

Lecture 7: Generalised Method of Moments

Econometric Methods – Warsaw School of Economics

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Outline

- 1 Introduction
- 2 GMM estimator
- 3 Application: Gali&Gertler's hybrid Phillips curve (1999)

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Instrumental variables: general idea

- OLS estimation based on the general underlying assumption that $E(\mathbf{X}_{T \times k}^T \boldsymbol{\varepsilon}_{T \times 1}) = \mathbf{0}_{k \times 1}$ (by Gauss-Markov).
- It may be broken i.a. for the following reasons:
 - just by construction of the economic model;
 - two-way causality between \mathbf{y}_t and a subset of \mathbf{x}_t ;
 - non-random sample selection.
- Solution: find variables that are truly orthogonal to $\boldsymbol{\varepsilon}_{T \times 1}$ (“instrumental variables”).

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OLS and IV: alternative approach

- OLS: residuals uncorrelated with regressors, $E(\mathbf{X}_{T \times k}^T \boldsymbol{\varepsilon}_{T \times 1}) = \mathbf{0}_{k \times 1}$
 - these are k “moment conditions” from which we infer the estimates
 - $\mathbf{X}^T \mathbf{y} = \mathbf{X}^T \mathbf{X} \boldsymbol{\beta} + \underbrace{\mathbf{X}^T \boldsymbol{\varepsilon}}_0 \Rightarrow \hat{\boldsymbol{\beta}}^{OLS} = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}$
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- what if we get more than $I(> k)$ “moment conditions”, i.e. more than we actually need? $\mathbf{Z}_{T \times I}^T \boldsymbol{\varepsilon}_{T \times 1} = \mathbf{0}_{I \times 1}$
 - there are $I(> k)$ “moment conditions”, so not all of them can be fulfilled by modifying $\boldsymbol{\beta}$ (only k parameters)
 - $\mathbf{Z}^T \mathbf{y} = \mathbf{Z}^T \mathbf{X} \boldsymbol{\beta} + \underbrace{\mathbf{Z}^T \boldsymbol{\varepsilon}}_0 \Rightarrow$ but we won't invert $\mathbf{Z}^T \mathbf{X}$ (not a square matrix this time!)

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Limitations of IV estimation

- $\beta^{IV} = (\mathbf{Z}^T \mathbf{X})^{-1} \mathbf{Z}^T \mathbf{y}$
 - \mathbf{Z} must have the same number of columns as \mathbf{X} for this operation to be feasible
- in IV, there must be as many instrumental variables as regressors (some of which can instrumentalise themselves)
- what are \mathbf{Z} ?
 - uncorrelated with ε , correlated with \mathbf{X}
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So let's take yet another perspective...

- OLS minimises the quadratic form $[\mathbf{X}^T \boldsymbol{\varepsilon}(\boldsymbol{\beta})]^T [\mathbf{X}^T \boldsymbol{\varepsilon}(\boldsymbol{\beta})]$
 - wrt. $\boldsymbol{\beta}_{k \times 1}$ (down to zero!)
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 - wrt. $\boldsymbol{\beta}_{k \times 1}$ with $\mathbf{Z}_{l \times T}$ down to zero, i.e. it is impossible to perfectly fulfil all the moment conditions
 - cannot solve l equations (moment conditions) for $k < l$ unknowns
- ...so take quadratic form $(\mathbf{Z}^T \boldsymbol{\varepsilon})^T \mathbf{W} (\mathbf{Z}^T \boldsymbol{\varepsilon})$ instead!: $\mathbf{W}_{l \times l}$
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How to obtain \mathbf{W} ?

- in general: any symmetric, positive-definite matrix sized $I \times I$
- according to [Hansen \(1982\)](#), you should assign the lowest weights to the moment conditions with lowest precision of estimation, which implies

$$\mathbf{V} = \frac{1}{T} \mathbf{Z}^T \mathbf{\Omega} \mathbf{Z}$$

- with $\mathbf{\Omega}$ – variance-covariance matrix of the estimator
 - under white noise: $\hat{\mathbf{\Omega}} = \hat{\sigma}^2 \mathbf{I}$
 - under non-spherical disturbances – [White](#) or [Newey-West](#) versions (see: serial correlation / heteroskedasticity lectures)

GMM as iterative procedure

- $\hat{\varepsilon}$ unknown in advance!
 - 1 estimate under the assumption of white noise
 - 2 calculate the residuals
 - 3 update \hat{W}
- various implementations of this algorithm in the software (i.a. as a two-step or iterative estimator – cf. GLS)

Derivation of GMM estimator (1)

Let us minimise the quadratic form with respect to β :

$$\begin{aligned}
 \epsilon^T \mathbf{Z} \mathbf{W} \mathbf{Z}^T \epsilon &= (\mathbf{y} - \beta \mathbf{X})^T \mathbf{Z} \mathbf{W} \mathbf{Z}^T (\mathbf{y} - \beta \mathbf{X}) = \\
 &= \mathbf{y}^T \mathbf{Z} \mathbf{W} \mathbf{Z}^T \mathbf{y} - \mathbf{y}^T \mathbf{Z} \mathbf{W} \mathbf{Z}^T \mathbf{X} \beta \\
 &\quad - \beta^T \mathbf{X}^T \mathbf{Z} \mathbf{W} \mathbf{Z}^T \mathbf{y} + \beta^T \mathbf{X}^T \mathbf{Z} \mathbf{W} \mathbf{Z}^T \mathbf{X} \beta \\
 &\rightarrow \min
 \end{aligned}$$

From matrix calculus:

$$\frac{\partial \mathbf{x}^T \mathbf{A} \mathbf{x}}{\partial \mathbf{x}} = \mathbf{x}^T (\mathbf{A}^T + \mathbf{A})$$

$$\frac{\partial \mathbf{x}^T \mathbf{a}}{\partial \mathbf{x}} = \mathbf{a}^T$$

Compute the derivative and set it equal to 0.

Derivation of GMM estimator (2)

- $-y^T ZWZ^T X - y^T ZWZ^T X + \beta^T (X^T ZWZ^T X + X^T ZWZ^T X) = 0$
- $\beta^T (X^T ZWZ^T X) = y^T ZWZ^T X$
- $(X^T ZWZ^T X) \beta = X^T ZWZ^T y$
- $\hat{\beta}^{GMM} = (X^T ZWZ^T X)^{-1} X^T ZWZ^T y$

Nonlinear GMM:

- you can also (numerically) minimise the quadratic form for a nonlinear model;
- it is advisable to write the moment conditions “as close to linear as possible”.

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GMM: special cases

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GMM: special cases

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$$\hat{\beta}^{OLS}$$

Variance-covariance of GMM estimator

- GMM estimator: consistent, asymptotically normally distributed
- asymptotic variance-covariance estimator in a linear model:

$$\text{Var}(\beta^{GMM}) = \frac{1}{T} \left[\frac{1}{T} (\mathbf{X}^T \mathbf{Z}) \mathbf{W}^{-1} \frac{1}{T} (\mathbf{Z}^T \mathbf{X}) \right]^{-1}$$

J statistic

- the non-zero value of the minimised quadratic form is interpretable
 - how far are we from fulfilling the (excessive) orthogonality conditions?
 - the lower the value, the lower the “distance” to perfect fulfilment of excessive conditions
- divided by T , it is χ^2 -distributed with degrees of freedom equal to $l - k$ (instruments in excess of parameters)

J -test of orthogonality

$$J(\hat{\beta}, \hat{\Omega}^{-1}) = \frac{1}{T} \varepsilon(\hat{\beta})^T \mathbf{Z} \hat{\Omega}^{-1} \mathbf{Z}^T \varepsilon(\hat{\beta})$$

H_0 : all orthogonality conditions fulfilled

H_1 : some orthogonality conditions not fulfilled

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Hybrid Phillips curve

- *Gali, Gertler (1999):*

$$\pi_t = \gamma_b \pi_{t-1} + \gamma_f E_t \pi_{t+1} + \lambda mc_t + \varepsilon_t$$

where π_t – inflation rate, mc_t – real marginal cost, ε_t – error term.

- Why GMM?

- there is an unobservable variable on the right-hand side, $E_t \pi_{t+1}$
- we can just **replace** it with observable $\pi_{t+1} = E_t \pi_{t+1} + v_{t+1}$,
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- $$\pi_t = \gamma_b \pi_{t-1} + \gamma_f \pi_{t+1} (v_{t+1}) + \lambda mc_t + \varepsilon_t$$
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Orthogonality conditions

- $E_t [\varepsilon_t \mathbf{z}_t] = E_t [(\pi_t - \gamma_b \pi_{t-1} - \gamma_f E_t \pi_{t+1} - \lambda m c_t) \mathbf{z}_t] = E_t [(\pi_t - \gamma_b \pi_{t-1} - \gamma_f \pi_{t+1} - \lambda m c_t) \mathbf{z}_t] = \mathbf{0}$
 - the expected value on the orthogonality condition allows to drop the expected value on π_{t+1}
- in this context, we interpret the instrument set \mathbf{Z} as **variables that allow to forecast inflation one period ahead without systematic errors**
- there should be more than 3 instruments to use GMM here

Example

(Imperfect) replication of Gali-Gertler results:

- Read the paper by Gali and Gertler.
- Consider the following set of variables as linear one-period-ahead predictors:
 - *inflation (4 lags), log real ULC (1 lag), output gap (1 lag), short- vs long-term interest rate spread (1 lag), log-differences of wage index (1 lag), log-differences of commodity price index (1 lag)*
- Define the instruments and the initial weight matrix \mathbf{W} .
- Estimate the model. Discuss the results. Do they confirm Your intuition? If not, look for the (economic) solution in the article.