## Next Generation Statistical Methods for Genome-Wide Association Studies

Course Overview and Intro
06 September 2023

## Agenda

- Course format
- Intro to genetic epidemiology
- Course scope
- A few basic definitions

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## DCEG Statistical Genetics Workshop

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Course format Intended audience

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#### Next Generation Statistical Methods for Genome Wide Association Studies: A Hands-On Course

#### Background

Genome-wide association studies (GWAS) have revolutionized our understanding of the genetic basis of complex traits and diseases. In the early years of GWAS, data analysis primarily relied on relatively simplistic methods, such as running millions of univariate linear or logistic regressions, one for each genetic variant. However, as the sample sizes for some GWAS have become extremely large and various types of other genomic data have become widely available, analysis of such data has

#### Latest Posts

2023

DCEG Statistical Genetics Workshop schedule for fall

Statistical Genetics Workshop Announcement

Location: Rm 1106-A/B at the CRL Building, 9615 Medical Center Drive, Rockville, MD 20892/online

Time: 9:30-12:30 EST

Link: https://nih.zoomgov.com/j/1600232059? pwd=aW1NTmRCWXAwajZ0bFN0ZEtTQWhiUT09

Read more

Published: Jul 5, 2023

#### Course format

The course will consist of nine sessions held from September to December of 2023. Sessions will be held on Wednesdays from 9:30 to 12:30 will include a lecture (1.25 hours, including Q&A) and a 1.5-hour practical tutorial. (See schedule below for specific dates.) Participants are expected to complete background reading before each session (estimated out-of-class time: < 2 hours) and hands-on exercises after each session (estimated out-of-class time: < 2 hours). The course will be hybrid with both in-person and online participants, and all lectures will be recorded and archived for future use. Practical tutorials will be in-person at the Shady Grove NCI campus.

#### Intended audience

Researchers and analysts with strong quantitative background who are involved or anticipate being involved in analysis of large-scale genome-wide genotyping data. Participants should have basic knowledge of epidemiologic study designs, genetic concepts and terminologies, and statistical methodologies (e.g., hypothesis testing, parameter estimation, regression models, Bayes probability), as well as familiarity with R and command-line interfaces.

By the end of the course, participants will have gained a deep understanding of advanced statistical methods and computational tools for analyzing GWAS data, and will be able to apply these methods to their own research. They will also be familiar with best practices for data management and sharing in GWAS, and will be able to produce reproducible and FAIR-compliant pipelines.

Published: Jul 5, 2023 by Wendy Wong

## DCEG Statistical Genetics Workshop schedule for fall 2023

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Time: 9:30-12:30 EST

Link: https://nih.zoomgov.com/j/1600232059?pwd=aW1NTmRCWXAwajZ0bFN0ZEtTQWhiUT09

- Session 1: Introduction 09/06/23
- Session 2: Basic GWAS analyses 09/20/23
- Session 3: Fine-mapping and colocalization 09/27/23
- Session 4: Heritability, functional enrichment, polygenic scores 10/25/23
- Session 5: Rare variants 11/08/23
- Session 6: Integrative analyses and Mendelian Randomization 11/15/23
- Session 7: GWAS, fine-mapping and PRS in diverse-genetic-ancestry and admixed samples 11/29/23
- Session 8: Genetic mosaicism and clonal hematopoiesis 12/06/23
- Session 9: Functional genomics 12/13/23

### Instructors



Kevin Brown, Ph.D.

DCEG, NCI

Dr. Brown received a Ph.D. in Genetics from the George Washington University in Washington, D.C., in 2003. He conducted his postdoctoral training in the Laboratory of Dr. Jeffrey Trent at the Translational Genomic Research Institute (TGen) in Phoenix, Arizona. He subsequently went on to direct his own research program at TGen as an investigator from 2005 to 2010, and served as an adjunct professor in basic medical sciences at the Mayo Clinic Cancer Center, the University of Arizona College of Medicine, and Arizona State University from 2008 to 2010. His work at TGen involved the application of whole-genome familial linkage, candidate gene, and genome-wide association study (GWAS) approaches to identify genetic variants associated with melanoma susceptibility. In 2010, Dr. Brown joined the Laboratory of Translational Genomics (LTG) in the Division of Cancer Epidemiology and Genetics (DCEG) as a tenure-track investigator. He was awarded NIH scientific tenure and promoted to senior investigator in 2018. His research focuses on the genetic underpinnings of melanoma susceptibility.

Kevin Brown's profile







#### Introduction

Basic GWAS analyses







+ Supplemental

#### **Topics Covered**

- · Why genetic epidemiology?
- Why GWAS and RVAS?
- What we have learned.
- What remains to be done.
- · Overview of the course

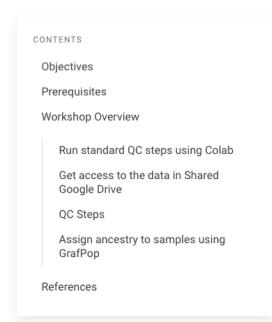
#### Practical

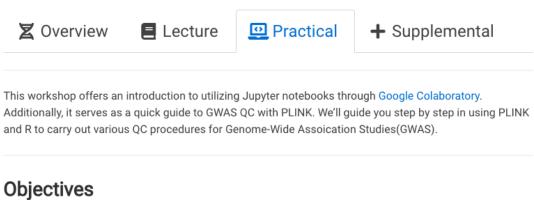
- · To introduce the Google Colaboratory platform.
- · To introduce participants the importance of Quality Control in the analysis pipeline.
- To provide participants with the basic practical skills needed to perform QC of GWAS data using plink.

#### Recommended Readings/Videos

- · Readings on genome-wide association study design and analysis
  - Uffelmann, E., Huang, Q. Q., Munung, N. S., de Vries, J., Okada, Y., Martin, A. R., Martin, H. C., & Lappalainen, T. (2021). Genome-wide association studies. Nature Reviews Methods Primers, 1(1), 1–21. https://doi.org/10.1038/s43586-021-00056-9
- · Overview of findings from genome-wide association studies
  - Penney, K. L., Michailidou, K., Carere, D. A., Zhang, C., Pierce, B., Lindström, S., & Kraft, P. (2017).
     Genetic Epidemiology of Cancer. https://doi.org/10.1093/oso/9780190238667.003.0005
  - Visscher, P. M., Wray, N. R., Zhang, Q., Sklar, P., McCarthy, M. I., Brown, M. A., & Yang, J. (2017). 10
     Years of GWAS Discovery: Biology, Function, and Translation. American Journal of Human
     Genetics, 101(1), 5–22. https://doi.org/10.1016/j.ajhg.2017.06.005

## Session 1:Introduction





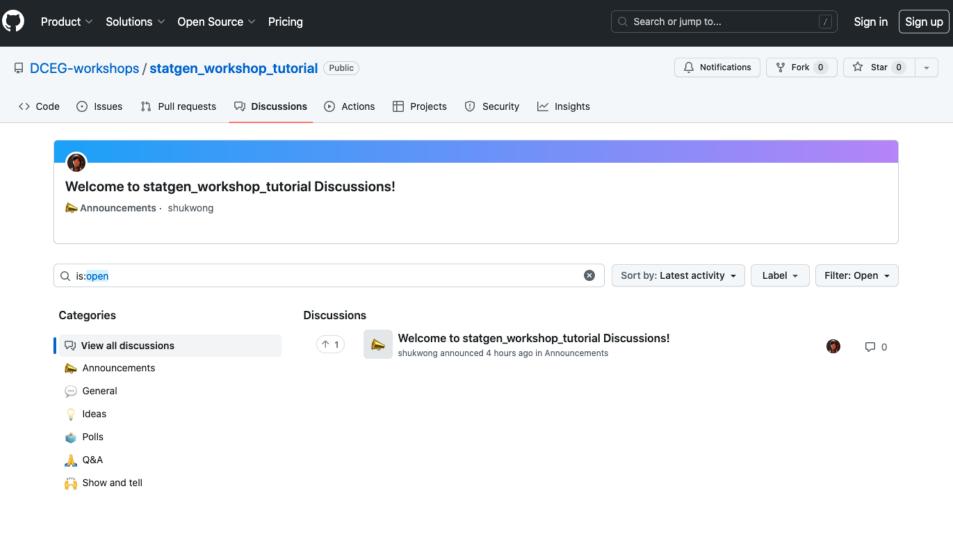
- To introduce participants the importance of Quality Control in the analysis pipeline.
- To provide participants with the basic practical skills needed to perform QC of GWAS data using plink.

#### **Prerequisites**

Before starting this workshop, we recommend that you have:

· To introduce the Google Colaboratory platform.

- A Google Account
- · A basic understanding of genetics, genome-wide association studies (GWAS), and the role of quality control (QC) in GWAS.
- · Familiarity with command-line tools, as we will use the command-line tool plink for the QC steps.



Keep the discussion going! Post questions, comments, resources to the workshop discussion page.



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- A few basic definitions

Epidemiology is the study of the distribution and determinants of health-related states in specified populations, and the application of this study to control health problems.

Genetic variation is one of the determinants of health-related states in populations, so...

Genetic epidemiology is the study of the role of genetic factors in determining health and disease in families and in populations, and the interplay of such genetic factors with social and environmental factors

# How can genetic epidemiology improve human health?

- Genetic epidemiology has contributed to a deeper understanding of disease biology, leading to new therapeutics.
- Genetic epidemiology also has direct clinical and public health applications.

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- Genetic epidemiology also has direct clinical and public health applications.

HEALTH

## Rare Gene Mutations Inspire New Heart Drugs

By GINA KOLATA MAY 24, 2017









HEALTH

## Rare Mutation Ignites Race for Cholesterol Drug

**Genetic Connections** 

By GINA KOLATA JULY 9, 2013









**BUSINESS DAY** 

## Aiming to Push Genomics Forward in New Study

By ANDREW POLLACK JAN. 13, 2014









Researchers identify a healthy individual with extremely low levels of LDL cholesterol.

Three siblings of this proband also had very low LDL levels.

DNA sequencing determined that these siblings all carried two copies of a rare loss-offunction variant in the ANGPTL3 gene.

Drugs that mimic the effect of this mutation are more likely to be safe and effective.

## Rare Gene Mutations Inspire New Heart Drugs

By GINA KOLATA MAY 24, 2017









HEALTH

## Rare Mutation Ignites Race for Cholesterol Drug

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## Aiming to Push Genomics Forward in New Study

By ANDREW POLLACK JAN. 13, 2014









Researchers determine that genetic variation in PCSK9 influences LDL levels.

They screen study subjects for new variants that have large effects and identify a healthy individual homozygous for an extremely rare loss-offunction variant with extremely low LDL levels.

Drugs that mimic the effect of this mutation are more likely to be safe and effective.





Genome-wide association study shows BCL11A associated with persistent fetal hemoglobin and amelioration of the phenotype of  $\beta$ -thalassemia

REPOR





Fetal hemoglobin (HbF) contributes to variability in the expression of hemoglobinopathies. A genome-wide association study identified a single nucleotide polymorphism near the *BCL11A* associated with HbF levels.

Silencing BCL11A prevents sickle cell disease in mice.

Trials of several approaches to silencing *BCL11A* are now underway.

Uda M(2008) PNAS Xu J (2011) Science Frangoul (2021) NEJM

- Genetic epidemiology has contributed to a deeper understanding of disease biology, leading to new therapeutics.
- Genetic epidemiology also has direct clinical and public health applications.

## Genetics in Public Health

- Predictive: Screening those with rare risk variants
  - BRCA1/BRCA2, Lynch Syndrome/HNPCC
- Diagnostic: screening newborns
  - Amer. College of Med. Genetics: screen for 29 conditions

 Table 2

 Newborn screening panel: core panel and secondary targets

		01	, .	,
		MS/MS		
Acylcarnitines		Amino acids		
9 OA	5 FAO	6 AA	3 Hb Pathies	6 Others
		CORE PANEL		
IVA	MCAD	PKU	Hb SS*	СН
GA I	VLCAD	MSUD	Hb S/βTh*	BIOT
HMG	LCHAD	HCY*	Hb S/C*	CAH*
MCD	TFP	CIT		GALT
MUT*	CUD	ASA		HEAR
3MCC*		TYR I*		CF

List includes congenital hypothyroidism, phenylketonuria, hemoglobinopathies, maple syrup urine disease, cystic fibrosis.

Cbl A,B\*

PROP

Brice and Zimmern (2010) Khoury et al (ed); Watson (2006) Genet Med

## Genetics in Public Health

Risk-stratified screening for breast cancer, combining data on questionnaire risk factors (QRF), mammographic density (MD), common risk variants (polygenic risk scores [PRS]), family history, and rare risk variants.

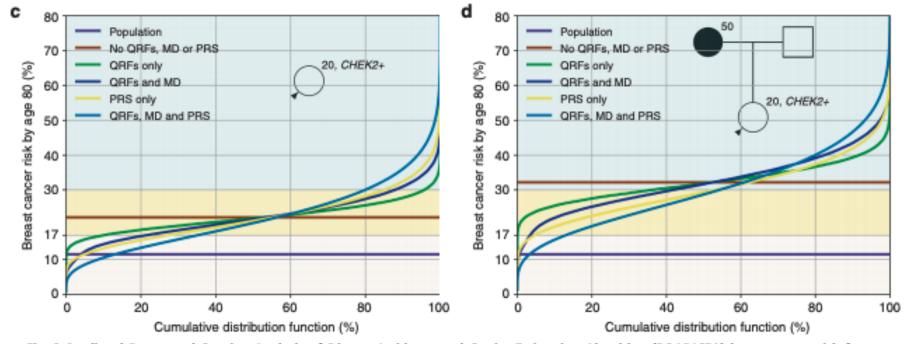
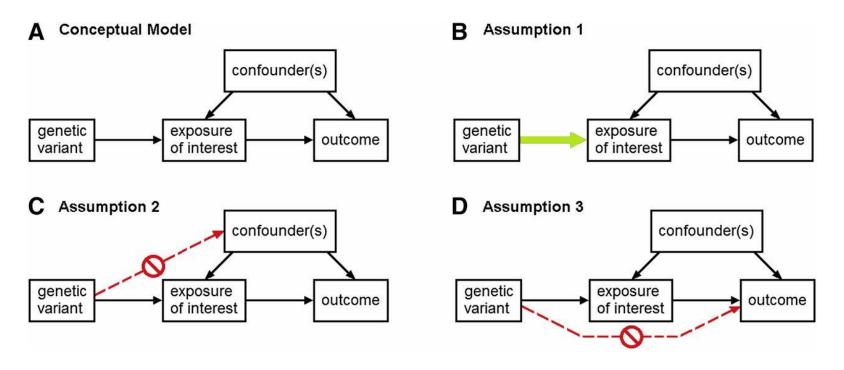


Fig. 3 Predicted Breast and Ovarian Analysis of Disease Incidence and Carrier Estimation Algorithm (BOADICEA) breast cancer risk for a female intermediate-risk rare pathogenic variant carrier, on the basis of the different predictors of risk (questionnaire-based risk factors [QRFs], mammographic density [MD], and polygenic risk scores [PRS]). (a, c) Lifetime risk (age 20 to 80 years) for a CHEK2 1100delC carrier with unknown family history; (b, d) lifetime risk for a CHEK2 1100delC carrier with her mother affected at age 50. (e, f) Risk for a PALB2 and an ATM rare pathogenic variant carrier respectively, both with unknown family history. The backgrounds of the graphs are shaded to indicate the familial breast cancer risk categories based on the National Institute for Health and Care Excellence (NICE) guidelines:<sup>3</sup> (1) near-population risk shaded in pink (<17%), (2) moderate risk shaded in yellow (≥17% and <30%), and (3) high risk shaded in blue (≥30%). Predictions based on UK breast cancer incidence.

- Genetic epidemiology has contributed to a deeper understanding of disease biology, leading to new therapeutics.
- Genetic epidemiology also has direct clinical and public health applications.
- But wait, there's more!

## Mendelian Randomization

Use genetic variation as an instrumental variable to assess the causal relationship between an exposure and an outcome.



## Mendelian Randomization

Mendelian Randomization

studies suggest CRP and HLD

levels are not causally associated

with coronary heart disease,

consistent with clinical trials.

ORIGINAL CONTRIBUTION



#### Genetic Loci Associated With C-Reactive Protein Levels and Risk of Coronary Heart Disease

Paul Elliott, FRCP John C. Chambers, PhD Weihua Zhang, PhD Robert Clarke, MD Jemma C. Hopewell, PhD John F. Peden, PhD

Jeanette Erdmann, PhD and genotyping between 2003 and Peter Braund, MSc James C. Engert, PhD Derrick Bennett, PhD

Lachlan Coin, PhD Deborah Ashby, PhD Ioanna Tzoulaki, PhD Main Outcome Measure Disk Ian J. Brown, PhD

Shahrul Mt-Isa, BSc Mark I. McCarthy, FRCP

Leena Peltonen, MD, PhD Nelson B. Freimer, MD

Martin Farrall, FRCPath Aimo Ruokonen, MD, PhD

Anders Hamsten, MD Noha Lim, PhD

Philippe Froguel, MD Dawn M. Waterworth, PhD

Peter Vollenweider, MD Gerard Waeber, MD

Marjo-Riitta Jarvelin, MD

Vincent Mooser, MD James Scott, FRS

Alistair S. Hall, FRCP Heribert Schunkert, MD

Sonia S. Anand, MD Rory Collins, FRCP

Nilesh J. Samani, FRCP Hugh Watkins, FRCP

Jaspal S. Kooner, FRCP See also pp 49 and 92.

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Context Plasma levels of C-reactive protein (CRP) are independently associated with risk of coronary heart disease, but whether CRP is causally associated with coronary heart disease or merely a marker of underlying atherosclerosis is uncertain.

Objective To investigate association of genetic loci with CRP levels and risk of coro-

Design, Setting, and Participants We first carried out a genome-wide association (n=17967) and replication study (n=13615) to identify genetic loci associated with plasma CRP concentrations. Data collection took place between 1989 and 2008

study of the most closely associate locus and published data on other 100823 controls, to Investigate th disease. We compared our finding vational studies of CRP levels and associated with CRP levels, we sel against coronary heart disease amo

Results Polymorphisms in 5 gen (% difference per minor allele): SN Interval [CI], -17.6% to -12.0%; I Cl, -14.4% to -8.5%;  $P=1.3\times10$ Cl, -23.4% to -17.9%;  $P=1.3\times$ -16.6% to -10.9%; P=1.9×10-1 CL =25.3% to =18.1% · P=8.1 × 1 cus with coronary heart disease g 1.01) per 20% lower CRP level. Ou CRP locus showed no association w to 1.02; per 20% lower CRP level, predicted from meta-analysis of the heart disease (z score, -3.45; P < .0

1.02 to 1.09; per minor allele), rs4537545 in исык (ОК, 0.94; 95% Ст, 0.91 to 0. and rs4420638 in the APOE-CI-CII cluster (OR, 1.16; 95% CI, 1.12 to 1.21) were all associated with risk of coronary heart disease.

Conclusion The lack of concordance between the effect on coronary heart disease risk of CRP genotypes and CRP levels argues against a causal association of CRP with coronary heart disease.

ORONARY HEART DISEASE death worldwide.1 Inflammation plays a key role in the pathogenesis of CHD at every stage CRP has been advocated as a means of from initiation to progression and rupture of the atherosclerotic plaque.2 Creactive protein (CRP), an acutephase protein synthesized primarily by the liver, is currently the most widely

used biomarker of inflammation.3 Ob-

servational studies have consistently (CHD) is the leading cause of demonstrated that higher plasma levels of CRP are associated with higher risk of CHD.4,5 and measurement of improving risk prediction.6 There is

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(Reprinted) JAMA, July 1, 2009-Vol 302, No. 1 37

#### Plasma HDL cholesterol and risk of myocardial infarction: a mendelian randomisation study

Benjamin F Volght\*, Gina M Peloso\*, Marju Orho-Melander, Ruth Frikke-Schmidt, Maja Barbalic, Majken K Jensen, George Hindy, Hilma Hölm, Eric L Ding, Toby Johnson, Herbert Schunkert, Nilesh J Samani, Robert Clarke, Jemma C Hopewell, John F Thompson, Mingyro Li Gudmar Thorleifsson, Christopher Newton-Cheh, Kiran Musunuru, James P Pirruccello, Danish Saleheen, Li Chen, Alexandre F R Stew Ame Schillert, Unnur Thorsteinsdottir, Gudmundur Thorgeirsson, Sonia Anand, James C Engert, Thomas Morgan, John Spertus, Monika Stall, Klaus Berger, Nicola Martinelli, Domenico Girelli, Pascal P McKeown, Christopher C Patterson, Stephen E Epstein, Joseph Devaney, Mary-Susan Burnett, Vincent Mooser, Samuli Ripotti, Ida Suvakka, Markku S Nieminen, Juha Sinisala, Marja-Litsa Lakki, Markus Perola Aki Havulinna, Ulf de Faire, Bruna Gigante, Erik Ingelsson, Tanja Zeller, Philipp Wild, Paul I W de Bakker, Olaf H Klungel thonius de Boer, Diederick E Grobbee, Pieter W Komphuisen, Vera H M Deneer, Clara C Elbers,

nen W.M.Maninus Verschuren, Jolando M.A. Raer, Vuonne Tuan der Schauss, Asif Roshoed Do, Jase M Ordovas, Gançalo R Abecasis, Michael Baehnke, Karen I, Mahlke, Mark J Daly lez, Shaun Purcell, Stacey Gabriel, Jaume Mamugat, John Peden, Jeanette Erdmann, ous Fischer, Christian Hengstenberg, Andreas Ziegler, Ian Buysschaert, Diether Lambrecht ri. Diene Rubin, Türgen Schrezenmeir, Stefan Schreiber, Arne Schäfer, John Danesh ugh Watkins, Alistair S Hall, Kim Overvad, Eric Rimm, Eric Boerwinkle, Anne Tybjaerg-Hansen Pier M Mannucci, Diego Ardissino, David Siscovick, Roberto Elosua, Kari Stefansson Leena Peltonen, Stephen M Schwartz, David Altshuler, Seker Kethiresan

ol is associated with reduced risk of myocardial infarction, but whether this ng the fact that genotypes are randomly assigned at meiosis, are independent odified by disease processes, mendelian randomisation can be used to test sma biomarker with disease is causal.

andomisation analyses. First, we used as an instrument a single nucleotide lial lipase gene (LIPG Asn396Ser) and tested this SNP in 20 studies controls). Second, we used as an instrument a genetic score consisting of ate with HDL cholesterol and tested this score in up to 12482 cases of As a positive control, we also tested a genetic score of 13 common SNPs

ele (2.6% frequency) had higher HDL cholesterol (0.14 mmol/L higher, p=8×10<sup>-13</sup>) but similar levels of other lipid and non-lipid risk factors for myocardial infarction compared with noncarriers. This difference in HDL cholesterol is expected to decrease risk of myocardial infarction by 13% (odds ratio Mensolynoth General Hospital, [OR] 0-87, 95% CI 0-84-0-91). However, we noted that the 396Ser allele was not associated with risk of myocardial infarction (OR 0.99, 95% CI 0.88-1.11, p=0.85). From observational epidemiology, an increase of 1 SD in HDL cholesterol was associated with reduced risk of myocardial infarction (OR 0 · 62, 95% CI 0 · 58-0 · 66). However, a 1 SD increase in HDL cholesterol due to genetic score was not associated with risk of myocardial infarction (OR 0.93, 95% CI 0 · 68-1 · 26, p=0 · 63). For LDL cholesterol, the estimate from observational epidemiology (a 1 SD increase in LDL cholesterol associated with OR 1 · 54, 95% CI 1 · 45-1 · 63) was concordant with that from genetic score (OR 2 · 13, 95% CI 1-69-2-69, p=2x10<sup>-10</sup>).

> Interpretation Some genetic mechanisms that raise plasma HDL cholesterol do not seem to lower risk of myocardial infarction. These data challenge the concept that raising of plasma HDL cholesterol will uniformly translate into reductions in risk of myocardial infarction

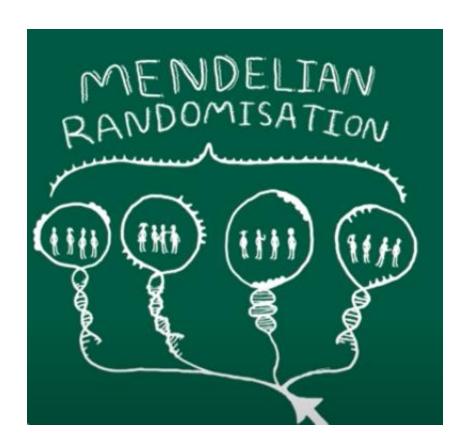
> Funding US National Institutes of Health, The Wellcome Trust, European Union, British Heart Foundation, and the German Federal Ministry of Education and Research.

Cholesterol fractions such as LDL and HDL cholesterol are among the most commonly measured biomarkers in clinical medicine. Observational studies have shown that LDL and HDL cholesterol have opposing associations

with risk of myocardial infarction, with LDL cholesterol being positively associated and HDL cholesterol being inversely associated.23 However, observational studies cannot distinguish between a causal role in the pathological process and a marker of the underlying

www.thelancet.com Vol 380 August 11, 2012

## Mendelian Randomization



A cute 00:02:15 introduction... that glosses over some important pesky details (to be discussed later).

## Genetic Epidemiology is Transdisciplinary

The history of genetic epidemiology is a tapestry of observational science, statistical developments, animal and plant breeding experiments, molecular experiments, medicine, epidemiology, technology, and...

Mendel
(discrete traits)
Galton, Pearson
(continuous traits)
Bateson, Punnett
(Mendel "rediscovered")
Fisher, Wright, Haldane
("the modern synthesis")

Mendelian and Statistical Genetics

Garrod
("inborn errors of metabolism,"
pharmacogenetics)
Neel
(hemoglobinopathies, radiation)

Clinical Genetics

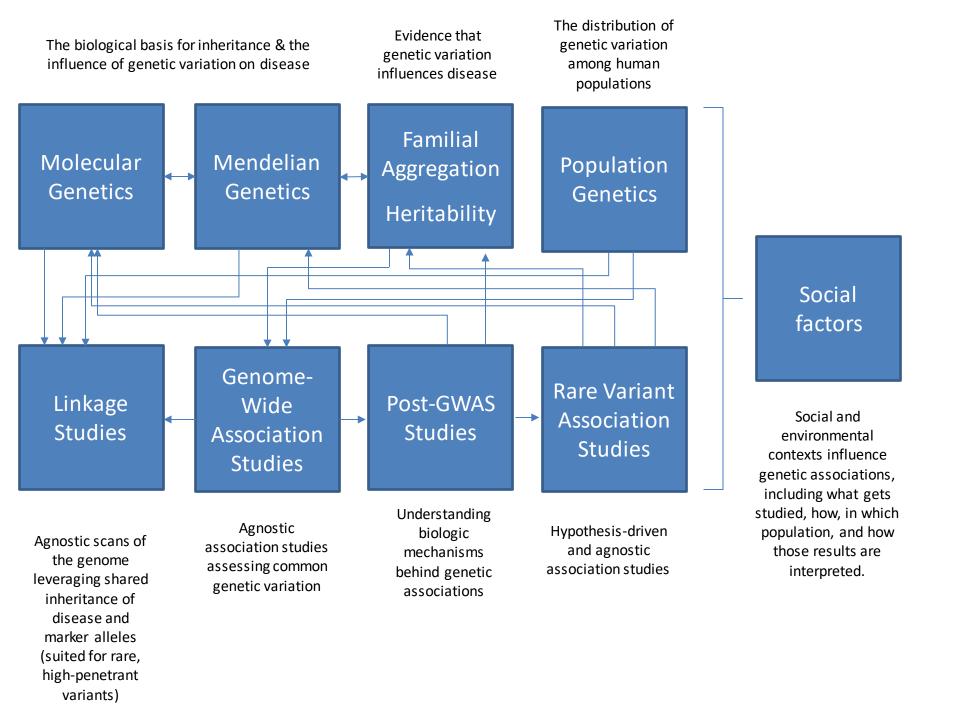
Morgan, Just, Franklin, Watson,
Crick, Hershey, Chase, Daly,
McClintock
(structure & role of DNA)
Collins, Lander
(the Human Genome Project, the
HapMap, high-throughput genotyping
and sequencing technologies)

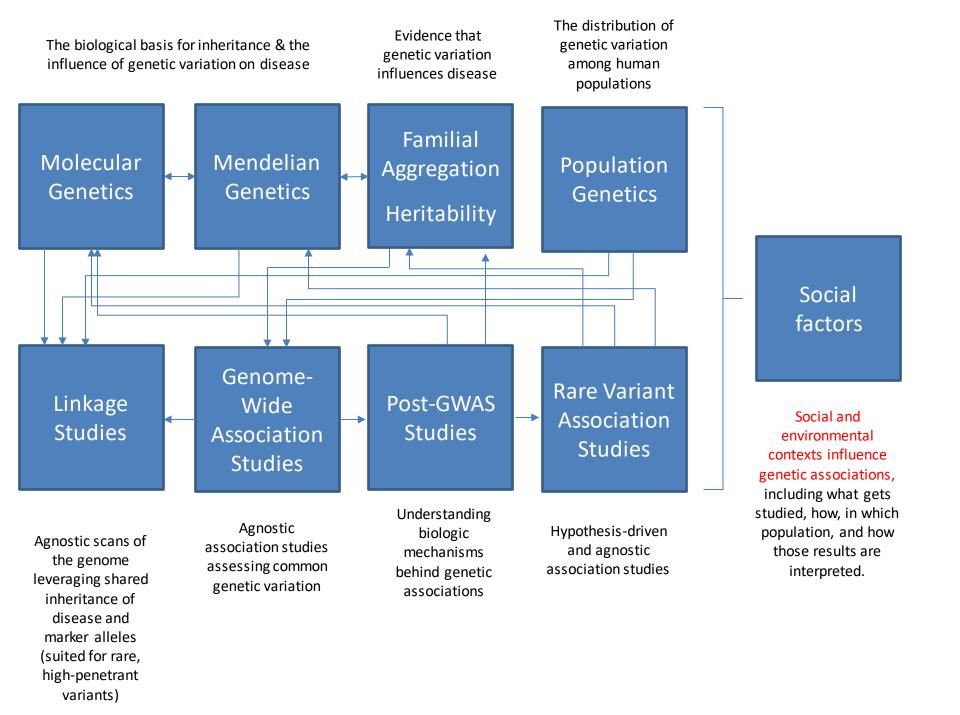
Molecular Genetics

## Genetic Epidemiology

"Understanding the function of much that is being revealed will not yield to classical Mendelian genetics but will require the epidemiological approach."

Genetic epidemiology is inseparable from "the concept of multifactorial causation."







#### At APOE

- ↑ LDL levels, statin use
- ↑ margarine use
- ↓ butter use

#### At LPA

- ↑ MI risk
- ↑ margarine use
- ↓ butter use

### At APOE

- 个 LDL levels, statin use
- ↑ margarine use
- ↓ butter use

### At LPA

- ↑ MI risk
- ↑ margarine use
- ↓ butter use

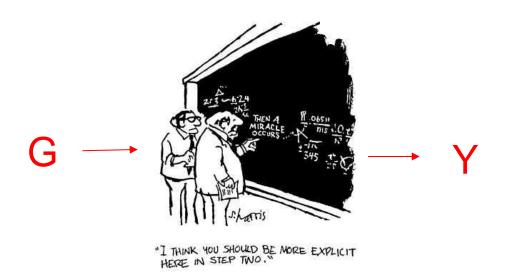
"Individuals who learn they have high cholesterol or have a family history of heart disease are more likely to switch to margarine instead of butter."

Take home messages

Social context matters

## Take home messages

- Social context matters
- Hypothesized biological mechanisms should be concrete & testable

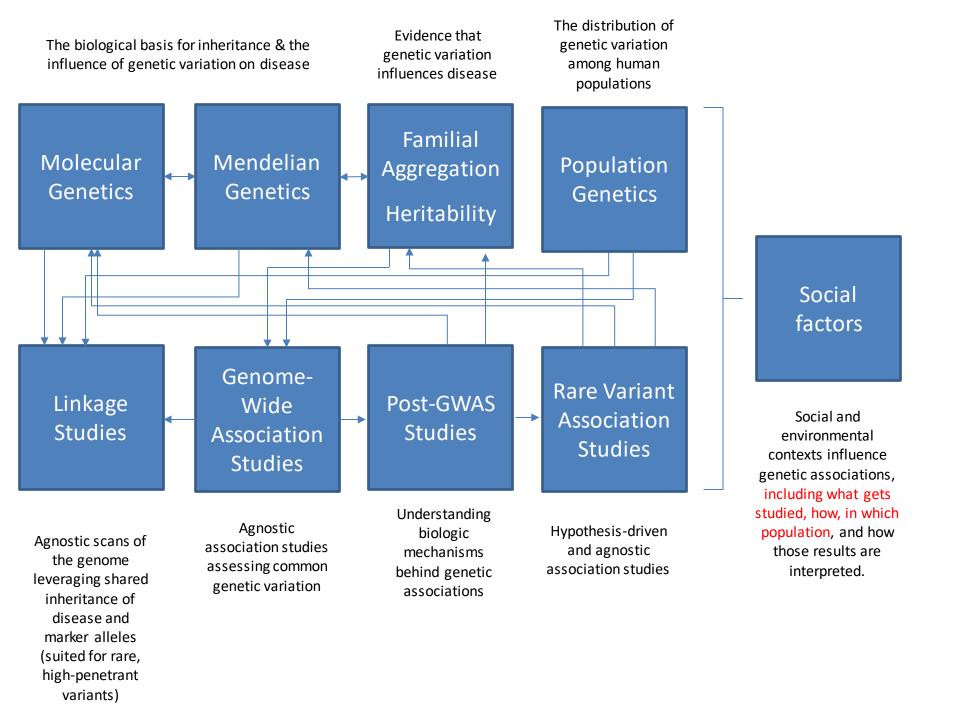


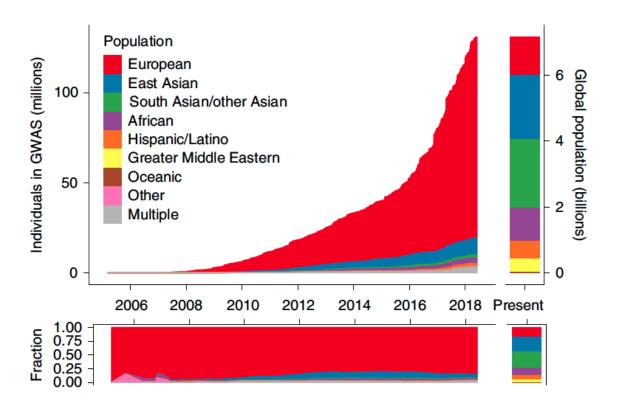
# Example: margarine vs butter use in the UK Biobank

### Take home messages

- Social context matters
- Hypothesized biological mechanisms should be concrete & testable

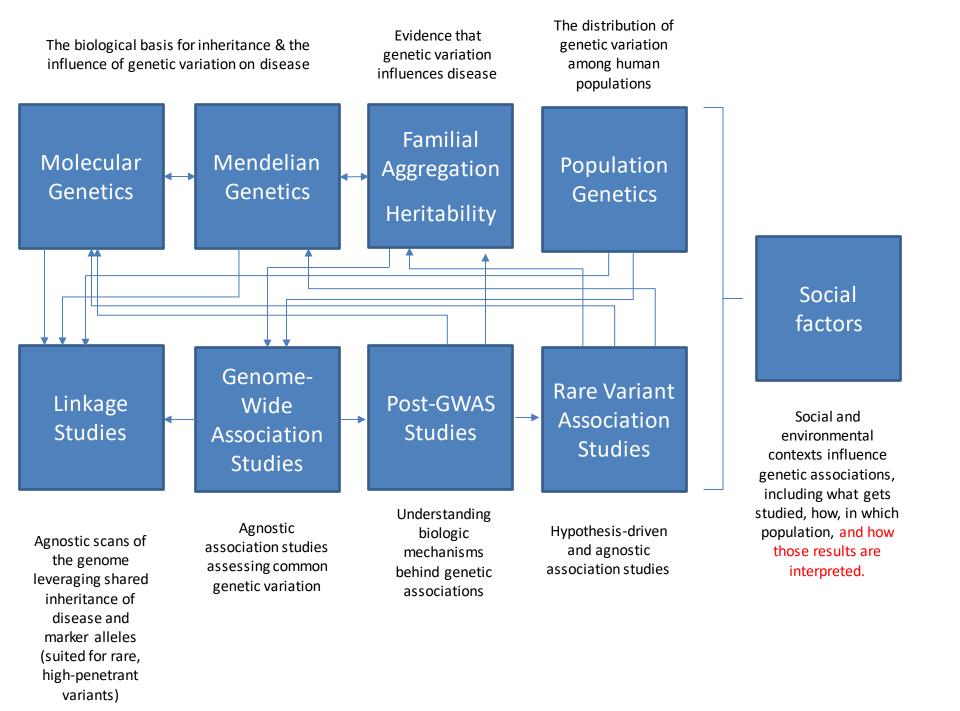
"Incorporation of disease biology involves a real understanding of the causal pathways leading from genes and environment to disease causation, and expression of these concepts in our mathematical models for penetrance."





**Fig. 1** | Ancestry of GWAS participants over time, as compared with the global population. Cumulative data, as reported by the GWAS catalog<sup>76</sup>. Individuals whose ancestry is 'not reported' are not shown.

This lack of representation limits equity and scientific opportunity.



The history of genetic epidemiology is a tapestry of observational science, statistical developments, animal and plant breeding experiments, molecular experiments, medicine, epidemiology, technology, and... prevailing ideas explaining and justifying social hierarchy.

## **Broader Scientific & Social Context**

Mendelian and Galton, Pearson, **Georges Cuvier** Statistical Genetic Blumenbach Genetics Fisher **Epidemiology** Louis Agassiz On the Natural Variety of many others many others "Understanding the Mankind function of much that is Clinical being revealed will not Genetics yield to classical Mendelian genetics but will require the epidemiological approach." Genetic Essentialist theories of "race" **Eugenics** epidemiology is Molecular inseparable from "the Genetics concept of multifactorial causation." 19<sup>th</sup> Century 18<sup>th</sup> Century 20<sup>th</sup> Century 21<sup>st</sup> Century **Imperialism Chattel Slavery** Revivial of debunked **Immigration Restrictions** eugenic theories: "Race Realism," Forced Sterilization "Human Biodiversity" Jim Crow & Segregation

### Why Talk about Eugenics?

- Shaky, pseudoscientific belief that empirical biological evidence exists to justify some individuals as inferior → Shouldn't forget this history → the "backdoor to eugenics" is slippery, and the history of eugenics is not "ancient history" at all
- Eugenic thinking is still very much alive → the historic time of civil unrest makes it all the more important that we keep our history in mind as we do our work
- Rather than assuming our data speaks for itself, we have a duty to consider the language, context, and framing of our work, so that it helps, not harms

Although founded on shaky arguments and weak empirical support, eugenics was firmly within the Western scientific establishment.

"The call is coming from inside the house!"

Setting the agenda in research

### Comment

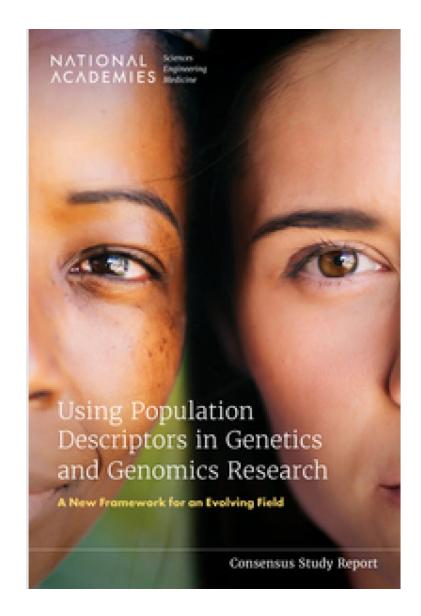


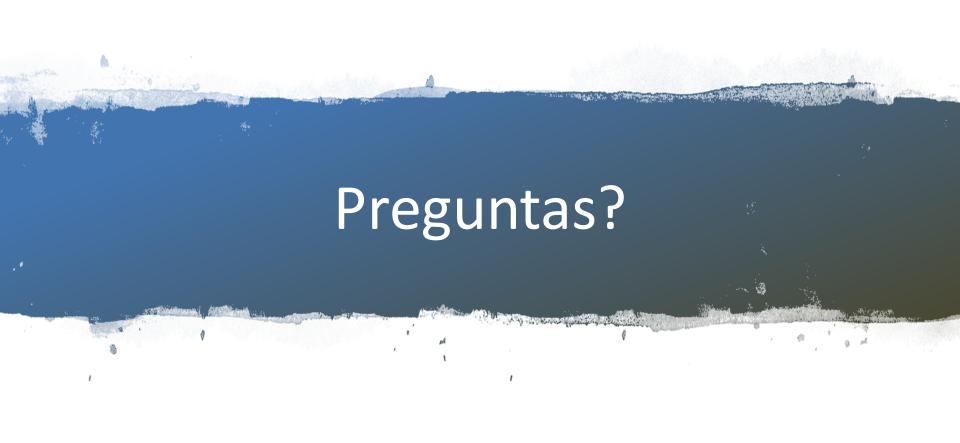
# Counter the weaponization of genetics research by extremists

Jedidiah Carlson, Brenna M. Henn, Dana R. Al-Hindi & Sohini Ramachandran



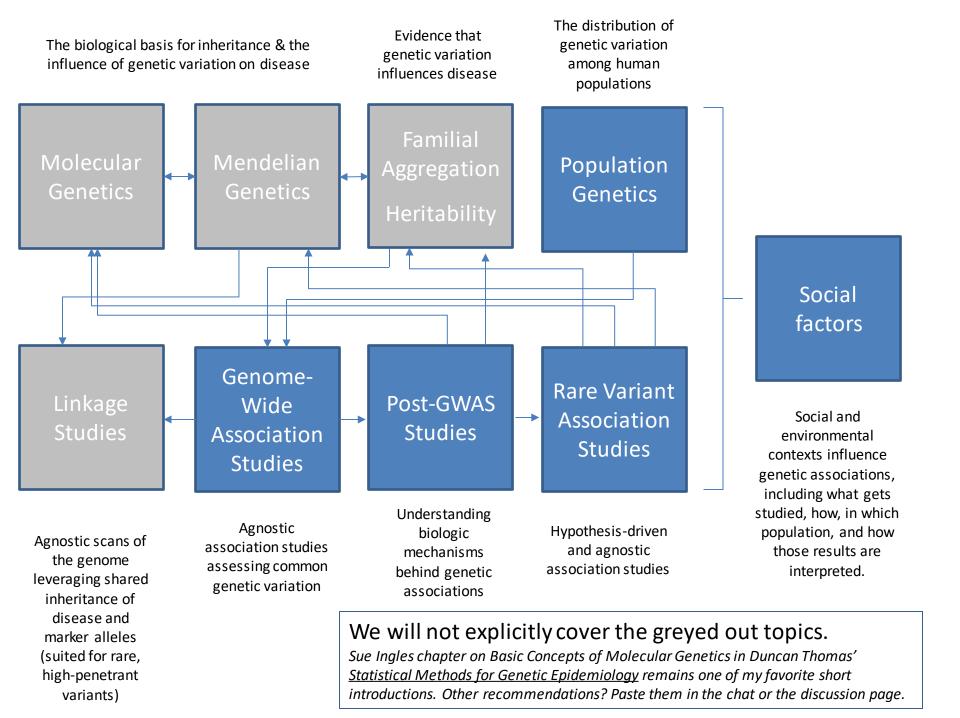
A memorial to the ten Black people who were killed by a shooter outside a shop in Buffalo, New York, in May 2022.

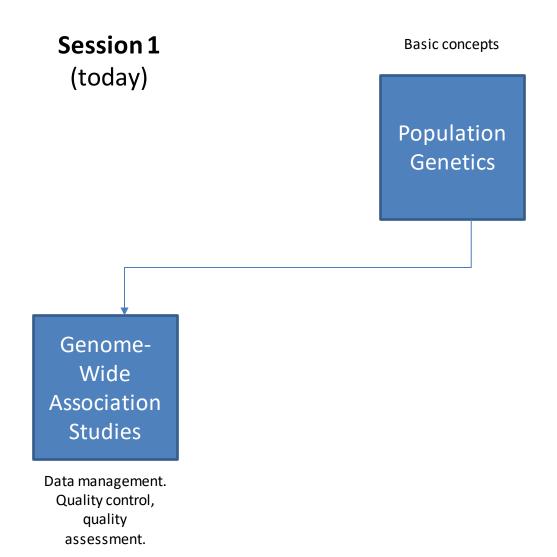




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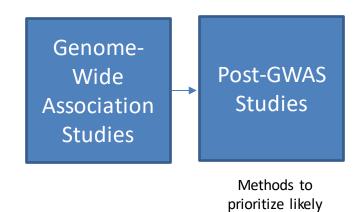


(September 20)

Genome-Wide Association Studies

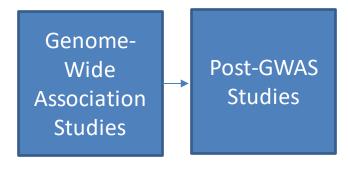
Basic GWAS testing. Meta-analyses. Visualizations. Annotating GWAS results.

(September 27)

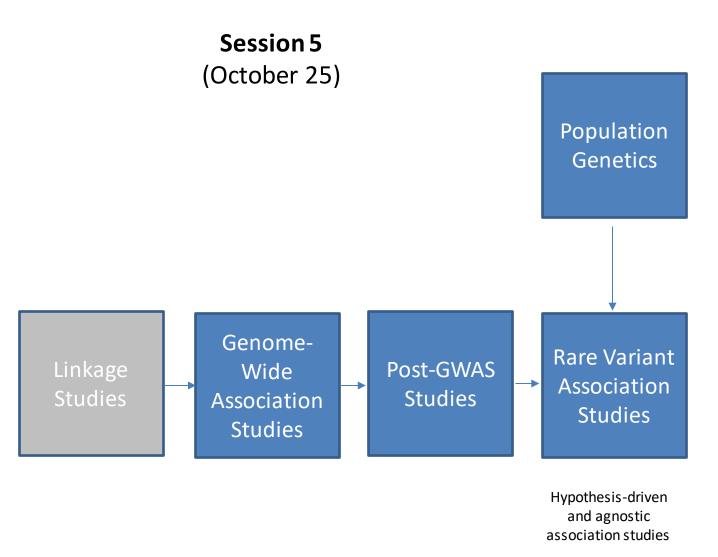


causal *variants*.
Statistical fine mapping, colocalization.

(October 25)

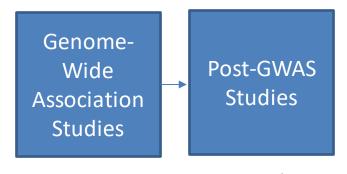


Genetic
architecture:
distribution of
effect sizes,
number of causal
alleles, functional
enrichment.
Polygenic risk
models.



### Xihao Li Aubrey Hubbard

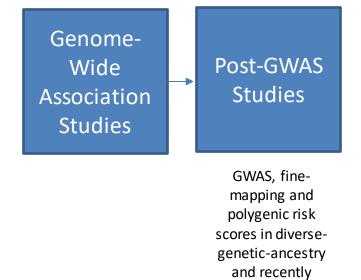
(November 15)



Integrative methods (TWAS, PWAS etc.) to prioritize causal genes.
Mendelian Randomization.

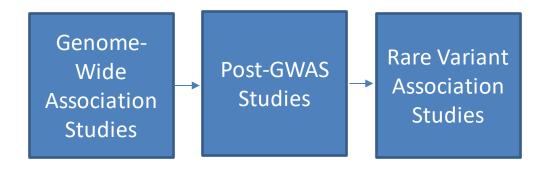
Guanghao Li Diptavo Dutta Sheila Rajagopal Aubrey Hubbard

(November 29)



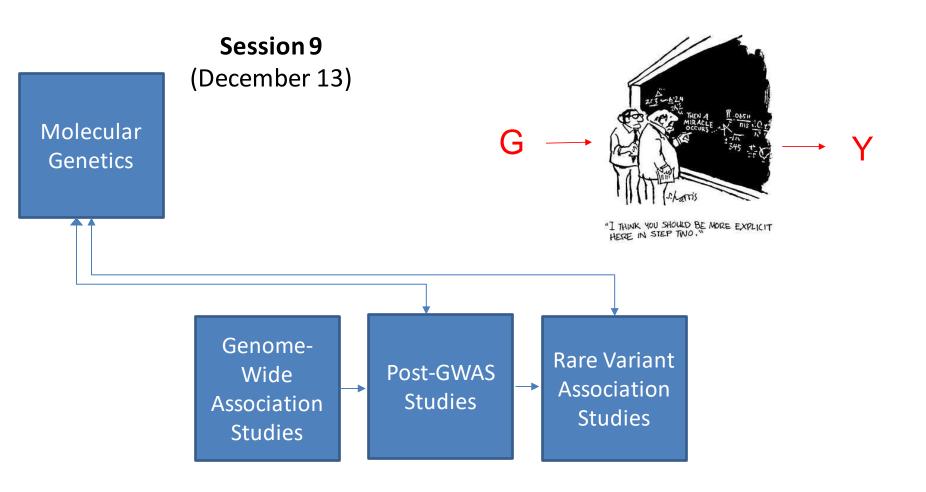
admixed samples

(December 6)



Genetic mosaicism and clonal hematopoiesis

Mitch Machiela Weiyin Zhou Sheila Rajagopal Aubrey Hubbard



**Functional Genomics** 



# Agenda

- Course format
- Intro to genetic epidemiology
- Course scope
- A few basic definitions

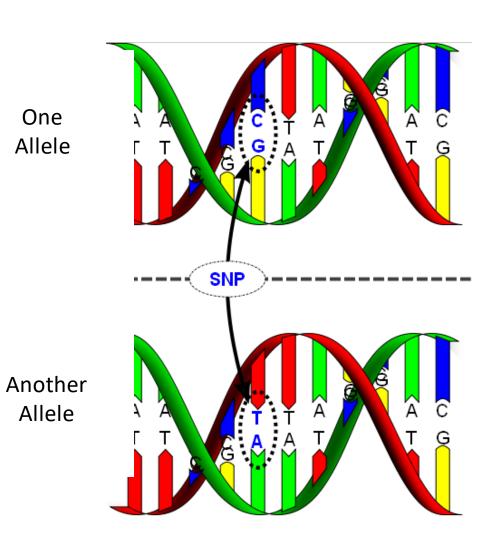
## A few basic definitions

- Single nucleotide polymorphisms (SNPs)
- Genotype and allele frequencies (heterozygosity)
- Hardy-Weinberg equilibrium
- Linkage disequilibrium (LD)
- Variation in allele frequencies and linkage disequilibrium

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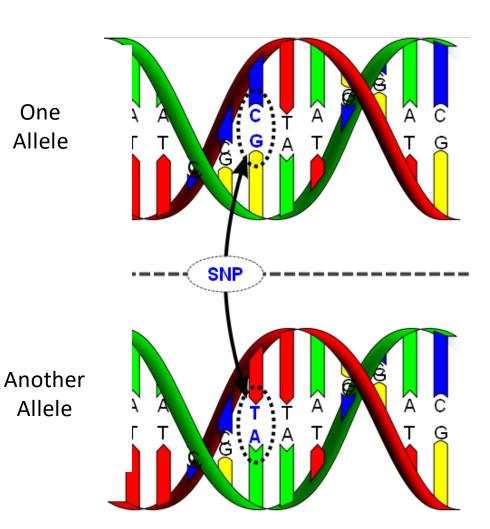
# Single Nucleotide Polymorphisms



## Single Nucleotide Polymorphisms

One Allele

Allele



SNPs almost always have only two alleles.

A SNP like this one is sometimes written as "a C/T SNP" ...

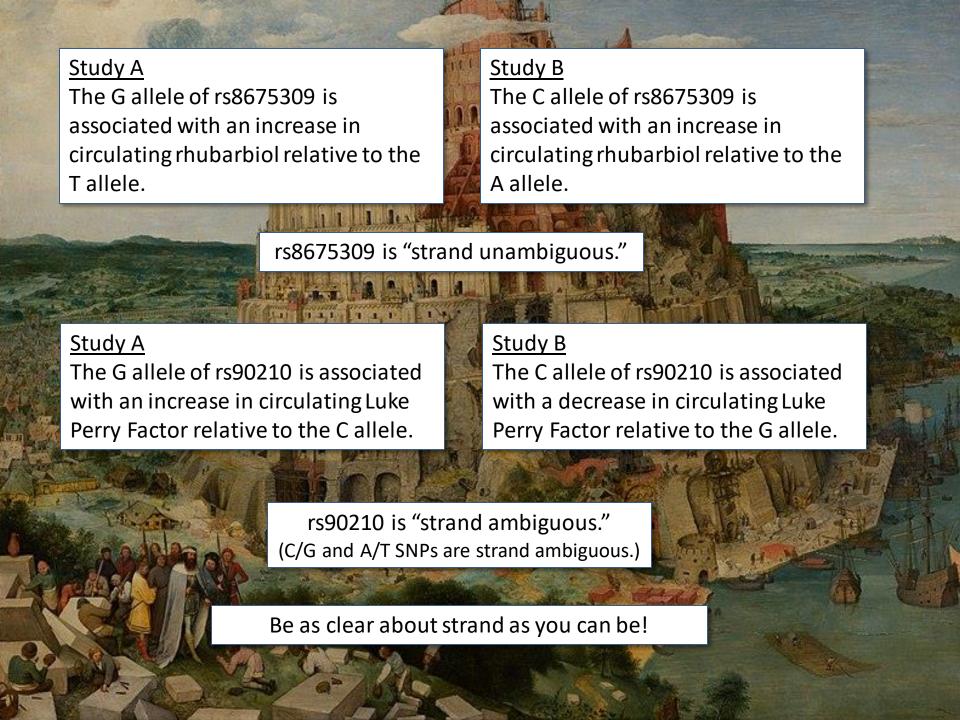
and sometimes written as "a G/A SNP."

To avoid ambiguity you have to specify which strand you are using to define the alleles.

Often the (+) strand of the human reference genome is used. The (+) strand runs 5' to 3,' starting at the telomere of the p arm of each chromosome. The current version of the reference genome is GRCh38/hg38.

SNPs are often referred to by their dbSNP Reference SNP (refSNP) number, or by chromosome and position. E.g. rs671 and chr12:111803962:G:A refer to the same variant.

> https://en.wikipedia.org/wiki/Reference\_genome https://www.ncbi.nlm.nih.gov/grc/human



## A few basic definitions

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# Definitions and an example: the ABO blood group

- An <u>allele</u> is the nucleotide sequence at a polymorphic locus (i.e. a small region of the genome, e.g. a single base or a gene) on a chromosome.
- While there are (typically) only at most 2 different alleles carried by an individual at a locus, there can be more than 2 alleles in the population.
- The combination of alleles at a locus carried by an individual is that individual's genotype.
- A familiar example is the ABO blood group, defined by presence of antigens (i.e. structures that induce an immune response) on red blood cells (RBCs).
- The genetics of ABO blood type maps to a single locus. Alleles A and B are codominant with each other and dominant over allele O (no antigen), such that the possible genotypes, i.e. the pairwise combinations of ABO alleles, give rise to phenotypes A, B, AB, or O.

	Group A	Group B	Group AB	Group O
Red blood cell type		<b>B</b>	AB	0
Antibodies in Plasma	Anti-B	Anti-A	None	Anti-A and Anti-B
Antigens in Red Blood Cell	P A antigen	† B antigen	A and B antigens	None

## Genotype frequencies in the population

A diploid individual (human) has two alleles (maternal and paternal) at an autosomal locus. Assuming there are K distinct alleles in the population,

Let  $n_{ij}^*$  be the number of individuals with genotype ij,

$$i=1..K$$
,  $j=i..K$ .

Then the genotype frequency  $P_{ij} = n_{ij}/n$ ,

n is the total number of individuals in the population

$$n=\sum n_{ij}$$

<sup>\*</sup>NB. The order of i and j is usually arbitrary in this notation such that  $n_{ij}$  and  $n_{ji}$  refer to the same individuals

## Multi allele example: ABO genotypes

$n_{11} = n_{OO}$	1,168
$n_{12} = n_{\mathrm{OA}}$	1,080
$n_{13} = n_{\mathrm{OB}}$	377
$n_{22} = n_{AA}$	262
$n_{23} = n_{AB}$	186
$n_{33} = n_{\mathrm{BB}}$	44
$n_{total}$	3,117

$P_{11} = P_{\text{OO}}$	0.37
$P_{12} = P_{\mathrm{OA}}$	0.35
$P_{13} = P_{\mathrm{OB}}$	0.12
$P_{22} = P_{\mathrm{AA}}$	0.08
$P_{23} = P_{AB}$	0.06
$P_{33} = P_{\mathrm{BB}}$	0.01
	1.00

## Two allele example: MTHFR\* C677T

MTHFR encodes methylene tetrahydrofolate reductase, an enzyme relevant to chemotherapy with 5-fluorouracil. There are two alleles, C or T, nucleotide position 677 in the gene. These alleles encode different amino acids in the protein at residue 222, either alanine (C) or valine (T), resulting in effects on protein stability, either stable (C) or thermolabile (T).

#### count of genotypes

$n_{11} = n_{\rm CC}$	334
$n_{12} = n_{\rm CT}$	350
$n_{22} = n_{\mathrm{TT}}$	82
$n_{total}$	766

#### frequency of genotypes

$P_{11} = P_{\rm CC}$	0.44
$P_{12} = P_{\mathrm{CT}}$	0.46
$P_{22} = P_{\mathrm{TT}}$	0.10
	1.00

## Allele frequencies in the population

"Allele frequencies" are the proportions of chromosomes in the population with each of the unique alleles at the variable locus.

Let  $m_i$  be the number of copies of allele i in the population.

Then the allele frequency  $p_i$  is  $m_i/m$ , where m=2n is the total number of chromosomes carrying the variable locus in the population of n individuals.

$$p_i = (2n_{ii} + \sum_{j \neq i} n_{ij})/(2n) = (2P_{ii} + \sum_{j \neq i} P_{ij})/2.$$

# Example: ABO genotypes to allele frequencies

$n_{11} = n_{OO}$	1,168
$n_{12} = n_{\mathrm{OA}}$	1,080
$n_{13} = n_{\mathrm{OB}}$	377
$n_{22} = n_{AA}$	262
$n_{23} = n_{AB}$	186
$n_{33} = n_{\mathrm{BB}}$	44
$n_{total}$	3,117

$p_1 = p_{\mathrm{O}}$	0.61
$p_2 = p_A$	0.29
$p_3 = p_B$	0.10
	1.00

# Example: MTHFR C677T genotypes to allele frequencies

$n_{11} = n_{\rm CC}$	334
$n_{12} = n_{\mathrm{CT}}$	350
$n_{22} = n_{\mathrm{TT}}$	82
$n_{total}$	766

$p_1 = p_C$	.66
$p_2 = p_{\mathrm{T}}$	.34
	1.00

# Notation for alleles for the most common type of variation: diallelic SNPs (K=2)

Single nucleotide polymorphisms (SNPs) – the most common type of variation, typically have 2 unique alleles, i.e. diallelic

For diallelic SNPs, we often use notation that substitutes p and q for  $p_1$  and  $p_2$ . Since  $p_1+p_2=1$ ,

$$p_1 = 1 - p_2 = p = 1 - q$$

However, we still have to specify to which allele p refers! Typically, p refers to the minor allele in the population, so p < 0.5.

People also often use "MAF" = minor allele frequency = p

#### The "N" for genotypes v. alleles

#### Genotypes

The number (i.e. N) of genotypes at a polymorphic locus in a sample is equal to the number of individuals.

#### Alleles

The number (i.e. N) of alleles is the number of chromosomal regions at a polymorphic locus in the population or, for autosomes, **2x the** number of individuals.

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#### Hardy-Weinberg Equilibrium (HWE)

In 1908, Godfrey Hardy and Wilhelm Weinberg independently derived a <u>formula</u> relating the allele frequencies in parents to genotype frequencies in offspring when they achieve <u>equilibrium</u> status.

Formally, HWE occurs under the following assumptions.

- Random mating
- No inbreeding
- Infinite population size
- Discrete generations
- Equal allele frequencies in males and females
- No mutation, migration, or selection.

Even when these assumptions do not hold exactly, HWE often provides a good (and useful) approximation for population genotype frequencies.

#### The Hardy-Weinberg Formula

Using notation from the "Alleles and Genotypes" vignette, the HW formula states that, in general,  $P_{ij} = 2 p_i p_j$  if  $i \neq j$ , and  $p_i^2$  if i = j. For the diallelic case:

$$P_{11} = p^2$$
 $P_{12} = 2 p q$ 
 $P_{22} = q^2$ 

That is, genotype probabilities are given by a multinomial distribution with parameters  $(2, p_1, ..., p_K)$ , or, in the special case of diallelic markers, a binomial\* distribution with parameters (2,p).

[\*Recall the binomial distribution:  $(p + q)^2 = p^2 + 2pq + q^2$ ]

#### We can use HWE to...

- Test for genotyping errors
- Simplify calculations, e.g. in simulations
- Test for underlying deviations from HWE assumptions
  - Random mating
  - No inbreeding
  - Infinite population size
  - Discrete generations
  - Equal allele frequencies in males and females
  - No mutation, migration or selection
- Test for marker-disease association in cases only for some inheritance modes (e.g. dominant, recessive) but low power

#### A few basic definitions

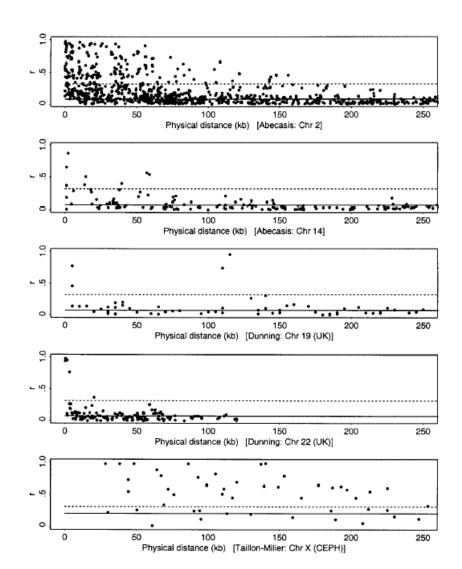
- Single nucleotide polymorphisms (SNPs)
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#### Linkage disequlibrium (LD)

- What are we talking about?
  - The correlation between genetic markers in a population, how it is measured or estimated, and how it is quantified.
  - Typically measured using r<sup>2</sup> (see appendix)
  - High r<sup>2</sup> between two SNPs requires both have similar MAFs
- What are the underlying biological mechanisms?
  - Mutation and recombination in meiosis.
  - Selection
  - Demographics
- Why study?
  - To assess the number of independent tests in the genome.
  - To identify a minimum number of variants that retain coverage of genetic variation in the genome, i.e. through correlation structure.
  - To impute markers that are not directly genotyped
  - To help understand the origins and biological properties of genetic variation in the genome.

#### Decay of LD: r as a function of physical distance

5 chromosomal regions (2001)



Am. J. Hum. Genet. 69:1-14, 2001

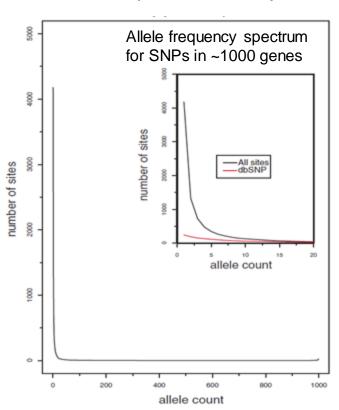
**Figure 3** Plots of  $\hat{r}$ , as a function of physical distance (in kb), for SNP data from five regions (Dunning et al. 2000; Taillon-Miller et al. 2000; Abecasis et al. 2001). On each plot, points above the unbroken line are in significant LD at the .05 level, and points above the dotted line correspond to what Kruglyak (1999) has called "useful LD"; these lines are set at r = .316, the equivalent of  $r^2 = .1$ .

#### A few basic definitions

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#### Distribution of allele frequencies

#### In a sample of 697 subjects



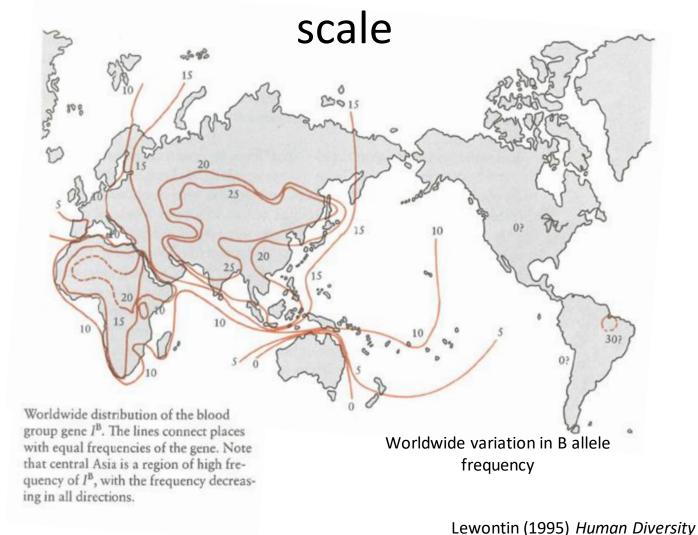
Most SNPs are rare (MAF<1%)

But, most of the differences between unrelated individuals are due to common variants (MAF>5%).

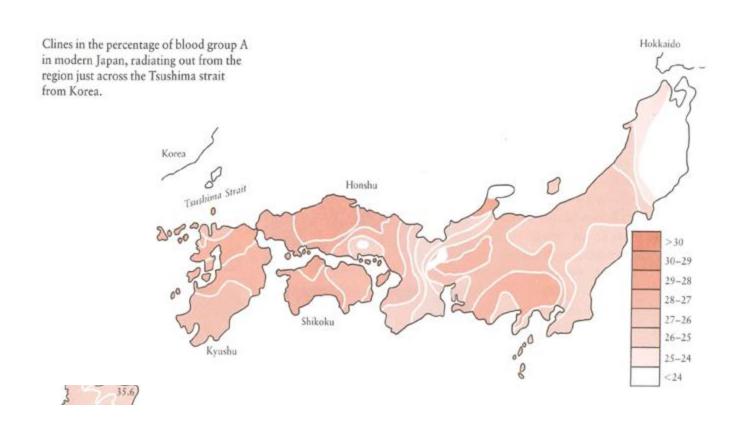
dbSNP is a database to record all SNPs seen in genetic studies (at NCBI/NIH)

# Allele frequencies at individual loci differ across geographically-defined populations, more or less in a smooth gradient

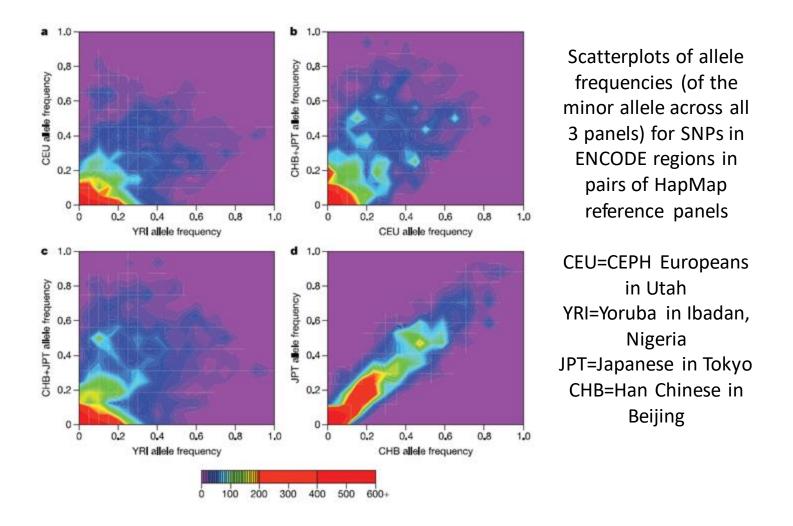
#### Allele frequency variation at the global



### Allele frequency variation at the (more) local scale



### Allele frequency across 4 reference populations by sequencing targeted regions of the genome\*



International HapMap Consortium (2005) Nature

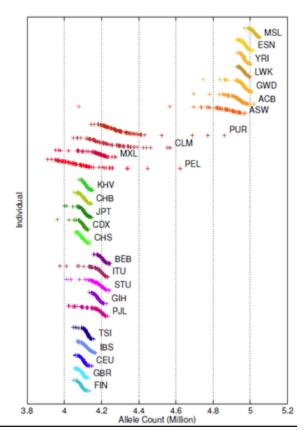
\*regions from the ENCODE project; reference populations from the HapMap project

#### <u>Critical concept for genetic epidemiology of humans</u>:

There is much greater variation, e.g. more SNPs, among African populations than others, consistent with out of Africa hypothesis

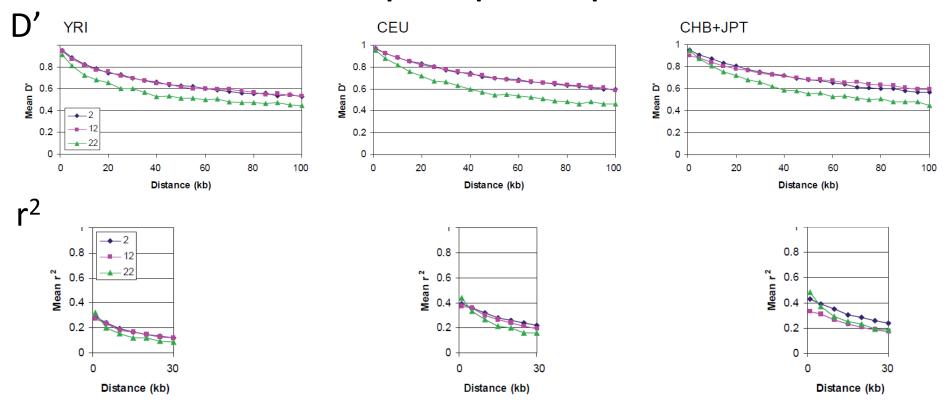
#### Variants per genome

Туре	Variant sites / genome
SNPs	3.8 * 10 <sup>6</sup>
Indels	5.7 * 10 <sup>5</sup>
Mobile Element Insertions	~1000
Large Deletions	~1000
CNVs	~150
Inversions	~11



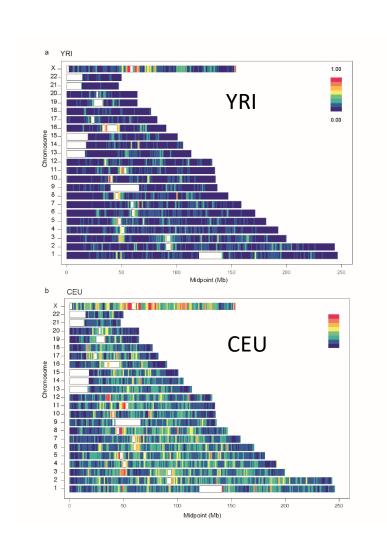
### Linkage disequilibrium also varies across populations

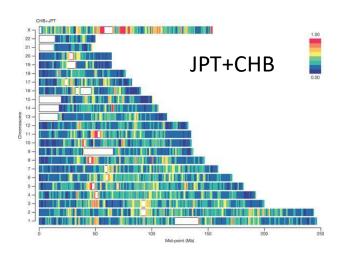
### Differences in decay of LD with distance in HapMap samples



- Mean pairwise D' (top) or r² (bottom) by SNP-SNP distance for chromosomes 2 (long),
   12 (medium), or 22 (short)
- LD persists longer in CEU (European) and CHB+JPT (Asian) than in YRI (African), especially visible for r<sup>2</sup> measure
- See Nature 437:1299 (2005), figure S6

#### Much less LD in African samples





Authors fit model for the local amount of LD decay (as r<sup>2</sup> in colors) in 30kb in YRI, CEU, JPT+CHB

Much lower values of r<sup>2</sup>, i.e. more LD decay per 30kb in YRI than in the others

See Nature 437:1299 (2005),

## What are some potential reasons/mechanisms for the wide variation in allele frequency and linkage disequilibrium?

- A. Stochastic drift with time
- B. Population migration, isolation, bottlenecks, or founder effects
- C. Selection phenomena positive or negative
- D. Non-random mating
- E. Mutation

The time scales for these effects are long, and they can be technically difficult to follow longitudinally. Nevertheless, we can infer something about the past from distributions of current allele frequency.

These differences in allele frequencies and linkage disequilibrium patterns can be used to map samples in terms of genetic similarity. These "map coordinates" can then be used to perform quality control or as covariates to account for potential confounding.

Model-based approaches: STRUCTURE, GRAF-pop Model-free approach: PCA

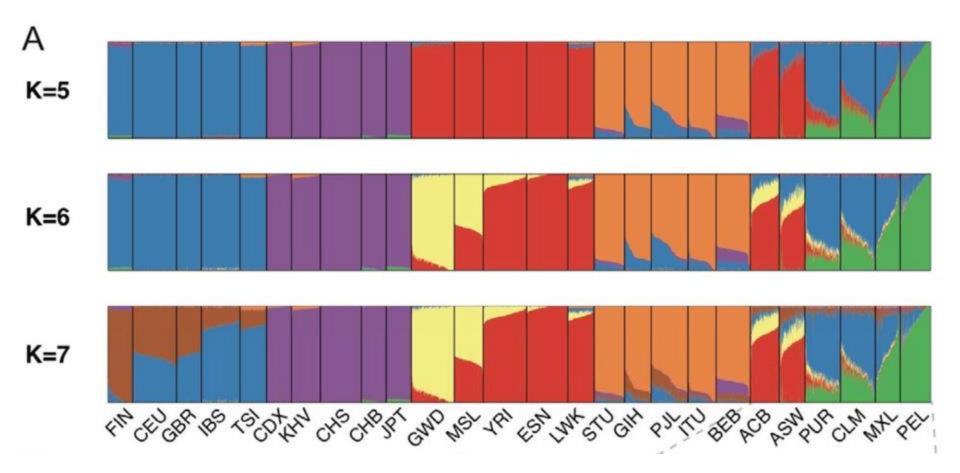
#### STRUCTURE (GRAF-Pop uses a similar conceptual model)

STRUCTURE calculates the proportion of each individual's genome that "comes from" each of K subpopulations (assuming each is in HWE).

- Underlying model: each allele in an individual can be traced back to one of K ancestral populations
- Under this model, we can estimate the proportion of the individual's alleles that come from each ancestral population.
- Can be run in a supervised fashion, where reference populations are the K subpopulations
- Can be run in an unsupervised fashion, where the subpopulations are inferred from the data, and these are calculated for different values of K (various heuristics for selecting the "optimal" K)

Rosenberg, 2002, Science

#### STRUCTURE on 1KG



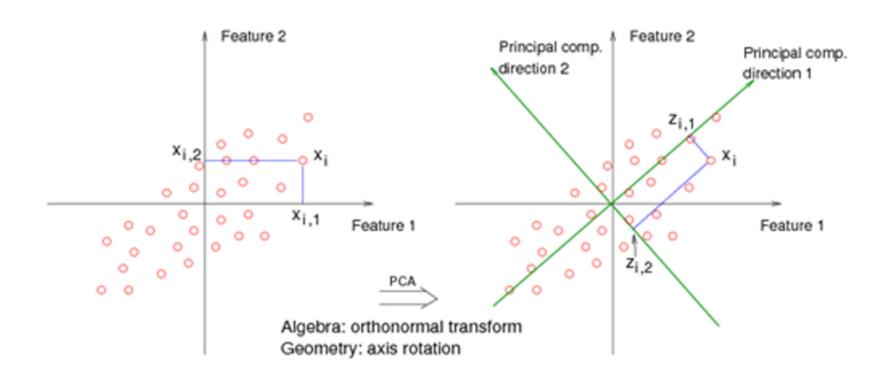
#### STRUCTURE

Don't overinterpret these proportions! They depend on the reference samples (or the diversity in the sample at hand). And the "populations" do not necessarily imply shared ancestry.

Further reading: A tutorial on how not to over-interpret STRUCTURE and ADMIXTURE bar plots, Lawson et al, Nature Communications, 2018

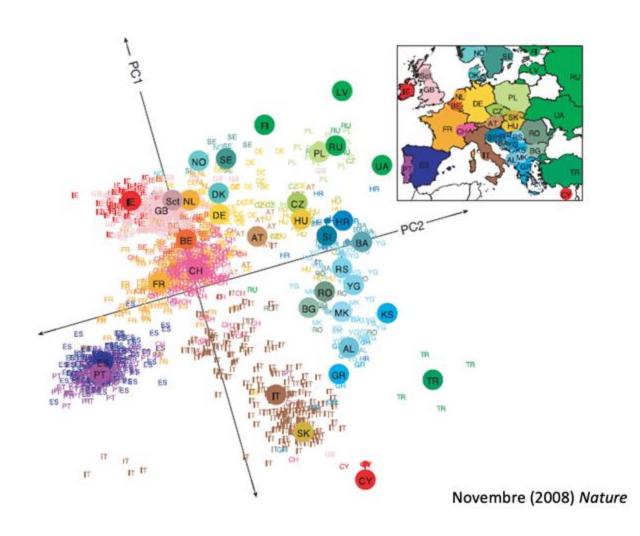
#### Principal Components (PCs)

Principal components analysis is a general statistical technique that can be used to summarize variation in many variables using a few key summary variables.



#### Principal Components (PCs) applied to genetic data

When applied to genetic data, PCA can capture latent structure that accounts for patterns of genetic similarity in a set of samples

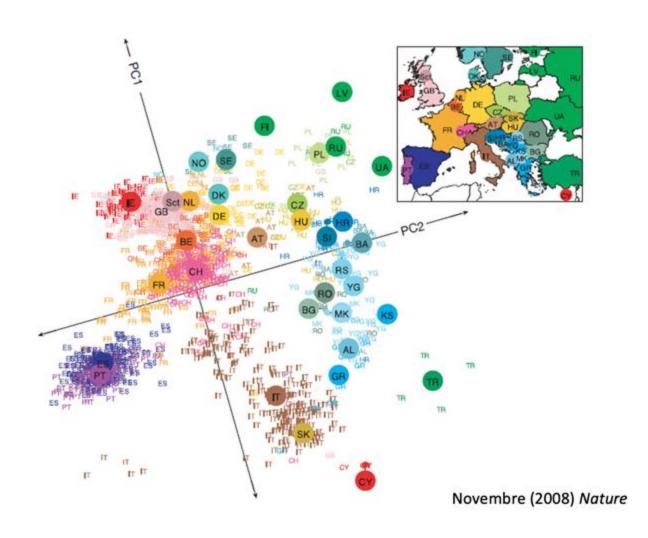


#### Principal Components (PCs) limitations

Nota bene: the picture from the Novembre paper is (by design) "too pretty"

#### Methods:

"We used a 'strict consensus' approach: if all observed grandparents originated from a single country, we used that country as the origin. If an individual's observed grandparents originated from different countries, we excluded the individual."



#### Principal Components (PCs) limitations

Remember PCs are just a summary of the data!

- Who is in the data?
- Who is not in the data?

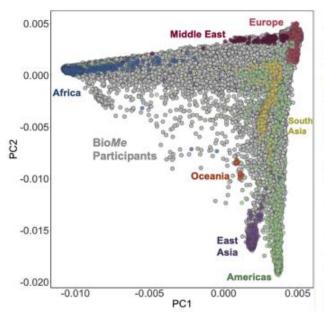


Figure 2. The continuous, category free, nature of genetic variation. Reproduced from Belbin et al. Towards a fine-scale population health monitoring system. This image shows individuals projected onto the first two principal components of genetic similarity. Colored dots are N=4149 reference panel individuals from 87 populations representing ancestry from 7 continental or subcontinental regions. Gray dots are N=31705 participants from BioMe, a diverse biobank based in New York City. Clearly delineated continental ancestry categories, the islands of color, are shown to be a by-product of sampling strategy. They are not reflective of the diversity in this real-world dataset, made evident by the continuous sea of gray.

#### Enough talk... on to hands-on practice!

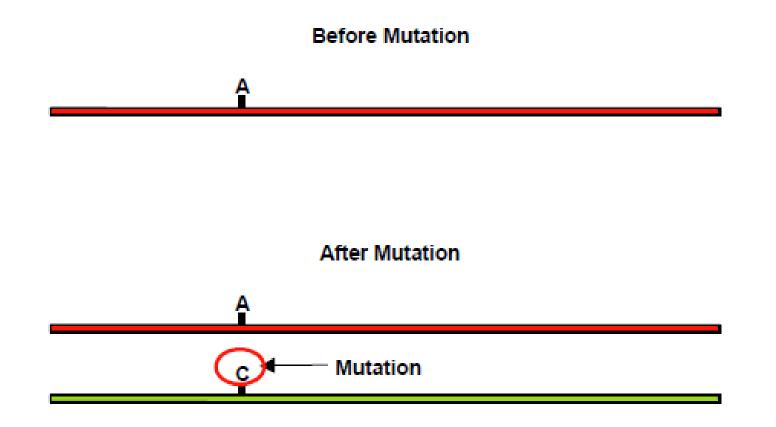


Appendix: details on measures of LD

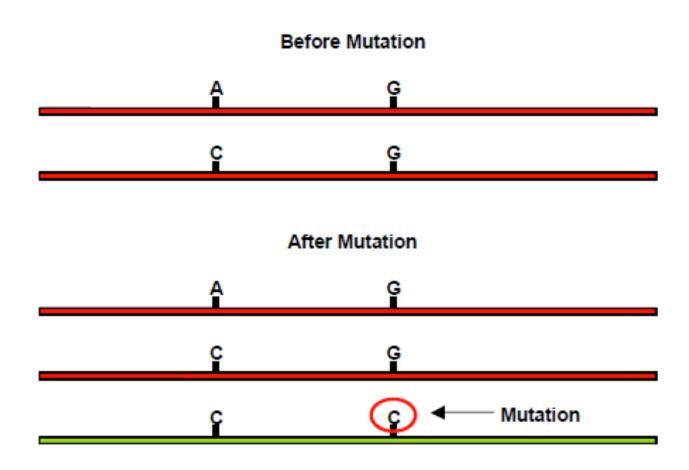
Part 1: Haplotypes and their relationship to LD

#### Origins of LD

Alleles that exist today reflect ancient genetic events ... first one mutation arises



#### ... then the another allele...

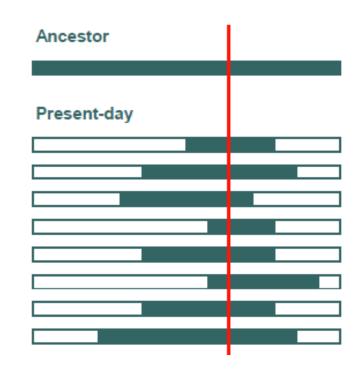


# ... while recombination (from meiosis) generates new arrangements for ancestral alleles



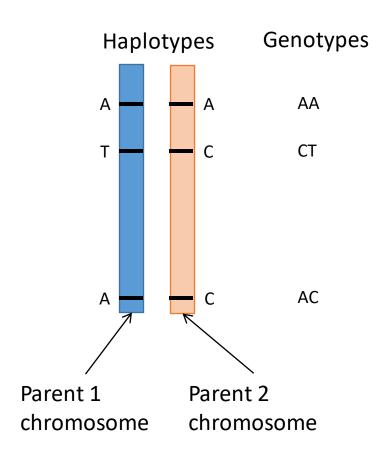
### Chromosomes are mosaics of ancestral haplotypes

- Extent and conservation of mosaic pieces depends on
  - Recombination rate
  - Mutation rate
  - Population size
  - Natural selection
- Combinations of alleles at very close markers tend to reflect individual ancestral <u>haplotypes</u>
- Longer range combinations of alleles reflect recombination of ancestral <u>haplotypes</u>



#### Haplotypes formally defined

Haplotypes are the ordered arrangements of alleles at specific loci <u>along the same chromosome</u>. They are <u>phased</u> genotypes, that is, the alleles as they are passed down from parents.



### Haplotypes are typically not directly observed by genotyping

- Genotype data typically does not include information about phase of alleles
- Can determine by long-range single molecule sequencing but experimentally challenging
- More typically, have to infer haplotypes from genotype data Example --
  - Genotypes Aa/Gg consistent with haplotype pairs
  - Can estimate haplotypes in a probabilistic framework

e.g. 
$$A \mid a$$
 with probability=0.9  $A \mid a$  with probability=0.1  $A \mid a$ 

- Most popular algorithm: Expectation Maximization<sup>1</sup>
  - Start with a guess for haplotype frequencies
  - Given this guess, calculate expected haplotype counts, assuming haplotypes are in HWE
  - Use these counts to update haplotype frequency estimates
  - Repeat until convergence

<sup>&</sup>lt;sup>1</sup> Thomas pp. 243-245

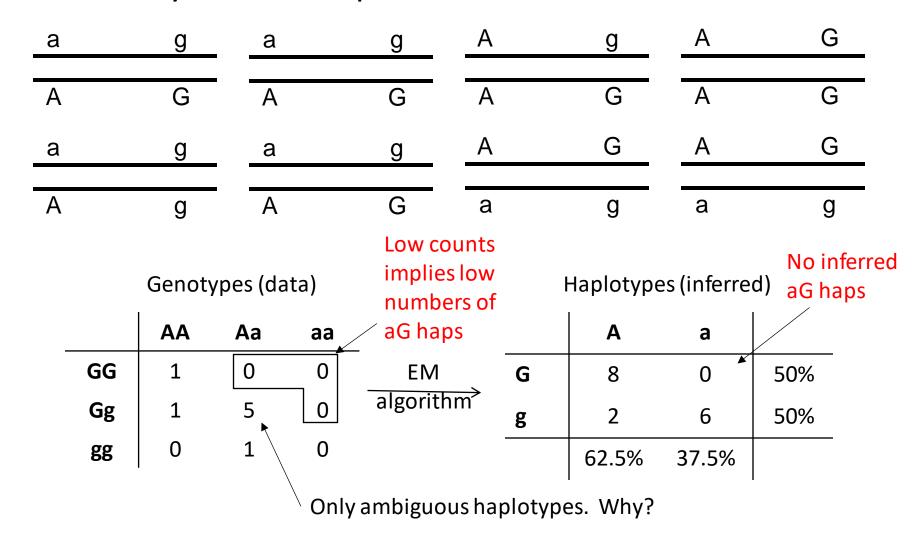
#### Comment about the definition LD

- Two loci are described as being LD if:
  - A. Alleles at the two loci are not independently distributed
  - B. The two loci are linked (recombination rate  $\theta < \frac{1}{2}$ )
- More precisely
  - Condition (A) is called "gametic disequilibrium"
  - Condition (B) is called "linkage"
  - Often people say LD when they mean gametic disequilibrium
- Thus, the standard measures of LD are really measures of gametic disequilibrium
- However, we usually do not compute LD for markers at different chromosomes, and the terms are equivalent for markers on the same chromosome

### Part 2: Estimating and quantifying LD

### First step in quantifying LD: From genotypes to haplotypes

Two nearby loci in sample of 8 individuals



### $\delta$ (aka D): a measure of LD related to marginal allele frequencies

• Consider  $\delta$  defined as follows:

	A	a	
G	$p_{AG} = p_A p_G + \delta$	$p_{aG} = p_a p_G - \delta$	$p_{G}$
g	$p_{Ag} = p_A p_g - \delta$	$p_{ag}=p_ap_g+\delta$	$p_g$
	p <sub>A</sub>	p <sub>a</sub>	1

- p<sub>A</sub> & p<sub>G</sub> are frequencies of alleles A & G
  - $p_a = 1-p_A$ ;  $p_g = 1-p_G$  allele frequencies of alleles a & g
- $\delta$  is a measure of departure from independence
  - No association between A and G  $\Rightarrow$   $\delta$  = 0
  - The four cells in this table must be between 0 and 1, so:
  - $\delta_{min}$  = -min( $p_A p_G$ ,  $p_a p_g$ ) and  $\delta_{max}$  = min( $p_A p_g$ ,  $p_a p_g$ )
- $\delta$  also sometimes termed "D"

### Why is $\delta$ a measure of dependence of alleles at two loci?

- Basic idea: two random variables are independent if knowing one gives you no information about the other, i.e. conditioning on one does not change the distribution of the other
- The mathematical definition of independence is Pr(A,G)=Pr(A)Pr(G)
- Why is this a good definition? Well, if independent

$$Pr(A \mid G) = \frac{Pr(A,G)}{Pr(G)} = \frac{Pr(A)Pr(G)}{Pr(G)} = Pr(A)$$

•  $\delta$  is just Pr(A,G)-Pr(A)Pr(G), which equals 0 if no LD.

### D' and $r^2$ : more interpretable LD measures related to $\delta$

 $\delta$  depends on allele frequencies and may be positive or negative making it difficult to interpret.

Since the marginal frequencies  $p_A = p_{AG} + p_{Ag}$  and  $p_G = p_{AG} + p_{aG}$ :  $\delta = p_{AG} - p_A p_G = p_{AG} p_{ag} - p_{Ag} p_{aG}$ 

#### Consider:

Measure	Formula	
$D' = \delta/\max(\delta)$	$\frac{p_{AG}p_{ag}-p_{Ag}p_{aG}}{\min(p_Ap_G,p_ap_g)}$	if δ<0
	$\frac{p_{AG}p_{ag}-p_{Ag}p_{aG}}{\min(p_Ap_g,p_ap_G)}$	if δ>0
$r^2$ (aka $\Delta^2$ )	$\frac{\left(p_{AG}p_{ag} - p_{Ag}p_{aG}\right)^2}{p_Aq_Ap_Gq_G} =$	$\left(\frac{p_{AG} - p_A p_G}{\sqrt{p_A q_A p_G q_G}}\right)^2$

#### Comments on D' and r<sup>2</sup>

- abs(D') = 1 is taken as evidence for no recombination, in the absence of other influences on LD.
- The LD measure "r" is **identical** to the Pearson correlation between alleles at locus X (e.g. A or a) and Y (e.g. G or g), i.e.

$$r = \sqrt{\frac{1}{N} \sum \frac{(X_i - \overline{X})(Y_i - \overline{Y})}{\sigma_X \sigma_y}},$$

where:

$$X_i = allele \ at \ locus \ X = \begin{cases} 0 \ if \ X = A \\ 1 \ if \ X = a \end{cases}$$
 $Y_i = allele \ at \ locus \ Y = \begin{cases} 0 \ if \ Y = G \\ 1 \ if \ Y = g \end{cases}$ 

• Computation of LD requires estimating frequency of two-locus haplotypes  $p_{AG}$ ,  $p_{Ag}$ ,  $p_{aG}$ ,  $p_{ag}$  as on previous slide.

### More comments: |D'| and r<sup>2</sup> reflect different properties of LD

- |D'|, absolute value of D prime ...
  - ranges from 0 [no LD] to 1 [complete LD]
  - is less sensitive to marginal allele frequencies (D is sensitive but D' is much less sensitive)
  - is directly related to recombination fraction, e.g. D'=1 implies no evidence for historical recombination
- r<sup>2</sup>, squared correlation ...
  - also ranges from 0 [no LD] to 1
  - r<sup>2</sup>=1 if complete LD and allele frequencies the same at two loci [perfect LD]
  - is correlation between alleles on the same chromosome
  - is very sensitive to marginal allele frequencies
    - Far away low frequency variants can have higher r<sup>2</sup> than close common variants (Delvin 1995 Genomics, table 4)
  - is directly related to power for association

### D' and r<sup>2</sup> from haplotype counts instead of frequencies

	Α	а	
G	n <sub>11</sub>	n <sub>10</sub>	n <sub>1•</sub>
g	n <sub>01</sub>	n <sub>00</sub>	n <sub>0∙</sub>
	n <sub>•1</sub>	$n_{ullet 0}$	n

$$\delta = (1/n)^2 (n_{11} n_{00} - n_{10} n_{01})$$

Measure	Formula	
D'	$\frac{n_{11}n_{00} - n_{10}n_{01}}{\min(n_{\bullet 1}n_{1\bullet}, n_{\bullet 0}n_{0\bullet})} \text{ if } \delta < 0$ $\frac{n_{11}n_{00} - n_{10}n_{01}}{\min(n_{\bullet 1}n_{0\bullet}, n_{\bullet 0}n_{1\bullet})} \text{ if } \delta > 0$	
$\Delta^2 = r^2$	$\frac{\left(n_{11}n_{00} - n_{10}n_{01}\right)^{2}}{n_{\bullet 1}n_{\bullet 0}n_{1\bullet}n_{o\bullet}}$	

Again r is also Pearson correlation between alleles G and A.

#### Worked example

Same haplotype table as on previous slide

	A	а	
G	8	0	50%
g	2	6	50%
	62.5%	37.5%	

$$\delta = (1/16)^2 (8 \times 6 - 2 \times 0) = 48/256 = 0.1875 > 0.$$

$$D' = (8 \times 6 - 0) / \min(8 \times 6, 10 \times 8) = 1$$

$$r^2 = (8 \times 6 - 0)^2 / (10 \times 6 \times 8 \times 8) = 0.6$$

\*\* Tip: If any of the four cells is 0, then |D'|=1. Why? \*\*