Tutorial 4: MAP

1 Theory

Statistical approach to classification, assuming input and output are random variables with some known probability distribution.

Notation

 Ω - Finite set of classes $\omega_i \in \Omega$, i = 1, 2, ..., N.

X - Input space $x \in X$.

We assume that the following distributions are known:

- $\{p(\omega_1), p(\omega_2), ..., p(\omega_n)\}$ A posteriori probabilities the classes.
- $\{p(x|\omega_1), p(x|\omega_2), ..., p(x|\omega_n)\}$ Probability of an input x given a class ω_i (if X is a continuous space, then $p(x|\omega_i)$ is a probability density).

Maximum A Posteriori (MAP)

$$\hat{\omega}_{\text{MAP}} \triangleq \underset{\omega}{\text{arg max }} p(\omega|x).$$

Recall Bayes' rule:

$$p(\omega|x) = \frac{p(x|\omega)p(\omega)}{p(x)}.$$

Using the latter, we can rewrite our MAP estimator as

$$\begin{split} \hat{\omega}_{\text{MAP}} &\triangleq \arg\max_{\omega} \, p(\omega|x) = \arg\max_{\omega} \, \frac{p(x|\omega)p(\omega)}{p(x)} = \arg\max_{\omega} \, p(x|\omega)p(\omega). \\ \\ &\Rightarrow \boxed{\hat{\omega}_{\text{MAP}} = \arg\max_{\omega} \, p(x|\omega)p(\omega)} \end{split}$$

As shown in the lecture, this classifier minimizes both the conditional error probability, and the average error probability.

2 Practice

Question 1

A transmitted message S is either 1 or 0 is with probability p and 1-p respectively. Unfortunately the message gets corrupted by a Gaussian noise N with zero mean $(\mu = 0)$ and unit variance $\sigma^2 = 1$. Hence, the received message Y is given by

$$Y = S + N$$
.

Compute the MAP estimate of S given Y = y.

Solution

The MAP estimator is given by

$$\begin{split} \hat{S}_{\text{MAP}} &= \underset{s = \{0,1\}}{\text{arg max}} \ P(S = s | Y = y) \\ &= \underset{s = \{0,1\}}{\text{arg max}} \ \left\{ f_{Y|S}(y|0)P(S = 0), \ f_{Y|S}(y|1)P(S = 1) \right\} \\ &= \underset{s = \{0,1\}}{\text{arg max}} \ \left\{ f_{Y|S}(y|0)(1 - p), \ f_{Y|S}(y|1)p \right\} \end{split}$$

Notice that given S = 0, Y becomes equal to the noise N, and therefore

$$f_{Y|S}(y|0) = \frac{1}{\sqrt{2\pi}}e^{-y^2/2}.$$

Given that S = 1, Y becomes Y = N + 1, hence

$$f_{Y|S}(y|1) = \frac{1}{\sqrt{2\pi}}e^{-(y-1)^2/2}.$$

Consequently, $\hat{S}_{\text{MAP}} = 1$ if

$$\frac{1}{\sqrt{2\pi}} e^{-(y-1)^2/2} p \ge \frac{1}{\sqrt{2\pi}} e^{-y^2/2} (1-p)$$

$$\Rightarrow -\frac{(y-1)^2}{2} + \log(p) \ge -\frac{y^2}{2} + \log(1-p)$$

$$\Rightarrow y \ge \frac{1}{2} + \log\left(\frac{1-p}{p}\right).$$

Hence,

$$\hat{S}_{\text{MAP}}(y) = \begin{cases} 1, & y \ge \frac{1}{2} + \log\left(\frac{1-p}{p}\right), \\ 0, & \text{o.w.} \end{cases}$$

As an example, for $p = 1 - p = \frac{1}{2}$ we get that S = 1 if $y \ge \frac{1}{2}$, i.e., the threshold is right in the middle of 0 and 1.

Question 2

Let X be a continuous random variable with the following PDF:

$$f_X(x) = \begin{cases} 2x, & 0 \le x \le 1, \\ 0, & \text{o.w.} \end{cases}$$

In addition, it is given that $Y|X=x\sim Geometric(x)$.

Given Y = 3, find the MAP estimate of X.

Remainder: The geometric distribution gives the probability that the first occurrence of success requires k independent trials, each with success probability p:

$$P(Y = k) = p(1 - p)^{k-1}$$
.

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Solution

We know that $Y|X = x \sim Geometric(x)$, so

$$P_{Y|X}(y|x) = x(1-x)^{y-1}.$$

Therefore,

$$P_{Y|X}(3|x) = x(1-x)^2.$$

The MAP estimate of X is given by

$$\hat{X}_{\text{MAP}} = \underset{x \in [0,1]}{\arg\max} \ P_{Y|X}(3|x) f_X(x) = x(1-x)^2 \cdot 2x = \underset{x \in [0,1]}{\arg\max} \ 2x^2(1-x)^2.$$

Setting the derivative to 0 we get

$$4x(1-x)^2 - 4x^2(1-x) = 0$$

Solving for x (and checking for maximization criteria), we obtain the MAP estimate as

$$\hat{X}_{\text{MAP}} = \frac{1}{2}.$$

Question 3

It is given that $X = \mathbb{R}^2$, and that $\Omega = \{\omega_1, \omega_2, \omega_3\}$.

The class prior probabilities are uniform: $p(\omega_1) = p(\omega_2) = p(\omega_3) = \frac{1}{3}$.

The class conditional probability is Gaussian, $p(x|\omega_i) \sim \mathcal{N}(\mu_i, \Sigma)$, where

$$\mu_1 = \begin{bmatrix} 0 \\ 3 \end{bmatrix}, \ \mu_2 = \begin{bmatrix} 3 \\ 0 \end{bmatrix}, \ \mu_3 = \begin{bmatrix} -3 \\ 0 \end{bmatrix}, \ \Sigma = \begin{bmatrix} 2 & 0 \\ 0 & 2 \end{bmatrix}.$$

- (a) What is the Bayes optimal decision rule? What are the decision boundaries in the plane \mathbb{R}^2 ?
- (b) Does the decision boundary for the Gaussian case always have the same shape? What it depends on?

Solution

(a) We need to compute the conditions for which $P(\omega_i|x) \ge P(\omega_j|x)$ $(i, j = 1, 2, 3, i \ne j)$:

$$P(\omega_{i}|x) \geq P(\omega_{j}|x)$$

$$\Leftrightarrow P(x|\omega_{i})p(\omega_{i}) \geq P(x|\omega_{j})p(\omega_{j})$$

$$\Leftrightarrow P(x|\omega_{i}) \geq P(x|\omega_{j})$$

$$\Leftrightarrow \frac{1}{\sqrt{4\pi^{2}|\Sigma|}} \exp\left(-\frac{1}{2}(x-\mu_{i})^{T}\Sigma^{-1}(x-\mu_{i})\right) \geq \frac{1}{\sqrt{4\pi^{2}|\Sigma|}} \exp\left(-\frac{1}{2}(x-\mu_{j})^{T}\Sigma^{-1}(x-\mu_{j})\right)$$

$$\Leftrightarrow \frac{1}{2}(x-\mu_{i})^{T}\Sigma^{-1}(x-\mu_{i}) \leq \frac{1}{2}(x-\mu_{j})^{T}\Sigma^{-1}(x-\mu_{j})$$

$$\Leftrightarrow ||x-\mu_{i}||_{2}^{2} \leq ||x-\mu_{j}||_{2}^{2}$$

$$\Leftrightarrow ||x-\mu_{i}||_{2} \leq ||x-\mu_{j}||_{2}$$

Thus, the decision boundary is a line in the plane \mathbb{R}^2 .

(b) Consider the following inequality

$$\frac{1}{2}(x-\mu_i)^T \Sigma^{-1}(x-\mu_i) \le \frac{1}{2}(x-\mu_j)^T \Sigma^{-1}(x-\mu_j).$$

Therefore, in general, the decision boundaries are a quadratic polynomial and their shape depends on Σ_i . Thus, in \mathbb{R}^2 the decision boundaries can be a circle, ellipse, parabola, hyperbola or a line (as in our case since all vectors have the same Σ).

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