

# **MTH 9831: Homework 2**

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## Problem 1

### 'Only if' part:

According to the tower rule,

$$E \left[ \sum_{t=1}^N H_t \Delta X_t \right] = E \left[ E \left[ \sum_{t=1}^N H_t \Delta X_t | \mathcal{F}_{t-1} \right] \right] \quad (1)$$

according to linearity of expectation,

$$E \left[ E \left[ \sum_{t=1}^N H_t \Delta X_t | \mathcal{F}_{t-1} \right] \right] = E \left[ \sum_{t=1}^N E [H_t \Delta X_t | \mathcal{F}_{t-1}] \right] \quad (2)$$

since  $H_t$  is predictable,

$$E \left[ \sum_{t=1}^N E [H_t \Delta X_t | \mathcal{F}_{t-1}] \right] = E \left[ \sum_{t=1}^N H_t E [X_t - X_{t-1} | \mathcal{F}_{t-1}] \right] \quad (3)$$

since  $X_t$  is an adapted stochastic process

$$E \left[ \sum_{t=1}^N H_t E [X_t - X_{t-1} | \mathcal{F}_{t-1}] \right] = E \left[ \sum_{t=1}^N H_t (X_{t-1} - X_{t-1}) \right] = 0 \quad (4)$$

### 'If' part:

Pick an event  $A_s \in \mathcal{F}_{s-1}$  for  $A$  where  $s \in \{1, 2, \dots, N\}$ , then let

$$H_t = \begin{cases} 0, & \text{if } s \neq t \\ 1_{A_s}, & \text{if } s = t \end{cases} \quad (5)$$

then

$$\sum_{t=1}^N H_t \Delta X_t = 1_{A_s} (X_s - X_{s-1}) \quad (6)$$

We know  $E(1_{A_s}(X_s - X_{s-1})) = 0$ , therefore

$$E(1_{A_s} X_s) = E(1_{A_s} X_{s-1}) \quad (7)$$

that is,

$$\int_{A_s} X_s dP = \int_{A_s} X_{s-1} dP \quad (8)$$

Since  $A_s \in \mathcal{F}_{s-1}$  is arbitrary, then

$$E(X_s | \mathcal{F}_{s-1}) = E(X_{s-1} | \mathcal{F}_{s-1}) = X_{s-1} \quad (9)$$

$s$  is also arbitrary, therefore  $X_t$  is a martingale.

## Problem 2

(a)

According to the definition,

$$\{S \wedge T \leq t\} = \{S \leq t \text{ or } T \leq t\} = \{S \leq t\} \cup \{T \leq t\} \in \mathcal{F}_t \quad (10)$$

so  $S \wedge T$  is stopping time.

Similarly,

$$\{S \vee T \leq t\} = \{S \leq t \text{ and } T \leq t\} = \{S \leq t\} \cap \{T \leq t\} \in \mathcal{F}_t \quad (11)$$

so  $S \vee T$  is stopping time.

(b)

For each fixed  $\omega$ , I claim  $S(\omega) + T(\omega) < t$  iff there are positive rationals  $p, q$  with  $p + q \leq t$  and  $S(\omega) \leq p, T(\omega) \leq q$ . Suppose  $S(\omega) + T(\omega) \leq t$ ; we can find a rational  $r$  with  $S(\omega) + T(\omega) \leq r \leq t$ . Then  $S(\omega) \leq r - T(\omega)$ , so we can find  $p$  with  $S(\omega) \leq p \leq r - T(\omega)$ . Setting  $q = r - p$  we see that we have  $T(\omega) \leq q$ . The reverse implication is obvious.

Thus we have,

$$\{S + T \leq t\} = \bigcup_{p, q \in \mathbb{Q}^+, p+q \leq t} (\{S \leq p\} \cap \{T \leq q\}) \quad (12)$$

Since  $\{S \leq p\} \in \mathcal{F}_p \subseteq \mathcal{F}_t$  and  $\{T \leq q\} \in \mathcal{F}_q \subseteq \mathcal{F}_t$ . Thus  $\{S + T < t\}$  is a countable union of events from  $\mathcal{F}_t$ , and so it itself in  $\mathcal{F}_t$ .

Similarly for  $S \cdot T$ , we have (only in discrete case)

$$\{S \cdot T \leq t\} = \bigcup_{p, q \in \mathbb{Q}^+, pq \leq t} (\{S \leq p\} \cap \{T \leq q\}) \quad (13)$$

Same reason  $\{S \cdot T < t\}$  is a countable union of events from  $\mathcal{F}_t$ , and so it itself in  $\mathcal{F}_t$ .

## Problem 3

(a)

Since  $X_0 < a$  and  $\tau_a = \inf\{t \geq 0 : X_t \geq a\}$ , then

$$\{\tau_a = n\} = \{X_0, X_1, \dots, X_{n-1} < a, X_n \geq a\} = \bigcap_{j=0}^{n-1} \{X_j < a\} \cap \{X_n \geq a\} \quad (14)$$

by 2(a), it is a countable intersection of events from  $\mathcal{F}_t$ , and so it itself in  $\mathcal{F}_t$ .

(b)

Since  $X_0 > b$  and  $\tau_b = \inf\{t \geq 0 : X_t \leq b\}$ , then

$$\{\tau_b = n\} = \{X_0, X_1, \dots, X_{n-1} > b, X_n \leq b\} = \bigcap_{j=0}^{n-1} \{X_j > b\} \cap \{X_n \leq b\} \quad (15)$$

by 2(a), it is a countable intersection of events from  $\mathcal{F}_t$ , and so it itself in  $\mathcal{F}_t$ .

## Problem 4

(a)

$X_t = -B_t$  is a Brownian motion by the following facts:

- $X_0 = -B_0 = 0$
- $X_t$  is almost surely continuous since  $B_t$  is almost surely continuous.
- $X_t$  has independent increments since  $B_t$  has independent increments.
- $X_t - X_s = B_s - B_t \sim N(0, s - t)$  since  $B_t - X_s \sim N(0, t - s)$ .

(b)

$Y_t = \frac{1}{c}B_{c^2t}$  is a Brownian motion by the following facts:

- Clearly  $Y_0 = \frac{1}{c}B_0 = 0$
- If  $0 \leq t_1 \leq t_2 \leq \dots \leq t_n$ , then  $0 \leq c^2t_1 \leq \dots \leq c^2t_n$ . By the definition of Brownian motion,

$$B(c^2t_n) - B(c^2t_{n-1}), \dots, B(c^2t_2) - B(c^2t_1) \quad (16)$$

are independent random variables. It then follows that

$$\frac{1}{c}B(c^2t_n) - \frac{1}{c}B(c^2t_{n-1}), \dots, \frac{1}{c}B(c^2t_2) - \frac{1}{c}B(c^2t_1) \quad (17)$$

are independent random variables, i.e. the stochastic process  $\{Y_t | 0 \leq t\}$  has independent increments.

- The fact that  $Y_{t+h} - Y_t$  is normally distributed follows immediately from the fact that  $B(c^2t + c^2h) - B(c^2t)$  is normally distributed. Furthermore,

$$E[Y_{t+h} - Y_t] = E\left[\frac{1}{c}B(c^2t + c^2h) - \frac{1}{c}B(c^2t)\right] = \frac{1}{c}E[B(c^2t + c^2h) - B(c^2t)] = 0 \quad (18)$$

where the last equality follows from the definition of Brownian Motion. To show that the variance equals  $h$ , observe that

$$\text{var}[Y_{t+h} - Y_t] = \text{var}\left[\frac{1}{c}B(c^2t + c^2h) - \frac{1}{c}B(c^2t)\right] = \frac{1}{c^2}[B(c^2t + c^2h) - B(c^2t)] = \frac{1}{c^2}c^2h = h \quad (19)$$

- Because the function  $t \rightarrow B(t)$  is almost surely continuous, the function  $t \rightarrow Y_t = \frac{1}{c}B(c^2h)$  is the composition of (almost surely) continuous functions and is therefore almost surely continuous.

(c)

$Z_t = tB_{1/t}$  is a Brownian motion by the following facts:

- $Z_0 = 0$  given
- For Brownian motions,

$$\text{Cov}(B_t, B_{t+s}) = \text{Cov}(B_t, B_{t+s}B_t) + \text{Cov}(B_t, B_t) = t \quad (20)$$

for all  $t, s \geq 0$ . For our process  $Z_t$ , we compute the Covariance function for  $s < t$ ,

$$\text{Cov}[Z_t, Z_{t+s}] = \text{Cov}[tB_{1/t}, (t+s)B_{1/(t+s)}] = t(t+s)\text{Cov}(B_{1/t}, B_{1/(t+s)}) = t \quad (21)$$

So

$$\text{Cov}[Z_t, Z_{t+s} - Z_t] = t - t = 0 \quad (22)$$

Because the random variables  $Z_{t+s}$  and  $Z_t$  are normal,  $\text{Cov}(Z_t, Z_{t+s} - Z_t) = 0$  implies that  $Z_{t+s} - Z_t$  and  $Z_t$  are independent.

- And

$$\text{Var}(Z_{t+s} - Z_t) = \text{Var}(Z_{t+s}) + \text{Var}(Z_t) - 2\text{Cov}(Z_{t+s}, Z_t) = (t+s) + t - 2t = s \quad (23)$$

so our increments are independent and have the right variances.

- Continuity is clear for  $t > 0$ . We know that  $Z_t$  has the distribution of a Brownian motion on  $Q$ , so

$$0 = \lim_{n \rightarrow \infty} Z\left(\frac{1}{n}\right) = \lim_{t \rightarrow 0} Z_t \quad (24)$$

and we conclude that  $Z_t$  is continuous at  $t = 0$ , so  $Z_t$  satisfies the properties of a standard Brownian motion.

## Problem 5

(1)

To prove  $Z_t$  is a martingale, we need to show

- Integrability:

$$E|Z_t| = E|e^{\sigma B_t - \frac{\sigma^2 t}{2}}| \leq E|e^{\sigma B_t}| E|e^{-\frac{\sigma^2 t}{2}}| \leq M E|e^{-\frac{\sigma^2 t}{2}}| < \infty \quad (25)$$

- Adapted: Since  $B_t$  and  $t$  are adapted, so is their combination  $Z_t$ .
- The conditional expected value of the next generation:

$$E(Z_t | \mathcal{F}_s) = E(e^{\sigma(B_t - B_s)} e^{\sigma B_s} | \mathcal{F}_s) E(e^{-\frac{\sigma^2(t-s)}{2}} e^{-\frac{\sigma^2 s}{2}} | \mathcal{F}_s) \quad (26)$$

since  $B_t - B_s$  is independent of  $\mathcal{F}_s$  and  $B_s$  is measurable by  $\mathcal{F}_s$ , the equation above becomes

$$e^{\sigma B_s} E(e^{\sigma(B_t - B_s)}) e^{-\frac{\sigma^2(t-s)}{2}} e^{-\frac{\sigma^2 s}{2}} \quad (27)$$

Since  $B_t - B_s \sim N(0, t-s)$ , let  $B_t - B_s = x$ , we can show that

$$\begin{aligned} E(e^{\sigma x}) &= \frac{1}{\sqrt{2\pi(t-s)}} \int_{-\infty}^{\infty} e^{\sigma x} e^{-\frac{x^2}{2(t-s)}} \\ &= e^{\frac{\sigma^2(t-s)}{2}} \frac{1}{\sqrt{2\pi(t-s)}} \int_{-\infty}^{\infty} e^{-\frac{(x - \sigma(t-s))^2}{2(t-s)}} \\ &= e^{\frac{\sigma^2(t-s)}{2}} \end{aligned} \quad (28)$$

Therefore,

$$E(Z_t | \mathcal{F}_s) = e^{\sigma B_s} e^{\frac{\sigma^2(t-s)}{2}} e^{-\frac{\sigma^2(t-s)}{2}} e^{-\frac{\sigma^2 s}{2}} = e^{\sigma B_s - \frac{\sigma^2 s}{2}} = Z_s \quad (29)$$

$Z_t$  is a martingale.

(2)

To prove  $X_t$  is a martingale, we need to show

- Integrability:

$$E|X_t| = E|B_t^2 - t| \leq E|B_t^2| + E(t) < \infty \quad (30)$$

- Adapted: since  $B_t$  is adapted then so is  $X_t$ .
- The conditional expected value of the next generation:

$$\begin{aligned} E[B_t^2 - t | \mathcal{F}_s] &= E[(B_t - B_s) + B_s]^2 - ((t - s) + s) | \mathcal{F}_s \\ &= E[(B_t - B_s)^2 + 2(B_t - B_s)B_s + B_s^2 | \mathcal{F}_s] - [(t - s) + s] \end{aligned} \quad (31)$$

since  $B_t - B_s$  is independent of  $\mathcal{F}_s$  and  $B_s$  is measurable by  $\mathcal{F}_s$ , the equation above becomes

$$\begin{aligned} E[B_t^2 - t | \mathcal{F}_s] &= E[(B_t - B_s)^2] + 2B_s E(B_t - B_s) + B_s^2 - [(t - s) + s] \\ &= [t - s] + B_s^2 - [(t - s) + s] \\ &= B_s^2 - s \\ &= X_s \end{aligned} \quad (32)$$

$X_t$  is a martingale.

## Problem 6

Since  $f$  is of finite quadratic variation, then

$$V_f^2(T) < \infty \quad (33)$$

$g$  is of finite variation and continuous, then

$$V_g(T) < \infty \quad (34)$$

Furthermore,

$$V_f^p(T) = \lim_{\|\pi\| \rightarrow 0} \sum_{j=1}^n |f(t_j) - f(t_{j-1})|^p < \infty \quad (35)$$

where

$$\|\pi_n\| = \max_{j=1, \dots, n} \{t_j - t_{j-1}\}, \pi_n = \{0 = t_0 < t_1 < t_2 < \dots < t_n = T\} \quad (36)$$

The proof that the covariation of continuous finite variation process and finite quadratic variation is zero follows from the following inequality. Here,  $\pi_n$  is a partition of the interval  $[0, t]$ , and  $V_t(X)$  is the variation of  $X$  over  $[0, t]$ . Using Cauchy-Schwarz inequality,

$$\begin{aligned} [f, g](T) &= \lim_{\|\pi\| \rightarrow 0} \sum_{j=1}^n (f(X_{t_j}) - f(X_{t_{j-1}})) (g(X_{t_j}) - g(X_{t_{j-1}})) \\ &\leq \lim_{\|\pi\| \rightarrow 0} \sqrt{\sum_{j=1}^n (f(X_{t_j}) - f(X_{t_{j-1}}))^2} \sqrt{\sum_{j=1}^n (g(X_{t_j}) - g(X_{t_{j-1}}))^2} \\ &\leq \lim_{\|\pi\| \rightarrow 0} \sqrt{\sum_{j=1}^n (f(X_{t_j}) - f(X_{t_{j-1}}))^2} \sqrt{\sum_{j=1}^n (g(X_{t_j}) - g(X_{t_{j-1}}))^2} \end{aligned} \quad (37)$$

By the continuity and finite variation of  $g$ ,

$$\lim_{\|\pi\| \rightarrow 0} \sqrt{\sum_{j=1}^n (g(X_{t_j}) - g(X_{t_{j-1}}))^2} = 0 \quad (38)$$

and  $f$  is of finite quadratic variation,

$$\lim_{\|\pi\| \rightarrow 0} \sqrt{\sum_{j=1}^n (f(X_{t_j}) - f(X_{t_{j-1}}))^2} < \infty \quad (39)$$

So  $[f, g](T)$  vanishes in the limit as  $\|\pi_n\|$  goes to zero, i.e.,

$$[f, g](T) = \lim_{\|\pi\| \rightarrow 0} \sum_{j=1}^n (f(X_{t_j}) - f(X_{t_{j-1}})) (g(X_{t_j}) - g(X_{t_{j-1}})) = 0 \quad (40)$$

## Problem 7

(a)

Using Ito's lemma, in general suppose we have  $X_t = f(t, B_t)$ , for some function  $f(t, x) \in C^{1,2}$ , then

$$dX_t = f_t dt + f_x dB_t + \frac{1}{2} f_{xx} d[B, B]_t = (f_t + \frac{1}{2} f_{xx}) dt + f_x dB_t \quad (41)$$

In our case,

$$X_t = B_t^6 = f(t, B_t), \text{ i.e. } f(t, x) = x^6 \quad (42)$$

then

$$f_t = 0, f_x = 6x^5, f_{xx} = 30x^4 \quad (43)$$

therefore with  $X_t - X_0 = \int_0^t dX_s$  and  $dX_t = 6B_t^5 dB_t + 15B_t^4 dt$

$$X_t - X_0 = 6 \int_0^t B_s^5 dB_s + 15 \int_0^t B_s^4 ds \quad (44)$$

since  $B_s$  is Martingale. so  $(*)$  is also a martingale, which says it's 0.

Finally,

$$E(X_t) = 15 \int_0^t \left( \int_0^s B_s^4 ds \right) dP = 15 \int_0^t \left( \int_0^s B_s^4 dP \right) ds = 15 \int_0^t E(B_s^4) ds = 15t^3 \quad (45)$$

where we use Fubini's theorem and the kurtosis  $E(B_s^4) = 3s^2$ . Similarly, we can apply to the other one

$$E(|X|^3) = \sigma^3(2)!! \sqrt{\frac{2}{\pi}} = 2\sqrt{\frac{2}{\pi}} t^{3/2} \quad (46)$$

(b)

Given the expression of  $|C_B^n|^2$ , we have its expectation

$$E|C_B^n|^2 = E \left( \sum_{i=1}^n |B_{t_i} - B_{t_{i-1}}|^6 \right) + E \left( \sum_{i \neq j} |B_{t_i} - B_{t_{i-1}}|^3 |B_{t_j} - B_{t_{j-1}}|^3 \right) \quad (47)$$

Since  $B_{t_i} - B_{t_{i-1}}$  and  $B_{t_j} - B_{t_{j-1}}$  is independent for all  $i \neq j$ , therefore,

$$E|C_B^n|^2 = \sum_{i=1}^n E\left(|B_{t_i} - B_{t_{i-1}}|^6\right) + \sum_{i \neq j} E|B_{t_i} - B_{t_{i-1}}|^3 E|B_{t_j} - B_{t_{j-1}}|^3 \quad (48)$$

since  $B_{t_i} - B_{t_{i-1}} \sim B(0, t_i - t_{i-1})$  and using conclusion of (a),

$$E|C_B^n|^2 = 15 \sum_{i=1}^n (t_i - t_{i-1})^3 + \frac{8}{\pi} \sum_{i \neq j} (t_i - t_{i-1})^{3/2} (t_j - t_{j-1})^{3/2} \quad (49)$$

(c)

Since

$$\|\Pi_n\| = \max_{1 \leq i \leq n} \{t_i - t_{i-1}\} \quad (50)$$

then

$$\begin{aligned} E|C_B^n|^2 &\leq 15\|\Pi_n\|^2 \sum_{i=1}^n (t_i - t_{i-1}) + \frac{8}{\pi} \|\Pi_n\| \sum_{i \neq j} (t_i - t_{i-1})(t_j - t_{j-1}) \\ &\leq 15\|\Pi_n\|^2 \sum_{i=1}^n (t_i - t_{i-1}) + \frac{8}{\pi} \|\Pi_n\| \sum_i (t_i - t_{i-1}) \sum_j (t_j - t_{j-1}) \\ &\leq 15\|\Pi_n\|^2 T + \frac{8}{\pi} \|\Pi_n\| T^2 \end{aligned} \quad (51)$$

where obviously  $T = \sum_{i=1}^n (t_i - t_{i-1})$ , then

$$\lim_{\|\Pi_n\| \rightarrow 0} E|C_B^n|^2 \leq 15\|\Pi_n\|^2 T + \frac{8}{\pi} \|\Pi_n\| T^2 = 0 \quad (52)$$

which implies that  $\lim_{\|\Pi_n\| \rightarrow 0} E|C_B^n|^2 = 0$ . We can conclude that the cubic variation of Brownian motion in  $[0, T]$  is 0.

## Problem 8

According to the given expression,

$$\frac{S_{t_i}}{S_{t_{i-1}}} = e^{\sigma(B_{t_i} - B_{t_{i-1}}) + (\mu - \frac{\sigma^2}{2})(t_i - t_{i-1})} \quad (53)$$

Therefore,

$$\begin{aligned} \left(\log \frac{S_{t_i}}{S_{t_{i-1}}}\right)^2 &= \left[\sigma(B_{t_i} - B_{t_{i-1}}) + \left(\mu - \frac{\sigma^2}{2}\right)(t_i - t_{i-1})\right]^2 \\ &= \sigma^2(B_{t_i} - B_{t_{i-1}})^2 + 2\sigma(B_{t_i} - B_{t_{i-1}})\left(\mu - \frac{\sigma^2}{2}\right)(t_i - t_{i-1}) + \left(\mu - \frac{\sigma^2}{2}\right)^2(t_i - t_{i-1})^2 \end{aligned} \quad (54)$$

Furthermore it is easy to check that

$$\lim_{\|\Pi_n\| \rightarrow 0} \left(\log \frac{S_{t_i}}{S_{t_{i-1}}}\right)^2 = \sigma^2[B, B](T) + 2\sigma\left(\mu - \frac{\sigma^2}{2}\right)[B, t](T) + \left(\mu - \frac{\sigma^2}{2}\right)^2[t, t](T) \quad (55)$$

given the previous expression. Since  $B$  is of finite quadratic variation and  $t$  is of finite variation and continuous. According to the conclusion of problem 6, we have

$$\lim_{\|\Pi_n\| \rightarrow 0} \left(\log \frac{S_{t_i}}{S_{t_{i-1}}}\right)^2 = \sigma^2[B, B](T) = \sigma^2 T \quad (56)$$