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For replication, go to: <https://github.com/DamonCharlesRoberts/white-identity-sources>.

Economic concerns appear to be weak predictors of white political identity*

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ABSTRACT Do economic or political threats explain reported white identity? Overall, the social identity literature would suggest that white identity would increase in response to economic threats. However, a number of those that study white identity, specifically, argue that it results from concern about political influence. Considering what whiteness means historically and contemporaneously, I argue that we should expect that political threats reflect stronger associations with white identity. Using data from the 2012, 2016, and 2020 American National Election Study, I consider a single model using penalized regression containing proxies of economic and political threats. I find evidence suggesting that in the post-Trump era, white political identity is strongly associated with reported feelings of Whites' loss of political influence as opposed to economic threats, as some suggest and may expect.

KEYWORDS white identity; economic threat; cultural threat

Introduction

Those conducting post-mortem analyses of the 2016 Presidential election of Donald J. Trump, a presidential candidate that personifies white grievance politics, contend that the substantial presence of non-college-educated white supporters

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explains his electoral success (Griffin, Teixeira, and Halpin 2017). Analyses by political scientists largely attribute Trump's success in attracting voters to those who see their whiteness as a politically relevant identity where they are distinct from Americans of other racial groups (Jardina, Kalmoe, and Gross 2020). For a long time, political scientists saw white political behavior and attitudes through the lens of prejudice. These analyses exclude another dimension of whites' racial attitudes: their affinity for, feelings of linked fate with other whites, and an expression that their race is a politically-relevant identity. More recently, scholars using the Tajfel and Turner (1986) social psychological framework of identity challenge this tendency to focus on whites' prejudice by arguing that some whites may see their race as an identity essential and relevant to politics. As a result, while some may express out-group attitudes, they also express an in-group preference for other whites (Jardina 2019). Despite the importance of these attitudes, the source of these in-group preferences remains somewhat unclear.

There appears to be a mismatch between those who study sources of a variety of social identities and those that study white political identity. Scholars studying identity, more broadly, often focus on economic threats as a critical motivator for an expressed connection to a social in-group (see Huddy 2001). However, some scholars studying white political attitudes and racial identity tend to focus on the role of political threats (see Bartels 2020). This mismatch leads to some conceptualizing white identity as originating from economic views in the tradition of the identity literature. In contrast, others conceptualize white identity as coming from political and cultural threats.

I agree with those who focus on the connection between white political identity and political threats. Considering identity's historical and contextual roots (Trawalter, Higginbotham, and Henderson 2022), I argue that white identifiers are concerned about their status in politics and society. Though I acknowledge that economic threat is a potent activation of identity and encourages engagement in group-based behavior, white identity holds decisive cultural and institutional significance in the U.S. (Rile Hayward 2013). One's whiteness is not a descriptive characteristic, but it reflects a history of allowances by institutions and advantages in one's ability to influence politics. Though race and economics are undoubtedly correlated, we must be careful in specifying the data-generating process.

Along with providing a narrative that scholars might use to identify potential mechanisms connecting political threats to white identity, I examine the correlates of white identity. Using data from the 2012, 2016, and 2020 American National Election Study, I utilize penalized regression to test the claim that a handful of predictors that reflect attitudes about political power are more influential than general (sociotropic) and pocketbook (egotropic) attitudes about the economy. I find that proxies of political threat appear stronger than economic attitudes in the post-Trump era. As this particular analysis is not a causal test of the mechanism connecting threat to white identity, I close by arguing that scholars should continue to examine the mechanisms that explain this connection.

Identity and white political identity

Social identity theory suggests that individuals form attitudes based on groups they associate with, even if those groups are subjectively defined (see Tajfel and Turner 1986). In the famous Robber's Cave experiment, Sherif (1988) divided a group of adolescent boys into two groups arbitrarily. After a week of spending time with others in the group, the boys came up with group names that defined the two groups and distinguished themselves; one group called themselves the Eagles, and the other group referred to themselves as the Rattlers. With the introduction of competition, these boys reported fierce hostility toward the members of the other group and expressed positive attitudes toward those within the group.

This example illustrates several features of social identities. The first feature of social identity theory is that these groups we may associate ourselves with are often subjectively defined (see Huddy 2001). The second feature is that these subjective definitions of the group make distinctions between the two groups that come to matter to members of each group - some are part of our group, and some are part of that other group. Many cognitive theories suggest that this occurs as a result of the tendency for humans to reduce cognitive load by grouping things together (Huddy 2001). The third feature is that introducing competition between these groups shapes attitudes toward those within and without your group. The motivator is that groups with higher status often want to keep that status, and individuals

in the high-status group are much more likely to support inequality to satisfy that motivation (Burns and Kinder 2012; Morrison, Fast, and Ybarra 2009). Fourth, individuals firmly hold these attitudes, which tend to be long-lasting. An additional feature of theories of social identity is that the degree to which one might identify with one group or another is fluid and that whether one identity that one might hold is active or if it is another often depends on the context and the identity's relevance to that context (Huddy 2012).

Racial groups in the U.S. contain a number of these features. As a result, a burgeoning literature in political science that has made this connection takes the racial identity of whites more seriously as one that informs their political attitudes (see Jardina 2019). Race in America is subjectively defined (see Rile Hayward 2013), and these definitions are deeply embedded and reinforced by our institutions (López 2006). Whiteness and one's willingness to identify with that group is important (Wong and Cho 2005) as it reflects a desire for an association with the history and affordances of being white (McDermott and Samson 2005).

As an identity, much of the literature view these attitudes as coming from the presence or focus on economic threats. Some studying white political identity find economic predictors as substantively and statistically significant (see Pérez et al. 2022). From non-experimental evidence, others make similar arguments about what predicts white political identity. Bunyasi (2019) provides evidence suggesting that Trump's messages highlighting increased job market competition harming whites because of affirmative action was a successful strategy and helped build his base of support. After all, the identity literature, more broadly, often focuses on economic threats. Evolutionary psychologists often take the position that identity and the resulting attitudes and behaviors directed toward out-groups are the result of competition over material resources (see Huddy 2001). Many political scientists use this as an explanation for negative attitudes directed toward immigrants among those who feel insecure about their job prospects (Pardos-Prado and Xena 2019). Others argue that deindustrialization and job loss activate whites' concerns about the racial hierarchy and predict their support for Republican challengers that express support for keeping the racial status quo (Baccini and Weymouth 2021).

As much of the literature on whites' racial attitudes rely heavily on social identity theory as a paradigm, a significant

portion of it presents the source of these group-based views as responses to economic threat. However, some disagree with this position (see Jardina 2019) and view these attitudes as the result of changes to the racial hierarchy. These other perspectives provide explanations for whites' attitudes and conceptualize the definition of whiteness and white identity as a desire to retain political power.

Historical accounts that contextualize attitudes stemming from white racial identity seem to suggest that, though economics correlates with racial attitudes, what it meant to be white was no longer about class. However, it became about what it meant for access and privileges afforded by institutions (Rile Hayward 2013). In the 19th century, privilege came from one's freedom and wealth. Those who were not enslaved but were poor were not considered white and were of a lower strata than whites who owned land (Rile Hayward 2013). With the abolition of slavery, the expressed (explicit) concern became not about wealth but now about race. This motivated a shift in how the institutions and how whites defined themselves. It became the focus for whites to be a higher status group, regardless of whether or not they were poorer than formerly enslaved Black Americans (Rile Hayward 2013).

Many conceptualizations of white political identity align with its primary source as concern for the group's political power. While a critical component of identity is a feeling of your fate as linked with others in your group (Leach et al. 2008), this linked fate is the connection between attitudes and material forms of threat but that identity, in and of itself, is better conceptualized as a result of cultural or ideational threat (Donnelly 2020). Aligning with the contextual account of racial identity for whites, many conceptualize white identity as an expressed desire for a recognition of privilege and status by institutions (López 2006) and a connection with the history that connects whites to their status (McDermott and Samson 2005). The prominent account of white political identity and its implications in contemporary American politics argues for a strong connection between white political identity and a desire to protect this racial hierarchy and status. This conceptualization reflects the tendency to report in-group-oriented, not necessarily out-group-oriented, attitudes (Jardina 2019).

As this is an open question and is rather a complex one, I attempt to wade into this by parsing through which of these

two perspectives are more plausible. In the next section, I examine whether attitudes expressing concern with economic or political threats are more robust correlates of the degree to which someone feels that being white is important to their identity. The goal is not to test particular causal mechanisms but to examine whether my belief and others' that white identity reflects reactions to political threats more than it does economic.

Data

To examine whether whites' attitudes expressing concern about political threats are more robust correlates than their perceptions of the economy, I use the 2012, 2016, and 2020 American National Election Studies (ANES). The studies are nationally representative questions that capture several questions that lend themselves to testing the relationship between economic and political concerns while also asking white respondents to report how important being white is to their identity.

Of course, I cannot *prove* that these questions reflect the two types of threats I purport to represent. This leaves any evidence of the particular mechanism connecting these predictors to reported levels of white identity – among other reasons – suspect. However, the goal of this paper is not necessarily to test a particular causal mechanism – this is a much larger project that we should continue to work on collectively. Nevertheless, the paper aims to offer an argument that contextualizes white identity in history and how that leads to different predictions about the correlates of white identity than many might assume from understanding the literature on the origins of social identity.

Each study asks white respondents, “How important is being white to your Identity” and offers them a five-point Likert scale with the options reflecting their level of agreement with the question ¹. They do not ask white respondents to report on items from the white political identity measures of Jardina (2019) or the more complex battery of items that Helms and Carter (1990) discusses. However, this question does capture the degree to which one sees their whiteness

1. Details about question-wording, original ANES coding, and my coding for all variables I include in my models are included in Table 1 located in the Appendix. I include descriptive statistics for the measures in these samples in Table 2, Table 3, and Table 4 in the Appendix

as part of who they are and reflects the conceptualization that others have used (see Schildkraut 2017; Jardina 2019; Pérez et al. 2022).

The ANES does not ask questions that explicitly capture the attitudes of respondents about perceptions of the economy for different groups. That is, there are no questions asking respondents about their perceptions of the economy for Black Americans relative to white Americans. The literature suggests, however, that those who have lower incomes are more likely to feel insecure about their financial status and are more likely to use this economic threat as a motivation for their negative attitudes toward immigrants (Pardos-Prado and Xena 2019). Therefore, one measure of economic threat that I use is the respondent's reported level of total family income. Those with lower levels of education are also likely to feel similar levels of insecurity about their economic status, as their ability to seek higher-paying jobs is likely more limited. Whites who report that they feel that their family is doing worse relative to the previous year and knowing others who lost a job also reflect a degree of insecurity about their economic status and may reflect perceptions of economic threat (Jardina 2019). I include an item reflecting respondents' perceptions of economic mobility today relative to the past. In 2012, whites harboring negative attitudes toward Obama rooted in racial animosity were more likely to express more negative views of the economy relative to how it had been in the past (Sides, Tesler, and Vavreck 2018). Others demonstrate that those who are reacting to economic threats and are concerned about their and other whites' status evaluate the economy and their economic position as worse (Baccini and Weymouth 2021). Therefore, I additionally include questions about retrospective sociotropic evaluations of the economy.

The ANES includes several questions that reflect perceptions of whites losing political power. First, I include a question that asks respondents to report whether or not they felt that whites currently have too little or too much influence on politics today. I also want to examine the degree to which respondents feel that the institutions benefit non-whites rather than whites by including the question that whites face discrimination and that the government is biased against whites. I include a question asking respondents about their belief that traditions in this country

are breaking². While these, like the economic attitudes, are just proxies of threat, they reflect the extent to which respondents may feel insecure about the political influence of whites.

Furthermore, the ANES includes the racial resentment battery, which asks questions reflecting whites' beliefs that Blacks are individually responsible for their circumstances and undeserving of assistance. I control for racial resentment as the individual items likely confound the relationship between both economic and political threat measures on expressed white identity. As economic evaluations often reflect partisans reporting a better economy when the government is run by co-partisans and a worse economy when the other party runs the government (Achen and Bartels 2017), I additionally control for a 7-point item measure of partisanship. Finally, I control for whether or not the respondent is female, as Sex affects one's socioeconomic status.

Evidence from a penalized regression

The summary of the literature points to a large number of potential predictors that influence the degree to which one takes on white political identity. To adjudicate between hypotheses, there is value in a comprehensive model that considers the plausibility of each. Most models consider those predictors they believe are most plausible and do not include others that the researcher believes to be less plausible (see Clarke and Primo 2012). They understandably do this out of concern with increases in the inefficiency of one's regression estimates: the more predictors you include, the risk of overfitting your model to your particular sample and from multicollinearity increases, thereby erroneously generalizing conclusions drawn from the fitted statistical model. Though this concern is reasonable, such a model neglects to examine the effects of the predictors a researcher believes to be less plausible.

Therefore, I use a penalized regression model to balance these concerns. The literature surveyed above suggests that there are a small number of variables that do predict white political identity, and there are many others that have no effect. Imagining an assumed hypothesis about the distribution of coefficients, this implies that the L2 regularization

2. The ANES only included this question in the 2012 and 2016 questionnaires

parameter equals 0. This is because the L2 regularization parameter shrinks coefficients asymptotically toward zero. The L1 regularization parameter, however, does allow for a shrinkage parameter to pull coefficients to 0. Given this, it suggests that the more appropriate model to capture this debate would be the LASSO penalized regression.

One option for implementation is to use the “classical” LASSO, which relies on an optimization procedure through cross-validation. It is a frequentist approach to LASSO penalized regression. However, there are several positive benefits to the second implementation of the LASSO penalized regression in a Bayesian framework (see Erp, Oberski, and Mulder 2019). From my review of the literature, I have *apriori* expectations for the distribution of effects on my predictors in the form of a Laplace distribution which I explicitly define in my prior density function. I also have the additional benefit of having the ability to define its hyperparameters. Meaning I suspect that there is a peak around zero and that there are relatively large tails. This characterization of my prior fits with that of the LASSO - a large number of included predictors have zero effect, with a relatively smaller proportion of included predictors that do have an effect but are relatively unsure as to how large those coefficients might be. I can specify the scale of that uncertainty by specifying a hyperparameter for this LASSO prior. Without having to perform cross-validation as I would with the classical LASSO, my posterior distribution naturally reflects the uncertainty around the size of my L2 shrinkage parameter and can estimate models with different L2 shrinkage parameters simultaneously. As **stan** includes implementations of this model, I rely on Hamiltonian Monte Carlo, a type of Markov Chain Monte Carlo sampling which allows for non-convexity in the L2 shrinkage parameters. As classical LASSO models are challenging to interpret due to cross-validation, using credible intervals offers a much less error-prone interpretation.

My measure of white identity is ordinal. Though white identity as a concept is likely continuous, through the specification of a cumulative density function with a logit link, my model bins this latent continuous space based on the number of response options. Though it creates a more complicated computational task, it is a more appropriate model specification given the available measure and my assumptions about the concept (Liddell and Kruschke 2018).

To account for potential problems stemming from missing data, I impute my data by using multiple imputation with

chained equations (MICE) where the particular model I use is a random forest model to allow for more flexibility and to reduce model dependence (Buuren and Groothuis-Oudshoorn 2011; Marbach 2022)³. I report the results of the model fitted using listwise deletion in Table 5, Table 6, and Table 7.

Results

Figure 1 presents the estimates of my fitted ordinal models with the LASSO shrinkage prior⁴. These models are fit on 6 chains with 2000 iterations each. To examine the evidence of whether a particular predictor “matters” for predicting white political identity, I examine the estimated credible intervals, which report the probability that the true estimate falls within the estimated range. I construct these credible intervals by reporting estimates between the 2.5% and 97.5% quantiles. This means that for a given credible interval, there is a two-tailed 95% probability that the true value falls within that range. For values outside the range of the credible interval, it is relatively implausible that the given value is the true value.

Examining the figure, it is readily apparent that in the face of indicators of attitudes about political status, economic attitudes inconsistently associate with white political identity. In all three years, reported attitudes about the nation’s economy being better relative to the previous year, reported feelings about the nation’s unemployment rate, income, concern about one’s family finances, and knowing someone who lost a job are not meaningfully correlated with white political identity. For other indicators of economic threat, none appear to be consistently meaningful. Retrospective evaluations of one’s financial situation and education matter in 2012, but these effects were not present in 2016 and 2020.

Turning to indicators of political threat, we still see significant inconsistency. In both 2012 and 2016, whether someone feels that new lifestyles are breaking society positively affects reported levels of white political identity. This

3. I impute 10 datasets and simultaneously fit my Bayesian LASSO on each dataset and pool the results. Uncertainty reflected in my reported credible intervals not only reflects the uncertainty from one model, but my reported uncertainty also reflects the uncertainty generated from the imputation procedure. One drawback with this particular approach is that my normalized split- \hat{R} will likely be greater than the widely recognized 1.01 cutoff that indicates model convergence due to the pooling of my multiply imputed datasets.

4. Table 8 presents the table of results.

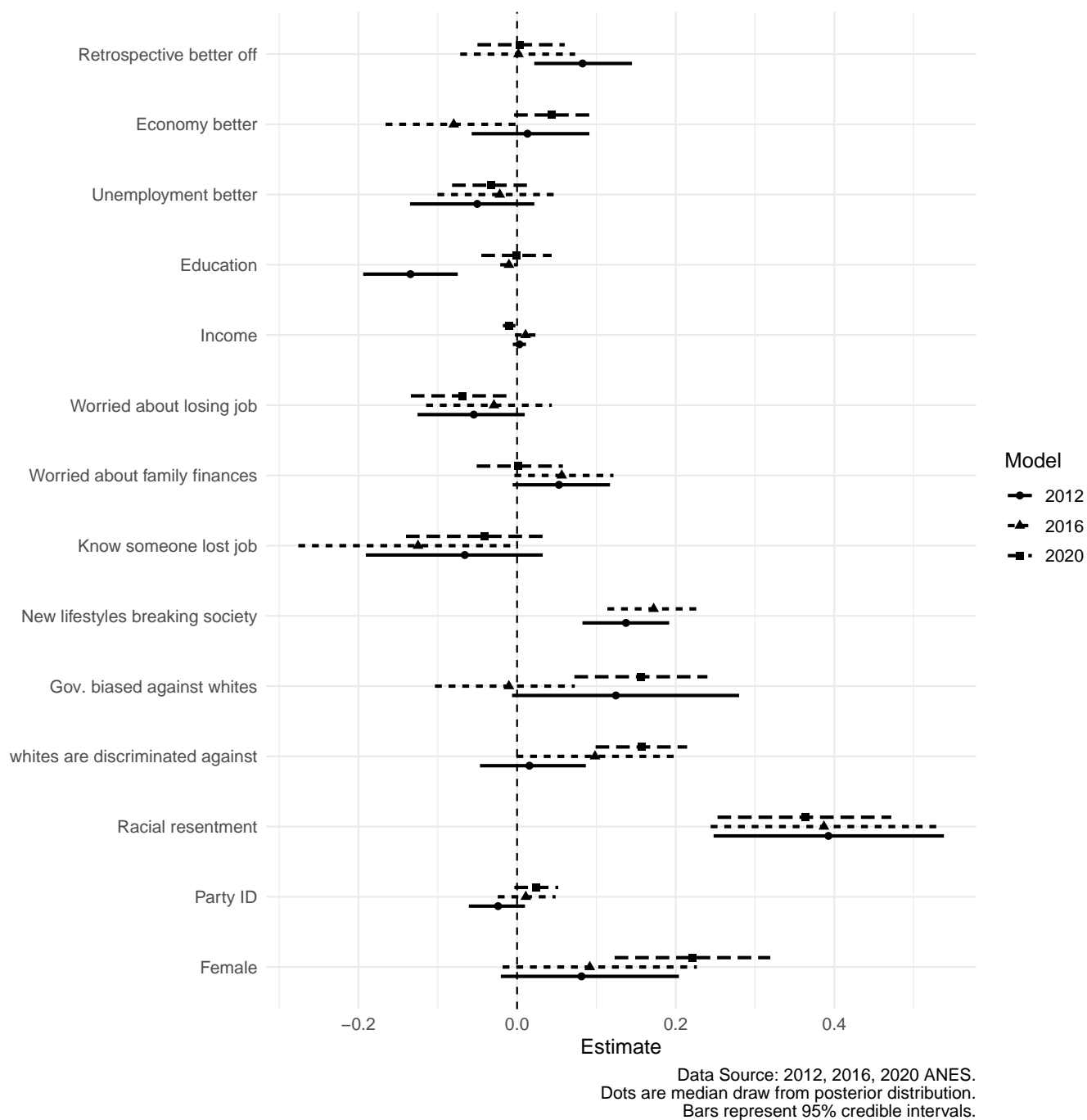


Figure 1: Economic factors are weak predictors of white identity

relationship is meaningful in 2012 and 2016 (this question is not present in the 2020 questionnaire). In 2020, those who reported that the government favored Black Americans also reported a higher white political identity, and those who feel that whites are discriminated against; however, neither of these effects are meaningful in 2012 or 2016.

Discussion

Overall, these results suggest that the effects of economic threat on white political identity are more robust in the pre-Trump era. Again, this is not confirmatory as there is no test of a causal mechanism for either economic or political threats. The goal of this analysis is to examine whether, in an omnibus test, either of these threats appears to be a more plausible predictor of white political identity. Though the relationship between political threats is inconsistent, their relationships appear more likely; particularly when we consider how many of these attitudes are explained by a rather complex data-generating process.

Economic threats tend to play a more significant part in explaining white identification in 2012. Political threats, however, appear to be a stronger predictor of white identification in the post-Trump era. In 2012, the U.S. economy was recovering from a massive recession and housing market collapse. However, as Obama's administration continued, conservatives became more outspoken about Obama's race (Tesler 2016). This shift to focus on Obama's race continued through the 2016 Presidential election and may reflect a much more explicit message to whites about political threats (Sides, Tesler, and Vavreck 2018). Though in 2020, the global economy faced significant uncertainty due to the COVID-19 pandemic, economic factors appear to be relatively less important. These results suggest that white political identity results from a political threat in the post-Trump era.

Conclusion

The literature on social identity suggests that the centrality of one's identity and group-based attitudes and behaviors result from economic threat. However, white political identity has a historical and institutional significance that, while

tied up in economics, reacts to changes in whites' influence over these institutions.

As pitting two theories against one another is a methodological challenge, I use data from the 2012, 2016, and 2020 ANES to demonstrate whether the relationships between attitudes proxying political and economic threats are more plausible. While this exercise is not and should not be interpreted as causal, I do demonstrate that economic factors appear to have a weaker relationship with white identity than political attitudes - particularly in the post-Trump era.

The narrative I advocate for suggests that white political identity is most responsive to threats of whites' institutional influence. The descriptive analysis of the relationships between political and economic threat fall short of causal identification. With the results of these analyses, scholars have evidence suggesting that economic threats appear to weakly associate with one's expression of a white political identity when including political threats in the same model.

The task for scholars is to examine whether political threats have any causal bearing on explaining the source of white political identity. This undoubtedly is a difficult task as threats are hard to treat experimentally and are reasonably suspect with observational data.

Appendix

Measures

Table 1: ANES Measures

Abbreviated				
	question	ANES Coding	Author coding	Years
white	How	1 = Extremely important; 2 = Very	1 = Not at all important; 2 = A little	2012,
identity	important is	important; 3 = Moderately important;	important; 3 = Moderately important; 4	2016,
	being white to	4 = A little important; 5 = Not at all	= Very important; 5 = Extremely	2020
	your identity	important; <1 Missing, not asked, etc	important; NA = Missing, not asked,	
			etc	

Economics

Retrospective	How much	1 = Much better; 2 = Somewhat	-2 = Much worse; -1 = Somewhat worse;	2012,
better off	better worse	better; 3 = The same; 4 = Somewhat	0 = The same; 1 = Somewhat better; 2	2016,
	off than 1	worse; 5 = Much worse, < 1 Missing,	= Much better; NA = Missing, not	2020
	year ago	not asked, etc	asked, etc	
Economy	U.S. economy	1 = Much better; 2 = somewhat	-2 = Much worse; -1 = Somewhat worse;	2012,
better	better worse	better; 3 = Stayed about the same, 4	0 = The same; 1 = Somewhat better; 2	2016,
	off than 1	= Somewhat worse; 5 = Much worse;	= Much better; NA = Missing, not	2020
	year ago	< 1 Missing, not asked, etc	asked, etc	


















Abbreviated		ANES Coding	Author coding	Years
	question			
Unemployment	Unemployment	1 = Much better; 2 = somewhat	-2 = Much worse; -1 = Somewhat worse;	2012,
better	better or	better; 3 = Stayed about the same, 4	0 = The same; 1 = Somewhat better; 2	2016,
	worse than 1	= Somewhat worse; 5 = Much worse;	= Much better; NA = Missing, not	2020
	year ago	< 1 Missing, not asked, etc	asked, etc	
Education	Educational	1 = < than High school; 2 = High	1 = < than High school; 2 = High	2012,
	attainment	school; 3 = Some post-High school; 4	school; 3 = Some post-High school; 4 =	2020
		= Bachelor's degree; < 1 Missing, not	Bachelor's degree; NA = Missing, not	
		asked, etc	asked, etc	
Education	Educational	See ANES codebook	See ANES codebook; NA = Missing,	2016
(2016)	attainment		not asked, etc	
Income	Total family	See ANES codebook; < Missing, not	See ANES codebook; NA = Missing,	2012,
	income	asked, etc	not asked, etc	2016,
				2020
Worried	How worried	1 = Not at all; 2 = A little; 3 =	-2 = Not at all; -1 = A little; 0 =	2012,
about	about losing	Moderately; 4 = Very; 5 = Extremely;	Moderately; 1 = Very; 2 = Extremely	2016,
losing job	job in near	< 1 Missing, not asked, etc		2020
	future			
Immigrants	How likely	1 = Extremely; 2 = Very; 3 =	1 = Not at all; 2 = Somewhat; 3 =	2012,
take jobs	immigration	Somewhat; 4 = Not at all; < 1	Very; 4 = Extremely; NA = Missing,	2016,
	will take away	Missing, not asked, etc	not asked, etc	2020
	jobs			

Abbreviated				
	question	ANES Coding	Author coding	Years
Worried	Worry about	1 = Extremely; 2 = Very; 3 =	1 = Not at all, 2 = A little; 3 =	2012,
about	family	Moderately; 4 = A little; 5 = Not at	Moderately; 4 = Very; 5 = Extremely;	2016,
family	financial	all; < 1 Missing, not asked, etc	NA = Missing, not asked, etc	2020
finances	situation			
Know	Know	1 = Someone lost job; 2 = No one lost	0 = No one lost job; 1 = Someone lost	2012,
someone	someone who	job; < 1 Missing, not asked, etc	job; NA = Missing, not asked, etc	2016,
lost job	lost job			2020
Political				
New	Newer	1 = Agree strongly; 2 = Agree	-2 = Disagree strongly; -1 = Disagree	2012,
lifestyles	lifestyles	somewhat; 3 = Neither; 4 = Disagree	somewhat; 0 = Neither, 1 = Agree	2016
breaking	breaking	somewhat; 5 = Disagree strongly; < 1	somewhat; 2 = Agree strongly	
society	down society	Missing, not asked, etc		

Abbreviated				
question		ANES Coding	Author coding	Years
Government Does the Ad- biased against whites (2012 only) Does the Federal Gov treat Blacks or whites Better	ministration favor Blacks or whites	1 = Favors whites; 2 = Favors Blacks; 3 = Treats both the same; < 1 Missing, not asked, etc	-1 = Favors whites; 0 = Treats both the same; 1 = Favors Blacks; NA = Missing, not asked, etc	2012, 2016, 2020
whites influence politics on U.S. politics	How much influence do whites have on U.S. politics	1 = Too much; 2 = Just about right; 3 = Too little; < 1 Missing, not asked, etc	-1 = Too little; 0 = Just about right; 1 = Too much; NA = Missing, not asked, etc	2012, 2016, 2020
whites are discrimi- nated against Controls	Discrimination in U.S. against whites	1 = A great deal; 2 = A lot; 3 = A moderate amount; 4 = A little; 5 = None at all; < 1 Missing, not asked, etc	1 = None at all; 2 = A little; 3 = A moderate amount; 4 = A lot; 5 = A great deal; NA = Missing, not asked, etc	2012, 2016, 2020

Abbreviated				
	question	ANES Coding	Author coding	Years
Racial Re- sentment	Standard, 4-item, battery	5-point Likert scale	Avg of four questions; NA = Missing, not asked, etc	2012, 2016, 2020
Party ID	Standard battery	1 = Strong Democrat; 2 = Democrat; 3 = Leans Democrat; 4 = Independent; 5 = Leans Republican; 6 = Republican; 7 = Strong Republican; < 1 Missing, not asked, etc	1 = Strong Democrat; 2 = Democrat; 3 = Leans Democrat; 4 = Independent/Neither; 5 = Leans Republican; 6 = Republican; 7 = Strong Republican; < 1 Missing, not asked, etc	2012, 2016, 2020
Female	Sex of respondent	1 = Male, 2 = Female; < 1 Missing, not asked, etc	0 = Male, 1 = Female; NA = Missing, not asked, etc	2012, 2016, 2020


















Table 2: 2012 ANES Descriptive Statistics

	Unique (#)	Missing (%)	Mean	SD	Min	Median	Max	
white identity	6	7	2.8	1.3	1.0	3.0	5.0	
Retrospective better off	6	1	-0.2	1.2	-2.0	-1.0	2.0	
Economy better	6	1	-0.3	1.1	-2.0	0.0	2.0	
Unemployment better	6	1	-0.3	1.1	-2.0	0.0	2.0	
Education	6	1	3.1	1.1	1.0	3.0	5.0	
Income	29	3	14.9	8.1	1.0	15.0	28.0	
Worried about losing job	6	45	-1.2	1.0	-2.0	-1.0	2.0	
immtakejobs	5	7	2.4	1.0	1.0	2.0	4.0	
Worried about family finances	6	7	2.7	1.2	1.0	3.0	5.0	
Know someone lost job	3	7	0.4	0.5	0.0	0.0	1.0	
New lifestyles breaking society	6	7	0.5	1.3	-2.0	1.0	2.0	
Government biased against whites	4	8	0.3	0.5	-1.0	0.0	1.0	
whites influence politics	4	7	0.1	0.5	-1.0	0.0	1.0	
whites are discriminated against	6	7	2.2	0.9	1.0	2.0	5.0	
Racial resentment	18	7	0.0	0.5	-2.0	0.0	2.0	
Party ID	8	0	4.2	2.1	1.0	4.0	7.0	
Female	2	0	0.5	0.5	0.0	1.0	1.0	

Data source: 2012 American National Election Study, unweighted.
 NAs are included in Unique column.

Descriptive statistics

Table 3: 2016 ANES Descriptive Statistics

	Unique (#)	Missing (%)	Mean	SD	Min	Median	Max	
white identity	6	15	2.6	1.3	1.0	3.0	5.0	
Retrospective better off	6	0	0.0	1.0	-2.0	0.0	2.0	
Economy better	6	0	-0.2	1.0	-2.0	0.0	2.0	
Unemployment better	6	0	0.0	1.0	-2.0	0.0	2.0	
Education	18	0	11.9	7.1	1.0	11.0	95.0	
Income	29	4	16.3	7.9	1.0	17.0	28.0	
Worried about losing job	6	36	-1.3	1.0	-2.0	-2.0	2.0	
immtakejobs	5	14	2.3	1.0	1.0	2.0	4.0	
Worried about family finances	6	14	2.6	1.2	1.0	3.0	5.0	
Know someone lost job	3	13	0.4	0.5	0.0	0.0	1.0	
New lifestyles breaking society	6	14	0.3	1.4	-2.0	1.0	2.0	
Government biased against whites	4	15	-0.1	0.7	-1.0	0.0	1.0	
whites influence politics	4	15	0.2	0.5	-1.0	0.0	1.0	
whites are discriminated against	5	18	2.0	0.8	1.0	2.0	4.0	
Racial resentment	18	14	0.0	0.5	-2.0	0.0	2.0	
Party ID	8	0	4.2	2.1	1.0	4.0	7.0	
Female	3	1	0.5	0.5	0.0	1.0	1.0	

Data source: 2016 American National Election Study, unweighted.

NAs are included in Unique column.

Table 4: 2020 ANES Descriptive Statistics

	Unique (#)	Missing (%)	Mean	SD	Min	Median	Max	
white identity	6	11	2.4	1.3	1.0	2.0	5.0	
Retrospective better off	6	0	0.1	1.0	-2.0	0.0	2.0	
Economy better	6	0	-0.7	1.3	-2.0	-1.0	2.0	
Unemployment better	6	0	-1.1	1.3	-2.0	-2.0	2.0	
Education	6	1	3.4	1.1	1.0	3.0	5.0	
Income	23	3	12.3	6.6	1.0	13.0	22.0	
Worried about losing job	6	33	-1.3	1.1	-2.0	-2.0	2.0	
immtakejobs	5	10	2.0	0.9	1.0	2.0	4.0	
Worried about family finances	6	0	2.1	1.1	1.0	2.0	5.0	
Know someone lost job	3	0	0.4	0.5	0.0	0.0	1.0	
Government biased against whites	4	12	-0.3	0.7	-1.0	0.0	1.0	
whites influence politics	4	12	0.3	0.6	-1.0	0.0	1.0	
whites are discriminated against	6	11	2.1	1.0	1.0	2.0	5.0	
Racial resentment	18	10	0.0	0.5	-2.0	0.0	2.0	
Party ID	8	0	4.2	2.3	1.0	4.0	7.0	
Female	3	0	0.5	0.5	0.0	1.0	1.0	

Data source: 2020 American National Election Study, unweighted.

NAs are included in Unique column.

Model results

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Table 5: Predictors of white Identity in 2012

	Model 1
Economic	
Retrospective better off	0.101 [0.016, 0.186]
Economy better	0.001 [−0.089, 0.097]
Unemployment better	−0.071 [−0.182, 0.021]
Education	−0.156 [−0.244, −0.071]
Income	0.004 [−0.008, 0.016]
Worried about losing job	−0.060 [−0.153, 0.018]
Worried about family finances	0.084 [0.000, 0.173]
Know someone lost job	−0.030 [−0.184, 0.087]
New lifestyles breaking society	0.114
Non-material	[0.037, 0.189]
Gov. biased against whites	0.050 [−0.080, 0.240]
whites are discriminated against	0.042 [−0.040, 0.144]
Racial resentment	0.271 [0.056, 0.487]
Demographics	
Party ID	−0.017 [−0.068, 0.029]
Female	0.141 [−0.011, 0.318]
Thresholds	
Threshold 1	−1.166 [−1.686, −0.630]
Threshold 2	−0.167 [−0.684, 0.362]
Threshold 3	1.028 [0.514, 1.560]
Threshold 4	2.173 [1.653, 2.719]
N	1664

Data source: 2012 American National Election Study.

95 percent credible intervals in brackets.

Median estimate from fitted model with 6 chains and 2000 iterations.

Table 6: Predictors of white identity in 2016

	Model 1
Economic	
Retrospective better off	0.538 [−0.534, 3.341]
Economy better	−0.128 [−2.108, 1.524]
Unemployment better	0.382 [−0.602, 1.873]
Education	−1.274 [−4.073, −0.096]
Income	0.315 [0.067, 0.880]
Worried about losing job	1.060 [−0.163, 4.003]
Worried about family finances	0.331 [−0.835, 1.976]
Know someone lost job	0.234 [−1.368, 3.292]
New lifestyles breaking society	0.361
Non-material	[−0.429, 1.811]
Gov. biased against whites	−3.138 [−9.409, −0.165]
whites are discriminated against	1.716 [−0.248, 6.738]
Racial resentment	0.250 [−1.428, 2.542]
Demographics	
Party ID	−0.089 [−0.724, 0.534]
Female	0.116 [−1.499, 2.720]
Thresholds	
Threshold 1	−8.243 [−34.758, 6.935]
Threshold 2	−7.129 [−33.375, 8.028]
Threshold 3	−4.645 [−29.365, 10.884]
Threshold 4	0.400 [−18.602, 17.738]
N	23

Data source: 2016 American National Election Study.

95 percent credible intervals in brackets.

Median estimate from fitted model with 6 chains and 2000 iterations.

Table 7: Predictors of white identity in 2020

	Model 1
Economic	
Retrospective better off	0.009 [−0.051, 0.077]
Economy better	0.107 [0.043, 0.170]
Unemployment better	−0.060 [−0.123, 0.000]
Education	0.024 [−0.031, 0.090]
Income	−0.008 [−0.019, 0.002]
Worried about losing job	−0.076 [−0.145, −0.012]
Worried about family finances	0.026 [−0.036, 0.102]
Know someone lost job	0.000 [−0.106, 0.108]
Gov. biased against whites	0.136 [0.029, 0.251]
whites are discriminated against	0.192 [0.115, 0.267]
Racial resentment	0.204 [0.057, 0.352]
Non-material	
Party ID	0.024
Demographics	
Female	[−0.011, 0.060] 0.297 [0.174, 0.428]
Thresholds	
Threshold 1	0.205 [−0.191, 0.623]
Threshold 2	1.121 [0.724, 1.546]
Threshold 3	2.400 [2.002, 2.830]
Threshold 4	3.549 [3.129, 3.999]
N	3377

Data source: 2020 American National Election Study.

95 percent credible intervals in brackets.

Median estimate from fitted model with 6 chains and 2000 iterations.

Table 8: Predictors of white identity in 2012, 2016, and 2020

	2012 ANES	2016 ANES	2020 ANES
Economic			
Retrospective better off	0.083 [0.022, 0.145]	0.002 [−0.072, 0.073]	0.004 [−0.050, 0.060]
Economy better	0.013 [−0.057, 0.091]	−0.080 [−0.166, 0.000]	0.044 [−0.004, 0.094]
Unemployment better	−0.050 [−0.135, 0.022]	−0.022 [−0.100, 0.046]	−0.033 [−0.082, 0.012]
Education	−0.134 [−0.194, −0.075]	−0.010 [−0.021, 0.000]	−0.001 [−0.045, 0.044]
Income	0.003 [−0.005, 0.011]	0.011 [−0.003, 0.026]	−0.010 [−0.018, −0.002]
Worried about losing job	−0.055 [−0.126, 0.010]	−0.029 [−0.115, 0.044]	−0.069 [−0.134, −0.009]
Worried about family finances	0.053 [−0.006, 0.117]	0.056 [−0.003, 0.121]	0.002 [−0.051, 0.058]
Know someone lost job	−0.066 [−0.191, 0.032]	−0.125 [−0.276, 0.001]	−0.041 [−0.140, 0.039]
New lifestyles breaking society	0.137	0.172	
Non-material			
Gov. biased against whites	[0.082, 0.192] 0.125 [−0.006, 0.280]	[0.114, 0.233] −0.010 [−0.104, 0.073]	0.156 [0.072, 0.240]
whites are discriminated against	0.016 [−0.047, 0.087]	0.098 [−0.001, 0.199]	0.157 [0.099, 0.214]
Racial resentment	0.393 [0.248, 0.538]	0.387 [0.244, 0.529]	0.364 [0.253, 0.472]
Demographics			
Party ID	−0.024 [−0.061, 0.010]	0.011 [−0.024, 0.049]	0.024 [−0.003, 0.052]
Female	0.081 [−0.021, 0.204]	0.092 [−0.018, 0.227]	0.221 [0.123, 0.319]
Thresholds			
Threshold 1	−1.392 [−1.746, −1.027]	−0.434 [−0.837, −0.003]	−0.162 [−0.475, 0.153]
Threshold 2	−0.447 [−0.796, −0.087]	0.341 [−0.053, 0.768]	0.692 [0.379, 1.008]
Threshold 3	0.747 [0.394, 1.108]	1.502 [1.098, 1.941]	1.855 [1.540, 2.175]
Threshold 4	1.988 [1.631, 2.365]	2.684 [2.257, 3.138]	3.030 [2.705, 3.356]
N	3509	3038	5963

Data Source: 2012, 2016, and 2020 American National Election Study.

95 percent credible intervals in brackets.

Median estimate from fitted model with 6 chains and 2000 iterations.

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