# Conceptualizing and building Markov models (with emphasis on rates and probabilities)

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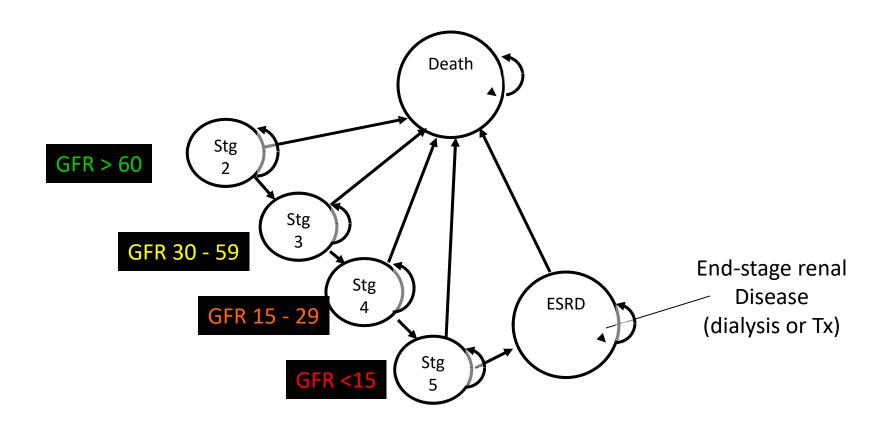
- Enumerate possible health states
- From each state consider possible stochastic events that could lead to transitions
  - Construct the cycle trees
- Populate model parameters
  - Incremental utilities
  - Costs
  - Transition probabilities

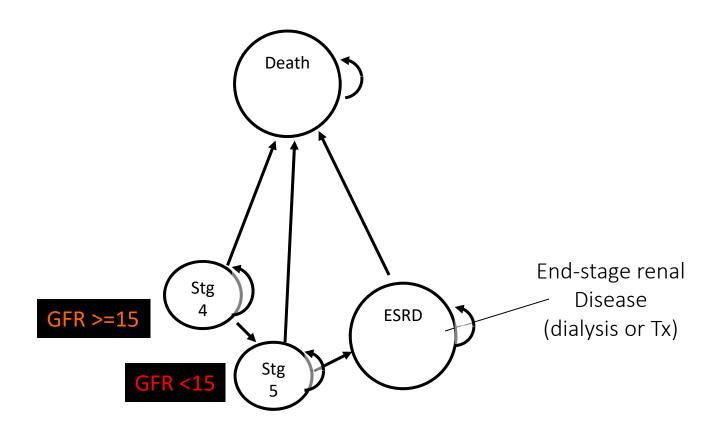
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- Enumerate possible health states
- From each state consider possible stochastic events that could lead to transitions
  - Construct the cycle trees
- Populate model parameters
- Debug and validate

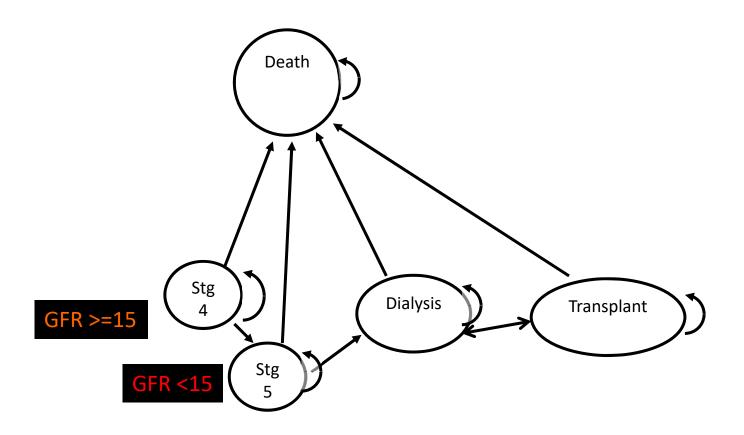
#### Example Markov conceptualization

- The objective is to model a cohort with chronic kidney disease (CKD)
- Severity of CKD is defined by the value of the glomerular filtration rate (GFR)
- The lower the GFR, the worse the kidney function
- Normal ~ 100 ml/min, at < 10 we consider dialysis</li>
- CKD progresses through stages:

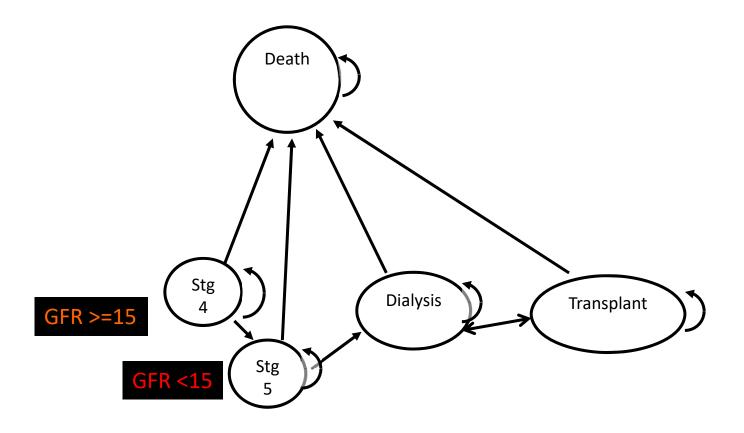




For now, let's consider only late stage CKD: Stg4, stg5, ESRD and death

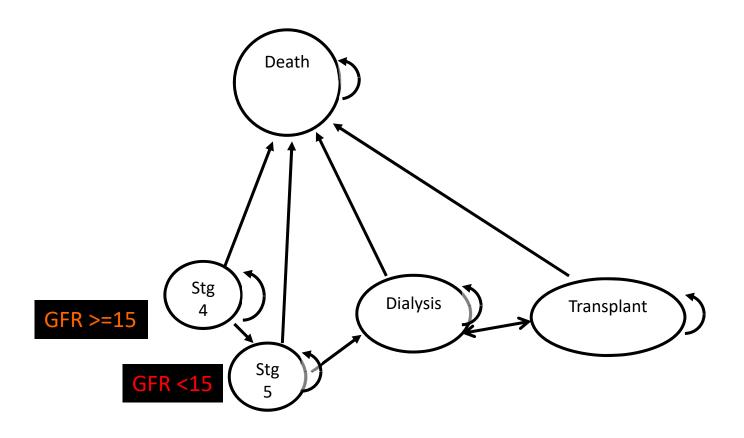


Let's split ESRD into dialysis and renal transplant states. What does this state structure and allowed transitions imply?

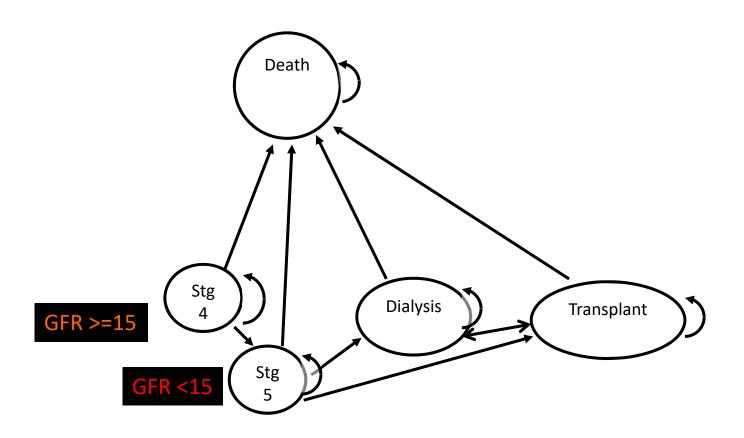


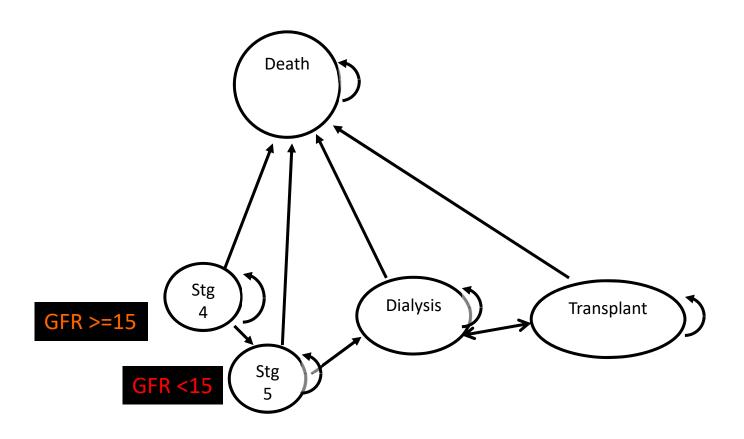
Let's split ESRD into dialysis and renal transplant states. What does this state structure and allowed transitions imply?

This is an unfair question for non-nephrologists

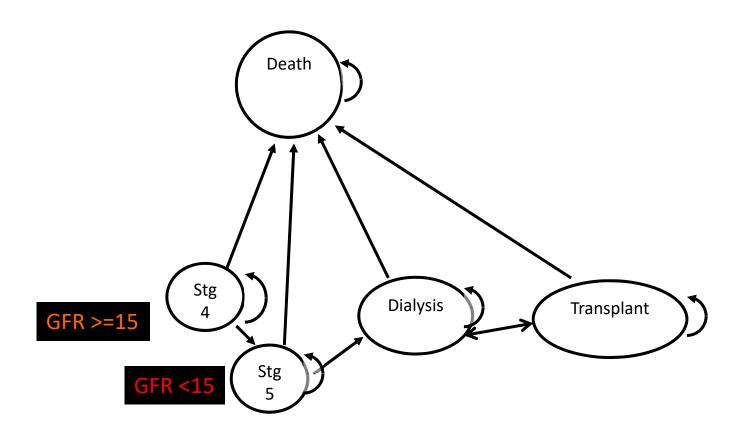


This structure does not allow for pre-emptive transplantation. How could it be adjusted to do so?

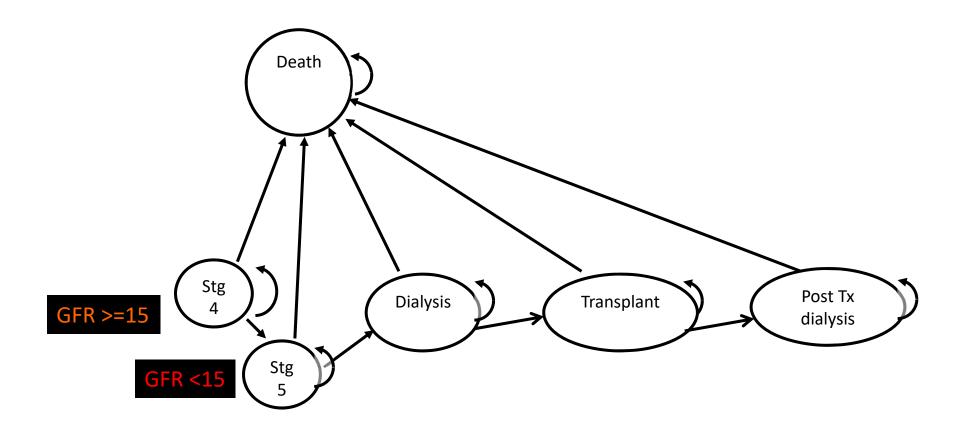




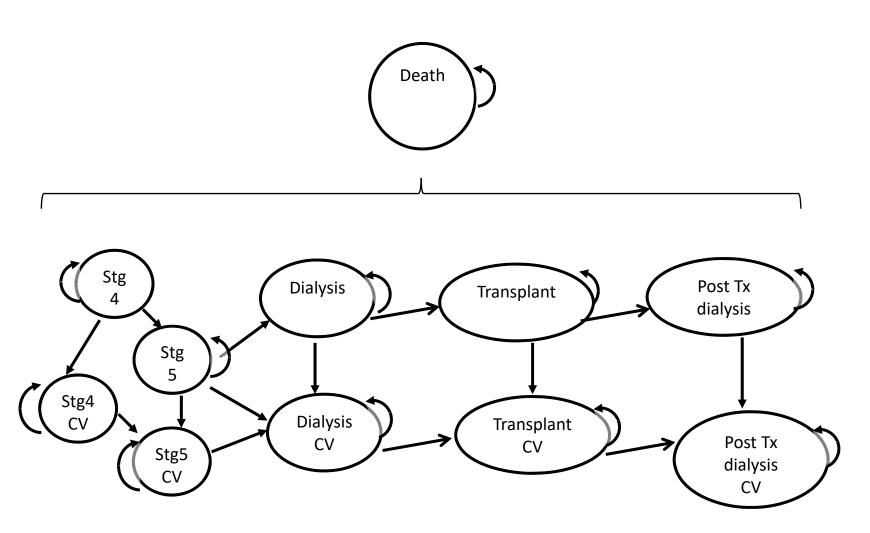
For now, let's NOT consider pre-emptive transplantation. But....., how many transplants could happen?



How should this be structured so that only ONE transplant can happen? (and respects the Markovian assumption)



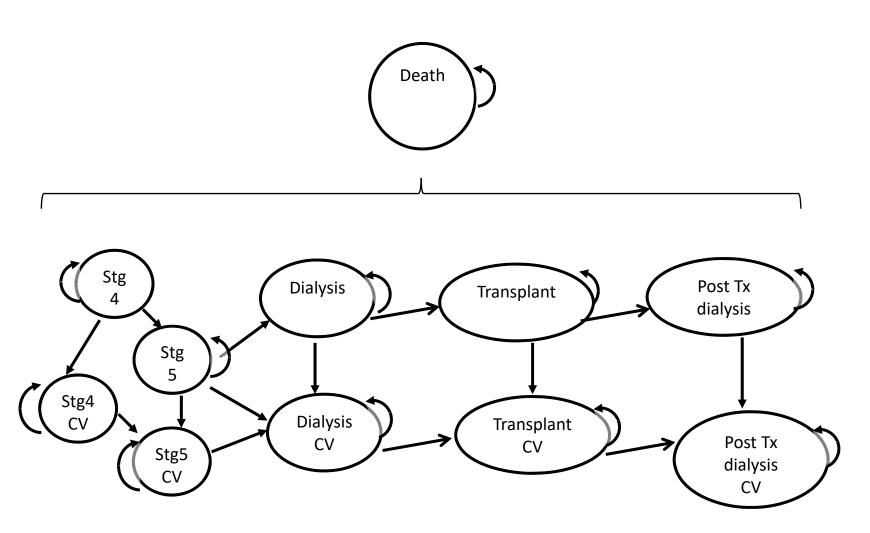
How should this be structured so that cardiovascular events are considered? (and respects the Markovian assumption)



Structured so that cardiovascular events are considered (respecting the Markovian assumption)

In general, when are additional states required?

- When costs or incremental utilities differ
- When subsequent disease evolution differs
- In order to 'remember' prior history



Structured so that cardiovascular events are considered (respecting the Markovian assumption)

# Are additional states required for CV events?

- Do costs or incremental utilities differ?
  - Yes
- Does subsequent disease evolution differ?
  - Yes
- Do we 'remember' prior history?
  - Yes, being in a 'X CV' state implies that there must have been a prior CV event

#### Distinguishing between states and shortterm, transient events

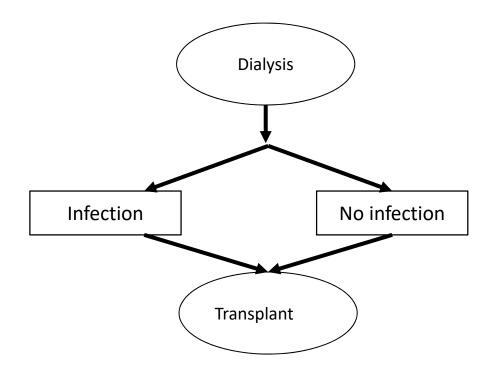
- Renal transplant recipients may have an infection immediately after the transplant procedure lasting 1 – 2 days (cycle length = 1 month)
- Imagine that the latter does not effect the ultimate function of the transplant
- Should infection be modeled as a state?

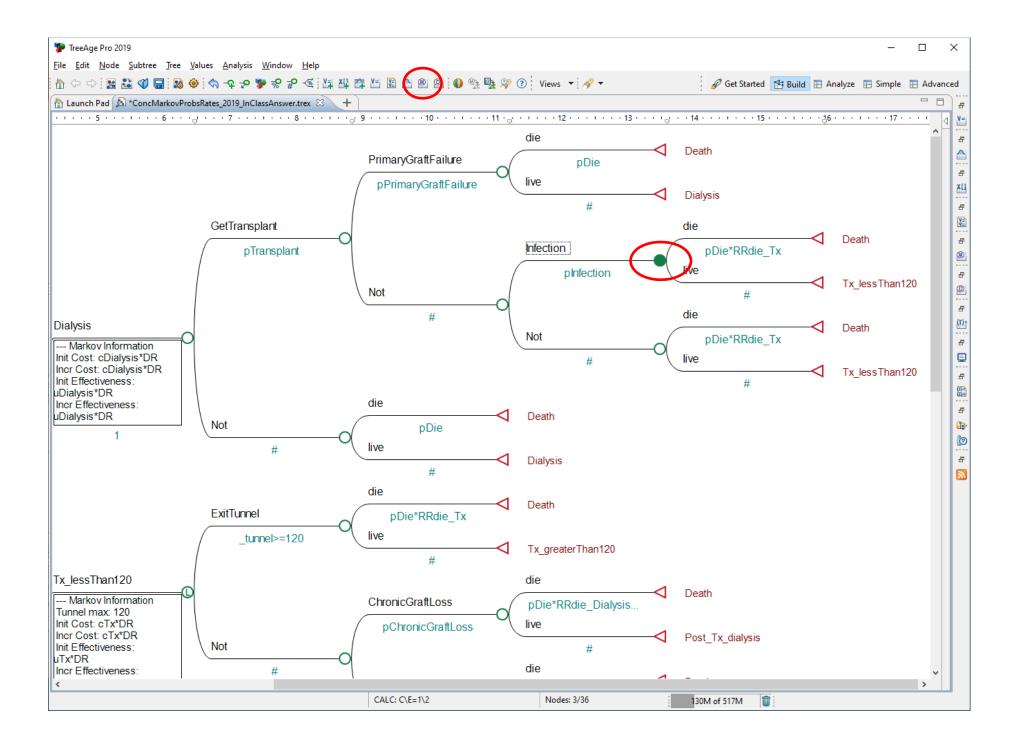
# Are additional states required for post Tx infection?

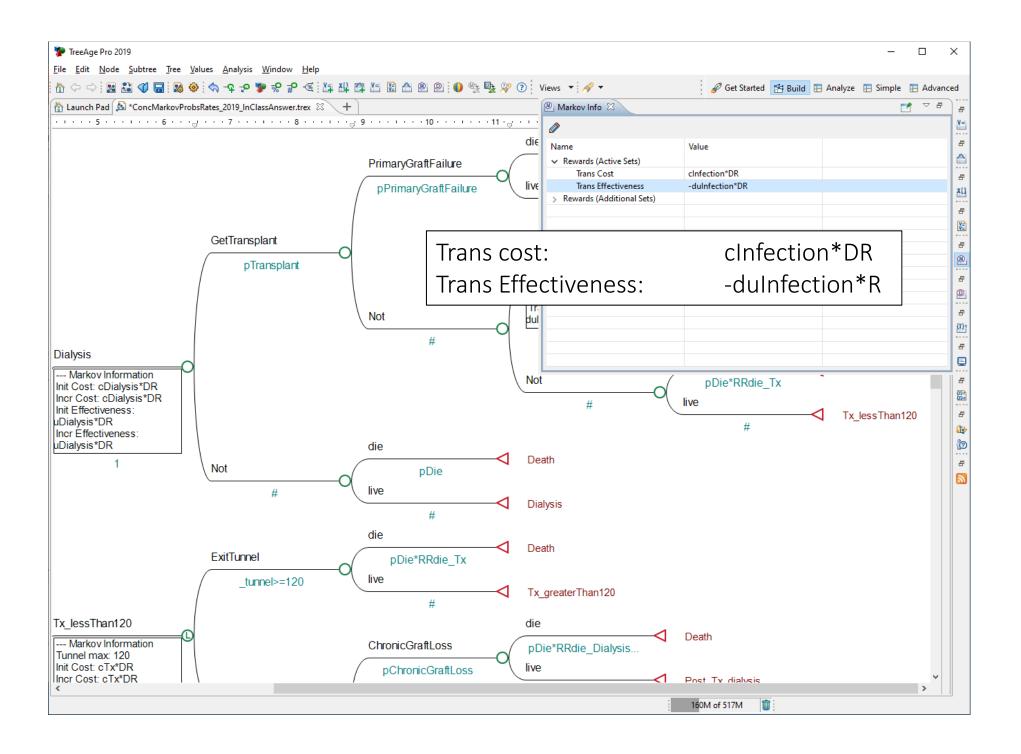
- Do costs or incremental utilities differ?
  - Yes, but only transiently, not permanently
- Does subsequent disease evolution differ?
  - No
- Do we need to 'remember' prior history?
  - No

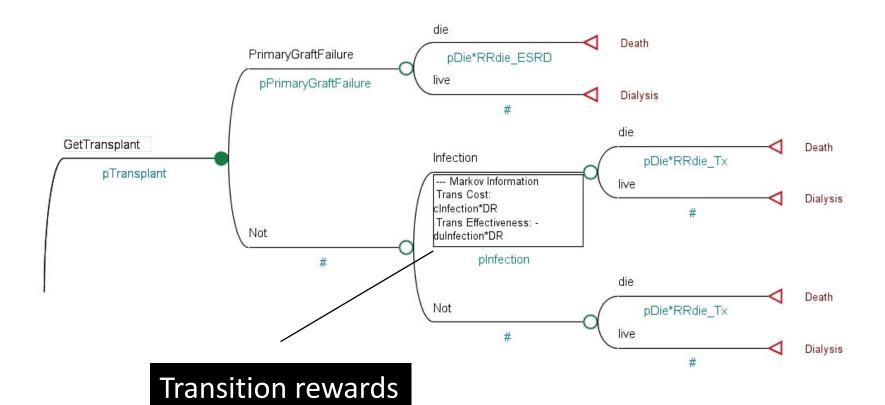
#### Transient events

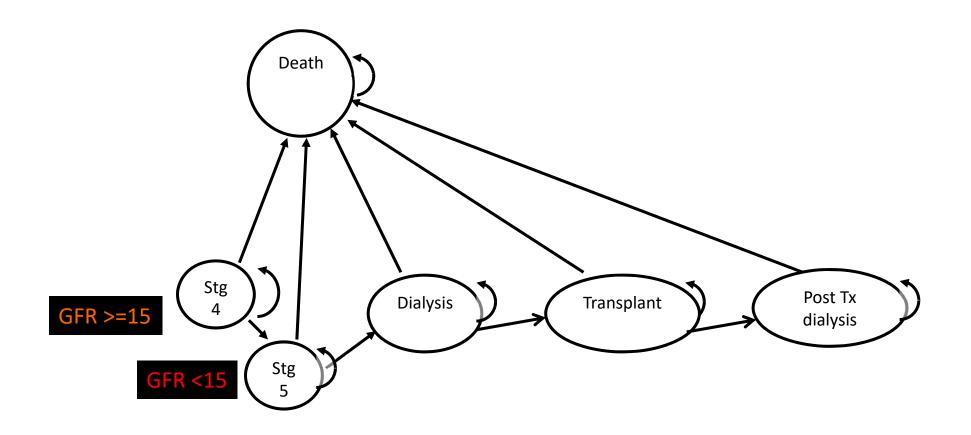
- If
  - Final and end states identical
  - Duration < 1 cycle</li>length
- Then
  - Capture effects as transition rewards (tolls)







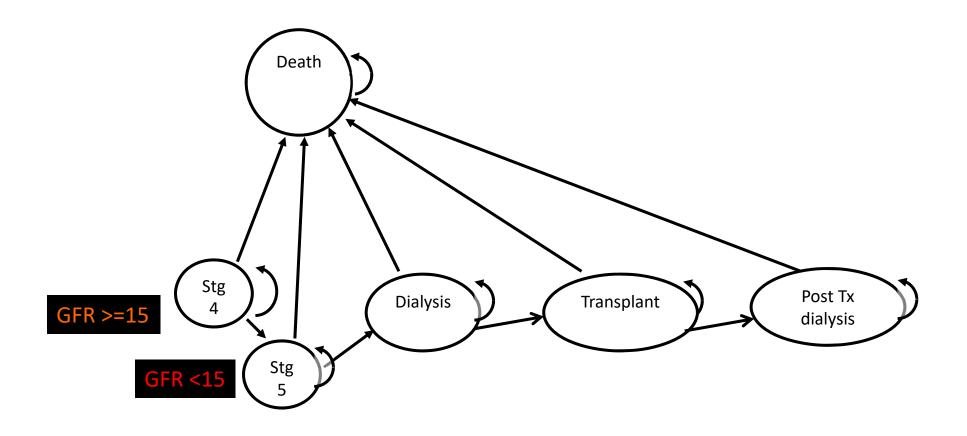




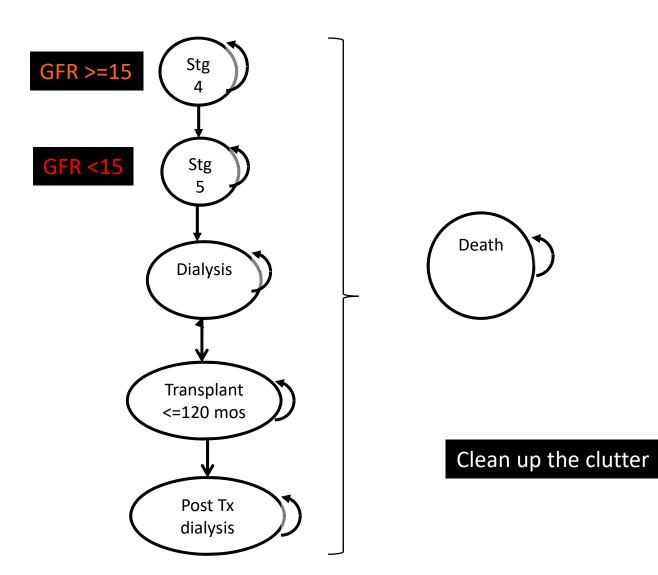
Our state structure and allowed transitions thus far.

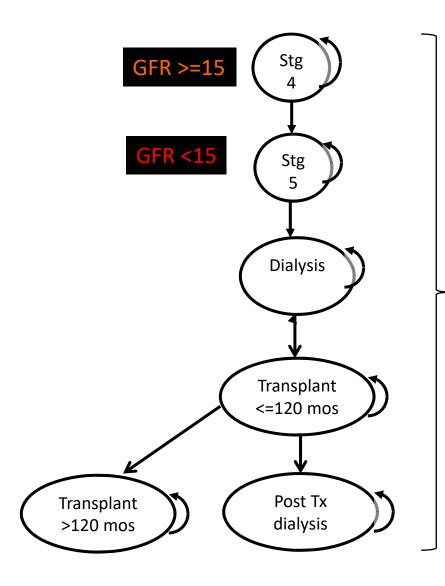
For simplicity,

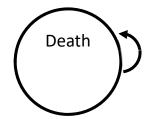
let's ignore cardiovascular events



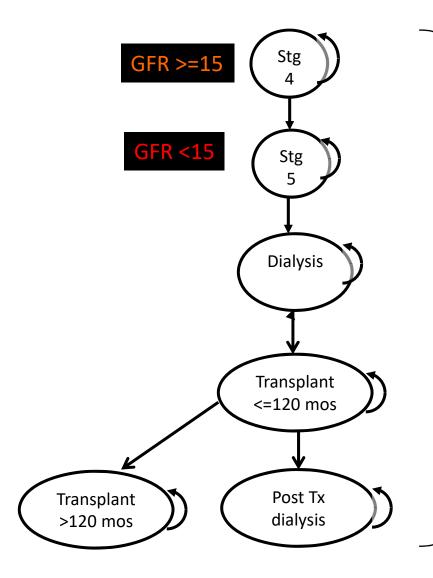
Suppose there is an early phase, post transplant, where there is a possibility of transplant failure, and a late phase where failure is impossible

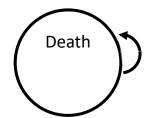






Let's imagine that the risk of transplant failure stops after 10 years

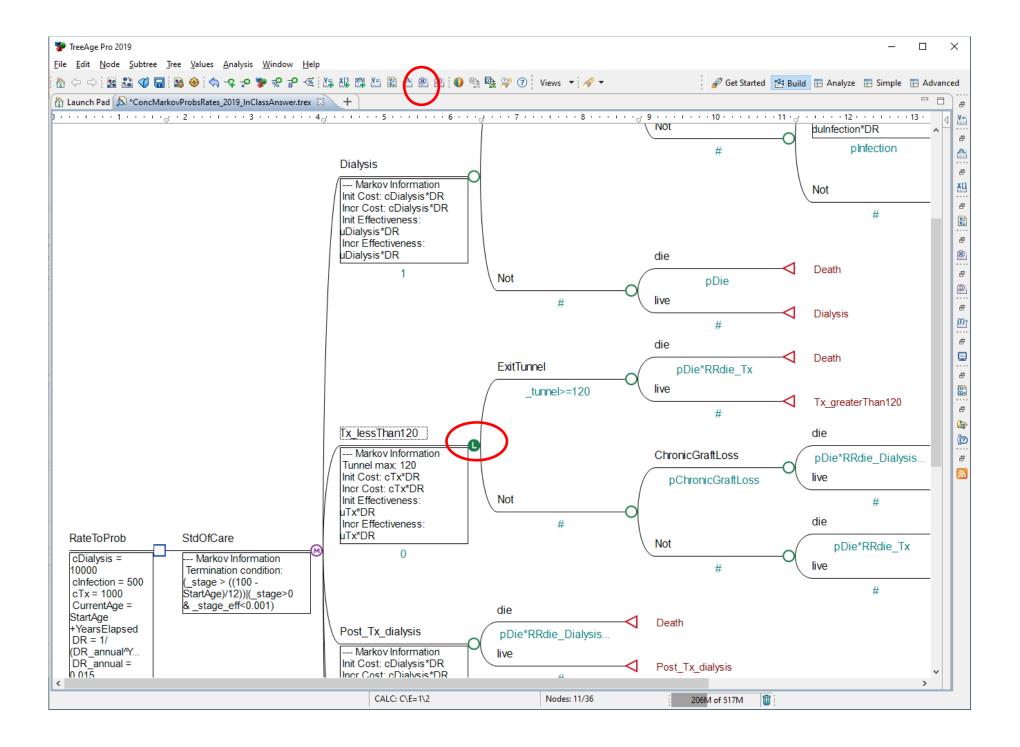


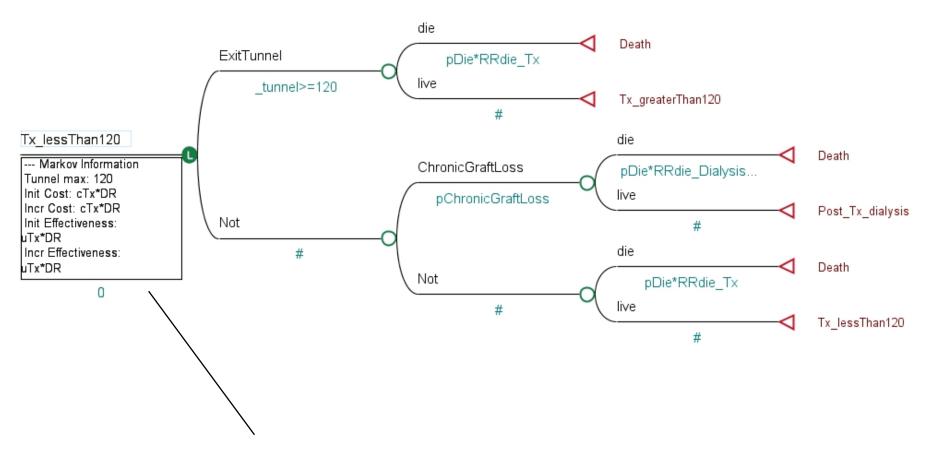


Let's imagine that the risk of transplant failure stops after 10 years .....but this is 10 years after transplant, NOT since the beginning of the process

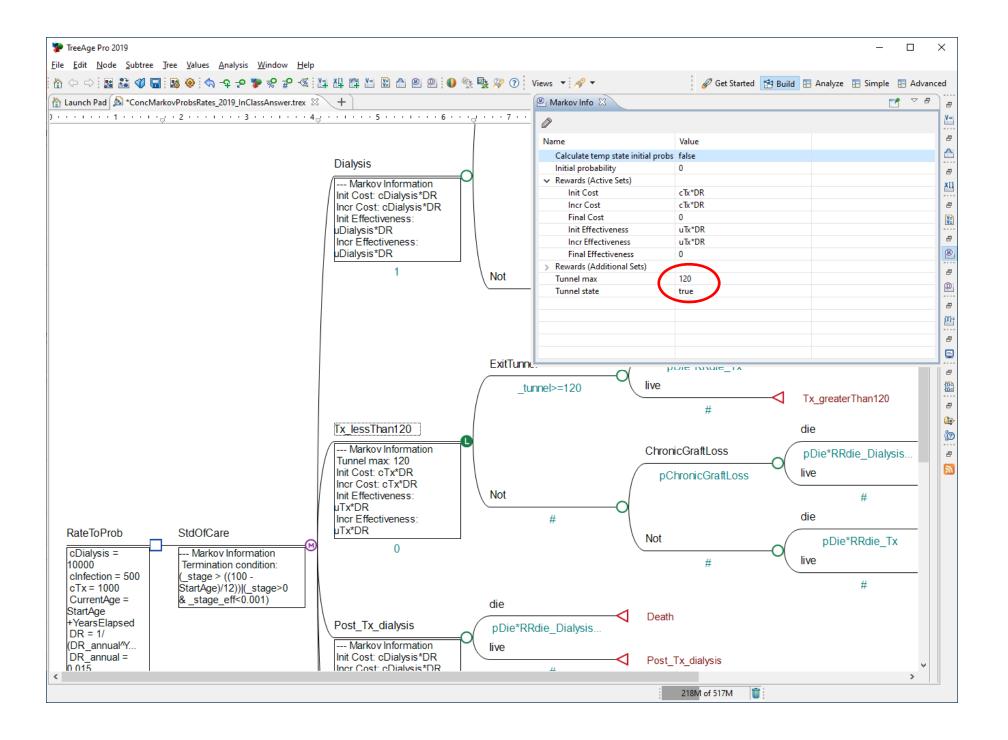
#### The "two clock" problem

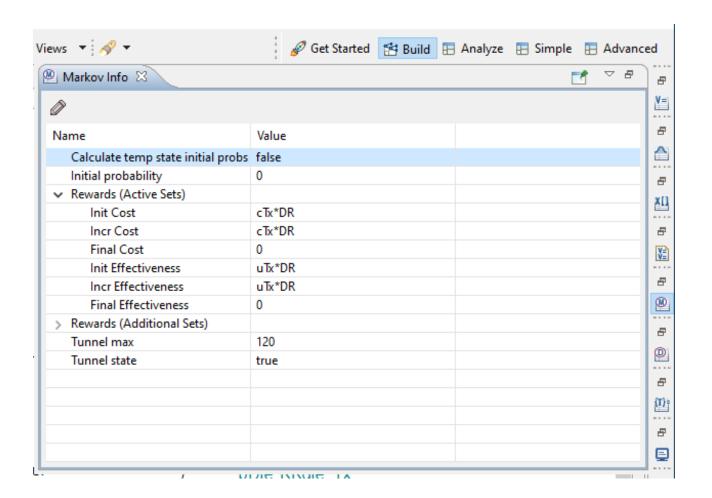
- Because of the Markovian assumption there is no explicit memory of when past events occured
- But, often want to base the probability of an event on the time since the occurrence of an interim event, not since the beginning of the process
- E.g. the probability of chronic rejection and transplant failure is a function of time since transplant, not since the start of the simulation
- Use tunnel states and the \_tunnel counter

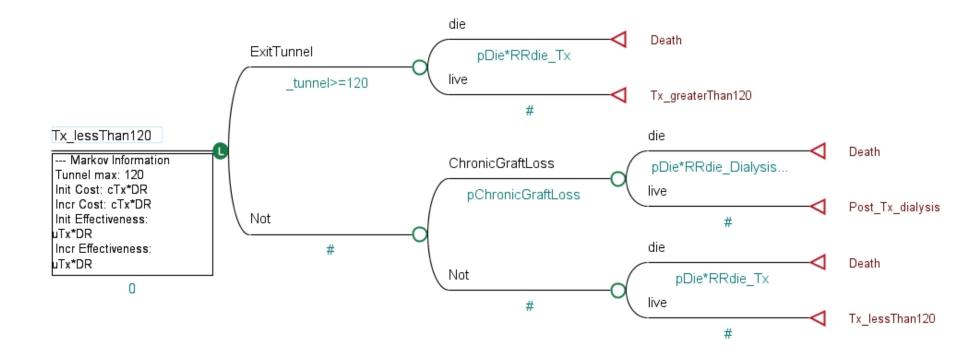




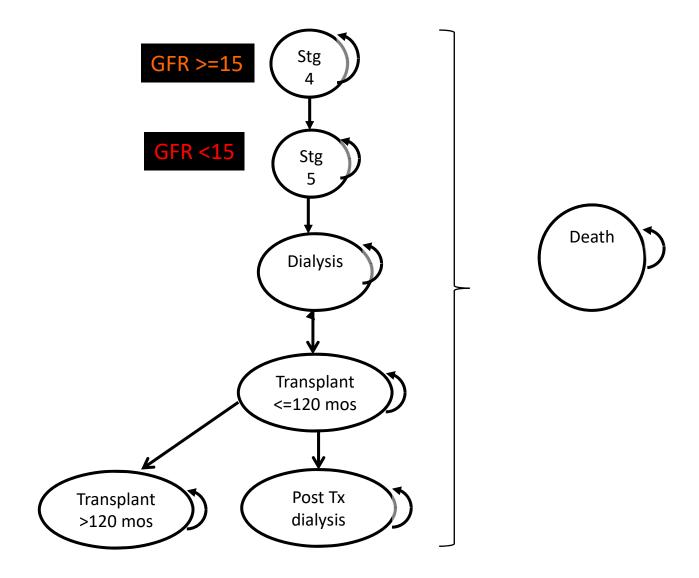
Set the max tunnel value to a number > 0 in the Markov dialogue box

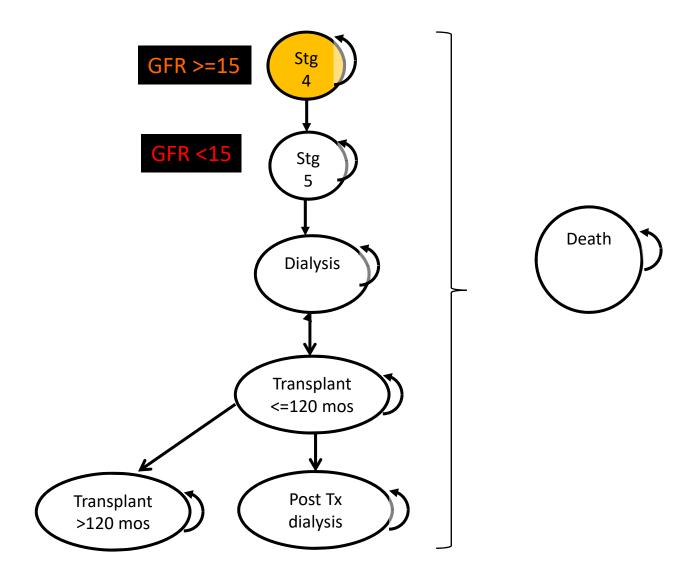


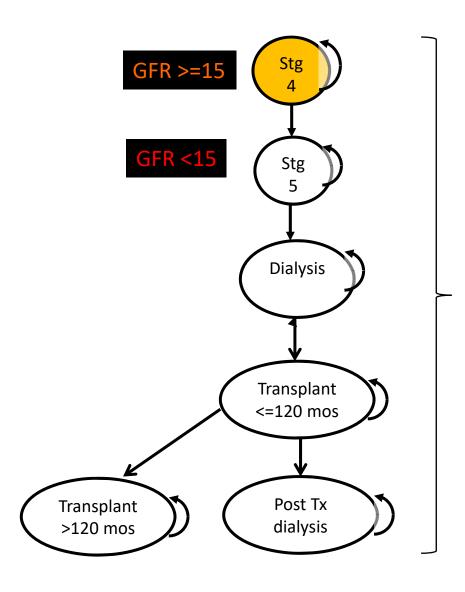




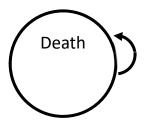
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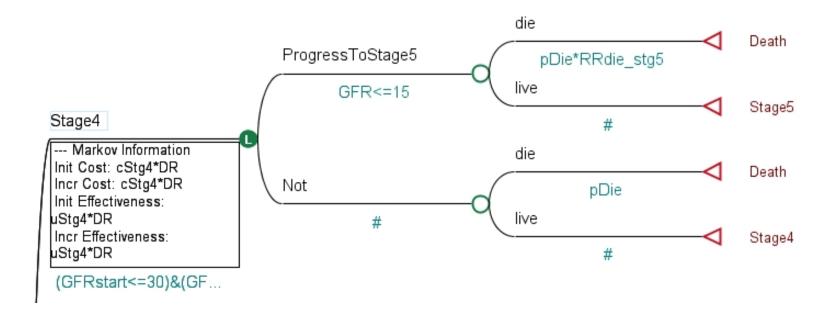


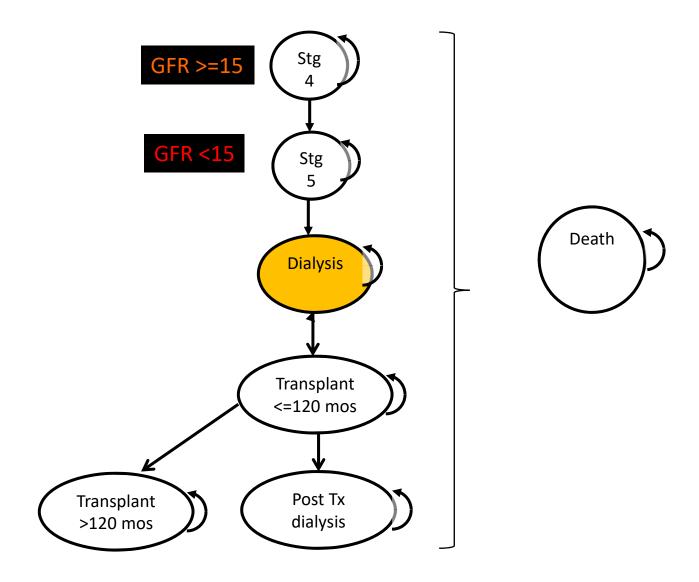


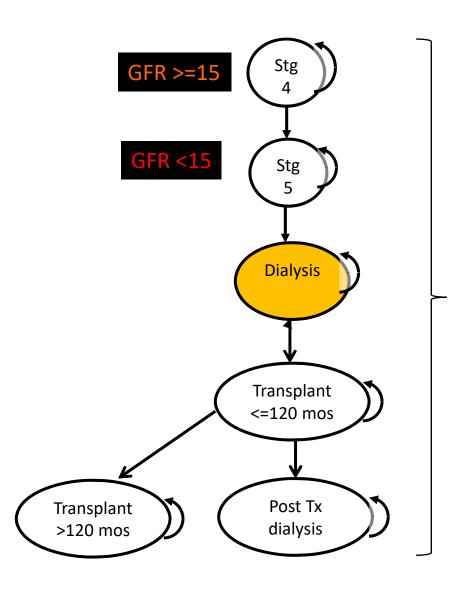


- Stay in stage 4 & live
- Stay in stage 4 & die
- Transition to stage 5 & live
- Transition to stage 5 & die

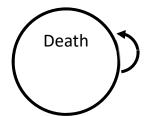


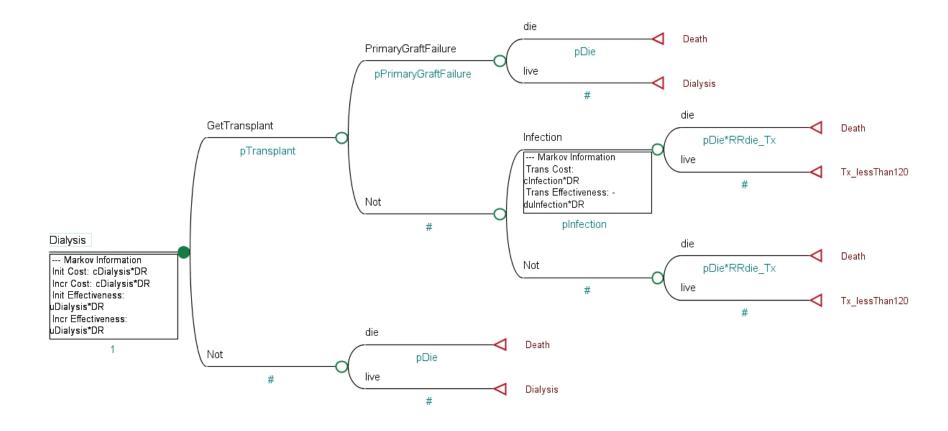






- Receive a transplant:
  - Immediately fails & back to dialysis
  - Succeeds
    - Infection
    - No infection
- No transplant, stay on dialysis
- For all of these, live or die





# Steps to conceptualizing and building a Markov model

- Enumerate possible health states
- From each state consider possible stochastic events that could lead to transitions
  - Construct the cycle trees
- Populate model parameters
  - Incremental utilities
  - Costs
  - Transition probabilities

#### Probabilities come in two flavors

- Bernoulli probabilities
  - Flip-of-a-coin events, do or do not happen
  - For example, immediate transplant failure
     (a.k.a. primary graft failure)

#### Probabilities come in two flavors

- Bernoulli probabilities
  - Why do I use this terminology?

# Consider the binomial probability distribution for a moment

$$P(x = k | n, p) = \binom{n}{k} \cdot p^k \cdot (1 - p)^{n - k}$$

$$P(x = 4|10, 0.3) = {10 \choose 4} \cdot 0.3^4 \cdot (0.7)^6 = \frac{10!}{4! \, 6!} \cdot 0.3^4 \cdot (0.7)^6$$

$$P(x = 4|10, 0.3) = \frac{10 \cdot 9 \cdot 8 \cdot 7}{4 \cdot 3 \cdot 2 \cdot 1} \cdot 0.3^{4} \cdot (0.7)^{6} = 210 \cdot 0.0081 \cdot 0.11765 = 0.2$$

### What if n = 1?

• What values can k have?

#### What if n = 1?

- What values can k have?
  - Well, 0 or 1:

$$P(x = 0|1, p) = {1 \choose 0} \cdot p^0 \cdot (1 - p)^{1 - 0} = 1 - p$$

$$P(x = 1|1, p) = {1 \choose 1} \cdot p^1 \cdot (1 - p)^{1 - 1} = p$$

# The Bernoulli distribution is a special case of the binomial where n = 1

- The percentage chance of drawing a '0' is (1 − p)\*100%
- The percentage chance of drawing a '1' is p\*100%
- This is exactly like flipping a weighted coin such that probability of a heads is p

# Bernoulli probabilities

May have a simple estimate of p from the literature

## Bernoulli probabilities

- Or something a little more sophisticated:
- Say we have a baseline probability estimate
- And an odds ratio from a published logistic regression
- For example:
  - pPimaryGraftFail = 0.025 if age < 55</p>
  - Odds Ratio = 2.1 if age >= 55
- How to model this?

#### Probabilities and Odds

- 1. Convert probability to odds  $p=0.025 \rightarrow odds = (0.025/0.975) = 0.0256$
- 2. Apply odds ratio  $0.0256 \times 2.1 = 0.054$
- 3. Convert back to probability 0.054/1.054 = 0.051

#### Probabilities and Odds

1. Convert probability to odds

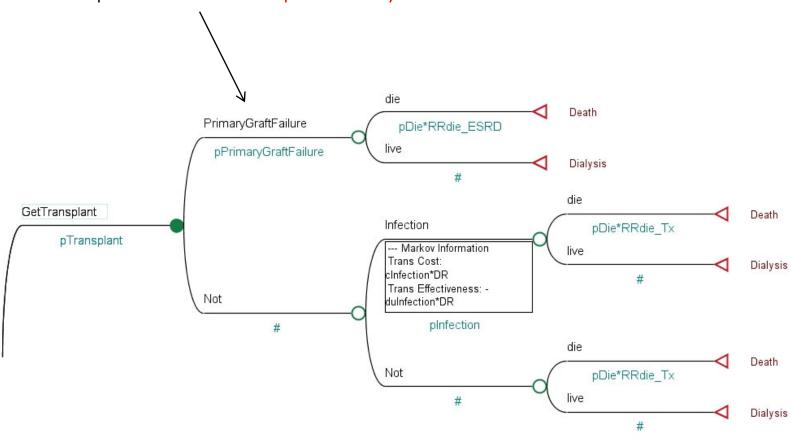
$$p=0.025 \rightarrow odds = (0.025/0.975) = .0256$$

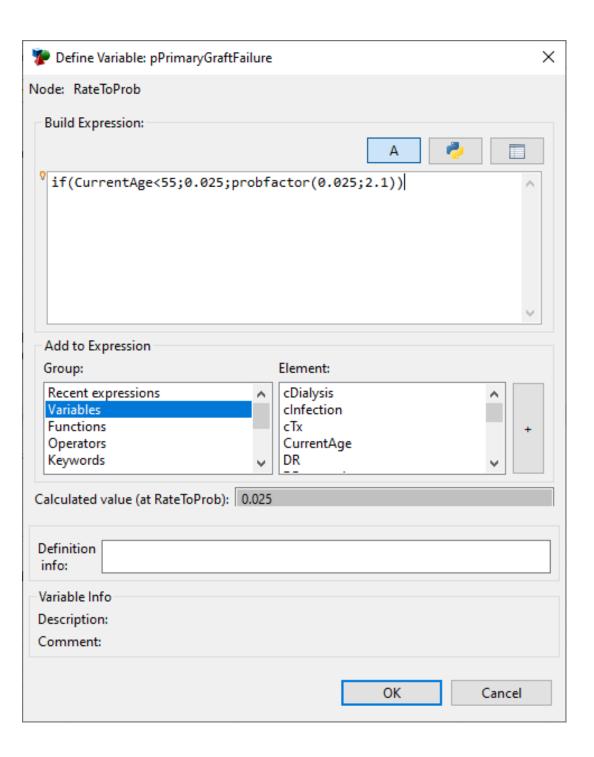
2. Apply odds ratio  $.0256 \times 2.1 = .054$ 

3. Convert back to probability .054/1.054 = .051

There is a built-in TreeAge function that does this for you: probfactor(pBaseline; OR)

#### Example of a Bernoulli probability





#### Probabilities come in two flavors

- Bernoulli probabilities
- Time-to-event probabilities
  - For example, risk of chronic transplant rejection and failure

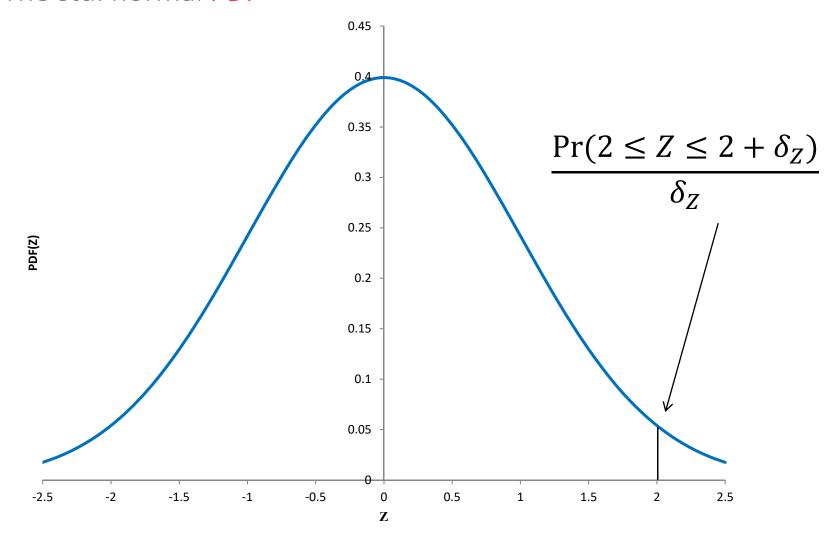
### Time to event probabilities

- The situation when risk of an event accumulates over time
- Data is obtained from a time-to-event (TTE) a.k.a survival analysis
- Our job is to take such data and use it to estimate the per cycle probability of the event – i.e. what is the cumulative risk of the event over one cycle's worth of time?
- For this we need to consider rates and probabilities

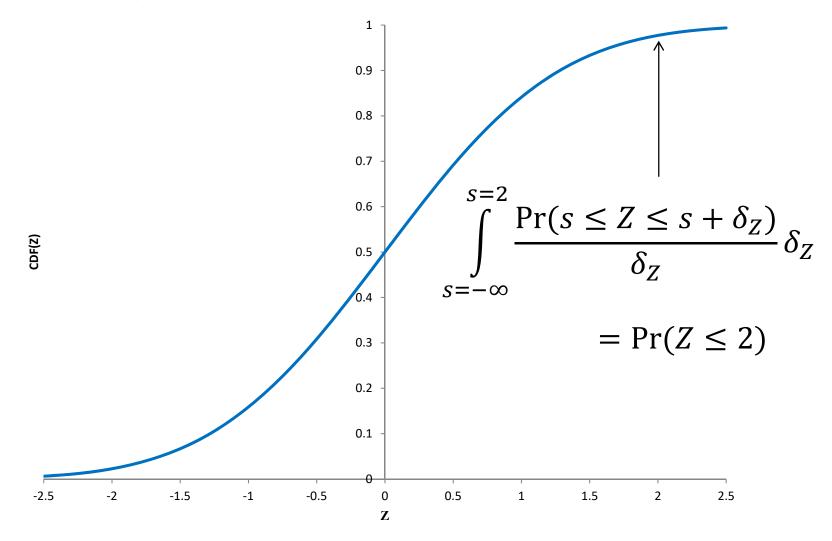
Probability = number of events / number of people at risk

Rate = number of events / time

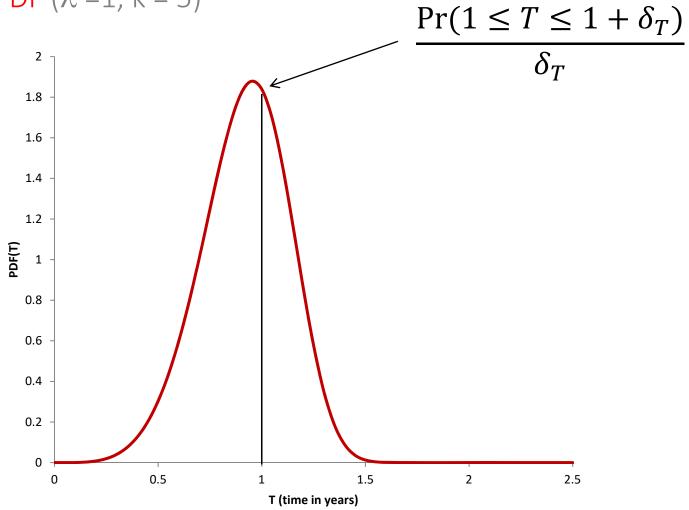
The std. normal PDF



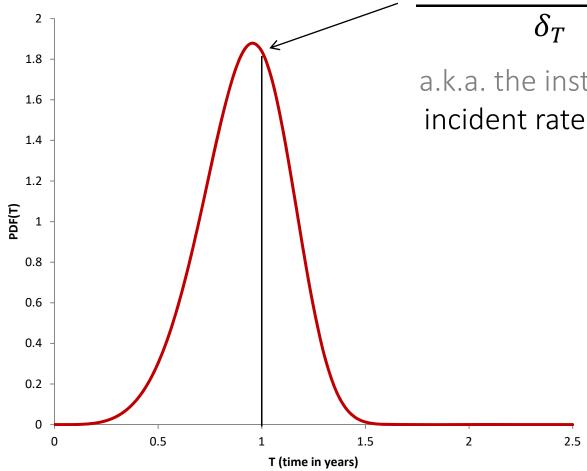
#### The std. normal CDF



### A Weibull PDF ( $\lambda = 1$ , k = 5)



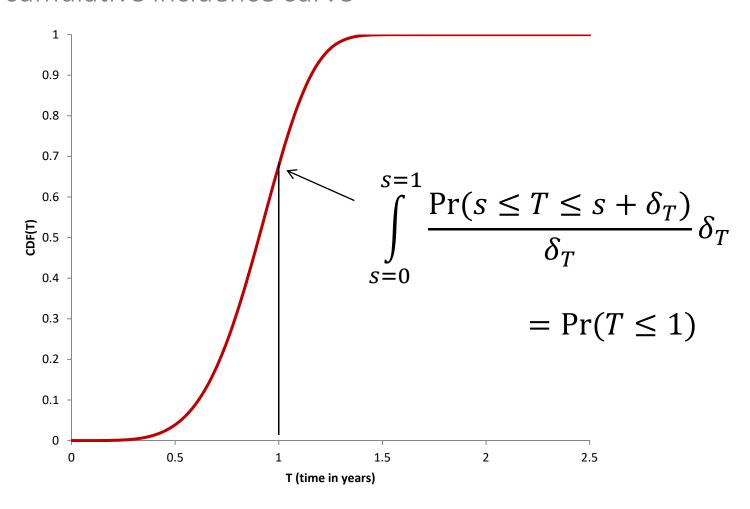
#### A Weibull PDF ( $\lambda = 1$ , k = 5)



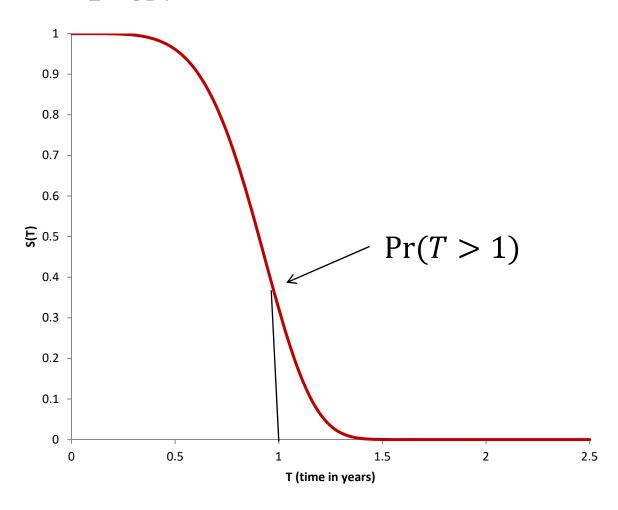
# $\Pr(1 \le T \le 1 + \delta_T)$

a.k.a. the instantaneous incident rate at T = 1 yr

# A Weibull CDF ( $\lambda$ =1, k = 5) a.k.a the cumulative incidence curve



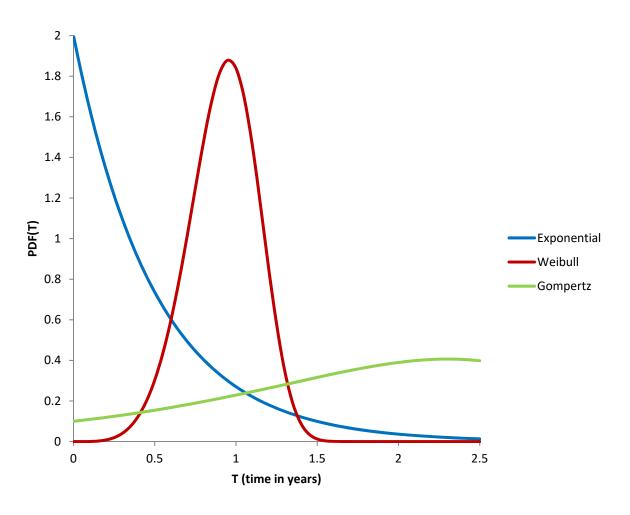
#### A Weibull Survival Function ( $\lambda = 1$ , k = 5)



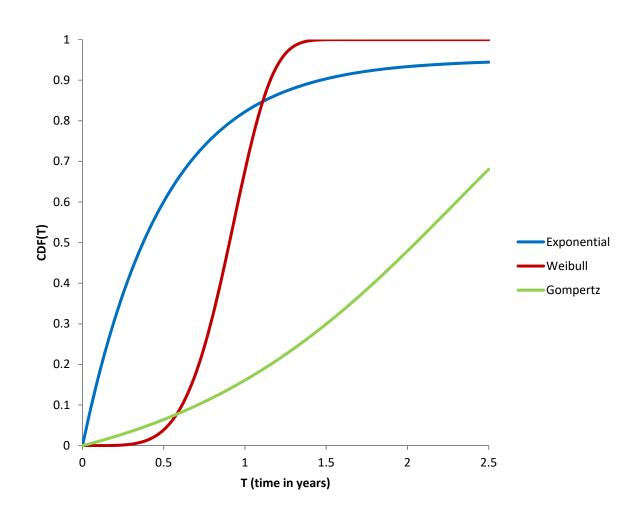
# Types of TTE distributions/survival functions

- Fully parametric
  - Exponential
  - Weibull
  - Gompertz
- Semi-parametric
  - Cox
- Non-parametric
  - Kaplan-Meier

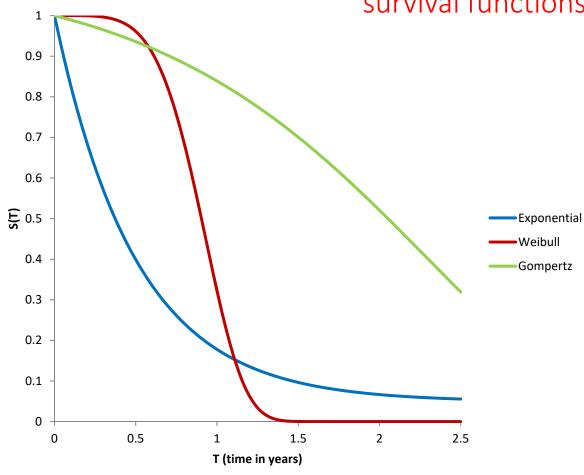
### Various parametric TTE PDFs



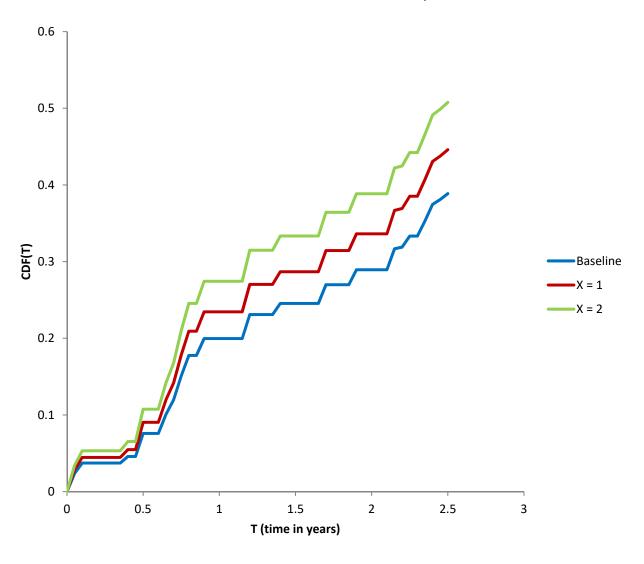
### Various parametric TTE CDFs

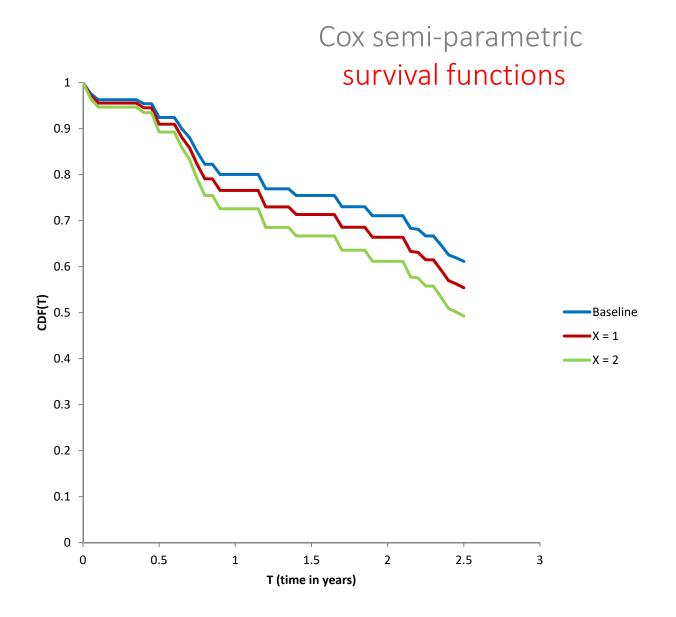


# Various parametric survival functions

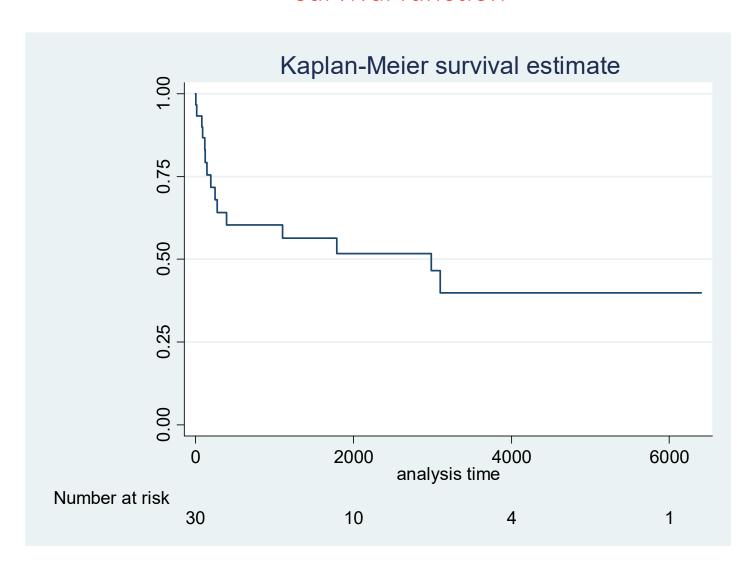


### Cox semi-parametric TTE CDFs

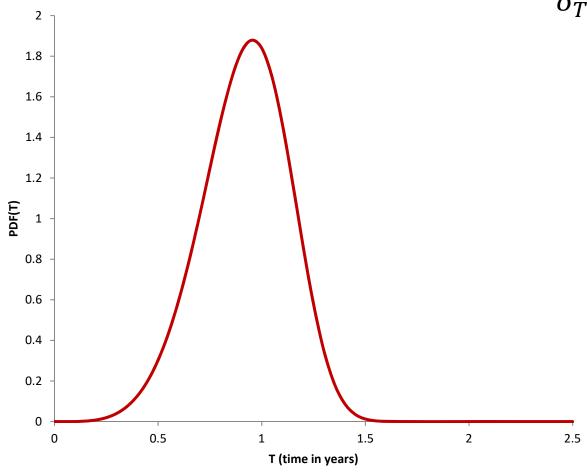


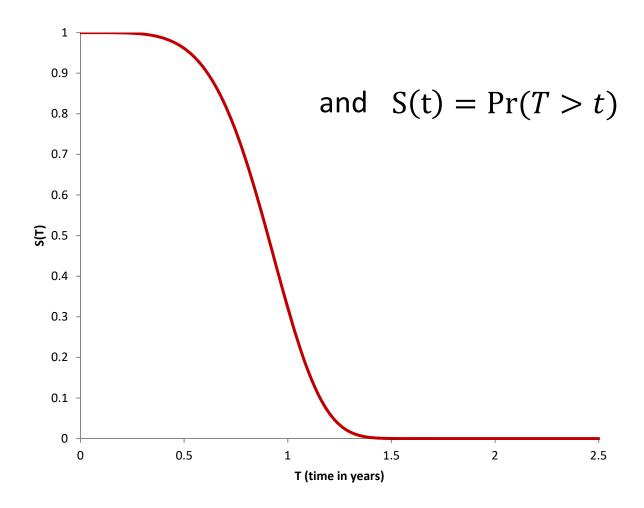


# Kaplan-Meier completely non-parametric survival function

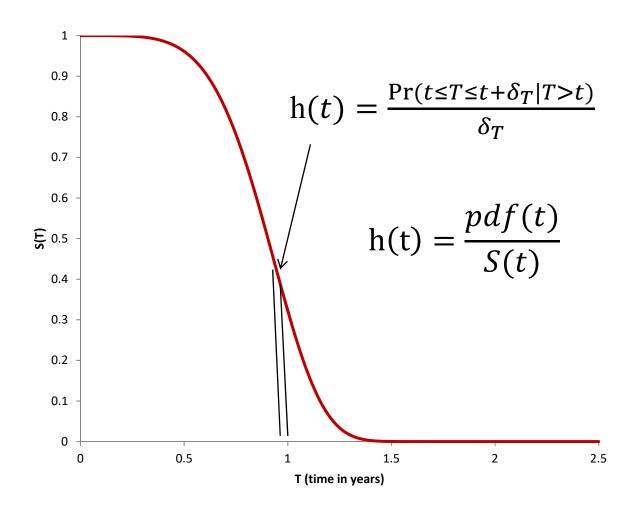


Recall 
$$pdf(t) = \frac{\Pr(t \le T \le t + \delta_T)}{\delta_T}$$

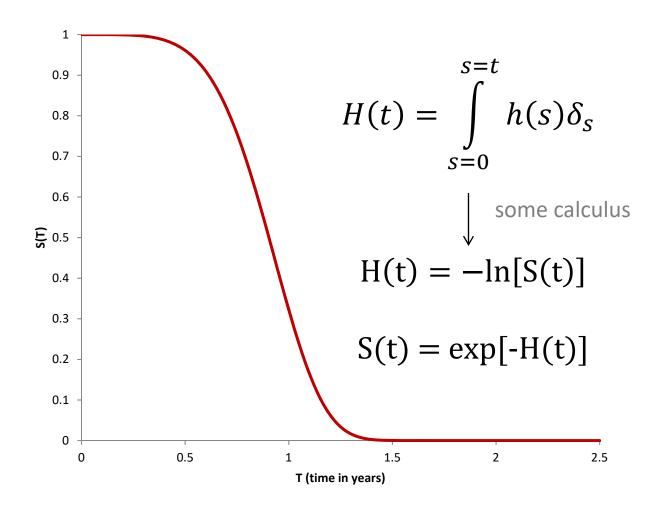




#### The instantaneous hazard rate or function, h(t)



#### The cumulative hazard function, H(t)



#### The relevance of this for decision models:

Probability of no event during a cycle =  $S(t_{end})/S(t_{beginning})$ 

Probability of event during a cycle =  $1 - [S(t_e)/S(t_b)]$ 

= 1- 
$$[exp(-H(t_e))/exp(-H(t_b))]$$

$$= 1 - \exp[H(t_b) - H(t_e)]$$

# The exponential survival model

$$pdf(t) = \lambda * exp(-\lambda t)$$

$$cdf(t) = 1 - exp(-\lambda t)$$

$$S(t) = \exp(-\lambda t)$$

$$h(t) = \lambda$$

$$H(t) = \lambda t$$

# The exponential survival model

pdf(t) = 
$$\lambda * \exp(-\lambda t)$$
  
cdf(t) = 1 -  $\exp(-\lambda t)$   
S(t) =  $\exp(-\lambda t)$   
h(t) =  $\lambda$ 

 $H(t) = \lambda t$ 

Probability of event during a cycle =  $1 - [S(t_e)/S(t_b)]$ 

= 1- 
$$[\exp(-H(t_e))/\exp(-H(t_b))]$$

$$= 1 - \exp[H(t_b) - H(t_e)]$$

$$= 1 - \exp[\lambda t_b - \lambda t_e] = 1 - \exp[\lambda(t_b - t_e)]$$

$$t_{b} - t_{e} = -1$$

So per cycle probability =  $1 - \exp(-\lambda)$ 

Probability of event during a cycle =  $1 - [S(t_e)/S(t_b)]$ 

= 1- 
$$[exp(-H(t_e))/exp(-H(t_b))]$$

$$= 1 - \exp[H(t_b) - H(t_e)]$$

So per cycle probability =  $1 - \exp(-\lambda)$ 

There is a built in TreeAge function for this:

ratetoprob( $\lambda$ ;1)

#### Probability of event during a cycle = $1 - [S(t_e)/S(t_b)]$

$$= 1 - \left[ \exp(-H(t_e)) / \exp(-H(t_b)) \right]$$

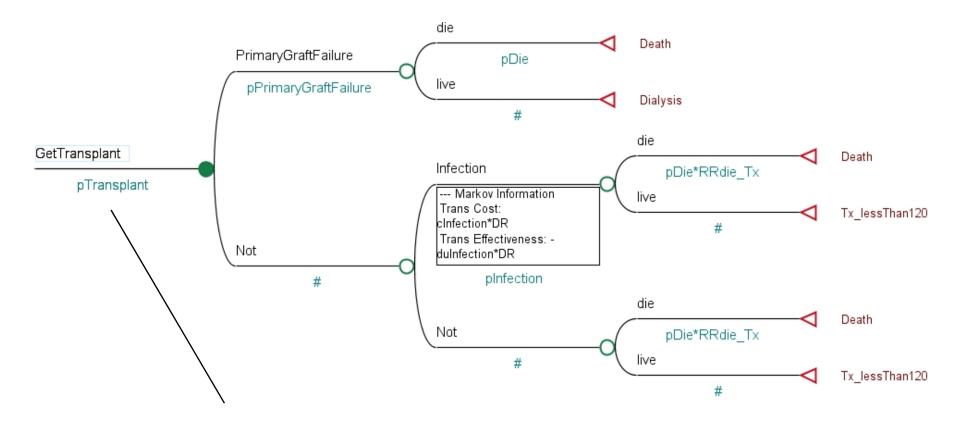
$$= 1 - \exp[H(t_b) - H(t_e)]$$

So per cycle probability =  $1 - \exp(-\lambda)$ 

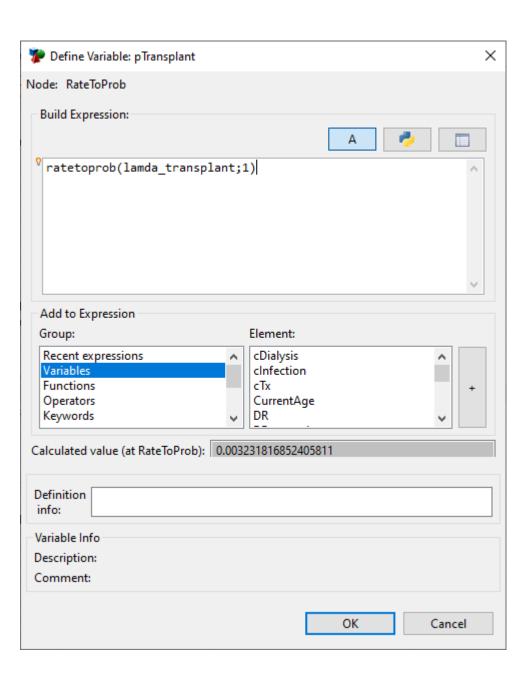
There is a built in TreeAge function for this:

ratetoprob( $\lambda$ ;1)

Note that  $\lambda$  here is the rate per cycle



Example of a time-to-event probability



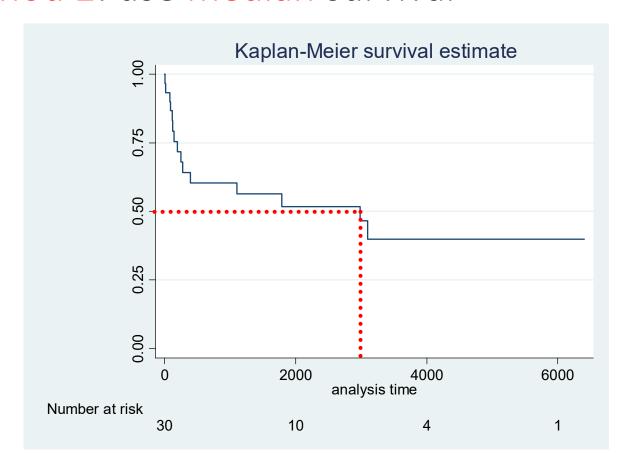
### How to estimate the hazard rate, $\lambda$

- Method 1: use mean survival
  - If a cohort is observed to 100% having the event
  - Assume an exponential distribution
  - Rate  $(\lambda)$  is the reciprocal of the mean survival time
  - e.g. if the mean is 8000 days = 22 years
  - Then the hazard rate,  $\lambda = 1/22 = 0.046$  deaths/year

#### Method 2: use median survival

Median survival is about 3000 days or 8.213 years

$$\lambda = -\ln[1-S(t)]/t^*$$
  
= -\ln(.5)/8.213  
= 0.085 deaths /  
year



<sup>\*</sup>Rearrange  $S(t) = \exp(-\lambda t)$ 

## Method 3: Run an exponential model

```
. streq, d(exp) nohr
        failure d: dead
  analysis time t: followup/365.25
Iteration 0: \log \text{ likelihood} = -57.014599
Iteration 1:
             log likelihood = -57.014599 (backed up)
Exponential regression -- log relative-hazard form
No. of subjects =
                                              Number of obs =
No. of failures =
Time at risk = 142.4229979
                                              LR chi2(0) = 0.00
Log likelihood = -57.014599
                                              Prob > chi2 =
                  Coef. Std. Err. z P>|z| [95% Conf. Interval]
                                  -8.72 0.000 -2.756812
. di \exp(b[1,1])
                                    \ln(\lambda), \lambda = \exp(-2.25) = 0.105
.10532007
```

## Method 4: use cumulative probability (cdf):

E.g. if the 3 year cumulative probability of the event is 11%, i.e. the cdf(3) = 0.11

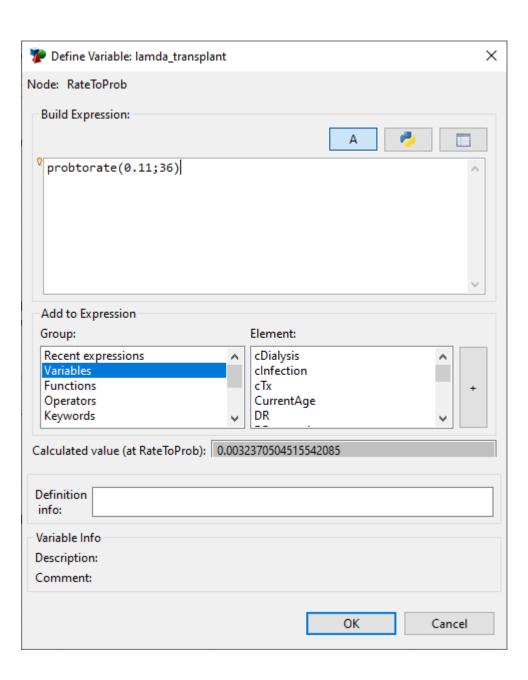
$$cdf(3) = 0.11 = 1 - exp(-\lambda 3); \lambda = -ln[1-0.11]/3$$

TreeAge has a built in function for this:

If the cycle length is a month then

 $\lambda = \text{probtorate}(0.11; 36)$ 

Per cycle probability is then ratetoprob( $\lambda$ ;1)



# Let's build a CKD Markov employing rates and probabilities

- Build a 1-strategy Markov model where all patients start in dialysis with transplant (Tx) and death states considered.
- Cycle length = 1 month, DR = 1.5%/yr, no HCC/WCC, terminate if age > 100 years or the number of QALMs generated by the cohort per cycle is < 0.001</li>
- Only 1 Tx is allowed. Write an expression for pTransplant using probtorate() and ratetoprob() functions from a study where the 3-year cumulative probability for dialysis patients receiving a transplant was 11% assuming an exponential distribution of transplant times

# Let's build a CKD Markov employing rates and probabilities

- In the transition from dialysis to Tx, consider the possibility of immediate transplant failure (primary graft loss). Use probfactor() for the probability of primary graft loss, baseline 0.025, OR = 2.1 if current age >= 55
- Also, model a transient infection event (given that primary graft loss doesn't happen) with a one-time cost and disutility using transition rewards
- Transplants may fail due to chronic graft loss requiring a return to dialysis. These follow an exponential distribution at a rate of 0.005 per month
- After 120 months, risk of Tx failure due to chronic graft loss = 0 (use a tunnel state and \_tunnel counter)

# The end.

