calzone

Release develop

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CONTENTS:

1	1.1 Installation 1.2 Command line interface	3 4 2
2	Summary and guide for calzone 2.1 Guide to calzone and calibration metrics	
3	Reliability diagram13.1 References2	1 7
4		23 24 24 24 26
5	5.1Theoretical Background	30 30 31
6	6.1Theoretical Background36.2Pros of Cox calibration analysis36.3Cons of Cox calibration analysis36.4Calculating Cox slope and intercept with calzone3	35 36 37
7		1 1 2 2

	7.6 Reference	44
9	Spiegelhalter's Z-test 8.1 Theoretical background 8.2 Pros of Spiegelhalter's Z test 8.3 Cons of Spiegelhalter's Z test 8.4 Calculating the Spiegelhalter Z score and p-value using calzone 8.5 P-value of Spiegelhalter's z test 8.6 Reference Prevalence adjustment	45 45 46 46 46 48
	Preform prevalence adjustment in calzone	50 52 52
10	Subgroup analysis	53
11	Multiclass extension	57
12	Running the GUI	61
13	Validating metrics with external packages 13.1 Reliability diagram	63 64 65 66
14	calzone 14.1 calzone package	67 67
Py	hon Module Index	81
Inc	ex	83

calzone is a Python package for calculation of various calibration metrics. This work is credited to Kwok Lung (Jason) Fan and Qian Cao.

The calzone package provides a suite of tools for assessing and improving the calibration of machine learning models, particularly for binary classification tasks. It offers calibration metrics and visualization tools for displaying reliability diagrams. Please read the summary and guide section in this documentation first before using the package.

Key features of calzone include:

- Calculation of calibration metrics (ECE, MCE, Hosmer-Lemeshow test, spiegelhalter z test, etc.)
- · Visualization functions for reliability diagrams
- Bootstrapping capabilities for confidence interval estimation
- Subgroup analysis for calibration metrics
- Command line interface scripts for batch processing
- Multi-class extension by using 1-vs-rest or top-class only

To accurately assess the calibration of machine learning models, it is essential to have a comprehensive and reprensative dataset with sufficient coverage of the prediction space. The calibration metrics is not meaningful if the dataset is not representative of true intended population.

We hope you find calzone useful in your machine learning projects!

CONTENTS: 1

2 CONTENTS:

CHAPTER

ONE

QUICK START

This tutorial provides a quick start to the calzone package, including a discussion of the installation and the basic command line interface useage.

1.1 Installation

Calzone dependencies are numpy, scipy, statsmodels and matplotlib. If you are experienced developer, you probably already have numpy, scipy and matplolib installed. If you don't have a package install, you can install them with conda

```
[]: conda install numpy
conda install scipy
conda install matplotlib
conda install statsmodels
```

Alternatively, you can install the dependencies with pip

```
[]: pip install numpy
pip install scipy
pip install matplotlib
pip install statsmodels
```

Then, you can install the calzone package from github if you only want to use the calculator inside you python script:

```
[]: pip install -e "git+https://github.com/DIDSR/calzone.git"
```

If you want to run the command line interface or GUI interface, you need to clone the github repository. Notice that the GUI interface require install nicegui using pip install nicegui.

```
[]: git clone https://github.com/DIDSR/calzone.git
cd calzone
# install calzone
pip install -e .
```

1.2 Command line interface

First, you need to prepare your dataset in a specific format. The dataset should be a CSV file with the following columns:

```
proba_0, proba_1, ..., proba_n, label
```

where n >= 1. The proba_i columns are the probabilities of the i-th class, and the label column is the true class label.

Or if you have subgroups, the dataset should be a CSV file with the following columns:

```
proba_0, proba_1, ..., proba_n, subgroup_1, subgroup_2, ..., subgroup_m, label
```

where $n \ge 1$ and $m \ge 1$. The subgroup_j columns are the values of the j-th subgroup type, and its values should be categorical.

In the case of multi-class, you need to specify the class-of-interest, and the problem will be treated as 1-vs-all binary classification. To test the full calibration of the whole model, you need to test the calibration of each class.

The program also works if your csv file has no header. It will assume the first [:-1] columns are the probabilities and the last column is the label.

We provides examples dataset in csv format in the e examples folder and it illustrate the input format of dataset.

```
[6]: ### For illuration purpose , I will use the functions in the helper.py in the local_directory
### The data is generated using beta-binomial distribution

from helper import * #import local helper functions for example data generation

generate_wellcal_data(5000,"../../example_data/simulated_welldata.csv",alpha_val=0.5,_
___beta_val=0.5,random_seed=123)

generate_miscal_data(5000,"../../example_data/simulated_misdata.csv", miscal_scale=2,
___alpha_val=0.5, beta_val=0.5,random_seed=123)

generate_subgroup_data(5000,"../../example_data/simulated_data_subgroup.csv",miscal_
__scale=2,alpha_val=0.5, beta_val=0.5,random_seed=123)
```

```
[3]: ### Snippet of the csv file format
    #Without subgroups
    import numpy as np
    print("Without subgroups")
    print(np.loadtxt('../../example_data/simulated_welldata.csv',dtype=str)[:5]) #first 5_
     → lines of the without subgroups csv files
    #With subgroups
    print("With subgroups")
    print(np.loadtxt('../../example_data/simulated_data_subgroup.csv',dtype=str)[:5])
     →#first 5 lines of the with subgroups csv files
    Without subgroups
    ['proba_0,proba_1,label'
     '1.444156178040510996e-01,8.555843821959489004e-01,0.00000000000000000e+00'
     '8.552048445812980848e-01,1.447951554187018874e-01,0.000000000000000000e+00'
     '2.569696048872897043e-01,7.430303951127102957e-01,0.000000000000000000e+00'
     '3.993130565553012490e-01,6.006869434446987510e-01,1.0000000000000000000e+00']
    With subgroups
    ['proba_0,proba_1,subgroup_1,label'
      '0.1444156178040511,0.8555843821959489,A,0'
```

(continues on next page)

```
'0.8552048445812981,0.1447951554187019,A,O'
'0.2569696048872897,0.7430303951127103,A,O'
'0.39931305655530125,0.6006869434446988,A,1']
```

To use the command line interface (CLI), you can use the script in the calzone directory. The program will save the metrics into the output csv file with the CI (if you turn on boostrap). The program will also save the relibaility diagram if you apply –plot flag. There is an optional flag –prevalence_adjustment which tries to derive the original model prevalence and apply prevalence adjustment. See more on prevalence adjustment in the prevalecne adjustment notebook.

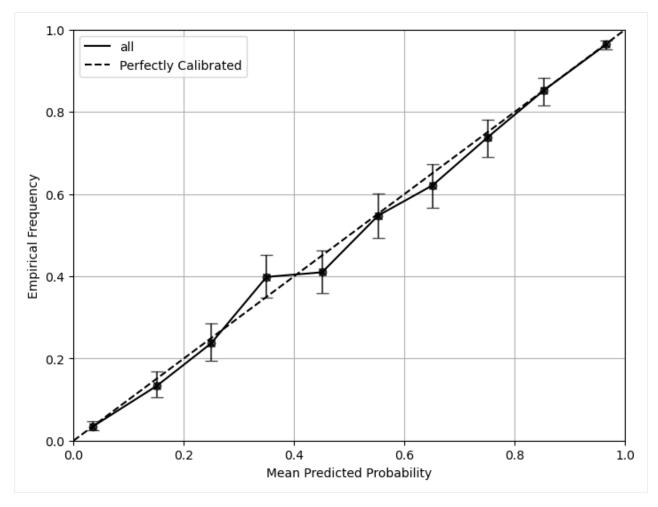
```
[1]: # Use -h to see the help message and options
    %run ../../cal_metrics.py -h
    usage: cal_metrics.py [-h] [--csv_file CSV_FILE] [--metrics METRICS]
                           [--prevalence_adjustment] [--n_bootstrap N_BOOTSTRAP]
                           [--bootstrap_ci BOOTSTRAP_CI]
                           [--class_to_calculate CLASS_TO_CALCULATE]
                           [--num_bins NUM_BINS]
                           [--hl_test_validation HL_TEST_VALIDATION] [--topclass]
                           [--save_metrics SAVE_METRICS] [--plot]
                           [--plot_bins PLOT_BINS] [--save_plot SAVE_PLOT]
                           [--save_diagram_output SAVE_DIAGRAM_OUTPUT] [--verbose]
    Calculate calibration metrics and visualize reliability diagram.
    options:
      -h, --help
                             show this help message and exit
      --csv_file CSV_FILE
                             Path to the input CSV file. (If there is header, it
                             must be in: proba_0,proba_1,...,subgroup_1(optional),s
                             ubgroup_2(optional),...label. If no header, then the
                             columns must be in the order of
                             proba_0,proba_1,...,label)
                             Comma-separated list of specific metrics to calculate
      --metrics METRICS
                             (SpiegelhalterZ, ECE-H, MCE-H, HL-H, ECE-C, MCE-C, HL-
                             C,COX,Loess,all). Default: all
      --prevalence_adjustment
                             Perform prevalence adjustment (default: False)
      --n_bootstrap N_BOOTSTRAP
                             Number of bootstrap samples (default: 0)
      --bootstrap_ci BOOTSTRAP_CI
                             Bootstrap confidence interval (default: 0.95)
      --class_to_calculate CLASS_TO_CALCULATE
                             Class to calculate metrics for (default: 1)
      --num_bins NUM_BINS
                             Number of bins for ECE/MCE/HL calculations (default:
                             10)
      --hl_test_validation HL_TEST_VALIDATION
                             Using nbin instead of nbin-2 as HL test DOF. Use it if
                             the dataset is validation set.
      --topclass
                             Whether to transform the problem to top-class problem.
      --save_metrics SAVE_METRICS
                             Save the metrics to a csv file
      --plot
                             Plot reliability diagram (default: False)
       --plot_bins PLOT_BINS
```

(continues on next page)

```
--save_plot SAVE_PLOT
                             Save the plot to a file. Must end with valid image
                             formats.
      --save_diagram_output SAVE_DIAGRAM_OUTPUT
                             Save the reliability diagram output to a file
      --verbose
                            Print verbose output
[8]: %run ../../cal_metrics.py \
    --csv_file '../../example_data/simulated_welldata.csv' \
    --metrics all \
    --n_bootstrap 1000 \
    --bootstrap_ci 0.95 \
    --class_to_calculate 1 \
    --num_bins 10 \
    --save_metrics '../../example_data/simulated_welldata_result.csv' \
    --plot \
    --plot_bins 10 \
    --save_plot '../../example_data/simulated_welldata_result.png' \
    --verbose \
    --save_diagram_output '../../example_data/simulated_welldata_diagram_output.csv'
    ### save_diagram_output only when you want to save the reliability diagram output
    #--prevalence_adjustment # only when you want to apply prevalence adjustment
    #--hl_test_validation #use it only when the data is from validation set
    Metrics with bootstrap confidence intervals:
    SpiegelhalterZ score: 0.376 (-1.581, 2.396)
    SpiegelhalterZ p-value: 0.707 (0.017, 0.971)
    ECE-H topclass: 0.01 (0.006, 0.021)
    ECE-H: 0.012 (0.011, 0.025)
    MCE-H topclass: 0.039 (0.016, 0.076)
    MCE-H: 0.048 (0.034, 0.107)
    HL-H score: 8.885 (7.392, 34.058)
    HL-H p-value: 0.352 (0.000, 0.495)
    ECE-C topclass: 0.009 (0.007, 0.021)
    ECE-C: 0.009 (0.008, 0.022)
    MCE-C topclass: 0.021 (0.018, 0.072)
    MCE-C: 0.023 (0.020, 0.071)
    HL-C score: 3.695 (4.778, 28.666)
    HL-C p-value: 0.884 (0.000, 0.781)
    COX coef: 0.994 (0.940, 1.054)
    COX intercept: -0.045 (-0.126, 0.031)
    COX coef lowerci: 0.937 (0.886, 0.994)
    COX coef upperci: 1.051 (0.994, 1.114)
    COX intercept lowerci: -0.123 (-0.205, -0.049)
    COX intercept upperci: 0.034 (-0.048, 0.109)
```

Number of bins for reliability diagram

COX ICI: 0.006 (0.001, 0.016) Loess ICI: 0.006 (0.004, 0.016)

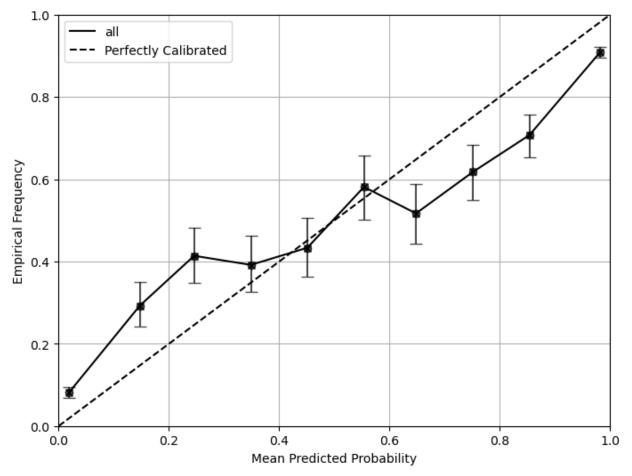


We use the CLI to compute the metrics and reliability diagram for the example dataset. The metrics and its 95% confidence interval will be printed and saved in a file. The reliability diagram will be saved in a separate file.

We can also test it on a miscalibrated dataset. The miscalibration is introduced by mutliply the log of probabilities by 2 and convert it back to probabilities.

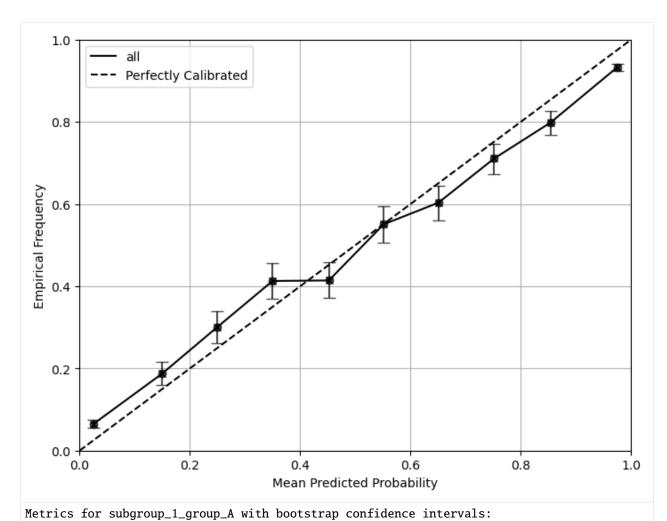
```
[9]: %run ../../cal_metrics.py \
    --csv_file '../../example_data/simulated_misdata.csv' \
    --metrics all \
    --n_bootstrap 1000 \
    --bootstrap_ci 0.95 \
    --class_to_calculate 1 \
    --num_bins 10 \
    --save_metrics '../../example_data/simulated_misdata_result.csv' \
    --plot \
    --plot_bins 10 \
    --save_plot '../../example_data/simulated_misdata_result.png' \
    --verbose
    Metrics with bootstrap confidence intervals:
    SpiegelhalterZ score: 29.626 (26.37, 33.185)
    SpiegelhalterZ p-value: 0. (0., 0.)
    ECE-H topclass: 0.081 (0.073, 0.091)
                                                                                (continues on next page)
```

```
ECE-H: 0.081 (0.073, 0.092)
MCE-H topclass: 0.151 (0.127, 0.202)
MCE-H: 0.168 (0.145, 0.244)
HL-H score: 1027.940 (818.524, 1292.977)
HL-H p-value: 0. (0., 0.)
ECE-C topclass: 0.079 (0.069, 0.09)
ECE-C: 0.08 (0.07, 0.091)
MCE-C topclass: 0.168 (0.141, 0.204)
MCE-C: 0.158 (0.139, 0.203)
HL-C score: 1857.584 (1355.637, 3341.319)
HL-C p-value: 0. (0., 0.)
COX coef: 0.497 (0.470, 0.524)
COX intercept: -0.045 (-0.125, 0.038)
COX coef lowerci: 0.469 (0.443, 0.494)
COX coef upperci: 0.526 (0.497, 0.554)
COX intercept lowerci: -0.123 (-0.203, -0.041)
COX intercept upperci: 0.034 (-0.046, 0.116)
COX ICI: 0.078 (0.071, 0.085)
Loess ICI: 0.074 (0.065, 0.083)
```

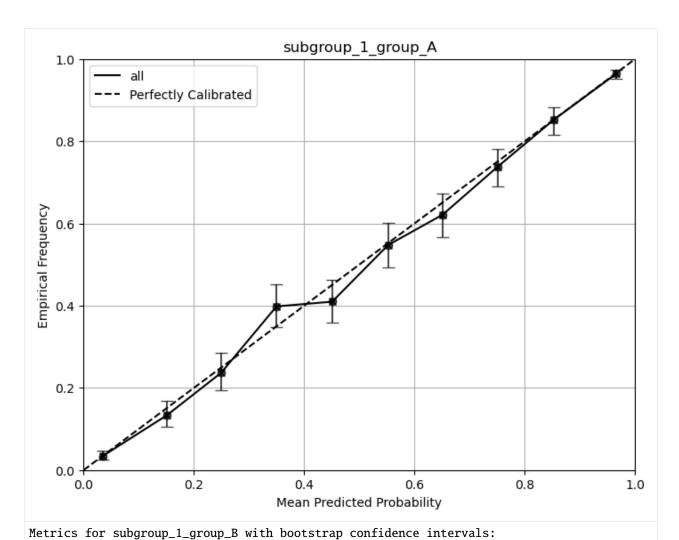


If your data has subgroups in it, simply run the script with the same argument as the one above. It will automatically detect the subgroup and generate the corresponding plots and metrics for each subgroup as well as the overall plot and metrics.

```
[10]: %run ../../cal_metrics.py \
     --csv_file '../../example_data/simulated_data_subgroup.csv' \
     --metrics all \
      --n_bootstrap 1000 \
     --bootstrap_ci 0.95 \
     --class_to_calculate 1 \
     --num_bins 10 \
      --save_metrics '../../example_data/simulated_data_subgroup_result.csv' \
     --plot \
     --plot_bins 10 \
      --save_plot '../../example_data/simulated_data_subgroup_result.png' \
      --verbose
     Metrics with bootstrap confidence intervals:
     SpiegelhalterZ score: 18.327 (15.794, 21.009)
     SpiegelhalterZ p-value: 0. (0., 0.)
     ECE-H topclass: 0.042 (0.035, 0.049)
     ECE-H: 0.042 (0.036, 0.049)
     MCE-H topclass: 0.055 (0.043, 0.087)
     MCE-H: 0.063 (0.055, 0.109)
     HL-H score: 429.732 (335.116, 584.729)
     HL-H p-value: 0. (0., 0.)
     ECE-C topclass: 0.042 (0.035, 0.049)
     ECE-C: 0.038 (0.032, 0.046)
     MCE-C topclass: 0.065 (0.055, 0.091)
     MCE-C: 0.064 (0.052, 0.086)
     HL-C score: 1138.842 (779.577, 1844.15)
     HL-C p-value: 0. (0., 0.)
     COX coef: 0.668 (0.638, 0.698)
     COX intercept: -0.02 (-0.073, 0.03)
     COX coef lowerci: 0.641 (0.611, 0.67)
     COX coef upperci: 0.696 (0.664, 0.727)
     COX intercept lowerci: -0.074 (-0.127, -0.025)
     COX intercept upperci: 0.034 (-0.019, 0.084)
     COX ICI: 0.049 (0.044, 0.055)
     Loess ICI: 0.037 (0.032, 0.043)
```



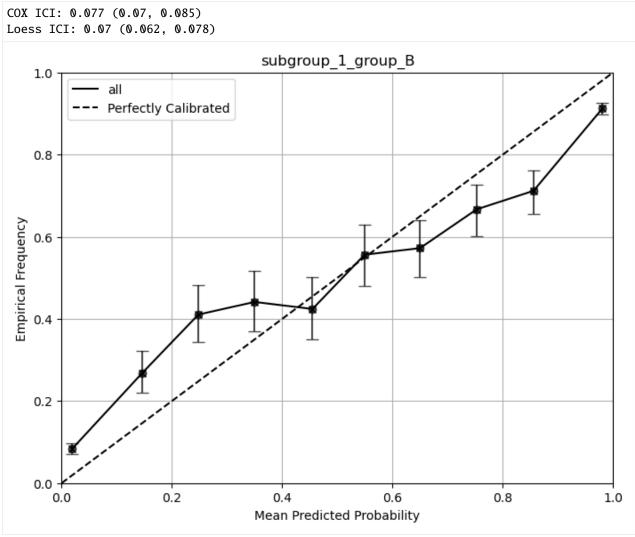
```
SpiegelhalterZ score: 0.376 (-1.536, 2.249)
SpiegelhalterZ p-value: 0.707 (0.018, 0.981)
ECE-H topclass: 0.01 (0.006, 0.021)
ECE-H: 0.012 (0.011, 0.025)
MCE-H topclass: 0.039 (0.017, 0.077)
MCE-H: 0.048 (0.034, 0.107)
HL-H score: 8.885 (7.273, 34.6)
HL-H p-value: 0.352 (0.000, 0.507)
ECE-C topclass: 0.009 (0.007, 0.022)
ECE-C: 0.009 (0.007, 0.023)
MCE-C topclass: 0.021 (0.018, 0.072)
MCE-C: 0.023 (0.018, 0.075)
HL-C score: 3.695 (4.463, 29.686)
HL-C p-value: 0.884 (0.000, 0.813)
COX coef: 0.994 (0.942, 1.050)
COX intercept: -0.045 (-0.135, 0.028)
COX coef lowerci: 0.937 (0.888, 0.990)
COX coef upperci: 1.051 (0.996, 1.111)
COX intercept lowerci: -0.123 (-0.214, -0.051)
COX intercept upperci: 0.034 (-0.056, 0.107)
COX ICI: 0.006 (0.001, 0.017)
Loess ICI: 0.006 (0.003, 0.016)
```



```
SpiegelhalterZ score: 27.936 (24.631, 31.175)
SpiegelhalterZ p-value: 0. (0., 0.)
ECE-H topclass: 0.077 (0.068, 0.086)
ECE-H: 0.077 (0.069, 0.087)
MCE-H topclass: 0.133 (0.108, 0.175)
MCE-H: 0.163 (0.13, 0.232)
HL-H score: 910.439 (716.681, 1156.971)
HL-H p-value: 0. (0., 0.)
ECE-C topclass: 0.074 (0.066, 0.084)
ECE-C: 0.075 (0.065, 0.085)
MCE-C topclass: 0.141 (0.124, 0.182)
MCE-C: 0.140 (0.115, 0.182)
HL-C score: 2246.171 (1393.175, 3900.391)
HL-C p-value: 0. (0., 0.)
COX coef: 0.507 (0.481, 0.539)
COX intercept: 0.000 (-0.073, 0.077)
COX coef lowerci: 0.478 (0.454, 0.508)
COX coef upperci: 0.536 (0.508, 0.569)
COX intercept lowerci: -0.078 (-0.152, -0.002)
COX intercept upperci: 0.079 (0.005, 0.155)
```

(continues on next page)





1.3 Using calzone in python

Instead of running the command line tool, you can also use calzone in python directly

Chapter 1. Quick Start

```
'HL-H score': 8.884991559088098,
'HL-H p-value': 0.35209071874348785,
'ECE-C topclass': 0.009458033653818828,
'ECE-C': 0.008733966945443138,
'MCE-C topclass': 0.020515047600205505,
'MCE-C': 0.02324031223486256,
'HL-C score': 3.694947603203135,
'HL-C p-value': 0.8835446575708198,
'COX coef': 0.9942499557748269,
'COX intercept': -0.04497652296600376,
'COX coef lowerci': 0.9372902801721911,
'COX coef upperci': 1.0512096313774626,
'COX intercept lowerci': -0.12348577118577644,
'COX intercept upperci': 0.03353272525376893,
'COX ICI': 0.005610391483826338,
'Loess ICI': 0.00558856942568957}
```

SUMMARY AND GUIDE FOR CALZONE

We provide a summary of the calibration metrics provides by calzone, including the pros and cons of each metrics. For a more detailed explanation of each metrics and how to calculate them using calzone, please refer to the specific notebook.

Meaning	Cons	Pros	Description	Metrics
Average deviation from true probability	Depend on binning Depend on class-by-class/top-class	• Intuitive • Easy to calculate	Using binned reliability diagram (equal-width or equal-count binning), sum of absolute difference, weighted by bin count.	Expected calibration error (ECE)
Maximum deviation from true probability	Depend on binning Depend on class-by-class/top-class	• Intuitive • Easy to calculate	Using binned reliability diagram (equal-width or equal-count binning), Maximum absolute difference.	Maximum calibration error (MCE)
Test of calibration	 Depend on binning Low power Wrong coverage 	Intuitive Statistical meaning	Using binned reliability diagram (equal-width or equal-count binning), Chi-squared based test using expected and observed.	Hosmer-Lemeshow test
Test of calibration	Doesn't detect prevalence shift	 Doesn't rely on binning Statistical meaning 	Decomposition of brier score. Normal distributed	Spiegelhalter's z test
A logit fit to the calibration curve	• Failed to capture some miscalibration	• Doesn't rely on binning • Hints at miscalibration type	Logistic regression of the logits	Cox's analysis
	Depend on the choice of curve fitting Depend on fitting parameters	Doesn't rely on binning Capture all kind of miscalibration	Similar to ECE, using smooth fit (usually losse) instead of binning to get the calibration curve	Integrated calibration index (ICI)

2.1 Guide to calzone and calibration metrics

calzone aims to access whether a model achieves moderate calibration, meaning whether $\mathbb{P}(\hat{Y}=Y|\hat{P}=p)=p$ for all $p\in[0,1]$.

To accurately assess the calibration of machine learning models, it is essential to have a comprehensive and reprensative dataset with sufficient coverage of the prediction space. The calibration metrics is not meaningful if the dataset is not representative of true intended population.

calzone takes in a csv dataset which contains the probability of each class and the true label. Most metrics in calzone only work with binary classification and which transforms the problem into 1-vs-rest when calcualte the metrics. Therefore, you need to specify the class-of-interest when using the metrics. The only exception is the Top-class Expected calibration error (ECE_{top}) and Top-class Maximum calibration error (MCE_{top}) metrics which only measure the calibration of the class with highest predicted probability hence works for multi-class problems. See the corresponding documentation for more details.

We recommend visualizing calibration using reliability diagrams. If you observe general over- or under-estimation of probabilities for a given class, consider applying a prevalence adjustment to determine if it's solely due to prevalence shift. After prevalence adjustment, plot the reliability diagrams again and examine the results of calibration metrics.

For a general sense of average probability deviation, we recommend using the Cox and Loess integrated calibration index (ICI) as they don't depend on binning. Alternativly, ECE can be used to measure the same but the result will depend on the binning scheme you used. If the probabilities distribution is highly skewed toward 0 and 1, use equal-count binning for ECE.

Please refer to the notebooks for detailed descriptions of each metric.

CHAPTER

THREE

RELIABILITY DIAGRAM

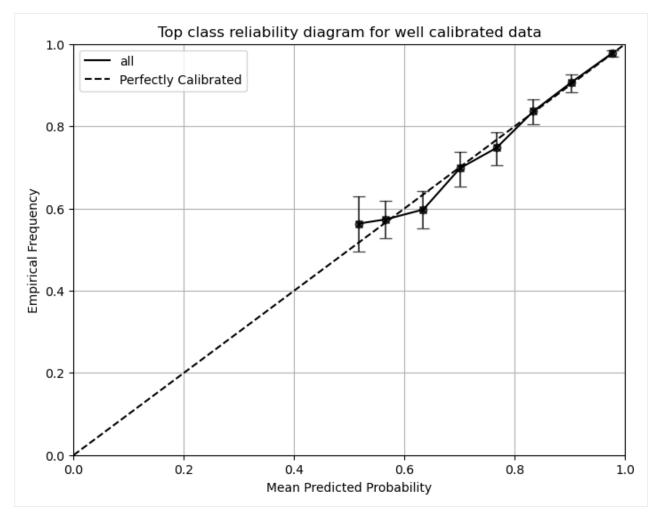
Reliability Diagram is a tool to visualize the calibration of a model given a set of data. It groups the data into bins and plots the accuracy of each bin against the average predicted value for that bin. The reliability diagram can be plotted for top-class prediction only or for a given class. The calzone package provides a function to calculate and plot the reliability diagram.

```
[13]: ### Create and plot the top-class well calibrated data
reliability,confindence,bin_edges,bin_counts = reliability_diagram(wellcal_dataloader.

→labels, wellcal_dataloader.probs, num_bins=15,class_to_plot=None) #clasto plot is None_

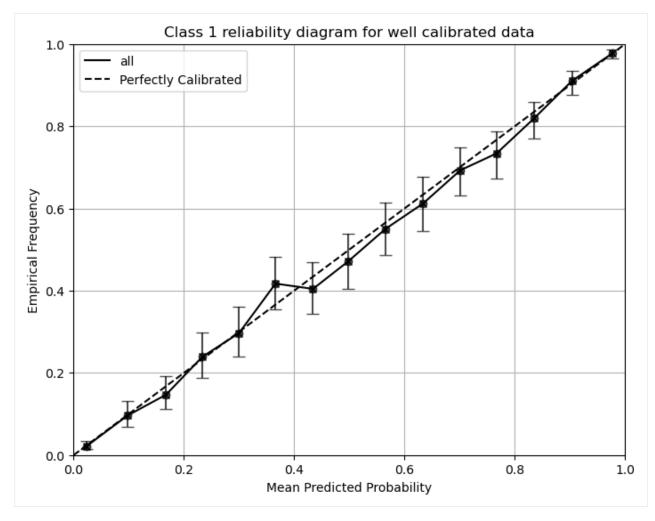
→mean calcuate for top-class
# Plot the reliability diagram
plot_reliability_diagram(reliability,confindence,bin_counts,error_bar=True,title='Top_

→class reliability diagram for well calibrated data')
```



The error bar in the reliability diagram is the 95% confidence interval calculated using wilson score interval which assume that samples in a bin is a series of Bernoulli trials with the success probability equal to the mean predicted probability. The confidence interval is only for reference and might not be exact.

Since we have a binary classification problem, The mean predicted probability will not go below 0.5 for the top-class reliability diagram. We will proceed to plot the class 1 reliability diagram.



Class-by-class reliability diagrams reveal more information about the model's calibration. The top-class reliability diagram could be misleading as it could shows reasonable calibration for the top-class, but the model could be overconfident for the other classes. We can demonstrate in the following example.

```
[40]: ### We will artificially drop the prevalence of class 1
# The top-class reliability diagram will still look good
# But the class-1 reliability diagram will be very bad
from calzone.utils import softmax_to_logits
from scipy.special import softmax
import numpy as np

test_dataloader = data_loader(data_path="../../../example_data/simulated_welldata.csv")
class_1_index = (test_dataloader.labels==1)

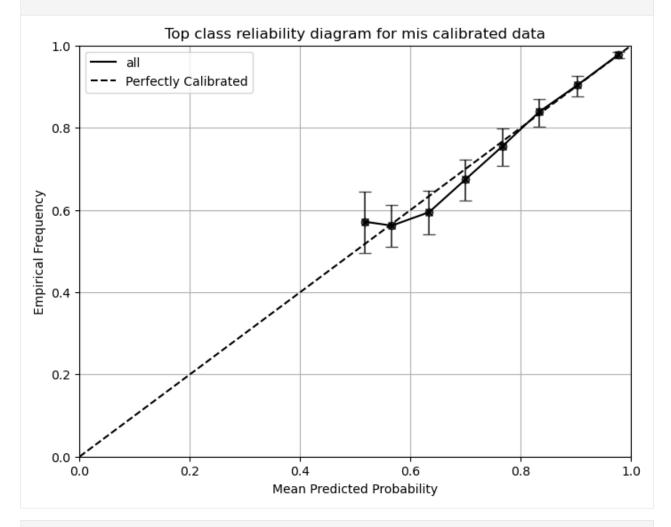
# We will drop 50% of class 1 samples
class_1_samples = np.where(class_1_index)[0]
drop_indices = np.random.choice(class_1_samples, size=int(len(class_1_samples)/2),__
--replace=False)

mask = np.ones(len(test_dataloader.labels), dtype=bool)
mask[drop_indices] = False

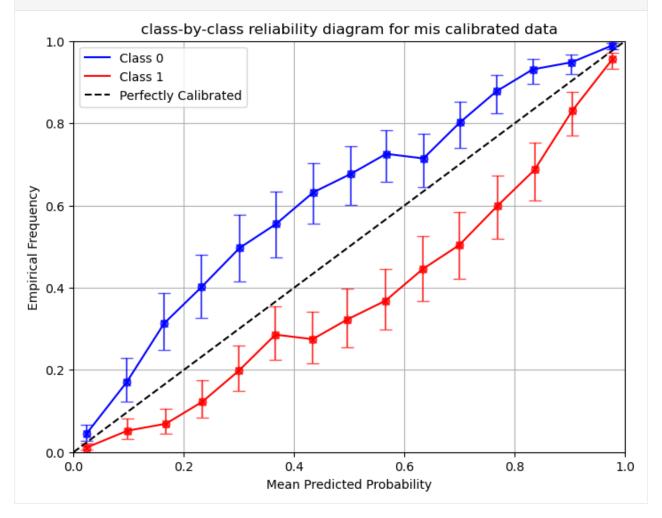
(continues on next page)
```

```
test_dataloader.labels = test_dataloader.labels[mask]
test_dataloader.probs = test_dataloader.probs[mask]
test_dataloader.data = test_dataloader.data[mask]
```

[42]: ### Create and plot the top-class reliability diagram



(continues on next page)



The example shows us that the prevalence shift in the testing data could lead to mis-calibration in a posterior sense. Calzone provides a method to calibrate it. The prevalence is not related to the model and the model could still have the correct likelihood ratio. See more discussion in the prevalence adjustment section.

3.1 References

Bröcker, J., & Smith, L. A. (2007). Increasing the Reliability of Reliability Diagrams. Weather and Forecasting, 22(3), 651–661. https://doi.org/10.1175/WAF993.1

3.1. References 21

EXEPECTED CALIBRATION ERROR(ECE) AND MAXIMUM CALIBRATION ERROR (MCE)

4.1 Theoretical Background

(True) Expected Calibration Error is defined as the average difference between the predicted probability and the true probability for a particular class. We will label the predicted probability as \hat{P} and the true probability as P and drop the class of interest label for simplicity.

$$ECE = \mathbb{E}_{\hat{P}}[|\mathbb{P}(\hat{Y} = Y | \hat{P} = p) - p|] = \int_{0}^{1} |\mathbb{P}(\hat{Y} = Y | \hat{P} = p) - p| \, dF_{\hat{P}}(p)$$

Some also define the ECE as the average difference between the predicted probability and the true probability of the predicted class. This is similar to the top-class vs class-by-class reliability diagram. We will refer to it as ECE_{top} .

$$\mathrm{ECE}_{top} = \mathbb{E}_{\hat{P_{top}}}[|\mathbb{P}(\hat{Y_{top}} = Y_{top}|\hat{P_{top}} = p_{top}) - p_{top}|])$$

Similarly, we can also define the Maximum Calibration Error (MCE) as the maximum difference between the predicted and true probabilities:

$$MCE = \max_{p} |\mathbb{P}(\hat{Y} = Y | \hat{P} = p) - p|$$

$$MCE_{top} = \max_{p_{top}} |\mathbb{P}(\hat{Y_{top}} = Y_{top} | \hat{P} = p_{top}) - p_{top}|$$

4.1.1 Estimated ECE and MCE

We can't compute ECE and MCE directly from the data, but we can group the data into bins and compute an Estimated ECE and MCE from the grouped data. In most literature, the estimated ECE and MCE are simply referred to as ECE and MCE. We adopt the same convention but want to remind the reader that the estimated ECE and MCE are not the same as the true ECE and MCE and they heavily depend on the binning method. In the text below, we will use the terms ECE and MCE to refer to the estimated ECE and MCE, and true ECE and MCE to refer to the true ECE and MCE.

$$ECE = \sum_{m=1}^{M} \frac{|B_m|}{n} |acc(B_m) - conf(B_m)|$$

where M is the number of bins, B_m is the m-th bin, $acc(B_m)$ is the accuracy of the m-th bin, $conf(B_m)$ is the mean predicted probability of the m-th bin, and n is the total number of samples. The ECE-topclass is a simple extension where only the top-class probability is used to group the data. We can define MCE similarly.

$$MCE = \max_{m} |acc(B_m) - conf(B_m)|$$

Both ECE and MCE group the data based on the predicted probability. This means the resulting ECE and MCE depend on the binning method used. Traditionally, the data are grouped into 10 equal-width bins. However, if the data is not evenly distributed, the resulting ECE and MCE may not be a good representation of the model's calibration.

Nixon et al. (2019) proposed a new binning method which bins the data into equal-count bins. They refer to the resulting ECE as Adaptive Calibration Error (ACE). The ACE also accounts for all the predictions for all classes. This is equivalent to a sum of ECE for all classes.

We adopt a modified version of the ACE and aim to measure the true ECE with equal-count binning but ignore the part about summing over all classes. This is equivalent to the ECE but only with equal-count binning. We refer to the equal-width binning ECE as ECE-H and the equal-count binning ECE as ECE-C. Similarly, we will refer to the equal-width binning MCE as MCE-H and the equal-count binning MCE as MCE-C.

4.2 Pros of ECE and MCE

ECE and MCE are perhaps the most intuitive metrics for measuring the calibration of a probabilistic model. They are simply the average deviation of the predicted probability from the true probability and the maximum deviation of the predicted probability from the true probability. The ECE and MCE are also easy to compute and interpret. ECE could be computed by doing a weighted average of the reliability diagram based on bin counts and MCE is just the maximum difference between the accuracy and the confidence. Because of that, ECE and MCE are widely used in the machine learning literature.

4.3 Cons of ECE and MCE

The biggest disadvantage of using ECE and MCE is that they rely on the binning scheme, and results will depend on the binning scheme and the number of bins. It can be shown that the expected ECE and MCE will always increase with the number of bins and the expected ECE and MCE will change with the number of samples for a fixed number of bins. This causes problems when interpreting the results. We will do a simple experiment in the following section to show this. Because of the above reasons, we recommend using ECE and MCE with other metrics that are not dependent on the binning scheme.

4.4 Calculating ECE and MCE with calzone

There is two way to calculate the ECE and MCE in calzone. The first way is by calling the function explicitly. Notice ECE_H is the equal-width binning and ECE_C is the equal-count binning.

```
→counts=bin_counts)
print("Top-class ECE-H: ",ece_h_top_class)
print("Top-class MCE-H: ",mce_h_top_class)

Top-class ECE-H: 0.009608653731328977
Top-class MCE-H: 0.03926468843081976
```

```
[3]: ### Similarily we can calculate the class 1 ECE-C and top-class ECE-C by setting is_
     →equal_freq=True
    reliability,confindence,bin_edges,bin_counts = reliability_diagram(wellcal_dataloader.
     →labels, wellcal_dataloader.probs,num_bins=10, class_to_plot=None, is_equal_freq=True)
    ece_c_top_class,mce_c_top_class = calculate_ece_mce(reliability,confindence,bin_
     reliability,confindence,bin_edges,bin_counts = reliability_diagram(wellcal_dataloader.
     →labels, wellcal_dataloader.probs, num_bins=10, class_to_plot=1, is_equal_freq=True)
    ece_c_classone,mce_c_classone = calculate_ece_mce(reliability,confindence,bin_counts=bin_
     →counts)
    print("Top-class ECE-C: ",ece_c_top_class)
    print("Top-class MCE-C: ",mce_c_top_class)
    print("Class 1 ECE-C: ",ece_c_classone)
    print("Class 1 MCE-C: ",mce_c_classone)
    Top-class ECE-C: 0.009458033653818828
    Top-class MCE-C: 0.020515047600205505
    Class 1 ECE-C: 0.008733966945443138
    Class 1 MCE-C: 0.02324031223486256
```

The second method is much simpler. We can use the calzone.metrics.CalibrationMetrics class to calculate all type of metrics.

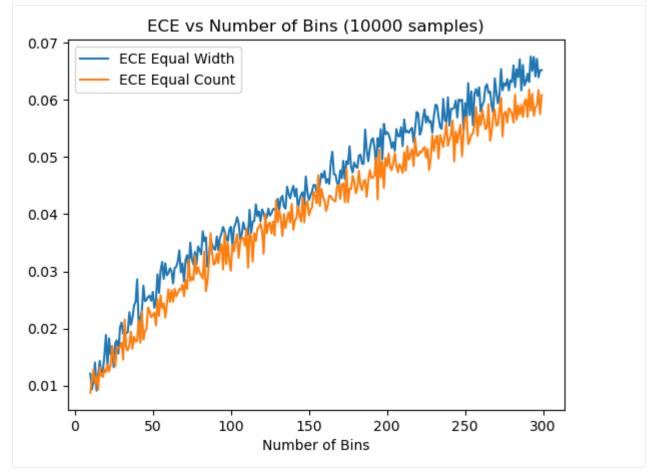
4.5 ECE and MCE as function of bin size

In this section, we want to quickly demonstrate how binning could affect the ECE and MCE.

```
[5]: range_of_binning = np.arange(10,300,1)
    result = np.zeros((len(range_of_binning),8))
    for i in range(len(range_of_binning)):
        calmetrics = CalibrationMetrics(class_to_calculate=1,num_bins=range_of_binning[i])
        result[i,:] = calmetrics.calculate_metrics(wellcal_dataloader.labels, wellcal_
        dataloader.probs, metrics=['ECE-H', 'MCE-H', 'ECE-C', 'MCE-C'],return_numpy=True)
```

```
[6]: import matplotlib.pyplot as plt
plt.plot(range_of_binning,result[:,1], label='ECE Equal Width')
plt.plot(range_of_binning,result[:,5], label='ECE Equal Count')
plt.title('ECE vs Number of Bins (10000 samples)')
#plt.ylim(0, np.max(ece_equal_width)+0.1)
plt.xlabel('Number of Bins')
plt.legend()
```

[6]: <matplotlib.legend.Legend at 0x7fe048e8b470>



We can see that the error goes up as the number of bins increases while equal count and equal width always return different results. Opposite effects can be observed for the number of sample. Therefore, ECE and MCE can only give us a rough estimate of the calibation error of the model.

4.6 Reference

Guo, C., Pleiss, G., Sun, Y., & Weinberger, K. Q. (2017). On Calibration of Modern Neural Networks (No. arXiv:1706.04599). arXiv. http://arxiv.org/abs/1706.04599

Pakdaman Naeini, M., Cooper, G., & Hauskrecht, M. (2015). Obtaining Well Calibrated Probabilities Using Bayesian Binning. Proceedings of the AAAI Conference on Artificial Intelligence, 29(1). https://doi.org/10.1609/aaai.v29i1. 9602.

Arrieta-Ibarra, I., Gujral, P., Tannen, J., Tygert, M., & Xu, C. (2022). Metrics of Calibration for Probabilistic Predictions.

Nixon, J., Dusenberry, M. W., Zhang, L., Jerfel, G., & Tran, D. (2020). Measuring Calibration in Deep Learning.

4.6. Reference 27



HOSMER-LEMESHOW TEST (HL TEST)

5.1 Theoretical Background

The Hosmer-Lemeshow test (HL Test) is a statistical test that can be used to assess the calibration of a probabilistic model. The test works by dividing the predicted probabilities into groups (typically deciles) and comparing the observed and expected frequencies of events in each group. A non-significant p-value (usually greater than 0.05) indicates we cannot reject the hypothesis that the model is well-calibrated, while a significant p-value suggests the opposite. The Hosmer-Lemeshow test is widely used in the literature and industry since it is simple to implement and interpret.

In order to calculate the Hosmer-Lemeshow test statistic, we need to first determine the binning scheme used to divide the predicted probabilities into groups. Conventionally, the predicted probabilities are divided into 10 equal-width bins. We will label the equal-width binning Hosmer-Lemeshow test as HL-H and equal-count binning Hosmer-Lemeshow test as HL-C. The Hosmer-Lemeshow test statistic is then calculated as follows:

$$\mathrm{HL} = \sum_{m=1}^{M} \left[\frac{(O_{1,m} - E_{1,m})^2}{E_{1,m}} + \frac{(O_{0,m} - E_{0,m})^2}{E_{0,m}} \right] = \sum_{m=1}^{M} \frac{(O_{1,m} - E_{1,m})^2}{E_{1,m} (1 - \frac{E_{1,m}}{N_m})} \sim \chi_{M-2}^2$$

where $E_{1,m}$ is the expected number of class 1 events in the \mathbf{m}^{th} bin, $O_{1,m}$ is the observed number of class 1 events in the \mathbf{m}^{th} bin, N_m is the total number of observations in the \mathbf{m}^{th} bin, and M is the number of bins. The HL test statistic is distributed as a chi-squared distribution with M-2 degrees of freedom. We can then use this test statistic to calculate the p-value for the test and determine whether we can reject the null hypothesis that the model is well-calibrated. Notice that the degree of freedom of HL test is M-2 by default but the degree of freedom should be M when the samples is not used for training (Hosmer Jr et. al., 2013). We provides the option to specify the degree of freedom in the calzone. The default value is still M-2.

5.2 Pros of HL test

The Hosmer-Lemeshow test offers several advantages in assessing calibration. It is a non-parametric test, meaning it does not require any assumptions about the distribution of the predicted probabilities, and it provides statistical meaning to the resulting test statistic. It is also very intuitive and easy to understand since it is just a chi-square based test. It can be calculated from the result of the reliability diagram. The HL test is widely used in the literature as a validation method for model calibration.

5.3 Cons of HL Test

Many studies have shown that the HL test is not an ideal way to examine the calibration of a model. The biggest problem is that the HL test depends on the binning scheme used. Whether equal-width or equal-count binning and the number of bins used can affect the results of the HL test. It is shown that the standard 10 equal-width bins often have the wrong size and low statistical power. Therefore, it is recommended not to use the HL test to examine the calibration of a model (Hosmer et. al., 1997). However, the HL test is still a useful tool to quickly check the calibration of a model and provide a reference for the calibration of a model.

5.4 Calculating HL test statistics and p-value with calzone

There are again two ways to calculate the HL test statistics and p-value with calzone. One is to call the function explicitly, and the other is to use the calzone.metrics.CalibrationMetrics class.

```
[1]: from calzone.utils import reliability_diagram,data_loader
    from calzone.metrics import hosmer_lemeshow_test
    import numpy as np
    ### loading the data
    wellcal_dataloader = data_loader(data_path="../../example_data/simulated_welldata.csv
    ")
    ### calculating the HL-H TS
    reliability_diagram(wellcal_dataloader.
    →labels, wellcal_dataloader.probs, num_bins=10, class_to_plot=1, is_equal_freq=False)
    #df = 10 instead 8 since it is validation data
    HL_H_ts,HL_H_p,df = hosmer_lemeshow_test(reliability,confindence,bin_count=bin_counts,_
    \rightarrowdf=10)
    print("HL-H Test Statistic: ",HL_H_ts)
    print("HL-H p-value: ",HL_H_p)
    HL-H Test Statistic: 8.884991559088098
    HL-H p-value: 0.5430520576015005
[2]: ### similar for HL-C model
    reliability_diagram(wellcal_dataloader.
    →labels, wellcal_dataloader.probs, num_bins=10, class_to_plot=1, is_equal_freq=True)
```

We can see the result from equal-width binning and equal-count binning are different. We will demostrate how to use the calzone.metrics.CalibrationMetrics class.

HL_C_ts,HL_C_p,df = hosmer_lemeshow_test(reliability,confindence,bin_count=bin_counts,_

```
[3]: ### using the CalibrationMetrics class
from calzone.metrics import CalibrationMetrics

(continues on next page)
```

 \rightarrow df=10)

print("HL-C Test Statistic: ",HL_C_ts)

HL-C Test Statistic: 3.694947603203135 HL-C p-value: 0.9600610057855211

print("HL-C p-value: ",HL_C_p)

31

5.5 P-value of HL test

We will show the statistical size of HL test. The statistical size refer to the type 1 error rate of a statistical test and idealy should match the p-value. Notice that the size of HL test had been shown to depend on sample size, number of bin and binning scheme (Hosmer et. al. 1997). We will generate fake data to show the size of HL test.

```
[4]: ### The size of HL Test
    from calzone.utils import fake_binary_data_generator
    np.random.seed(123)
    fakedata_generator = fake_binary_data_generator(alpha_val=0.5, beta_val=0.5)
    cal_metrics = CalibrationMetrics()
    sample_size = 1000
    simulation size = 10000
    results = []
     # generate data
    for i in range(simulation_size):
         X, y = fakedata_generator.generate_data(sample_size)
             tempresult = cal_metrics.calculate_metrics(y, X, ['HL-H', 'HL-C'],return_
     \rightarrownumpy=False, df=10)
             keys = list(tempresult.keys())
             results.append(np.array(list(tempresult.values())))
         else:
             tempresult = cal_metrics.calculate_metrics(y, X, ['HL-H', 'HL-C'], return_
     \rightarrownumpy=True, df=10)
             results.append(tempresult)
    results = np.array(results)
```

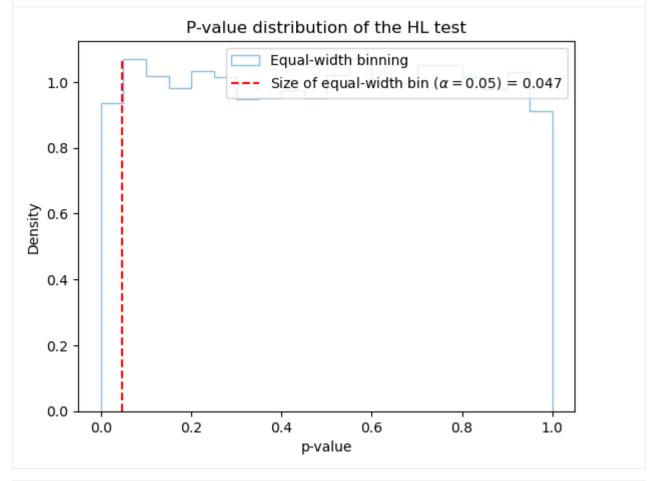
```
[5]: ### Showing the size of the model
import matplotlib.pyplot as plt
hl_h_pvalue = results[:,1]
hl_c_pvalue = results[:,3]
size_h = np.mean(hl_h_pvalue < 0.05)
size_c = np.mean(hl_c_pvalue < 0.05)
print("The size of HL-H is :", round(size_h,3))
print("The size of HL-C is :", round(size_c,3))</pre>
The size of HL-H is : 0.047
The size of HL-C is : 0.055
```

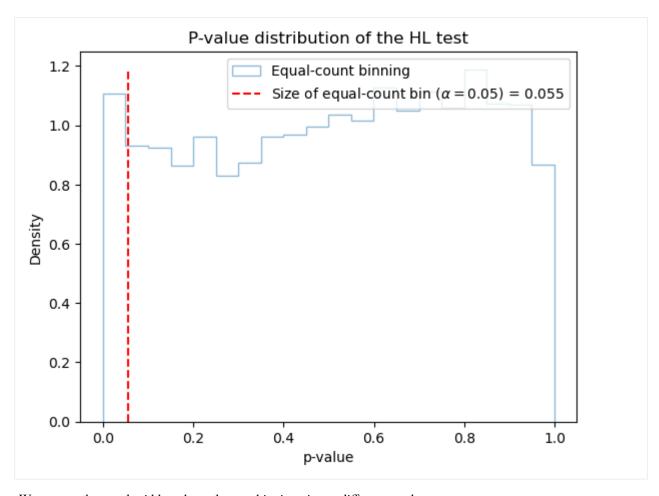
```
[6]: value,_,_=plt.hist(hl_h_pvalue, bins=20, density=True, alpha=0.5, label='Equal-width_

__binning',histtype='step')

(continues on next page)
```

[6]: Text(0.5, 1.0, 'P-value distribution of the HL test')





We can see the equal-width and equal-count binning gives a different result.

5.6 Reference

Hosmer, D. W., & Lemesbow, S. (1980). Goodness of fit tests for the multiple logistic regression model. Communications in statistics-Theory and Methods, 9(10), 1043-1069.

Hosmer, D. W., Hosmer, T., Cessie, S. L., & Lemeshow, S. (1997). A COMPARISON OF GOODNESS-OF-FIT TESTS FOR THE LOGISTIC REGRESSION MODEL. 16.

Hosmer Jr, David W., Stanley Lemeshow, and Rodney X. Sturdivant. Applied logistic regression. John Wiley & Sons, 2013.

5.6. Reference 33

SIX

COX CALIBRATION ANALYSIS

6.1 Theoretical Background

Cox calibration analysis is both a logistic recalibration technique and a method to examine the current calibration of a model. To perform the analysis, we first to fit a logistic regression model using logit (log odds, aka $\log \frac{\hat{p}}{1-\hat{p}}$) as the predictor variable and the outcome as the target variable.

$$p_{new} = \frac{1}{1 + e^{-(a+b \cdot \log \frac{\hat{p}}{1-\hat{p}})}}$$

In the case of perfect calibration, $P(Y=1|p=\hat{p})=\hat{p}$ and the new probability p_{new} is equal to the original probability \hat{p} . That means a=0 and b=1. If b>1, the model is under-confident at high probabilities and over-confident at low probabilities for the class-of-interest. If b<1, the model is over-confident at high probabilities and under-confident at low probabilities for the class-of-interest. If a>0, the model is over-confident at all probabilities for the class-of-interest. If a<0, the model is under-confident at all probabilities for the class-of-interest. The confidence interval of a and b can be used to guide the calibration of the model. The user can also choose to fix a=0 and fit for b only and vice versa, then there will be no interaction between a and b and the confidence interval can be used as a statistical test to test for perfect calibration.

6.2 Pros of Cox calibration analysis

Cox calibration analysis doesn't depend on binning of data, which is a big advantage since common metrics such as ECE/MCE and HL test all depend on binning and we have shown that changing binning can lead to different results. We can also use it to perform statistical tests by fixing a to 0 and test whether b=1 and the other way around to test for perfect calibration. Also, the fitted values of a and b can tell us how the model is miscalibrated, whether it is an overall under- or over-confidence or if it is over-confident in some ranges and under-confident in others. For example, if a is not close to 0 while b is close to 1, it likely indicates a prevalence shift. See more details in the prevalence adjustment notebook.

6.3 Cons of Cox calibration analysis

Cox Calibration analysis can only assess weak calibration , meaning whether $P(Y=1|\hat{p}=p)=p$ for all p. It only captures certain types of miscalibration (general over/under-confidence). A model can have a=0 and b=1 and still be miscalibrated. For example, a model is over-confident at low and high probabilities and under-confident at intermediate probabilities could has a=0 and b=1 but still be miscalibrated.

6.4 Calculating Cox slope and intercept with calzone

There are two ways to calculate the Cox slope and intercept. Calling the Cox function gives you more control over the calculation, including fixing a = 0 or b = 1.

```
[1]: from calzone.utils import reliability_diagram,data_loader
    from calzone.metrics import cox_regression_analysis
    import numpy as np
    ### loading the data
    wellcal_dataloader = data_loader(data_path="../../example_data/simulated_welldata.csv
    ")
    ### calculating cox slope and intercept
    cox_slope, cox_intercept,cox_slope_ci,cox_intercept_ci = cox_regression_analysis(wellcal_
    →dataloader.labels, wellcal_dataloader.probs,class_to_calculate=1,print_results=True)
                            Logit Regression Results
    Dep. Variable:
                                      No. Observations:
                                                                      5000
                                Logit Df Residuals:
    Model:
                                                                      4998
                                  MLE Df Model:
    Method:
                                                                         1
    Date:
                      Wed, 02 Oct 2024 Pseudo R-squ.:
                                                                    0.4438
                              15:53:27 Log-Likelihood:
    Time:
                                                                    -1927.5
                                 True LL-Null:
                                                                    -3465.6
    converged:
    Covariance Type:
                             nonrobust LLR p-value:
                                                                     0.000
    coef
                          std err
                                                P>|z|
                                                          [0.025
                                                                     0.975]
                            0.040 -1.123
0.029 34.212
                -0.0450
                                                0.262
    const
                                                          -0.123
                                                                     0.034
                 0.9942
                                                0.000
    x1
                                                                     1.051
```

You can also fix the intercept a=0 by using the fix_intercept=True option. Similarly, you can fix the slope b=1 by using the fix_slope=True option.

```
[2]: ### fixing intercept and calculating cox slope
    cox_slope, cox_intercept,cox_slope_ci,cox_intercept_ci = cox_regression_analysis(wellcal_
     →dataloader.labels, wellcal_dataloader.probs,class_to_calculate=1, fix_intercept=True,
     →print_results=True)
    Warning: Maximum number of iterations has been exceeded.
             Current function value: 0.385628
             Iterations: 0
                               Logit Regression Results
                                            No. Observations:
    Dep. Variable:
                                                                              5000
    Model:
                                    Logit
                                            Df Residuals:
                                                                              4999
    Method:
                                           Df Model:
                                      MLE
    Date:
                         Wed, 02 Oct 2024
                                           Pseudo R-squ.:
                                                                            0.4436
    Time:
                                 15:53:27 Log-Likelihood:
                                                                           -1928.1
    converged:
                                    False LL-Null:
                                                                           -3465.6
    Covariance Type:
                                nonrobust
                                            LLR p-value:
                                                                               nan
```

	coef	std err	Z	P> z	[0.025	0.975]
const	0	0	nan	nan	0	0
x1	0.9939	0.029	34.210	0.000	0.937	1.051

Alternatively, we can use the CalibrationMetrics class to compute the COX slope and intercept.

```
[3]: from calzone.metrics import CalibrationMetrics
    calmetrics = CalibrationMetrics()
    calmetrics.calculate_metrics(wellcal_dataloader.labels, wellcal_dataloader.probs,_
    →metrics=['COX'],print_results=True)
                             Logit Regression Results
    Dep. Variable:
                                         No. Observations:
                                                                        5000
                                 Logit Df Residuals:
    Model:
                                                                        4998
    Method:
                                   MLE Df Model:
                       Wed, 02 Oct 2024 Pseudo R-squ.:
    Date:
                                                                      0.4438
    Time:
                              15:53:27 Log-Likelihood:
                                                                     -1927.5
                                  True LL-Null:
    converged:
                                                                     -3465.6
    Covariance Type:
                             nonrobust LLR p-value:
                                                                       0.000
    ______
                                                                      0.975]
                   coef std err
                                                 P>|z|
                 -0.0450
                                      -1.123
    const
                             0.040
                                                 0.262
                                                           -0.123
                                                                       0.034
                  0.9942
                             0.029
                                      34.212
                                                 0.000
                                                            0.937
                                                                       1.051
    x1
[3]: {'COX coef': 0.9942499557748269,
     'COX intercept': -0.04497652296600376,
     'COX coef lowerci': 0.9372902801721911,
     'COX coef upperci': 1.0512096313774626,
     'COX intercept lowerci': -0.12348577118577644,
     'COX intercept upperci': 0.03353272525376893,
     'COX ICI': 0.005610391483826338}
```

The resulting COX slope and intercept can be used to calibrate the model but it is beyond the scope of this package.

6.5 P-value of COX slope and intecept test

Although Cox calibration analysis is usually only used to estimate the overall calibration trend, the resulting estimates of the slope and intercept can also be used to test whether the model is well calibrated (moderate calibration). We will do a demostrate on the statistical size of the slope and intercept test below

```
[]: ### The size of slope test
from calzone.utils import fake_binary_data_generator
np.random.seed(123)
fakedata_generator = fake_binary_data_generator(alpha_val=0.5, beta_val=0.5)
cal_metrics = CalibrationMetrics()

(continues on next page)
```

```
sample_size = 1000
simulation_size = 1000
results = []
# generate data
for i in range(simulation_size):
   X, y = fakedata_generator.generate_data(sample_size)
   if i == 0:
        tempresult = cal_metrics.calculate_metrics(y, X, ['COX'],return_numpy=False,fix_
→intercept=True) #we need to fix the intercept to be 0
        keys = list(tempresult.keys())
       results.append(np.array(list(tempresult.values())))
    else:
        tempresult = cal_metrics.calculate_metrics(y, X, ['COX'],return_numpy=True,fix_
→intercept=True) #we need to fix the intercept to be 0
        results.append(tempresult)
results = np.array(results)
```

```
[5]: Cox_slope = results[:,0]
   Cox_slope_lowerci = results[:,2]
   Cox_slope_upperci = results[:,3]
   chance = np.logical_and(Cox_slope_lowerci<=1, Cox_slope_upperci>=1)
   print('The size of the Cox slope test is: ', 1-np.mean(chance))

The size of the Cox slope test is: 0.039000000000000035
```

We can also do the intercept test:

```
[]: ### The size of intercept test
    from calzone.utils import fake_binary_data_generator
    np.random.seed(123)
    fakedata_generator = fake_binary_data_generator(alpha_val=0.5, beta_val=0.5)
    cal_metrics = CalibrationMetrics()
    sample_size = 1000
    simulation_size = 1000
    results = []
    # generate data
    for i in range(simulation_size):
        X, y = fakedata_generator.generate_data(sample_size)
        if i == 0:
            tempresult = cal_metrics.calculate_metrics(y, X, ['COX'],return_numpy=False,fix_
     →slope=True) #we need to fix the slope to be 1
            keys = list(tempresult.keys())
            results.append(np.array(list(tempresult.values())))
        else:
            tempresult = cal_metrics.calculate_metrics(y, X, ['COX'],return_numpy=True,fix_
     →slope=True) #we need to fix the slope to be 1
            results.append(tempresult)
    results = np.array(results)
```

```
[7]: Cox_intercept = results[:,1]
Cox_intercept_lowerci = results[:,4]

(continues on next page)
```

```
Cox_intercept_upperci = results[:,5]
chance = np.logical_and(Cox_intercept_lowerci<=0, Cox_intercept_upperci>=0)
print('The size of the Cox intercept test is: ', 1-np.mean(chance))
The size of the Cox intercept test is: 0.05600000000000000
```

The test results in different size values but they are closer than with the HL test.

6.6 References

Cox, D. R. (1958). Two Further Applications of a Model for Binary Regression.

Calster, B. V., & Steyerberg, E. W. (2018). Calibration of Prognostic Risk Scores. In R. S. Kenett, N. T. Longford, W. W. Piegorsch, & F. Ruggeri (Eds.), Wiley StatsRef: Statistics Reference Online (1st ed., pp. 1–10). Wiley. https://doi.org/10.1002/9781118445112.stat08078

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6.6. References 39

INTEGRATED CALIBRATION INDEX (ICI)

7.1 Theoretical Background

Integrated Calibration Index (ICI) is essentially the same as expected calibration error (ECE) in terms of the idea. They both try to measure the average deviation of the predicted probabilities from the true probabilities. However, ECE is calculated by grouping the samples into bins and then calculating the weighted average of the deviation of the mean predicted probabilities from the empirical accuracy. ICI, on the other hand, is calculated by fitting a smooth curve using the samples itself and therefore doesn't require binning. However, the choice of the curve fitting method can affect the result and is arbitrary. The most common choice is locally estimated scatterplot smoothing (LOESS) (Cleveland, 1979). People also use other methods such as polynomial fitting and spline fitting. Interestingly, not many people have looked into using COX regression results to calculate ICI, which is implemented in calzone. Notice that the Cox-ICI can be way off from the truth if the logistic regression does not result in a good fit.

The formula for ICI is:

$$ICI = \int_0^1 |\mathbb{S}(p) - p| \, dF_{\hat{P}}(p)$$

where $\mathbb{S}(p)$ is the fitted function and $F_{\hat{P}}(p)$ is the cumulative distribution function of the empirical probabilities. In the actual implementation, we calculate

$$ICI = \frac{1}{N} \sum_{i=1}^{N} |\mathbb{S}(p_i) - p_i|$$

where p_i is the predicted probability of the *i*-th sample.

7.2 Pros of ICI

The main advantage of ICI is that it skips the need for binning while still giving an easily interpretable metric. It is essentially the same as ECE and can be interpreted as the average deviation from the true probability. ICI can capture any type of miscalibration if the calibration curve is well-described by the fitting method.

7.3 Cons of ICI

The main disadvantage of ICI is the need for a fitting method. Locally estimated scatterplot smoothing (LOESS) is a non-parametric regression method that fits a smooth line through the data. It is the most common method used in ICI. However, it still requires hyperparameters like the span (window width) which could affect the fitting result and ICI greatly.

7.4 Calculating LOESS ICI and COX ICI using calzone

To calculate LOESS ICI and COX ICI using calzone, we can call the function directly

```
[1]: from calzone.utils import reliability_diagram,data_loader
    from calzone.metrics import cox_regression_analysis,lowess_regression_analysis,cal_ICI_
     \hookrightarrowCOX
    ### loading the data
    wellcal_dataloader = data_loader(data_path="../../example_data/simulated_welldata.csv
    ### calculating cox ICI
    cox_slope, cox_intercept,cox_slope_ci,cox_intercept_ci = cox_regression_analysis(wellcal_
     →dataloader.labels, wellcal_dataloader.probs,class_to_calculate=1,print_results=False)
    cox_ici = cal_ICI_cox(cox_slope, cox_intercept, wellcal_dataloader.probs, class_to_
     ⇔calculate=1)
    ### calculating loess ICI
    loess_ici, lowess_fit_p, lowess_fit_p_correct = lowess_regression_analysis(wellcal_
     →dataloader.labels, wellcal_dataloader.probs, class_to_calculate=1, span=0.5, delta=0.
     \rightarrow001, it=0)
    print(f"Cox ICI: {cox_ici}")
    print(f"Loess ICI: {loess_ici}")
    Cox ICI: 0.005610391483826338
    Loess ICI: 0.00558856942568957
```

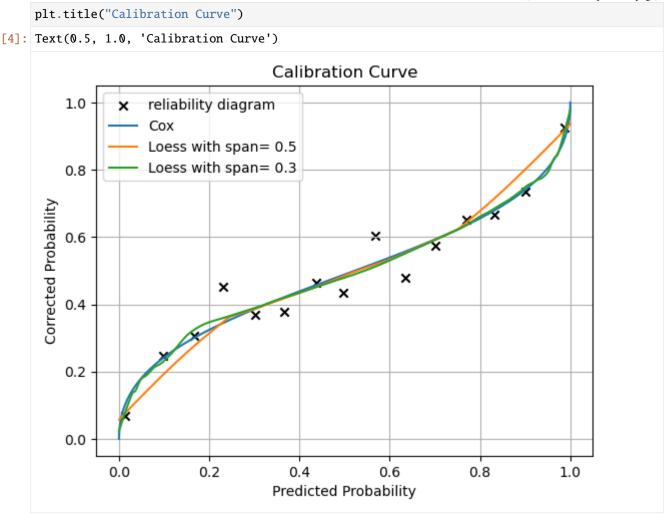
Alternatively, we can use the CalibrationMetrics class to compute the COX and Loess ICI

7.5 Visualization of the fitted curve

We can also plot the loess curve and the COX curve.

```
[3]: ### We will use linear miscalibrated data to demostrate the plot. Notice that the COX_
     → should capture the miscalibration perfectly in the example case.
    miscal_dataloader = data_loader(data_path=".../.../example_data/simulated_misdata.csv")
    ### calculating cox ICI
    cox_slope, cox_intercept,cox_slope_ci,cox_intercept_ci = cox_regression_analysis(miscal_
     →dataloader.labels, miscal_dataloader.probs,class_to_calculate=1,print_results=False)
    cox_ici = cal_ICI_cox(cox_slope, cox_intercept, wellcal_dataloader.probs, class_to_
     →calculate=1)
    ### calculating loess ICI
    loess_ici, lowess_fit_p, lowess_fit_p_correct = lowess_regression_analysis(miscal_

→dataloader.labels, miscal_dataloader.probs, class_to_calculate=1)
    ### We also try a different span
    loess_ici2, lowess_fit_p2, lowess_fit_p2_correct = lowess_regression_analysis(miscal_
     →dataloader.labels, miscal_dataloader.probs, class_to_calculate=1, span=0.3, delta=0.
     \rightarrow001, it=0)
    print("Cox ICI: ", cox_ici)
    print("Loess ICI (span = 0.5): ", loess_ici)
    print("Loess ICI (span = 0.3): ", loess_ici2)
    Cox ICI: 0.0984122810555748
    Loess ICI (span = 0.5): 0.07356445428053172
    Loess ICI (span = 0.3): 0.07692218401743334
[4]: ### ploting the curve
    from calzone.metrics import logit_func
    import matplotlib.pyplot as plt
    from calzone.utils import reliability_diagram
    from calzone.vis import plot_reliability_diagram
    import numpy as np
    cox_func= logit_func(coef=cox_slope, intercept=cox_intercept)
    proba_class1 = np.sort(np.clip(miscal_dataloader.probs[:,1],1e-10,1-1e-10))
    cox_p_correct=cox_func(proba_class1)
    reliability.confindence,bin_edges,bin_counts = reliability_diagram(miscal_dataloader.
     →labels, miscal_dataloader.probs, num_bins=15,class_to_plot=1)
    plt.scatter(confindence,reliability, label="reliability diagram", marker="x",color='k')
    #ig = plot_reliability_diagram(reliability, confindence, bin_counts)
    #plt.close()
    #print(fig)
    plt.plot(proba_class1, cox_p_correct, label="Cox")
    plt.plot(lowess_fit_p, lowess_fit_p_correct, label="Loess with span= 0.5")
    plt.plot(lowess_fit_p2, lowess_fit_p2_correct, label="Loess with span= 0.3")
    plt.xlabel("Predicted Probability")
    plt.ylabel("Corrected Probability")
    plt.legend()
    plt.grid()
```



7.6 Reference

Cleveland, W.S. (1979) "Robust Locally Weighted Regression and Smoothing Scatterplots". Journal of the American Statistical Association 74 (368): 829-836.

Austin, P. C., & Steyerberg, E. W. (2019). The Integrated Calibration Index (ICI) and related metrics for quantifying the calibration of logistic regression models. Statistics in Medicine, 38(21), 4051–4065. https://doi.org/10.1002/sim.8281

Huang, Y., Li, W., Macheret, F., Gabriel, R. A., & Ohno-Machado, L. (2020). A tutorial on calibration measurements and calibration models for clinical prediction models. Journal of the American Medical Informatics Association, 27(4), 621–633. https://doi.org/10.1093/jamia/ocz228

EIGHT

SPIEGELHALTER'S Z-TEST

8.1 Theoretical background

Spiegelhalter's Z-test is a statistical test that tests whether a probabilistic model is calibrated. It is named after the statistician David Spiegelhalter, who proposed it in 1986. It is a non-parametric test that does not require any binning.

The Spiegelhalter's Z-test was inspired by the fact that the Brier score (mean squared error) can be decomposed into reliability and resolution. In fact, any proper scoring rule can be decomposed into reliability and resolution, as shown by Brocker (2008). For example, the cross-entropy can be decomposed into KL-divergence (reliability) and entropy (resolution).

The Brier score can be decomposed into reliability and resolution as follows:

$$B = \frac{1}{N} \sum_{i=1}^{N} (x_i - p_i)^2 = \frac{1}{N} \sum_{i=1}^{N} (x_i - p_i)(1 - 2p_i) + \frac{1}{N} \sum_{i=1}^{N} p_i(1 - p_i)$$

Where the first term measure the reliability/calibration and the second term measure the resolution/discrimination.

The Variance of the Brier score is:

$$Var(B) = \frac{1}{N^2} \sum_{i=1}^{N} (1 - 2p_i)^2 p_i (1 - p_i)$$

and the Speigelhalter's Z-test is defined as:

$$Z = \frac{B - E(B)}{\sqrt{\text{Var}(B)}} = \frac{\sum_{i=1}^{N} (x_i - p_i)(1 - 2p_i)}{\sum_{i=1}^{N} (1 - 2p_i)^2 p_i (1 - p_i)}$$

and Z is approximately standard normal distributed under the null hypothesis of calibration. Spiegelhalter's Z-test has the right size in many situations and it is powerful in many situations.

8.2 Pros of Spiegelhalter's Z test

Spiegelhalter's Z test is a statistical test which can provide statistical evidence that the null hypothesis (well-calibrated) is true or false. It is a non-parametric test and doesn't require any hyperparameter tuning. It also doesn't require any binning of data, which is extremely useful compared to the Hosmer-Lemeshow test.

8.3 Cons of Spiegelhalter's Z test

The power of Spiegelhalter's Z test is limited for some cases of miscalibration, such as prevalence shift. However, it is a very powerful test for many other cases of miscalibration.

8.4 Calculating the Spieegelhalter Z score and p-value using calzone

We can call functions from the calzone package to calculate the Spiegelhalter Z score and p-value directly.

We can also use the CalibrationMetrics class

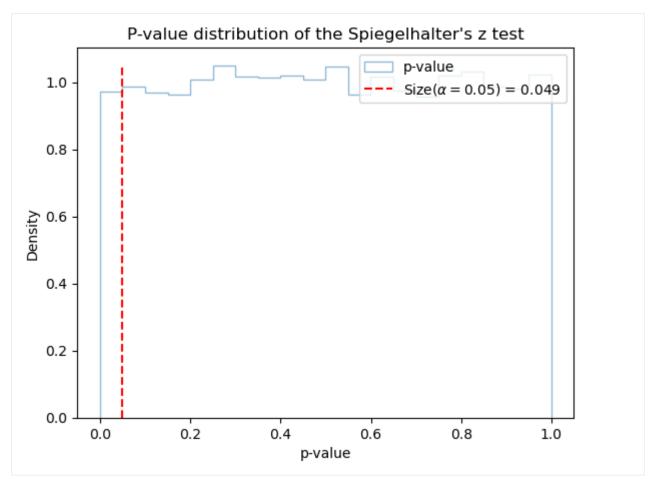
8.5 P-value of Spiegelhalter's z test

Like to HL test, we can check whether the Spiegelhalter's z test has the correct size.

```
results.append(np.array(list(tempresult.values())))
else:
    tempresult = cal_metrics.calculate_metrics(y, X, ['SpiegelhalterZ'],return_
    numpy=True)
    results.append(tempresult)
results = np.array(results)
[4]: ### Showing the size of the model
import matplotlib.pyplot as plt
z scores = results[: 0]
```

```
[4]: ### Showing the size of the model
import matplotlib.pyplot as plt
z_scores = results[:,0]
p_values = results[:,1]
size = np.mean(p_values < 0.05)
print("The size of Spiegelhalter's z test is :", round(size,3))
The size of Spiegelhalter's z test is : 0.049</pre>
```

[5]: Text(0.5, 1.0, "P-value distribution of the Spiegelhalter's z test")



We can see that Spiegelhalter's Z test has a accurate size.

8.6 Reference

Spiegelhalter, D. J. (1986). Probabilistic prediction in patient management and clinical trials.

Bröcker, J. (2009). Reliability, Sufficiency, and the Decomposition of Proper Scores. Quarterly Journal of the Royal Meteorological Society, 135(643), 1512–1519. https://doi.org/10.1002/qj.456

PREVALENCE ADJUSTMENT

In this notebook, we will discuss how prevalence will affect the calibration of the model in a binary classification problem and how to adjust for prevalence differences.

When we discuss calibration, we usually refer to whether the probability output by the model matches the posterior probability of the true outcome.

$$P(D = 1 | \hat{p} = p) = p, \forall p \in [0, 1]$$

where \hat{p} is the predicted probability of the true outcome being 1.

However, the posterior probability of the true outcome being 1 depends on the prevalence of the outcome 1. Using Bayes' theorem, we can derive the following relationship:

$$P(D = 1|\hat{p} = p) = \frac{P(\hat{p} = p|D = 1)P(D = 1)}{P(\hat{p} = p)}$$

The term $P(\hat{p}=p|D=1)$ is independent of prevalence for a given model. The term P(D=1) is the prevalence of the outcome 1. The term $P(\hat{p}=p)$ is the marginal probability of the predicted probability being p and implicitly depends on the prevalence of the true outcome. We can expand the denominator using the fact that $P(\hat{p}=p)=P(\hat{p}=p|D=1)\eta+P(\hat{p}=p|D=0)(1-\eta)$. Further rearranging the above equation will lead to the following equation:

$$P(D=1|\hat{p}=p) = \frac{\mathsf{LR}(p) \times \eta}{\mathsf{LR}(p) \times \eta + 1 - \eta}$$

where $LR(p) = \frac{P(\hat{p} = p \mid D = 1)}{P(\hat{p} = p \mid D = 0)}$ is the likelihood ratio of the predicted probability being p given the true outcome being 1 and 0 respectively, and η is the prevalence of the outcome 1.

The likelihood ratio is independent of the prevalence, so that the model can be calibrated for a specific prevalence but will become mis-calibrated for a different prevalence. We can say such a model is "intrinsically calibrated", meaning that the likelihood ratio of the model with a specific prevalence produced a correct posterior probability of the true outcome being 1.

An intrinsically calibrated model can be adapted to a population with a different prevalence but the same probability distribution within class. To adjust for prevalence differences, we rely on the fact that the likelihood ratio is independent of the prevalence. We can use the following equation to adjust the predicted probability of the true outcome being 1 for a different prevalence:

$$P(D=1|\hat{p}=p) = \frac{\eta L R(p)}{\eta L R(p) + (1-\eta)} = p$$

$$LR(p) = \frac{p}{1-p} \cdot \frac{1-\eta}{\eta}$$

$$P'(D=1|\hat{p}=p) = \frac{\eta' L R(p)}{\eta' L R(p) + (1-\eta')} = \frac{\eta'/(1-\eta')}{(1/p-1)(\eta/(1-\eta))} = p'$$

where η is the prevalence of the derivation population (aka the population for which the model is calibrated) and η' is the prevalence of the outcome 1 in the new population. We will refer to p' as the adjusted probability.

In practice, we might have a dataset with the true label (which we can use to calculate the prevalence η) and predicted probability of the true outcome being 1. We can search for the derivation prevalence η that minimizes cross-entropy loss between the adjusted probability p' and the posterior probability of the true outcome being 1.

$$\min_{\eta} \sum_{i=1}^{N} (y_i \log(p_i') + (1 - y_i) \log(1 - p_i'))$$

Notice that minimizing cross-entropy loss with respect to η is equivalent to minimizing the KL divergence since the prevalence adjustment is a monotonic transformation and doesn't affect the resolution component of the cross-entropy loss.

9.1 Preform prevalence adjustment in calzone

We will demonstrate how to perform prevalence adjustment in calzone. The first method is to find optimal prevalence first and apply the adjustment.

```
[1]: ##
    from calzone.utils import find_optimal_prevalence,apply_prevalence_adjustment,data_
     →loader, fake_binary_data_generator
    import numpy as np
    # We generate data and drop the prevalence
    np.random.seed(123)
    fakedata_generator = fake_binary_data_generator(alpha_val=0.5, beta_val=0.5)
    X, y = fakedata_generator.generate_data(5000)
    ### drop half the outcome 1 prevalence
    class_1_index = (y==1)
    class_1_samples = np.where(class_1_index)[0]
    drop_indices = np.random.choice(class_1_samples, size=int(len(class_1_samples)/2),__
     →replace=False)
    mask = np.ones(len(y), dtype=bool)
    mask[drop_indices] = False
    y = y[mask]
    X = X[mask]
    optimal_prevalence.adjusted_p = find_optimal_prevalence(y, X, class_to_calculate=1)
    print("Dataset prevalence: ", np.mean(y))
    print("Derived prevalence: ", optimal_prevalence)
    Dataset prevalence: 0.3300531914893617
    Derived prevalence: 0.49863799264980607
```

The function return both the derived prevalence and the adjusted probability. We can also use the derived prevalence adjustment factor to perform the adjustment mannually.

```
[2]: ### Prevalence Adjustment
from calzone.metrics import lowess_regression_analysis
proba_adjust = apply_prevalence_adjustment(optimal_prevalence, y, X, class_to_
```

```
    →calculate=1)
print('Loess ICI before prevalence adjustment: ', lowess_regression_analysis(y, X, class_
    →to_calculate=1)[0])
print('Loess ICI after prevalence adjustment: ', lowess_regression_analysis(y, proba_
    →adjust, class_to_calculate=1)[0])

Loess ICI before prevalence adjustment: 0.07961758926734244
Loess ICI after prevalence adjustment: 0.008745511902314453
```

calzone also provides a argument to perform prevalence adjustment directly from the CalibrationMetrics class.

```
[4]: for key in before_prevalence.keys():
        print(key)
        print('before adjustment:',before_prevalence[key],', after adjustment:',after_
     →prevalence[key])
    ECE-H topclass
    before adjustment: 0.014081013182402267, after adjustment: 0.010355911839501922
    ECE-H
    before adjustment: 0.0841517729106883 , after adjustment: 0.013671230516636386
    COX coef
    before adjustment: 0.9400481147756811 , after adjustment: 0.9400481147756811
    COX intercept
    before adjustment: -0.6897839569176842 , after adjustment: -0.029403495083063648
    COX coef lowerci
    before adjustment: 0.8754203499121679 , after adjustment: 0.8754203499121678
    COX coef upperci
    before adjustment: 1.0046758796391944, after adjustment: 1.0046758796391944
    COX intercept lowerci
    before adjustment: -0.7837388214288888 , after adjustment: -0.12775157222121533
    COX intercept upperci
    before adjustment: -0.5958290924064796 , after adjustment: 0.06894458205508802
    COX ICI
    before adjustment: 0.0841517733462589 , after adjustment: 0.007508966220374058
    Loess ICI
    before adjustment: 0.07961758926734244 , after adjustment: 0.008745511902314453
```

9.2 Prevalence adjustment and constant shift in logit of class-ofinterest

In the section, we will prove that the prevalence shift is equivalent to a constant shift in logit of class-of-interest. In other words, prevalence adjustment can be done by addint a constant to the logit of class-of-interest. For the calibrated case, the likelihood ratio of the two classes is:

$$LR(p) = \frac{\frac{e^{x_2}}{e^{x_1} + e^{x_2}}}{\frac{e^{x_1}}{e^{x_1} + e^{x_2}}} \cdot \frac{1 - \eta}{\eta} = e^{x_2 - x_1} \cdot \frac{1 - \eta}{\eta}$$

Assumer we add a constant c to the logit of class-of-interest (x_2 here), the likelihood ratio becomes:

$$LR'(p) = e^{x_2 - x_1 + c} \cdot \frac{1 - \eta}{\eta}$$

And the posterior probability becomes:

$$P'(D=1|\hat{p}=p) = \frac{\eta L R'(p)}{\eta L R'(p) + (1-\eta)} = \frac{\eta L R(p) \cdot e^c}{\eta L R(p) \cdot e^c + (1-\eta)}$$

Which is equivalent to the posterior probability after prevalence adjustment:

$$\frac{\eta' LR(p)}{\eta' LR(p) + (1 - \eta')}$$

By setting

$$\eta' = \frac{1}{1 + e^a \left(\frac{1 - \eta}{\eta}\right)}$$

Therefore, prevalence adjustment is equivalent to a constant shift in logit of class-of-interest.

9.3 References

Chen, W., Sahiner, B., Samuelson, F., Pezeshk, A., & Petrick, N. (2018). Calibration of medical diagnostic classifier scores to the probability of disease. Statistical Methods in Medical Research, 27(5), 1394–1409. https://doi.org/10.1177/0962280216661371

Gu, W., & Pepe, M. S. (2011). Estimating the diagnostic likelihood ratio of a continuous marker. Biostatistics, 12(1), 87–101. https://doi.org/10.1093/biostatistics/kxq045

Tian, J., Liu, Y.-C., Glaser, N., Hsu, Y.-C., & Kira, Z. (2020). Posterior Re-calibration for Imbalanced Datasets (No. arXiv:2010.11820). arXiv. http://arxiv.org/abs/2010.11820

Horsch, K., Giger, M. L., & Metz, C. E. (2008). Prevalence scaling: applications to an intelligent workstation for the diagnosis of breast cancer. Academic radiology, 15(11), 1446–1457. https://doi.org/10.1016/j.acra.2008.04.022

TEN

SUBGROUP ANALYSIS

In many real-world applications, we are not just interested in the calibration of the overall population, but also interested in the calibration for subgroups within the population. calzone provides a simple way to perform subgroup analysis given some data input format. In order to perform subgroup analysis, the input csv file should contain the following columns:

```
proba_0, proba_1, ..., proba_n, subgroup_1, subgroup_2, ..., subgroup_m, label where n \ge 1 and m \ge 1.
```

In this example, we will use the example simulated dataset in the calzone package with only one subgroup field and two subgroups. See quickstart for more details.

```
[]: ### import the packages and read the data
    import numpy as np
    from calzone.utils import data_loader
    from calzone.metrics import CalibrationMetrics
    dataset = data_loader('../../example_data/simulated_data_subgroup.csv')
    print(np.loadtxt('../../example_data/simulated_data_subgroup.csv',dtype=str)[:5])
     →#first 5 lines of the csv files
    print("Whether the dataset has subgroup:",dataset.have_subgroup)
    ### Create the CalibrationMetrics class
    metrics_cal = CalibrationMetrics(class_to_calculate=1)
    ['proba_0,proba_1,subgroup_1,label'
      '0.1444156178040511,0.8555843821959489,A,0'
     '0.8552048445812981,0.1447951554187019,A,O'
     '0.2569696048872897.0.7430303951127103.A.0'
     '0.39931305655530125,0.6006869434446988,A,1']
    Whether the dataset has subgroup: True
```

```
for metric in result:
            print(f"{metric}: {result[metric]}")
subgroup 1
subgroup 1 class A
SpiegelhalterZ score: 0.3763269161877356
SpiegelhalterZ p-value: 0.7066738713391099
ECE-H topclass: 0.009608653731328977
ECE-H: 0.01208775955804901
MCE-H topclass: 0.03926468843081976
MCE-H: 0.04848338618970194
HL-H score: 8.884991559088098
HL-H p-value: 0.35209071874348785
ECE-C topclass: 0.009458033653818828
ECE-C: 0.008733966945443138
MCE-C topclass: 0.020515047600205505
MCE-C: 0.02324031223486256
HL-C score: 3.694947603203135
HL-C p-value: 0.8835446575708198
COX coef: 0.9942499557748269
COX intercept: -0.04497652296600376
COX coef lowerci: 0.9372902801721911
COX coef upperci: 1.0512096313774626
COX intercept lowerci: -0.12348577118577644
COX intercept upperci: 0.03353272525376893
COX ICI: 0.005610391483826338
Loess ICI: 0.00558856942568957
subgroup 1 class B
SpiegelhalterZ score: 27.93575342117766
SpiegelhalterZ p-value: 0.0
ECE-H topclass: 0.07658928982434714
ECE-H: 0.0765892898243467
MCE-H topclass: 0.1327565894838103
MCE-H: 0.16250572519432438
HL-H score: 910.4385762101924
HL-H p-value: 0.0
ECE-C topclass: 0.07429481165606829
ECE-C: 0.07479369479609524
MCE-C topclass: 0.14090872416947742
MCE-C: 0.14045600565696226
HL-C score: 2246.1714434139853
HL-C p-value: 0.0
COX coef: 0.5071793536874274
COX intercept: 0.00037947714112375366
COX coef lowerci: 0.47838663128188996
COX coef upperci: 0.5359720760929648
COX intercept lowerci: -0.07796623141885761
COX intercept upperci: 0.07872518570110512
COX ICI: 0.07746407648179383
Loess ICI: 0.06991428582761099
```

[3]: ### An alernative way to do the same thing is through command line interface

```
%run ../../cal_metrics.py \
--csv_file '../../example_data/simulated_data_subgroup.csv' \
--metrics all \
--class_to_calculate 1 \
--num_bins 10 \
--verbose
Metrics:
SpiegelhalterZ score: 18.327
SpiegelhalterZ p-value: 0.
ECE-H topclass: 0.042
ECE-H: 0.042
MCE-H topclass: 0.055
MCE-H: 0.063
HL-H score: 429.732
HL-H p-value: 0.
ECE-C topclass: 0.042
ECE-C: 0.038
MCE-C topclass: 0.065
MCE-C: 0.064
HL-C score: 1138.842
HL-C p-value: 0.
COX coef: 0.668
COX intercept: -0.02
COX coef lowerci: 0.641
COX coef upperci: 0.696
COX intercept lowerci: -0.074
COX intercept upperci: 0.034
COX ICI: 0.049
Loess ICI: 0.037
Metrics for subgroup subgroup_1_group_A:
SpiegelhalterZ score: 0.376
SpiegelhalterZ p-value: 0.707
ECE-H topclass: 0.01
ECE-H: 0.012
MCE-H topclass: 0.039
MCE-H: 0.048
HL-H score: 8.885
HL-H p-value: 0.352
ECE-C topclass: 0.009
ECE-C: 0.009
MCE-C topclass: 0.021
MCE-C: 0.023
HL-C score: 3.695
HL-C p-value: 0.884
COX coef: 0.994
COX intercept: -0.045
COX coef lowerci: 0.937
COX coef upperci: 1.051
COX intercept lowerci: -0.123
COX intercept upperci: 0.034
```

```
COX ICI: 0.006
Loess ICI: 0.006
Metrics for subgroup subgroup_1_group_B:
SpiegelhalterZ score: 27.936
SpiegelhalterZ p-value: 0.
ECE-H topclass: 0.077
ECE-H: 0.077
MCE-H topclass: 0.133
MCE-H: 0.163
HL-H score: 910.439
HL-H p-value: 0.
ECE-C topclass: 0.074
ECE-C: 0.075
MCE-C topclass: 0.141
MCE-C: 0.140
HL-C score: 2246.171
HL-C p-value: 0.
COX coef: 0.507
COX intercept: 0.000
COX coef lowerci: 0.478
COX coef upperci: 0.536
COX intercept lowerci: -0.078
COX intercept upperci: 0.079
COX ICI: 0.077
Loess ICI: 0.07
```

[]:

ELEVEN

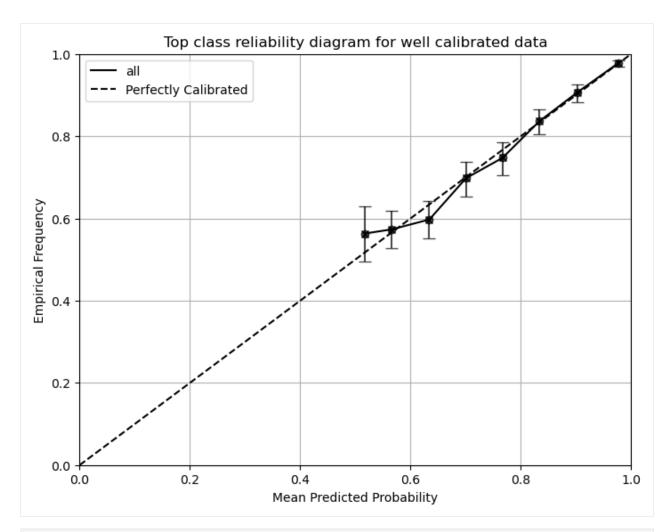
MULTICLASS EXTENSION

In the previous notebooks, most metrics only consider binary classification except ECE_{top} and MCE_{top} . To handle multi-class classification, the most common way is to convert it to a series of binary classification problems by using 1-vs-rest approach. Alternativly, if the user only care about whether the predicted class of each sample is having the correct predicted probability, the user transform the data to a top-class binary classification problem. In top-class problem, the class 0 probability is the $1-p_{top}$ and class 1 probability is p_{top} and the label is whether true label is equal to the top-class prediction. In calzone, user can transform the data using data_loader.transform_topclass() function. Most metrics will still works but the interpretation of the metrics will be different.

```
[3]: from calzone.utils import data_loader,reliability_diagram
    from calzone.vis import plot_reliability_diagram
    from calzone.metrics import CalibrationMetrics

dataset = data_loader('../../example_data/simulated_welldata.csv')
    cal_metrics = CalibrationMetrics(class_to_calculate=1)
    ### Transform to top-class
    transformed_data = dataset.transform_topclass()
[4]: ### Now the transformed data become a top-class problem when we specify the class-of-
```

```
[4]: ### Now the transformed data become a top-class problem when we specify the class-of-
interest to be 1
reliability,confindence,bin_edges,bin_counts = reliability_diagram(transformed_data.
labels, transformed_data.probs, num_bins=15,class_to_plot=1)
# Plot the reliability diagram
plot_reliability_diagram(reliability,confindence,bin_counts,error_bar=True,title='Top_class reliability diagram for well calibrated data')
```



```
[5]: cal_metrics.calculate_metrics(transformed_data.labels, transformed_data.probs,metrics=
     → 'all')
[5]: {'SpiegelhalterZ score': 0.37632691618773545,
      'SpiegelhalterZ p-value': 0.7066738713391101,
      'ECE-H topclass': 0.009608653731328977,
      'ECE-H': 0.009608653731329372,
      'MCE-H topclass': 0.03926468843081976,
      'MCE-H': 0.03926468843081932,
      'HL-H score': 6.029652031234071,
      'HL-H p-value': 0.11017646753618837,
      'ECE-C topclass': 0.009458033653818828,
      'ECE-C': 0.009458033653818974,
      'MCE-C topclass': 0.020515047600205505,
      'MCE-C': 0.020515047600205394,
      'HL-C score': 5.0830845954424,
      'HL-C p-value': 0.7486601568004448,
      'COX coef': 0.9990721119864874,
      'COX intercept': -0.009313116424641145,
      'COX coef lowerci': 0.9097229915817346,
      'COX coef upperci': 1.0884212323912401,
```

```
'COX intercept lowerci': -0.1322110153978738,
'COX intercept upperci': 0.1135847825485915,
'COX ICI': 0.0012819521292332472,
'Loess ICI': 0.0038637011857438034}
```

You can see the top-class ECE is the same as regular ECE.

TWELVE

RUNNING THE GUI

GUI is currently under development and not all features from the command line are supported. To run the GUI, you need to install nicegui using the following command:

[]: pip install nicegui

Then simply run the following command on calzone directory:

[]: python GUI_cal_metric.py

VALIDATING METRICS WITH EXTERNAL PACKAGES

In this notebook, we want to validate the metrics we have implemented in the calzone package using other packages or programs.

13.1 Reliability diagram

We will use scikit-learn's calibration_curve function to calculate the reliability diagram

```
[2]: from calzone.utils import reliability_diagram,data_loader
    from calzone.metrics import calculate_ece_mce,hosmer_lemeshow_test
    import numpy as np
    from sklearn.calibration import calibration_curve
    ### loading the data
    wellcal_dataloader = data_loader(data_path="../../example_data/simulated_welldata.csv
     '')
    ###scikit-learn implementation
    scikit_reliability_H,scikit_confidence_H = calibration_curve(wellcal_dataloader.labels,
     wellcal_dataloader.probs[:,1],n_bins=15,strategy='uniform',pos_label=1)
    scikit_reliability_C,scikit_confidence_C = calibration_curve(wellcal_dataloader.labels,
     →wellcal_dataloader.probs[:,1],n_bins=15,strategy='quantile',pos_label=1)
    ### calzone implementation
    calzone_reliability_H,calzone_confindence_H,bin_edge_H,bin_count_H = reliability_
     →diagram(wellcal_dataloader.labels,wellcal_dataloader.probs,num_bins=15, class_to_
     →plot=1, is_equal_freq=False)
    calzone_reliability_C,calzone_confindence_C,bin_edge_C,bin_count_C = reliability_
     →diagram(wellcal_dataloader.labels,wellcal_dataloader.probs,num_bins=15, class_to_
     →plot=1, is_equal_freq=True)
    ###showing the difference between the two implementations
    print("Difference for equal-width binning:")
    print("Reliability difference:", np.round(np.abs(scikit_reliability_H - calzone_
     →reliability_H), 4))
    print("Confidence difference:", np.round(np.abs(scikit_confidence_H - calzone_

→confindence_H), 4))
    print("\nDifference for equal-count binning:")
    print("Reliability difference:", np.round(np.abs(scikit_reliability_C - calzone_
     →reliability_C), 4))
```

We can see that the calzone package return the same Reliability diagram as the sckit-learn package. We will move to test the expected calibration error.

13.2 Expected calibration error and Z test

We will use mapie package to validate some of the metrics in calzone. Description of mapie can be found here.

```
[4]: ### compare MAPIE and calzone equal-count binning

print("MAPIE topclass ECE-C:",top_label_ece(wellcal_dataloader.labels,wellcal_dataloader.

→probs,num_bins = 15,split_strategy='quantile'))

print("calzone topclass ECE-C:",calculate_ece_mce(calzone_reliability_topclass_C,calzone_

→confindence_topclass_C,bin_count_topclass_C)[0])

MAPIE topclass ECE-C: 0.016227424850494457

calzone topclass ECE-C: 0.016263864264387196
```

We can see that both package return a very similar result for ECE. We will also move to validate spiegelhalter's z test. We found out that the mapie package incorrectly calculates the p-value by using a one-sided test. Therefore, we will only compare the test statistic but not the p-value.

For other metrics, we could not find any external packages that could be used to cross validate our implementation. However, we perform some simulations to test the type 1 error rate for cox slope and intercept and both return sensible values. For the ICI, we compare it with empirical ECE and found they are very similar.

13.3 HL test

Lastly, we will use a quick r code to calculate the HL test test statistic to make sure it is correct.

```
[1]: ### This section is r code
    library(ResourceSelection)
     # Read the CSV file
    data <- read.csv("../../example_data/simulated_welldata.csv")</pre>
    predicted_prob <- data[,2] # First column with predicted probabilities</pre>
                               # Third column with actual labels
    labels <- data[,3]</pre>
     # Perform Hosmer-Lemeshow test
    hltest <- function(observed, predicted) {</pre>
      hl_test <- hoslem.test(observed, predicted,g=15)</pre>
       cat("Hosmer-Lemeshow Test Results:\n")
       cat("Chi-square statistic:", round(hl_test$statistic, 4), "\n")
      cat("Degrees of freedom:", hl_test$parameter, "\n")
       cat("p-value:", round(hl_test$p.value, 4), "\n")
       return(hl_test)
    }
    result <- hltest(labels, predicted_prob)</pre>
    ResourceSelection 0.3-6
                                       2023-06-27
    Hosmer-Lemeshow Test Results:
    Chi-square statistic: 6.299
    Degrees of freedom: 13
    p-value: 0.9346
```

13.3. HL test 65

We see that the test statistics are the same. The R package doesn't allow user input degree of freedom so the p-value is different as expected.

13.4 reference

Taquet, V., Blot, V., Morzadec, T., Lacombe, L., & Brunel, N. (2022). MAPIE: an open-source library for distribution-free uncertainty quantification. arXiv preprint arXiv:2207.12274.

Pedregosa, F., Varoquaux, Ga"el, Gramfort, A., Michel, V., Thirion, B., Grisel, O., ... others. (2011). Scikit-learn: Machine learning in Python. Journal of Machine Learning Research, 12(Oct), 2825–2830.

Lele, S. R., Keim, J. L., & Solymos, P. (2017). Resource selection (probability) functions for use-availability data. Package 'ResourceSelection', Version 0.3-2.

FOURTEEN

CALZONE

14.1 calzone package

14.1.1 Submodules

14.1.2 calzone.metrics module

Metrics calculation functions for the Calibration Measure package.

class calzone.metrics.CalibrationMetrics(class_to_calculate=1, num_bins=10)

Bases: object

A class for calculating calibration metrics for classification models.

__init__(class_to_calculate=1, num_bins=10)

Initialize the CalibrationMetrics class.

Parameters

- class_to_calculate (int, optional) The class index to calculate the metrics for. Defaults to 1.
- num_bins (int, optional) Number of bins to use for the ECE/MCE/HL calculations. Defaults to 10.

bootstrap(*y_true*, *y_proba*, *metrics*, *perform_pervalance_adjustment=False*, *n_samples=1000*, **kwargs)
Run bootstrap and return a numpy structured array with correct field names.

This function performs bootstrap resampling to estimate the distribution of calibration metrics. It generates multiple samples with replacement from the input data and calculates the specified metrics for each sample.

Parameters

- y_true (array-like) True labels.
- **y_proba** (*array-like*) Predicted probabilities.
- metrics (list of str) List of metric names to calculate.
- **perform_pervalance_adjustment** (*bool*, *optional*) Whether to perform prevalence adjustment for each bootstrap sample. Defaults to False.
- **n_samples** (*int*, *optional*) Number of bootstrap samples to generate. Defaults to 1000.
- **kwargs Additional keyword arguments to pass to the metric calculation functions.

Returns

numpy.ndarray – A structured array containing the bootstrapped metrics. Each field in the array corresponds to a metric, and each row represents a bootstrap sample.

Calculate the specified calibration metrics.

This function computes various calibration metrics for binary classification models. It supports multiple metrics and can perform prevalence adjustment if needed.

List of available metrics:

- SpiegelhalterZ: Spiegelhalter's Z-test for calibration
- ECE-H: Expected Calibration Error with equal-space binning
- MCE-H: Maximum Calibration Error with equal-space binning
- HL-H: Hosmer-Lemeshow test with equal-space binning
- ECE-C: Expected Calibration Error with equal-count binning
- MCE-C: Maximum Calibration Error with equal-count binning
- HL-C: Hosmer-Lemeshow test with equal-count binning
- COX: Cox regression analysis
- Loess: Locally Estimated Scatterplot Smoothing regression analysis

Parameters

- **y_true** (*numpy.ndarray*) True labels.
- **y_proba** (*numpy.ndarray*) Predicted probabilities.
- **metrics** (*list of str or 'all'*) List of metric names to calculate. If 'all', calculates all available metrics.
- **perform_pervalance_adjustment** (*bool*, *optional*) Whether to perform prevalence adjustment. Defaults to False.
- **return_numpy** (bool, optional) Whether to return the results as a numpy array. Defaults to False.
- **kwargs Additional keyword arguments to pass to the metric calculation functions.

Returns

dict or numpy.ndarray – A dictionary containing the calculated metrics, or a numpy array if return_numpy is True.

optimal_prevalence_adjustment(y_true, y_proba)

Perform optimal prevalence adjustment and return adjusted probabilities.

This function finds the optimal prevalence value that minimizes the difference between the predicted and actual positive rates, and then adjusts the input probabilities accordingly.

Parameters

- y_true (array-like) True labels.
- **y_proba** (*array-like*) Predicted probabilities.

Returns

- **optimal_prevalence** (*float*) Optimal prevalence value.
- adjusted_proba (array-like) Adjusted probabilities. First column is the adjusted probabilities for the other class, second column is the adjusted probabilities for the class of interest.

calzone.metrics.cal_ICI(y_adjust, y_proba)

Calculate the Integrated Calibration Index (ICI) for given adjusted probabilities.

Parameters

- **y_adjust** (*array-like*) Adjusted probabilities. Shape (n_samples,).
- **y_proba** (*array-like*) Original predicted probabilities. Shape (n_samples,).

Returns

float – The Integrated Calibration Index (ICI) value.



1 Note

The ICI is calculated by calculating the mean absolute difference between predicted probabilities and the adjusted probabilities.

calzone.metrics.cal_ICI_cox(coef, intercept, y_proba, class_to_calculate=1, epsilon=1e-07, **kwargs) Calculate the Integrated Calibration Index (ICI) for a given Cox regression model.

The ICI measures the average absolute difference between the predicted probabilities and the probabilities transformed by the fitted Cox regression model.

Parameters

- **coef** (*float*) The coefficient (slope) from the Cox regression.
- **intercept** (*float*) The intercept from the Cox regression.
- y_proba (array-like) Predicted probabilities. Should be of shape (n_samples, n_classes).
- class_to_calculate (int) The class to calculate the ICI for in multi-class problems. Default is 1.
- epsilon (float) Small value to avoid numerical instability when clipping probabilities. Default is 1e-7.

Returns

ICI (*float*) – The Integrated Calibration Index.



- Lower ICI values indicate better calibration.
- The function applies the inverse logit transformation to the predicted probabilities using the coefficients from the Cox regression.

calzone.metrics.cal_ICI_func(func, y_proba, class_to_calculate=1)

Calculate the Integrated Calibration Index (ICI) for a given calibration function.

Parameters

- **func** (*callable*) The calibration function to evaluate.
- y_proba (array-like) Predicted probabilities for each class. Shape (n_samples, n classes).
- class_to_calculate (int, optional) The class index to calculate the ICI for. Defaults to 1.

float – The Integrated Calibration Index (ICI) value.



Note

The ICI is calculated by calculating the mean absolute difference between predicted probabilities and the calibration function evaluated at predicted probabilities.

calzone.metrics.calculate_ece_mce(reliability, confindence, bin_counts)

Calculate Expected Calibration Error (ECE) and Maximum Calibration Error (MCE).

These metrics assess the calibration of a classification model by comparing predicted probabilities to observed frequencies.

Parameters

- **reliability** (*array-like*) Array of observed frequencies for each bin.
- **confindence** (*array-like*) Array of predicted probabilities for each bin.
- bin_counts (array-like) Array of sample counts in each bin.

Returns

- **ece** (*float*) The Expected Calibration Error.
- **mce** (*float*) The Maximum Calibration Error.

1 Note

- ECE is a weighted average of the absolute differences between confidence and reliability.
- MCE is the maximum absolute difference between confidence and reliability across all bins.
- Lower values of both ECE and MCE indicate better calibration.

calzone.metrics.cox_regression_analysis(y_true, y_proba, epsilon=1e-07, class_to_calculate=1, print_results=False, fix_intercept=False, fix_slope=False, **kwargs)

Perform Cox regression analysis for classification calibration.

This function fits a logistic regression model to the logit of predicted probabilities to assess the calibration of classification predictions.

Parameters

- **y_true** (array-like) True binary labels. If multi-class, will be converted to binary.
- **y_proba** (array-like) Predicted probabilities. Should be of shape (n_samples, n_classes).
- **epsilon** (*float*) Small value to avoid log(0) errors. Default is 1e-7.

- **class_to_calculate** (*int*) The class to treat as the positive class in binary classification. Default is 1.
- **print_results** (*boo1*) If True, prints the summary of the logistic regression results. Default is False.
- **fix_intercept** (*bool*) If True, fixes the intercept to 0. Can't be used with fix_slope. Default is False.
- fix_slope (bool) If True, fixes the coefficient to 1. Can't be used with fix_intercept.
 Default is False.

- **coef** (*float*) The coefficient (slope) of the Cox regression.
- **intercept** (*float*) The intercept of the Cox regression.
- **coef_ci** (*tuple*) The confidence interval for the coefficient.
- **intercept_ci** (*tuple*) The confidence interval for the intercept.

1 Note

- A well-calibrated model should have a coefficient close to 1 and an intercept close to 0.
- The function clips probabilities to avoid numerical instability.
- For multi-class problems, the function converts the problem to binary classification based on the specified class_to_calculate.

calzone.metrics.get_CI(result, alpha=0.05)

Calculate confidence intervals for each field in the result.

Parameters

- **result** (*numpy.ndarray*) Structured array containing the results for which to calculate confidence intervals.
- **alpha** (*float*, *optional*) The significance level for the confidence interval calculation. Defaults to 0.05.

Returns

dict – A dictionary where keys are field names from the input array and values are tuples containing the lower and upper bounds of the confidence interval.

Note

This function calculates percentile-based confidence intervals for each field in the input structured array. It's useful for bootstrap or Monte Carlo simulations.

calzone.metrics.hosmer_lemeshow_test(reliability, confidence, bin_count, df=None, **kwargs)

Compute the Hosmer-Lemeshow test for goodness of fit.

This test is used to assess the calibration of binary classification models with full probability outputs. It compares observed and expected frequencies of events in groups of the data.

Parameters

• **reliability** (*array-like*) – Observed proportion of positive samples in each bin.

- **confidence** (*array-like*) Predicted probabilities for each bin.
- bin_count (array-like) Number of samples in each bin.
- df (int, optional) Degrees of freedom for the test. Defaults is nbins 2.

- **chi_squared** (*float*) The chi-squared statistic of the Hosmer-Lemeshow test.
- **p_value** (*float*) The p-value associated with the chi-squared statistic.
- **df** (*int*) The degrees of freedom for the test.

1 Note

- The Hosmer-Lemeshow test is widely used for assessing calibration in probabilitiest models.
- A small p-value (typically < 0.05) suggests that the model is a poor fit to the data.
- This test can be sensitive to the number of groups and sample size.
- It is recommended to use the Hosmer-Lemeshow test in conjunction with other metrics.

calzone.metrics.logit_func(coef, intercept)

Create a logistic function with given coefficient and intercept.

Parameters

- **coef** (*float*) The coefficient (slope) of the logistic function.
- **intercept** (*float*) The intercept of the logistic function.

Returns

callable – A function that takes an input x and returns the logistic function value.

1 Note

The returned function applies the logistic transformation: $f(x) = 1 / (1 + \exp(-(\cos f * \log(x / (1 - x)) + \operatorname{intercept})))$

calzone.metrics.lowess_regression_analysis(y_true, y_proba, epsilon=1e-07, class_to_calculate=1, span=0.5, delta=0.001, it=0, **kwargs)

Perform Lowess regression analysis for classification calibration.

This function applies Locally Weighted Scatterplot Smoothing (LOWESS) to assess the calibration of classification predictions.

Parameters

72

- **y_true** (*array-like*) True binary labels. If multi-class, will be converted to binary.
- **y_proba** (array-like) Predicted probabilities. Should be of shape (n_samples, n classes).
- **epsilon** (*float*, *optional*) Small value to avoid numerical instability when clipping probabilities. Defaults to 1e-10.
- **class_to_calculate** (*int*, *optional*) The class to treat as the positive class in binary classification. Defaults to 1.

Chapter 14. calzone

- **span** (*float*, *optional*) The fraction of the data used when estimating each y-value. Defaults to 0.5.
- **delta** (*float*, *optional*) Distance within which to use linear-interpolation instead of weighted regression. Defaults to 0.001.
- it (int, optional) The number of residual-based reweightings to perform. Defaults to 0.

- ICI (*float*) The Integrated Calibration Index.
- **sorted_proba** (*array-like*) Sorted predicted probabilities.
- **smoothed_proba** (*array-like*) Corresponding LOWESS-smoothed actual probabilities.

1 Note

- The function clips probabilities to avoid numerical instability.
- For multi-class problems, the function converts the problem to binary classification based on the specified class_to_calculate.
- The Integrated Calibration Index (ICI) provides a measure of calibration error, with lower values indicating better calibration.

calzone.metrics.spiegelhalter_z_test(y_true, y_proba, class_to_calculate=1)

Perform Spiegelhalter's Z-test for calibration of probabilistic predictions.

This test assesses whether predicted probabilities are well-calibrated by comparing them to observed outcomes.

Parameters

- **y_true** (*array-like*) True labels of the samples.
- **y_proba** (*array-like*) Predicted probabilities for each class. Shape should be (n_samples, n_classes).
- **class_to_calculate** (*int*) Index of the class to calculate the test for. Default is 1.

Returns

- **z** score (*float*) The z-score of the Spiegelhalter's Z-test.
- **p_value** (*float*) The p-value associated with the z-score.

1 Note

- This test is used to assess the calibration of a classification model.
- A small p-value (typically < 0.05) suggests that the model is poorly calibrated.
- The test assumes that predictions are independent and identically distributed.

14.1.3 calzone.utils module

Uitlity functions for the Calibration Measure package.

calzone.utils.apply_prevalence_adjustment(adjusted_prevalence, y_true, y_proba, class_to_calculate=1) Apply the prevalence adjustment method.

Parameters

- adjusted_prevalence (float) The adjusted prevalence to test.
- **y_true** (array-like) True labels.
- **y_proba** (*array-like*) Predicted probabilities.
- $class_to_calculate(int)$ The class index to adjust. Default is 1.

Returns

numpy.ndarray - Adjusted probabilities.

```
class calzone.utils.data_loader(data_path)
```

Bases: object

A class for loading and preprocessing data from a CSV file.

This class handles various data formats, including those with or without subgroup information and headers.

data_path

Path to the CSV file containing the data.

```
Type
```

str

Header

Array of column headers from the CSV file.

```
Type
```

numpy.ndarray

subgroups

List of subgroup column names, if present.

```
Type
```

list

subgroup_indices

List of indices for subgroup columns, if present.

```
Type
list
```

have_subgroup

Flag indicating whether subgroup information is present.

```
Type
bool
```

data

Raw data loaded from the CSV file.

Type

numpy.ndarray

probs

Probability values extracted from the data.

Type

numpy.ndarray

labels

Label values extracted from the data.

Type

numpy.ndarray

subgroups_class

List of unique subgroup classes for each subgroup, if present.

Type

list

subgroups_index

List of indices for each subgroup class, if present.

Type

list

```
__init__(self, data_path)
```

Initializes the data_loader object and loads data from a CSV file.

transform_topclass(self)

Transforms the data to top class binary problem.

```
__init__(data_path)
```

Initializes the data_loader object and loads data from the specified file.

Parameters

data_path (*str*) – Path to the CSV file containing the data.

The method performs the following steps: 1. Loads the header from the CSV file. 2. Checks for the presence of subgroup information. 3. Loads the data based on the presence or absence of subgroup information. 4. Extracts probability values and labels from the loaded data. 5. If subgroups are present, extracts subgroup classes and their indices.

Note: - If there is a header, it must be in the format: proba_0,proba_1,...,subgroup_1(optional),subgroup_2(optional),...,label - If there is no header, the columns must be in the order of proba_0,proba_1,...,label

transform_topclass()

Transforms the data to top class binary problem

Returns

data_loader - A new data_loader object with transformed data

class calzone.utils.fake_binary_data_generator(alpha_val, beta_val)

Bases: object

A class for generating fake binary data and applying miscalibration.

This class provides methods to generate binary classification data and apply different types of miscalibration to the probabilities.

alpha_val

Alpha parameter for the beta distribution.

Type

float

beta_val

Beta parameter for the beta distribution.

Type

float

__init__(alpha_val, beta_val)

Initialize the fake binary data generator.

Parameters

- **alpha_val** (*float*) Alpha parameter for the beta distribution.
- **beta_val** (*float*) Beta parameter for the beta distribution.

abraitary_miscal(logits, miscal_function)

Apply arbitrary miscalibration to the input logits.

This function allows for the application of any custom miscalibration function to the input logits.

Parameters

- **logits** (numpy.ndarray) Input logits of shape (n_samples, 2).
- **miscal_function** (*callable*) Function to apply miscalibration to the logits.

Returns

numpy.ndarray – Miscalibrated probabilities of shape (n_samples, 2).

generate_data(sample_size)

Generate fake binary classification data.

Parameters

sample_size (*int*) – Number of samples to generate.

Returns

- **X** (numpy.ndarray) Array of shape (sample_size, 2) containing probabilities for each class
- y_true (numpy.ndarray) Array of shape (sample_size,) containing true binary labels.

linear_miscal(X, miscal_scale)

Apply linear miscalibration to the input probabilities.

This function transforms the input probabilities to logits, applies a linear scaling, and then converts back to probabilities.

Parameters

- **X** (numpy.ndarray) Input probabilities of shape (n samples, 2).
- **miscal_scale** (*float*) Scale factor for miscalibration.

Returns

numpy.ndarray – Miscalibrated probabilities of shape (n_samples, 2).

calzone.utils.find_optimal_prevalence(y_true, y_proba, class_to_calculate=1, epsilon=1e-07)

Find the optimal adjustment prevalence using scipy.optimize.

Parameters

76

• y_true (array-like) - True labels.

Chapter 14. calzone

- y_proba (array-like) Predicted probabilities.
- class_to_calculate (int) The class index to optimize for. Default is 1.
- **epsilon** (*float*) Small value to avoid numerical instability. Default is 1e-7.

- **optimal prevalence** (*float*) The optimal prevalence.
- adjusted_probabilities (numpy.ndarray) The adjusted probabilities using the optimal prevalence.

calzone.utils.loss(adjusted_prevalence, y_true, y_proba, class_to_calculate=1)

Calculate the loss function for prevalence adjustment.

Parameters

- adjusted_prevalence (float) The adjusted prevalence.
- y_true (array-like) True labels.
- y_proba (array-like) Predicted probabilities.
- **class_to_calculate** (*int*) The class index to calculate loss for. Default is 1.

Returns

float – Calculated loss value.

calzone.utils.make_roc_curve(y_true, y_proba, class_to_plot=None)

Compute the Receiver Operating Characteristic (ROC) curve for binary or multiclass classification.

Parameters

- **y_true** (*array-like*) True labels of the data. Shape (n_samples,).
- **y_proba** (*array-like*) Predicted probabilities of the positive class. Shape (n_samples, n classes).
- class_to_plot (int, optional) The class to plot the ROC curve for. If None, plots the ROC curve for each class. Default is None.

Returns

- **fpr** (*array*) False Positive Rate for the selected class or each class. Shape (n_points,).
- tpr (array) True Positive Rate for the selected class or each class. Shape (n_points,).
- **roc_auc** (*float or array*) Area Under the ROC Curve (AUC) for the selected class or each class. If class_to_plot is not None, returns a float. If class_to_plot is None, returns an array of shape (n_classes,).

1 Note

- The input arrays y_true and y_proba must have the same number of samples.
- The input array y_proba must have probabilities for each class in a multiclass problem.
- The input array y_proba must not contain any NaN values.

Example

```
>>> y_true = [0, 1, 1, 0, 1]
>>> y_proba = [[0.2, 0.8], [0.6, 0.4], [0.3, 0.7], [0.9, 0.1], [0.4, 0.6]]
>>> fpr, tpr, roc_auc = roc_curve(y_true, y_proba, class_to_plot=1)
```

calzone.utils.reliability_diagram(y_true, y_proba, num_bins=10, class_to_plot=None, is_equal_freq=False, save_path=None)

Compute the reliability diagram for a binary or multi-class classification model.

Parameters

- **y_true** (*array-like*) True labels of the samples. Can be a binary array or a one-hot encoded array.
- **y_proba** (array-like) Predicted probabilities for each class. Shape should be (n_samples, n_classes).
- num_bins (int) Number of bins to divide the predicted probabilities into. Default is 10.
- **class_to_plot** (*int or None*) Index of the class to plot the reliability diagram for. If None, the diagram will be computed for all classes. Default is None.
- **is_equal_freq** (*bool*) If True, the bins will be equally frequent. If False, the bins will be equally spaced. Default is False.

Returns

- **reliabilities** (*array-like*) Array of accuracies for each bin. Shape depends on the value of class_to_plot.
- **confidences** (*array-like*) Array of average confidences for each bin. Shape depends on the value of class to plot.
- **bin_edges** (*array-like*) Array of bin edges.
- **bin_counts** (*array-like*) Array of counts for each bin.

1 Note

- The reliability diagram is a graphical tool to assess the calibration of a classification model. It plots the average predicted probability against the observed accuracy for each bin of predicted probabilities.
- If y_true is a binary array, it will be converted to a one-hot encoded array internally.
- If class_to_plot is not None, the reliability diagram will be computed only for the specified class. Otherwise, it will be computed for all classes.
- The number of bins determines the granularity of the reliability diagram. Higher values result in more bins and a more detailed diagram.

calzone.utils.removing_nan(y_true, y_predict, y_proba)

Remove rows containing NaN values from input arrays.

Parameters

- **y_true** (array-like) True labels.
- y_predict (array-like) Predicted labels.
- y_proba (array-like) Predicted probabilities.

78 Chapter 14. calzone

- y_true (array-like) Cleaned version of y_true with NaN rows removed.
- y_predict (array-like) Cleaned version of y_predict with NaN rows removed.
- y_proba (array-like) Cleaned version of y_proba with NaN rows removed.

calzone.utils.softmax_to_logits(probabilities, epsilon=1e-07)

Convert softmax probabilities to logits.

Parameters

- **probabilities** (array-like) Input probabilities.
- **epsilon** (*float*) Small value to avoid log(0). Default is 1e-7.

Returns

numpy.ndarray - Computed logits.

calzone.utils.transform_topclass(probs, labels)

Transforms the data to top class binary problem

Parameters

- probs (numpy.ndarray) Array of probability values
- labels (numpy.ndarray) Array of label values

Returns

tuple - (transformed_probs, transformed_labels)

14.1.4 calzone.vis module

Visualization functions for the Calibration Measure package.

```
calzone.vis.plot_reliability_diagram(reliabilities, confidences, bin_counts, bin_edges=None, line=True, error_bar=False, z=1.96, title='Reliability Diagram', save_path=None, return_fig=False, custom_colors=None, dni=150)
```

Plot a reliability diagram to visualize the calibration of a model.

Parameters

- **reliabilities** (*array-like*) Empirical frequencies for each bin.
- **confidences** (*array-like*) Mean predicted probabilities for each bin.
- **bin_counts** (*array-like*) Number of samples in each bin.
- **bin_edges** (*array-like*, *optional*) Edges of the bins. If None, assumes equal-spaced bins.
- line (bool, optional) If True, plot lines connecting points. If False, plot as a bar chart. Defaults to True.
- error_bar (bool, optional) If True, add error bars to the plot. Defaults to False.
- **z** (*float*, *optional*) Z-score for calculating Wilson score interval. Defaults to 1.96.
- **title** (*str*, *optional*) Title of the plot. Defaults to 'Reliability Diagram'.
- **save_path** (*str*, *optional*) Path to save the figure. If None, figure is not saved. Defaults to None.

- return_fig (bool, optional) If True, return the figure object. Defaults to False.
- **custom_colors** (*list*, *optional*) List of custom colors for multi-class plots. Defaults to None.
- **dpi** (*int*, *optional*) DPI for saving the figure. Defaults to 150.

matplotlib.figure.Figure, optional – The figure object if return_fig is True.

calzone.vis.plot_roc_curve(fpr, tpr, roc_auc, class_to_plot=None, title='ROC Curve', save_path=None, dpi=150, return_fig=False)

Plots the Receiver Operating Characteristic (ROC) curve.

Parameters

- **fpr** (array-like) False Positive Rate values.
- **tpr** (*array-like*) True Positive Rate values.
- roc_auc (float or array-like) Area Under the ROC Curve (AUC) value(s).
- **class_to_plot** (*int*, *optional*) The class to plot. If None, plots all classes. Defaults to None.
- **title** (*str*, *optional*) Title of the plot. Defaults to 'ROC Curve'.
- **save_path** (*str*, *optional*) Path to save the figure. If None, the figure is not saved. Defaults to None.
- dpi (int, optional) The resolution in dots per inch for saving the figure. Defaults to 150.
- return_fig (bool, optional) If True, returns the figure object instead of displaying it. Defaults to False.

Returns

matplotlib.figure.Figure or None – The figure object if return_fig is True, otherwise None.

This function creates a matplotlib figure showing the ROC curve(s).

14.1.5 Module contents

PYTHON MODULE INDEX

С

calzone, 80
calzone.metrics, 67
calzone.utils, 74
calzone.vis, 79

82 Python Module Index

INDEX

Symbols	D
init() (calzone.metrics.CalibrationMetrics	data (calzone.utils.data_loader attribute), 74
method), 67	data_loader (class in calzone.utils), 74
init() (calzone.utils.data_loader method), 75init() (calzone.utils.fake_binary_data_generator	data_path (calzone.utils.data_loader attribute), 74
method), 76	F
A	<pre>fake_binary_data_generator(class in calzone.utils),</pre>
abraitary_miscal() (cal- zone.utils.fake_binary_data_generator method), 76	find_optimal_prevalence() (in module calzone.utils), 76
alpha_val (calzone.utils.fake_binary_data_generator	G
attribute), 75	generate_data() (cal-
apply_prevalence_adjustment() (in module cal- zone.utils), 74	zone.utils.fake_binary_data_generator method), 76
В	<pre>get_CI() (in module calzone.metrics), 71</pre>
beta_val (calzone.utils.fake_binary_data_generator at-	Н
tribute), 76	have_subgroup (calzone.utils.data_loader attribute), 74
bootstrap() (calzone.metrics.CalibrationMetrics	Header (calzone.utils.data_loader attribute), 74
method), 67	hosmer_lemeshow_test() (in module calzone.metrics),
С	71
cal_ICI() (in module calzone.metrics), 69	L
cal_ICI_cox() (in module calzone.metrics), 69	labels (calzone.utils.data_loader attribute), 75
cal_ICI_func() (in module calzone.metrics), 69	linear_miscal() (cal-
calculate_ece_mce() (in module calzone.metrics), 70	zone.utils.fake_binary_data_generator
calculate_metrics() (cal-	method), 76
zone.metrics.CalibrationMetrics method),	<pre>logit_func() (in module calzone.metrics), 72</pre>
68	loss() (in module calzone.utils), 77
CalibrationMetrics (class in calzone.metrics), 67 calzone	lowess_regression_analysis() (in module cal-
module, 80	zone.metrics), 72
calzone.metrics	M
module, 67	make_roc_curve() (in module calzone.utils), 77
calzone.utils	module
module, 74	calzone, 80
calzone.vis	calzone.metrics, 67
module, 79	calzone.utils,74
<pre>cox_regression_analysis() (in module cal- zone.metrics), 70</pre>	calzone.vis,79

```
0
optimal_prevalence_adjustment()
                                                (cal-
         zone.metrics.CalibrationMetrics
                                            method),
Ρ
plot_reliability_diagram()
                                      module
                                                cal-
                                 (in
         zone.vis), 79
plot_roc_curve() (in module calzone.vis), 80
probs (calzone.utils.data_loader attribute), 74
R
reliability_diagram() (in module calzone.utils), 78
removing_nan() (in module calzone.utils), 78
S
softmax_to_logits() (in module calzone.utils), 79
spiegelhalter_z_test() (in module calzone.metrics),
         73
subgroup_indices
                      (calzone.utils.data_loader
         tribute), 74
subgroups (calzone.utils.data_loader attribute), 74
subgroups_class (calzone.utils.data_loader attribute),
         75
subgroups_index (calzone.utils.data_loader attribute),
Т
transform_topclass()
                            (calzone.utils.data_loader
         method), 75
```

transform_topclass() (in module calzone.utils), 79

84 Index