

# CS559-B HW1 Solution

Due: Sep. 25th, 2019

## Problem 1 (10pt): Independence and un-correlation

(1) Suppose  $X$  and  $Y$  are two continuous random variables, show that if  $X$  and  $Y$  are independent, then they are uncorrelated.

Solution: suppose  $X$  has density function  $f_X(x)$ ,  $Y$  has density  $f_Y(y)$ . If  $X$  and  $Y$  independent, then  $f(x, y) = f_X(x)f_Y(y)$ . Therefore:

$$\begin{aligned}\text{Cov}(X, Y) &= \mathbb{E}[(X - \mu_X)(Y - \mu_Y)] \\&= \int_x \int_y (x - \mu_X)(y - \mu_Y) f(x, y) dx dy \\&= \int_x \int_y (x - \mu_X)(y - \mu_Y) f_X(x) f_Y(y) dx dy \\&= \int_x (x - \mu_X) f_X(x) dx \int_y (y - \mu_Y) f_Y(y) dy \\&= \left( \int_x x f_X(x) dx - \mu_X \right) \left( \int_y y f_Y(y) dy - \mu_Y \right) = 0.\end{aligned}$$

(2) Suppose  $X$  and  $Y$  are uncorrelated, can we conclude  $X$  and  $Y$  are independent? If so, prove it, otherwise, give one counterexample. (Hint: consider  $X \sim \text{Uniform}[-1, 1]$  and  $Y = X^2$ )

Solution: consider  $X \sim \text{Uniform}[-1, 1]$  and  $Y = X^2$ . Then  $X$  and  $Y$  are not independent. However,  $\mathbb{E}(XY) = \mathbb{E}(X^3) = 0$  and  $\mathbb{E}(X) = 0$ , thus,  $\text{Cov}(X, Y) = \mathbb{E}(XY) - \mathbb{E}(X)\mathbb{E}(Y) = 0$

**Problem 2 (15pt):** [Minimum Error Rate Decision] Let  $\omega_{max}(x)$  be state of nature for which  $P(\omega_{max}|x) \geq P(\omega_i|x)$  for all  $i = 1, \dots, c$ .

(1) Show that  $P(\omega_{max}|x) \geq \frac{1}{c}$

Solution: since  $P(\omega_{max}|x) \geq P(\omega_i|x)$ , then:

$$\sum_{i=1}^c P(\omega_{max}|x) \geq \sum_{i=1}^c P(\omega_i|x) = 1$$

So  $P(\omega_{max}|x) \geq \frac{1}{c}$ .

(2) Show that for minimum-error-rate decision rule, the average probability of error is given by

$$P(\text{error}) = 1 - \int P(\omega_{max}|x)p(x)dx$$

Solution: by definition of averaged probability of error, we have:

$$\begin{aligned}
P(error) &= \int P(error|x)p(x)dx \\
&= \int (1 - P(\omega_{max}|x))p(x)dx \\
&= 1 - \int P(\omega_{max}|x)p(x)dx
\end{aligned}$$

(3) Show that  $P(error) \leq \frac{c-1}{c}$

Solution: from (2) and (1) we have:

$$P(error) = 1 - \int P(\omega_{max}|x)p(x)dx \leq 1 - \int \frac{1}{c}p(x)dx = 1 - \frac{1}{c} = \frac{c-1}{c}$$

**Problem 3 (10pt):** [Likelihood Ratio] Suppose we consider two category classification, the class conditionals are assumed to be Gaussian, i.e.,  $p(x|\omega_1) = N(4, 1)$  and  $p(x|\omega_2) = N(8, 1)$ , based on prior knowledge, we have  $P(\omega_2) = \frac{1}{4}$ . We do not penalize for correct classification, while for misclassification, we put 1 unit penalty for misclassifying  $\omega_1$  to  $\omega_2$  and put 3 unit for misclassifying  $\omega_2$  to  $\omega_1$ . Derive the bayesian decision rule using likelihood ratio.

Solution: we have the likelihood distribution:  $p(x|\omega_1) = \frac{1}{\sqrt{2\pi}}e^{-\frac{(x-4)^2}{2}}$ ,  $p(x|\omega_2) = \frac{1}{\sqrt{2\pi}}e^{-\frac{(x-8)^2}{2}}$ , and we could also easily obtain the prior  $p(\omega_2) = \frac{1}{4}$ ,  $p(\omega_1) = 1 - p(\omega_2) = \frac{3}{4}$ . Based on the problem, we could get the loss matrix as:

$$\lambda = \begin{bmatrix} 0 & 3 \\ 1 & 0 \end{bmatrix}$$

For bayesian decision based on likelihood ratio, we decide  $\omega_1$  if the following:

$$\frac{p(x|\omega_1)}{p(x|\omega_2)} > \frac{\lambda_{12} - \lambda_{22} P(\omega_2)}{\lambda_{21} - \lambda_{11} P(\omega_1)}$$

For RHS, we have  $RHS = \frac{3-0}{1-0} * \frac{1}{3} = 1$ , For LHS, we have  $\frac{p(x|\omega_1)}{p(x|\omega_2)} = e^{-\frac{(x-4)^2}{2} + \frac{(x-8)^2}{2}}$ . Therefore we decide  $\omega_1$  if:

$$e^{-\frac{(x-4)^2}{2} + \frac{(x-8)^2}{2}} > 1$$

Take ln and solve the equation, we get bayesian decision rule which is decide  $\omega_1$  if  $x < 6$ , decide  $\omega_2$  otherwise.

**Problem 4 (15pt):** [Minimum Risk, Reject Option] In many machine learning applications, one has the option either to assign the pattern to one of  $c$  classes, or to reject it as being unrecognizable.

If the cost for reject is not too high, rejection may be a desirable action. Let

$$\lambda(\alpha_i|\omega_j) = \begin{cases} 0, & i = j \text{ and } i, j = 1, \dots, c \\ \lambda_r, & i = c + 1 \\ \lambda_s, & \text{otherwise} \end{cases}$$

where  $\lambda_r$  is the loss incurred for choosing the  $(c+1)$ -th action, rejection, and  $\lambda_s$  is the loss incurred for making any substitution error.

(1) Derive the decision rule with minimum risk.

Solution: for  $i = 1, \dots, c$ :

$$\begin{aligned} R(\alpha_i|x) &= \sum_{j=1}^c \lambda(\alpha_i|\omega_j) P(\omega_j|x) \\ &= \lambda_s \sum_{j=1, j \neq i} P(\omega_j|x) \\ &= \lambda_s [1 - P(\omega_i|x)] \end{aligned}$$

For  $i = c + 1$ :

$$R(\alpha_{c+1}|x) = \lambda_r$$

Thus, decision rule based on minimum risk would be: decide  $\omega_i$  if  $R(\alpha_i|x) \leq R(\alpha_{c+1}|x)$ , i.e.,  $P(\omega_i|x) \geq 1 - \frac{\lambda_r}{\lambda_s}$ , and reject otherwise.

(2) What happens if  $\lambda_r = 0$ ?

Solution: if  $\lambda_r = 0$ , we always reject.

(3) What happens if  $\lambda_r > \lambda_s$ ?

Solution: if  $\lambda_r > \lambda_s$ , we never reject.

**Problem 5 (10pt):** [Maximum Likelihood Estimation (MLE)] Suppose we have training samples  $\{x_1, x_2, \dots, x_n\}$ . Consider the following distributions:

(1) Exponential density:  $f(x; \theta) = \theta e^{-\theta x}$ ,  $x \geq 0$ ,  $\theta > 0$ , find MLE for  $\theta$

Solution: the likelihood function is:

$$L(\theta; x_1, \dots, x_n) = \theta^n \exp(-\theta \sum_{i=1}^n x_i)$$

The log-likelihood function would be:

$$l(\theta; x_1, \dots, x_n) = n \ln(\theta) - \theta \sum_{i=1}^n x_i$$

Let  $\frac{dl}{d\theta} = 0$ , we have the MLE for  $\theta$ :

$$\hat{\theta} = \frac{n}{\sum_{i=1}^n x_i}$$

(2) Uniform distribution  $Unif[\theta_1, \theta_2]$ , find the MLE for  $\theta_1$  and  $\theta_2$ .

Solution: we do not need to use derivative calculus to find the MLE in this case. The density for  $Unif[\theta_1, \theta_2]$  is  $\frac{1}{\theta_2 - \theta_1}$  on  $[\theta_1, \theta_2]$ . Therefore, the likelihood function is:

$$L(\theta_1, \theta_2 | x_1, \dots, x_n) = \begin{cases} (\frac{1}{\theta_2 - \theta_1})^n, & \text{if all } x_i \text{ in the interval } [\theta_1, \theta_2] \\ 0, & \text{otherwise} \end{cases}$$

We maximize the likelihood function by making  $\theta_2 - \theta_1$  as small as possible. The only constraint is that the interval  $[\theta_1, \theta_2]$  should include all data. Therefore, the MLE for  $\theta_1$  and  $\theta_2$  is:

$$\begin{aligned} \hat{\theta}_1 &= \min(x_1, \dots, x_n) \\ \hat{\theta}_2 &= \max(x_1, \dots, x_n) \end{aligned}$$