

Problem 1

As the prisoner cannot remember his former choice, he faces the same situation when he makes choice. Assume expected time to get free is $E(t)$ which is composed by three situations, he chooses the tunnel leads to freedom within 3 hours, others two are lead to freedom within $7 + E(t)$ or $5 + E(t)$. Each of them has a probability of $\frac{1}{3}$. Therefore, the $E(t)$ is calculated by equation 1

$$E(t) = \frac{1}{3} \cdot 3 + \frac{1}{3} \cdot [E(t) + 5] + \frac{1}{3} \cdot [E(t) + 7] \quad (1)$$

Solve equation 1 the expected time it will take the prisoner to obtain freedom is 15 hours.

Problem 2

Question a

Model: $i.i.d$ samples $\{x_i\}_{i=1}^n$ are normally distributed, $p(x|y) \sim \mathcal{N}(y, 1)$

Parameter: y is the mean of the distribution.

Goal: Get the y when samples are given, which leads the samples have the max probability to happen.

Probability density of X is calculated by equation 2

$$p(X = x_i | y = \hat{y}_{MLE}) = \frac{1}{\sqrt{2\pi}\sigma^2} e^{-\frac{(x_i - \hat{y}_{MLE})^2}{2\sigma^2}} \quad (2)$$

Likelihood Function is given by equation

$$\begin{aligned} L(\hat{y}_{MLE}) &= \prod_{i=1}^n p(X = x_i | y = \hat{y}_{MLE}) \\ &= \prod_{i=1}^n \frac{1}{\sqrt{2\pi}} e^{-\frac{(x_i - \hat{y}_{MLE})^2}{2}} \\ &= \frac{n}{\sqrt{2\pi^2}} e^{-\frac{1}{2} \sum_{i=1}^n (x_i - \hat{y}_{MLE})^2} \end{aligned} \quad (3)$$

Then use \ln function to simplify,

$$\ln(L(\hat{y}_{MLE})) = \ln\left(\frac{n}{\sqrt{2\pi}}\right) - \frac{1}{2} \sum_{i=1}^n (x_i - \hat{y}_{MLE})^2 \quad (4)$$

Derivative,

$$\frac{d \ln (L(\hat{y}_{MLE}))}{d \hat{y}_{MLE}} = \frac{1}{2} \sum_{i=1}^n (x_i - \hat{y}_{MLE}) \quad (5)$$

Set the derivative equal to 0. We can get \hat{y}_{MLE}

$$\hat{y}_{MLE} = \frac{1}{n} \sum_{i=1}^n x_i \quad (6)$$

Question b

MAP function: Use the \hat{y}_{MLE} in equation 3 to

$$\begin{aligned} M(\hat{y}) &= p(\hat{y}) \cdot L(\hat{y}_{MLE}) \\ &= \frac{\sqrt{s}}{\sqrt{2\pi}} e^{-\frac{s(\hat{y}-z)^2}{2}} \cdot \frac{n}{\sqrt{2\pi}} e^{-\frac{1}{2} \sum_{i=1}^n (x_i - \hat{y})^2} \end{aligned} \quad (7)$$

Use $\ln(x)$ function to simplify:

$$L(\hat{y}) = \ln \left(\frac{\sqrt{s}}{\sqrt{2\pi}} \right) - \frac{s(\hat{y}-z)^2}{2} + \ln \left(\frac{n}{\sqrt{2\pi}} \right) - \frac{1}{2} \sum_{i=1}^n (x_i - \hat{y})^2 \quad (8)$$

Set $\frac{dL(\hat{y})}{d\hat{y}}$ equals to 0 to get the $\arg \max_{\hat{y} \in R} L(\hat{y})$

$$\frac{dL(\hat{y})}{d\hat{y}} = -s(\hat{y}-z) + \sum_{i=1}^n (x_i - \hat{y}) = 0 \quad (9)$$

Solve equation 9 then get the \hat{y}_1

$$\hat{y}_1 = \frac{n}{n+s} \hat{y}_{MLE} + \frac{sz}{n+s} \quad (10)$$

Question c

The density function of y is

$$p(y) = \begin{cases} \frac{1}{2}, & -1 \leq y \leq 1 \\ 0, & \text{o.w} \end{cases} \quad (11)$$

MAP function

$$\begin{aligned} M(\hat{y}) &= p(\hat{y}) \cdot L(\hat{y}_{MLE}) \\ &= \frac{1}{2} \cdot \frac{n}{\sqrt{2\pi}} e^{-\frac{1}{2} \sum_{i=1}^n (x_i - \hat{y})^2} \end{aligned} \quad (12)$$

Use \ln function to simplify:

$$L(\hat{y}) = \ln[M(\hat{y})] = \ln\left(\frac{1}{2}\right) + \ln\left(\frac{n}{\sqrt{2\pi}}\right) - \frac{1}{2} \sum_{i=1}^n (x_i - \hat{y})^2 \quad (13)$$

Set $\frac{dL(y)}{dy}$ equals to 0 to get the $\arg \max_{\hat{y} \in R} L(y)$ when $-1 \leq y_{MLE} \leq 1$, namely

$$\frac{dL(y)}{dy} = \sum_{i=1}^n (x_i - \hat{y}) = 0 \quad (14)$$

Solve equation 14, we can get $\hat{y}_2 = y_{MLE}$ when $-1 \leq y_{MLE} \leq 1$, while when $y_{MLE} < -1$, $\hat{y}_2 = -1$, when $y_{MLE} > 1$, $\hat{y}_2 = 1$, therefore

$$\hat{y}_2 = \begin{cases} -1, & y_{MLE} < -1 \\ y_{MLE}, & -1 \leq y_{MLE} \leq 1 \\ 1, & y_{MLE} > 1 \end{cases} \quad (15)$$

Problem 3

Question a

Random variables of Poisson distribution:

$$P(X = k) = \frac{\lambda^k}{k!} e^{-\lambda} \quad (16)$$

a set of iid samples $D = \{x_1, x_2, \dots, x_N\}$ drawn according to the Poisson probability mass:

$$L(\lambda) = P(x_1, x_2, \dots, x_N) = P(x_1)P(x_2) \cdots P(x_N) \quad (17)$$

We use \ln to simplify

$$\begin{aligned} \ln L(\lambda) &= \ln P(x_1) + \ln P(x_2) + \cdots + \ln P(x_N) \\ &= \sum_{i=1}^N (x_i \ln \lambda - \sum_{j=1}^{x_i} \ln j - \lambda) \end{aligned} \quad (18)$$

ML condition tells:

$$\frac{\partial \ln L(\lambda)}{\partial \lambda} = \sum_{i=1}^N \frac{x_i}{\lambda} - N = 0 \therefore \lambda_{ML} = \frac{1}{N} \sum_{i=1}^N x_i \quad (19)$$

$$\begin{aligned} E(\hat{\lambda}) &= E\left(\frac{1}{N} \sum_{i=1}^N x_i | \lambda\right) \\ &= \frac{1}{N} \sum_{i=1}^N E(x_i | \lambda) \\ &= \frac{1}{N} \sum_{i=1}^N \lambda \\ &= \lambda \end{aligned} \tag{20}$$

note that $E[X] = \lambda$:

$$\begin{aligned} E(\hat{x}_k) &= \sum_{i=1}^{\infty} x_i \frac{\lambda^{x_i}}{x_i!} e^{-\lambda} \\ &= \lambda \sum_{i=0}^{\infty} \frac{\lambda^{x_i}}{x_i!} e^{-\lambda} \end{aligned}$$

So the MLE is not biased.

Question b

Use y and k to reexpress λ :

$$y = P(X = 0)^2 = e^{-2\lambda} \tag{21}$$

So $\lambda = -\frac{1}{2} \ln y$. Then replace λ into the formula.

$$P(X = k) = \frac{\lambda^k}{k!} e^{-\lambda} = \frac{(-\frac{1}{2} \ln y)^k}{k!} \sqrt{y} \tag{22}$$

to a single k , here is maximum likelihood estimator:

$$\hat{y}_{MLE} = \arg \max_y P(X = k) = \frac{1}{n} \sum_{i=1}^n x_i \tag{23}$$

Because $n = 1$, $\hat{y}_{MLE} = e^{-2\lambda}$

$$b(y) = E(y) - \hat{y}_{MLE} = \sum_{k=0}^{\infty} e^{-2k} \frac{\lambda^k}{k!} e^{-\lambda} = e^{e^{-2}\lambda - \lambda} - e^{-2\lambda} \neq 0 \tag{24}$$

Obviously, \hat{y}_{MLE} is biased.

Question c

Assume that y_U is biased:

$$E(y_{MLE}) = \sum_{k=0}^{\infty} y_U P(X = k) = \sum_{k=0}^{\infty} y_U \frac{\lambda^k}{k!} e^{-\lambda} = y = e^{-2\lambda} \quad (25)$$

Because $e^{\lambda} = \sum_{k=0}^{\infty} \frac{\lambda^k}{k!}$, $y_U = (-1)^k$

Question d

$$\begin{aligned} MSE_{ML} &= E[(p_{MLE} - p)^2] \\ &= \sum_{k=1}^{\infty} (e^{-2k} - e^{-2\lambda})^2 \frac{\lambda^k}{k!} e^{-\lambda} \\ &= e^{(e^{-4}-1)\lambda} - 2e^{(e^{-2}-2)\lambda} + e^{-4\lambda} MSE_U \\ &= E[(p_U - p)^2] \\ &= \sum_{k=1}^{\infty} ((-1)^k - e^{-2\lambda})^2 \frac{\lambda^k}{k!} e^{-\lambda} \\ &= 1 - e^{-4\lambda} \end{aligned} \quad (26)$$

calculate it with Python, set:

$$\begin{aligned} f(x) &= MSE_{ML} - MSE_U \\ &= [e^{(e^{-4}-1)x} - 2e^{(e^{-2}-2)x} + e^{-4x}] - [1 - e^{-4x}], x > 0 \end{aligned} \quad (27)$$

```
import matplotlib.pyplot as plt
import numpy as np
import math

e = math.e

x = np.linspace(0, 20, 500)
y = (e**((e**(-4)-1)*x) - 2*e**((e**(-2)-2)*x) + e**(-4*x)) - (1 - e**(-4*x))

plt.title("compare(ML vs U)")
plt.plot(x, y)
plt.show()
```

and then draw the image of $f(x)$, it shows in figure1

$f(x) < 0, x > 0 \therefore MSE_{ML} < MSE_U$

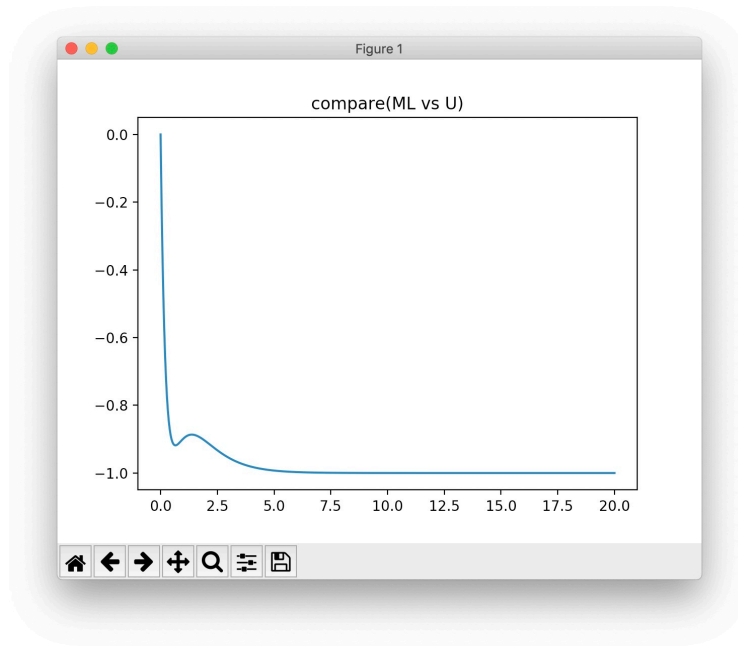


Figure 1: Figure1

Problem 4

Question a

$$\hat{\mu} = \frac{1}{n} \sum_{i=1}^n x_i \quad (28)$$

$$\hat{\Sigma} = \frac{1}{n} \sum_{i=1}^n (x_i - \hat{\mu})(x_i - \hat{\mu})^T \quad (29)$$

$$\begin{aligned}
 E[\hat{\Sigma}_{MLE}] &= E\left[\frac{1}{n} \sum_{i=1}^n (x_i - \hat{\mu}_{MLE})(x_i - \hat{\mu}_{MLE})^T\right] \\
 &= \frac{1}{n} \sum_{i=1}^n E\left[x_i \cdot x_i^T - x_i \cdot \hat{\mu}_{MLE}^T - \hat{\mu}_{MLE} \cdot x_i^T + \hat{\mu}_{MLE} \cdot \hat{\mu}_{MLE}^T\right] \\
 &= \frac{1}{n} \sum_{i=1}^n E\left[x_i \cdot x_i^T - x_i \cdot \frac{1}{n} \sum_{j=1}^n x_j^T - \frac{1}{n} \sum_{j=1}^n x_j \cdot x_i^T + \frac{1}{n} \sum_{j=1}^n x_j \cdot \frac{1}{n} \sum_{j=1}^n x_j^T\right] \\
 &= \frac{1}{n} \sum_{i=1}^n \left(E[x_i \cdot x_i^T] - \frac{2}{n} \sum_{j=1, j \neq i}^n E[x_i \cdot x_j^T] - \frac{2}{n} E[x_i \cdot x_i^T] + \frac{1}{n^2} \sum_{j=i}^n E[x_j \cdot x_j^T] + \frac{1}{n^2} \sum_{j=1}^n \sum_{k=1, k \neq j}^n E[x_j \cdot x_k^T] \right) \\
 &= \frac{1}{n} \sum_{i=1}^n \left(\frac{n-2}{n} E[x_i \cdot x_i^T] - \frac{2n-2}{n} \mu_{MLE} \cdot \mu_{MLE} + \frac{1}{n} E[x_i \cdot x_i^T] + \frac{n^2-n}{n^2} \mu_{MLE} \cdot \mu_{MLE}^T \right) \\
 &= \frac{n-1}{n^2} \sum_{i=1}^n E[x \cdot x^T - x \cdot \mu^T - \mu \cdot x^T + \mu \cdot \mu^T] \\
 &= \frac{n-1}{n^2} \sum_{i=1}^n E[(x - \mu)(x - \mu)^T] \\
 &= \frac{n-1}{n^2} \sum_{i=1}^n \Sigma \\
 &= \frac{n-1}{n} \Sigma
 \end{aligned} \tag{30}$$

Obviously, $\hat{\Sigma}_{MLE}$ is biased

Question b

$$\hat{\Sigma}' = \frac{n}{n-1} \hat{\Sigma}_{MLE} = \frac{1}{n-1} \sum_{i=1}^n (x_i - \hat{\mu})(x_i - \hat{\mu})^T \tag{31}$$

Problem 5

Question a

We choose k samples from range $[1, 2, \dots, N]$ So the **ML Function** is

$$M(N) = p(x_1, x_2, \dots, x_k | N) = \prod_{i=0}^{i=k-1} \frac{1}{N-i} \tag{32}$$

When we get the $\arg \max_{N \in R} M(N)$ Then define a new function $t(N)$ to simplify, we just need to consider about denominator of $M(x)$

$$t(N) = N \cdot (N-1) \cdots (N-k+1) = \frac{N!}{k! (N-k)!} \cdot k! = C_N^k \cdot k! \tag{33}$$

It is easy to find that $\arg \max_{N \in R} M(N) = \arg \min_{N \in R} M(N)$ Then we find the $\arg \min_{N \in R} M(N)$

$$\begin{aligned}
 \arg \min_{N \in R} M(N) &= \arg \min_{N \in R} C_N^k \cdot k! \\
 &= \arg \min_{N \in R} \frac{N!}{k!(N-k)!} \\
 &= \arg \min_{N \in R} N \\
 &= \arg \max_{x \in \{x_i\}_1^k} x_i
 \end{aligned} \tag{34}$$

Therefore, $\hat{N}_{MLE} = \max\{x_i | x_i \in \{x_i\}_1^k\}$

Question b

We should prove $E(\hat{N}_{MLE}) \neq N$

$$\begin{aligned}
E(\hat{N}_{MLE}) &= \sum_{i=1}^N P(\hat{N}_{MLE} = i) \cdot i \\
&= \sum_{i=1}^N [P(\hat{N}_{MLE} \leq i) - P(\hat{N}_{MLE} \leq i-1)] \cdot i \\
&= \underbrace{\left(\frac{C_N^k}{C_N^k} - 0\right)}_{i \leq k} \cdot k + \underbrace{\sum_{i=k+1}^N \left[\frac{C_i^k}{C_N^k} - \frac{C_{i-1}^k}{C_N^k}\right]}_{k+1 \leq i \leq N} \cdot i + \underbrace{\left(1 - \frac{C_N^k}{C_N^k}\right)}_{i=N+1} + \underbrace{(1-1)}_{i > N+1} \\
&= \frac{k}{\frac{N!}{k!(N-k)!}} + \sum_{i=k+1}^N \left[\frac{\frac{i!}{k!(i-k)!}}{\frac{N!}{k!(N-k)!}} - \frac{\frac{(i-1)!}{k!(i-k-1)!}}{\frac{N!}{k!(N-k)!}} \right] \cdot i \\
&= \frac{k \cdot k! (N-k)!}{N!} + \sum_{i=k+1}^N \left[\frac{i! (N-k)!}{(i-k)! N!} - \frac{(i-1)! (N-k)!}{(i-k-1)! N!} \right] \cdot i \\
&= \frac{k \cdot k! (N-k)!}{N!} + \frac{(N-k)!}{N!} \cdot \sum_{i=k+1}^N \left[\frac{i!}{(i-k)!} + \frac{(i-1)!}{(i-k-1)!} \right] \cdot i \\
&= \frac{(N-k)!}{N!} \left(k \cdot k! + \sum_{i=k+1}^N \left[\frac{i!}{(i-k)!} + \frac{i!}{(i-k)} \cdot \frac{i-k}{i} \right] \cdot i \right) \\
&= \frac{(N-k)!}{N!} \left(k \cdot k! + \sum_{i=k+1}^N \frac{i!}{(i-k)!} (2i-k) \right) \\
&= \frac{(N-k)!}{N!} \left(k \cdot k! + (k+1)! \sum_{i=k+1}^N \frac{i!}{(i-k-1)! (k+1)!} \cdot \frac{2i-k}{i-k} \right) \\
&= \frac{(N-k)!}{N!} \left(k \cdot k! + (k+1)! \sum_{i=k+1}^N \frac{i!}{(i-k-1)! (k+1)!} \cdot \left(1 + \frac{i}{i-k}\right) \right) \\
&= \frac{(N-k)!}{N!} \left(k \cdot k! + (k+1)! \cdot \left(\sum_{i=k+1}^N \frac{i!}{(i-k-1)! (k+1)!} + \sum_{i=k+1}^N \frac{i!}{(i-k-1)! (k+1)!} \cdot \frac{i}{i-k} \right) \right) \\
&= \frac{(N-k)!}{N!} \left(k \cdot k! + (k+1)! \cdot \left(\sum_{i=k+1}^N \frac{i!}{(i-k-1)! (k+1)!} + \sum_{i=k+1}^N \frac{i!}{(i-k-1)! (k+1)!} \cdot \frac{i}{i-k} \right) \right) \\
&= \frac{(N-k)!}{N!} \left(k \cdot k! + (k+1)! \cdot \left(\frac{(N+1)!}{(N-k+1)! (k+1+1)!} + \sum_{i=k+1}^N \frac{i!}{(i-k-1)! (k+1)!} \cdot \left(1 + \frac{k}{i-k}\right) \right) \right) \\
&= \frac{(N-k)!}{N!} \left(k \cdot k! + (k+1)! \cdot \left(\frac{2(N+1)!}{(N-k+1)! (k+1+1)!} + \sum_{i=k+1}^N \frac{i!}{(i-k)! (k+1)!} \cdot k \right) \right) \\
&= \frac{(N-k)!}{N!} \left(k \cdot k! + (k+1)! \cdot \left(\frac{2(N+1)!}{(N-k+1)! (k+1+1)!} + \frac{k}{(k+1)} \cdot \left(\sum_{i=k}^N \frac{i!}{(i-k)! k!} - 1 \right) \right) \right) \\
&= \frac{(N-k)!}{N!} \left(k \cdot k! + (k+1)! \cdot \left(\frac{2(N+1)!}{(N-k+1)! (k+1+1)!} + \frac{k}{(k+1)} \cdot \left(\frac{(N+1)!}{(N-k)! (k+1)!} - 1 \right) \right) \right) \\
&= \frac{(N-k)!}{N!} \left(k \cdot k! + \left(\frac{2(N+1)!}{(N-k+1)! (k+2)} + \frac{k}{(k+1)} \cdot \left(\frac{(N+1)!}{(N-k)!} - 1 \right) \right) \right) \\
&= \frac{k}{k+1} (N+1) \neq N
\end{aligned}$$

Question c

According to question b unbiased estimator is $N_{ub} = \frac{k+1}{k} \hat{N}_{MLE} - 1$

Question d

The simulation result is shown in figure [2](#)

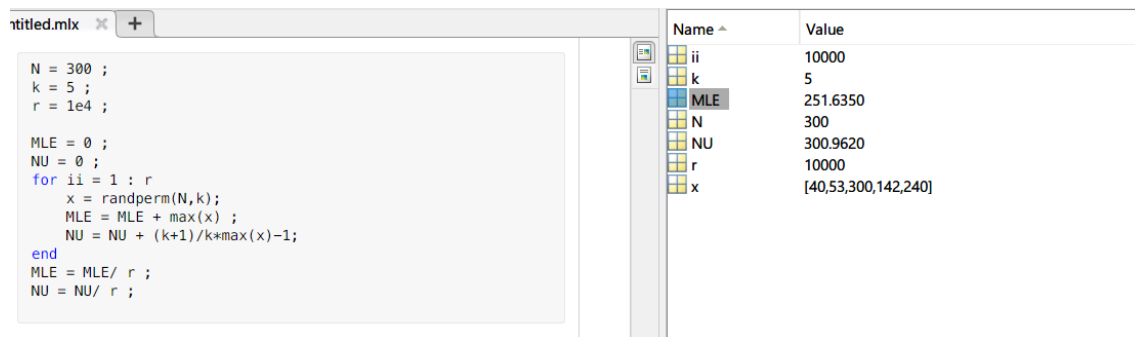


Figure 2: Figure2