BM2_HW2

Yangyang Chen

2024-02-16

Problem 1

(a) Fill out the table and give comments.

```
# load data
dose=c(0:4)## x_i
num=rep(30, 5) ## m i
killed=c(2, 8, 15, 23, 27) ## y_i
data=data.frame(dose,num,killed) ## (x_i, m_i, y_i)
# data preparation
x=data$dose
v=data$killed
m=data$num
resp=cbind(y,m-y) #### counts of success (death=1), failure (surv=0)
# Model fitting
# logit link
glm_logit=glm(resp~x, family=binomial(link='logit'))
# probit link
glm_probit=glm(resp~x, family=binomial(link='probit'))
# complementary log-log link
glm_clog=glm(resp~x, family=binomial(link='cloglog')) # asymmetric
results_func = function(fit,model){
  fit summary = fit |> summary()
  beta = fit_summary$coefficients[2]
  CI = fit$coefficients +
    kronecker(t(c(0,qnorm(0.025),-qnorm(0.025))), t(t(sqrt(diag(vcov(fit))))))
  CI lower = CI[2,2]
  CI\_upper = CI[2,3]
  dev = fit_summary$deviance
  p_new = predict.glm(fit, newdata = tibble(x = 0.01), type = "response")
  return(tibble(model, beta, CI_lower, CI_upper, dev, p_new))
}
results_func(glm_logit, "logit") |>
  rbind(results_func(glm_probit, "probit")) |>
  rbind(results_func(glm_clog, "c-log-log")) |>
  knitr::kable(digits = 4)
```

model	beta	CI_lower	CI_upper	dev	p_new
logit probit	1.1619 0.6864	0.8063 0.4967	1.5175 0.8760	0.3787 0.3137	0.0901 0.0853
c-log-log	0.7468	0.5323	0.9613	2.2305	0.0033 0.1282

- (a) All models showed a significant relationship between dose and death with significant p-values less than 0.001.
- (b) As the dose increases, the probability of death decreases because all CIs for the dose estimates are positive and do not include zero.
- (c) The probit link model fits the data best since it minimizes residual deviance.
- (d) The probability of death estimated using the probit link model at a dose level of 0.01 was 0.0853.

(b) Suppose that the dose level is in natural logarithm scale, estimate LD50 with 90% confidence interval based on the three models.

```
LD50_func = function(fit, model){
       alpha = fit$coefficients[1]
       beta = fit$coefficients[2]
       betacov = vcov(fit)
       if(model == "c-log-log"){
              x0fit = (log(-log(0.5)) - alpha) /beta
              varx0 = betacov[1,1]/(beta^2)+
                      betacov[2,2]*(log(-log(0.5)) - alpha) ^ 2 / (beta ^ 4) + 2 * betacov[1,2] * (log(-log(0.5)) - alpha) ^ 2 / (beta ^ 4) + 2 * betacov[1,2] * (log(-log(0.5)) - alpha) ^ 2 / (beta ^ 4) + 2 * betacov[1,2] * (log(-log(0.5)) - alpha) ^ 2 / (beta ^ 4) + 2 * betacov[1,2] * (log(-log(0.5)) - alpha) ^ 2 / (beta ^ 4) + 2 * betacov[1,2] * (log(-log(0.5)) - alpha) ^ 2 / (beta ^ 4) + 2 * betacov[1,2] * (log(-log(0.5)) - alpha) ^ 2 / (beta ^ 4) + 2 * betacov[1,2] * (log(-log(0.5)) - alpha) ^ 2 / (beta ^ 4) + 2 * betacov[1,2] * (log(-log(0.5)) - alpha) ^ 2 / (beta ^ 4) + 2 * betacov[1,2] * (log(-log(0.5)) - alpha) ^ 2 / (beta ^ 4) + 2 * betacov[1,2] * (log(-log(0.5)) - alpha) ^ 2 / (beta ^ 4) + 2 * betacov[1,2] * (log(-log(0.5)) - alpha) ^ 2 / (beta ^ 4) + 2 * betacov[1,2] * (log(-log(0.5)) - alpha) ^ 2 / (beta ^ 4) + 2 * betacov[1,2] * (log(-log(0.5)) - alpha) ^ 2 / (beta ^ 4) + 2 * betacov[1,2] * (log(-log(0.5)) - alpha) ^ 2 / (beta ^ 4) + 2 * betacov[1,2] * (log(-log(0.5)) - alpha) ^ 2 / (beta ^ 4) + 2 * betacov[1,2] * (log(-log(0.5)) - alpha) ^ 2 / (beta ^ 4) + 2 * betacov[1,2] * (log(-log(0.5)) - alpha) ^ 2 / (beta ^ 4) + 2 * betacov[1,2] * (log(-log(0.5)) - alpha) ^ 2 / (beta ^ 4) + 2 * betacov[1,2] * (log(-log(0.5)) - alpha) ^ 2 / (beta ^ 4) + 2 * betacov[1,2] * (log(-log(0.5)) - alpha) ^ 2 / (beta ^ 4) + 2 * betacov[1,2] * (log(-log(0.5)) - alpha) ^ 2 / (beta ^ 4) + 2 * betacov[1,2] * (log(-log(0.5)) - alpha) ^ 2 / (beta ^ 4) + 2 * betacov[1,2] * (log(-log(0.5)) - alpha) ^ 2 / (beta ^ 4) + 2 * betacov[1,2] * (log(-log(0.5)) - alpha) ^ 2 / (beta ^ 4) + 2 * betacov[1,2] * (log(-log(0.5)) - alpha) ^ 2 / (beta ^ 4) + 2 * betacov[1,2] * (log(-log(0.5)) - alpha) ^ 2 / (beta ^ 4) + 2 * betacov[1,2] * (log(-log(0.5)) - alpha) ^ 2 / (beta ^ 4) + 2 * betacov[1,2] * (log(-log(0.5)) - alpha) ^ 2 / (beta ^ 4) + 2 * betacov[1,2] * (log(-log(0.5)) - alpha) ^ 2 / (beta ^ 4) + 2 * betacov[1,2] * (log(-log(0.5)) - alpha) ^ 2 / (beta ^ 4) + 2 * betacov[1,2] * (log(-log(0.5)) - alpha) ^ 2 / (beta ^ 4) + 2 * betacov[1,2] * (log(-log(0.5))
       }
       else
       {
              x0fit = -alpha / beta
              varx0 = betacov[1,1] / (beta ^ 2) +
                     betacov[2,2] * (alpha ^ 2) / (beta ^ 4) -
                      2 * betacov[1,2] *alpha / (beta ^ 3)
       }
       estimate = exp(x0fit)
       CI_lower = exp(x0fit - c(qnorm(0.95)) * sqrt(varx0))
       CI_upper = exp(x0fit + c(qnorm(0.95))*sqrt(varx0))
       return(tibble(model, estimate, CI_lower, CI_upper))
}
LD50_func(glm_logit, "logit") |>
       rbind(LD50_func(glm_probit, "probit")) |>
       rbind(LD50_func(glm_clog, "c-log-log")) |>
       knitr::kable(digits = 2)
```

model	estimate	CI_lower	CI_upper
logit	7.39	5.51	9.91
probit	7.44	5.58	9.90
c-log-log	8.84	6.53	11.98

Problem 2

```
amount=seq(10, 90, by = 5)
offers=c(4, 6, 10, 12, 39, 36, 22, 14, 10, 12, 8, 9, 3, 1, 5, 2, 1)
enrolls=c(0, 2, 4, 2, 12, 14, 10, 7, 5, 5, 3, 5, 2, 0, 4, 2, 1)
data=data.frame(amount, offers, enrolls)
glm_logit_2 = glm(cbind(enrolls, offers - enrolls) ~ amount,
                family = binomial(link = "logit"))
glm_logit_2 |> summary()
##
## Call:
## glm(formula = cbind(enrolls, offers - enrolls) ~ amount, family = binomial(link = "logit"))
##
## Coefficients:
              Estimate Std. Error z value Pr(>|z|)
##
## (Intercept) -1.64764
                           0.42144 -3.910 9.25e-05 ***
               0.03095
                                    3.197 0.00139 **
## amount
                           0.00968
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for binomial family taken to be 1)
##
##
       Null deviance: 21.617 on 16 degrees of freedom
## Residual deviance: 10.613 on 15 degrees of freedom
## AIC: 51.078
##
## Number of Fisher Scoring iterations: 4
```

(a) How does the model fit the data?

Since most of the offers are less than 30, we can assume that the data is sparse. Therefore, we should use the Hosmer-Lemeshow statistic to evaluate the goodness of fit.

```
library(ResourceSelection)
```

```
## ResourceSelection 0.3-6 2023-06-27
hoslem_stat = hoslem.test(glm_logit_2$y, fitted(glm_logit_2), g = 10)
hoslem_stat

##
## Hosmer and Lemeshow goodness of fit (GOF) test
##
## data: glm_logit_2$y, fitted(glm_logit_2)
## X-squared = 1.6111, df = 8, p-value = 0.9907
```

Because p = 0.9907 > 0.05, we fail to reject the null and conclude that the model fits the data well at 95% significant level.

(b) How do you interpret the relationship between the scholarship amount and enrollment rate? What is 95% CI?