

# Probability and Statistics

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## Section 6.3

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### P136 Q3

$$\frac{\bar{X} - \mu}{\frac{\sigma}{\sqrt{n}}} = \frac{\bar{X}}{\frac{1}{\sqrt{16}}} = 4\bar{X} \sim N(0, 1)$$

因此有

$$\begin{aligned} P(|\bar{X}| < c) &= P(-c < \bar{X} < c) \\ &= P(-4c < 4\bar{X} < 4c) \\ &= \Phi(4c) - \Phi(-4c) \\ &= 2\Phi(4c) - 1 \\ &= 0.5 \end{aligned}$$

即  $\Phi(4c) = 0.75$ , 查表得  $c \approx 0.1686$ 。

### P136 Q6

已知  $T \sim t_n$ , 则有

$$T = \frac{X}{\sqrt{\frac{Y}{n}}} \text{ 且 } X \sim N(0, 1), Y \sim \chi_n^2$$

因此有

$$T^2 = \frac{X^2}{\frac{Y}{n}} = \frac{\frac{X^2}{1}}{\frac{Y}{n}} \text{ 且 } X^2 \sim \chi_1^2, Y \sim \chi_n^2$$

因此,  $T^2 \sim F_{1,n}$ 。

**P136 Q8**

若  $X, Y$  为  $\lambda = 1$  的独立指数分布, 令  $Z = \frac{X}{Y}$ , 则有

$$\begin{aligned}
 f_Z(z) &= \int_0^\infty f_X(zy)f_Y(y)|y|dy \\
 &= \int_0^\infty e^{-zy}e^{-y}ydy \\
 &= \int_0^\infty ye^{-(z+1)y}dy \\
 &= \left[ -\frac{y}{z+1}e^{-(z+1)y} \right] \Big|_0^\infty + \int_0^\infty \frac{1}{z+1}e^{-(z+1)y}dy \\
 &= \frac{1}{z+1} \int_0^\infty e^{-(z+1)y}dy \\
 &= \frac{1}{z+1} \left[ -\frac{1}{z+1}e^{-(z+1)y} \right] \Big|_0^\infty \\
 &= \frac{1}{(z+1)^2}
 \end{aligned}$$

即

$$\begin{aligned}
 f_Z(z) &= \frac{1}{(z+1)^2} \\
 &= 4 \cdot \frac{1}{(2z+2)^2} \\
 &= \frac{1}{1 \cdot 1} \cdot 2^1 \cdot 2^1 \cdot \frac{z^{1-1}}{(2z+2)^2} \\
 &= \frac{\Gamma(\frac{2+2}{2})}{\Gamma(\frac{2}{2}) \cdot \Gamma(\frac{2}{2})} \cdot 2^{\frac{2}{2}} \cdot 2^{\frac{2}{2}} \cdot \frac{z^{\frac{2}{2}-1}}{(2z+2)^2}
 \end{aligned}$$

因此,  $Z \sim F(2, 2)$ 。

**补充 1**

$$\frac{\bar{X} - \mu}{\frac{\sigma}{\sqrt{n}}} = \frac{\bar{X} - \mu}{\frac{4}{\sqrt{n}}} \sim N(0, 1)$$

因此有

$$\begin{aligned}
 P(|\bar{X} - \mu| < 1) &= P(-1 < \bar{X} - \mu < 1) \\
 &= P\left(-\frac{\sqrt{n}}{4} < \frac{\bar{X} - \mu}{\frac{4}{\sqrt{n}}} < \frac{\sqrt{n}}{4}\right) \\
 &= \Phi\left(\frac{\sqrt{n}}{4}\right) - \Phi\left(-\frac{\sqrt{n}}{4}\right) \\
 &= 2\Phi\left(\frac{\sqrt{n}}{4}\right) - 1 \\
 &\geq 0.95
 \end{aligned}$$

即  $\Phi\left(\frac{\sqrt{n}}{4}\right) \geq 0.975$ , 查表得  $\frac{\sqrt{n}}{4} \geq 1.96$ , 解得  $n \geq 61.47$ 。

## 补充 2

$$\frac{(\bar{X} - \bar{Y}) - (\mu_1 - \mu_2)}{\sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}} = \frac{\bar{X} - \bar{Y}}{20\sqrt{\frac{1}{36} + \frac{1}{49}}} \sim N(0, 1)$$

因此有

$$\begin{aligned}
 P(|\bar{X} - \bar{Y}| \leq 10) &= P(-10 \leq \bar{X} - \bar{Y} \leq 10) \\
 &= P\left(-\frac{1}{2\sqrt{\frac{1}{36} + \frac{1}{49}}} \leq \frac{\bar{X} - \bar{Y}}{20\sqrt{\frac{1}{36} + \frac{1}{49}}} \leq \frac{1}{2\sqrt{\frac{1}{36} + \frac{1}{49}}}\right) \\
 &= 2\Phi\left(\frac{1}{2\sqrt{\frac{1}{36} + \frac{1}{49}}}\right) - 1 \\
 &\approx 0.9772
 \end{aligned}$$

## 补充 3

$$\sum_{i=1}^n \frac{(X_i - \mu)^2}{\sigma^2} = \frac{\sum_{i=1}^{10} X_i^2}{0.3^2} \sim \chi_{10}^2$$

因此有

$$\begin{aligned}
 P\left\{\sum_{i=1}^{10} X_i^2 \leq c\right\} &= P\left\{\frac{\sum_{i=1}^{10} X_i^2}{0.3^2} \leq \frac{c}{0.3^2}\right\} \\
 &= P\left\{\chi_{10}^2 \leq \frac{c}{0.3^2}\right\} \\
 &= 0.95
 \end{aligned}$$

查表得  $\frac{c}{0.3^2} \approx 18.307$ , 解得  $c \approx 1.64763$ 。

## 补充 4

(1)

可知  $X_1 - X_2 \sim N(0, 2\sigma^2)$ ,  $X_1 + X_2 \sim N(0, 2\sigma^2)$ , 因此有

$$\frac{X_1 - X_2}{\sqrt{2}\sigma} \sim \chi_1^2 \text{ 且 } \frac{X_1 + X_2}{\sqrt{2}\sigma} \sim \chi_1^2$$

两者相互独立, 其非奇异线性组合也相互独立, 因此有

$$\left( \frac{X_1 - X_2}{X_1 + X_2} \right)^2 = \frac{\left( \frac{X_1 - X_2}{\sqrt{2}\sigma} \right)^2}{\left( \frac{X_1 + X_2}{\sqrt{2}\sigma} \right)^2} \sim F_{1,1}$$

(2)

$$\frac{1}{\frac{(X_1 + X_2)^2}{(X_1 + X_2)^2 + (X_1 - X_2)^2}} = \frac{(X_1 + X_2)^2 + (X_1 - X_2)^2}{(X_1 + X_2)^2} = 1 + \frac{(X_1 - X_2)^2}{(X_1 + X_2)^2}$$

可将原式改为

$$P\left\{ \frac{(X_1 + X_2)^2}{(X_1 + X_2)^2 + (X_1 - X_2)^2} > k \right\} = 0.10 \rightarrow P\left\{ \frac{(X_1 - X_2)^2}{(X_1 + X_2)^2} < \frac{1}{k} - 1 \right\} = 0.10$$

查表可知  $F_{0.10}(1, 1) = \frac{1}{F_{0.90}(1, 1)} \approx \frac{1}{39.9}$ 。因此有  $\frac{1}{k} - 1 \approx \frac{1}{39.9}$ , 解得  $k \approx 0.9755$ 。

## 补充 5

已知  $\bar{X}_n \sim N(\mu, \frac{\sigma^2}{n})$ ,  $X_{n+1} \sim N(\mu, \sigma^2)$ , 则有

$$\begin{aligned} X_{n+1} - \bar{X}_n &\sim N(0, \frac{n+1}{n}\sigma^2) \\ \frac{X_{n+1} - \bar{X}_n}{\sigma\sqrt{\frac{n+1}{n}}} &\sim N(0, 1) \end{aligned}$$

又有

$$\frac{(n-1)S_n^2}{\sigma^2} \sim \chi_{n-1}^2$$

因此有

$$\frac{\frac{X_{n+1} - \bar{X}_n}{\sigma\sqrt{\frac{n+1}{n}}}}{\sqrt{\frac{(n-1)S_n^2}{\sigma^2} \cdot \frac{1}{n-1}}} = \frac{X_{n+1} - \bar{X}_n}{S_n\sqrt{\frac{n+1}{n}}} \sim t_{n-1}$$

因此  $c = \sqrt{\frac{n}{n+1}}$ ,  $t_c \sim t_{n-1}$ 。