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$$\tilde{f} \sim R(0, \theta) \quad \theta > 0$$

вер. модель

1) \tilde{x}_n выборка

$$\tilde{\theta}_1 = 2\bar{x} = 2 \frac{1}{n} \sum_{i=1}^n x_i$$

$$M \tilde{\theta}_1 = M \left[\frac{2}{n} \sum x_i \right] = \frac{2}{n} \sum M x_i = \frac{2}{n} n \frac{\theta}{2} = \theta$$

$$M \tilde{f} = \frac{\theta}{2}$$

не смещ.

$$D \tilde{\theta}_1 = D \frac{2}{n} \sum x_i = \frac{4}{n^2} \sum D x_i = \frac{4}{n^2} n D \tilde{f} = \frac{4}{n} \frac{\theta^2}{12} = \frac{\theta^2}{3n} \rightarrow 0$$

асмпт.

$$2) \tilde{\theta}_2 = x_{\min}$$

$$M \tilde{\theta}_2 = M [x_{\min}]$$

$$\min(\tilde{f}_1, \dots, \tilde{f}_n) \sim 1 - (1 - F(x))^n$$

$$P(y) = n(1 - F(y))^{n-1} f(y)$$

$$= n \left(1 - \frac{y}{\theta}\right)^{n-1} \frac{1}{\theta} (0, \theta)$$

$$M [x_{\min}] = \int_0^\theta y n \frac{1}{\theta} \left(1 - \frac{y}{\theta}\right)^{n-1} dy = \langle t = \frac{y}{\theta} \rangle =$$

$$= \int_0^1 t \theta n (1-t)^{n-1} dt = n \theta B(2, n) =$$

$$= n \theta \frac{\Gamma(2) \Gamma(n)}{\Gamma(n+2)} = n \theta \frac{\Gamma(2) \Gamma(n)}{(n+1)n \Gamma(n)} = \frac{\theta}{n+1}$$

смещен.

не асимпт. не смещ.

$$B(p, q) = \int_0^1 x^{p-1} (1-x)^{q-1} dx$$

Аналогично: $\tilde{\theta}_2' = (n+1) \kappa_{\min}$ - несмещ

$$M[\kappa_{\min}^2] = \int_0^{\theta} y^2 n \left(1 - \frac{y}{\theta}\right)^{n-1} \frac{1}{\theta} dy =$$

$$= \int_0^1 \theta^2 t^2 n (1-t)^{n-1} dt = \theta^2 n B(3, n) =$$

$$= \frac{2\theta^2}{(n+2)(n+1)}$$

$$D[\kappa_{\min}] = \frac{2\theta^2}{(n+2)(n+1)} - \frac{\theta^2}{(n+1)^2} = \frac{\theta^2 [(n+1)^2 - 2(n+1)]}{(n+1)^2 (n+2)} =$$

$$= \frac{\theta^2 n}{(n+1)^2 (n+2)}$$

$$D[\theta_2'] = D[(n+1) \kappa_{\min}] = \frac{\theta^2 n}{n+2} \xrightarrow[n \rightarrow \infty]{} 0$$

проверка: $\text{опр. } \tilde{\theta}_2' \xrightarrow{P} \theta$

$$\theta_2' = (n+1) \kappa_{\min}; \quad \forall \varepsilon > 0 \quad P(|\theta_2' - \theta| \geq \varepsilon) \xrightarrow[n \rightarrow \infty]{} 0$$

$$P(|\theta_2' - \theta| \geq \varepsilon) \geq P(\theta_2' \geq \theta + \varepsilon) = P$$

$$= P\left(\kappa_{\min} \geq \frac{\theta + \varepsilon}{n+1}\right) = P(\kappa_1 \geq \dots, \kappa_2 \geq \dots, \kappa_n \geq \dots) =$$

$$= \prod_{i=1}^n P\left(\kappa_i \geq \frac{\theta + \varepsilon}{n+1}\right) = \left(1 - F\left(\frac{\theta + \varepsilon}{n+1}\right)\right)^n =$$

$$= \left(1 - \frac{\theta + \varepsilon}{\theta(n+1)}\right)^n \xrightarrow[n \rightarrow \infty]{} e^{-\frac{\theta + \varepsilon}{\theta}} > 0$$

неполн.

$$3) \tilde{\theta}_3 = x_{\max}; \quad x_i = F(y), \quad x_{\max} = (F(y))^n$$

$$f_{x_{\max}}(y) = n(F(y))^{n-1} f(y) \quad (0, \theta)$$

$$M[\tilde{\theta}_3] = \int_0^\theta y^n \left(\frac{y}{\theta}\right)^{n-1} \frac{1}{\theta} dy = \frac{n}{\theta^n} \int_0^\theta y^n dy = \frac{n\theta}{n+1}$$

liney., acc. nececy.

$$\theta_3' = \frac{n+1}{n} x_{\max} - \text{nececy.}$$

$$M[\tilde{\theta}_3^2] = \int_0^\theta y^2 n \left(\frac{y}{\theta}\right)^{n-1} \frac{1}{\theta} dy = \frac{n}{\theta^n} \int_0^\theta y^{n+1} dy = \frac{\theta^2 n}{n+2}$$

$$D[\tilde{\theta}_3] = \frac{\theta^2 n}{n+2} - \frac{n^2 \theta^2}{(n+1)^2} = \theta^2 \frac{n}{(n+2)(n+1)^2}$$

$$D[\theta_3'] = \frac{(n+1)^2}{n^2} \frac{n \theta^2}{(n+2)(n+1)^2} = \frac{\theta^2}{n(n+2)} \xrightarrow{n \rightarrow \infty} 0$$

como x max

$$4) \tilde{\theta}_4 = x_{\min} + x_{\max}$$

$$M[\tilde{\theta}_4] = M[x_{\min}] + M[x_{\max}] = \frac{\theta}{n+1} + \frac{n\theta}{n+1} = \theta$$

$$D[\tilde{\theta}_4] = D[x_{\min}] + D[x_{\max}] + 2 \text{cov}(x_{\max}, x_{\min})$$

$$\text{cov}(x_{\max}, x_{\min}) = M[x_{\min} x_{\max}] - M[x_{\min}] \cdot M[x_{\max}]$$

$$p(y, z) = n(n-1) (F(z) - F(y))^{n-2} p(z) p(y)$$

$$M[x_{\min}, x_{\max}] = \iint yz \cdot n(n-1) \left(\frac{z}{\theta} - \frac{y}{\theta}\right)^{n-2}$$

$$\frac{1}{\theta} \frac{1}{\theta} dz dy = \int_0^\theta dz \int_0^z yz (n-1)n \left(\frac{z}{\theta} - \frac{y}{\theta}\right)^{n-2}$$

$$\frac{1}{\theta^2} dy = \left\langle t = \frac{y}{z} \right\rangle = \frac{n(n-1)}{\theta^n} \int_0^\theta dz \int_0^1 t z^2$$

$$(z - tz)^{n-2} dt = \frac{n(n-1)}{\theta^n} \int_0^\theta z^{n+1} dz \int_0^1 t (1-t)^{n-2} dt$$

$$= \frac{n(n-1)}{\theta^n} \int_0^\theta z^{n+1} \frac{1}{n(n-1)} dz = \frac{\theta^{n+2}}{(n+2)\theta^n} = \frac{\theta^2}{n+2}$$

$$\text{COV}(x_{\min}, x_{\max}) = \frac{\theta^2}{n+2} - \frac{\theta}{n+1} \cdot \frac{n\theta}{n+1} = \frac{\theta^2}{(n+2)(n+1)}$$

$$D[\tilde{\theta}_n] = \frac{\theta^2 n}{(n+1)^2 (n+2)} + \frac{\theta^2 n}{(n+1)^2 (n+2)} + \frac{2\theta^2}{(n+1)^2 (n+2)}$$

$$= \frac{2\theta^2}{(n+2)(n+1)} \xrightarrow{h \rightarrow \infty} 0 \quad \text{Cochran's theorem}$$

5) ~~$M[\tilde{\theta}_5]$~~ $\tilde{\theta}_5 = x_1 + \frac{\sum_{k=2}^n x_k}{(n-1)}$

$$M[\tilde{\theta}_5] = M[x_1] + \frac{n}{n-1} + \frac{n}{n-1} M[x_2]$$

$$z = \frac{\theta}{2} \left(1 + \frac{n-1}{n-1} \right) = \theta$$

$$x_i \sim R(0; \theta)$$

$$D[\tilde{\theta}_5] = D[x_1] + \frac{n-1}{(n-1)^2} D[x_2] = \frac{\theta^2}{12} \left(1 + \frac{1}{n-1} \right) \xrightarrow{n \rightarrow \infty} 0$$

$$x_1 + \frac{1}{n-1} \sum_{m=2}^n x_m \xrightarrow{P} x_1 + \frac{\theta}{2} \text{ — не сбл. закон.}$$

Формулы в скобках: $\theta_2, \theta_3, \theta_5$ — базирующиеся
 и смещен $\tilde{\theta}_1, \tilde{\theta}_3, \tilde{\theta}_4$: $D\tilde{\theta}_1 = \frac{\theta^2}{3n}$, $D\tilde{\theta}_3' = \frac{\theta^2}{n(n+2)}$
 $D\tilde{\theta}_4 = \frac{2\theta^2}{(n+1)(n+2)}$

$$\frac{\theta^2}{3n} > \frac{\theta^2}{n(n+2)} \text{ для } n > 1$$

$$\frac{\theta^2}{n(n+2)} < \frac{2\theta^2}{(n+1)(n+2)} \text{ для } n > 1 \quad \tilde{\theta}_3' \text{ лучше}$$

1/2
 а) $p = \begin{cases} e^{-x}, & x \geq 0 \\ 0, & x < 0 \end{cases}$

$$\gamma = \frac{\tilde{\mu}_3}{\tilde{\mu}_2^{3/2}}$$

$$L_1 = \int_0^{\infty} x e^{-x} dx = -x e^{-x} \Big|_0^{\infty} + \int_0^{\infty} e^{-x} dx = -e^{-x} \Big|_0^{\infty} = 1$$

$$\mu_2 = \int_0^{\infty} (x-1)^2 e^{-x} dx = \int_0^{\infty} x^2 e^{-x} dx - 2 \int_0^{\infty} x e^{-x} dx + \int_0^{\infty} e^{-x} dx = 2 - 2 + 1 = 1$$

$$\mu_3 = \int_0^{\infty} (x-1)^3 e^{-x} dx = \int_0^{\infty} x^3 e^{-x} dx - 3 \int_0^{\infty} x^2 e^{-x} dx + 3 \int_0^{\infty} x e^{-x} dx - \int_0^{\infty} e^{-x} dx = 6 - 3 \cdot 2 + 3 \cdot 1 - 1 = 2$$

$$f = \frac{\mu_3}{(\mu_2)^{3/2}} = 2$$

г) Проверка гипотезы о нормальности

$$h = 2,344 \frac{s}{\sqrt{n}}; \quad s^2 = \frac{1}{n-1} \sum_{i=1}^n (x_i - \bar{x})^2$$

$$\mu_2 = \frac{1}{n} \sum (x_i - \bar{x})^2; \quad \mu_2 n = \sum_{i=1}^n (x_i - \bar{x})^2$$

$$s^2 = \frac{\mu_2 n}{n-1} \quad s = \sqrt{\frac{\mu_2 n}{n-1}}$$

$$\hat{p}(z) = \frac{1}{nh} \sum_{i=1}^n q\left(\frac{z - x_i}{h}\right) = \frac{1}{nh} \sum_{i=1}^n \left(\frac{3}{4} \left(1 - \left(\frac{z - x_i}{h}\right)^2\right)\right)$$

$$\hat{p}(z) = \frac{3}{4nh} \sum_{i=1}^n \left(1 - \left(\frac{z - x_i}{h}\right)^2\right)$$

д) У ПТ: $\sqrt{n} \frac{\tilde{Z}_n - Z_n}{\sqrt{\tilde{Z}_n - Z_n^2}} \sim N(0,1)$

$$\tilde{Z}_1 = \frac{1}{n} \sum_{i=1}^n x_i \quad \tilde{Z}_2 = \frac{1}{n} \sum_{i=1}^n x_i^2$$

$$L_1 = \int_0^{+\infty} x e^{-x} dx \quad (\bar{L}_1 - L_1) \sim \frac{\sqrt{\bar{L}_1 - L_1^2}}{\sqrt{n}} N(0, 1)$$

$$\bar{L}_1 - L_1 \rightsquigarrow N(0, \frac{\sqrt{\bar{L}_1 - L_1^2}}{n}); \quad L_1 \rightsquigarrow N(\bar{L}_1, \frac{\sqrt{\bar{L}_1 - L_1^2}}{n})$$

$$N_3$$

$$p(x) = \begin{cases} \frac{e^{-\frac{x}{\theta}}}{\theta}, & x \geq 0 \\ 0, & x < 0 \end{cases}, \quad \theta > 0$$

$$n=3 \quad \bar{K}_n = (K_1, \dots, K_n) \quad F(x) = 1 - e^{-\frac{x}{\theta}}$$

$$\bar{K}_3 = (K_1, K_2, K_3)$$

$$\tilde{\theta}_1 = \bar{K}_1$$

$$\tilde{\theta}_1 = \bar{K}_1$$

$$\tilde{\theta}_2 = \frac{K_{\min} + K_{\max}}{2}$$

$$\tilde{\theta}_3 = K_{(1)}$$

$$Z \sim p(x)$$

$$M[Z] = \int_0^{\infty} x \cdot \frac{e^{-\frac{x}{\theta}}}{\theta} dx = \frac{1}{\theta} \int_0^{\infty} x e^{-\frac{x}{\theta}} dx = \frac{1}{\theta} \cdot \theta^2 = \theta$$

$$M[Z^2] = \frac{1}{\theta} \int_0^{\infty} x^2 e^{-\frac{x}{\theta}} dx = \frac{1}{\theta} \cdot 2 \cdot \theta^3 = 2\theta^2$$

$$D[Z] = M[Z^2] - M[Z]^2 = \theta^2$$

$$a) M[\tilde{\theta}_1] = M\left[\frac{1}{n} \sum K_i\right] = \frac{1}{n} M[\sum K_i] = \frac{1}{n} \cdot n M[Z] =$$

$$= \theta$$

нел. метод

$$\mathcal{M}[\tilde{\Theta}_1]: \tilde{\Theta}_3 = \kappa_{(2)}$$

$$\tilde{\kappa}_n \sim n \binom{n-1}{n-1} \frac{e^{-\frac{\kappa}{\theta}}}{\theta} (1 - e^{-\frac{\kappa}{\theta}})^{n-\kappa}$$

$$\begin{aligned} k \geq 2 \quad \tilde{\kappa}_{(2)} &\sim n \binom{n-1}{n-1} \frac{e^{-\frac{\kappa}{\theta}}}{\theta} (1 - e^{-\frac{\kappa}{\theta}})^1 (e^{-\frac{\kappa}{\theta}})^{n-2} = \\ &= \frac{n(n-1)}{\theta} (1 - e^{-\frac{\kappa}{\theta}}) (e^{-\frac{\kappa}{\theta}})^{n-1} \end{aligned}$$

$$\begin{aligned} \mathcal{M}[\tilde{\Theta}_3] &= \frac{n(n-1)}{\theta} \int_0^\infty \kappa (1 - e^{-\frac{\kappa}{\theta}}) e^{-\frac{\kappa}{\theta}(n-1)} d\kappa = \\ &= \frac{n(n-1)}{\theta} \left(\int_0^\infty \kappa e^{-\frac{\kappa}{\theta}(n-1)} d\kappa + \int_0^\infty \kappa e^{-\frac{\kappa}{\theta}n} d\kappa \right) = \end{aligned}$$

$$= \frac{n(n-1)}{\theta} \left(\frac{\theta^2}{(n-1)^2} - \frac{\theta^2}{n^2} \right) = \frac{n\theta}{n-1} - \frac{(n-1)\theta}{n} = \frac{2n-1}{n(n-1)} \theta$$

check.

$$\tilde{\Theta}_3' = \frac{n(n-1)}{2n-1} \kappa_{(2)} - \text{неверно}$$