

Probability and Statistics

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Section 4.1

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P116 Q6

a.

$$\begin{aligned} E(x) &= \int_{-\infty}^{\infty} x f(x) dx \\ &= \int_0^1 x \cdot 2x dx \\ &= \frac{2}{3} \end{aligned}$$

b.

Y 的分布函数为

$$\begin{aligned} F_Y(y) &= P\{Y \leq y\} \\ &= P\{X^2 \leq y\} \\ &= P\{X \leq \sqrt{y}\} \\ &= \int_0^{\sqrt{y}} 2x dx \\ &= y \quad (0 \leq y \leq 1) \end{aligned}$$

Y 的密度函数为

$$f_Y(y) = \begin{cases} \frac{dF_Y(y)}{dy} = 1 & (0 \leq y \leq 1) \\ 0 & \text{otherwise} \end{cases}$$

Y 的期望为

$$\begin{aligned} E(Y) &= \int_0^1 y \cdot 1 dy \\ &= \frac{1}{2} \end{aligned}$$

c.

直接计算 X^2 的期望为

$$\begin{aligned} E(X^2) &= \int_0^1 x^2 \cdot 2x dx \\ &= \frac{1}{2} x^4 \Big|_0^1 \\ &= \frac{1}{2} \end{aligned}$$

d.

根据方差的定义有

$$\begin{aligned} Var(x) &= \int_{-\infty}^{\infty} (x - E(x))^2 f(x) dx \\ &= \int_0^1 \left(x - \frac{2}{3}\right)^2 \cdot 2x dx \\ &= \frac{1}{18} \end{aligned}$$

由定理 4.2.2 计算方差为

$$\begin{aligned} Var(x) &= E(x^2) - [E(x)]^2 \\ &= \frac{1}{2} - \left(\frac{2}{3}\right)^2 \\ &= \frac{1}{18} \end{aligned}$$

P117 Q15

从一种彩票中购买两张的期望为

$$\begin{aligned} E(X) &= \sum_{x=0}^{C_n^2} c \cdot P(X = x) \\ &= c \cdot \frac{n-1}{C_n^2} \\ &= \frac{2c}{n} \end{aligned}$$

从两种彩票中各购买一张的期望为

$$\begin{aligned} E(Y) &= 2 \cdot \sum_{x=0}^n c \cdot P(Y = x) \\ &= 2 \cdot c \cdot \frac{1}{n} \\ &= \frac{2c}{n} \end{aligned}$$

可知，两种彩票的期望相同。

P117 Q20

当 X 为泊松随机变量时, 有

$$\begin{aligned}
 E\left(\frac{1}{1+X}\right) &= \sum_{x=0}^{\infty} \frac{1}{1+x} \cdot \frac{e^{-\lambda} \lambda^x}{x!} \\
 &= \sum_{x=0}^{\infty} \frac{e^{-\lambda} \lambda^x}{(x+1)!} \\
 &= \frac{e^{-\lambda}}{\lambda} \sum_{x=0}^{\infty} \frac{\lambda^{x+1}}{(x+1)!} \\
 &= \frac{e^{-\lambda}}{\lambda} \sum_{x=1}^{\infty} \frac{\lambda^x}{x!} \\
 &= \frac{e^{-\lambda}}{\lambda} \cdot (e^{\lambda} - 1)
 \end{aligned}$$

P117 Q21

由题可知 $X \sim U(0, 1)$, 则 $S = X^2$ 的期望为

$$\begin{aligned}
 E(S) &= \int_0^1 x^2 P(x) dx \\
 &= \int_0^1 x^2 dx \\
 &= \frac{1}{3}
 \end{aligned}$$

P117 Q31

由题可知 $X \sim U(1, 2)$, 则 $Y = \frac{1}{X}$ 的期望为

$$\begin{aligned}
 E(Y) &= \int_1^2 \frac{1}{x} \cdot \frac{1}{2-1} dx \\
 &= \ln 2
 \end{aligned}$$

易知 X 的期望为 $\frac{1+2}{2} = \frac{3}{2}$ 。明显有

$$E\left(\frac{1}{X}\right) \neq \frac{1}{E(X)}$$

补充 1

(1)

$$\begin{aligned}
 E(Y) &= \int_{-\infty}^{\infty} 2x \cdot P\{X = x\} dx \\
 &= \int_0^{\infty} 2x \cdot e^{-x} dx \\
 &= -2(x+1)e^{-x} \Big|_0^{\infty} \\
 &= 2
 \end{aligned}$$

(2)

$$\begin{aligned}
 E(Y) &= \int_{-\infty}^{\infty} e^{-2x} \cdot P\{X = x\} dx \\
 &= \int_0^{\infty} e^{-2x} \cdot e^{-x} dx \\
 &= -\frac{1}{3}e^{-3x} \Big|_0^{\infty} \\
 &= \frac{1}{3}
 \end{aligned}$$

补充 2

$$\begin{aligned}
 E(X) &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} x \cdot f(x, y) dx dy \\
 &= \int_0^1 \int_y^1 x \cdot 12y^2 dx dy \\
 &= \int_0^1 12y^2 \int_y^1 x dx dy \\
 &= \int_0^1 12y^2 \cdot \frac{1}{2}(1 - y^2) dy \\
 &= \int_0^1 6y^2 - 6y^4 dy \\
 &= 2y^3 - \frac{6}{5}y^5 \Big|_0^1 \\
 &= \frac{4}{5}
 \end{aligned}$$

$$\begin{aligned}
E(Y) &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} y \cdot f(x, y) dx dy \\
&= \int_0^1 \int_y^1 y \cdot 12y^2 dx dy \\
&= \int_0^1 12y^3 \cdot (1 - y) dy \\
&= \int_0^1 12y^3 - 12y^4 dy \\
&= 3y^4 - \frac{12}{5}y^5 \Big|_0^1 \\
&= \frac{3}{5}
\end{aligned}$$

$$\begin{aligned}
E(XY) &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} xy \cdot f(x, y) dx dy \\
&= \int_0^1 \int_y^1 xy \cdot 12y^2 dx dy \\
&= \int_0^1 12y^3 \int_y^1 x dx dy \\
&= \int_0^1 12y^3 \cdot \frac{1}{2}(1 - y^2) dy \\
&= \int_0^1 6y^3 - 6y^5 dy \\
&= \frac{3}{2}y^4 - y^6 \Big|_0^1 \\
&= \frac{1}{2}
\end{aligned}$$

$$\begin{aligned}
E(X^2 + Y^2) &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} (x^2 + y^2) \cdot f(x, y) dx dy \\
&= \int_0^1 \int_y^1 (x^2 + y^2) \cdot 12y^2 dx dy \\
&= \int_0^1 \int_y^1 12x^2y^2 + 12y^4 dx dy \\
&= \int_0^1 4y^2 + 12y^4 - 16y^5 dy \\
&= \frac{4}{3}y^3 + \frac{12}{5}y^5 - \frac{8}{3}y^6 \Big|_0^1 \\
&= \frac{16}{15}
\end{aligned}$$