

# Foundations of Machine Learning

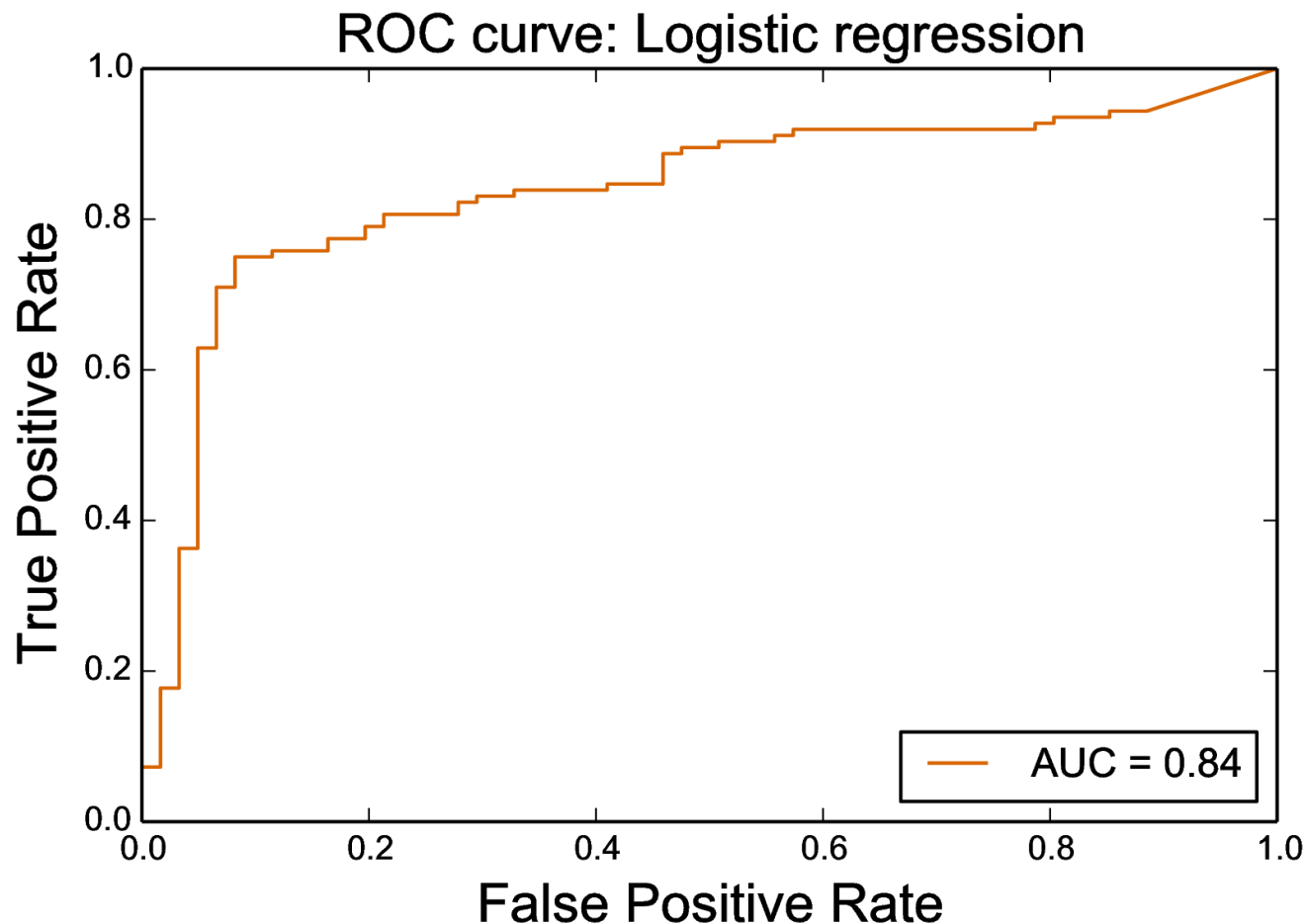
## CentraleSupélec — Fall 2016

### 6. Regularized linear regression

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# Logistic regression on the Endometrium vs Uterus data



# Learning objectives

- Understand **regularization** as a means to control model complexity.
- Define **Lasso**, **ridge regression**, **elastic net**.
- Understand the role of the  **$l_1$  and  $l_2$  norms** in regularization
- Interpret **solution paths** for Lasso and ridge regression.

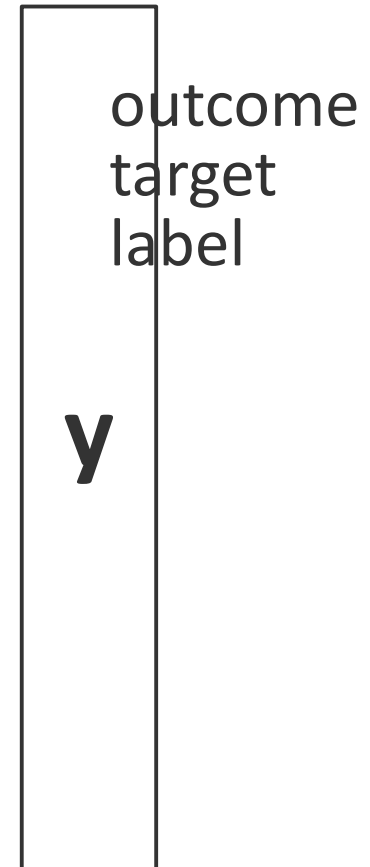
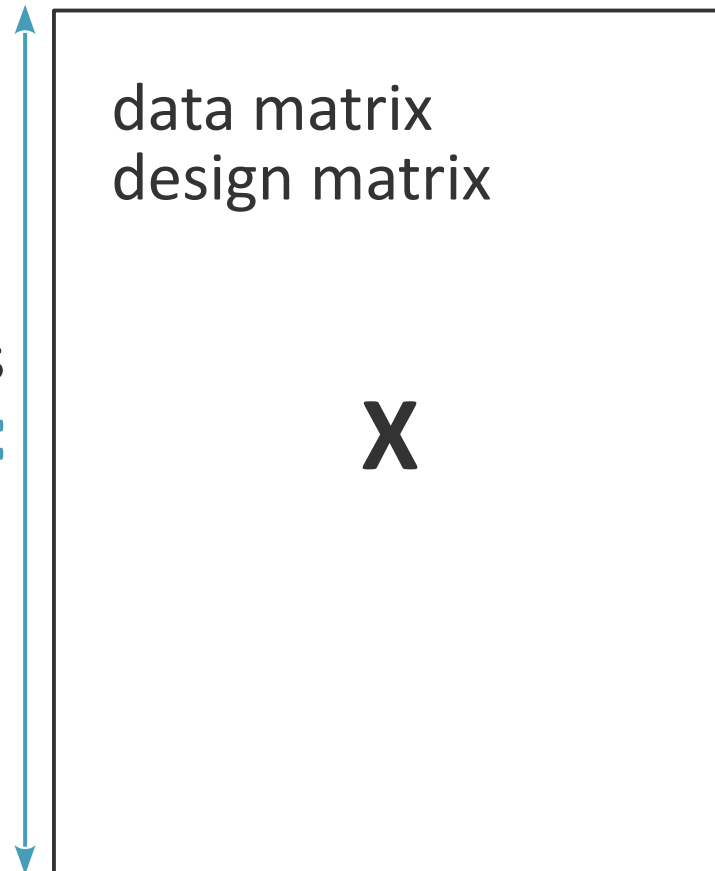
# Regression setting

$$x_j^i \in \mathbb{R}$$

$$y^i \in \mathbb{R}$$

features variables  
descriptors regressors  
attributes **p**

observations  
samples **n**  
data points

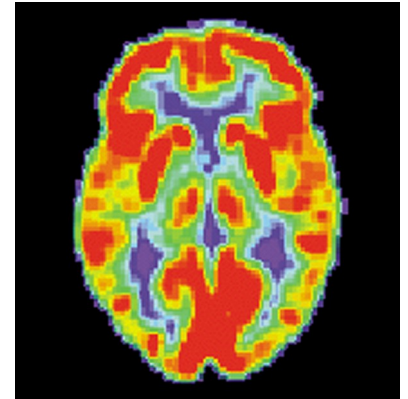


# Large p, small n

E.g.

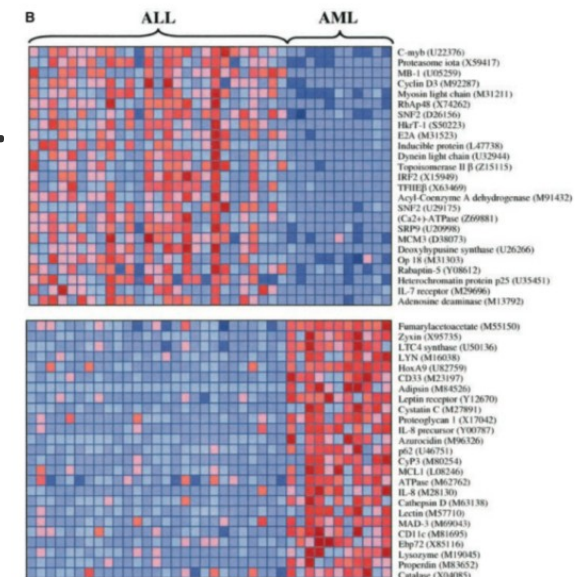
- **neuroimaging**

thousands of brain regions / pixels / voxels  
much fewer patients

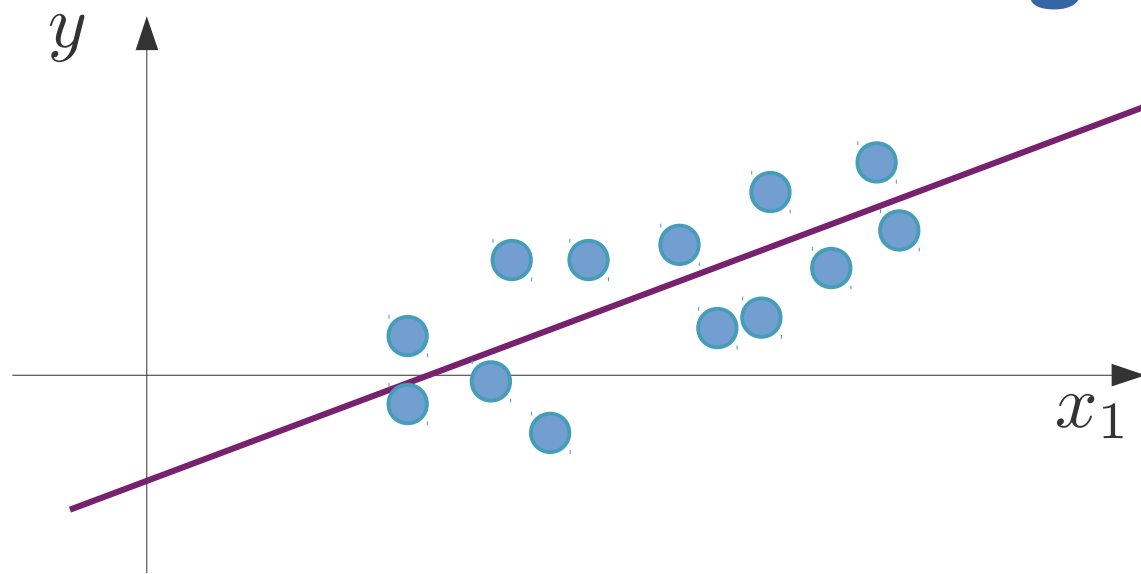


- genetics and genomics

thousands of genes, millions of SNPs...  
usually, at best thousands of patients



# Linear regression



$$\mathbf{x} \in \mathbb{R}^p$$
$$f(\mathbf{x}|\beta) = \sum_{j=1}^p \beta_j x_j + \beta_0$$

**Least-squares fit** (equivalent to MLE under the assumption of Gaussian noise):

$$\hat{\beta} = \arg \min_{\beta} (y - X\beta)^{\top} (y - X\beta) = (X^{\top} X)^{-1} X^{\top} y$$

The solution is uniquely defined when **n > p** and **X<sup>⊤</sup>X** invertible.

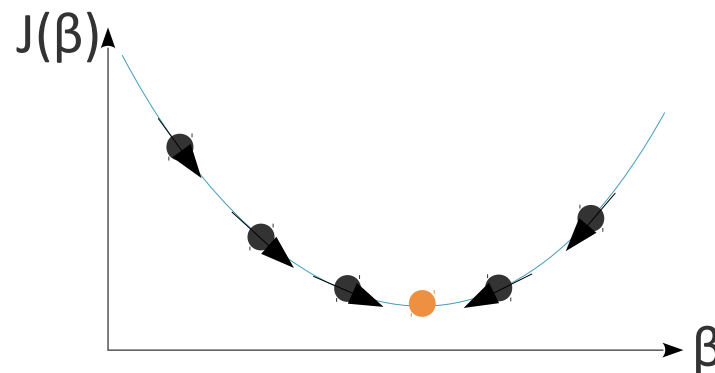
# When $X^T X$ not invertible

$$(X^T X)\hat{\beta} = X^T y$$

- Pseudo-inverse
- Linear system of  $p$  equations:

## Numerical methods

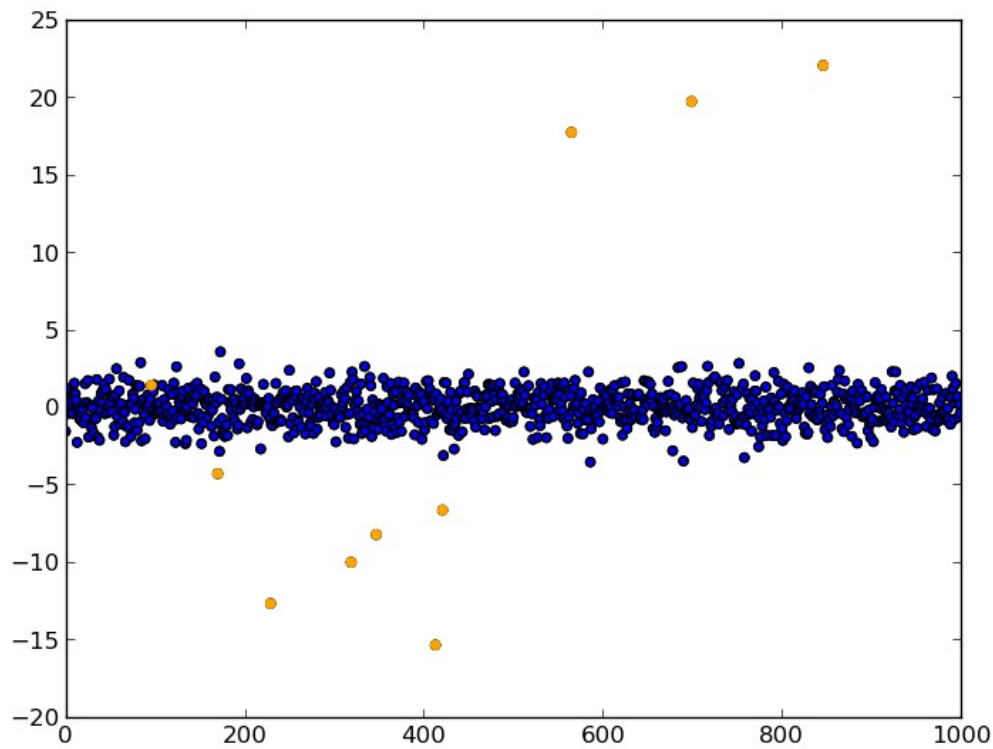
- Gaussian elimination
- LU decomposition
- Gradient descent



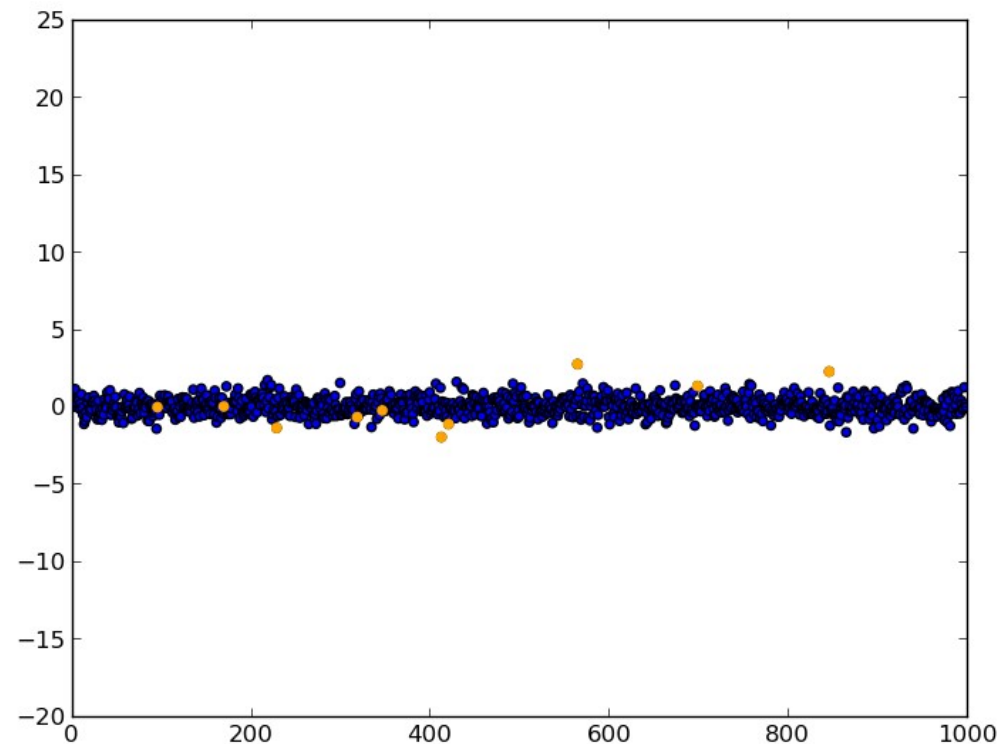
# Linear regression when $p \gg n$

Simulated data:  $p=1000$ ,  $n=100$ , 10 causal features

## True coefficients



## Predicted coefficients





# Advantages of least-squares fit

- Unbiased  $E[\hat{\beta}] = \beta$
- Explicit form
- Computational time?

# Advantages of least-squares fit

- **Unbiased**  $E[\hat{\beta}] = \beta$
- **Explicit form**
- **Computational time:**  $O(np^2 + p^3)$

compute  $X^T X$     invert  $X^T X$



computation of  $X^T y$ :  $O(np)$

computation of  $(X^T X)^{-1} X^T y$ :  $O(np)$

# Cons of least-squares fit

- **Multicollinearity** leads to high variance of the estimator
- Requires  $n > p$
- Prediction error increases linearly as a function of  $p$
- Hard to **interpret** when  $p$  is large

**Would prefer a small subset with strong effects.**

# Regularization

# Regularization

- Minimize

SSE +  $\lambda$  penalty on model complexity

- **Biased estimator** when  $\lambda \neq 0$ .
- Trade bias for a smaller variance.
- $\lambda$  can be set by cross-validation.
- Simpler model  $\approx$  fewer parameters  
→ **shrinkage**: drive the coefficients of the parameters towards 0.

# Ridge regression

# Ridge regression

- **Sum-of-squares penalty**

$$\hat{\beta}_{\text{ridge}} = \arg \min_{\beta} ||y - X\beta||_2^2 + \lambda ||\beta||_2^2$$

- **Compute the ridge regression estimator.**

# Ridge regression

- **Sum-of-squares penalty**

$$\hat{\beta}_{\text{ridge}} = \arg \min_{\beta} \|y - X\beta\|_2^2 + \lambda \|\beta\|_2^2$$

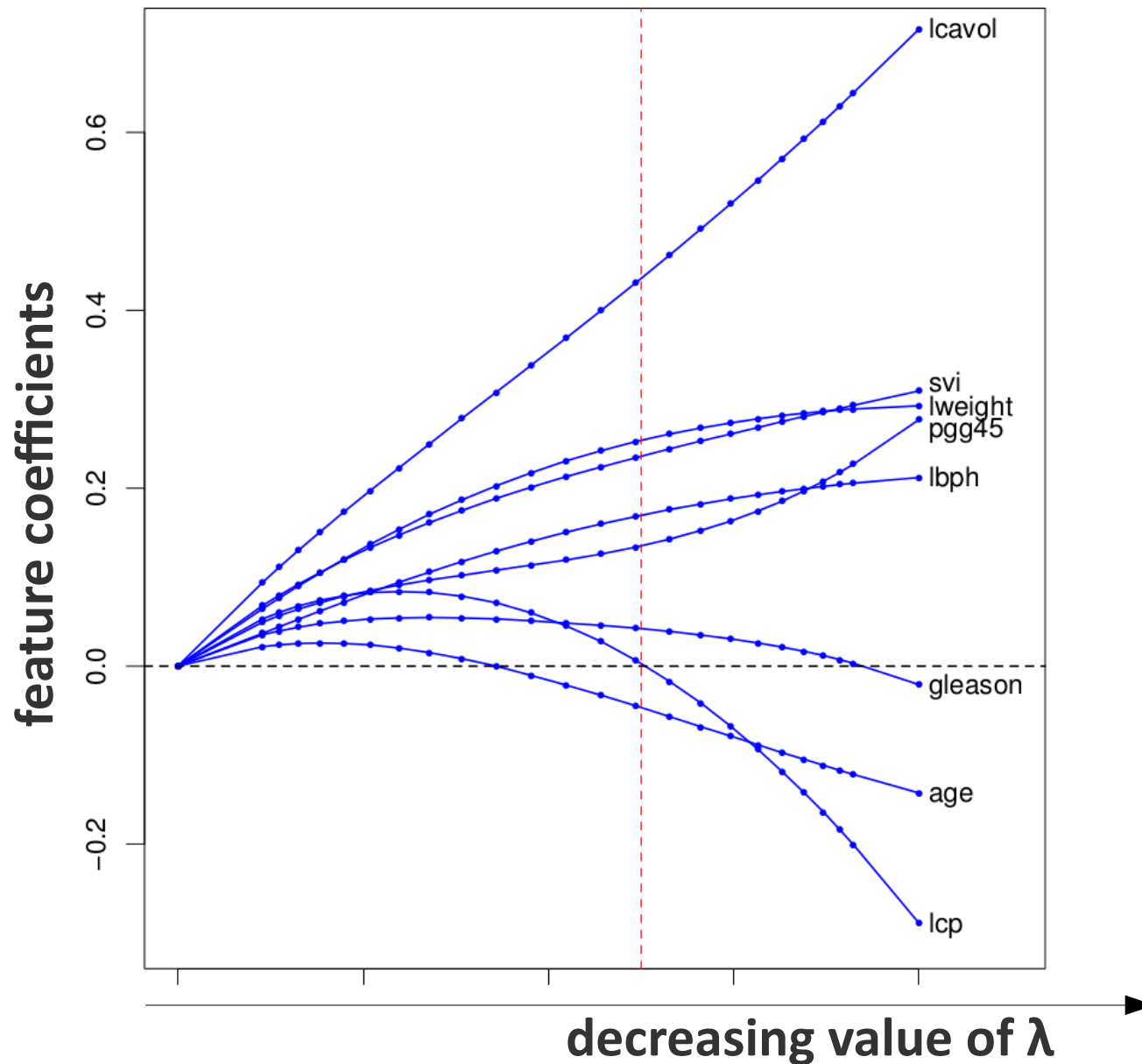
- **Compute the ridge regression estimator.**

$$\hat{\beta}_{\text{ridge}} = (X^\top X + \lambda I)^{-1} X^\top y$$

if  $(X^\top X + \lambda I)$  invertible.



# Ridge regression solution path



# Standardization

- What happens if we multiply  $x_j$  by a constant?
  - For standard linear regression
  - For ridge regression

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- What happens if we multiply  $x_j$  by a constant?

- For **standard linear regression**:

$$\hat{\beta}_j \rightarrow \frac{1}{c} \hat{\beta}_j$$

- For **ridge regression**:

Not so clear, because of the penalization term  $\lambda \beta_j^2$

- Need to **standardize** the features

$$\tilde{x}_j^i = \frac{x_j^i}{\sqrt{\frac{1}{n} \sum_{i=1}^n (x_j^i - \bar{x}_j)^2}}$$

average value of  $x_j$

# Ridge regression

- **Grouped selection:**
  - correlated variables get similar weights
  - identical variables get identical weights
- Ridge regression shrinks coefficients towards 0 but does not result in a **sparse model**.
- **Sparsity:**
  - many coefficients get a weight of 0
  - they can be eliminated from the model.

# Lasso

# Lasso

- **L1 penalty**

$$||\beta||_1 = \sum_{j=1}^p |\beta_j|$$

$$\hat{\beta}_{\text{lasso}} = \arg \min_{\beta} ||y - X\beta||_2^2 + \lambda ||\beta||_1$$

- aka **basis pursuit** (signal processing)
- no closed-form solution
- **quadratic programming** problem: equivalent to

$$\hat{\beta}_{\text{lasso}} = \arg \min_{\beta} ||y - X\beta||_2^2 \quad \text{s.t.} \quad ||\beta||_1 \leq t$$

for a unique one-to-one match between  $t$  and  $\lambda$ .

**QP:** maximize a quadratic form under linear constraints.

# Equivalence between the formulations

$$\hat{\beta}_{\text{lasso}} = \arg \min_{\beta} ||y - X\beta||_2^2 \quad \text{s.t.} \quad ||\beta||_1 \leq t$$

- **minimize  $f(\beta)$  under the constraint  $g(\beta) \leq 0$**

$$f(\beta) = ||y - X\beta||_2^2$$

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**Case 1:** the unconstrained minimum lies in the **feasible region**.  $\{\beta : g(\beta) \leq 0\}$



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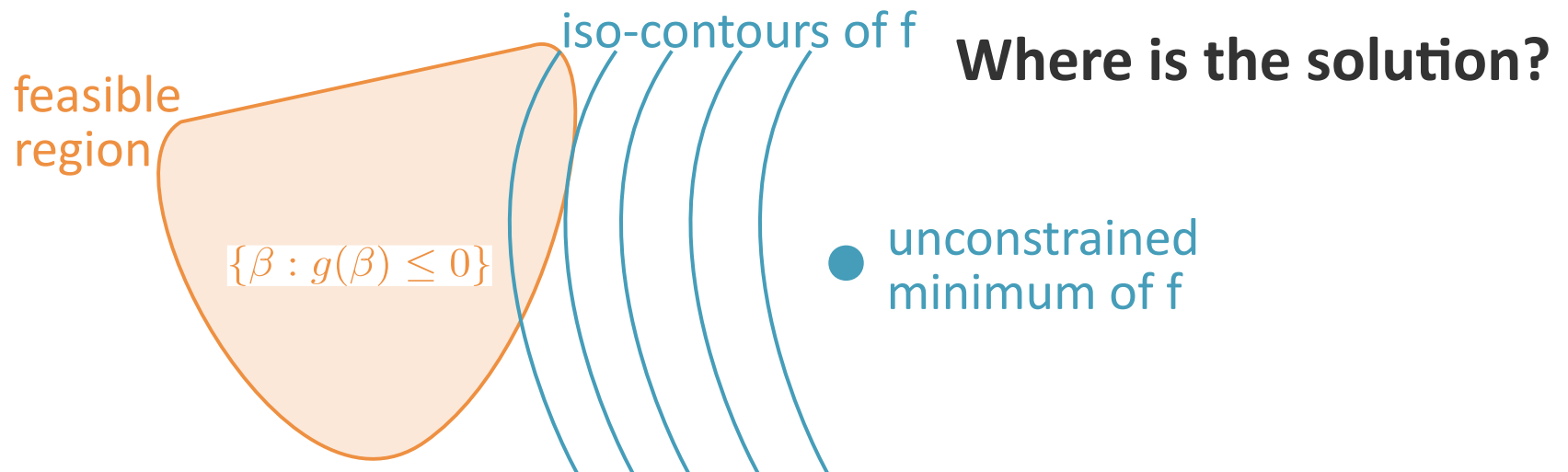
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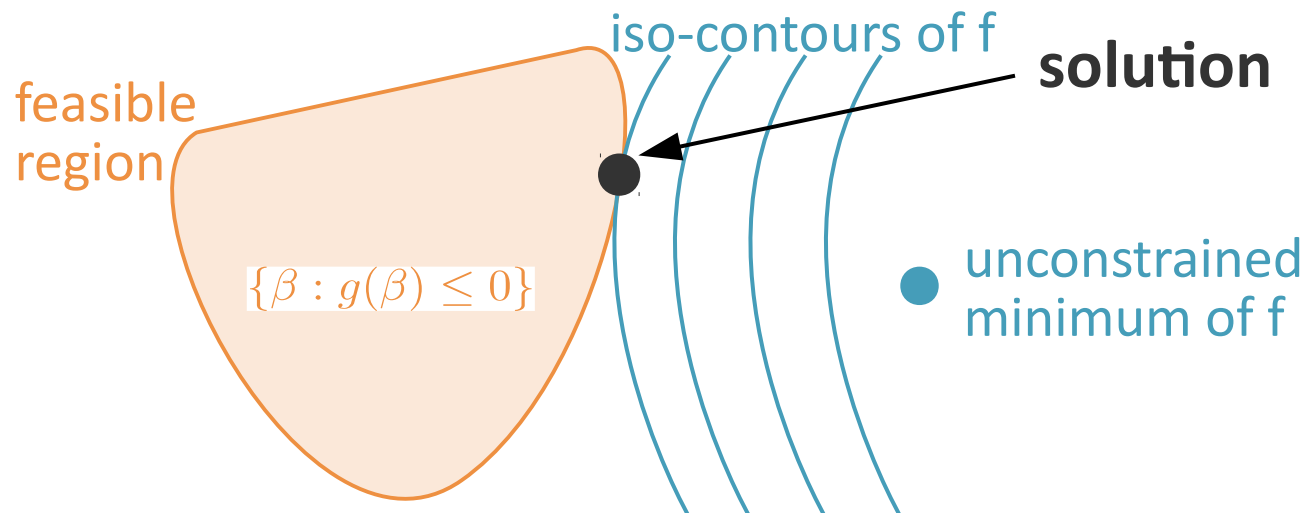
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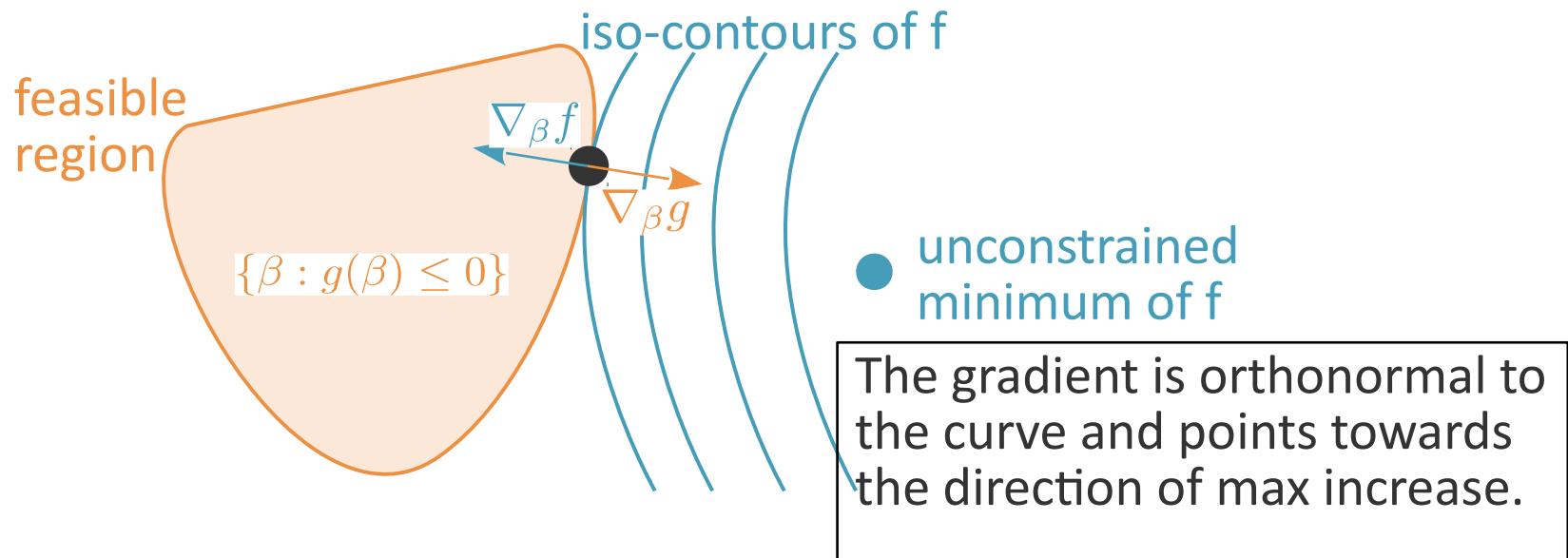
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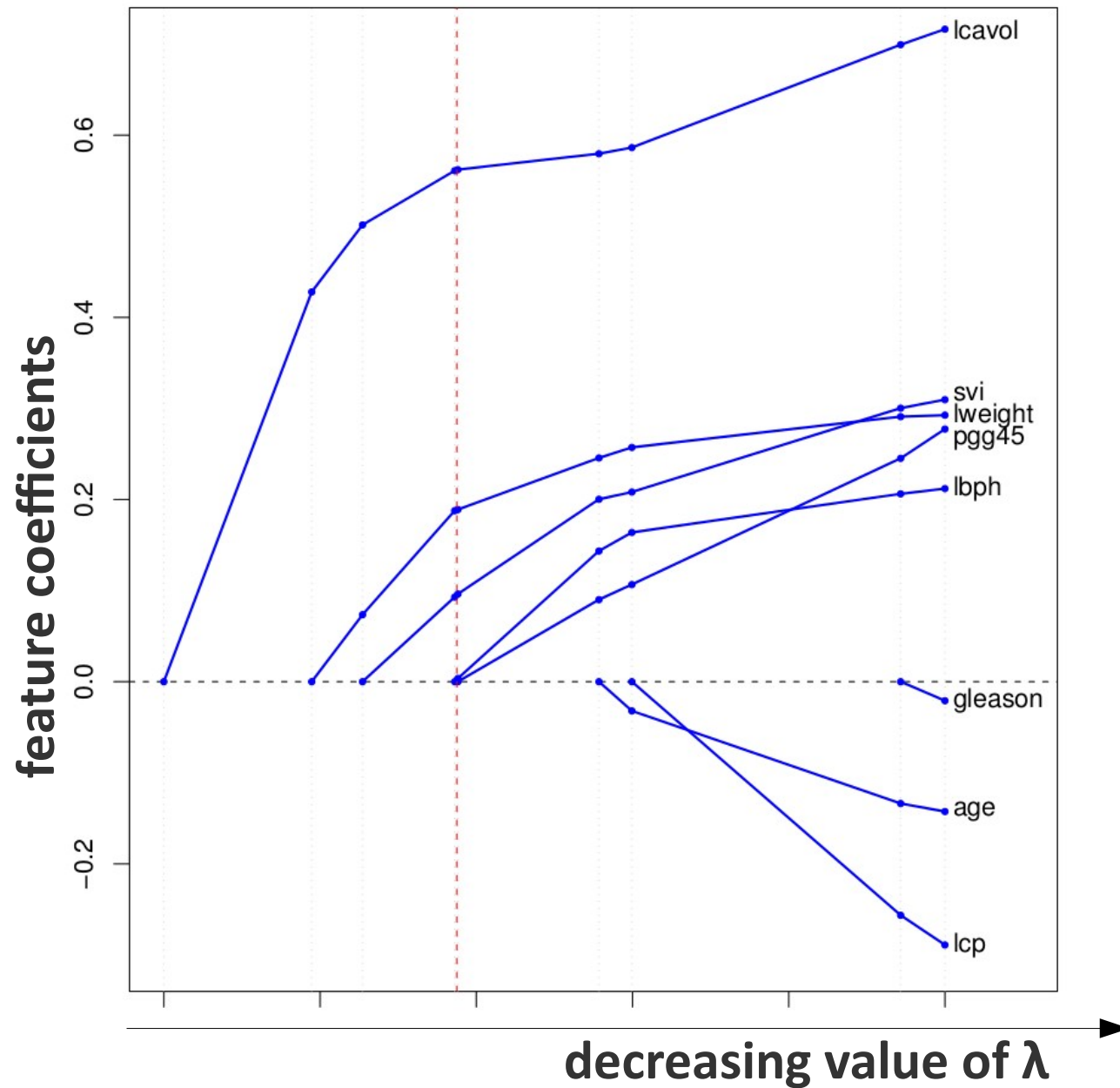
**Case 1:** the unconstrained minimum lies in the feasible region.

**Case 2:** it does not. Then it lies at a point where the feasible region and the iso-contours of  $f$  are tangent and the gradients are in opposite directions.

The **Lagrangian  $f(\beta) + \lambda g(\beta)$  must be minimized ( $\lambda \geq 0$ )**

$$\hat{\beta}_{\text{lasso}} = \arg \min_{\beta} ||y - X\beta||_2^2 + \lambda ||\beta||_1$$

# Lasso solution path



# Forward stepwise regression

- Build model **sequentially**, adding one variable at a time
  - Start with the intercept
  - At each step, add the variable that **most improves the fit**
  - **Stop when**  $||\beta||_1 \leq t$
- Greedy solution

# Least Angle Regression

At each step, add “only as much of a variable as needed”

1. Standardize the predictors to have mean zero and unit norm. Start with the residual  $\mathbf{r} = \mathbf{y} - \bar{\mathbf{y}}$ ,  $\beta_1, \beta_2, \dots, \beta_p = 0$ .

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3. Move  $\beta_j$  from 0 towards its least-squares coefficient  $\langle \mathbf{x}_j, \mathbf{r} \rangle$ , until some other competitor  $\mathbf{x}_k$  has as much correlation with the current residual as does  $\mathbf{x}_j$ .

$$\begin{aligned}\beta_j &\leftarrow \beta_j + \alpha \frac{1}{\sum_{i=1}^n (x_j^i)^2} \sum_{i=1}^n x_j^i r^i \\ &= \beta_j + \alpha (x_j^\top x_j)^{-1} x_j^\top r \\ &= \beta_j + \alpha \langle x_j^\top, x_j \rangle^{-1} \langle x_j, r \rangle\end{aligned}$$

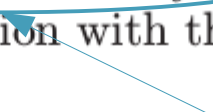
step size

$$r = (y - \bar{y}) - \beta_j x_j$$

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$$r = (y - \bar{y}) - \beta_j x_j - \beta_k x_k$$

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**Maximum number of steps:**  
 **$\max(n-1, p)$**

# Elastic Net

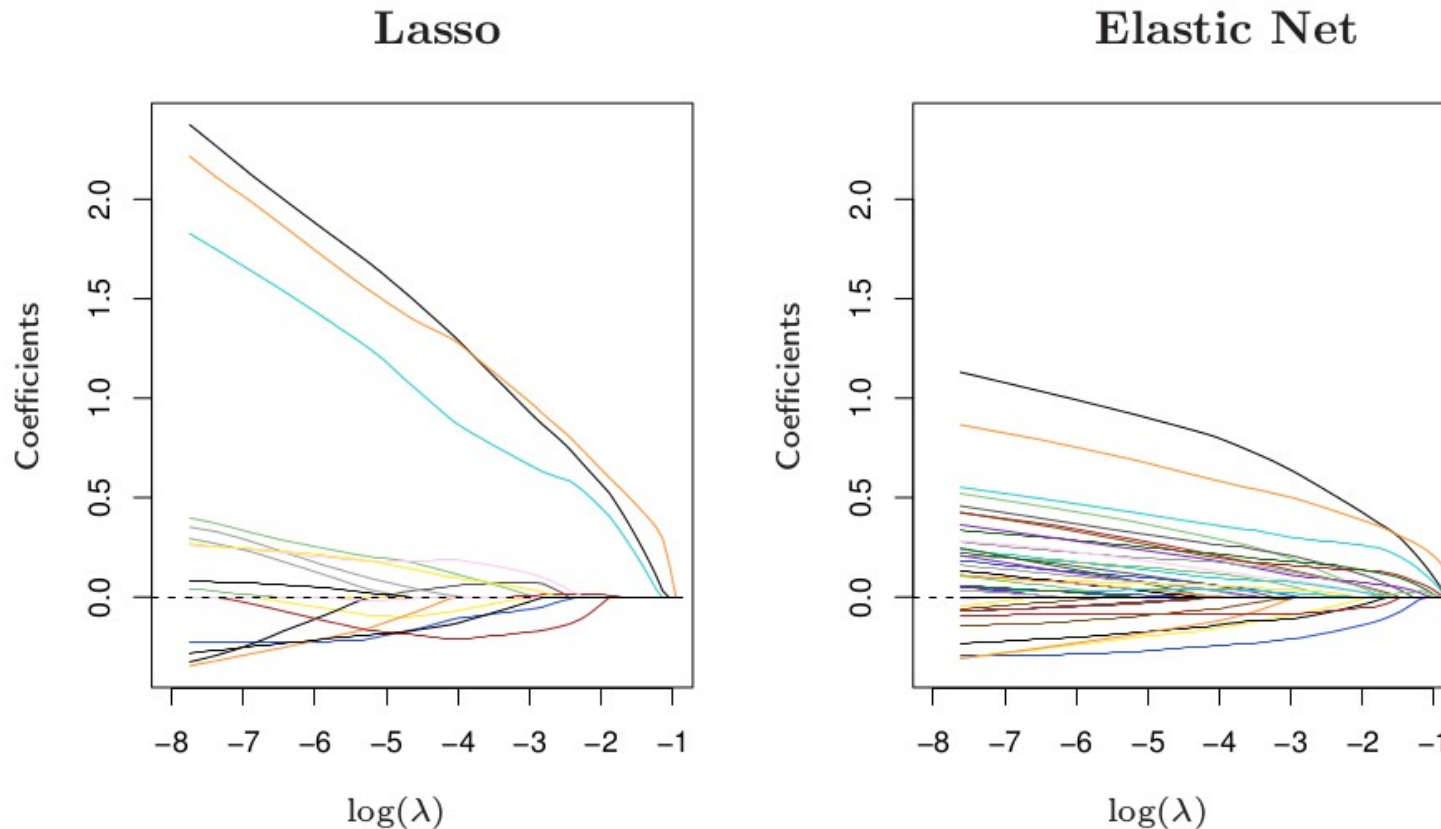
# Elastic Net

- **Combine lasso** and **ridge regression**

$$\hat{\beta}_{\text{enet}} = \arg \min_{\beta} ||y - X\beta||_2^2 + \lambda (\alpha ||\beta||_2^2 + (1 - \alpha) ||\beta||_1)$$

- **Select variables** like the lasso.
- **Shrinks together coefficients of correlated variables** like the ridge regression.

# E.g. Leukemia data



Elastic Net results in more non-zero coefficients than Lasso, but with smaller amplitudes.



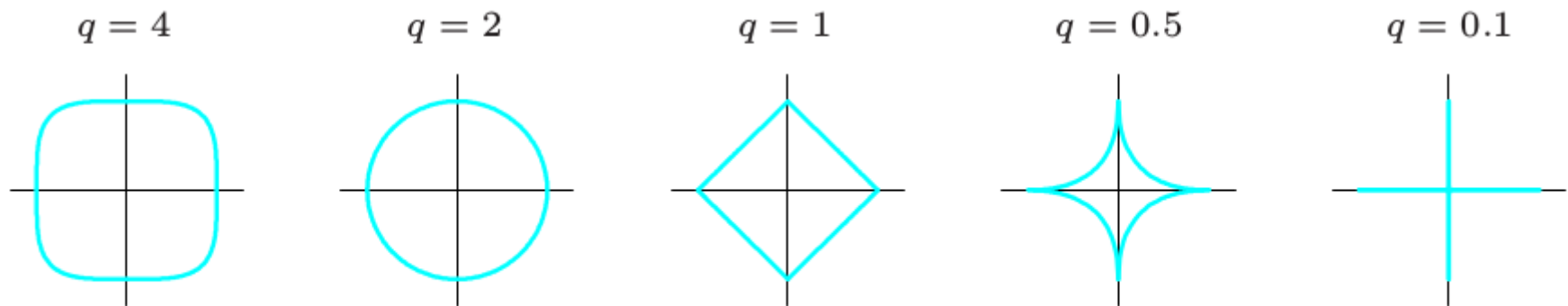
# Lq-norm regularization

# Lq-norm regularization

$$\hat{\beta} = \arg \min_{\beta} ||Y - X\beta||_2^2 + \lambda ||\beta||_q^q \quad ||\beta||_q = \left( \sum_{j=1}^p |\beta_j|^q \right)^{1/q}$$

Equivalently:

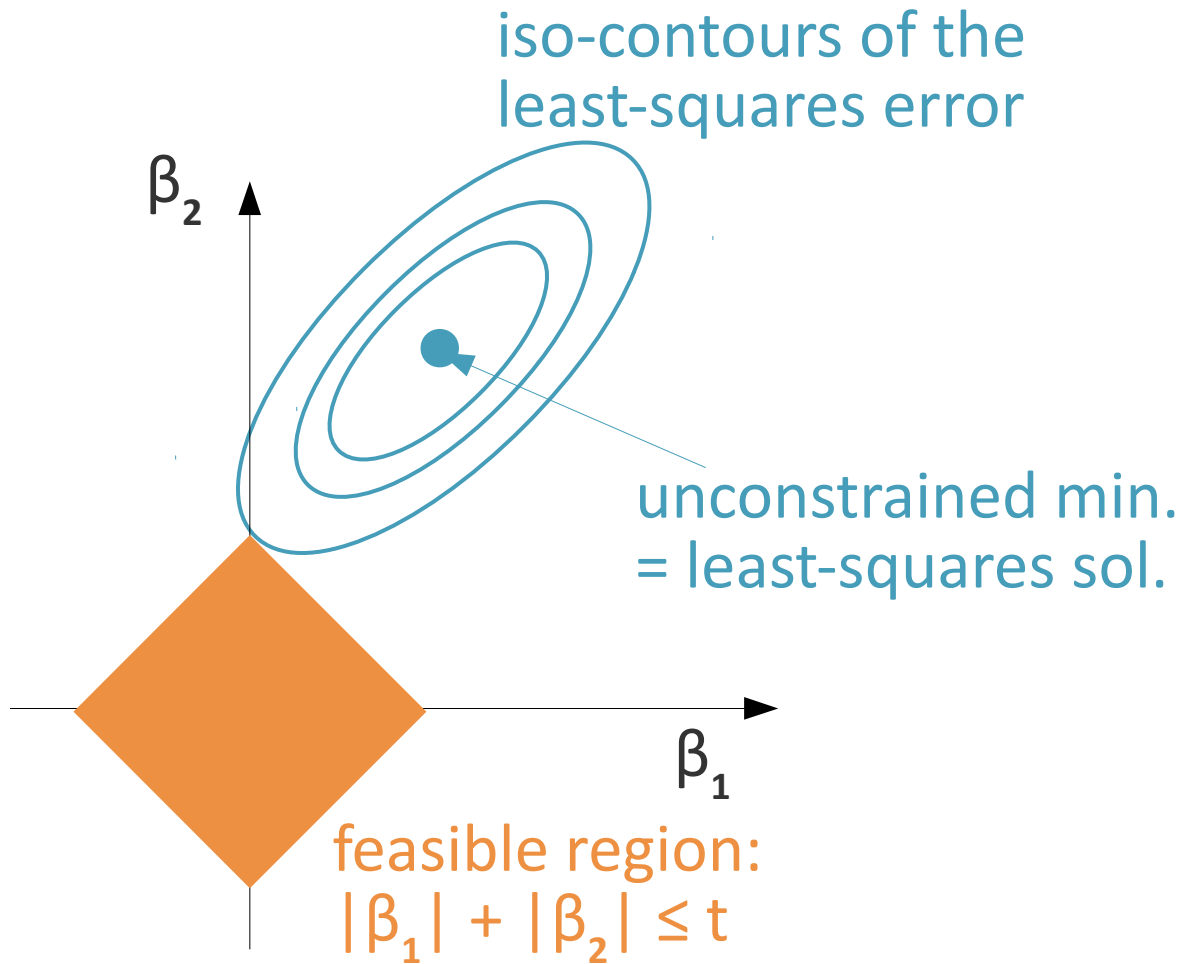
$$\hat{\beta} = \arg \min_{\beta} ||Y - X\beta||_2^2 \text{ s. t. } ||\beta||_q^q \leq s$$



**FIGURE 3.12.** Contours of constant value of  $\sum_j |\beta_j|^q$  for given values of  $q$ .

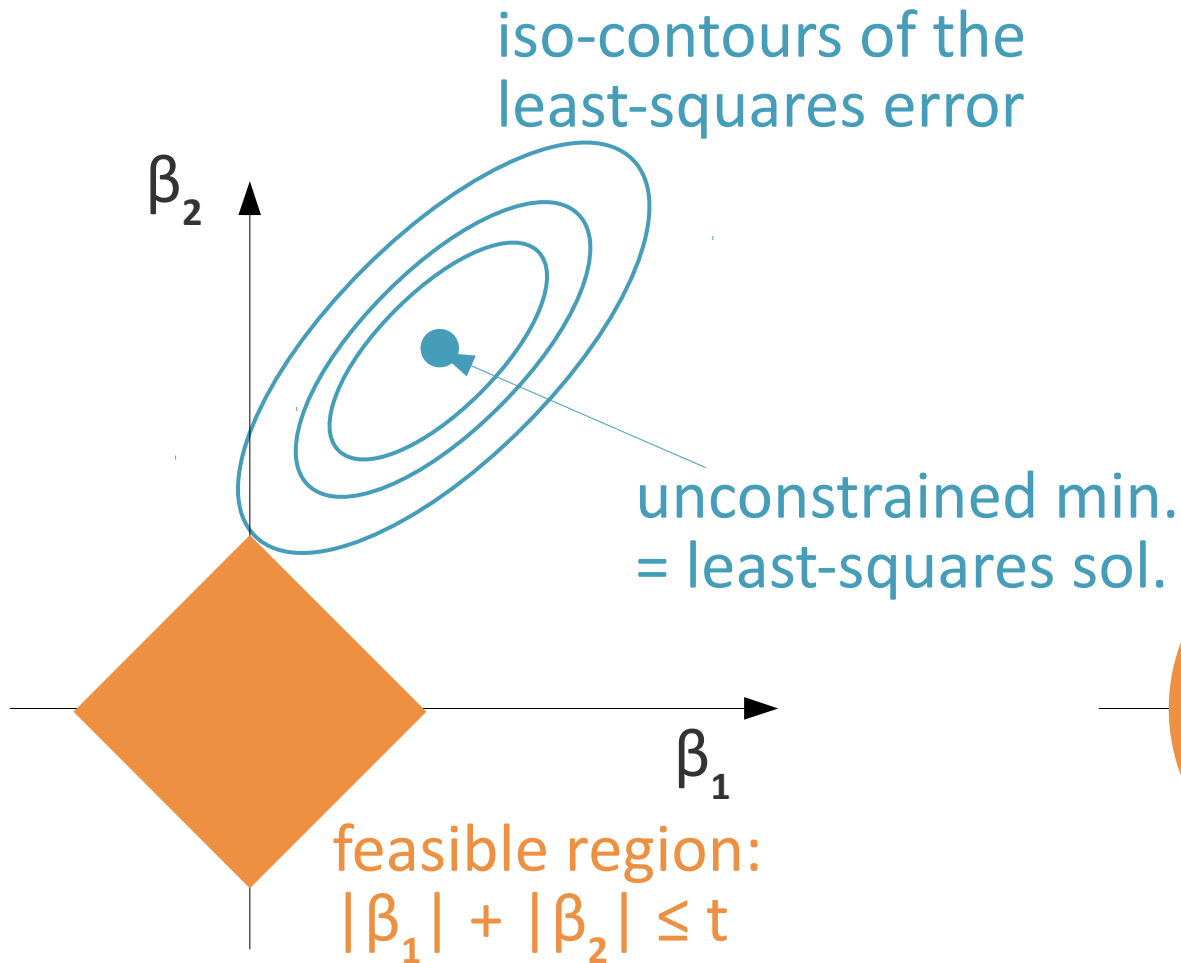
# Lasso vs. ridge

## L1 norm

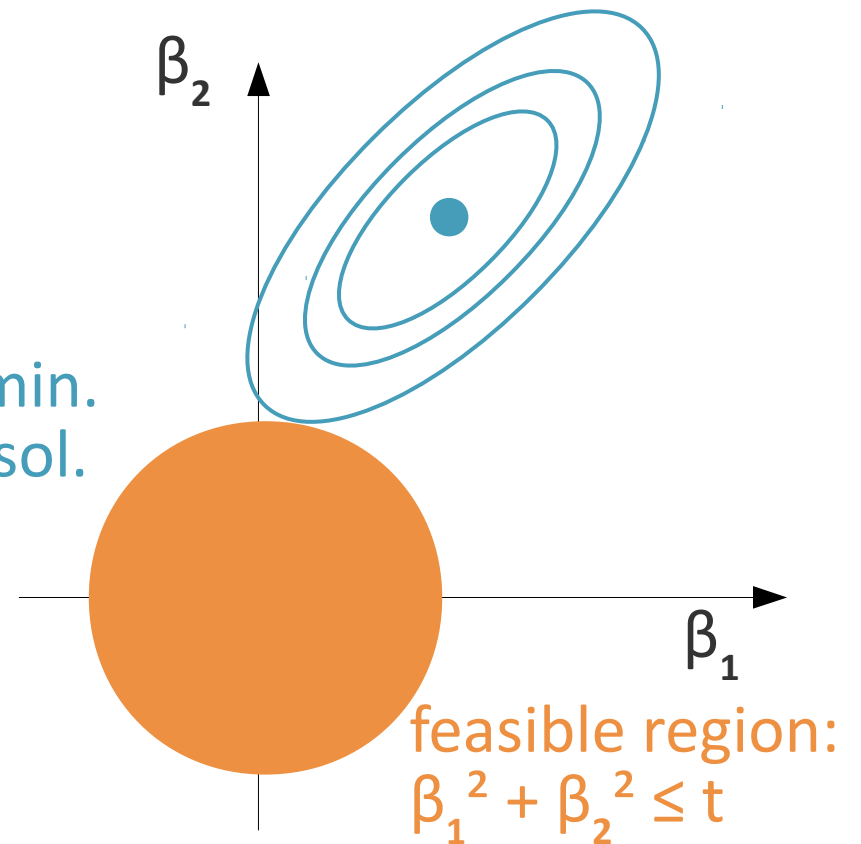


# Lasso vs. ridge

## L1 norm



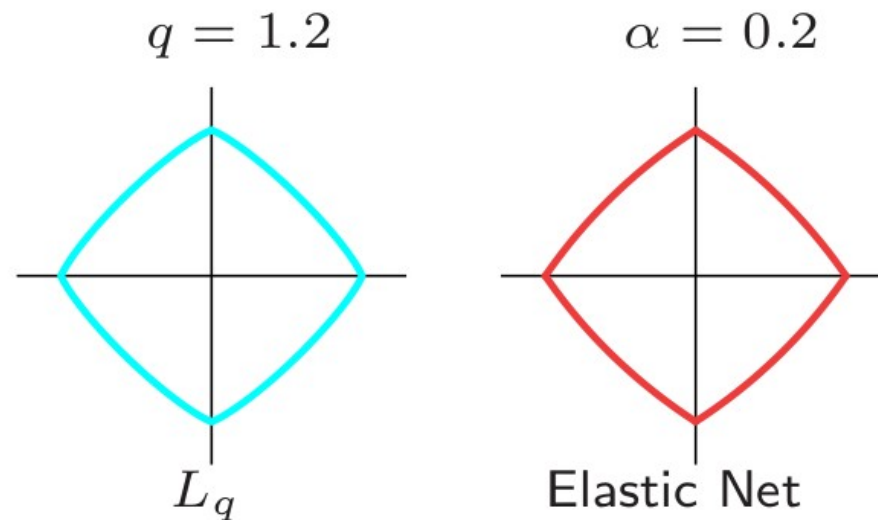
## L2 norm



# Elastic net

- Elastic penalty

$$\hat{\beta} = \arg \min_{\beta} ||y - X\beta||_2^2 + \lambda (\alpha ||\beta||_2^2 + (1 - \alpha) ||\beta||_1)$$



# Structured regularization

# Group lasso

Use  $K$  predefined groups of variables that are known to “work” together and expected to be either all active or all inactive together.

E.g.

- genes belonging to the same biological pathway.

$$\hat{\beta} = \arg \min_{\beta} \left\| y - \sum_{k=1}^K X_k \beta_k \right\|_2^2 + \lambda \sum_{k=1}^K \sqrt{p_k} \|\beta_k\|_2$$

Features belonging to group  $k$

Size of group  $k$

# Other examples of structured penalties

- **Overlapping groups**

Jacob et al. (2009). Group lasso with overlap and graph lasso. *ICML*.

- **Graphs**

Li & Li (2010). Variable selection and regression analysis for graph-structured covariates with an application to genomics. *Ann. App. Stats.*

- **Trees**

Zhao et al. (2006). Grouped and hierarchical model selection through composite absolute penalties. *Ann. Stat.*

- **Multiple related tasks**

Obozinski et al. (2006). Multitask feature selection. *Technical Report, UC Berkeley*.



# Minimize SSE + $\lambda$ x regularizer

- **Ridge**
  - gives similar weights to similar variables
  - not very sparse
  - analytical solution
- **Lasso**
  - randomly picks one of several correlated variables
  - sparse
  - LAR algorithm
- **Elastic net**
  - selects variables like the lasso
  - shrinks together the coefficients of correlated variables.
- **Many other regularizers** are possible
  - Lp norms, groups, graphs, trees...