

Online Appendix: Gibbs Sampler and VB

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1 The Model

$$y = Z\theta + \epsilon \quad (1)$$

where y is $T \times 1$ vector of responses, Z is the $T \times k$ matrix of standardized regressors, ϵ is the $T \times 1$ vector of *i.i.d* Normal errors with mean 0 and unknown variance σ^2 .

2 The Priors

Following Bhattacharya, Pati, Pillai, and Dunson (2015) (BPPD), we elicit hierarchical DL prior as following:

$$\theta_j | \phi, \tau \sim DE(\phi_j \tau), \quad \phi_j \sim Dir(a, \dots, a) \quad (2)$$

Conditional prior for θ_j is $DE(\phi_j \tau)$ implies a zero mean Double Exponential or Lapalace distribution with the density $f(\theta_j) = (2\phi_j \tau)^{-1} \exp(-\frac{|\theta_j|}{\phi_j \tau})$ for $\theta_j \in \mathbb{R}$.

Next, we set a Gamma priors for τ and σ^{-2} :

$$\tau \sim G(ka, 1/2), \quad \sigma^{-2} \sim G(\nu, S). \quad (3)$$

The above hierarchical *DL* prior for θ_j can be expressed as:

$$\theta_j \sim N(0, \psi_j \phi_j^2 \tau^2), \quad \psi_j \sim Exp(1/2) \quad (4)$$

Hence, the prior of θ is $N(\mathbf{0}, V)$ where $V = \text{diag}(\psi_1 \phi_1^2 \tau^2, \dots, \psi_k \phi_k^2 \tau^2)$.

3 The Posteriors

Multiply the likelihood and the priors, we have

$$\begin{aligned}
& L(y|\theta, \sigma^2) p(\theta|\phi, \tau, \sigma^2) p(\tau) p(\phi) p(\sigma^2) \\
& \propto \frac{1}{(\sqrt{\sigma^2})^T} \exp\left[-\frac{1}{2\sigma^2} (y - Z\theta)'(y - Z\theta)\right] \times \prod_{j=1}^k \frac{1}{2\phi_j \tau} \exp\left[-\sum_{j=1}^k \frac{|\theta_j|}{\phi_j \tau}\right] \times \\
& \tau^{ka-1} \exp\left(-\frac{\tau}{2}\right) \times \prod_{j=1}^k \phi_j^{a-1} \times (\sigma^{-2})^{\nu-1} \exp(-S/\sigma^2)
\end{aligned} \tag{5}$$

or

$$\begin{aligned}
& L(y|\theta, \sigma^2) p(\theta|\phi, \tau, \psi, \sigma^2) p(\tau) p(\phi) p(\psi) p(\sigma^2) \\
& \propto \frac{1}{(\sqrt{\sigma^2})^T} \exp\left[-\frac{1}{2\sigma^2} (y - Z\theta)'(y - Z\theta)\right] \times \prod_{j=1}^k \frac{1}{\sqrt{\psi_j \phi_j^2 \tau^2}} \exp\left[-\sum_{j=1}^k \frac{\theta_j^2}{2\psi_j \phi_j^2 \tau^2}\right] \times \\
& \tau^{ka-1} \exp\left(-\frac{\tau}{2}\right) \times \prod_{j=1}^k \phi_j^{a-1} \times \prod_{j=1}^k \exp(-\psi_j/2) \times (\sigma^{-2})^{\nu-1} \exp(-S/\sigma^2)
\end{aligned} \tag{6}$$

In (6), the items involving θ are $\exp\left[-\frac{1}{2\sigma^2} (y - Z\theta)'(y - Z\theta)\right] \times \prod_{j=1}^k \frac{1}{\sqrt{\psi_j \phi_j^2 \tau^2}} \exp\left[-\sum_{j=1}^k \frac{\theta_j^2}{2\psi_j \phi_j^2 \tau^2}\right]$. Thus, the conditional posterior of θ is $N(\bar{b}, \bar{V})$, with $\bar{V} = [\sigma^{-2} Z'Z + V^{-1}]^{-1}$ and $\bar{b} = \sigma^{-2} \bar{V} Z'y$.

In (5), the items involving τ are $\frac{1}{\tau^k} \exp\left[-\sum_{j=1}^k \frac{|\theta_j|}{\phi_j \tau}\right] \times \tau^{ka-1} \exp\left(-\frac{\tau}{2}\right)$, which can be written as $\tau^{ka-k-1} \exp\left[-\frac{1}{2}\left(\tau + \frac{1}{\tau} \left(\sum_{j=1}^k \frac{2|\theta_j|}{\phi_j}\right)\right)\right]$. Thus the conditional posterior of τ is $giG(ka - k, 1, \sum_{j=1}^k \frac{2|\theta_j|}{\phi_j})$.¹

To find the conditional posteriors of ϕ_j , we integrate τ out following BPPD. Collecting the terms involving ϕ_j , we have $\prod_{j=1}^k (\phi_j^{a-1} \frac{1}{\phi_j}) \int_{\tau=0}^{\infty} \tau^{ka-k-1} \exp\left(-\frac{\tau}{2}\right) \exp\left[-\sum_{j=1}^k \frac{|\theta_j|}{\phi_j \tau}\right] d\tau$. Thus, we can derive the conditional posterior of ϕ_j as following: First, we have $\xi_j \sim giG(a-1, 1, 2|\theta_j|)$. Next let $\Xi = \sum_{j=1}^k \xi_j$. The conditional posterior of ϕ_j can then be found to be ξ_j/Ξ .

In (6), the terms involving ψ_j s are $\prod_{j=1}^k \psi_j^{-\frac{1}{2}} \exp\left[-\sum_{j=1}^k \frac{\theta_j^2}{2\psi_j \phi_j^2 \tau^2}\right] \times \prod_{j=1}^k \exp(-\psi_j/2)$. Thus

¹ $y \sim giG(p, a, b)$ if $f(y) \propto y^{p-1} \exp\left[-\frac{1}{2}(ay + b/y)\right]$.

the conditional posterior of $1/\psi_j$ is Inverse Gaussian with mean $\sqrt{\frac{\phi_j^2 \tau^2}{\theta_j^2}}$ and scale parameter 1.

In (6), the terms involving σ^{-2} are

$$\frac{1}{(\sqrt{\sigma^2})^T} \exp\left[-\frac{1}{2\sigma^2}(y - Z\theta)'(y - Z\theta)\right] \times (\sigma^{-2})^{\nu-1} \exp(-S/\sigma^2)$$

Thus the conditional posterior of σ^{-2} is $G(\frac{T}{2} + \nu, \frac{1}{2}(y - Z\theta)'(y - Z\theta) + S)$.

4 Variational Bayes

The optimal VB q densities are as follows:

4.1 $q(\theta)$

$$q(\theta) \sim N(\bar{\theta}, \bar{V}), \tag{7}$$

where

$$\bar{V} = \left(\frac{T}{2} + \nu\right) Z'Z + V^{-1})^{-1}$$

$$\bar{\theta} = \left(\frac{T}{2} + \nu\right) \bar{V} Z' y$$

$$V^{-1} = \text{diag}(\overline{\psi_1^{-1} \phi_1^{-2} \tau^{-2}}, \dots, (\overline{\psi_k^{-1} \phi_k^{-2} \tau^{-2}}))$$

$$E[\log(|V^{-1}|)] = \sum_{j=1}^k E[\log(\psi_j) + 2\log(\phi_j)] + 2kE[\log(\tau)],$$

where

$$E[\log(\psi_j) + 2\log(\phi_j)] = \int_0^\infty q(\psi_j) \log(\psi_j) d\psi_j + 2 \int_0^\infty q(\phi_j) \log(\phi_j) d\phi_j$$

and

$$E[\log(\tau)] = \int_0^\infty q(\tau) \log(\tau) d\tau$$

4.2 $q(\sigma^{-2})$

$$q(\sigma^{-2}) \sim G\left(\frac{T}{2} + \nu, \bar{S}\right), \quad (8)$$

where

$$\bar{S} = \frac{1}{2}[\|y - Z\bar{\theta}\|^2 + \text{tr}(Z'Z\bar{V})] + S$$

Hence

$$\bar{\sigma}^{-2} = \frac{\frac{T}{2} + \nu}{\bar{S}},$$

and

$$E[\log(\sigma^{-2})] = \psi\left(\nu + \frac{T}{2}\right) - \log(\bar{S})$$

where $\psi(\bullet)$ is the Digamma function.

4.3 $q(\tau)$

$$q(\tau) \sim giG\left[ka - k, 1, \sum_{j=1}^k 2(\bar{\theta}_j^2 + \bar{V}_{jj})^{1/2} \frac{1}{\phi_j}\right], \quad (9)$$

Let $\chi = \sum_{j=1}^k 2(\bar{\theta}_j^2 + \bar{V}_{jj})^{1/2} \frac{1}{\phi_j}$, we have

$$\bar{\tau} = \frac{\sqrt{\chi} K_{ka-k+1}(\sqrt{\chi})}{K_{ka-k}(\sqrt{\chi})}$$

and

$$\overline{\tau^2} = \bar{\tau}^2 + \chi \left[\frac{K_{ka-k+2}(\sqrt{\chi})}{K_{ka-k}(\sqrt{\chi})} - \left(\frac{K_{ka-k+1}(\sqrt{\chi})}{K_{ka-k}(\sqrt{\chi})} \right)^2 \right]$$

where $K_*[\bullet]$ is the modified Bessel functions of the second kind.

4.4 $q(\psi_j)$

$$q\left(\frac{1}{\psi_j}\right) \sim iG\left(\sqrt{\frac{\phi_j^2 \bar{\tau}^2}{\bar{\theta}_j^2 + \bar{V}^{jj}}}, 1\right), \quad (10)$$

Let $\rho = \sqrt{\frac{\phi_j^2 \tau^2}{\bar{\theta}_j^2 + \bar{V}_{jj}}}$,

$$\frac{1}{\bar{\psi}_j} = \rho$$

and

$$\bar{\psi}_j = 1 + 1/\rho$$

Note that to calculate *ELBO*, we need to use the following optimal q density of ψ_j .²

$$q(\psi_j) = \left(\frac{\psi_j}{2\pi}\right)^{1/2} \exp\left\{-\frac{(\psi_j - \rho)^2}{2\rho^2\psi_j}\right\} \quad (11)$$

4.5 $q(\phi_j)$

$$q(\xi_j) \sim giG(a-1, 1, 2\sqrt{\bar{\theta}_j^2 + (\bar{V}_{jj})^2}) \quad (12)$$

Let $\varpi = 2\sqrt{\bar{\theta}_j^2 + (\bar{V}_{jj})^2}$, we have

$$\bar{\xi}_j = \frac{\sqrt{\varpi} K_a(\sqrt{\varpi})}{K_{a-1}(\sqrt{\varpi})},$$

and

$$var(\xi_j) = \varpi \left\{ \frac{K_{a+1}(\sqrt{\varpi})}{K_{a-1}(\sqrt{\varpi})} - \left[\frac{K_a(\sqrt{\varpi})}{K_{a-1}(\sqrt{\varpi})} \right]^2 \right\}.$$

where $var(\bullet)$ denotes the variance.

Scaling ξ , we have

$$\bar{\phi}_j = \frac{\bar{\xi}_j}{\sum^k \bar{\xi}_j},$$

and

$$\bar{\phi}_j^2 = \bar{\phi}_j + \frac{var(\xi_j)}{(\sum^k \bar{\xi}_j)^2}$$

Thus, the optimal q density of ϕ_j takes the following form:

$$q(\phi_j) \sim giG[a-1, \sum^k \bar{\xi}_j, (2\sqrt{\bar{\theta}_j^2 + (\bar{V}_{jj})^2})/(\sum^k \bar{\xi}_j)] \quad (13)$$

²If x is distributed as $f(x)$, then $y = 1/x$ is distributed as $\frac{1}{y^2} f(\frac{1}{y})$.

4.6 ELBO

The evidence lower bound (*ELBO*) is as follows:³

$$\begin{aligned}
ELBO &= E\{\log p(\mathbf{y}, \theta, \sigma^2, \phi, \tau, \psi)\} - E\{\log q(\theta, \sigma^2, \phi, \tau, \psi)\} \\
&= E\{\log p(\mathbf{y}|\theta, \sigma^2, \phi, \tau, \psi)\} + E\{\log p(\theta)\} + E\{\log p(\sigma^2)\} \\
&\quad + E\{\log p(\phi)\} + E\{\log p(\psi)\} + E\{\log p(\tau)\} \\
&\quad - E\{\log q(\theta)\} - E\{\log q(\sigma^2)\} - E\{\log q(\tau)\} - E\{\log q(\phi)\} - E\{\log q(\psi)\} + \dots
\end{aligned} \tag{14}$$

where

$$E\{\log p(\mathbf{y}|\theta, \sigma^2, \phi, \tau, \psi)\} = -\frac{T}{2} \log(2\pi) - \frac{T}{2} [-\psi(\nu + \frac{T}{2}) + \log(\bar{S})] - \frac{\bar{S} - S}{(\frac{\bar{S}}{\nu + \frac{T}{2}})}, \tag{15}$$

$$E\{\log p(\theta)\} = -\frac{k}{2} \log(2\pi) - \frac{1}{2} E[\log |\mathbf{V}|] - \frac{1}{2} [\bar{\theta}' \mathbf{V}^{-1} \bar{\theta} + \text{tr}(\mathbf{V}^{-1} \bar{\mathbf{V}})], \tag{16}$$

$$E\{\log p(\sigma^{-2})\} = \nu \log S - \log \Gamma(\nu) - (\nu - 1) [-\psi(\nu + \frac{T}{2}) + \log(\bar{S})] - (\frac{S}{\frac{\bar{S}}{\nu + \frac{T}{2}}}), \tag{17}$$

$$E\{\log q(\theta)\} = -(\frac{k}{2} + \frac{k}{2} \log(2\pi) + \frac{1}{2} \log |\bar{\mathbf{V}}_i|), \tag{18}$$

$$E\{\log q(\sigma^{-2})\} = -[\nu + \frac{T}{2} - \log(\bar{S}) + \log(\Gamma(\nu + \frac{T}{2})) - (\nu + \frac{T}{2} - 1) \psi(\nu + \frac{T}{2})]. \tag{19}$$

³It is seen that the *ELBO* described below contains many constant terms. These terms need to be dropped from *ELBOs* in VB iterations to speed up the process.

Notice that

$$\begin{aligned}
& E\{\log p(\mathbf{y}|\theta, \sigma^2, \phi, \tau, \psi)\} + E\{\log p(\theta)\} + E\{\log p(\sigma^2)\} - E\{\log(q(\theta))\} - E\{\log q(\sigma^2)\} \\
&= \frac{k}{2} - \frac{T}{2} \log(2\pi) + \frac{1}{2} \log(|\bar{V}|) - \frac{1}{2} E[\log(|V|)] - \frac{1}{2} [\bar{\theta}' V^{-1} \bar{\theta} + \text{tr}(V^{-1} \bar{V})] + \nu \log(\underline{s}) - \log \Gamma(\nu) \\
&- (\nu + \frac{T}{2}) \log(\bar{s}) + \log(\nu + \frac{T}{2}) \\
&= \frac{1}{2} \log(|\bar{V}|) - \frac{1}{2} E[\log(|V|)] - \frac{1}{2} [\bar{\theta}' V^{-1} \bar{\theta} + \text{tr}(V^{-1} \bar{V})] - (\nu + \frac{T}{2}) \log(\bar{s}) + \text{Const.}
\end{aligned} \tag{20}$$

Above terms are the baseline *ELBO* discussed in GKP2018. The additional terms are as follows:

$$\begin{aligned}
E\{\log p(\tau)\} &= -\log \Gamma(ka) - ka \left[\int_0^\infty (q(\tau) \log \tau) d\tau \right] - 0.5\bar{\tau} \\
&= -ka \left[\int_0^\infty (q(\tau) \log \tau) d\tau \right] - 0.5\bar{\tau} + \text{Const.}
\end{aligned} \tag{21}$$

$$E\{\log p(\psi)\} = \sum_{j=1}^k \left(\log \frac{1}{2} - 0.5\bar{\psi}_j \right) = -\frac{1}{2} \sum_{j=1}^k \bar{\psi}_j + \text{Const.} \tag{22}$$

$$\begin{aligned}
E\{\log p(\phi)\} &= \log \Gamma(ka) - k \log \Gamma(a) - (a-1) \sum_{j=1}^k \left[\int_0^\infty (q(\phi_j) \log \phi_j) d\phi_j \right] \\
&= -(a-1) \sum_{j=1}^k \left[\int_0^\infty (q(\phi_j) \log \phi_j) d\phi_j \right]
\end{aligned} \tag{23}$$

$$E\{\log q(\tau)\} = \int_0^\infty q(\tau) \log q(\tau) d\tau. \tag{24}$$

$$E\{\log q(\psi)\} = \sum_{j=1}^k \int_0^\infty q(\psi_j) \log q(\psi_j) d\psi_j \tag{25}$$

$$E\{\log q(\phi)\} = \sum_{j=1}^k \int_0^\infty q(\phi_j) \log q(\phi_j) d\phi_j \tag{26}$$

Terms in *ELBO* that do not have clear analytical forms can be numerically estimated.

References

- [1] Bhattacharya, A., Pati, D., Pillai, N., and Dunson, D. (2015), Dirichlet-laplace priors for optimal shrinkage, *Journal of the American Statistical Association*, 110, 1479-1490.