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In **natural language processing**, the **latent Dirichlet allocation (LDA)** is a **generative statistical model** that allows sets of observations to be explained by **unobserved** groups that explain why some parts of the data are similar. For example, if observations are words collected into documents, it posits that each document is a mixture of a small number of topics and that each word's presence is attributable to one of the document's topics. LDA is an example of a **topic model** and belongs to the **machine learning** toolbox and in wider sense to the **artificial intelligence** toolbox.

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Overview [ [edit](#) ]

In evolutionary biology and bio-medicine, the model is used to detect the presence of structured genetic variation in a group of individuals. The model assumes that alleles carried by individuals under study have origin in various extant or past populations. The model and various inference algorithms allow scientist to estimate the allele frequencies in those source populations and the origin of alleles carried by individuals under study. The source populations can be interpreted ex-post in terms of various evolutionary scenarios. In [association studies](#), detecting the presence of genetic structure is considered a necessary preliminary step to avoid [confounding](#).

One example of LDA in engineering is to automatically classify documents and estimate their relevance to various topics.

In LDA, each document may be viewed as a *mixture* of various topics where each document is considered to have a set of topics that are assigned to it via LDA. This is identical to [probabilistic latent semantic analysis](#) (pLSA), except that in LDA the topic distribution is assumed to have a sparse [Dirichlet prior](#). The sparse Dirichlet priors encode the intuition that documents cover only a small set of topics and that topics use only a small set of words frequently. In practice, this results in a better disambiguation of words and a more precise assignment of documents to topics. LDA is a generalization of the [pLSA](#) model, which is equivalent to LDA under a uniform Dirichlet prior distribution.<sup>[4]</sup>

For example, an LDA model might have topics that can be classified as **CAT\_related** and **DOG\_related**. A topic has probabilities of generating various words, such as *milk*, *meow*, and *kitten*, which can be classified and interpreted by the viewer as "CAT\_related". Naturally, the word *cat* itself will have high probability given this topic. The **DOG\_related** topic likewise has probabilities of generating each word: *puppy*, *bark*, and *bone* might have high probability. Words without special relevance, such as "*the*" (see [function word](#)), will have roughly even probability between classes (or can be placed into a separate category). A topic is neither [semantically](#) nor [epistemologically](#) strongly defined. It is identified on the basis of automatic detection of the likelihood of term co-occurrence. A lexical word may occur in several topics with a different probability, however, with a different typical set of neighboring words in each topic.

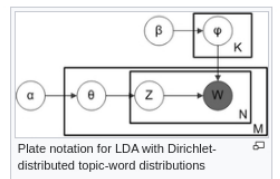
Each document is assumed to be characterized by a particular set of topics. This is similar to the standard [bag of words model](#) assumption, and makes the individual words [exchangeable](#).

With **plate notation**, which is often used to represent **probabilistic graphical models** (PGMs), the dependencies among the many variables can be captured concisely. The boxes are “plates” representing replicates, which are repeated entities. The outer plate represents documents, while the inner plate represents the repeated word positions in a given document; each position is associated with a choice of topic and word. The variable names are defined as follows:

$M$  denotes the number of documents  
 $N$  is number of words in a given document (document  $i$  has  $N_i$  words)  
 $\alpha$  is the parameter of the Dirichlet prior on the per-document topic distributions  
 $\beta$  is the parameter of the Dirichlet prior on the per-topic word distribution  
 $\theta_i$  is the topic distribution for document  $i$   
 $\varphi_k$  is the word distribution for topic  $k$   
 $z_{ij}$  is the topic for the  $j$ -th word in document  $i$   
 $w_{ij}$  is the specific word.

The fact that  $W$  is grayed out means that words  $w_{ij}$  are the only **observable variables**, and the other variables are **latent variables**. As proposed in the original paper<sup>[3]</sup>, a sparse Dirichlet prior can be used to model the topic-word distribution, following the intuition that the probability distribution over words in a topic is skewed, so that only a small set of words have high probability. The resulting model is the most widely applied variant of LDA today. The plate notation for this model is shown on the right, where  $K$  denotes the number of topics and  $\varphi_1, \dots, \varphi_K$  are  $V$ -dimensional vectors storing the parameters of the Dirichlet-distributed topic-word distributions ( $V$  is the number of words in the vocabulary).

It is helpful to think of the entities represented by  $\theta$  and  $\varphi$  as matrices created by decomposing the original document-word matrix that represents the corpus of documents being modeled. In this view,  $\theta$  consists of rows defined by documents and columns defined by topics, while  $\varphi$  consists of rows defined by topics and columns defined by words. Thus,  $\varphi_1, \dots, \varphi_K$  refers to a set of rows, or vectors, each of which is a distribution over words, and  $\theta_1, \dots, \theta_M$  refers to a set of rows, each of which is a distribution over topics.



**Generative process** [ edit ]

To actually infer the topics in a corpus, we imagine a generative process whereby the documents are created, so that we may infer, or reverse engineer, it. We imagine the generative process as follows. Documents are represented as random mixtures over latent topics, where each topic is characterized by a distribution over all the words. LDA assumes the following generative process for a corpus  $D$  consisting of  $M$  documents each of length  $N_i$ :

1. Choose  $\theta_i \sim \text{Dir}(\alpha)$ , where  $i \in \{1, \dots, M\}$  and  $\text{Dir}(\alpha)$  is a [Dirichlet distribution](#) with a symmetric parameter  $\alpha$  which typically is sparse ( $\alpha < 1$ )
2. Choose  $\varphi_k \sim \text{Dir}(\beta)$ , where  $k \in \{1, \dots, K\}$  and  $\beta$  typically is sparse
3. For each of the word positions  $i, j$ , where  $i \in \{1, \dots, M\}$ , and  $j \in \{1, \dots, N_i\}$ 
  - (a) Choose a topic  $z_{i,j} \sim \text{Multinomial}(\theta_i)$ .
  - (b) Choose a word  $w_{i,j} \sim \text{Multinomial}(\varphi_{z_{i,j}})$ .

(Note that *multinomial distribution* here refers to the [multinomial](#) with only one trial, which is also known as the [categorical distribution](#).)

The lengths  $N_i$  are treated as independent of all the other data generating variables ( $w$  and  $z$ ). The subscript is often dropped, as in the plate diagrams shown here.

**Definition** [ edit ]

A formal description of LDA is as follows:

Definition of variables in the model		
Variable	Type	Meaning
$K$	integer	number of topics (e.g. 50)
$V$	integer	number of words in the vocabulary (e.g. 50,000 or 1,000,000)
$M$	integer	number of documents
$N_{d=1 \dots M}$	integer	number of words in document $d$
$N$	integer	total number of words in all documents; sum of all $N_d$ values, i.e. $N = \sum_{d=1}^M N_d$
$\alpha_{k=1 \dots K}$	positive real	prior weight of topic $k$ in a document; usually the same for all topics; normally a number less than 1, e.g. 0.1, to prefer sparse topic distributions, i.e. few topics per document
$\alpha$	$K$ -dimensional vector of positive reals	collection of all $\alpha_k$ values, viewed as a single vector
$\beta_{w=1 \dots V}$	positive real	prior weight of word $w$ in a topic; usually the same for all words; normally a number much less than 1, e.g. 0.001, to strongly prefer sparse word distributions, i.e. few words per topic
$\beta$	$V$ -dimensional vector of positive reals	collection of all $\beta_w$ values, viewed as a single vector
$\varphi_{k=1 \dots K, w=1 \dots V}$	probability (real number between 0 and 1)	probability of word $w$ occurring in topic $k$
$\varphi_{k=1 \dots K}$	$V$ -dimensional vector of probabilities, which must sum to 1	distribution of words in topic $k$
$\theta_{d=1 \dots M, k=1 \dots K}$	probability (real number between 0 and 1)	probability of topic $k$ occurring in document $d$
$\theta_{d=1 \dots M}$	$K$ -dimensional vector of probabilities, which must sum to 1	distribution of topics in document $d$
$z_{d=1 \dots M, w=1 \dots N_d}$	integer between 1 and $K$	identity of topic of word $w$ in document $d$
<b>Z</b>	$N$ -dimensional vector of integers between 1 and $K$	identity of topic of all words in all documents
$w_{d=1 \dots M, w=1 \dots N_d}$	integer between 1 and $V$	identity of word $w$ in document $d$
<b>W</b>	$N$ -dimensional vector of integers between 1 and $V$	identity of all words in all documents

We can then mathematically describe the random variables as follows:

$$\begin{aligned}\varphi_{k=1 \dots K} &\sim \text{Dirichlet}_V(\beta) \\ \theta_{d=1 \dots M} &\sim \text{Dirichlet}_K(\alpha) \\ z_{d=1 \dots M, w=1 \dots N_d} &\sim \text{Categorical}_K(\theta_d) \\ w_{d=1 \dots M, w=1 \dots N_d} &\sim \text{Categorical}_V(\varphi_{z_{dw}})\end{aligned}$$

**Inference** [ edit ]

See also: [Dirichlet-multinomial distribution](#)

Learning the various distributions (the set of topics, their associated word probabilities, the topic of each word, and the particular topic mixture of each document) is a problem of [statistical inference](#).

**Monte Carlo simulation** [ edit ]

The original paper by Pritchard et al.<sup>[1]</sup> used approximation of the posterior distribution by Monte Carlo simulation. Alternative proposal of inference techniques include [Gibbs sampling](#).<sup>[5]</sup>

**Variational Bayes** [ edit ]

The original ML paper used a [variational Bayes](#) approximation of the [posterior distribution](#).<sup>[3]</sup>

**Likelihood maximization** [ edit ]

A direct optimization of the likelihood with a block relaxation algorithm proves to a fast alternative to MCMC.<sup>[6]</sup>

**Unknown number of populations/topics** [ edit ]

In practice, the most adequate number of populations or topics is not known beforehand. It can be estimated by estimation of the posterior distribution with [Reversible jump Markov chain Monte Carlo]<sup>[7]</sup>

**Alternative approaches** [ edit ]

Alternative approaches include [expectation propagation](#).<sup>[8]</sup>

Recent research has been focused on speeding up the inference of latent Dirichlet Allocation to support the capture of a massive number of topics in a large number of documents. The update equation of the collapsed Gibbs sampler mentioned in the earlier section has a natural sparsity within it that can be taken advantage of. Intuitively, since each document only contains a subset of topics  $K_d$ , and a word also only appears in a subset of topics  $K_w$ , the above update equation could be rewritten to take advantage of this sparsity.<sup>[9]</sup>

$$p(Z_{d,n} = k) \propto \frac{\alpha \beta}{C_k^{-n} + V \beta} + \frac{C_k^d \beta}{C_k^{-n} + V \beta} + \frac{C_k^w (\alpha + C_k^d)}{C_k^{-n} + V \beta}$$

In this equation, we have three terms, out of which two of them are sparse, and the other is small. We call these terms  $a$ ,  $b$  and  $c$  respectively. Now, if we normalize each term by summing over all the topics, we get:

$$\begin{aligned}A &= \sum_{k=1}^K \frac{\alpha \beta}{C_k^{-n} + V \beta} \\ B &= \sum_{k=1}^K \frac{C_k^d \beta}{C_k^{-n} + V \beta}\end{aligned}$$

$$C = \sum_{k=1}^K \frac{C_k^w (\alpha + C_k^d)}{C_k^w + V\beta}$$

Here, we can see that  $\mathbf{B}$  is a summation of the topics that appear in document  $\mathbf{d}$ , and  $\mathbf{C}$  is also a sparse summation of the topics that a word  $w$  is assigned to across the whole corpus.  $\mathbf{A}$  on the other hand, is dense but because of the small values of  $\alpha$  &  $\beta$ , the value is very small compared to the two other terms.

Now, while sampling a topic, if we sample a random variable uniformly from  $s \sim U(s \mid \mathbf{A} + \mathbf{B} + \mathbf{C})$ , we can check which bucket our sample lands in. Since  $\mathbf{A}$  is small, we are very unlikely to fall into this bucket; however, if we do fall into this bucket, sampling a topic takes  $O(K)$  time (same as the original Collapsed Gibbs Sampler). However, if we fall into the other two buckets, we only need to check a subset of topics if we keep a record of the sparse topics. A topic can be sampled from the  $\mathbf{B}$  bucket in  $O(K_d)$  time, and a topic can be sampled from the  $\mathbf{C}$  bucket in  $O(K_w)$  time where  $K_d$  and  $K_w$  denotes the number of topics assigned to the current document and current word type respectively.

Notice that after sampling each topic, updating these buckets is all basic  $O(1)$  arithmetic operations.

#### Aspects of computational details [\[ edit \]](#)

Following is the derivation of the equations for [collapsed Gibbs sampling](#), which means  $\varphi$ s and  $\theta$ s will be integrated out. For simplicity, in this derivation the documents are all assumed to have the same length  $N$ . The derivation is equally valid if the document lengths vary.

According to the model, the total probability of the model is:

$$P(\mathbf{W}, \mathbf{Z}, \boldsymbol{\theta}, \boldsymbol{\varphi}; \alpha, \beta) = \prod_{i=1}^K P(\varphi_i; \beta) \prod_{j=1}^M P(\theta_j; \alpha) \prod_{t=1}^N P(Z_{j,t} \mid \theta_j) P(W_{j,t} \mid \varphi_{Z_{j,t}}),$$

where the bold-font variables denote the vector version of the variables. First,  $\boldsymbol{\varphi}$  and  $\boldsymbol{\theta}$  need to be integrated out.

$$\begin{aligned} P(\mathbf{Z}, \mathbf{W}; \alpha, \beta) &= \int_{\boldsymbol{\theta}} \int_{\boldsymbol{\varphi}} P(\mathbf{W}, \mathbf{Z}, \boldsymbol{\theta}, \boldsymbol{\varphi}; \alpha, \beta) d\boldsymbol{\varphi} d\boldsymbol{\theta} \\ &= \int_{\boldsymbol{\theta}} \prod_{i=1}^K P(\varphi_i; \beta) \prod_{j=1}^M \prod_{t=1}^N P(W_{j,t} \mid \varphi_{Z_{j,t}}) d\boldsymbol{\varphi} \int_{\boldsymbol{\theta}} \prod_{j=1}^M P(\theta_j; \alpha) \prod_{t=1}^N P(Z_{j,t} \mid \theta_j) d\boldsymbol{\theta}. \end{aligned}$$

All the  $\theta$ s are independent to each other and the same to all the  $\varphi$ s. So we can treat each  $\theta$  and each  $\varphi$  separately. We now focus only on the  $\theta$  part.

$$\int_{\boldsymbol{\theta}} \prod_{j=1}^M P(\theta_j; \alpha) \prod_{t=1}^N P(Z_{j,t} \mid \theta_j) d\boldsymbol{\theta} = \prod_{j=1}^M \int_{\theta_j} P(\theta_j; \alpha) \prod_{t=1}^N P(Z_{j,t} \mid \theta_j) d\theta_j.$$

We can further focus on only one  $\theta$  as the following:

$$\int_{\theta_j} P(\theta_j; \alpha) \prod_{t=1}^N P(Z_{j,t} \mid \theta_j) d\theta_j.$$

Actually, it is the hidden part of the model for the  $j^{\text{th}}$  document. Now we replace the probabilities in the above equation by the true distribution expression to write out the explicit equation.

$$\int_{\theta_j} P(\theta_j; \alpha) \prod_{t=1}^N P(Z_{j,t} \mid \theta_j) d\theta_j = \int_{\theta_j} \frac{\Gamma(\sum_{i=1}^K \alpha_i)}{\prod_{i=1}^K \Gamma(\alpha_i)} \prod_{i=1}^K \theta_{j,i}^{\alpha_i-1} \prod_{t=1}^N P(Z_{j,t} \mid \theta_j) d\theta_j.$$

Let  $n_{j,r}^i$  be the number of word tokens in the  $j^{\text{th}}$  document with the same word symbol (the  $r^{\text{th}}$  word in the vocabulary) assigned to the  $i^{\text{th}}$  topic. So,  $n_{j,r}^i$  is three dimensional. If any of the three dimensions is not limited to a specific value, we use a parenthesized point  $(\cdot)$  to denote. For example,  $n_{j,(\cdot)}^i$  denotes the number of word tokens in the  $j^{\text{th}}$  document assigned to the  $i^{\text{th}}$  topic. Thus, the right most part of the above equation can be rewritten as:

$$\prod_{t=1}^N P(Z_{j,t} \mid \theta_j) = \prod_{i=1}^K \theta_{j,i}^{n_{j,i}^{(\cdot)}}.$$

So the  $\theta_j$  integration formula can be changed to:

$$\int_{\theta_j} \frac{\Gamma(\sum_{i=1}^K \alpha_i)}{\prod_{i=1}^K \Gamma(\alpha_i)} \prod_{i=1}^K \theta_{j,i}^{\alpha_i-1} \prod_{i=1}^K \theta_{j,i}^{n_{j,i}^{(\cdot)}} d\theta_j = \int_{\theta_j} \frac{\Gamma(\sum_{i=1}^K \alpha_i)}{\prod_{i=1}^K \Gamma(\alpha_i)} \prod_{i=1}^K \theta_{j,i}^{n_{j,i}^{(\cdot)} + \alpha_i - 1} d\theta_j.$$

Clearly, the equation inside the integration has the same form as the [Dirichlet distribution](#). According to the [Dirichlet distribution](#),

$$\int_{\theta_j} \frac{\Gamma(\sum_{i=1}^K n_{j,i}^{(\cdot)} + \alpha_i)}{\prod_{i=1}^K \Gamma(n_{j,i}^{(\cdot)} + \alpha_i)} \prod_{i=1}^K \theta_{j,i}^{n_{j,i}^{(\cdot)} + \alpha_i - 1} d\theta_j = 1.$$

Thus,

$$\begin{aligned} \int_{\theta_j} P(\theta_j; \alpha) \prod_{t=1}^N P(Z_{j,t} \mid \theta_j) d\theta_j &= \int_{\theta_j} \frac{\Gamma(\sum_{i=1}^K \alpha_i)}{\prod_{i=1}^K \Gamma(\alpha_i)} \prod_{i=1}^K \theta_{j,i}^{n_{j,i}^{(\cdot)} + \alpha_i - 1} d\theta_j \\ &= \frac{\Gamma(\sum_{i=1}^K \alpha_i)}{\prod_{i=1}^K \Gamma(\alpha_i)} \frac{\prod_{i=1}^K \Gamma(n_{j,i}^{(\cdot)} + \alpha_i)}{\Gamma(\sum_{i=1}^K n_{j,i}^{(\cdot)} + \alpha_i)} \int_{\theta_j} \frac{\Gamma(\sum_{i=1}^K n_{j,i}^{(\cdot)} + \alpha_i)}{\prod_{i=1}^K \Gamma(n_{j,i}^{(\cdot)} + \alpha_i)} \prod_{i=1}^K \theta_{j,i}^{n_{j,i}^{(\cdot)} + \alpha_i - 1} d\theta_j \\ &= \frac{\Gamma(\sum_{i=1}^K \alpha_i)}{\prod_{i=1}^K \Gamma(\alpha_i)} \frac{\prod_{i=1}^K \Gamma(n_{j,i}^{(\cdot)} + \alpha_i)}{\Gamma(\sum_{i=1}^K n_{j,i}^{(\cdot)} + \alpha_i)}. \end{aligned}$$

Now we turn our attention to the  $\boldsymbol{\varphi}$  part. Actually, the derivation of the  $\boldsymbol{\varphi}$  part is very similar to the  $\boldsymbol{\theta}$  part. Here we only list the steps of the derivation:

$$\begin{aligned} \int_{\boldsymbol{\varphi}} \prod_{i=1}^K P(\varphi_i; \beta) \prod_{j=1}^M \prod_{t=1}^N P(W_{j,t} \mid \varphi_{Z_{j,t}}) d\boldsymbol{\varphi} \\ &= \prod_{i=1}^K \int_{\varphi_i} P(\varphi_i; \beta) \prod_{j=1}^M \prod_{t=1}^N P(W_{j,t} \mid \varphi_{Z_{j,t}}) d\varphi_i \\ &= \prod_{i=1}^K \int_{\varphi_i} \frac{\Gamma(\sum_{r=1}^V \beta_r)}{\prod_{r=1}^V \Gamma(\beta_r)} \prod_{r=1}^V \varphi_{i,r}^{\beta_r-1} \prod_{r=1}^V \varphi_{i,r}^{n_{(\cdot),r}^i} d\varphi_i \\ &= \prod_{i=1}^K \int_{\varphi_i} \frac{\Gamma(\sum_{r=1}^V \beta_r)}{\prod_{r=1}^V \Gamma(\beta_r)} \prod_{r=1}^V \varphi_{i,r}^{n_{(\cdot),r}^i + \beta_r - 1} d\varphi_i \\ &= \prod_{i=1}^K \frac{\Gamma(\sum_{r=1}^V \beta_r)}{\prod_{r=1}^V \Gamma(\beta_r)} \frac{\prod_{r=1}^V \Gamma(n_{(\cdot),r}^i + \beta_r)}{\Gamma(\sum_{r=1}^V n_{(\cdot),r}^i + \beta_r)}. \end{aligned}$$

For clarity, here we write down the final equation with both  $\boldsymbol{\phi}$  and  $\boldsymbol{\theta}$  integrated out:

$$P(\mathbf{Z}, \mathbf{W}; \alpha, \beta) = \prod_{i=1}^K \frac{\Gamma(\sum_{i=1}^K \alpha_i)}{\prod_{i=1}^K \Gamma(\alpha_i)} \frac{\prod_{i=1}^K \Gamma(n_{j,i}^{(\cdot)} + \alpha_i)}{\Gamma(\sum_{i=1}^K n_{j,i}^{(\cdot)} + \alpha_i)} \times \prod_{i=1}^K \frac{\Gamma(\sum_{r=1}^V \beta_r)}{\prod_{r=1}^V \Gamma(\beta_r)} \frac{\prod_{r=1}^V \Gamma(n_{(\cdot),r}^i + \beta_r)}{\Gamma(\sum_{r=1}^V n_{(\cdot),r}^i + \beta_r)}.$$

$$j=1 \prod_{i=1}^n \Gamma(\alpha_i) \prod_{i=1}^n \Gamma(\sum_{j=1}^n n_{j,(\cdot)}^i + \alpha_i) \prod_{i=1}^n \Gamma(\sum_{r=1}^V \Gamma(\beta_r) \prod_{i=1}^n \Gamma(\sum_{r=1}^V n_{(\cdot),r}^i + \beta_r)$$

The goal of Gibbs Sampling here is to approximate the distribution of  $P(\mathbf{Z} \mid \mathbf{W}; \alpha, \beta)$ . Since  $P(\mathbf{W}; \alpha, \beta)$  is invariable for any of  $\mathbf{Z}$ , Gibbs Sampling equations can be derived from  $P(\mathbf{Z}, \mathbf{W}; \alpha, \beta)$  directly. The key point is to derive the following conditional probability:

$$P(Z_{(m,n)} \mid \mathbf{Z}_{-(m,n)}, \mathbf{W}; \alpha, \beta) = \frac{P(Z_{(m,n)}, \mathbf{Z}_{-(m,n)}, \mathbf{W}; \alpha, \beta)}{P(\mathbf{Z}_{-(m,n)}, \mathbf{W}; \alpha, \beta)},$$

where  $Z_{(m,n)}$  denotes the  $Z$  hidden variable of the  $n^{th}$  word token in the  $m^{th}$  document. And further we assume that the word symbol of it is the  $v^{th}$  word in the vocabulary.  $\mathbf{Z}_{-(m,n)}$  denotes all the  $Z$ s but  $Z_{(m,n)}$ . Note that Gibbs Sampling needs only to sample a value for  $Z_{(m,n)}$ , according to the above probability, we do not need the exact value of

$$P\left(Z_{m,n} \mid \mathbf{Z}_{-(m,n)}, \mathbf{W}; \alpha, \beta\right)$$

but the ratios among the probabilities that  $Z_{(m,n)}$  can take value. So, the above equation can be simplified as:

$$\begin{aligned} P(Z_{(m,n)} = v \mid \mathbf{Z}_{-(m,n)}, \mathbf{W}; \alpha, \beta) & \\ \propto P(Z_{(m,n)} = v, \mathbf{Z}_{-(m,n)}, \mathbf{W}; \alpha, \beta) & \\ = \left(\frac{\Gamma\left(\sum_{i=1}^K \alpha_i\right)}{\prod_{i=1}^K \Gamma(\alpha_i)}\right)^M \prod_{j \neq m} \frac{\prod_{i=1}^K \Gamma\left(n_{j,(\cdot)}^i + \alpha_i\right)}{\Gamma\left(\sum_{i=1}^K n_{j,(\cdot)}^i + \alpha_i\right)} \left(\frac{\Gamma\left(\sum_{r=1}^V \beta_r\right)}{\prod_{r=1}^V \Gamma(\beta_r)}\right)^K \prod_{i=1}^K \prod_{r \neq v} \Gamma\left(n_{(\cdot),r}^i + \beta_r\right) \frac{\prod_{i=1}^K \Gamma\left(n_{m,(\cdot)}^i + \alpha_i\right)}{\Gamma\left(\sum_{i=1}^K n_{m,(\cdot)}^i + \alpha_i\right)} \prod_{i=1}^K \frac{\Gamma\left(n_{(\cdot),v}^i + \beta_v\right)}{\Gamma\left(\sum_{r=1}^V n_{(\cdot),r}^i + \beta_r\right)} & \\ \propto \frac{\prod_{i=1}^K \Gamma\left(n_{m,(\cdot)}^i + \alpha_i\right)}{\Gamma\left(\sum_{i=1}^K n_{m,(\cdot)}^i + \alpha_i\right)} \prod_{i=1}^K \frac{\Gamma\left(n_{(\cdot),v}^i + \beta_v\right)}{\Gamma\left(\sum_{r=1}^V n_{(\cdot),r}^i + \beta_r\right)} & \\ \propto \prod_{i=1}^K \Gamma\left(n_{m,(\cdot)}^i + \alpha_i\right) \prod_{i=1}^K \frac{\Gamma\left(n_{(\cdot),v}^i + \beta_v\right)}{\Gamma\left(\sum_{r=1}^V n_{(\cdot),r}^i + \beta_r\right)}. & \end{aligned}$$

Finally, let  $n_{j,r}^{i-(m,n)}$  be the same meaning as  $n_{j,r}^i$  but with the  $Z_{(m,n)}$  excluded. The above equation can be further simplified leveraging the property of [gamma function](#). We first split the summation and then merge it back to obtain a  $k$ -independent summation, which could be dropped:

$$\begin{aligned} & \propto \prod_{i \neq k} \Gamma\left(n_{m,(\cdot)}^{i-(m,n)} + \alpha_i\right) \prod_{i \neq k} \frac{\Gamma\left(n_{(\cdot),v}^{i-(m,n)} + \beta_v\right)}{\Gamma\left(\sum_{r=1}^V n_{(\cdot),r}^{i-(m,n)} + \beta_r\right)} \Gamma\left(n_{m,(\cdot)}^{k-(m,n)} + \alpha_k + 1\right) \frac{\Gamma\left(n_{(\cdot),v}^{k-(m,n)} + \beta_v + 1\right)}{\Gamma\left(\sum_{r=1}^V n_{(\cdot),r}^{k-(m,n)} + \beta_r + 1\right)} \\ & = \prod_{i \neq k} \Gamma\left(n_{m,(\cdot)}^{i-(m,n)} + \alpha_i\right) \prod_{i \neq k} \frac{\Gamma\left(n_{(\cdot),v}^{i-(m,n)} + \beta_v\right)}{\Gamma\left(\sum_{r=1}^V n_{(\cdot),r}^{i-(m,n)} + \beta_r\right)} \Gamma\left(n_{m,(\cdot)}^{k-(m,n)} + \alpha_k\right) \frac{\Gamma\left(n_{(\cdot),v}^{k-(m,n)} + \beta_v\right)}{\Gamma\left(\sum_{r=1}^V n_{(\cdot),r}^{k-(m,n)} + \beta_r\right)} \left(n_{m,(\cdot)}^{k-(m,n)} + \alpha_k + 1\right) \frac{n_{(\cdot),v}^{k-(m,n)} + \beta_v + 1}{\sum_{r=1}^V n_{(\cdot),r}^{k-(m,n)} + \beta_r + 1} \\ & = \prod_i \Gamma\left(n_{m,(\cdot)}^{i-(m,n)} + \alpha_i\right) \prod_i \frac{\Gamma\left(n_{(\cdot),v}^{i-(m,n)} + \beta_v\right)}{\Gamma\left(\sum_{r=1}^V n_{(\cdot),r}^{i-(m,n)} + \beta_r\right)} \left(n_{m,(\cdot)}^{k-(m,n)} + \alpha_k + 1\right) \frac{n_{(\cdot),v}^{k-(m,n)} + \beta_v + 1}{\sum_{r=1}^V n_{(\cdot),r}^{k-(m,n)} + \beta_r + 1} \\ & \propto \left(n_{m,(\cdot)}^{k-(m,n)} + \alpha_k + 1\right) \frac{n_{(\cdot),v}^{k-(m,n)} + \beta_v + 1}{\sum_{r=1}^V n_{(\cdot),r}^{k-(m,n)} + \beta_r + 1} \end{aligned}$$

Note that the same formula is derived in the article on the [Dirichlet-multinomial distribution](#), as part of a more general discussion of integrating [Dirichlet distribution](#) priors out of a [Bayesian network](#).

#### Related problems [ edit ]

#### Related models [ edit ]

Topic modeling is a classic solution to the problem of [information retrieval](#) using linked data and semantic web technology<sup>[**10**]</sup>. Related models and techniques are, among others, [latent semantic indexing](#), [independent component analysis](#), [probabilistic latent semantic indexing](#), [non-negative matrix factorization](#), and [Gamma-Poisson distribution](#).

The LDA model is highly modular and can therefore be easily extended. The main field of interest is modeling relations between topics. This is achieved by using another distribution on the simplex instead of the Dirichlet. The Correlated Topic Model<sup>[**11**]</sup> follows this approach, inducing a correlation structure between topics by using the [logistic normal distribution](#) instead of the Dirichlet. Another extension is the hierarchical LDA (hLDA),<sup>[**12**]</sup> where topics are joined together in a hierarchy by using the nested [Chinese restaurant process](#), whose structure is learnt from data. LDA can also be extended to a corpus in which a document includes two types of information (e.g., words and names), as in the [LDA-dual model](#).<sup>[**13**]</sup> Nonparametric extensions of LDA include the [hierarchical Dirichlet process](#) mixture model, which allows the number of topics to be unbounded and learnt from data.

As noted earlier, pLSA is similar to LDA. The LDA model is essentially the Bayesian version of pLSA model. The Bayesian formulation tends to perform better on small datasets because Bayesian methods can avoid overfitting the data. For very large datasets, the results of the two models tend to converge. One difference is that pLSA uses a variable *d* to represent a document in the training set. So in pLSA, when presented with a document the model hasn't seen before, we fix  $\Pr(\boldsymbol{w} \mid \boldsymbol{z})$ —the probability of words under topics—to be that learned from the training set and use the same EM algorithm to infer  $\Pr(\boldsymbol{z} \mid \boldsymbol{d})$ —the topic distribution under *d*. Blei argues that this step is cheating because you are essentially refitting the model to the new data.

#### Spatial models [ edit ]

In evolutionary biology, it is often natural to assume that the geographic locations of the individuals observed bring some information about their ancestry. This is the rational of various models for geo-referenced genetic data<sup>[**7**]</sup><sup>[**14**]</sup>

Variations on LDA have been used to automatically put natural images into categories, such as "bedroom" or "forest", by treating an image as a document, and small patches of the image as words;<sup>[**15**]</sup> one of the variations is called [Spatial Latent Dirichlet Allocation](#).<sup>[**16**]</sup>

#### See also [ edit ]

- [Variational Bayesian methods](#)
- [Pachinko allocation](#)
- [tf-idf](#)

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## External links



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- jLDADMM** A Java package for topic modeling on normal or short texts. jLDADMM includes implementations of the LDA topic model and the *one-topic-per-document* Dirichlet Multinomial Mixture model. jLDADMM also provides an implementation for document clustering evaluation to compare topic models.
- STTM** A Java package for short text topic modeling (<https://github.com/qiang2100/STTM>). STTM includes these following algorithms: Dirichlet Multinomial Mixture (DMM) in conference KDD2014, Bitern Topic Model (BTM) in journal TKDE2016, Word Network Topic Model (WNTM ) in journal KAIS2018, Pseudo-Document-Based Topic Model (PTM) in conference KDD2016, Self-Aggregation-Based Topic Model (SATM) in conference IJCAI2015, (ETM) in conference PAKDD2017, Generalized P'olya Urn (GPU) based Dirichlet Multinomial Mixturemodel (GPU-DMM) in conference SIGIR2016, Generalized P'olya Urn (GPU) based Poisson-based Dirichlet Multinomial Mixturemodel (GPU-PDMM) in journal TIS2017 and Latent Feature Model with DMM (LF-DMM) in journal TACL2015. STTM also includes six short text corpus for evaluation. STTM presents three aspects about how to evaluate the performance of the algorithms (i.e., topic coherence, clustering, and classification).
- Lecture that covers some of the notation in this article: [LDA and Topic Modelling Video Lecture by David Blei](#) or [same lecture on YouTube](#)
- [D. Mimno's LDA Bibliography](#) An exhaustive list of LDA-related resources (incl. papers and some implementations)
- [Gensim](#), a Python+NumPy implementation of online LDA for inputs larger than the available RAM.
- [topicmodels](#) and [lda](#) are two R packages for LDA analysis.
- "Text Mining with R" including [LDA methods](#), video presentation to the October 2011 meeting of the Los Angeles R users group
- [MALLET](#) Open source Java-based package from the University of Massachusetts-Amherst for topic modeling with LDA, also has an independently developed GUI, the [Topic Modeling Tool](#)
- [LDA in Mahout](#) implementation of LDA using [MapReduce](#) on the [Hadoop](#) platform
- [Latent Dirichlet Allocation \(LDA\) Tutorial for the Infer.NET Machine Computing Framework](#) Microsoft Research C# Machine Learning Framework
- [LDA in Spark](#): Since version 1.3.0, [Apache Spark](#) also features an implementation of LDA
- [LDA](#), [exampleLDA](#) MATLAB implementation

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Natural language processing

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