Learning Theory - Homework 3

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Exercise 1 1

1) By simple computation we get:

$$q(\underline{x}_{t+1}|\underline{x}_t)p(\underline{x}_t) = \tilde{q}(\underline{x}_{t+1}|\underline{x}_t)p(\underline{x}_t)A(\underline{x}_{t+1},\underline{x}_t) = \min\{\tilde{q}(\underline{x}_{t+1}|\underline{x}_t)p(\underline{x}_t), \tilde{q}(\underline{x}_t|\underline{x}_{t+1})p(\underline{x}_{t+1})\}$$

$$(1)$$

$$q(\underline{x}_t|\underline{x}_{t+1})p(\underline{x}_{t+1}) = \tilde{q}(\underline{x}_t|\underline{x}_{t+1})p(\underline{x}_{t+1})A(\underline{x}_t,\underline{x}_{t+1}) = \min\{\tilde{q}(\underline{x}_t|\underline{x}_{t+1})p(\underline{x}_{t+1}), \tilde{q}(\underline{x}_{t+1}|\underline{x}_t)p(\underline{x}_t)\}$$

Therefore, $q(\underline{x}_{t+1}|\underline{x}_t)p(\underline{x}_t) = q(\underline{x}_t|\underline{x}_{t+1})p(\underline{x}_{t+1})$ i.e. detailed balance holds.

2) If \underline{x}' and \underline{x} differ at more than the chosen coordinate i, then $\frac{\tilde{q}(\underline{x}|\underline{x}')}{\tilde{q}(x'|x)}$ is not defined, so $A(\underline{x}', \underline{x}) = 1$.

If they differ at at position i, $\tilde{q}(\underline{x}'|\underline{x}) = p(x_i'|\{x_j\}_{j\neq i})$ and $\tilde{q}(\underline{x}|\underline{x}') = p(x_i|\{x\}_{j\neq i})$, so $\frac{\tilde{q}(\underline{x}|\underline{x}')p(\underline{x}')}{\tilde{q}(\underline{x}'|\underline{x})p(\underline{x})} = \frac{p(x_i|\{x_j\}_{j\neq i})p(\underline{x}')}{p(x_i'|\{x_j\}_{j\neq i})p(\underline{x})} = \frac{p(x_i|\{x_j\}_{j\neq i})p(x_i'|\{x_j\}_{j\neq i})p(x_i'|\{x_j\}_{j\neq i})p(\{x_j\}_{j\neq i})}{p(x_i'|\{x_j\}_{j\neq i})p(x_i|\{x_j\}_{j\neq i})p(x_i|\{x_j\}_{j\neq i})p(x_j\}_{j\neq i})} = 1.$ In conclusion, Gibbs sampling yields an acceptance probability $A(\underline{x}',\underline{x}) = 1$.

3) For this distribution, given states \underline{s} and \underline{s}' which differ at most at i, we have that we transition from \underline{s} to \underline{s}' with probability $\tilde{q}(\underline{s}'|\underline{s}) = p(s_i'|\{s_j\}_{j\neq i}) =$ $\frac{p(\underline{s}')}{\sum\limits_{\underline{s},\underline{t}} p(\{s_j\})}$. But,

$$\sum_{s_i} p(\{s_j\}) = \frac{1}{Z} \left[\exp\left\{ \sum_{\substack{(k,l) \in E \\ k,l \neq i}} J_{kl} s_k s_l + \sum_{(k,i) \in E} J_{ki} s_k + \sum_{(i,l) \in E} J_{il} s_l + \sum_{\substack{k \in V \\ k \neq i}} h_k s_k + h_i \right\} + \exp\left\{ \sum_{\substack{(k,l) \in E \\ k,l \neq i}} J_{kl} s_k s_l - \sum_{(k,i) \in E} J_{ki} s_k - \sum_{(i,l) \in E} J_{il} s_l + \sum_{\substack{k \in V \\ k \neq i}} h_k s_k - h_i \right\} \right]$$
(3)

Therefore,
$$\tilde{q}(\underline{s}'|\underline{s}) = \frac{\exp\{\sum\limits_{(k,i)\in E}J_{ki}s_ks_i' + \sum\limits_{(i,l)\in E}J_{il}s_i's_l + h_is_i'\}}{\exp\{\sum\limits_{(k,i)\in E}J_{ki}s_k + \sum\limits_{(i,l)\in E}J_{il}s_l + h_i\} + \exp\{-\sum\limits_{(k,i)\in E}J_{ki}s_k - \sum\limits_{(i,l)\in E}J_{il}s_l - h_i\}}$$
 and this is just $\frac{1}{Z}[1 \pm tanh(\sum\limits_{(k,i)\in E}J_{ki}s_k + \sum\limits_{(i,l)\in E}J_{il}s_l + h_i)]$ for $s_i' = \pm 1$. The term $\sum\limits_{(k,i)\in E}J_{ki}s_k + \sum\limits_{(i,l)\in E}J_{il}s_l = \sum\limits_{(i,j)\in E}J_{ij}s_j$, as $(k,i)\in E$ and $(i,l)\in E$ form the Markov blanket of i through vertices k and l .

2 Exercise 2

The KL-divergence between p and q is:

$$KL(p||q) = \mathbb{E}_p[\log p] - \mathbb{E}_p[\log q] = \mathbb{E}_p[x^T W x - \log Z_p(W)] - \mathbb{E}_p[x^T U x - \log Z_q(U)] \Rightarrow$$

$$\Rightarrow \underset{U}{\operatorname{arg \, min}} KL(p||q) = \underset{U}{\operatorname{arg \, max}} \{\mathbb{E}_p[x^T U x] - \log Z_q(U)\} \quad (4)$$

We have that:

$$x^{T}Ux = \begin{bmatrix} x_{1} & x_{2} & \dots & x_{D} \end{bmatrix} \begin{bmatrix} U_{11} & U_{12} & \dots & U_{1D} \\ U_{21} & U_{22} & \dots & U_{2D} \\ \vdots & \vdots & \dots & \vdots \\ U_{n1} & U_{n2} & \dots & U_{DD} \end{bmatrix} \begin{bmatrix} x_{1} \\ x_{2} \\ \vdots \\ x_{D} \end{bmatrix} = \sum_{i=1}^{D} \sum_{j=1}^{D} U_{ij}x_{i}x_{j}$$
(5)

$$Tr(UC) = \sum_{i=1}^{D} U_{1i} \mathbb{E}_p[x_i x_1] + \dots + \sum_{i=1}^{D} U_{Di} \mathbb{E}_p[x_i x_D] = \sum_{i=1}^{D} \sum_{j=1}^{D} U_{ij} \mathbb{E}_p[x_i x_j]$$
 (6)

From this we get that $\mathbb{E}_p[x^TUx] = Tr(UC)$. Therefore $KL(p||q) = \arg\max_{U} \{Tr(UC) - \log Z_q(U)\}$.

3 Exercise 3

Note: When using the Naive Bayes Classifier, we should actually compare $p(\underline{x}^*, \text{class} = 0)$ and $p(\underline{x}^*, \text{class} = 1)$ to establish the class that the sample was extracted from. This is because $p(\underline{x}^*|\text{class} = 0)$ and $p(\underline{x}^*|\text{class} = 1)$ would only consider how well the sample would fit within the class, without accounting for how frequent the class itself is.

The inequality $p(\underline{x}^*, \text{class} = 0) > p(\underline{x}^*, \text{class} = 1)$ is equivalent to $\log p(\underline{x}^*, \text{class} = 0) > \log p(\underline{x}^*, \text{class} = 1)$. Expanding the probabilities, we get:

$$\log p_1 + \sum_{i=1}^K \log p(x_i^*|\text{class} = 1) > \log p_0 + \sum_{i=1}^K \log p(x_i^*|\text{class} = 0)$$
 (7)

Using the notations in the statement, we have that:

$$\log p(x_i^*|\text{class} = 1) = x_i^* \log \theta_i^1 + (1 - x_i^*) \log(1 - \theta_i^1) = x_i^* \log \frac{\theta_i^1}{1 - \theta_i^1} + \log(1 - \theta_i^1)$$

$$\log p(x_i^*|\text{class} = 0) = x_i^* \log \theta_i^0 + (1 - x_i^*) \log(1 - \theta_i^0) = x_i^* \log \frac{\theta_i^0}{1 - \theta_i^0} + \log(1 - \theta_i^0)$$
(8)

Therefore, we get:

$$\log p_{1} + \sum_{i=1}^{K} \log(1-\theta_{i}^{1}) + \sum_{i=1}^{K} x_{i}^{*} \log \frac{\theta_{i}^{1}}{1-\theta_{i}^{1}} > \log p_{0} + \sum_{i=1}^{K} \log(1-\theta_{i}^{0}) + \sum_{i=1}^{K} x_{i}^{*} \log \frac{\theta_{i}^{0}}{1-\theta_{i}^{0}} \Leftrightarrow \sum_{i=1}^{K} x_{i}^{*} \log \frac{\theta_{i}^{1}(1-\theta_{i}^{0})}{\theta_{i}^{0}(1-\theta_{i}^{1})} + \log \frac{p_{1}}{p_{0}} + \sum_{i=1}^{K} \log \frac{1-\theta_{i}^{1}}{1-\theta_{i}^{0}} > 0 \quad (9)$$

We therefore see that we can write the condition of classifying \underline{x}^* in class 1 as $w^T\underline{x}^*+b>0$, where:

$$w = \left[\log \frac{\theta_1^1 (1 - \theta_1^0)}{\theta_1^0 (1 - \theta_1^1)} \dots \log \frac{\theta_K^1 (1 - \theta_K^0)}{\theta_K^0 (1 - \theta_K^1)} \right], b = \log \frac{p_1}{p_0} + \sum_{i=1}^K \log \frac{1 - \theta_i^1}{1 - \theta_i^0}$$
(10)

4 Exercise 4

1) The Bayesian network looks as follows:

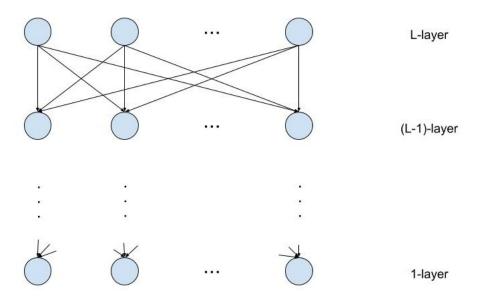


Figure 1: Sigmoid Belief Network

2) Computing $p(\mathbf{x}^0)$ implies computing $\sum_{\mathbf{x}^1} ... \sum_{\mathbf{x}^L} p(\mathbf{x}^0, \mathbf{x}^1, ..., \mathbf{x}^L)$ for a fixed value of \mathbf{x}^0 . The complexity of computing a probability $p(x_i^{l-1}|\mathbf{x}^l)$ is O(w) and we have w such probabilities to compute per layer. This gives a complexity of $O(w^2)$ per layer of computing $p(\mathbf{x}^{l-1}|\mathbf{x}^l)$. Due to the structure of the Bayesian

Network, we can decompose our computation into:

$$\sum_{\mathbf{x}^1} \dots \sum_{\mathbf{x}^L} p(\mathbf{x}^0, \mathbf{x}^1, \dots, \mathbf{x}^L) = \sum_{\mathbf{x}^1} p(\mathbf{x}^0 | \mathbf{x}^1) \sum_{\mathbf{x}^2} p(\mathbf{x}^1 | \mathbf{x}^2) \dots \sum_{\mathbf{x}^L} p(\mathbf{x}^{L-1} | \mathbf{x}^L) p(\mathbf{x}^L)$$
(11)

Each \mathbf{x}^i takes 2^w values. To compute $f_{L-1}(\mathbf{x}^{L-1}) = \sum_{\mathbf{x}^L} p(\mathbf{x}^{L-1}|\mathbf{x}^L) p(\mathbf{x}^L)$ for all values of \mathbf{x}^{L-1} requires a complexity of $O(w^2 2^{2w})$. The same applies then for computing $f_{L-2} = \sum_{\mathbf{x}^{L-1}} p(\mathbf{x}^{L-2}|\mathbf{x}^{L-1}) f_{L-1}(\mathbf{x}^{L-1})$ and so on.

The overall complexity is therefore $O(Lw^22^{2w})$.

3) The energy term for the Variational EM procedure is:

$$\mathbb{E}_{q}[\log p(\mathbf{x}^{1}, ..., \mathbf{x}^{L}, \mathbf{x}^{0})] = \sum_{\mathbf{x}^{1}, ..., \mathbf{x}^{L}} \prod_{l=1}^{L} \prod_{i=1}^{w} q(x_{i}^{l}) \log p(\mathbf{x}^{1}, ..., \mathbf{x}^{L}, \mathbf{x}^{0})$$
(12)

In this case, even if splitting $\log p(\mathbf{x}^1,...,\mathbf{x}^L,\mathbf{x}^0)$ into its constituents, we can not find a smart factorization, so we need to compute every term in the sum separately. There are 2^{Lw} terms. $\log p(\mathbf{x}^1,...,\mathbf{x}^L,\mathbf{x}^0)$ takes $O(Lw^2)$ operations to compute and $\prod_{l=1}^L \prod_{i=1}^w q(x_i^l)$ takes O(Lw). Therefore, the total complexity is $O(Lw^22^{Lw})$.

5 Exercise 5

Our objective is to maximize $\sum_{n=1}^{N}\log p(\mathbf{x}^{(n)}) = \sum_{n=1}^{N}\log \sum_{k=1}^{H}p(\mathbf{x}^{(n)}|\mathbf{m}_{k},\sigma_{k}^{2})p(k),$ where we have considered the training set $\{\mathbf{x}^{(n)}\}_{n=1}^{N}$. Given the probability distributions $\{\{q_{nk}^{(t)}\}_{k=1}^{H}\}_{n=1}^{N}$ at step t in the EM algorithm, by convexity we have $\sum_{n=1}^{N}\log \sum_{k=1}^{H}p(\mathbf{x}^{(n)}|\mathbf{m}_{k},\sigma_{k}^{2})p(k)\geq \sum_{n=1}^{N}\sum_{k=1}^{H}q_{nk}^{(t)}\log \frac{p(k)p(\mathbf{x}^{(n)}|\mathbf{m}_{k},\sigma_{k}^{2})}{q_{nk}^{(t)}}$ and we get equality for $q_{nk}^{(t)}=\frac{p(k)p(\mathbf{x}^{(n)}|\mathbf{m}_{k}^{(t)},(\sigma_{k}^{2})^{(t)})}{\sum\limits_{k=1}^{H}p(k)p(\mathbf{x}^{(n)}|\mathbf{m}_{k}^{(t)},(\sigma_{k}^{2})^{(t)})}$. This is the expectation step in the EM algorithm. We now fix $q_{nk}^{(t)}$ and optimize with respect to p(k), \mathbf{m}_{k} and σ_{k}^{2} , which means that we optimize $\sum_{n=1}^{N}\sum_{k=1}^{H}q_{nk}^{(t)}[\log p(k)-\log q_{nk}^{(t)}+\log p(\mathbf{x}^{(t)}|\mathbf{m}_{k},\sigma_{k}^{2})].$ Optimizing with respect to \mathbf{m}_{k} and σ_{k}^{2} implies maximizing $\sum_{n=1}^{N}\sum_{k=1}^{H}q_{nk}^{(t)}\log p(\mathbf{x}^{(n)}|\mathbf{m}_{k},\sigma_{k}^{2})=\sum_{n=1}^{N}\sum_{k=1}^{H}q_{nk}^{(t)}[-\frac{D}{2}\log(2\pi\sigma_{k}^{2})-\frac{1}{2\sigma_{k}^{2}}(\mathbf{x}^{(n)}-\mathbf{m}_{k})^{T}(\mathbf{x}^{(n)}-\mathbf{m}_{k})].$

Differentiating with respect to \mathbf{m}_k gives:

$$\sum_{n=1}^{N} q_{nk}^{(t)} \frac{1}{\sigma_k^2} (\mathbf{x}^{(n)} - \mathbf{m}_k) = 0 \Rightarrow \mathbf{m}_k^{(t+1)} = \frac{\sum_{n=1}^{N} q_{nk}^{(t)} \mathbf{x}^{(n)}}{\sum_{n=1}^{N} q_{nk}^{(t)}}$$
(13)

Differentiating with respect to σ_k^2 yields:

$$\sum_{n=1}^{N} q_{nk}^{(t)} \left[-\frac{D}{2\sigma_k^2} + \frac{1}{2\sigma_k^2} (\mathbf{x}^{(n)} - \mathbf{m}_k)^T (\mathbf{x}^{(n)} - \mathbf{m}_k) \right] \Rightarrow (\sigma_k^2)^{(t+1)} = \frac{\sum_{n=1}^{N} q_{nk}^{(t)} (\mathbf{x}^{(n)} - \mathbf{m}_k^{(t)})^T (\mathbf{x}^{(n)} - \mathbf{m}_k^{(t)})}{D\sum_{n=1}^{N} q_{nk}^{(t)}}$$

$$(14)$$

Putting it all together, the update rule of \mathbf{m}_k and σ_k^2 in the M-step of the EM algorithm is:

$$\mathbf{m}_{k}^{(t+1)} = \frac{\sum_{n=1}^{N} q_{nk}^{(t)} \mathbf{x}^{(n)}}{\sum_{n=1}^{N} q_{nk}^{(t)}}; (\sigma_{k}^{2})^{(t+1)} = \frac{\sum_{n=1}^{N} q_{nk}^{(t)} (\mathbf{x}^{(n)} - \mathbf{m}_{k}^{(t)})^{T} (\mathbf{x}^{(n)} - \mathbf{m}_{k}^{(t)})}{D \sum_{n=1}^{N} q_{nk}^{(t)}}$$
(15)

6 Exercise 6

1) The log-likelihood is:

$$L(W) = \log p(\underline{v}^{(1)}, \underline{v}^{(2)}, ..., \underline{v}^{(N)}|W) = \sum_{n=1}^{N} \log p(\underline{v}^{(n)}|W) = \sum_{n=1}^{N} \log \sum_{\underline{h}} p(\underline{h}, \underline{v}^{(n)}|W)$$

$$(16)$$

We have that $p(\underline{h}, \underline{v}|W) = \frac{1}{Z} \exp\{\sum_{i=1}^K \sum_{j=1}^M W_{ij} v_i h_j\}$, with $Z = \sum_{h,v} W_{ij} v_i h_j$, so:

$$\frac{\partial}{\partial W_{ij}} L(W) = \sum_{n=1}^{N} \left(\frac{\partial}{\partial W_{ij}} \log \sum_{\underline{h}} \exp\{\sum_{i=1}^{K} \sum_{j=1}^{M} W_{ij} v_i^{(n)} h_j\} - \frac{\partial}{\partial W_{ij}} \log Z \right)$$
(17)

We now compute each of the partial derivative terms:

$$\frac{\partial}{\partial W_{ij}} \log Z = \frac{1}{Z} \sum_{\underline{h},\underline{v}} v_i h_j \exp\{\sum_{i=1}^K \sum_{j=1}^M W_{ij} v_i h_j\} = \sum_{\underline{h},\underline{v}} v_i h_j p(\underline{h},\underline{v}|W) = \langle v_i h_j \rangle$$
(18)

$$\frac{\partial}{\partial W_{ij}} \log \sum_{\underline{h}} \exp\{\sum_{k=1}^{K} \sum_{l=1}^{M} W_{kl} v_{k}^{(n)} h_{l}\} = \frac{1}{\sum_{\underline{h}} \exp\{\sum_{k=1}^{K} \sum_{l=1}^{M} W_{kl} v_{k}^{(n)} h_{l}\}} \sum_{\underline{h}} v_{i}^{(n)} h_{j} \exp\{\sum_{k=1}^{K} \sum_{l=1}^{M} W_{kl} v_{k}^{(n)} h_{l}\} = \frac{1}{\sum_{\underline{h}} \exp\{\sum_{k=1}^{K} \sum_{l=1}^{M} W_{kl} v_{k}^{(n)} h_{l}\}} \sum_{\underline{h}} v_{i}^{(n)} h_{j} \exp\{\sum_{k=1}^{K} \sum_{l=1}^{M} W_{kl} v_{k}^{(n)} h_{l}\} \sum_{\underline{h}} \exp\{\sum_{k=1}^{K} \sum_{l=1}^{M} W_{kl} v_{k}^{(n)} h_{l}\} = \sum_{\underline{h}} v_{i}^{(n)} h_{j} p(h_{j} | \underline{v}^{(n)}, W) = \mathbb{E}_{p(h_{j} | \underline{v}^{(n)}, W)} [v_{i}^{(n)} h_{j}] \quad (19)$$

In conclusion,

$$\frac{\partial}{\partial W_{ij}} L(W) = \sum_{n=1}^{N} \mathbb{E}_{p(h_j|\underline{v}^{(n)},W)} [v_i^{(n)} h_j] - \langle v_i h_j \rangle$$
 (20)

2) We have that:

$$p(h_{j}, \underline{v}, W) = \frac{1}{Z} \sum_{\substack{h_{k} \\ k \neq j}} \exp\{\sum_{i=1}^{K} W_{ij} v_{i} h_{j} + \sum_{\substack{k=1 \\ k \neq j}}^{M} \sum_{i=1}^{K} W_{ik} v_{i} h_{k}\} \Rightarrow$$

$$\Rightarrow p(h_{j} | \underline{v}, W) = \frac{p(h_{j}, \underline{v} | W)}{\sum_{\underline{h}} p(h_{j}, \underline{v} | W)} = \frac{\exp\{\sum_{i=1}^{K} W_{ij} v_{i} h_{j}\}}{\sum_{\underline{h}, \underline{v} \in \{-1, 1\}} \exp\{\sum_{i=1}^{K} W_{ik} v_{i} h_{k}\}}$$
(21)

Therefore, we get:

$$\mathbb{E}_{p(h_{j}|v_{i}^{(n)},W)}[v_{i}^{(n)}h_{j}] = v_{i}^{(n)}[p(h_{j}=1|v_{i}^{(n)},W) - p(h_{j}=-1|v_{i}^{(n)},W)] =$$

$$= v_{i}^{(n)} \frac{\exp\{\sum_{k=1}^{K} W_{kj}v_{k}^{(n)}\} - \exp\{\sum_{k=1}^{K} -W_{kj}v_{k}^{(n)}\}}{\exp\{\sum_{k=1}^{K} W_{kj}v_{k}^{(n)}\} + \exp\{\sum_{k=1}^{K} -W_{kj}v_{k}^{(n)}\}} = v_{i}^{(n)} \tanh(\sum_{i=1}^{K} W_{kj}v_{k}^{(n)}) \quad (22)$$

Finally, using this, we get the desired result:

$$\frac{\partial}{\partial W_{ij}}L(W) = \sum_{n=1}^{N} \left(v_i^{(n)} \tanh(\sum_{k=1}^{K} W_{kj} v_k^{(n)}) - \langle v_i h_j \rangle \right)$$
(23)