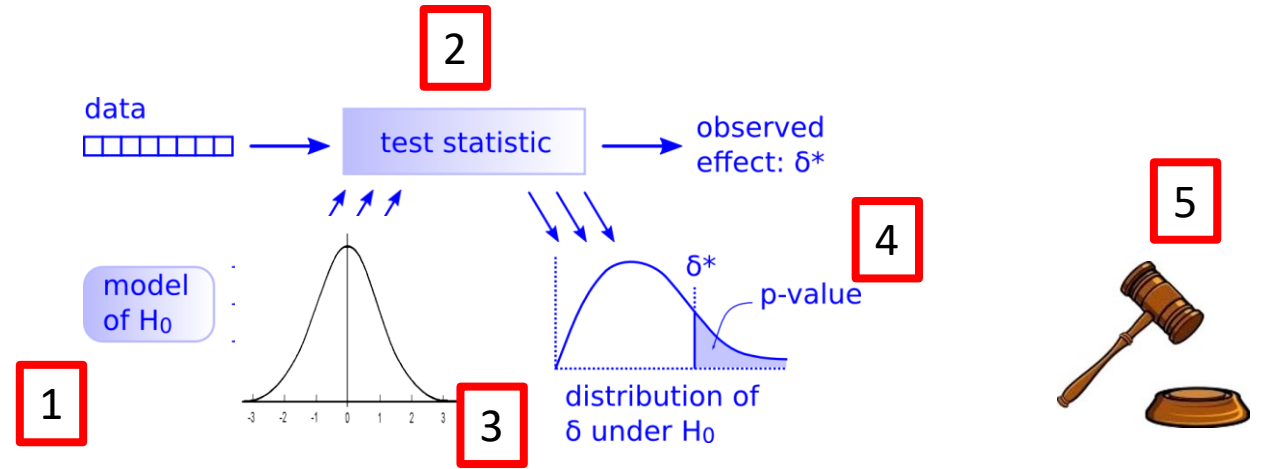
Theories of hypothesis testing

If there is time: hypothesis tests for a single mean and connections between hypothesis tests and confidence intervals

Randomization and parametric hypothesis tests

The big picture: There is only one hypothesis test!


ONE Test TO RULE THEM ALL

Just need to follow 5 steps!

Randomization and parametric hypothesis tests

The difference between randomization/permutation tests and parametric hypothesis tests is in how the null distribution is created (step 3):

- Randomization/permutation tests the null distribution is created through computational simulations

- In parametric tests, the null distribution is created using a parametric probability distribution
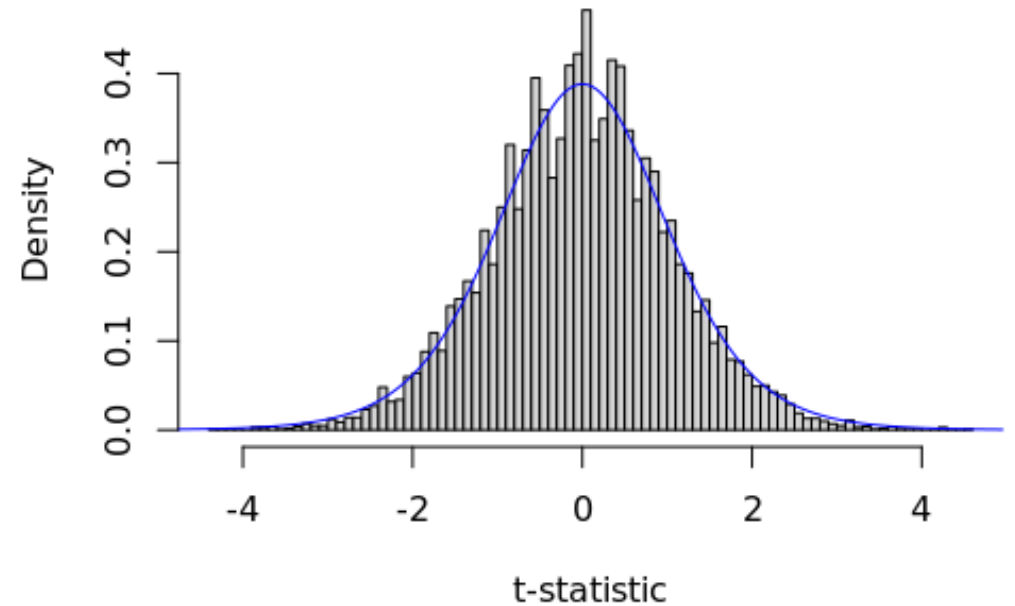
The big picture: There is only one hypothesis test!
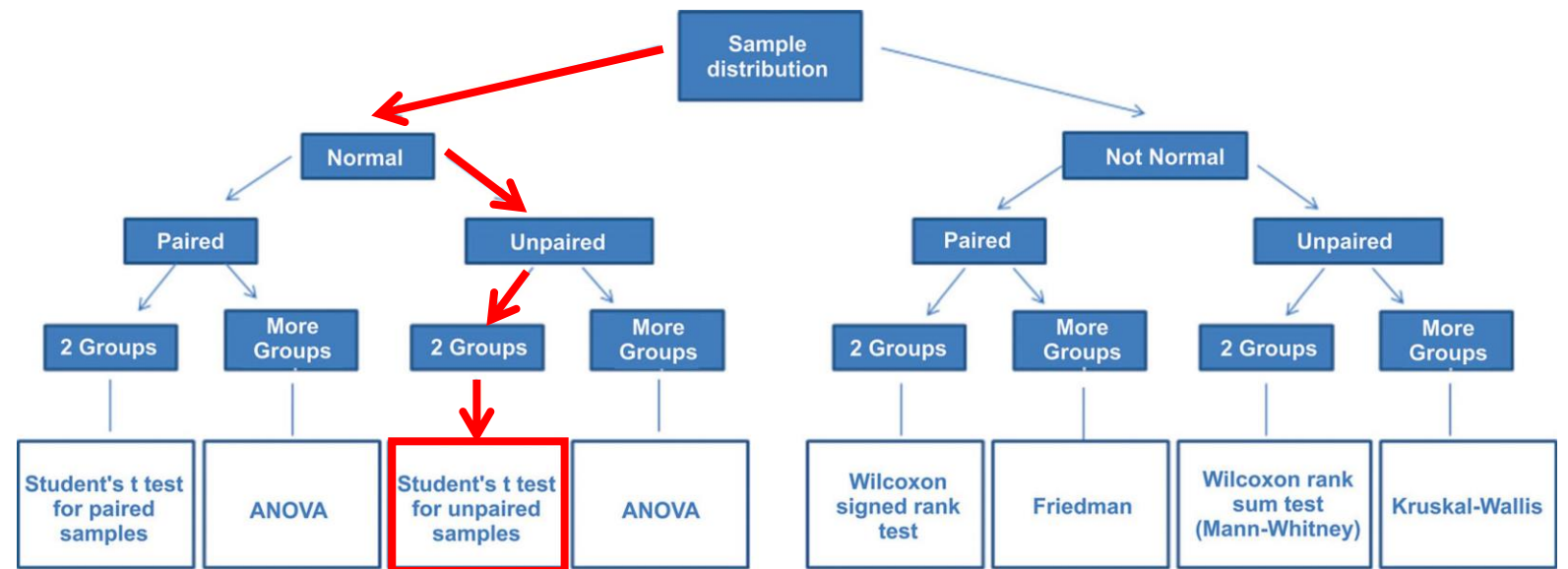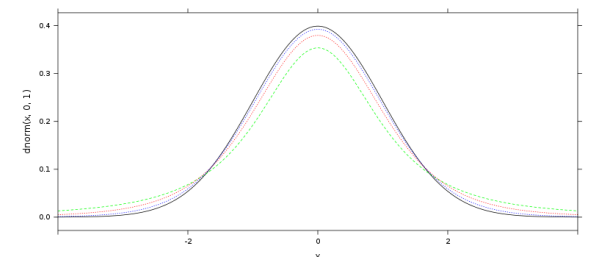

The hypothesis test zoo

ONE Test
TO RULE THEM ALL
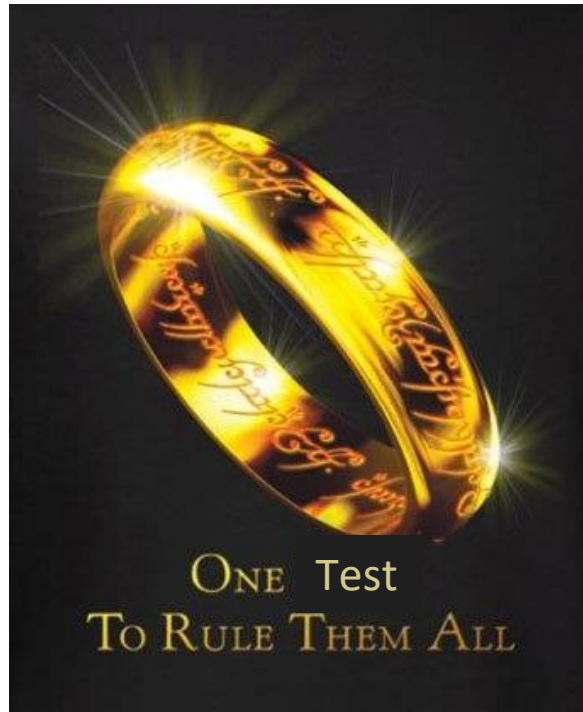
We can run a large number of additional hypothesis tests by following the 5 steps!
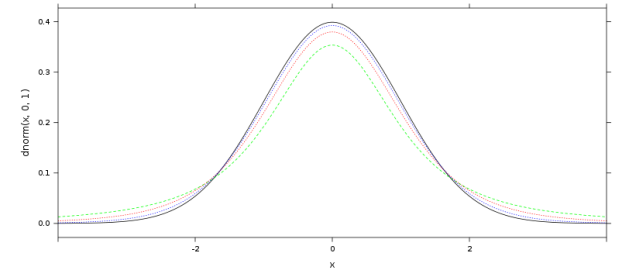
$$t = \frac{\bar{x}_t - \bar{x}_t}{s_p \cdot \sqrt{\frac{1}{n_t} + \frac{1}{n_c}}}$$

t-tests for comparing two means

Students' t-test assumes the variance in each population is the same, and uses an SE estimate of:

$$\hat{SE}_{\bar{x}_t - \bar{x}_c} = s_p \cdot \sqrt{\frac{1}{n_t} + \frac{1}{n_c}}$$

$$s_p = \sqrt{\frac{\sum_i^{n_t}(x_i - \bar{x}_c)^2 + \sum_j^{n_c}(x_j - \bar{x}_c)^2}{n_t + n_c - 2}}$$

$$t = \frac{\bar{x}_t - \bar{x}_t}{s_p \cdot \sqrt{\frac{1}{n_t} + \frac{1}{n_c}}}$$

Welch's t-test does **not** assume that the variance in each population is the same and uses an estimate of:

$$\hat{SE}_{\bar{x}_t - \bar{x}_c} = \sqrt{\frac{s_t^2}{n_t} + \frac{s_c^2}{n_c}}$$

$$t = \frac{\bar{x}_t - \bar{x}_c}{\sqrt{\frac{s_t^2}{n_t} + \frac{s_c^2}{n_c}}}$$

Since we have SE estimates, we can compute confidence intervals:
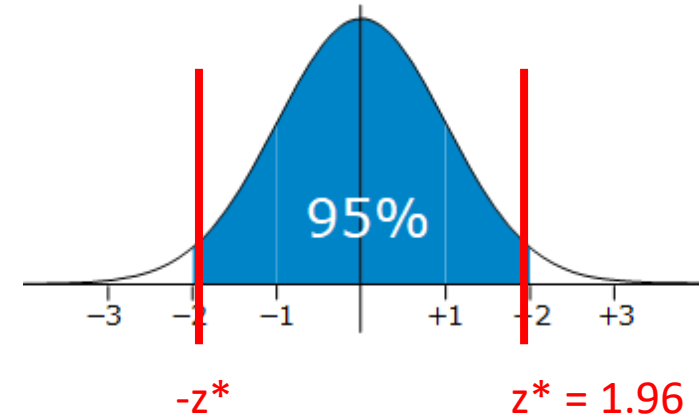
$$CI_{95} \approx stat \pm 2 \cdot SE$$

Confidence interval for the difference of two means

Confidence intervals for the bootstrap had the form:

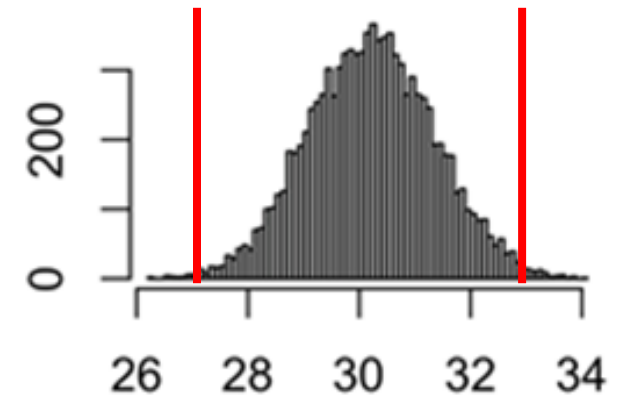$$CI_{95} \approx stat \pm 2 \cdot SE^*$$

qnorm(.975) = 1.96

-z*   z* = 1.96

Side note: one can also calculate 95% bootstrap confidence intervals using:

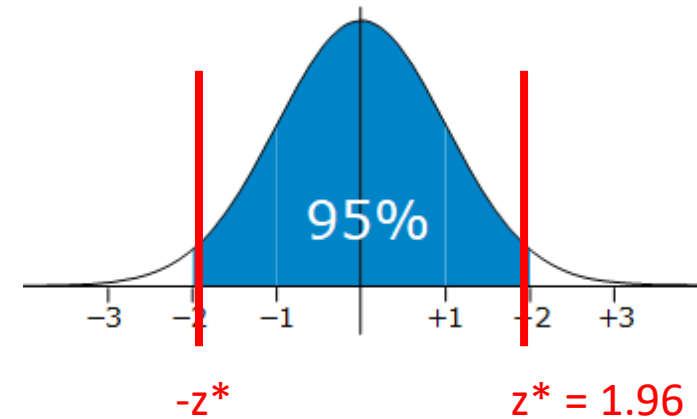quantile(bootstrap_distribution, c(.025, .975))

Confidence interval for the difference of two means

Confidence intervals for the bootstrap had the form:

$$CI_{95} \approx \text{stat} \pm 2 \cdot SE^*$$

qnorm(.975) = 1.96

95%

-z*   z* = 1.96
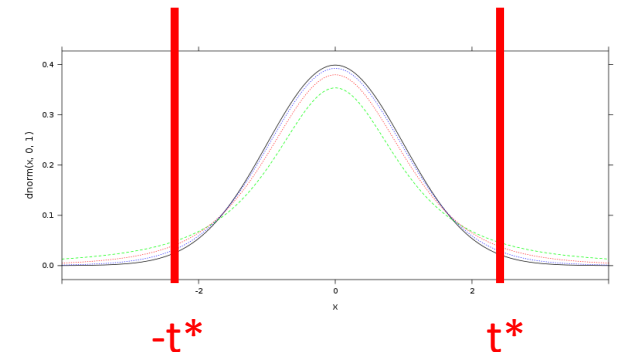
When creating confidence intervals using t-statistics we use:

$$CI_{95} \approx \text{statistic} \pm t^* \cdot \hat{SE}$$

df = min(n_t, n_c) - 1

qt(.975, df)

$$\hat{SE}_{\bar{x}_t - \bar{x}_c} = \sqrt{\frac{s_t^2}{n_t} + \frac{s_c^2}{n_c}}$$

-t*   t*

For a difference of means: $CI = (\bar{x}_t - \bar{x}_c) \pm t^* \cdot \sqrt{\frac{s_t^2}{n_t} + \frac{s_c^2}{n_c}}$

The big picture: There is only one hypothesis test!
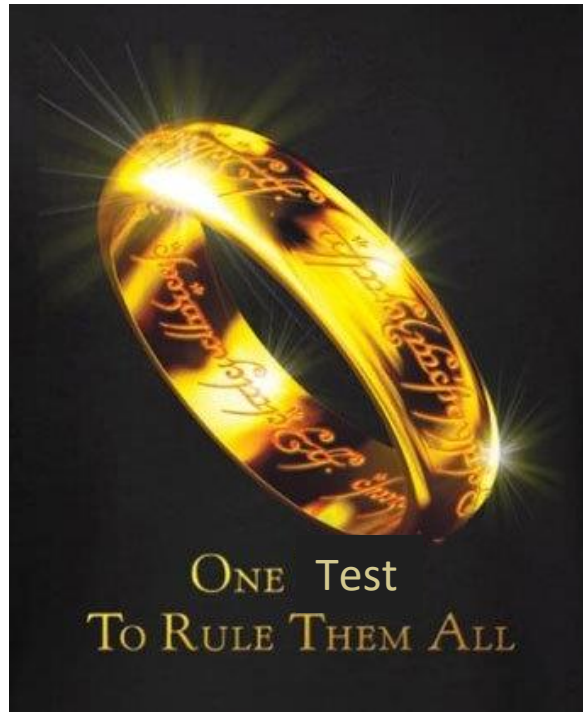
The hypothesis test zoo

ONE Test
TO RULE THEM ALL

We can run a large number of additional hypothesis tests by following the 5 steps!
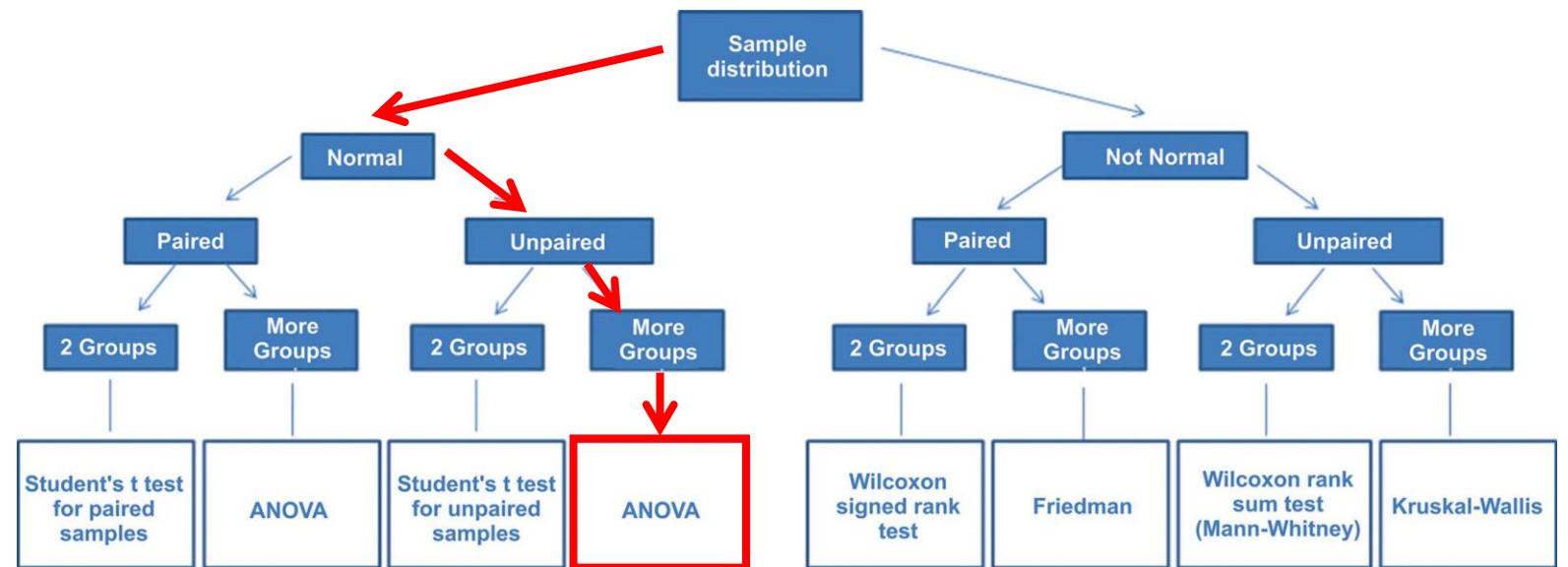
ANOVA: $H_0: \mu_1 = \mu_2 = \ldots = \mu_k$

$$F = \frac{\frac{1}{K-1}\sum_{i=1}^{K} n_i(\bar{x}_i - \bar{x}_{tot})^2}{\frac{1}{N-K}\sum_{i=1}^{K}\sum_{j=1}^{n_i}(x_{ij} - \bar{x}_i)^2}$$
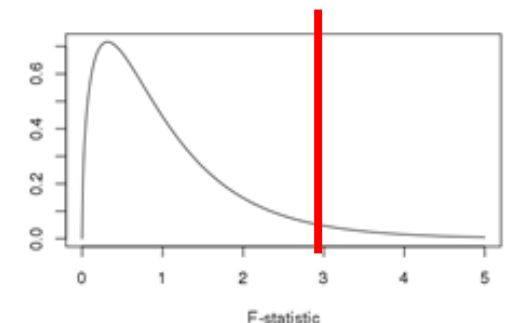
The big picture: There is only one hypothesis test!

The hypothesis test zoo

We can run a large number of additional hypothesis tests by following the 5 steps!

Nonparametric hypothesis tests

Brief mention: nonparametric hypothesis tests

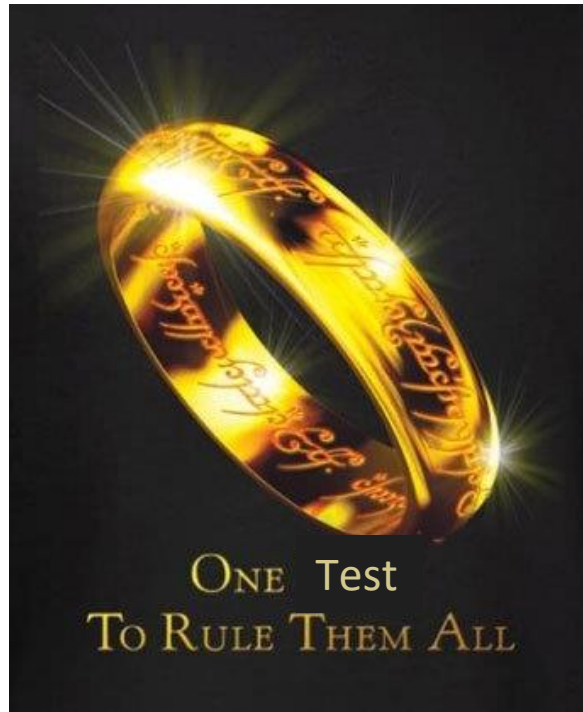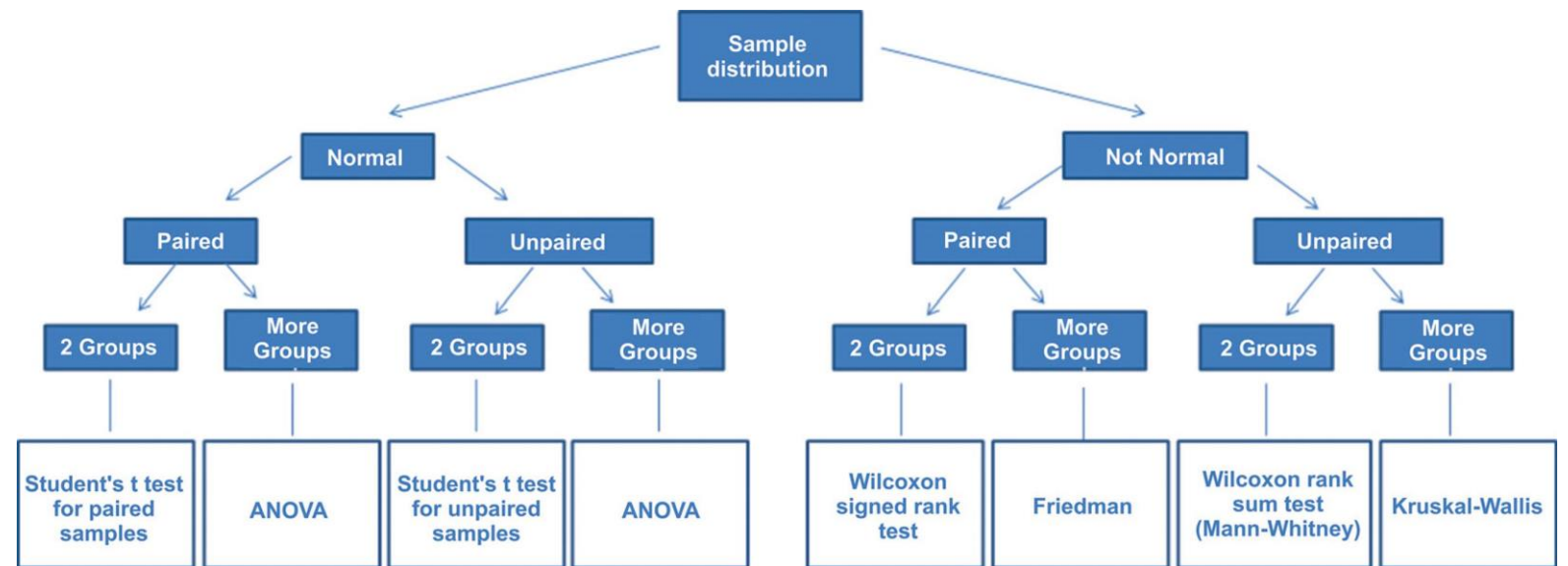Nonparametric hypothesis tests use null distributions that do not have a small fixed set of parameters
  - The null distribution changes/grows with the size of the data

Often nonparametric tests are based on converting the data to ranks
  - E.g., Mann-Whitney U test/Wilcoxon rank-sum test
    - Tests whether the probability of X being greater than Y is equal to the probability of Y being greater than X.
      - (where X and Y come from two populations)

Nonparametric tests have fewer assumptions than parametric tests so they are potentially more robust
  - e.g., they do not assume the data comes from a normal distribution, they are resistance to outliers, etc.

Mann-Whitney U Test

Is there a difference in the rank sum?
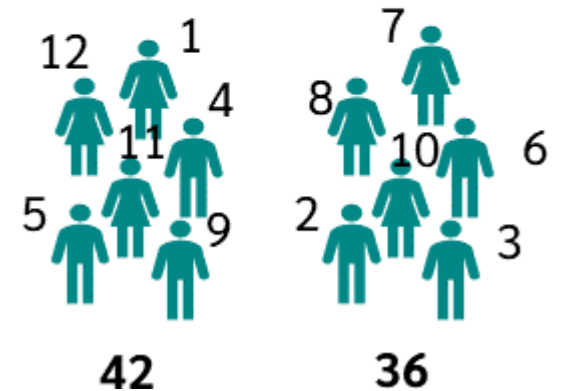
Questions

Question: When running a hypothesis test, is it better to…

1. Report the actual p-value

2. Just report if we reject/fail to reject the null hypothesis at the = 0.05 level?

Theories of hypothesis tests
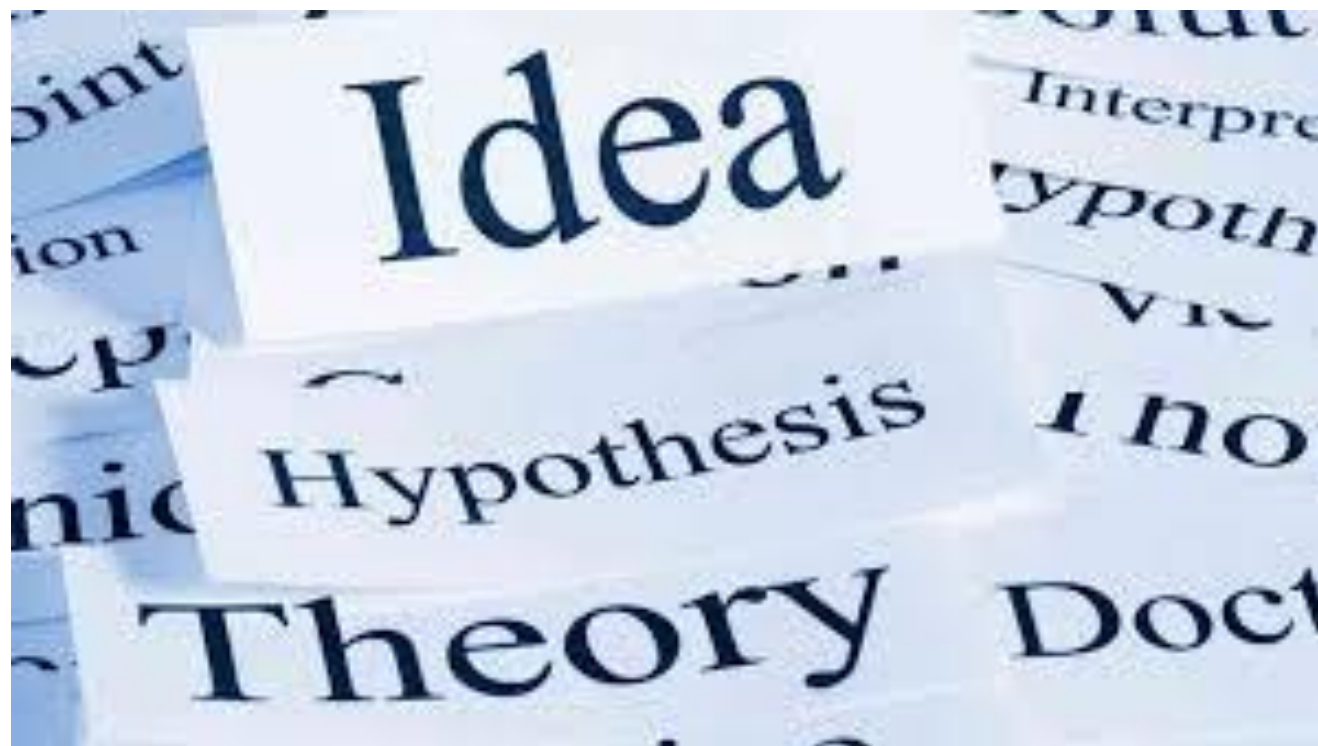
Two theories of hypothesis testing

Null-hypothesis significance testing (NHST) is a hybrid of two theories:

1. Significance testing of Ronald Fisher

2. Hypothesis testing of Jezy Neyman and Egon Pearson

Fisher (1890-1962)   Neyman (1894-1981)   Pearson (1895-1980)

Ronald Fisher's significance testing

Views the p-value as strength of evidence against the null hypothesis

- p-values part of an on-going scientific process:

  They tell the experimenter "what results to ignore"
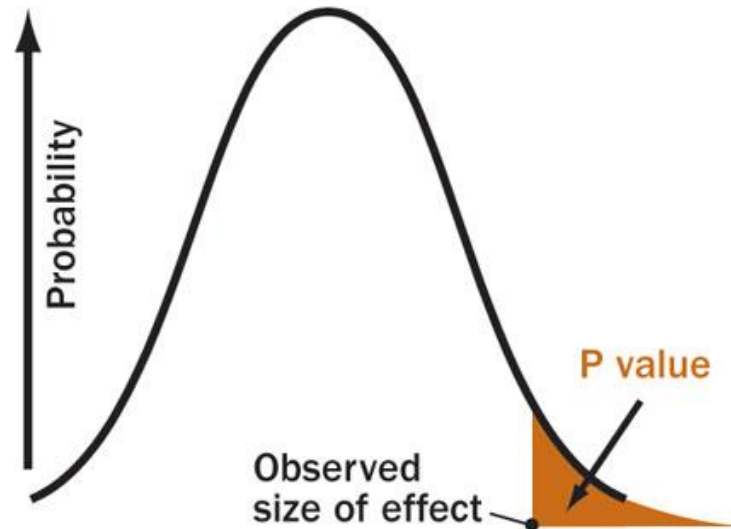
Neyman-Pearson null hypothesis testing

Makes **a formal decision** in statistical tests

Reject H$_0$: if the observed sample statistic is beyond a fixed value

- i.e., reject H$_0$ if the p-value is less than some predetermined **significance level** α

Do not reject H$_0$: if the observed sample statistic is not beyond a fixed value. This means the test is inconclusive.

Null distribution

α
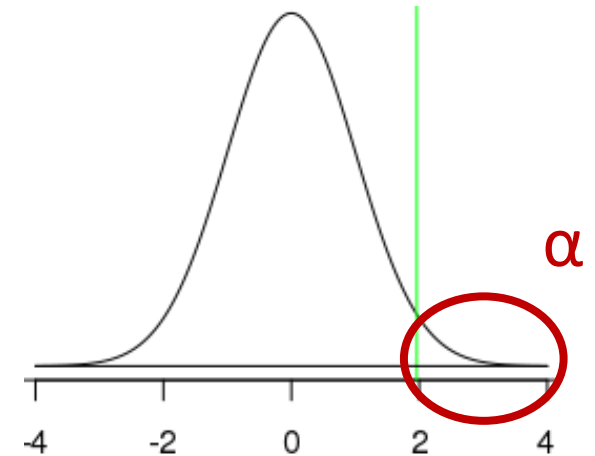
Neyman-Pearson frequentist logic

Type I error: incorrectly rejecting the null hypothesis when it is true

If Neyman-Pearson null hypothesis testing paradigm was followed perfectly, then only ~5% of the time would we falsely report an effect when null hypothesis was actually true (for $\alpha = 0.05$)

- i.e., we would only make type I errors 5% of the time
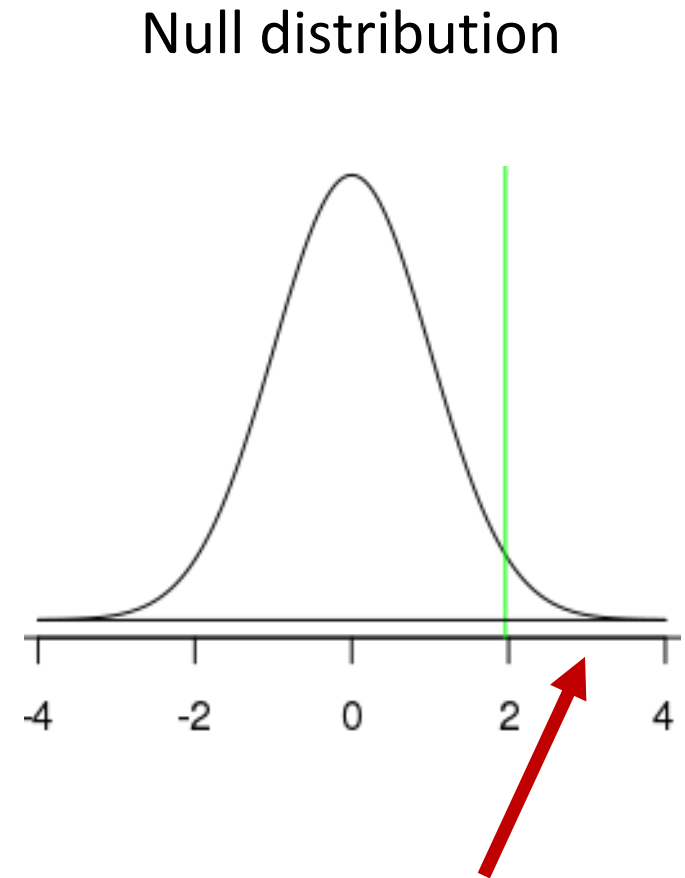
Null distribution

The null distribution is true but statistic landed here

Neyman-Pearson Frequentist logic
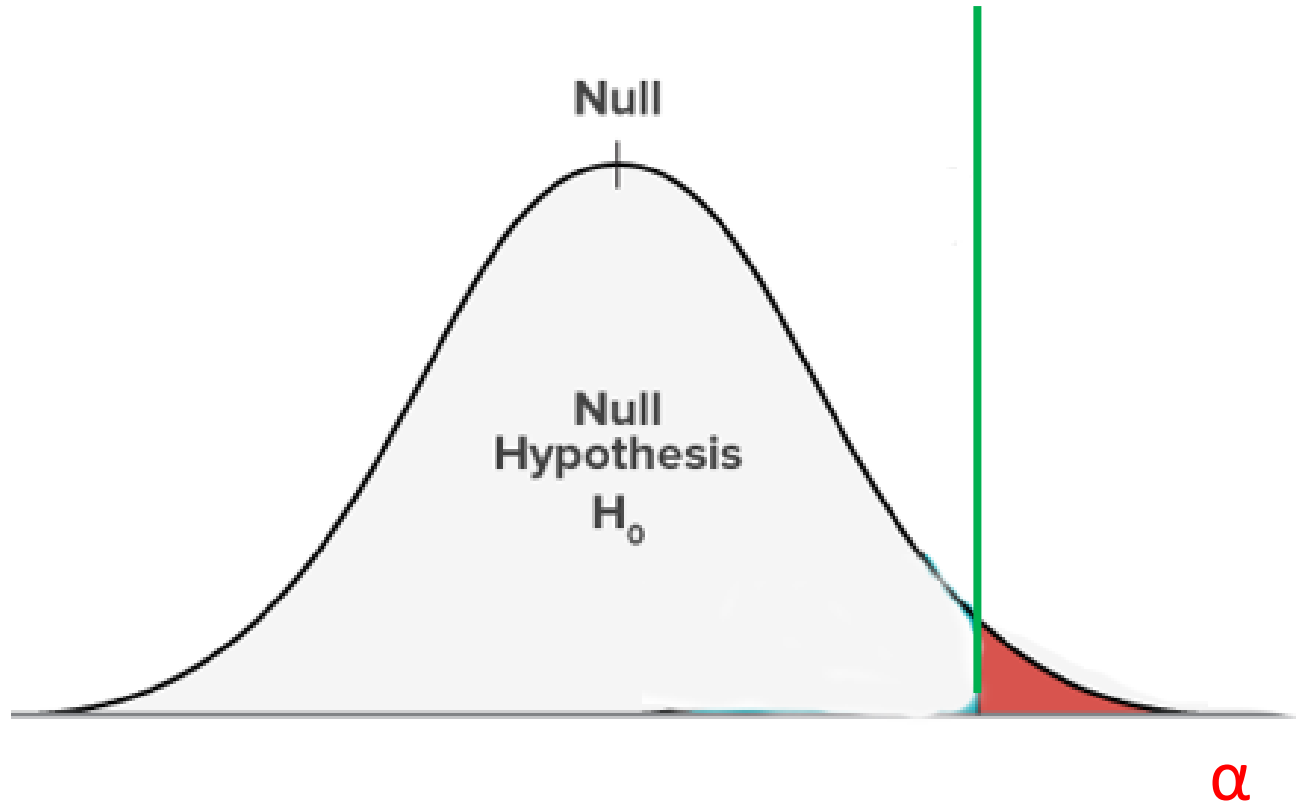
Neyman-Pearson Frequentist logic
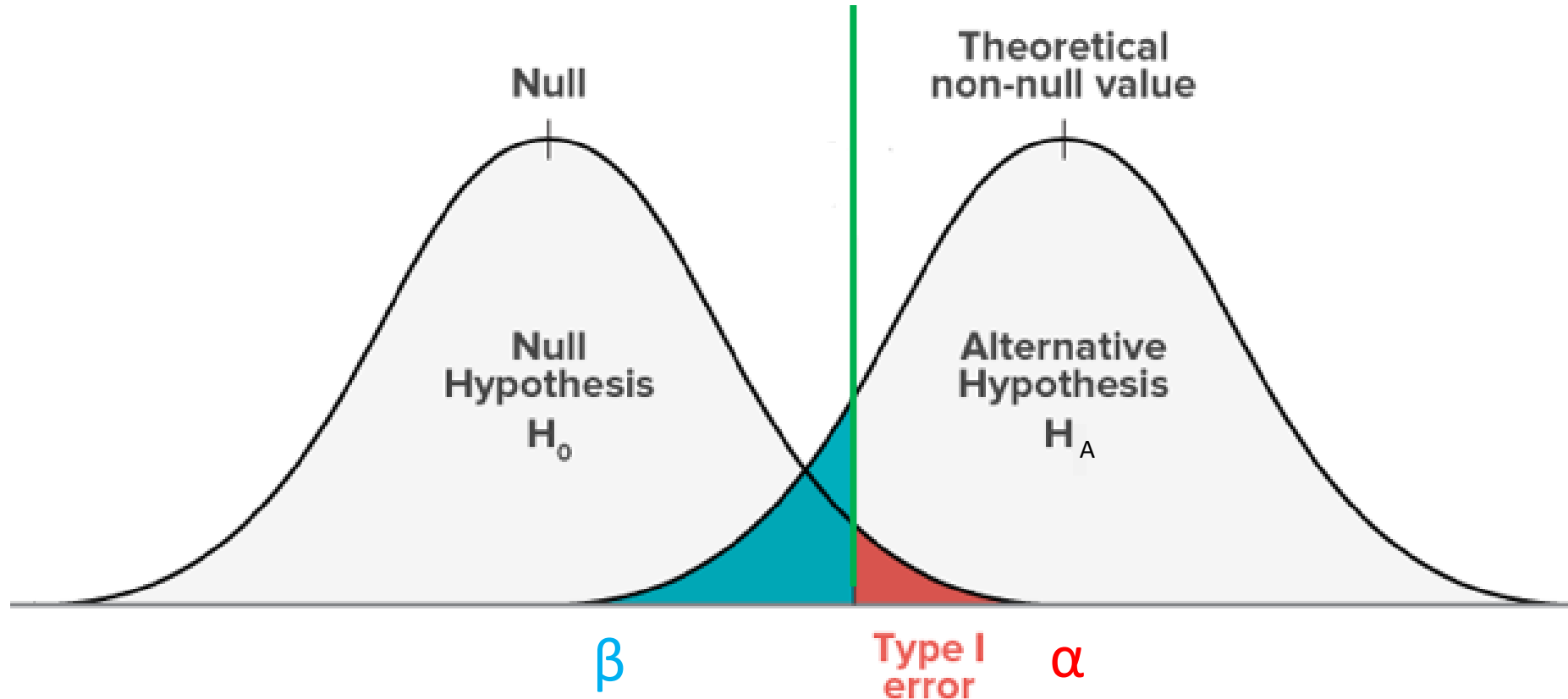
Type II error: incorrectly rejecting failing to reject H_0 when it is false
  • The rate at which we make type II errors is often denoted with the symbol β

Neyman-Pearson Frequentist logic
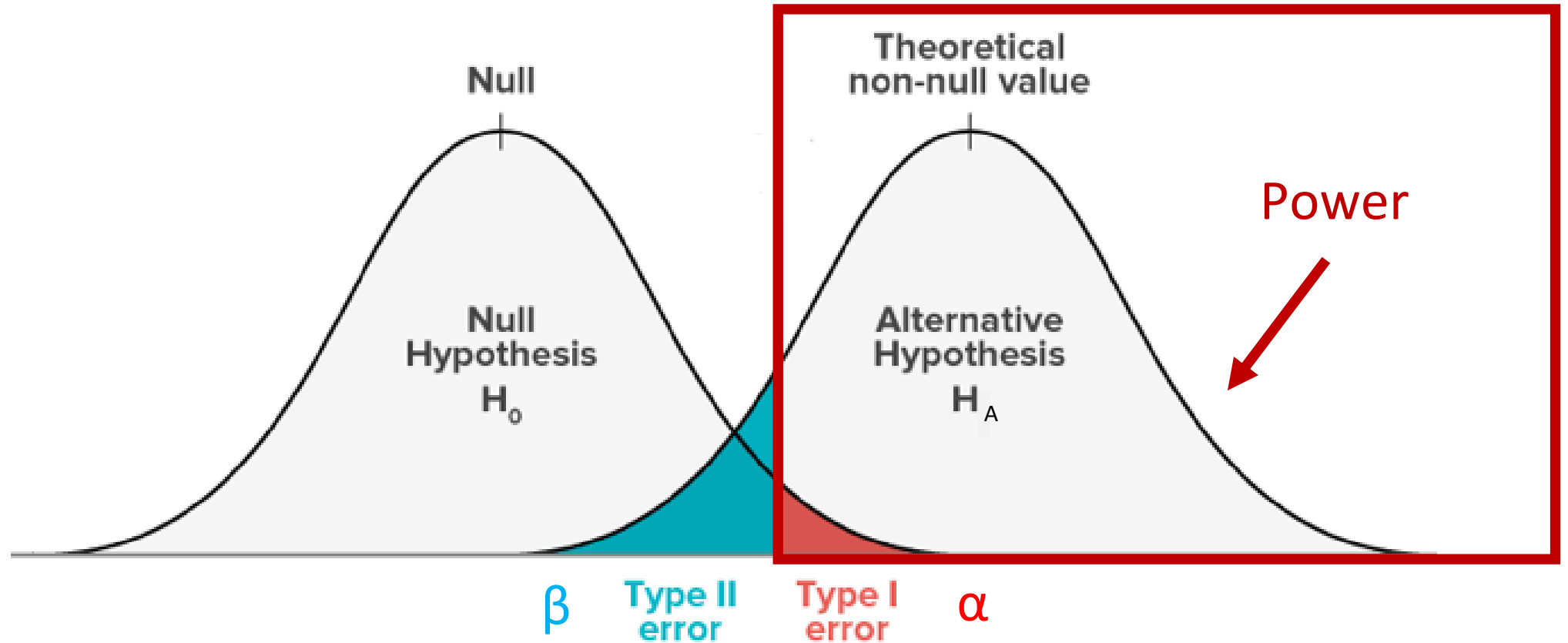
The **power** of a test is the probability we reject the H_0 when it is **false**
- 1 - β
- For a fixed α level, it would be best to use the most powerful test

Type I and Type II Errors

Null

Theoretical non-null value

Null Hypothesis H_0

Alternative Hypothesis H_A
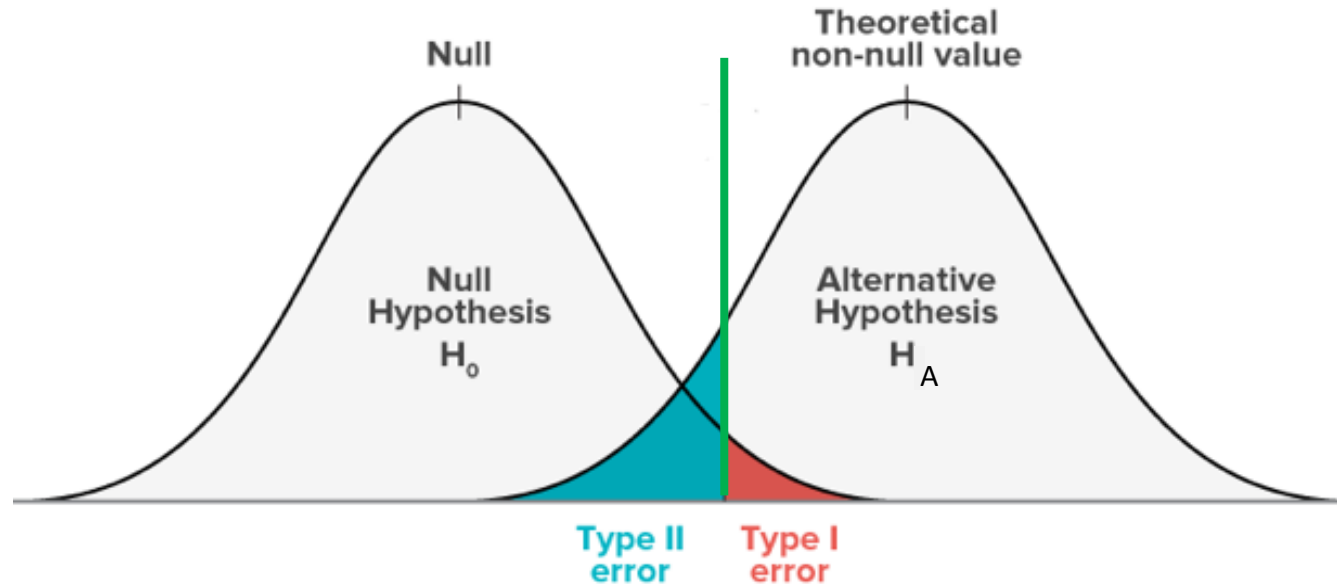
Type II error   Type I error

Decision

	Reject H_0	Do not reject H_0
H_0 is true	Type I error (α) (false positive)	No error
H_A is true (H₀ is false)	No error	Type II error (β) (false negative)

Truth

Problems with the NP hypothesis tests

Problem 1: we are interested in the results of a specific experiment, not whether we are right most of the time

- E.g., 95% of these statements are false:
  - Joy can't smell Parkinson's disease, Lawyers are left-handed at the same rate as the general population, Calcium is not beneficial for your heart, …

Problem 2: Arbitrary thresholds for alpha levels

- P-value = 0.051, we don't reject H_0

Problems with the NP hypothesis tests

Problem 1: we are interested in the results of a specific experiment, not whether we are right most of the time

- E.g., 95% of these statements are false:
  - Joy can't smell Parkinson's disease, Lawyers are left-handed at the same rate as the general population, Calcium is not beneficial for your heart, …

Problem 2: Arbitrary thresholds for alpha levels

- P-value = 0.051, we don't reject H_0?

Problem 3: running many tests can give rise to a high number of type 1 errors

Genes and leukemia example

Scientists collected 7129 gene expression levels from 38 patients to find genetic differences between two types leukemia (L1 and L2)

Suppose there was no genetic differences between the types of leukemia

- H_0: μ_{L1} = μ_{L2} is true for all genes

Q: If each gene was tested separately using a significance level of α = 0.05, approximately how many type I errors would be expected?
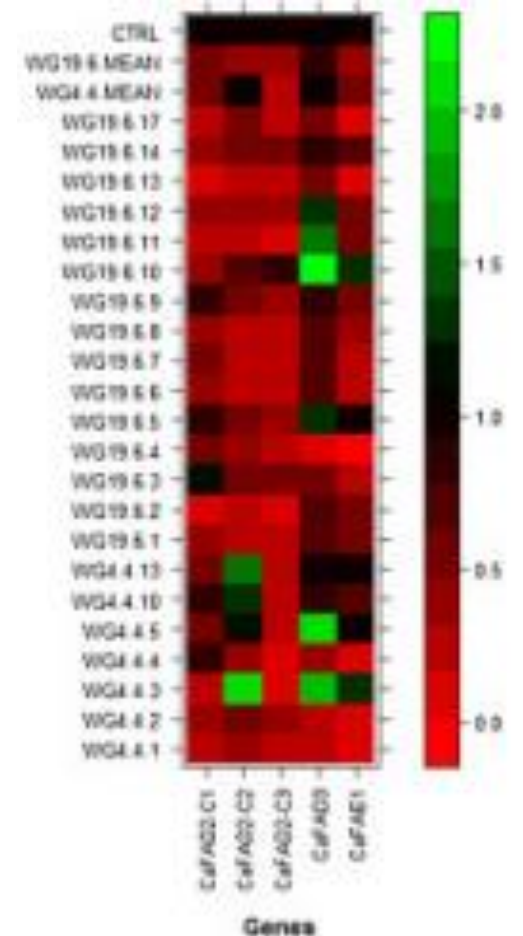
Genes and leukemia example

There are methods that try to correct for running multiple hypothesis tests

The **_Bonferroni correction_** is one way that controls the probability of **_any_** hypothesis test giving a type I error
- i.e., controls the familywise error rate   (no type I errors for any of the tests run)

It works by dividing the initial α level by the number of tests run
- E.g., $\alpha = 0.05/7129 = 0.000007$
- All p-values need to be below this level to be considered statistically significant
- This can lead to many type II errors
  - (Type II error: failure to reject H_0 when it is false)

The problem of multiple testing

For α = 0.05, ~5% of all published research findings should incorrectly reject the null hypothesis

Publication bias (file drawer effect): Generally positive results are more likely to be published, so if you read the literature, the proportion of incorrect results is likely to be greater than 5%.

Why Most Published Research Findings Are False

John P. A. Ioannidis

The Earth Is Round ($p < .05$)

Jacob Cohen

After 4 decades of severe criticism, the ritual of null hypothesis significance testing—mechanical dichotomous decisions around a sacred .05 criterion—still persists. This article reviews the problems with this practice, including sure how to test H_0, chi-square with Yates's (1951) correction or the Fisher exact test, and wonders whether he has enough power. Would you believe it? And would you believe that if he tried to publish this result without a

American Statistical Association's Statement on p-values

Some thoughts…

Better to have hypothesis tests than none at all. Just need to think carefully and use your judgment.

Report effect size in most cases – i.e., confidence intervals

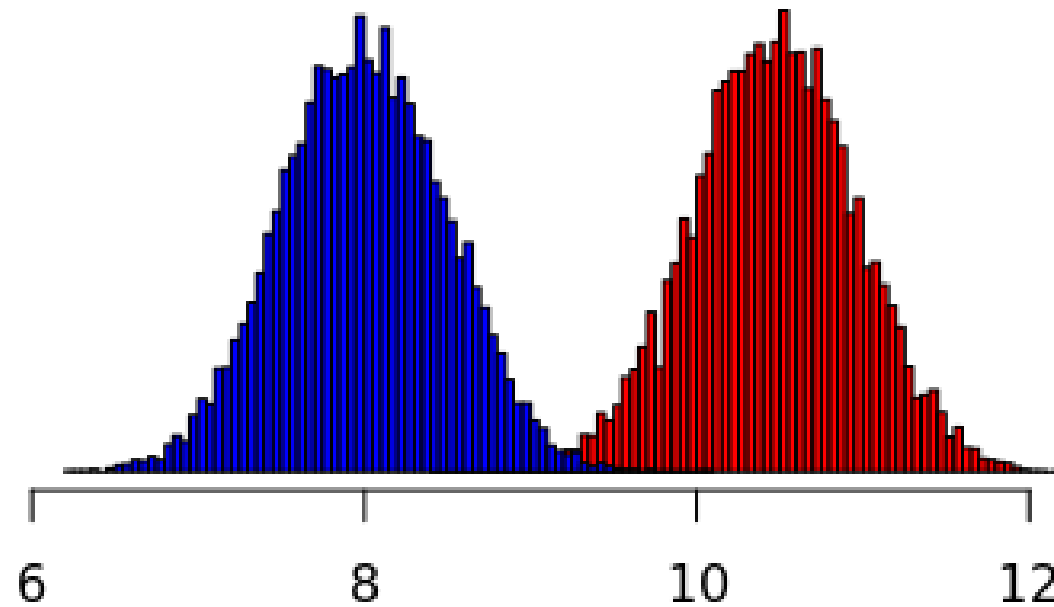Report the p-values rather than accept/reject H_0
- i.e., report   p = 0.23   not   p < 0.05

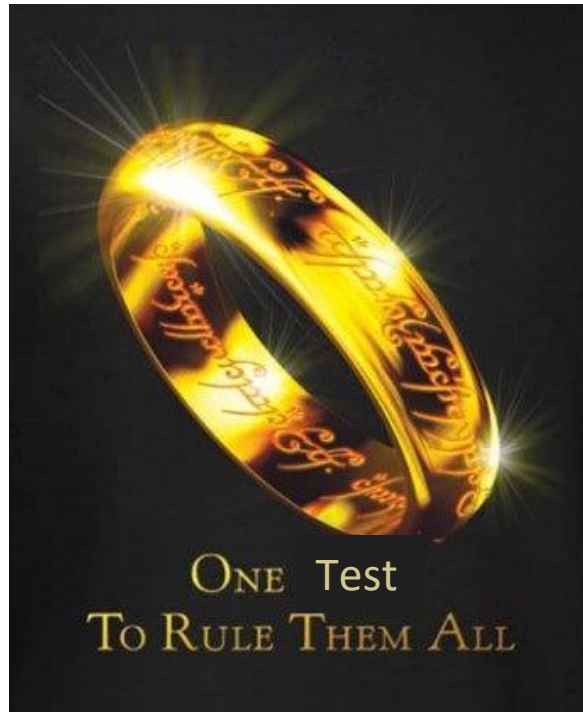Replicate findings (perhaps in different contexts) to make sure you get the same results

Be a good/honest scientists and try to get at the Truth!

THE TRUTH IS OUT THERE

Connections between null, alternative and bootstrap distribution using test of a single mean

The big picture: There is only one hypothesis test!

The hypothesis test zoo

ONE Test **TO RULE THEM ALL**

We can run a large number of additional hypothesis tests by following the 5 steps!

Single sample

Single sample t-test
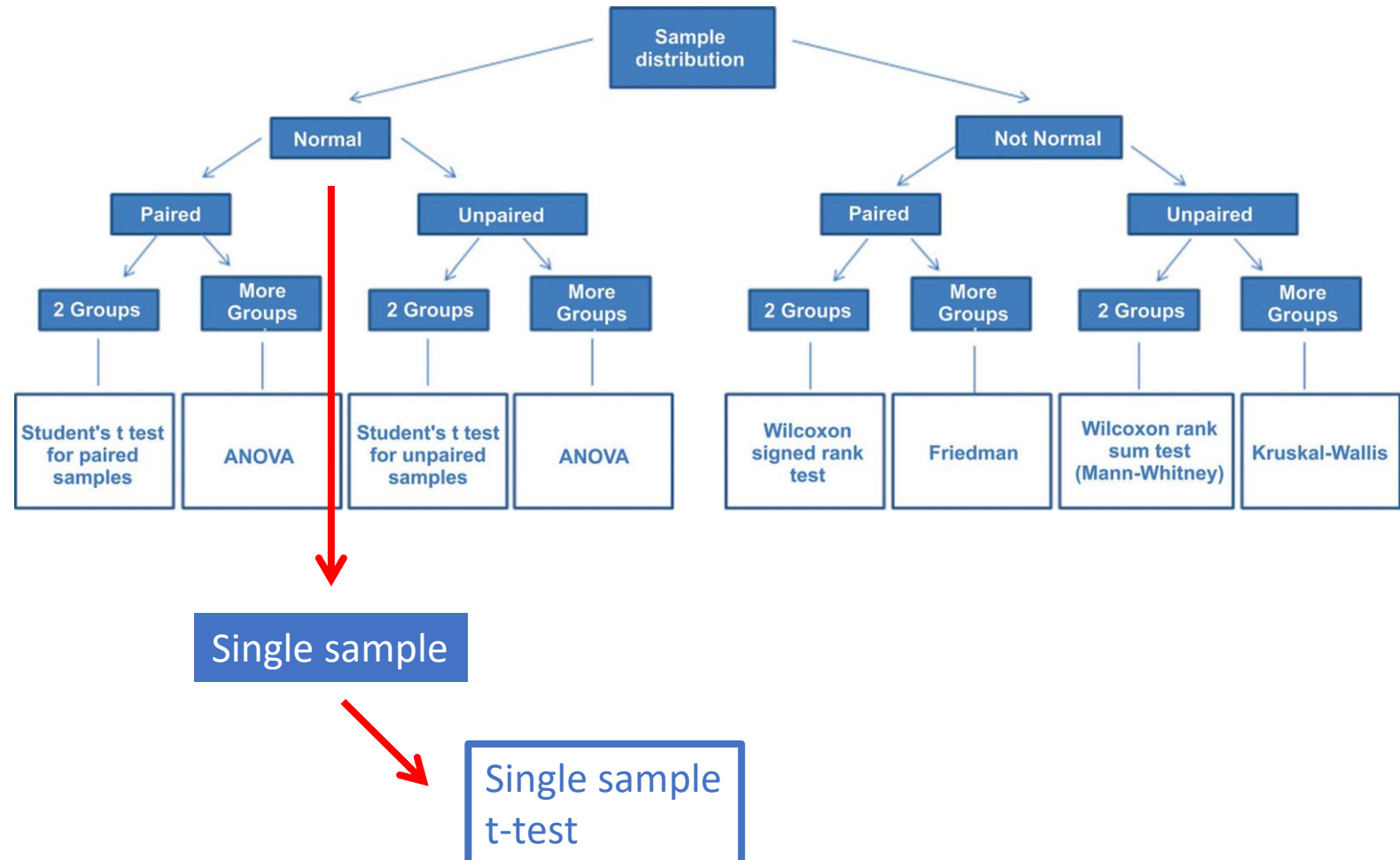
Example: Do mammals on average sleep more than humans?

According to a data set that comes with the ggplot package, humans sleep 8 hours a day

The data set also has the sleeping times of 82 other mammals

Let's test if the average sleep time of all mammals is different than 8 hours, based on the sample of 82 mammals.

- (warning: we obviously need to be careful drawing conclusions here because it's not clear whether this is a simple random sample)

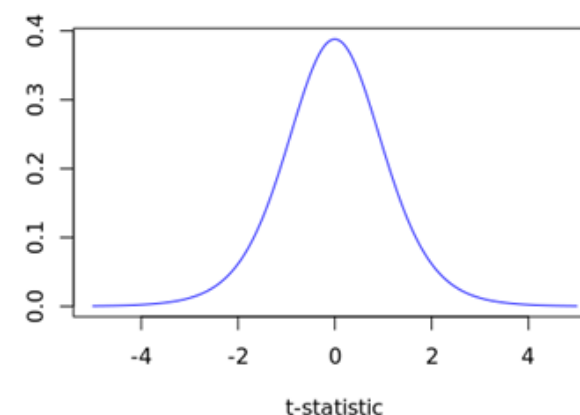# Parametric hypothesis test for a single mean

Step 1: state the null hypothesis:

Step 2: We can use a t-statistic:

$$t = \frac{estimate \ - \ param_0}{\hat{SE}} \qquad \hat{SE} = \frac{s}{\sqrt{n}}$$

$$t = \frac{\bar{x} \ - \ 8}{\frac{s}{\sqrt{n}}}$$

$\bar{x} = 10.46 \qquad n = 82$

$s = 4.47 \qquad t = 4.99$

Note: In a paired samples t-test we subtract the paired values in the two samples and run a one sample t-test on the differences.

Step 3: The null distribution is a t-distribution with n - 1 degrees of freedom

t-statistic

Step 4 and 5... ???

We can also get confidence intervals using:

$$CI = \bar{x} \pm t^* \cdot \frac{s}{\sqrt{n}}$$

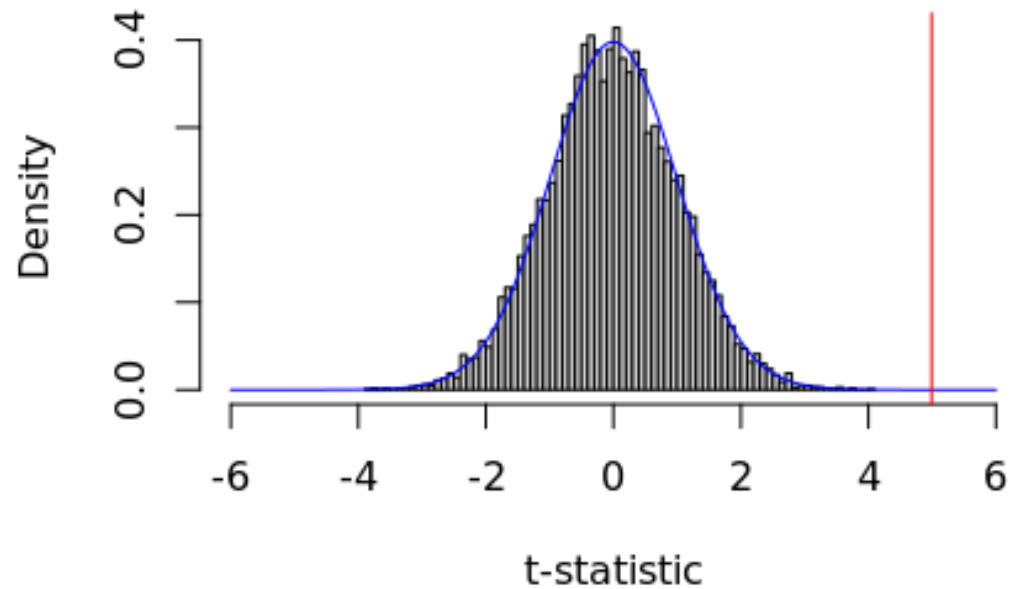# Randomization hypothesis test for a single mean

Step 1: Null hypothesis:

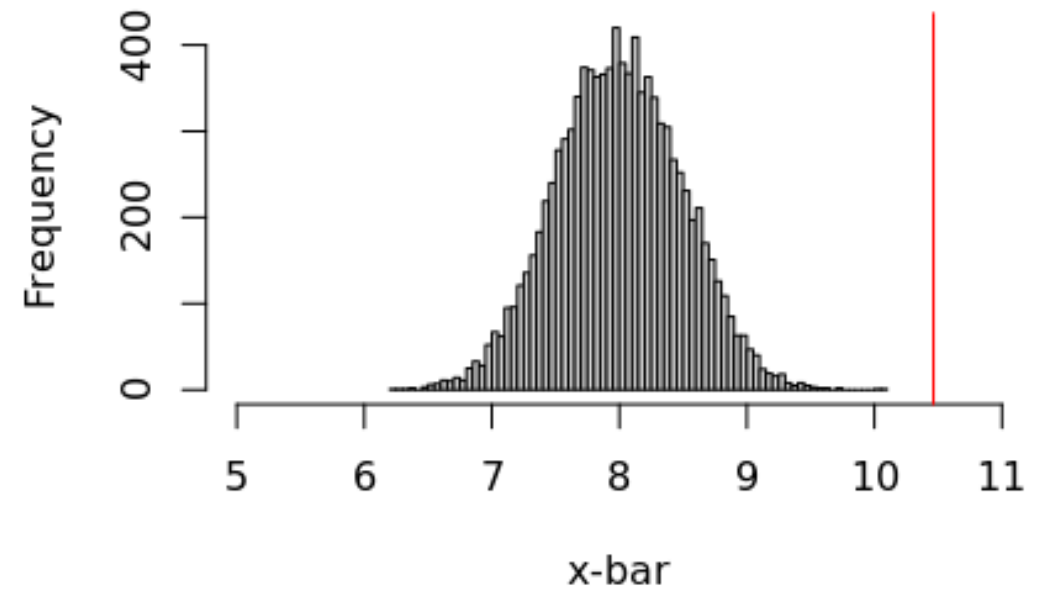**Step 2**: We could use…

Step 3: Any ideas how to create one point in our null distribution?

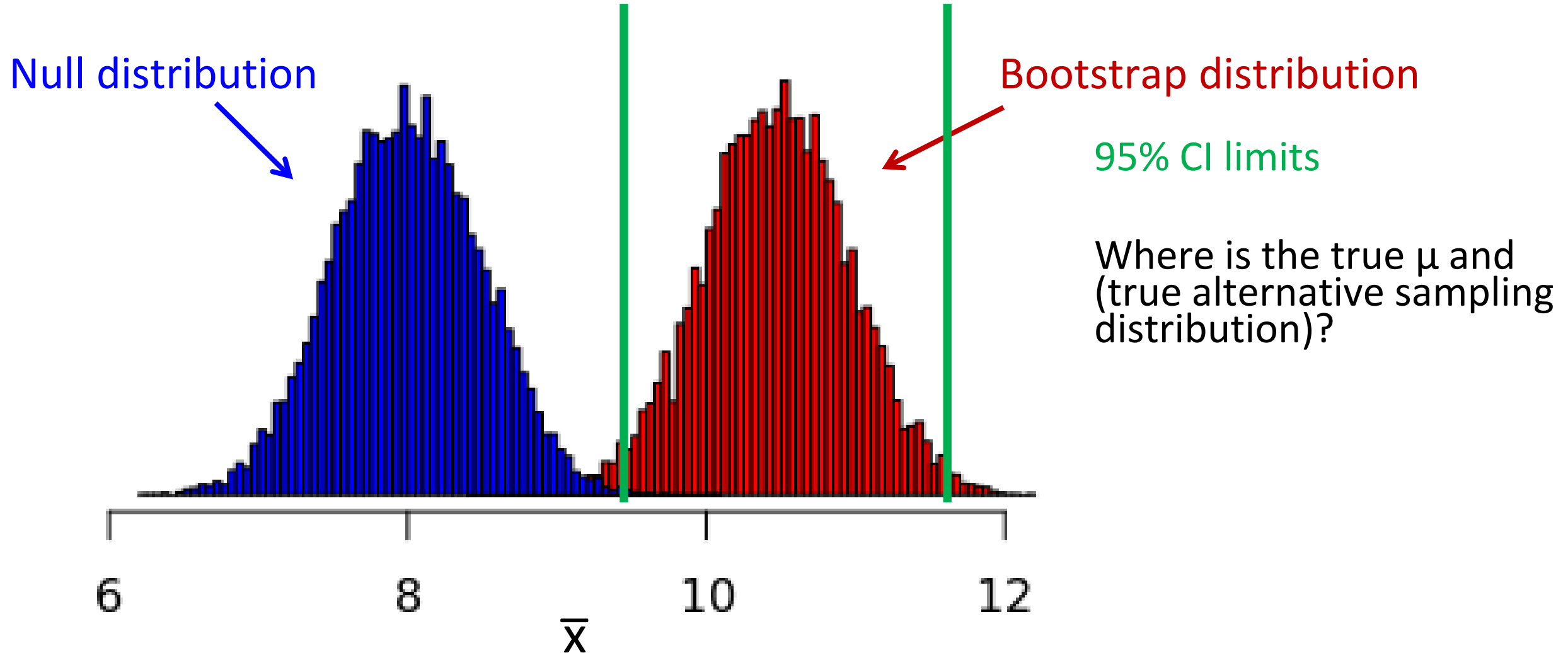# Null distributions

Null distribution using a t-statistic

Null distribution using x̄ statistic

The p-value in both cases is…

Relationship between null and bootstrap distributions

Null distribution

Bootstrap distribution

95% CI limits

Where is the true μ and (true alternative sampling distribution)?

$\overline{x}$

Next class: start on the tidyverse…