

yc4384_P8122_HW3

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Preparing mice data

```
set.seed(124)
n = 16
p_C = 1/5
C = rbinom(n, 1, p_C)
theta0 = 1/2
theta1 = -1/5
p_A = theta0 + theta1 * C
A = rbinom(n, 1, p_A)
beta0 = 110
beta1 = 20
beta2 = 5
sigma_Y = 1
mu_Y = beta0 + beta1 * C + beta2 * A
Y = rnorm(n, mu_Y, sigma_Y)

mice_data = data.frame(
  mouse_id = 1:n,
  obesity = C,
  light_exposure = A,
  glucose = Y
)

head(mice_data)
```

```
##   mouse_id obesity light_exposure  glucose
## 1         1       0              1 114.8490
## 2         2       0              1 113.7769
## 3         3       0              1 114.1312
## 4         4       0              0 108.9575
## 5         5       0              1 113.8964
## 6         6       0              1 115.4442
```

Problem 1: Interpretation of parameters:

p : The baseline covariate C (obesity) follows Bernoulli distribution, where p represents the probability of $C = 1$ (with obesity).

θ_0 : The probability of assigning the units to treatment group (light) if the units are not with obesity ($C = 0$).

θ_1 : The probability of assigning the units with obesity to treatment group (light) is 1/5 times lower than the units without obesity on average.

β_0 : The mean of baseline glucose when the mice are non-obese and unexposed to light.

β_1 : Intervening to increase C (obesity) by one unit will, on average, increase the outcome Y (glucose) by 20 units, holding other covariates constants.

β_2 : Intervening to increase A (light) by one unit will, on average, decrease the outcome Y (glucose) by 5 units, holding other covariates constants.

Problem 2: Marginal and Conditional PACE

Since $P(C = 1) = \frac{1}{5}$

and $E[Y | A, C] = \beta_0 + \beta_1 C + \beta_2 A = 110 + 20C - 5A$

The Marginal PACE is:

$$\begin{aligned}
& E[Y_1] - E[Y_0] \\
&= \sum E[Y | A = 1, C = c] \cdot P(C = c) - \sum E[Y | A = 0, C = c] \cdot P(C = c) \\
&= E[Y | A = 1, C = 1] \cdot P(C = 1) + E[Y | A = 1, C = 0] \cdot (1 - P(C = 1)) \\
&\quad - [E[Y | A = 0, C = 1] \cdot P(C = 1) + E[Y | A = 0, C = 0] \cdot (1 - P(C = 1))] \\
&= (\beta_0 + \beta_1 + \beta_2) \cdot P(C = 1) + (\beta_0 + \beta_2) \cdot (1 - P(C = 1)) - [(\beta_0 + \beta_1) \cdot P(C = 1) + \beta_0 \cdot (1 - P(C = 1))] \\
&= (\beta_0 + \beta_1 + \beta_2) \cdot P(C = 1) + (\beta_0 + \beta_2) \cdot (1 - P(C = 1)) - [(\beta_0 + \beta_1) \cdot P(C = 1) + \beta_0 \cdot (1 - P(C = 1))] \\
&= \beta_2 \\
&= -5
\end{aligned}$$

Similarly, the conditional PACE is:

$$\begin{aligned}
& E[Y_1 | C = c] - E[Y_0 | C = c] \\
&= E[Y | A = 1, C = c] - E[Y | A = 0, C = c] \\
&= (\beta_0 + \beta_1 \cdot C + \beta_2) - (\beta_0 + \beta_1 \cdot C) \\
&= \beta_2 \\
&= -5
\end{aligned}$$

I. Under the following assumptions Marginal PACE can be identified:

1) Consistency:

$$Y = Y_1 \text{ if treated, and } Y = Y_0 \text{ if not treated.}$$

The observed outcome is equal to the potential outcome for the treatment received.

2) Exchangeability:

$$Y_1, Y_0 \perp A | X.$$

Conditional on covariates X , treatment assignment is independent of the potential outcomes.

3) Positivity:

$$0 < P(A = 1 | X) < 1.$$

There must be a positive probability of receiving each treatment level for all values of covariates.

II. Under the following assumptions Conditional PACE can be identified:

1) Conditional Consistency:

$$Y = Y_1 \text{ if } A = 1, \quad Y = Y_0 \text{ if } A = 0, \quad \text{given } X.$$

Similar to marginal consistency, but conditional on X .

2) Conditional Exchangeability:

$$Y_1, Y_0 \perp A \mid X = x.$$

The treatment assignment is independent of the potential outcomes, conditional on covariates X .

3) Conditional Positivity:

$$0 < P(A = 1 \mid X = x) < 1.$$

Within each level of X , there must be a positive probability of receiving both the treatment and control.

Problem 3: G-formula for Randomized and Observational Studies

I. In a randomized study, the treatment assignment $A \perp C$. Therefore:

$$\begin{aligned} E[Y_a] &= \sum E[Y \mid A = a, C = c] \cdot P(C = c) \\ &= E[Y \mid A = a] \\ \text{Thus, } E[Y_1] &= E[Y \mid A = 1] \\ \text{and } E[Y_0] &= E[Y \mid A = 0] \end{aligned}$$

II. In an observational study, treatment assignment A is not independent of the covariates C . As a result, we need to adjust for the distribution of covariates:

$$\begin{aligned} E[Y_a] &= \sum E[Y \mid A = a, C = c] \cdot P(C = c) \\ \text{where} & \\ E[Y_1] &= \sum E[Y \mid A = 1, C = c] \cdot P(C = c) \\ &= E[Y \mid A = 1, C = 1] \cdot P(C = 1) + E[Y \mid A = 1, C = 0] \cdot (1 - P(C = 1)) \\ \text{and} & \\ E[Y(0)] &= \sum E[Y \mid A = 0, C = c] \cdot P(C = c) \\ &= E[Y \mid A = 0, C = 1] \cdot P(C = 1) + E[Y \mid A = 0, C = 0] \cdot (1 - P(C = 1)) \end{aligned}$$

Thus, the key difference between the two studies is how the treatment assignment interacts with the covariates, which affects the application of the g-formula.