

Principal Component Analysis and Linear Discriminant Analysis

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Decomposition and Components

- ▶ Decomposition is a great idea.
- ▶ Linear decomposition and linear basis, e.g., the Fourier transform
- ▶ The bases
 - ▶ construct the feature space
 - ▶ may be orthogonal bases, may be not
 - ▶ give the direction to find the components
 - ▶ specified vs. learnt?
- ▶ The features
 - ▶ are the “image” (or projection) of the original signal in the feature space
 - ▶ e.g., the orthogonal projection of the original signal onto the feature space
 - ▶ the projection does not have to be orthogonal
- ▶ Feature extraction

Outline

Principal Component Analysis

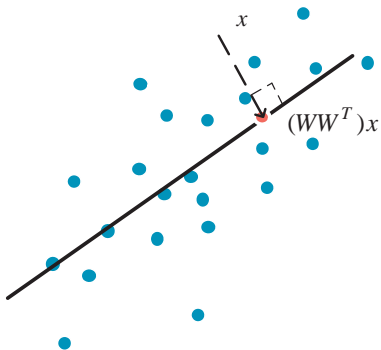
Linear Discriminant Analysis

Comparison between PCA and LDA

Principal Components and Subspaces

- ▶ Subspaces preserve part of the information (and energy, or uncertainty)
- ▶ Principal components
 - ▶ are orthogonal bases
 - ▶ and preserve the large portion of the information of the data
 - ▶ capture the major uncertainties (or variations) of data
- ▶ Two views
 - ▶ Deterministic: minimizing the distortion of projection of the data
 - ▶ Statistical: maximizing the uncertainty in data
 - ▶ are they the same?
 - ▶ under what condition they are the same?

View 1: Minimizing the MSE



- ▶ $\mathbf{x} \in \mathbb{R}^n$, and **assume centering** $E\{\mathbf{x}\} = \mathbf{0}$.
- ▶ m is the dim of the subspace, $m < n$
- ▶ orthonormal bases $\mathbf{W} = [\mathbf{w}_1, \mathbf{w}_2, \dots, \mathbf{w}_m]$
- ▶ where $\mathbf{W}^T \mathbf{W} = \mathbf{I}$, i.e., rotation
- ▶ orthogonal projection of \mathbf{x} :

$$\mathbf{P}\mathbf{x} = \sum_{i=1}^m (\mathbf{w}_i^T \mathbf{x}) \mathbf{w}_i = (\mathbf{W}\mathbf{W}^T) \mathbf{x}$$

- ▶ it achieves the minimum mean-square error (**prove it!**)

$$e_{MSE}^{min} = E\{||\mathbf{x} - \mathbf{P}\mathbf{x}||^2\} = E\{||\mathbf{P}^\perp \mathbf{x}||^2\}$$

PCA can be posed as: finding a subspace that minimizes the MSE:

$$\underset{\mathbf{W}}{\operatorname{argmin}} J_{MSE}(\mathbf{W}) = E\{||\mathbf{x} - \mathbf{P}\mathbf{x}||^2\}, \text{ s.t., } \mathbf{W}^T \mathbf{W} = \mathbf{I}$$

Let do it...

It is easy to see:

$$J_{MSE}(\mathbf{W}) = E\{\mathbf{x}^T \mathbf{x}\} - E\{\mathbf{x}^T \mathbf{P} \mathbf{x}\}$$

So,

$$\text{minimizing } J_{MSE}(\mathbf{W}) \longrightarrow \text{maximizing } E\{\mathbf{x}^T \mathbf{P} \mathbf{x}\}$$

Then we have the following constrained optimization problem

$$\max_{\mathbf{W}} E\{\mathbf{x}^T \mathbf{W} \mathbf{W}^T \mathbf{x}\} \text{ s.t. } \mathbf{W}^T \mathbf{W} = \mathbf{I}$$

The Lagrangian is

$$L(\mathbf{W}, \lambda) = E\{\mathbf{x}^T \mathbf{W} \mathbf{W}^T \mathbf{x}\} + \lambda^T (\mathbf{I} - \mathbf{W}^T \mathbf{W})$$

The set of KKT conditions gives:

$$\frac{\partial L(\mathbf{W}, \lambda)}{\partial \mathbf{w}_i} = 2E\{\mathbf{x} \mathbf{x}^T\} \mathbf{w}_i - 2\lambda_i \mathbf{w}_i, \quad \forall i$$

What is it?

Let's denote by $\mathbf{S} = E\{\mathbf{x}\mathbf{x}^T\}$ (note: $E\{\mathbf{x}\} = \mathbf{0}$).

The KKT conditions give:

$$\mathbf{S}\mathbf{w}_i = \lambda_i \mathbf{w}_i, \quad \forall i$$

or in a more concise matrix form:

$$\mathbf{S}\mathbf{W} = \lambda_i \mathbf{W}$$

What is this?

Then, the value of minimum MSE is

$$e_{MSE}^{min} = \sum_{i=m+1}^n \lambda_i$$

i.e., the sum of the eigenvalues of the orthogonal subspace to the PCA subspace.

View 2: Maximizing the Variation

Let's look it from another perspective:

- ▶ We have a linear projection of \mathbf{x} to a 1-d subspace $y = \mathbf{w}^T \mathbf{x}$
- ▶ **an important note:** $E\{y\} = 0$ as $E\{\mathbf{x}\} = 0$
- ▶ The first principal component of \mathbf{x} is such that the variance of the projection \mathbf{y} is maximized
- ▶ of course, we need to constrain \mathbf{w} to be a unit vector.
- ▶ so we have the following optimization problem

$$\max_{\mathbf{w}} J(\mathbf{w}) = E\{y^2\} = E\{(\mathbf{w}^T \mathbf{x})^2\}, \quad s.t. \mathbf{w}^T \mathbf{w} = 1$$

- ▶ **what is it?**

$$\max_{\mathbf{w}} J(\mathbf{w}) = \mathbf{w}^T \mathbf{S} \mathbf{w}, \quad s.t. \mathbf{w}^T \mathbf{w} = 1$$

Maximizing the Variation (cont.)

- ▶ The sorted eigenvalues of \mathbf{S} are $\lambda_1 \geq \lambda_2 \geq \dots \geq \lambda_n$, and eigenvectors are $\{\mathbf{e}_1, \dots, \mathbf{e}_n\}$.
- ▶ It is clearly that the first PC is $y_1 = \mathbf{e}_1^T \mathbf{x}$
- ▶ This can be generalized to m PCs (where $m < n$) with one more constraint

$$E\{y_m y_k\} = 0, \quad k < m$$

i.e., the PCs are uncorrelated with all previously found PCs

- ▶ The solution is:

$$\mathbf{w}_k = \mathbf{e}_k$$

- ▶ Sounds familiar?

The Two Views Converge

The two views lead to the same result!

- ▶ You should prove:

$\textit{uncorrelated components} \iff \textit{orthonormal projection bases}$

- ▶ What if we are more greedy, say needing independent components?
- ▶ Do we shall expect orthonormal bases?
- ▶ In which case, we still have orthonormal bases?
- ▶ We'll see it in next lecture.

The Closed-Form Solution

Learning the principal components from $\{\mathbf{x}_1, \dots, \mathbf{x}_N\}$:

1. calculating $\mathbf{m} = \frac{1}{N} \sum_{k=1}^N \mathbf{x}_k$
2. centering $\mathbf{A} = [\mathbf{x}_1 - \mathbf{m}, \dots, \mathbf{x}_N - \mathbf{m}]$
3. calculating $\mathbf{S} = \sum_{k=1}^N (\mathbf{x}_k - \mathbf{m})(\mathbf{x}_k - \mathbf{m})^T = \mathbf{A}\mathbf{A}^T$
4. eigenvalue decomposition

$$\mathbf{S} = \mathbf{U}^T \Sigma \mathbf{U}$$

5. sorting λ_i and \mathbf{e}_i
6. finding the bases

$$\mathbf{W} = [\mathbf{e}_1, \mathbf{e}_2, \dots, \mathbf{e}_m]$$

Note: The components for \mathbf{x} is

$$\mathbf{y} = \mathbf{W}^T(\mathbf{x} - \mathbf{m}), \text{ where } \mathbf{x} \in \mathbb{R}^n \text{ and } \mathbf{y} \in \mathbb{R}^m$$

A First Issue

- ▶ n is the dimension of input data, N is the size of the training set
- ▶ In practice, $n \gg N$
 - ▶ E.g., in image-based face recognition, if the resolution of a face image is 100×100 , when stacking all the pixels, we end up $n = 10,000$.
- ▶ Note that \mathbf{S} is a $n \times n$ matrix
- ▶ Difficulties:
 - ▶ \mathbf{S} is ill-conditioned, as in general $\text{rank}(\mathbf{S}) \ll n$
 - ▶ Eigenvalue decomposition of \mathbf{S} is too demanding
- ▶ So, what should we do?

Solution I: First Trick

- ▶ A is a $n \times N$ matrix, then $\mathbf{S} = \mathbf{A}\mathbf{A}^T$ is $n \times n$,
- ▶ but $\mathbf{A}^T\mathbf{A}$ is $N \times N$
- ▶ Trick
 - ▶ Let's do eigenvalue decomposition on $\mathbf{A}^T\mathbf{A}$
 - ▶ $\mathbf{A}^T\mathbf{A}\mathbf{e} = \lambda\mathbf{e} \longrightarrow \mathbf{A}\mathbf{A}^T\mathbf{A}\mathbf{e} = \lambda\mathbf{A}\mathbf{e}$
 - ▶ i.e., if \mathbf{e} is an eigenvector of $\mathbf{A}^T\mathbf{A}$, then $\mathbf{A}\mathbf{e}$ is the eigenvector of $\mathbf{A}\mathbf{A}^T$
 - ▶ and the corresponding eigenvalues are the same
- ▶ Don't forget to normalize $\mathbf{A}\mathbf{e}$

Note: This trick does not fully solved the problem, as we still need to do eigenvalue decomposition on a $N \times N$ matrix, which can be fairly large in practice.

Solution II: Using SVD

- ▶ Instead of doing EVD, doing SVD (singular value decomposition) is easier
- ▶ $\mathbf{A} \in \mathbb{R}^n \times \mathbb{R}^N$
- ▶ $\mathbf{A} = \mathbf{U}\mathbf{\Sigma}\mathbf{V}^T$
 - ▶ $\mathbf{U} \in \mathbb{R}^n \times \mathbb{R}^n$, and $\mathbf{U}^T\mathbf{U} = \mathbf{I}$
 - ▶ $\mathbf{\Sigma} \in \mathbb{R}^n \times \mathbb{R}^N$, is diagonal
 - ▶ $\mathbf{V} \in \mathbb{R}^N \times \mathbb{R}^N$, and $\mathbf{V}^T\mathbf{V} = \mathbf{V}\mathbf{V}^T = \mathbf{I}$

Solution III: Iterative Solution

- ▶ We can design an iterative procedure for finding \mathbf{W} , i.e.,
 $\mathbf{W} \leftarrow \mathbf{W} + \Delta\mathbf{W}$
- ▶ looking at the View of MSE minimization, our cost function:

$$\|\mathbf{x} - \sum_{i=1}^m (\mathbf{w}_i^T \mathbf{x}) \mathbf{w}_i\|^2 = \|\mathbf{x} - \sum_{i=1}^m y_i \mathbf{w}_i\|^2 = \|\mathbf{x} - (\mathbf{W}\mathbf{W}^T) \mathbf{x}\|^2$$

- ▶ we can stop updating if the KKT is met

$$\Delta \mathbf{w}_i = \gamma y_i [\mathbf{x} - \sum_{i=1}^m y_i \mathbf{w}_i]$$

- ▶ Its matrix form is: \leftarrow subspace learning algorithm

$$\Delta \mathbf{W} = \gamma (\mathbf{x}\mathbf{x}^T \mathbf{W} - \mathbf{W}\mathbf{W}^T \mathbf{x}\mathbf{x}^T \mathbf{W})$$

- ▶ **Two issues:**
 - ▶ The orthogonality is not reinforced
 - ▶ Slow convergence

Solution IV: PAST

To speed up the iteration, we can use recursive least squares (RLS). We can consider the following cost function

$$J(t) = \sum_{i=1}^t \beta_{t-i} \|\mathbf{x}(i) - \mathbf{W}(t)\mathbf{y}(i)\|^2$$

where β is the forgetting factor.

\mathbf{W} can be solved recursively by the following PAST algorithm

1. $\mathbf{y}(t) = \mathbf{W}^T(t-1)\mathbf{x}(t)$
2. $\mathbf{h}(t) = \mathbf{P}(t-1)\mathbf{y}(t)$
3. $\mathbf{m}(t) = \mathbf{h}(t)/(\beta + \mathbf{y}^T(t)\mathbf{h}(t))$
4. $\mathbf{P}(t) = \frac{1}{\beta} \text{Tri}[\mathbf{P}(t-1) - \mathbf{m}(t)\mathbf{h}^T(t)]$
5. $\mathbf{e}(t) = \mathbf{x}(t) - \mathbf{W}(t-1)\mathbf{y}(t)$
6. $\mathbf{W}(t) = \mathbf{W}(t-1) + \mathbf{e}(t)\mathbf{m}^T(t)$

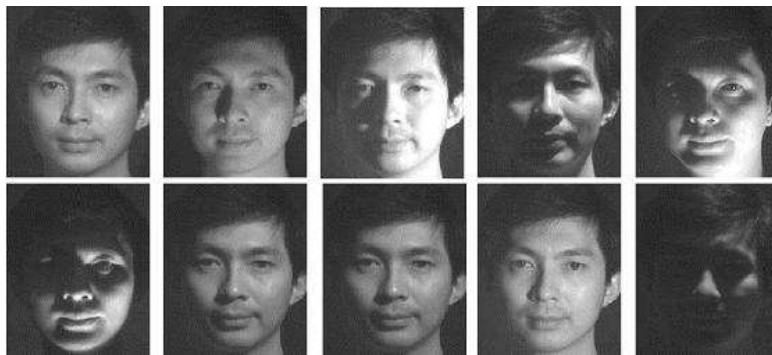
Outline

Principal Component Analysis

Linear Discriminant Analysis

Comparison between PCA and LDA

Face Recognition: Does PCA work well?

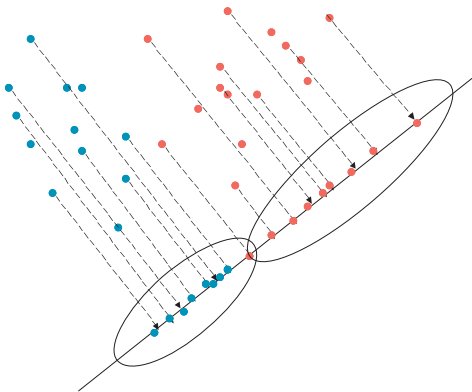


- ▶ The same face under different illumination conditions
- ▶ What does PCA capture?
- ▶ Is this what we really want?

From Descriptive to Discriminative

- ▶ PCA extracts features (or components) that well describe the pattern
- ▶ Are they necessarily good for distinguishing between classes and separating patterns?
- ▶ Examples?
- ▶ We need discriminative features.
- ▶ Supervision (or labeled training data) is needed.
- ▶ The issues are:
 - ▶ How do we define the discriminant and separability between classes?
 - ▶ How many features do we need?
 - ▶ How do we maximizing the separability?
- ▶ Here, we give an example of linear discriminant analysis.

Linear Discriminant Analysis



- ▶ Finding an optimal linear projection \mathbf{W}
- ▶ Catches major difference between classes and discount irrelevant factors
- ▶ In the projected discriminative subspace, data are clustered

Within-class and Between-class Scatters

We have two sets of labeled data: $\mathcal{D}_1 = \{\mathbf{x}_1, \dots, \mathbf{x}_{n_1}\}$ and $\mathcal{D}_2 = \{\mathbf{x}_1, \dots, \mathbf{x}_{n_2}\}$. Let's define some terms:

- ▶ The centers of two classes, $\mathbf{m}_i = \frac{1}{n_i} \sum_{\mathbf{x} \in \mathcal{D}_i} \mathbf{x}_i$
- ▶ Data scatter by definition

$$\mathbf{S} = \sum_{\mathbf{x} \in \mathcal{D}} (\mathbf{x} - \mathbf{m})(\mathbf{x} - \mathbf{m})^T$$

- ▶ **Within-class scatter:**

$$\mathbf{S}_w = \mathbf{S}_1 + \mathbf{S}_2$$

- ▶ **Between-class scatter:**

$$\mathbf{S}_b = (\mathbf{m}_1 - \mathbf{m}_2)(\mathbf{m}_1 - \mathbf{m}_2)^T$$

Fisher Liner Discriminant

Input: We have two sets of labeled data: $\mathcal{D}_1 = \{\mathbf{x}_1, \dots, \mathbf{x}_{n_1}\}$ and $\mathcal{D}_2 = \{\mathbf{x}_1, \dots, \mathbf{x}_{n_2}\}$.

Output: We want to find a 1-d linear projection \mathbf{w} that maximizes the separability between these two classes.

- ▶ Projected data: $\mathcal{Y}_1 = \mathbf{w}^T \mathcal{D}_1$ and $\mathcal{Y}_2 = \mathbf{w}^T \mathcal{D}_2$
- ▶ Projected class centers: $\tilde{\mathbf{m}}_i = \mathbf{w}^T \mathbf{m}_i$
- ▶ Projected within-class scatter (it is a scalar in this case)

$$\tilde{\mathbf{S}}_w = \mathbf{w}^T \mathbf{S}_w \mathbf{w} \quad \textit{prove it!}$$

- ▶ Projected between-class scatter (it is a scalar in this case)

$$\tilde{\mathbf{S}}_b = \mathbf{w}^T \mathbf{S}_b \mathbf{w} \quad \textit{prove it!}$$

- ▶ Fisher Linear Discriminant

$$J(\mathbf{w}) = \frac{|\tilde{m}_1 - \tilde{m}_2|^2}{\tilde{\mathbf{S}}_1 + \tilde{\mathbf{S}}_2} = \frac{|\tilde{\mathbf{S}}_b|}{|\tilde{\mathbf{S}}_w|} = \frac{\mathbf{w}^T \mathbf{S}_b \mathbf{w}}{\mathbf{w}^T \mathbf{S}_w \mathbf{w}}$$

Rayleigh Quotient

Theorem

$f(\lambda) = \|\mathbf{Ax} - \lambda\mathbf{Bx}\|_B$ where $\|\mathbf{z}\|_B \triangleq \mathbf{z}^T \mathbf{B}^{-1} \mathbf{z}$ is minimized by the Rayleigh quotient

$$\lambda = \frac{\mathbf{x}^T \mathbf{Ax}}{\mathbf{x}^T \mathbf{Bx}}$$

Proof.

$$\begin{aligned} \frac{\partial f(\lambda)}{\partial \lambda} &= (\mathbf{Bx})^T (\mathbf{Bz}) = \mathbf{x}^T \mathbf{B}^T \mathbf{B}^{-1} \mathbf{z} \\ &= \mathbf{x}^T (\mathbf{Ax} - \lambda \mathbf{Bx}) = \mathbf{x}^T \mathbf{Ax} - \lambda \mathbf{x}^T \mathbf{Bx} \end{aligned}$$

setting it to zero to see the result clearly.



Optimizing Fish Discriminant

Theorem

$J(\mathbf{w}) = \frac{\mathbf{w}^T \mathbf{S}_b \mathbf{w}}{\mathbf{w}^T \mathbf{S}_w \mathbf{w}}$ is maximized when

$$\mathbf{S}_b \mathbf{w} = \lambda \mathbf{S}_w \mathbf{w}$$

Proof.

Let $\mathbf{w}^T \mathbf{S}_w \mathbf{w} = c \neq 0$. We can construct the Lagrangian as

$$L(\mathbf{w}, \lambda) = \mathbf{w}^T \mathbf{S}_b \mathbf{w} - \lambda(\mathbf{w}^T \mathbf{S}_w \mathbf{w} - c)$$

Then KKT is

$$\frac{\partial L(\mathbf{w}, \lambda)}{\partial \mathbf{w}} = \mathbf{S}_b \mathbf{w} - \lambda \mathbf{S}_w \mathbf{w}$$

It is clearly that

$$\mathbf{S}_b \mathbf{w}^* = \lambda \mathbf{S}_w \mathbf{w}^*$$



An Efficient Solution

- ▶ A naive solution is $\mathbf{S}_w^{-1} \mathbf{S}_b \mathbf{w} = \lambda \mathbf{w}$
- ▶ Then we can do EVD on $\mathbf{S}_w^{-1} \mathbf{S}_b$, which needs some computation
- ▶ Is there a more efficient way?
- ▶ Facts:
 - ▶ $\mathbf{S}_b \mathbf{w}$ is along the direction of $\mathbf{m}_1 - \mathbf{m}_2$. why?
 - ▶ we don't care about λ the scalar factor
- ▶ So we can easily figure out the direction of \mathbf{w} by

$$\mathbf{w} = \mathbf{S}_w^{-1}(\mathbf{m}_1 - \mathbf{m}_2)$$

- ▶ **Note:** $\text{rank}(\mathbf{S}_b) = 1$

Multiple Discriminant Analysis

Now, we have c number of classes:

- ▶ within-class scatter $\mathbf{S}_w = \sum_{i=1}^c \mathbf{S}_i$ as before
- ▶ between-class scatter is a bit different from 2-class

$$\mathbf{S}_b \triangleq \sum_{i=1}^c n_i (\mathbf{m}_i - \mathbf{m})(\mathbf{m}_i - \mathbf{m})^T$$

- ▶ total scatter

$$\mathbf{S}_t \triangleq \sum_{\mathbf{x}} (\mathbf{x} - \mathbf{m})(\mathbf{x} - \mathbf{m})^T = \mathbf{S}_w + \mathbf{S}_b$$

- ▶ MDA is to find a subspace with bases \mathbf{W} that maximizes

$$J(\mathbf{W}) = \frac{|\tilde{\mathbf{S}}_b|}{|\tilde{\mathbf{S}}_w|} = \frac{|\mathbf{W}^T \mathbf{S}_b \mathbf{W}|}{|\mathbf{W}^T \mathbf{S}_w \mathbf{W}|}$$

The Solution to MDA

- ▶ The solution is obtained by G-EVD

$$\mathbf{S}_b \mathbf{w}_i = \lambda_i \mathbf{S}_w \mathbf{w}_i$$

where each \mathbf{w}_i is a generalized eigenvector

- ▶ In practice, what we can do is the following
 - ▶ find the eigenvalues as the root of the characteristic polynomial

$$|\mathbf{S}_b - \lambda_i \mathbf{S}_w| = 0$$

- ▶ for each λ_i , solve \mathbf{w}_i from

$$(\mathbf{S}_b - \lambda_i \mathbf{S}_w) \mathbf{w}_i = 0$$

- ▶ **Note:** \mathbf{W} is not unique (up to rotation and scaling)
- ▶ **Note:** $\text{rank}(\mathbf{S}_b) \leq (c - 1)$ (why?)

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Comparison between PCA and LDA

The Relation between PCA and LDA

