STAT4001 Homework 2

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 $\mathbf{Q}\mathbf{1}$

(a) Recall the OLS estimation for $\hat{\beta_0}$ and $\hat{\beta_1}$:

$$\hat{\beta}_0 = \bar{y} - \hat{\beta}_1 \bar{x}$$

$$\hat{\beta}_1 = \frac{\sum_{i=1}^n (x_i - \bar{x})(y_i - \bar{y})}{\sum_{i=1}^n (x_i - \bar{x})^2} = \frac{S_{XY}}{S_{XX}}$$

We first show that $\hat{\beta}_1$ is unbiased by finding the expectation of the estimator:

$$E(\hat{\beta}_{1}) = E(\frac{\sum_{i=1}^{n} (x_{i} - \bar{x})(y_{i} - \bar{y})}{\sum_{i=1}^{n} (x_{i} - \bar{x})^{2}})$$

$$= \frac{1}{\sum_{i=1}^{n} (x_{i} - \bar{x})^{2}} \sum_{i=1}^{n} (x_{i} - \bar{x}) E(y_{i} - \bar{y})$$

$$= \frac{1}{\sum_{i=1}^{n} (x_{i} - \bar{x})^{2}} \sum_{i=1}^{n} (x_{i} - \bar{x}) E(\beta_{0} + \beta_{1} x_{i} + \varepsilon - \beta_{0} - \beta_{1} \bar{x}_{i} - 0) \text{ as sum of the error term=0}$$

$$= \frac{1}{\sum_{i=1}^{n} (x_{i} - \bar{x})^{2}} \sum_{i=1}^{n} (x_{i} - \bar{x})^{2} \beta_{1}$$

$$= \beta_{1}$$

Hence, with the result of the unbiased estimator $\hat{\beta}_1$,

$$E(\hat{\beta}_0) = E(\bar{y} - \hat{\beta}_1 \bar{x})$$

$$= E(\beta_0 + \beta_1 \bar{x} - \beta_1 \bar{x})$$

$$= \beta_0$$

from this we can also see that $\hat{\beta_0}$ is unbiased if and only if $\hat{\beta_1}$ is unbiased

(b) Recall the Ridge regression estimator, $\hat{\beta}_0$ and $\hat{\beta}_1$:

$$\hat{\beta}_0 = \bar{y} - \hat{\beta}_1 \bar{x}$$

$$\hat{\beta}_1 = \frac{\sum_{i=1}^n (x_i - \bar{x})(y_i - \bar{y})}{\sum_{i=1}^n (x_i - \bar{x})^2 + \lambda} = \frac{S_{XY}}{S_{XX} + \lambda}$$

Same with the OLS part, we start with $\hat{\beta}_1$:

$$E(\hat{\beta}_1) = E(\frac{\sum_{i=1}^n (x_i - \bar{x})(y_i - \bar{y})}{\sum_{i=1}^n (x_i - \bar{x})^2 + \lambda})$$

$$= \frac{1}{\sum_{i=1}^n (x_i - \bar{x})^2 + \lambda} \sum_{i=1}^n (x_i - \bar{x}) E(y_i - \bar{y})$$

$$= \frac{1}{\sum_{i=1}^n (x_i - \bar{x})^2 + \lambda} \sum_{i=1}^n (x_i - \bar{x})^2 \beta_1$$

$$\neq \beta_1$$

And

$$E(\hat{\beta}_0) = E(\bar{y} - \hat{\beta}_1 \bar{x})$$

$$= E(\beta_0 + \beta_1 \bar{x} - \frac{1}{\sum_{i=1}^n (x_i - \bar{x})^2 + \lambda} \sum_{i=1}^n (x_i - \bar{x})^2 \beta_1 \bar{x})$$

$$\neq \beta_0$$

Hence $\hat{\beta}_1$ is biased for ridge regression setting, and by the OLS $\hat{\beta}_0$ part, we can also conclude the $\hat{\beta}_0$ is also biased as $\hat{\beta}_1$ is biased.

$\mathbf{Q2}$

Consider the objective function $f(\beta_j) = a\beta_j^2 - 2b\beta_j + \lambda |\beta_j|$

If
$$\beta_j < 0$$
, $f'(\beta_j) = 2a\beta_j - 2b - \lambda$, set it to 0 we have $\beta_j = \frac{2b + \lambda}{2a}$

Since $\beta_j < 0$ and we are given a>0, we must have $2b + \lambda < 0$ and hence $b < -\frac{-\lambda}{2} < 0$

Now we let β_j be the minimizer for the $f(\beta_j)$,

If
$$\beta_j > 0$$
, $f(-\beta_j) = a\beta_j^2 + 2b\beta_j + \lambda |\beta_j| < f(\beta_j)$ as we know when $\beta_j < 0$, $f(\beta_j) = a\beta_j^2 - 2b\beta_j + \lambda |\beta_j| > f(-\beta_j) = a\beta_j^2 + 2b\beta_j + \lambda |\beta_j|$ as $b < -\frac{-\lambda}{2} < 0$ for $\beta_j < 0$

Hence if β_i is the minimizer for the $f(\beta_i)$, $\beta_i < 0$ must to be true.

now if $\beta_j = 0$ such that $f(\beta_j)$ minimized, we first find f(0)=0 and

$$f(\beta_j) = a\beta_j^2 - 2b\beta_j + \lambda |\beta_j|$$

$$= a\beta_j^2 - 2b\beta_j - \lambda \beta_j , \text{ as we proved } \beta_j < 0$$

$$= a\beta_j (\beta_j - \frac{2b + \lambda}{a})$$

$$f(\frac{2b + \lambda}{2a}) = a(\frac{2b + \lambda}{2a})(\frac{2b + \lambda}{2a} - \frac{2b + \lambda}{a})$$

$$= -a(\frac{2b + \lambda}{2a})^2 < f(0) = 0$$

Hence we have β_j is not 0 as 0 will not minimize objective function Hence $\beta_j < 0$ and $\hat{\beta}_j = \frac{2b + \lambda}{2a}$ minimize objective function.

Q3

(a) For OLS estimator:

Bias =
$$\beta_0 + \beta_1 - E(\hat{y_0})$$

= $\beta_0 + \beta_1 x_0 - E(\hat{\beta_0} + \hat{\beta_1} x_0)$
= $\beta_0 + \beta_1 x_0 - \beta_0 - \beta_1 x_0 = 0$ as proved in Q1
variance = $\operatorname{Var}(\hat{\beta_0} + \hat{\beta_1} x_0 + \varepsilon)$
= $\operatorname{Var}(\hat{\beta_0}) + \operatorname{Var}(\beta_1 \hat{x_0}) + 2\operatorname{Cov}(\hat{\beta_0}, \beta_1 \hat{x_0}) + \sigma^2$ since noise and x are indepentant
= $\sigma^2(\frac{1}{n} + \frac{\bar{x}^2}{\sum_{i=1}^n (x_i - \bar{x})^2}) + \frac{x_0^2 \sigma^2}{\sum_{i=1}^n (x_i - \bar{x})^2} - \frac{2\bar{x}\sigma^2}{\sum_{i=1}^n (x_i - \bar{x})} + \sigma^2$
= $\sigma^2(\frac{1}{n} + 1 + \frac{(x_0 - \bar{x})^2}{\sum_{i=1}^n (x_i - \bar{x})^2})$

(b) For the Ridge regression setting:

$$\begin{aligned} \text{Bias} &= \beta_0 + \beta_1 x_0 - E(\hat{y_0}) \\ &= \beta_0 + \beta_1 x_0 - E(\hat{\beta_0} + \hat{\beta_1} x_0) \\ &= \beta_0 + \beta_1 x_0 - \beta_0 - \beta_1 \bar{x} + \frac{1}{\sum_{i=1}^n (x_i - \bar{x})^2 + \lambda} \sum_{i=1}^n (x_i - \bar{x})^2 \beta_1 \bar{x} - \frac{x_0}{\sum_{i=1}^n (x_i - \bar{x})^2 + \lambda} \sum_{i=1}^n (x_i - \bar{x})^2 \beta_1 \text{ from Q1} \\ &= -\beta_1 (\frac{\sum_{i=1}^n (x_i - \bar{x})^2 (\bar{x} - x_0)}{\sum_{i=1}^n (x_i - \bar{x})^2 + \lambda}) \\ \text{Bias}^2 &= [\beta_1 (\frac{\sum_{i=1}^n (x_i - \bar{x})^2 (x - x_0)}{\sum_{i=1}^n (x_i - \bar{x})^2 + \lambda})]^2 > 0 \\ \text{variance} &= \text{Var}(\hat{\beta_0} + \hat{\beta_1} x_0 + \varepsilon) \\ &= \text{Var}(\hat{\beta_0}) + \text{Var}(\beta_1 \hat{x}_0) + 2\text{Cov}(\hat{\beta_0}, \beta_1 \hat{x}_0) + \sigma^2 \text{since noise and x are indepentant} \\ &= \sigma^2 (\frac{1}{n} + 1 + \frac{(x_0 - \bar{x})^2}{\sum_{i=1}^n (x_i - \bar{x})^2 + \lambda}) \end{aligned}$$

Q4

(a) Consider the Lagrange multiplier applied to the second from equation:

$$L(\beta_j, \lambda) = \sum_{l=1}^{m} (\tilde{y} - \sum_{j=1}^{p} \beta_j \tilde{x}_{lj})^2 - \lambda g(\beta_j)$$
$$= \sum_{l=1}^{m} (\tilde{y} - \sum_{j=1}^{p} \beta_j \tilde{x}_{lj})^2 + \lambda g(\beta_j) \text{ subject to g=0}$$

Which is same as the first form of the objective function. Also recall that in the Lagrange multiplier setting we max/min L subject to $g(\beta_j)=0$. So here the sign of g is not important if we subject to g to be 0. Consider the following $g(\beta_j^*)$ for some constant c:

$$g(\beta_j) = (\beta_j - c)^2$$
 subject to g=0

The square here come from the property of the ridge regression that wont force the parameters to 0. From here we can see that we must have the condition that $\beta_j \leq c$ which is because $(\beta_j - c)^2 \geq 0$ for all constant c. So we can simply discard g in the argmin statuent once we have the new condition of the beta.

Hence the function L become:

we want to retain the origin

$$L(\beta_{j}, \lambda) = \sum_{l=1}^{m} (\tilde{y} - \sum_{j=1}^{p} \beta_{j} \tilde{x_{lj}}) \beta_{j} \text{ ective function form}$$
$$= \sum_{l=1}^{m} (\tilde{y} - \sum_{j=1}^{p} \beta_{j} \tilde{x_{lj}})^{2} \text{ subject to } \beta_{j} \leq c$$

Hence two form of the objective function is same with subject to $\beta_j \leq c$. Where $(m, \tilde{y}, \tilde{x_{lj}}) = (n, y, \tilde{y_{lj}})$

(b) The objective function of this question become:

$$g(\hat{\beta}_1, \hat{\beta}_2) = \sum_{i=1}^n (y - \hat{\beta}_1 x_{1i} - \hat{\beta}_2 x_{2i})^2 - \lambda(\hat{\beta}_1 + \hat{\beta}_2)$$
$$= \sum_{i=1}^n (y - x_{1i}(\hat{\beta}_1 + \hat{\beta}_2))^2 - \lambda(\hat{\beta}_1^2 + \hat{\beta}_2^2)$$

Partial differentiate g with respect to $\hat{\beta}_1$ and $\hat{\beta}_2$ with spectively, we will have two same differential equation for $\hat{\beta_1}$ and $\hat{\beta_2}$:

$$\frac{\partial g}{\partial \hat{\beta}_1} = 2 \sum_{i=1}^n (y - x_{1i}(\hat{\beta}_1 + \hat{\beta}_2))(x_{1i}) - 2\lambda \hat{\beta}_1$$

$$\frac{\partial g}{\partial \hat{\beta}_2} = 2 \sum_{i=1}^n (y - x_{1i}(\hat{\beta}_1 + \hat{\beta}_2))(x_{1i}) - 2\lambda \hat{\beta}_1$$
two x_i are different:
x_i1 and x_i2

Which proved that $\hat{\beta}_1 = \hat{\beta}_2$

 Q_5

(a)

X=rnorm(100,0,1); noise=rnorm(100,0,0.1)

(b)

$Y=1+X+X^2+X^3+noise$

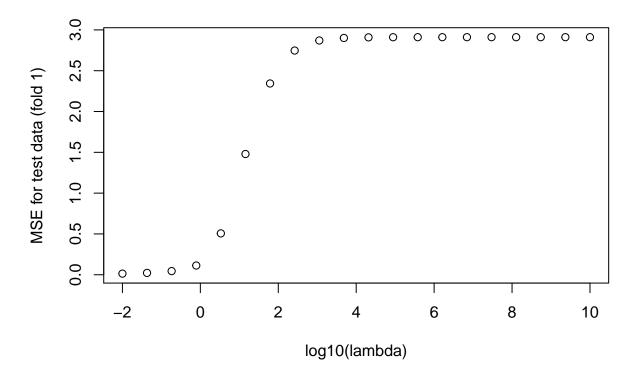
(c) We first start with setting the number of the folds and its labels, here we simply choose 5 fold.

```
set.seed(1155127616)
X_Q3=matrix(c(X,X^2,X^3,X^4,X^5,X^6,X^7,X^8,X^9,X^{10}),ncol=10,byrow=F)
Y_Q3=Y
nFold=5
cv_fold_label=sample(nFold, length(X), replace=T)
cv fold label
```

```
##
    [1] 5 2 5 3 2 1 4 5 1 5 1 1 1 4 4 4 1 5 1 2 1 4 5 4 1 3 2 5 4 3 1 3 2 5 2 5 4
   [38] 1531113321111122554254124251353413534
##
   [75] 2 3 3 2 1 3 1 1 4 2 3 1 4 5 3 5 2 4 1 5 2 2 2 2 4 3
```

Use fold 1 as test data, fold 2-5 as training data

```
library(MASS)
## Warning: package 'MASS' was built under R version 3.6.3
library(ISLR)
## Warning: package 'ISLR' was built under R version 3.6.3
library(glmnet)
## Warning: package 'glmnet' was built under R version 3.6.3
## Loading required package: Matrix
## Warning: package 'Matrix' was built under R version 3.6.2
## Loaded glmnet 4.0-2
fold=1
train_label <- which(cv_fold_label!=fold)</pre>
test_label <- which(cv_fold_label==fold)</pre>
allLambda=10^seq(10,-2,length=20)
allLambda
## [1] 1.000000e+10 2.335721e+09 5.455595e+08 1.274275e+08 2.976351e+07
## [6] 6.951928e+06 1.623777e+06 3.792690e+05 8.858668e+04 2.069138e+04
## [11] 4.832930e+03 1.128838e+03 2.636651e+02 6.158482e+01 1.438450e+01
## [16] 3.359818e+00 7.847600e-01 1.832981e-01 4.281332e-02 1.000000e-02
## sample size for test data
n_k <- length(test_label)</pre>
mse_test <- rep(NA, length(allLambda))</pre>
for (i in 1:length(allLambda)){
lambda <- allLambda[i]</pre>
ridge.mod <- glmnet(X_Q3[train_label,], Y_Q3[train_label],alpha=
ridge.pred <- predict(ridge.mod,s=lambda,newx=X_Q3[test_label,
mse_test[i] <- mean((ridge.pred-Y_Q3[test_label])^2)</pre>
}
plot(log10(allLambda), mse_test, xlab="log10(lambda)", ylab="MSE for test data (fold 1)")
```



We now Perform cross-validation

```
mse_allFold <- matrix(nrow=nFold, ncol=length(allLambda))
n_allFold <- rep(NA, nFold)
for (fold in 1:nFold){
    train_label <- which(cv_fold_label!=fold)
    test_label <- which(cv_fold_label==fold)

## sample size for test data
n_allFold[fold] <- length(test_label)
for (i in 1:length(allLambda)){
    lambda <- allLambda[i]
    ridge.mod <- glmnet(X_Q3[train_label,], Y_Q3[train_label],alpha=0_lambda=lambda)
    ridge.pred <- predict(ridge.mod,s=lambda,newx=X_Q3[test_label,])
    mse_allFold[fold, i] <- mean((ridge.pred-Y_Q3[test_label])^2)
}
n_allFold</pre>
```

[1] 27 20 17 17 19

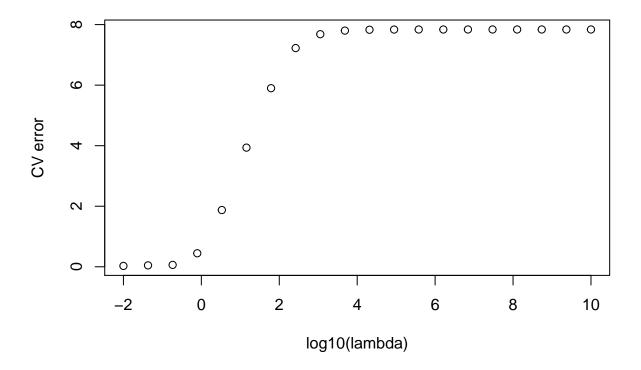
```
mse_allFold
```

```
##
             [,1]
                        [,2]
                                  [,3]
                                             [,4]
                                                        [,5]
                                                                  [,6]
                                                                             [,7]
                              2.910960
                                        2.910960
## [1,]
         2.910960
                   2.910960
                                                   2.910959
                                                             2.910954
                                                                        2.910932
## [2,]
         5.961949
                   5.961949
                              5.961949
                                        5.961948
                                                   5.961947
                                                              5.961940
                   5.863696
                              5.863696
                                        5.863695
                                                   5.863692
                                                                        5.863612
## [3,]
         5.863697
                                                             5.863677
```

```
## [4,] 15.427140 15.427140 15.427139 15.427135 15.427117 15.427041 15.426714
## [5,] 11.794624 11.794624 11.794624 11.794623 11.794621 11.794613 11.794578
##
             [,8]
                       [,9]
                                [,10]
                                          [,11]
                                                    [,12]
                                                               [,13]
        2.910837
                             2.908699
                                                 2.870287
## [1,]
                  2.910432
                                       2.901317
                                                           2.747081
                                                                     2.343606
## [2,]
        5.961781 5.961229
                             5.958869
                                       5.948801
                                                5.906247
                                                           5.733443
## [3,] 5.863334 5.862147 5.857064 5.835433
                                                5.744797
                                                          5.389875
## [4,] 15.425315 15.419328 15.393734 15.284946 14.832762 13.122760 8.627514
## [5,] 11.794427 11.793783 11.791026 11.779254 11.729345 11.524103 10.763795
##
           [,15]
                     [,16]
                               [,17]
                                          [,18]
                                                      [,19]
                                                                  [,20]
## [1,] 1.479257 0.5067345 0.1130800 0.04474375 0.022446537 0.013724742
## [2,] 3.567834 1.3707394 0.2461154 0.04490852 0.012348888 0.006075175
## [3,] 2.385451 0.8028509 0.1991804 0.05618344 0.009828093 0.004839453
## [4,] 4.662414 2.9152960 0.8615691 0.09925486 0.155020731 0.062348128
## [5,] 8.551733 4.3738901 0.9805412 0.07830419 0.049588754 0.061282999
```

Summarize and obtain cross-validation error:

```
cv_error <- matrix(n_allFold/sum(n_allFold), nrow=1)%*%mse_allFold
plot(log10(allLambda), cv_error, xlab="log10(lambda)", ylab="CV error")</pre>
```



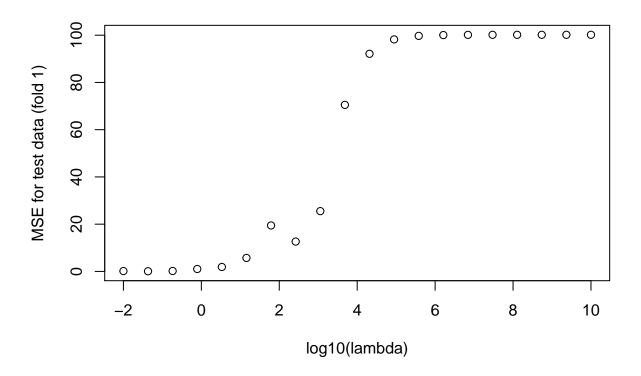
```
bestLambda <- which.min(cv_error) #lambda with lowerest CV error bestLambda
```

[1] 20

now we fit the ridge regression using this lambda

```
final ridge mod <- glmnet(X Q3, Y Q3,alpha=0,lambda=bestLambda)
final_ridge_mod$beta #coeff we want '
## 10 x 1 sparse Matrix of class "dgCMatrix"
##
## V1
        2.761663e-01
        1.259971e-02
## V2
        1.012256e-01
## V3
## V4 -6.613814e-03
       1.626607e-02
## V5
## V6 -2.150539e-03
## V7
        2.162037e-03
## V8 -4.440864e-04
## V9
        2.989636e-04
## V10 -8.098156e-05
 (d) We just follow (c) with different Y:
set.seed(1155127616)
X_Q4=matrix(c(X,X^2,X^3,X^4,X^5,X^6,X^7,X^8,X^9,X^10),ncol=10,byrow=F)
Y Q4=1+X^7+noise
nFold=5
cv_fold_label=sample(nFold, length(X), replace=T)
cv_fold_label
##
     [1] \ 5 \ 2 \ 5 \ 3 \ 2 \ 1 \ 4 \ 5 \ 1 \ 5 \ 1 \ 1 \ 1 \ 4 \ 4 \ 4 \ 1 \ 5 \ 1 \ 2 \ 1 \ 4 \ 5 \ 4 \ 1 \ 3 \ 2 \ 5 \ 4 \ 3 \ 1 \ 3 \ 2 \ 5 \ 2 \ 5 \ 4
## [38] 1 5 3 1 1 1 3 3 2 1 1 1 1 1 2 2 5 5 4 2 5 4 1 2 4 2 5 1 3 5 3 4 1 3 5 3 4
## [75] 2 3 3 2 1 3 1 1 4 2 3 1 4 5 3 5 2 4 1 5 2 2 2 2 4 3
Use fold 1 as test data, fold 2-5 as training data
library(MASS)
library(ISLR)
library(glmnet)
fold=1
train_label <- which(cv_fold_label!=fold)</pre>
test_label <- which(cv_fold_label==fold)</pre>
allLambda=10^seq(10,-2,length=20)
allLambda
  [1] 1.000000e+10 2.335721e+09 5.455595e+08 1.274275e+08 2.976351e+07
## [6] 6.951928e+06 1.623777e+06 3.792690e+05 8.858668e+04 2.069138e+04
## [11] 4.832930e+03 1.128838e+03 2.636651e+02 6.158482e+01 1.438450e+01
## [16] 3.359818e+00 7.847600e-01 1.832981e-01 4.281332e-02 1.000000e-02
## sample size for test data
n_k <- length(test_label)</pre>
mse test <- rep(NA, length(allLambda))</pre>
for (i in 1:length(allLambda)){
```

```
lambda <- allLambda[i]
ridge.mod <- glmnet(X_Q4[train_label,], Y_Q4[train_label],alpha=0,lambda=lambda)
ridge.pred <- predict(ridge.mod,s=lambda,newx=X_Q4[test_label]))
mse_test[i] <- mean((ridge.pred-Y_Q4[test_label])^2)
}
plot(log10(allLambda), mse_test, xlab="log10(lambda)", ylab="MSE for test data (fold 1)")</pre>
```



We now Perform cross-validation

```
mse_allFold <- matrix(nrow=nFold, ncol=length(allLambda))
n_allFold <- rep(NA, nFold)
for (fold in 1:nFold){
    train_label <- which(cv_fold_label!=fold)
    test_label <- which(cv_fold_label==fold)

## sample size for test data
n_allFold[fold] <- length(test_label)
for (i in 1:length(allLambda)){
    lambda <- allLambda[i]
    ridge.mod <- glmnet(X_Q4[train_label,], Y_Q4[train_label],alpha=0,lambda=lambda)
    ridge.pred <- predict(ridge.mod,s=lambda,newx=X_Q4[test_label])^2)
}
n_allFold</pre>
```

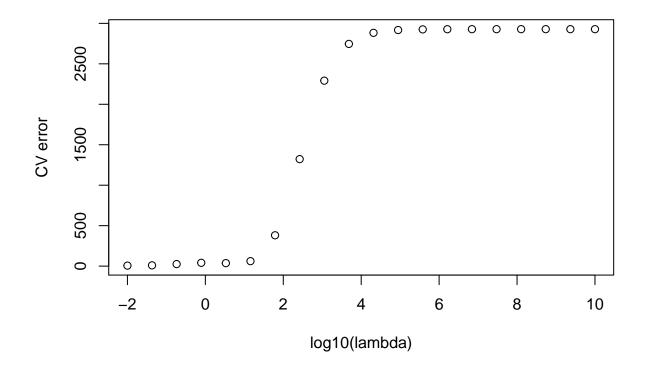
[1] 27 20 17 17 19

$mse_allFold$

```
##
               [,1]
                            [,2]
                                        [,3]
                                                     [,4]
                                                                 [,5]
                                                                              [,6]
## [1,]
          100.16485
                      100.16479
                                   100.16455
                                               100.16350
                                                            100.15900
                                                                        100.13976
## [2,]
           88.43757
                       88.43753
                                    88.43739
                                                88.43679
                                                             88.43423
                                                                         88.42325
## [3,]
         1288.68213
                    1288.68168 1288.67975
                                              1288.67149
                                                           1288.63614 1288.48482
## [4,] 15088.07532 15088.07412 15088.06899 15088.04700 15087.95289 15087.54996
## [5,]
          522.95737
                      522.95732
                                   522.95712
                                               522.95624
                                                            522.95249
                                                                        522.93642
##
                            [,8]
                                        [,9]
                                                    [,10]
                                                                [,11]
                                                                             [,12]
               [,7]
## [1,]
          100.05741
                       99.70545
                                    98.21648
                                                92.12241
                                                             70.50082
                                                                         25.52393
## [2,]
           88.37626
                       88.17540
                                    87.32454
                                                83.82631
                                                             71.16774
                                                                         42.37148
        1287.83713
## [3,]
                    1285.06742 1273.31282 1224.60198
                                                         1042.47014
                                                                        572.34333
## [4,] 15085.82506 15078.44300 15046.88961 14913.03881 14360.38663 12320.39483
## [5,]
          522.86767
                      522.57368
                                   521.32628
                                               516.16151
                                                            496.89716
                                                                         446.80610
##
             [,13]
                         [,14]
                                    [,15]
                                                [,16]
                                                           [,17]
                                                                        [,18]
## [1,]
          12.63366
                     19.44717
                                 5.704804
                                            1.878552
                                                        1.024262
                                                                   0.1793433
## [2,]
          23.99668
                     17.53961
                                 8.714659
                                            5.403025
                                                       2.505977
                                                                   0.5459149
## [3,]
         113.08893
                     23.24046
                                 6.737452
                                            2.762650
                                                       3.775238
                                                                   2.7035021
## [4,] 7184.93270 1805.29636 82.717365 49.383184 153.850001 118.0576641
  [5,]
##
         386.61846
                    323.74822 227.293204 140.372746 71.889885
                                                                  24.6493321
##
              [,19]
                          [,20]
         0.07438211
                     0.1503327
## [1,]
## [2,]
         0.06042782
                     0.1095662
## [3,]
         0.09429324
                     0.9489481
## [4,] 52.48516092 33.7615152
## [5,]
         3.83085800
                    1.7545692
```

Summarize and obtain cross-validation error:

```
cv_error <- matrix(n_allFold/sum(n_allFold), nrow=1)%*%mse_allFold
plot(log10(allLambda), cv_error, xlab="log10(lambda)", ylab="CV error")</pre>
```



```
bestLambda <- which.min(cv_error) #lambda with lowerest CV error
bestLambda</pre>
```

[1] 20

now we fit the ridge regression using this lambda

```
final_ridge_mod <- glmnet(X_Q4, Y_Q4,alpha=0,lambda=bestLambda)
final_ridge_mod$beta #coeff we want</pre>
```

```
## 10 x 1 sparse Matrix of class "dgCMatrix"
##
## V1
       -0.132740185
## V2
        1.277092476
        2.544079164
  VЗ
       -0.479377235
## V4
        0.888058141
## V5
## V6
       -0.215256504
        0.174522535
       -0.050553288
## V8
        0.029692561
## V9
## V10 -0.009736758
```