

① Standardised regression  
 $x$  and  $y \sim N(0, 1)$   
 $y$  - centered

ridge reg.  
 $\text{loss}(\hat{\beta}) = (y - \hat{y})^T (y - \hat{y})$

+  $\lambda \hat{\beta}^T \hat{\beta}$   
 ↪ regularization  
 $y = \hat{\beta}_1 x + \hat{\beta}_2 x = x(\hat{\beta}_1 + \hat{\beta}_2)$

(a) optimal  $\hat{\beta}_1, \hat{\beta}_2$

$$\frac{\partial \text{loss}}{\partial \beta_1} \stackrel{\lambda - \text{fixed}}{=} 0 \quad \frac{\partial \text{loss}}{\partial \beta_2} = 0$$

$$\begin{aligned} \text{loss}(\hat{\beta}) &= (y - x(\hat{\beta}_1 + \hat{\beta}_2))^T \\ (y - x(\hat{\beta}_1 + \hat{\beta}_2)) + \lambda \hat{\beta}^T \hat{\beta} &= 0 \quad (2) \\ \hat{\beta} &= \hat{\beta}_1 + \hat{\beta}_2 \end{aligned}$$

$$\widehat{\beta} + \widehat{\beta} = (\widehat{\beta}_1 + \widehat{\beta}_2)^\top (\widehat{\beta}_1 + \widehat{\beta}_2)$$

$$= (\gamma - x(\widehat{\beta}_1 + \widehat{\beta}_2))$$

$$+ x^\top (\widehat{\beta}_1 + \widehat{\beta}_2)^\top (\widehat{\beta}_1 + \widehat{\beta}_2) +$$

$$\begin{aligned}\beta^\top \beta &= ? \\ \beta &= \begin{pmatrix} \beta_1 \\ \beta_2 \\ \vdots \\ \beta_n \end{pmatrix} \quad \beta^\top = \begin{pmatrix} \beta_1 & \beta_2 & \dots & \beta_n \end{pmatrix} \\ &\quad [n \times 1]\end{aligned}$$

$$\begin{aligned}\widehat{\beta}^\top \cdot \widehat{\beta} &= (\widehat{\beta}_1 + \widehat{\beta}_2 + \dots + \widehat{\beta}_n)^\top \\ &= \widehat{\beta}_1^2 + \widehat{\beta}_2^2 + \dots + \widehat{\beta}_n^2\end{aligned}$$

$$S_0, (\hat{\beta}_1 + \hat{\beta}_2)^T (\hat{\beta}_1 + \hat{\beta}_2) = \\ = (\hat{\beta}_{12} + \hat{\beta}_{22})^T (\hat{\beta}_{12} + \hat{\beta}_{22}).$$

$$\begin{aligned} & \textcircled{=} (y - x(\hat{\beta}_1 + \hat{\beta}_2))^T \\ & (y - x(\hat{\beta}_1 + \hat{\beta}_2)) + \\ & + x(\hat{\beta}_{12} + \hat{\beta}_{22}) \\ & (y - x(\hat{\beta}_1 + \hat{\beta}_2))^T \cdot (y - \\ & - x(\hat{\beta}_1 + \hat{\beta}_2)) = \\ & = y^T y - y^T (x(\hat{\beta}_1 + \hat{\beta}_2)) \\ & - (x(\hat{\beta}_1 + \hat{\beta}_2))^T y + \\ & + (x(\hat{\beta}_1 + \hat{\beta}_2))^T x(\hat{\beta}_1 + \hat{\beta}_2) \\ & = y^T y - y^T (\hat{\beta}_1 + \hat{\beta}_2)^T x^T - \end{aligned}$$

$$\begin{aligned}
& - (\hat{\beta}_1 + \hat{\beta}_2) x^+ y^+ \\
& + (\hat{\beta}_1 + \hat{\beta}_2)^T x^T x (\hat{\beta}_1 + \hat{\beta}_2) \\
& = y^T y - y^T x (\hat{\beta}_1 + \hat{\beta}_2) - \\
& - x^T y (\hat{\beta}_1 + \hat{\beta}_2) + \\
& + (\hat{\beta}_1^2 + \hat{\beta}_2^2) x^T x
\end{aligned}$$

$$y^T x (\hat{\beta}_1 + \hat{\beta}_2) =$$

$$= x^T y (\hat{\beta}_1 + \hat{\beta}_2)$$

$$\begin{aligned}
y^T x &= (y_1, y_2, \dots, y_n) \begin{pmatrix} x_1 \\ x_2 \\ \vdots \\ x_n \end{pmatrix} \\
x^T y &= (x_1, x_2, \dots, x_n) \begin{pmatrix} y_1 \\ y_2 \\ \vdots \\ y_n \end{pmatrix}
\end{aligned}$$

$$= + y^T y - 2(\hat{\beta}_1 + \hat{\beta}_2)x^T y +$$

$$+ (\hat{\beta}_1 + \hat{\beta}_2)^2 x^T x$$

$$\text{loss}(\hat{\beta}_1, \hat{\beta}_2) = y^T y -$$

$$- 2(\hat{\beta}_1 + \hat{\beta}_2)x^T y +$$

$$+ (\hat{\beta}_1 + \hat{\beta}_2)^2 x^T x +$$

$$+ (\hat{\beta}_1 + \hat{\beta}_2)$$

$$\frac{\partial \text{loss}(\hat{\beta}_1, \hat{\beta}_2)}{\partial \hat{\beta}_1} =$$

$$\partial \hat{\beta}_1$$

$$= - 2x^T y +$$

$$+ 2(\hat{\beta}_1 + \hat{\beta}_2)x^T x +$$

$$+ 2x^T \hat{\beta}_1 = 0$$

$$\begin{aligned}
 & -2X^T y + 2\hat{\beta}_1 X^T X + \\
 & + 2\hat{\beta}_2 X^T X + 2X\hat{\beta}_2 \\
 \frac{\partial \text{loss}(\beta_1, \beta_2)}{\partial \beta_2} = 
 \end{aligned}$$

$$\begin{aligned}
 & = -2X^T y + 2(\hat{\beta}_1 + \hat{\beta}_2)X^T \\
 & \cdot X + 2\hat{\beta}_2 = \\
 & = -2X^T y + 2\hat{\beta}_1 X^T X + \\
 & + 2\hat{\beta}_2 X^T X + 2X\hat{\beta}_2
 \end{aligned}$$

$$\left\{
 \begin{array}{l}
 \frac{\partial \text{loss}(\beta_1, \beta_2)}{\partial \beta_1} = 0 \quad (1) \\
 \frac{\partial \text{loss}(\beta_1, \beta_2)}{\partial \beta_2} = 0 \quad (2)
 \end{array}
 \right.$$

$$(2) - (2) \rightarrow 2\hat{\beta}_1 - 2\hat{\beta}_2 = 0$$

$$\hat{\beta}_1 = \hat{\beta}_2$$

$$-2x^T y + 2\hat{\beta}_1 x^T x +$$

$$+ 2\hat{\beta}_1 x^T x + x^T \hat{\beta}_1 =$$

$$-2x^T y + 2\hat{\beta}_1 x^T x + x^T \hat{\beta}_1 =$$

$$-x^T y + 2\hat{\beta}_1 x^T x + x^T \hat{\beta}_1 =$$

$$\hat{\beta}_1 (2x^T x + 1) = x^T y$$

$$\hat{\beta}_1 = \frac{x^T y}{2x^T x + 1}$$

$$\boxed{\hat{\beta}_1 = \hat{\beta}_2 = \frac{x^T y}{2x^T x + 1}}$$

$$(b) \lambda \rightarrow \infty$$

$$\lim_{\lambda \rightarrow \infty} \hat{\beta}_1 = \lim_{\lambda \rightarrow \infty} \hat{\beta}_2 =$$

$$= \lim_{\lambda \rightarrow \infty} \frac{x^T y}{2x^T x + \lambda} =$$

$$= 0$$

$$(c) \lambda \rightarrow 0$$
$$\hat{\beta}_1, \hat{\beta}_2 \rightarrow$$

$$\hat{\beta}_1 = \hat{\beta}_2$$

$$x^T \hat{\beta}_1 + \hat{\beta}_2 = \frac{x^T y}{2x^T x + \lambda} +$$

$$\frac{2x^T y}{2x^T x + \lambda}$$

$$\begin{aligned}
 & \lim_{\gamma \rightarrow 0} \hat{\beta}_1 + \hat{\beta}_2 = \\
 &= \lim_{\gamma \rightarrow 0} \frac{2x^T y}{2x^T x + \gamma} \\
 &= \frac{2x^T y}{2x^T x} = \frac{x^T y}{x^T x}
 \end{aligned}$$

$$② y = X \beta + u$$

$\beta$  - non-random

$$E(u|X) = 0$$

$$X \quad \text{rank } X = n \\ [n \times k]$$

$$\text{Var}(u|X) = \sigma^2 W$$

w.t.t)

$\hat{\beta}$  standard OLS  
estimator

$$(a) E(\hat{\beta}|X) = (X^T X)^{-1} X^T y$$

$$\hat{\beta} = (X^T X)^{-1} X^T y$$

$$E((X^T X)^{-1} X^T y | X) =$$

$$\begin{aligned}
 &= (X^T X)^{-1} X^T E(y|X) = \\
 &= (X^T X)^{-1} X^T E(X\beta + \epsilon|X) = \\
 &= (X^T X)^{-1} X^T (X\beta + \\
 &\quad + \epsilon E(\overset{\text{blue}}{y}|X)) = 
 \end{aligned}$$

$$\begin{aligned}
 &= (X^T X)^{-1} X^T X\beta = 
 \end{aligned}$$

$$\begin{aligned}
 &= \underset{\text{brace}}{\underbrace{\beta}}
 \end{aligned}$$

$$\begin{aligned}
 E(\hat{\beta}) &= E(E(\hat{\beta}|X)) = \\
 &= E(\beta) = \underset{\text{brace}}{\underbrace{\beta}}
 \end{aligned}$$

$$\begin{aligned}
 (b) \text{Var}(\hat{\beta}|X) &= X^T y | X \\
 &= \text{Var}((X^T X)^{-1} X^T y | X)
 \end{aligned}$$

$$\begin{aligned}
 &= (X^T X)^{-1} X^T \text{Var}(y|X) \\
 &\quad \cdot ((X^T X)^{-1} X^T)^T = \\
 &= (X^T X)^{-1} X^T \text{Var}(X\beta + \epsilon|X) \\
 &\quad \cdot X (X^T X)^{-1} = \\
 &= (X^T X)^{-1} X^T \text{Var}(y|X) \\
 &\quad \cdot X (X^T X)^{-1} = \\
 &= (X^T X)^{-1} X^T \sigma^2 w \\
 &\quad \cdot X (X^T X)^{-1} = \\
 &= \sigma^2 (X^T X)^{-1} X^T w \\
 &\quad \cdot X (X^T X)^{-1}
 \end{aligned}$$

(c) CI for  $\beta$  will  
 not be valid  
 sc not correct (because new)

because  $\text{Var}(\hat{\beta}|X) = \sigma^2 I + \sigma^2 X(X^T X)^{-1} X^T$   
 not with  $\text{Var}$  of unbiased est

$\hat{\beta}$  is unbiased,  
 not efficient

$$(d) \text{Cov}(y, \hat{\beta}|X) = \\ = \text{Cov}(X\beta + u, (X^T X)^{-1} X^T y|X)$$

$$y = X\beta + u$$

$$(e) \text{Cov}(u, (X^T X)^{-1} X^T (X\beta + u)|X) = \\ = \text{Cov}(u, (X^T X)^{-1} X^T X\beta + \\ + (X^T X)^{-1} X^T u|X) = \\ = \text{Cov}(u, (X^T X)^{-1} X^T u|X) \\ = X^T (X^T X)^{-1} X^T \text{Cov}(u, u|X) = \text{Var}(u|X) =$$

$$= (X^T X)^{-1} X^T \sigma^2 w =$$

$$= \sigma^2 w \times (X^T X)^{-1}$$

③

$$X = \begin{pmatrix} 2 & 1 \\ -1 & 2 \\ 1 & 1 \end{pmatrix}$$

(a)  $X^T X$  and  $\omega$

diagonalize it

$$X^T = \begin{pmatrix} 2 & -1 & 1 \\ 1 & 2 & 1 \end{pmatrix}$$

$$X^T X = \begin{pmatrix} 2 & -1 & 1 \\ 1 & 2 & 1 \end{pmatrix} \begin{pmatrix} 2 & 1 \\ -1 & 2 \\ 1 & 1 \end{pmatrix}$$

$$= \begin{pmatrix} 2 \cdot 2 + 1 + 1 & 2 \cdot 1 - 2 + 1 \cdot 1 \\ 1 \cdot 2 + 2(-1) + 1 \cdot 1 & 1 \cdot 1 + 2 \cdot 2 + 1 \cdot 1 \end{pmatrix} =$$

$$= \begin{pmatrix} 6 & 1 \\ 1 & 5 \end{pmatrix}$$

Diagonalize  $X^T X$

$$\det(X^T X - \lambda I) = 0$$

$$\det \begin{pmatrix} 6-\lambda & 1 \\ 1 & 5-\lambda \end{pmatrix} =$$

$$= (6-\lambda)^2 - 1 = 0$$

$$36 - 12\lambda + \lambda^2 - 1 = 0$$

$$\lambda^2 - 12\lambda + 35 = 0$$

$$D = 36 - 4 \cdot 35 \cdot 1 =$$

$$= 7 \quad \lambda_{1,2} = \frac{12 \pm \sqrt{2}}$$

$$\lambda_1 = 5 \quad \lambda_2 = 7.$$

$$U_1 = \begin{pmatrix} 6 & 5 \\ 2 & 6 & 5 \end{pmatrix} U_1 =$$

$$= \begin{pmatrix} 2 & 2 \\ 2 & 2 & 2 \end{pmatrix} U_1 = 0$$

$$U_1 = \begin{pmatrix} 2 \\ -1 \end{pmatrix}$$

$$\lambda_1 = 7 \quad D = \begin{pmatrix} 7 & 0 \\ 0 & 5 \end{pmatrix}$$

$$U_2 \begin{pmatrix} 6 & 1 \\ 2 & 6 & 1 \end{pmatrix} = U_2 \begin{pmatrix} -1 & 1 \\ 2 & -1 \end{pmatrix}$$

$$U_2 = \begin{pmatrix} 1 \\ 1 \end{pmatrix}$$

(5) SVD of X

$$SVD = U \cdot D \cdot V^T$$

$[n \times n] [n \times n] [n \times n]$

$$X \cdot X^T = \begin{pmatrix} 6 & 1 \\ -1 & 6 \end{pmatrix}.$$

X is Standardized  
for this let us use  
Python code  
code is provided here  
 $X_{\text{standardized}} =$

D is diagonal

$$D = \begin{pmatrix} \sqrt{x_2} & 0 \\ 0 & \sqrt{\frac{1}{x_1}} \end{pmatrix} =$$

$$= \begin{pmatrix} \sqrt{7} & 0 \\ 0 & \sqrt{5} \end{pmatrix}$$

$$V^T = \begin{pmatrix} \overrightarrow{v_{11}} \\ \overrightarrow{v_{21}} \\ \overrightarrow{v_{12}} \\ \overrightarrow{v_{22}} \end{pmatrix} =$$

$$\|v_1\| = \sqrt{1^2 + 1^2} = \sqrt{2}$$

$$\|v_2\| = \sqrt{1^2 + (-1)^2} = \sqrt{2}$$

$$= \begin{pmatrix} \frac{1}{\sqrt{2}} & \frac{1}{\sqrt{2}} \\ -\frac{1}{\sqrt{2}} & -\frac{1}{\sqrt{2}} \end{pmatrix}$$

$$U = \begin{pmatrix} \frac{1}{\sqrt{x_2}} & 0 \\ 0 & \frac{1}{\sqrt{x_1}} \end{pmatrix} = \begin{pmatrix} \frac{1}{\sqrt{7}} & 0 \\ 0 & \frac{1}{\sqrt{5}} \end{pmatrix}$$

$$SVD = UDV^T$$

$$X^T X = \begin{pmatrix} 6 & 1 \\ -1 & 6 \end{pmatrix} =$$

$$= \begin{pmatrix} -\frac{1}{\sqrt{7}} & 0 \\ 0 & \frac{2}{\sqrt{5}} \end{pmatrix} \cdot \begin{pmatrix} \sqrt{7} & 0 \\ 0 & \sqrt{5} \end{pmatrix}$$

$$\cdot \begin{pmatrix} -\frac{1}{\sqrt{2}} \\ \frac{1}{\sqrt{2}} \end{pmatrix} \begin{pmatrix} \frac{1}{\sqrt{2}} \\ -\frac{1}{\sqrt{2}} \end{pmatrix} V^T$$

D

(c) best approxim  
to  $X$  with rank  $= i$

keep largest  $\rightarrow$  other

$$D = \begin{pmatrix} \sqrt{\lambda} & 0 \\ 0 & 0 \end{pmatrix}$$

$$X_{\text{standard}} = \frac{x - \bar{x}}{x_{\text{std}}}$$

$$x_{\text{approx}} \sim x + x_{\text{std}} + \bar{x}$$

$$SVD = UDV^T$$



$$SVD \text{ of } X = UDV^T$$

D positive from  
highest to lowest  
 $d_{11} > d_{22} > \dots > 0$

$$\|x_w\|^2 \rightarrow \max_{\|w\|=1} s.t. \|w_i\|^2$$

(d) Lagrangian

$$\left\{ \begin{array}{l} \|Xw\|^2 \rightarrow \max \\ \text{s.t.} \\ \|w\| = 1 \end{array} \right.$$

$$\begin{aligned} L &= \|Xw\|^2 - \lambda (\|w\|^2 - 1) \\ &= w^T X^T \cdot X \cdot w - \\ &\quad - \lambda (w^T \cdot w - 1) \end{aligned}$$

(e) F.O.C

$$\begin{aligned} \frac{\partial L}{\partial w} &= (w^T X^T \cdot X w)'_w - \\ &\quad - \lambda (w^T w - 1)'_w = \\ &= 2 X^T X w - 2 \lambda w = 0 \\ &2 X^T X w = 2 \lambda w \end{aligned}$$

$$Y = X^T X$$

(F) optimal  $\sim$  of  
columns of  $V$

$$V - ?$$

SVD

$$X^T \cdot X = (U D V^T)^T$$

$$\begin{aligned} U^T \cdot U \cdot D V^T &= V^T \cdot D^T \\ U^T \cdot U \cdot D V^T &= \\ &= V D^T D V^T \end{aligned}$$

$$2 X^T \cdot w = \lambda w$$

$$\begin{aligned} U D^T D V^T w &= \lambda w \\ (U D^T D \cdot V^T) w V^T &= X V^T w \end{aligned}$$

$$D^T D U^T w = \lambda U^T w$$

$$\begin{pmatrix} \delta_1 & & & \\ & \delta_2 & & \\ & & \ddots & \\ & & & \delta_r \end{pmatrix} U^T w = \lambda U^T w$$

$$\delta_i \cdot U^T w = \lambda U^T w$$

Optimal  $w$  is

the first col  
of  $U$