

# Project #2

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PHSX 815

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## Abstract

This project compared two hypotheses and attempted to distinguish between the two. The Null Hypothesis assumed that the rate parameter  $\lambda$  for a Poisson Distribution was constant, while the Alternative Hypothesis assumed that the rate parameter  $\lambda$  was distributed according to a Gamma Distribution. Data was generated for each hypothesis with  $N_{meas} = 1,000$  measurements per experiment, and  $N_{exp} = 10,000$  experiments total. A 95% Confidence Level was used for the Hypothesis comparison, which was done using the Log Likelihood Ratio. The power to distinguish between the two hypotheses was found to be 39.4%. Reasons for the low power of test will be discussed.

## 1 Null Hypothesis

The null hypothesis assumes that the rate parameter  $\lambda$  in a Poisson Distribution is constant. The Poisson Distribution is,

$$P(X|\lambda) = \frac{\lambda^X e^{-\lambda}}{X!}, \quad (1)$$

where  $\lambda$  is a positive, real number, representing the average rate that an event occurs at, and  $X$  is a whole number that represents the number of events that occurred within some fixed time interval.

The Poisson Distribution can be seen for various values of  $\lambda$  in Fig. 1. As  $\lambda$  gets sufficiently large, the distribution approaches that of a normal distribution, as predicted by the Central Limit Theorem.

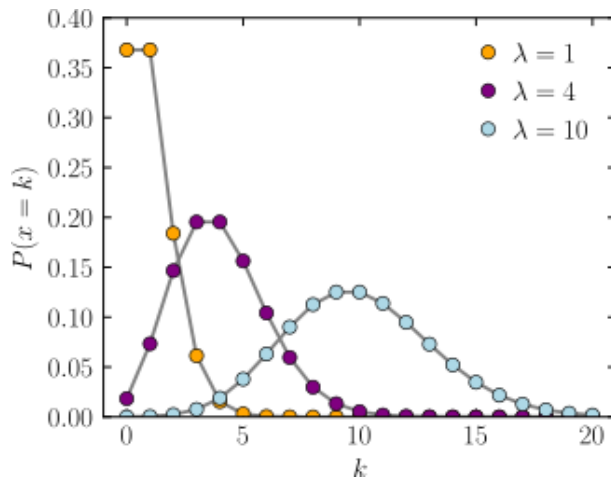


Figure 1: Poisson distributions with various values of  $\lambda$ . The expected peak in the distribution (the mean) occurs at the value of  $\lambda$ .

24 The likelihood of the a distribution is found by Baye's Theorem.

$$P(\lambda|X) \approx \prod_{i=1}^{N_{meas}} P(X_i|\lambda) = \prod_{i=1}^{N_{meas}} \frac{\lambda^{X_i} e^{-\lambda}}{X_i!} \quad (2)$$

25 where  $X_i$  is each individual measurement made in a single experiment and  
 26  $N_{meas}$  is the total number of measurements in the given experiment.

27 A more useful quantity is the logarithm of the likelihood, as it allows us to  
 28 turn the product into a summation using properties of logarithmic functions.  
 29 Going through the calculation, and using the Stirling approximation for the  
 30 factorial results in the following expression.

$$\ln(P(\lambda|X)) = \sum_{i=1}^{N_{meas}} [x_i \ln(\lambda) - \frac{1}{2} \ln(2\pi x_i) - x_i \ln(x_i) + x_i - \lambda] \quad (3)$$

31 Which is the expression needed to calculate the likelihood of the Null Hy-  
 32 pothesis given a sample data set.

33 For the simulation, random measurements from the Poisson Distribution  
 34 were done using a standard function from the numpy library. The rate param-  
 35 eter was chosen to be  $\lambda = 5.0$ , the number of measurements per experiment  
 36 was  $N_{meas} = 1,000$ , and the number of experiments was  $N_{exp} = 10,000$ . The  
 37 data for the Null Hypothesis can be seen in Fig. 2.

---

```

38
39 #Poisson Distribution
40 def Poisson(self, lamb):
41     return np.random.poisson(lam=lamb, size=1)
42

```

---

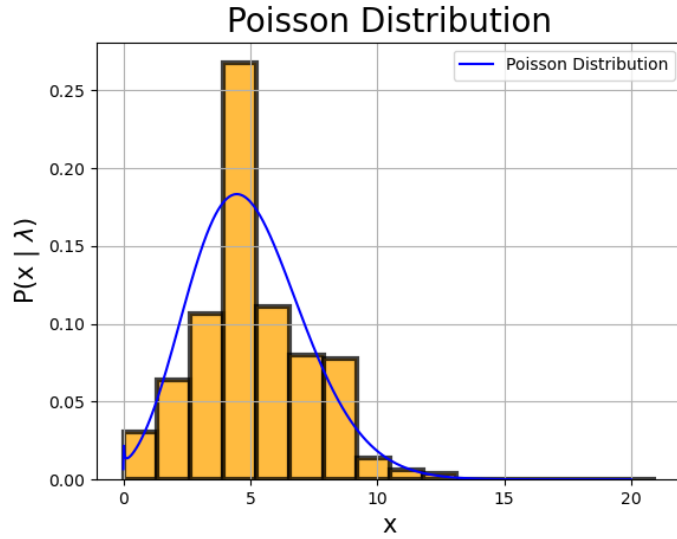


Figure 2: The data generated by the Null Hypothesis assuming a constant rate parameter  $\lambda = 5.0$ . The curve in blue is the actual distribution.

## 2 Alternative Hypothesis

The Alternative Hypothesis assumed that the rate parameter  $\lambda$  was distributed according to a Gamma Distribution. The Gamma Distribution is,

$$P(\lambda|\alpha, \beta) = \frac{\beta^\alpha}{\Gamma(\alpha)} \lambda^{\alpha-1} e^{-\beta\lambda}, \quad (4)$$

where  $\alpha$  is a positive, real number representing the shape of the distribution,  $\beta$  is a positive, real number representing the width of the distribution, and  $\Gamma(\alpha)$  is the Gamma function.

Fig. 3 shows the Gamma Distribution for various values of  $\alpha$  and  $\beta$ . As the value  $\frac{\alpha}{\beta}$  gets sufficiently large, the distribution approaches a normal distribution as predicted by the Central Limit Theorem.

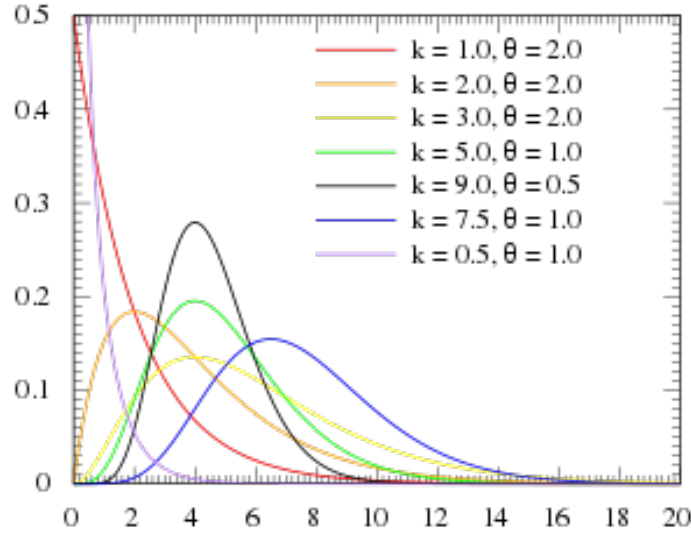


Figure 3: Gamma distributions with various values of  $\alpha = k$  and  $\beta = \frac{1}{\theta}$ . The expected peak in the distribution (the mean) occurs at the value of  $\frac{\alpha}{\beta}$ .

The simulation for the Alternative Hypothesis worked as follows. The parameters for the Gamma Distribution were chosen such that the most likely  $\lambda$  value remains the same as the Null Hypothesis:  $\alpha = 6.0$ , and  $\beta = 1.2$ . For each new experiment, a new value of the rate parameter  $\lambda$  was generated from this Gamma Distribution. Then a data set  $X$  was generated from a Poisson Distribution for that value of  $\lambda$ . The number of measurements per experiment and the number of experiments remains the same as that for the Null Hypothesis.

Fig. 4 shows the distribution of  $\lambda$  values according to the Gamma Distribution, and Fig. 5 shows the data generated under this multi-step hypothesis.

Random measurements of  $\lambda$  from the Gamma Distribution were done using a standard function from the numpy library.

---

```
#Gamma Distribution
def Gamma(self, alpha, beta):
```

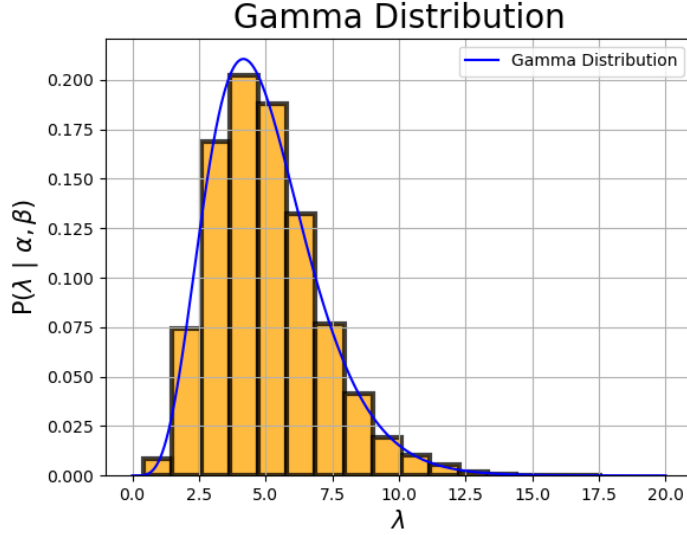


Figure 4: The distribution of  $\lambda$  values according to a Gamma distribution with parameters  $\alpha = 6.0$  and  $\beta = 1.2$ . The expected peak in the distribution (the mean) occurs at the value of  $\lambda = \frac{\alpha}{\beta} = 5.0$ . The curve in blue is the actual distribution.

```

67     #both alpha and beta must be positive
68     k = alpha
69     theta = 1/beta
70
71     return np.random.gamma(shape = k, scale = theta, size = 1)
72

```

---

### 3 Hypothesis Comparison

Hypothesis Comparison was done using the Log Likelihood Ratio (LLR).

$$LLR = \ln\left(\frac{P(X|H_0)}{P(X|H_1)}\right) = \ln(P(X|H_0)) - \ln(P(X|H_1)) \quad (5)$$

where  $H_0$  is the Null Hypothesis described by the parameter  $\lambda$ , and  $H_1$  is the Alternative Hypothesis described by the parameters  $\alpha$  and  $\beta$ .

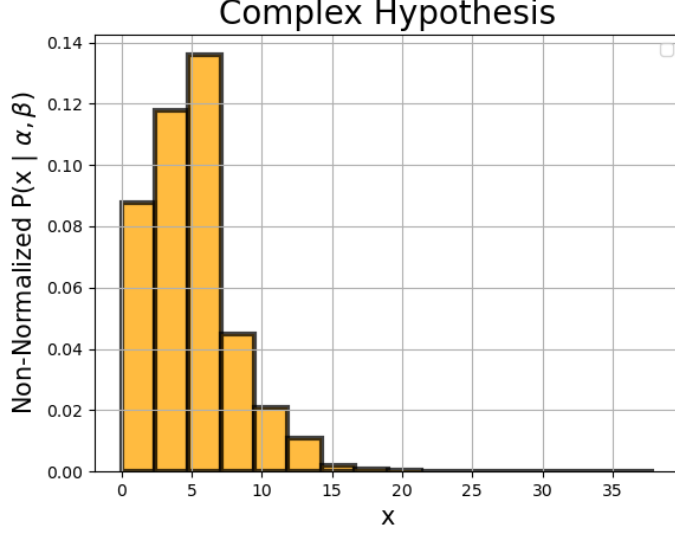


Figure 5: The data generated under the Alternative Hypothesis where the rate parameter of a Poisson Distribution  $\lambda$  was distributed according to a Gamma Distribution.

77 The log likelihood of a data set for the Null Hypothesis is described  
78 succinctly by Eqn. 3. The Alternative Hypothesis is more complex due  
79 to it's multi-step nature. The probability distribution shown in Fig. 5 is  
80 described by the following equation.

$$P(X|\alpha, \beta) = \int P(X|\lambda)P(\lambda|\alpha, \beta)d\lambda \quad (6)$$

81 where  $P(X|\lambda)$  is the Poisson Distribution, and  $P(\lambda|\alpha, \beta)$  is the Gamma Dis-  
82 tribution, and all  $\lambda$  values are summed over.

83 The likelihood is once again given by Baye's Theorem shown in Eqn.  
84 2. While this likelihood can be calculated analytically, we instead took a  
85 numerical approach using the histogram shown in Fig. 5. While useful to  
86 visualize the data under the Alternative Hypothesis, the histogram was really  
87 plotted in order to save two lists. The first list,  $n_{save}$ , contains the height of  
88 each bin. The second list,  $bins_{save}$ , has the location of each bin edge. Once  
89 the histogram has been normalized (by ensuring that the sum of all of the bin  
90 heights is equal to one), the height of each bin is the probability of measuring  
91 that data point (number of events).

92 Overall, the log likelihood function for the complex hypothesis is as fol-  
93 lows.

---

```
94
95     #Log Likelihood for Complex Hypothesis
96     logprob_min = np.log(1/len(hist_complex))
97
98     def ComplexLikelihood(data): #takes in one experiment
99         g = 0 #initialize value
100
101         for d in data: #measurements in an experiment
102             bin_current = 0
103
104             #protect against going over the farthest right bin edge
105             if d > bins_save[len(bins_save) - 1]:
106                 logprob = logprob_min
107
108             #find what bin the measurement falls in, starting from
109             the left
110             else:
111                 while d > bins_save[bin_current + 1]:
112                     bin_current += 1
113
114             #protect against bins w/ no counts
115             if n_save[bin_current] <= 0:
116                 logprob = logprob_min
117
118             else:
119                 logprob = np.log(n_save[bin_current]) #once
120                 normalized, height of the histogram =
121                 probability of measurement
122
123             g += logprob
124
125         return g
126
```

---

127 Generating the Log Likelihood Ratio, as shown in Eqn. 5, for each Hy-  
128 pothesis requires that you compute the likelihood of each hypothesis for each  
129 data set. The code is shown below.

---

```
130
131 #Log Likelihood Ratios
```

```

132 LL_simple_simple = [] #Log Likelihood of the Simple Hypothesis
133     using the data made from the Simple Hypothesis
134 LL_complex_simple = [] #Log Likelihood of the Complex
135     Hypothesis using the data made from the Simple Hypothesis
136
137 LL_simple_complex = [] #Log Likelihood of the Simple Hypothesis
138     using the data made from the Complex Hypothesis
139 LL_complex_complex = [] #Log Likelihood of the Complex
140     Hypothesis using the data made from the Complex Hypothesis
141
142 LLR_simple = []
143 LLR_complex = []
144
145 #Log Likelihood of the Simple Hypothesis using the data made
146     from the Simple Hypothesis
147 for exp in data_simple: #each experiment
148     LL = LogLikelihoodPoisson(lamb, exp)
149     LL_simple_simple.append(LL)
150
151 #Log Likelihood of the Complex Hypothesis using the data made
152     from the Simple Hypothesis
153 for exp in data_simple: #each experiment
154     LL = ComplexLikelihood(exp)
155     LL_complex_simple.append(LL)
156
157 #Log Likelihood of the Simple Hypothesis using the data made
158     from the Complex Hypothesis
159 for exp in data_complex: #each experiment
160     LL = LogLikelihoodPoisson(lamb, exp)
161     LL_simple_complex.append(LL)
162
163 #Log Likelihood of the Complex Hypothesis using the data made
164     from the Complex Hypothesis
165 for exp in data_complex: #each experiment
166     LL = ComplexLikelihood(exp)
167     LL_complex_complex.append(LL)
168
169 #LLR for simple hypothesis
170 for i in range(len(LL_simple_simple)):
171     LLR_simple.append(LL_simple_simple[i] -

```



```

172         LL_complex_simple[i])
173
174     #LLR for complex hypothesis
175     for i in range(len(LL_simple_complex)):
176         LLR_complex.append(LL_simple_complex[i] -
177                             LL_complex_complex[i])
178

```

---

179 The last portion of the code sorted the LLR lists in ascending order and  
180 then found the power of the test by finding the critical LLR value (determined  
181 by the confidence level) in both lists.

---

```

182
183     #Sort the LLRs
184     LLR_simple = Sorter.DefaultSort(LLR_simple)
185     LLR_complex = Sorter.DefaultSort(LLR_complex)
186
187     #Define Confidence Level and find critical LLR value
188     alpha_CL = 0.05 #Confidence Level = 95%
189
190     LLR_alpha_CL = LLR_simple[math.ceil(Nexp * alpha_CL)] #critical
191                     LLR value
192
193     #Find Beta and Power of Test
194     for i in range(len(LLR_complex)):
195         if LLR_complex[i] >= LLR_alpha_CL:
196             LLR_Beta = LLR_complex[i]
197             LLR_Beta_Position = i
198             break
199
200     Beta_CL = (len(LLR_complex) - LLR_Beta_Position)/Nexp #Beta_CL
201               = percent of entries in LLR_complex above LLR_alpha_CL
202
203     Power = 1 - Beta_CL
204

```

---

## 205 4 Results

206 Results are summarized in Fig. 6. The power of the test to discern  
207 between the two hypothesis was found to be 39%, which is quite low for  
208 a number of reasons. The first is that the number of measurements per

209 experiment,  $N_{meas}$  was only 1,000, which is quite low. Using a higher value  
 210 for  $N_{meas}$  would improve the power of the test. The second reason is that the  
 211 two hypothesis had the same Most Likely Estimate (MLE) for  $\lambda$ . Changing  
 212 the parameters of the Alternative Hypothesis,  $\alpha$  and  $\beta$  to create a different  
 213 MLE for the value of  $\lambda$  would make the two hypotheses more distinguishable.

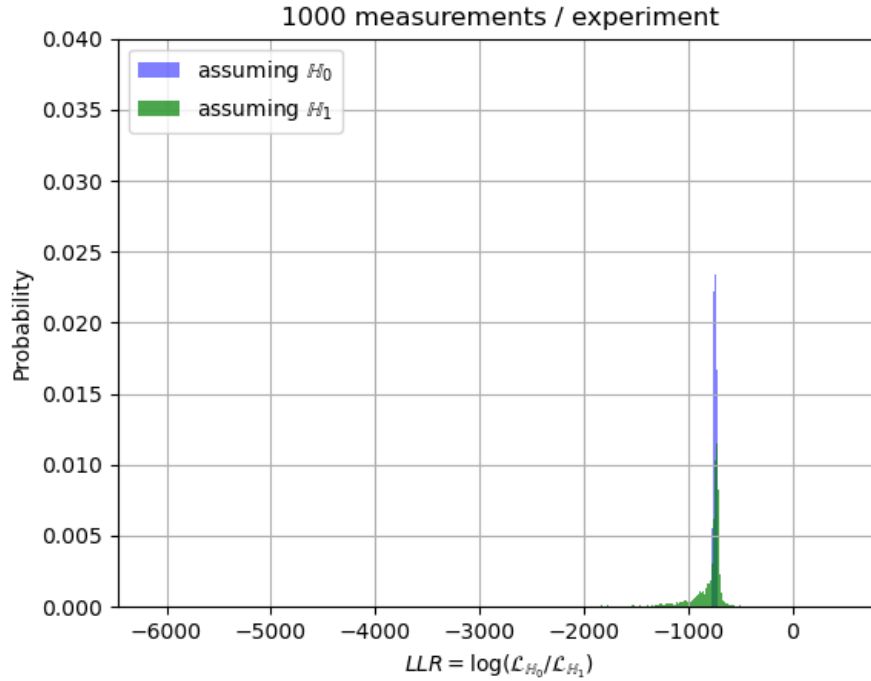


Figure 6: A graph of the log likelihood ratio for the two Hypotheses. The Null Hypothesis (blue) assumes a constant rate parameter  $\lambda$ , while the Alternative Hypothesis assumes that the rate parameter  $\lambda$  is distributed according to a Gamma Distribution. Parameter values were chosen such that both hypotheses had the same MLE for the value  $\lambda$ , which makes it hard to differentiate between the two hypotheses.

## 214 5 Summary

215 This paper described the process for how to use the Log Likelihood Ratio  
216 to distinguish between two hypotheses. The Null Hypothesis was quite sim-  
217 ple, assuming a constant rate parameter of  $\lambda = 5$  for the Poisson Distribution.  
218 The Alternative Hypothesis utilized a new rate parameter for each experi-  
219 ment, being randomly generated from a Gamma Distribution with  $\alpha = 6.0$ ,  
220 and  $\beta = 1.2$ . Due to the similarity of the two hypotheses, the power of the  
221 test was quite low, found to be 39.4%.

222 Improving the power of test should be done by simulating more measure-  
223 ments per experiment (which my computer could not handle for the given  
224 number of experiments), and by choosing values of  $\alpha$  and  $\beta$  such that the  
225 Maximum Likelihood Estimate of the most-likely rate parameter  $\lambda$  is distinct  
226 between the two hypotheses.

## 227 6 Hypothesis.py

228 The code file **Hypothesis.py** was used to do the generation of the data  
229 and the analysis of the data for this simulation. Since most of the analysis  
230 was shown in Section 3, only the data generation will be shown here. The  
231 entire code file can be viewed in the GitHub Repository.

---

```
232  
233 #import packages  
234 import math  
235 import numpy as np  
236 import matplotlib.pyplot as plt  
237 import sys  
238  
239 #import Random class  
240 sys.path.append(".")  
241 import Random as rng  
242  
243 #import Sorting class  
244 sys.path.append(".")  
245 import MySort as mys  
246  
247 # main function  
248 if __name__ == "__main__":  
249     # if the user includes the flag -h or --help print the options
```

```

250 if '-h' in sys.argv or '--help' in sys.argv:
251     print ("Usage: %s [options]" % sys.argv[0])
252     print (" options:")
253     print (" --help(-h)          print options")
254     print (" -seed [integer number] seed")
255     print (" -Nmeas [integer number] number of measurements
256           per experiments")
257     print (" -Nexp [integer number] number of experiments")
258     print (" -lambda [float number] Parameter for the Null
259           Hypothesis")
260     print (" -alpha [float number] Parameter for the
261           Alternative Hypothesis")
262     print (" -beta [float number] Parameter for the
263           Alternative Hypothesis")
264     print
265     sys.exit(1)
266
267 #Initialize
268 seed = 5555
269
270 Nmeas = 1 #number of measurements per experiment
271 Nexp = 1 #number of experiments
272
273 lamb = 1.0 #must be positive
274 alpha = 1.0 #must be positive
275 beta = 1.0 #must be positive
276
277 #System Inputs
278 if '-seed' in sys.argv:
279     p = sys.argv.index('-seed')
280     seed = sys.argv[p+1]
281
282 if '-Nmeas' in sys.argv:
283     p = sys.argv.index('-Nmeas')
284     ptemp = int(sys.argv[p+1])
285     Nmeas = ptemp
286
287 if '-Nexp' in sys.argv:
288     p = sys.argv.index('-Nexp')
289     ptemp = int(sys.argv[p+1])

```

```

290     Nexp = ptemp
291
292     if '-lambda' in sys.argv:
293         p = sys.argv.index('-lambda')
294         ptemp = float(sys.argv[p+1])
295         lamb = ptemp
296
297     if '-alpha' in sys.argv:
298         p = sys.argv.index('-alpha')
299         ptemp = float(sys.argv[p+1])
300         alpha = ptemp
301
302     if '-beta' in sys.argv:
303         p = sys.argv.index('-beta')
304         ptemp = float(sys.argv[p+1])
305         beta = ptemp
306
307     #class instance of Random and Sorting class
308     random = rng.Random(seed)
309     Sorter = mys.MySort()
310
311     #initialize data
312     data_simple = [] #[[exp1], [exp2], ...] each experiment in the
313                     simple hypothesis
314     data_graph = [] #[meas1, meas1, ...] used to plot data from
315                     simple hypothesis
316
317     data_complex = [] #[[exp1], [exp2], ...] each experiment in the
318                      complex hypothesis
319     hist_complex = [] #[meas1, meas1, ...] every measurement in the
320                      complex hypothesis from all lambdas
321     lamb_graph = [] #[lamb1, lamb2, ...] used to plot the
322                      distribution of lambdas
323
324     #Generate Data for Simple Hypothesis (fixed lambda)
325     for e in range(0, Nexp): #each experiment
326         data_exp_simple = [] #all measurements in a given experiment
327
328         for m in range(0, Nmeas): #each measurement
329             measurement_simple = float(random.Poisson(lamb))

```

```

330         data_exp_simple.append(measurement_simple)
331         data_graph.append(measurement_simple)
332
333     data_simple.append(data_exp_simple)
334
335     #Generate Data for Complex Hypothesis (lambda comes from a
336     Gamma Distribution)
337     for e in range(0, Nexp): #each experiment
338         data_exp_complex = [] #all measurements in a given
339         experiment
340
341         lamb_complex = float(random.Gamma(alpha, beta)) #new lambda
342         every experiment
343
344         lamb_graph.append(lamb_complex)
345
346         for m in range(0, Nmeas): #each measurement
347             measurement_complex = float(random.Poisson(lamb_complex))
348             hist_complex.append(measurement_complex)
349             data_exp_complex.append(measurement_complex)
350
351         data_complex.append(data_exp_complex)
352

```

---