

MLE ~~Maximum Likelihood Estimation~~

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Maximum Likelihood Estimation (MLE)

Let x_1, x_2, \dots, x_n be a random sample (iid x_i) with common pdf $f_x(x, \theta)$, $\theta \in \Theta$. Also assume that (x_1, x_2, \dots, x_n) be a observed value of (x_1, x_2, \dots, x_n) . Then likelihood function of θ given the data $x = (x_1, x_2, \dots, x_n)$ is defined as

$$L(\theta | x) = \prod_{i=1}^n f_x(x_i, \theta). \quad \text{--- ①}$$

A value of θ that maximizes $L(\theta | x)$ is called the Maximum Likelihood Estimator (MLE) of θ . Normally MLE of θ is denoted as $\hat{\theta} = \hat{\theta}(x)$. Thus

$$L(\hat{\theta} | x) = \max_{\theta \in \Theta} \{L(\theta | x)\}$$

* Maximum likelihood estimation method is most widely used estimation method in statistical literature.

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Before we discuss specific example, we must stress that when ever a maximum is sought by differentiation, the 2nd order derivatives must also be examined in search of a maximum.

It should also be mentioned that maximization of the likelihood function ① is equivalent to maximization of its logarithm function.

Note: let us recall that a function $y = f(x)$ attains a maximum at a point $x = x_0$ if $\frac{d}{dx} f(x) \Big|_{x=x_0} = 0$ and $\frac{d^2}{dx^2} f(x) \Big|_{x=x_0} < 0$.

Ex: let x_1, x_2, \dots, x_n be iid rvs from Bernoulli $\text{Ber}(\theta)$ distⁿ. Find MLE of θ .

Solⁿ: The likelihood funcⁿ of ~~given~~ θ given the observed data $x = (x_1, x_2, \dots, x_n)$ is

$$L(\theta|x) = \prod_{i=1}^n f_X(x_i, \theta) = \prod_{i=1}^n \theta^{x_i} (1-\theta)^{1-x_i} = \theta^{\sum_{i=1}^n x_i} (1-\theta)^{n - \sum_{i=1}^n x_i}$$

The log-likelihood function of θ is given by

$$\ln(\theta) = \log(L(\theta|x)) = \log\left(\theta^{\sum_{i=1}^n x_i} (1-\theta)^{n - \sum_{i=1}^n x_i}\right)$$

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$$l(\theta) = \left(\sum_{i=1}^n x_i \right) \log \theta + \left(n - \sum_{i=1}^n x_i \right) \log (1-\theta)$$

$$\frac{dl(\theta)}{d\theta} = \frac{\sum_{i=1}^n x_i}{\theta} - \frac{\left(n - \sum_{i=1}^n x_i \right)}{1-\theta}$$

$$\therefore \frac{dl(\theta)}{d\theta} = 0 \Rightarrow \theta = \frac{1}{n} \sum_{i=1}^n x_i \quad \text{--- (*)}$$

\therefore Now let us ~~check~~ check the 2nd derivate

$$\frac{d^2 l(\theta)}{d\theta^2} = - \frac{\sum_{i=1}^n x_i}{\theta^2} - \frac{\left(n - \sum_{i=1}^n x_i \right)}{(1-\theta)^2}$$

$< 0 \quad \forall \quad \theta$ such that $0 \leq \theta \leq 1$.

In particular if $\theta = \frac{1}{n} \sum_{i=1}^n x_i$ then also

$$\frac{d^2 l(\theta)}{d\theta^2} < 0.$$

Thus MLE of θ is

$$\hat{\theta} = \frac{1}{n} \sum_{i=1}^n x_i = \bar{X}$$

Ex: Determine the MLE of λ in Poisson $P(\lambda)$ distribution based on a random sample X_1, X_2, \dots, X_n taken from $P(\lambda)$ distⁿ.

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The likelihood function of λ is given by

$$L(\lambda|x) = \prod_{i=1}^n f(x_i, \lambda) = \prod_{i=1}^n \frac{e^{-\lambda} \lambda^{x_i}}{x_i!} = \frac{e^{-n\lambda} \lambda^{\sum_{i=1}^n x_i}}{\prod_{i=1}^n x_i!}$$

the log-likelihood function of λ is given by

$$l(\lambda) = \log(L(\lambda|x)) = \log \left[\frac{e^{-n\lambda} \lambda^{\sum_{i=1}^n x_i}}{\prod_{i=1}^n x_i!} \right]$$

$$= -n\lambda + \left(\sum_{i=1}^n x_i \right) \log \lambda - \log \left(\prod_{i=1}^n x_i! \right)$$

now

$$\frac{dl(\lambda)}{d\lambda} = -n + \frac{\sum_{i=1}^n x_i}{\lambda}$$

$$\text{Thus } \frac{dl(\lambda)}{d\lambda} = 0 \Rightarrow \lambda = \frac{1}{n} \sum_{i=1}^n x_i = \bar{x}.$$

$$\text{Also } \frac{d^2 l(\lambda)}{d\lambda^2} = -\frac{\left(\sum_{i=1}^n x_i \right)}{\lambda^2} < 0 \quad \forall \lambda > 0.$$

Therefore MLE of λ is

$$\boxed{\hat{\lambda} = \bar{x}}$$

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Ex: Let x_1, x_2, \dots, x_n be iid rvs from a $N(\mu, \sigma^2)$ distribution. find MLE when following conditions hold

- (i) μ is unknown, σ^2 known
- (ii) μ is known, σ^2 unknown
- (iii) μ unknown, σ^2 unknown.

Ans: (i) $\hat{\mu} = \bar{x}$

(ii) $\hat{\sigma}^2 = \frac{1}{n} \sum_{i=1}^n (x_i - \mu)^2$

(iii) $\hat{\mu} = \bar{x}, \hat{\sigma}^2 = \frac{1}{n} \sum_{i=1}^n (x_i - \bar{x})^2$

Result: Invariance Property of MLE

Maximum likelihood Estimators satisfy a very useful property known as invariance property.

Suppose MLE of a parameter θ is $\hat{\theta}$.

You are interested to know what is the MLE of a function of θ , say $g(\theta)$. Then MLE of $g(\theta)$ is $g(\hat{\theta})$. This is known as invariance property of MLE.

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In the previous example suppose we ask:
~~(i)~~ in part:

(ii) find MLE of σ then your answer
 will be $\hat{\sigma} = \sqrt{\frac{1}{n} \sum_{i=1}^n (x_i - \mu)^2}$

Similarly in part

(iii) $\hat{\sigma} = \sqrt{\frac{1}{n} \sum_{i=1}^n (x_i - \bar{x})^2}$.

In part (ii) suppose we ask what is the
 MLE of e^{μ} . Your answer should be
 it is $e^{\bar{x}}$.

[In part (iii) what is MLE of $\mu + k\sigma$
 where k is given constant.]

Ex: Let $x_1, x_2, \dots, x_n \stackrel{\text{iid}}{\sim} \text{Ber}(\theta)$. Find MLE of $\sqrt{\theta(1-\theta)}$.

Ex: Let $x_1, x_2, \dots, x_n \stackrel{\text{iid}}{\sim} \text{Poisson } P(\theta)$. Find MLE
 of $e^{-2\theta}$.

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Ex: Let x_1, x_2, \dots, x_n be a random sample from a exponential $\exp(\lambda)$ distⁿ. Find MLE of λ .

Ex: Let x_1, x_2, \dots, x_n be a random sample from Gamma $G(\alpha, \beta)$ distⁿ. Find MLE

(i) of α when β known (ii) of β when α known
(iii) of α and β when both unknown.

Many more examples you can try yourself.