## Lab 2: Mixed-effects models

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### Laws of probability

1. Axiom of conditional probability

$$Pr(X,Y) = Pr(Y|X) Pr(X)$$

Often easier to specify conditional probabilities than joint probabilities

2. Law of total probability

$$Pr(X) = \int Pr(X, Y) dY$$

Used when justifying hierarchical models

### Laplace approximation

– Define joint log-likelihood:

$$f(\theta, \varepsilon; y) = \log(\Pr(y|\theta_1, \varepsilon) \Pr(\varepsilon|\theta_2))$$

Taylor series expansion of joint log-likelihood

$$f(\varepsilon|\theta,y) \approx f(\hat{\varepsilon}|\theta) + f'(\hat{\varepsilon}|\theta)(\hat{\varepsilon}-\varepsilon) + \frac{1}{2}f''(\hat{\varepsilon}|\theta)(\hat{\varepsilon}-\varepsilon)^{2}$$

Evaluate Taylor series around "inner maximum"

$$\hat{\varepsilon} = \operatorname{argmax}_{\varepsilon}(f(\theta, \varepsilon))$$

- Implies that  $f'(\hat{\varepsilon}|\theta) = 0$
- Approximate joint likelihood via Taylor series expansion

$$\Pr(y|\theta_1,\varepsilon)\Pr(\varepsilon|\theta_2) = e^{f(\varepsilon|\theta)} \approx e^{f(\hat{\varepsilon}|\theta) - \frac{1}{2}|f''(\hat{\varepsilon})|(\hat{\varepsilon} - \varepsilon)^2}$$

#### Laplace approximation

Approximate joint likelihood via Taylor series expansion

$$\Pr(y|\theta_1,\varepsilon)\Pr(\varepsilon|\theta_2) = e^{f(\varepsilon|\theta)} \approx e^{f(\hat{\varepsilon}|\theta) - \frac{1}{2}|f''(\hat{\varepsilon})|(\hat{\varepsilon} - \varepsilon)^2}$$

Integrate both sides

$$\int \Pr(y|\theta_1, \varepsilon) \Pr(\varepsilon|\theta_2) d\varepsilon = \int e^{f(\varepsilon|\theta)} d\varepsilon$$

$$\int \Pr(y|\theta_1, \varepsilon) \Pr(\varepsilon|\theta_2) d\varepsilon \approx \int e^{f(\hat{\varepsilon}|\theta)} e^{-\frac{1}{2}|f''(\hat{\varepsilon})|(\hat{\varepsilon}-\varepsilon)^2} d\varepsilon$$

• And  $e^{f(\hat{\varepsilon}|\theta)}$  is a constant so:

$$\int e^{f(\hat{\varepsilon}|\theta)} e^{-\frac{1}{2}|f''(\hat{\varepsilon})|(\hat{\varepsilon}-\varepsilon)^2} \mathrm{d}\varepsilon = e^{f(\hat{\varepsilon}|\theta)} \int e^{-\frac{1}{2}|f''(\hat{\varepsilon})|(\hat{\varepsilon}-\varepsilon)^2} \mathrm{d}\varepsilon$$

- Looks like a normal distribution
  - $\hat{\varepsilon}$  is the mean of the normal distribution
  - $f''(\hat{\varepsilon})$  is the hessian of the normal distribution  $(f''(\hat{\varepsilon}) = \Sigma^{-1})$

Normal PDF: 
$$\Pr(\varepsilon|\mu,\Sigma) = \frac{1}{(2\pi)^{k/2}|\Sigma|^{1/2}} \exp\left(\frac{-(\varepsilon-\mu)^{\mathrm{T}}\Sigma^{-1}(\varepsilon-\mu)}{2}\right)$$

### **Chi-squared example**

 $X \sim Chi.Squared(k)$ 

$$\Pr(X=x) \propto x^{\frac{k}{2}-1} e^{\frac{-x}{2}}$$

Defining the log-likelihood

$$\log(\Pr(x)) \propto f(x)$$

Taking derivatives:

$$f(x) \propto \left(\frac{k}{2} - 1\right) \log(x) - \frac{x}{2}$$

$$f'(x) \propto \left(\frac{k}{2} - 1\right) x^{-1} - \frac{1}{2}$$

$$f''(x) \propto -\left(\frac{k}{2} - 1\right)x^{-2}$$

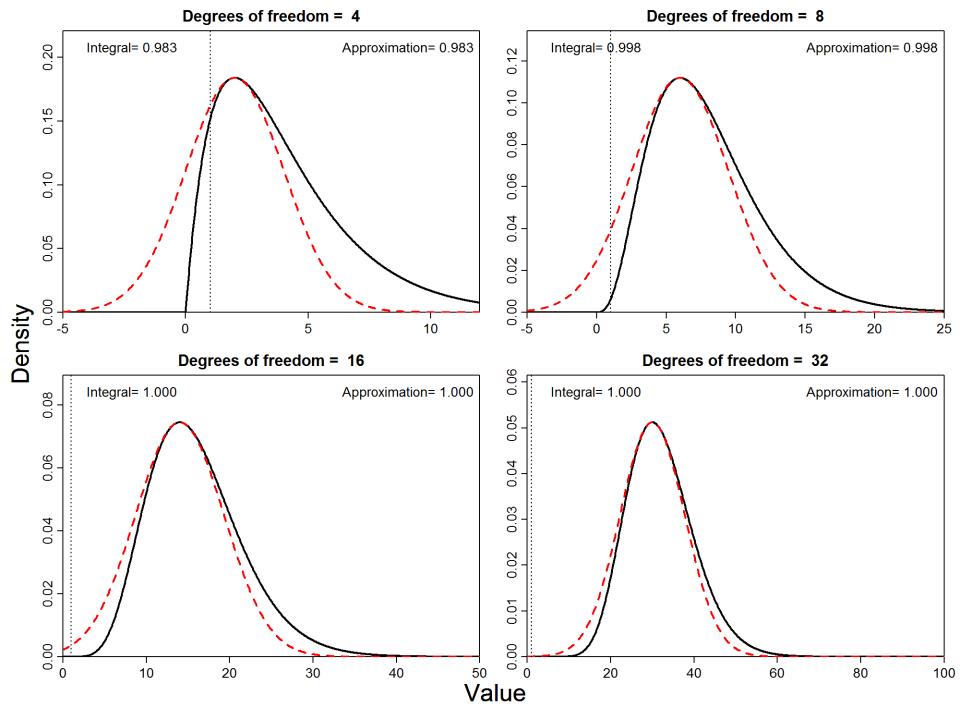
Solving for mode and Hessian:

$$f'(x) = 0 \quad \to \quad \hat{x} = k - 2$$

$$f''(\hat{x}) = -\left(\frac{1}{2(k-2)}\right)$$

Hence:

$$\Pr(x) \propto Normal(k-2, 2(k-2))$$



#### **Bottom line**

$$\ln L(\theta; y) \equiv \int \Pr(y, \varepsilon | \theta) \, d\varepsilon \cong \log(\Pr(y, \varepsilon | \theta)) - \frac{1}{2} \log(|\mathbf{H}|)$$

- Where

$$Pr(y, \varepsilon | \theta) = Pr(y | \theta_1, \varepsilon) Pr(\varepsilon | \theta_2)$$

And

$$\mathbf{H} = \frac{\partial^2}{\partial \varepsilon^2} (\log(\Pr(y, \varepsilon | \theta)))$$

- Definitions
  - $-\log(L(\theta;y))$  is the marginal log-likelihood
  - $Pr(y, \varepsilon | \theta)$  is the joint likelihood
  - |H| is the determinant of the Hessian matrix

#### Generalized linear mixed model

1. Specify distribution for response variable

$$c_i \sim Poisson(\lambda_i)$$

2. Specify function for expected value

$$g^{-1}(\lambda_i) = x_0 + \mathbf{x}_i^T \mathbf{\beta} + \mathbf{z}_i^T \mathbf{\epsilon}$$

Specify a link function

$$g^{-1}(a) = \log(a) \rightarrow g(a) = \exp(a)$$

4. Specify distribution for random effects

$$\varepsilon \sim Normal(0, \sigma_{\varepsilon}^2)$$

= General linear model + mixed effect(s)

#### How to estimate standard errors?

- Estimate the "Hessian" at the log-marginal likelihood

$$H(\mathbf{\theta}; \mathbf{y}) = \begin{bmatrix} \frac{\partial^2 \ln L(\mathbf{\theta}; \mathbf{y})}{\partial \theta_1^2} & \frac{\partial^2 \ln L(\mathbf{\theta}; \mathbf{y})}{\partial \theta_1 \partial \theta_2} \\ \frac{\partial^2 \ln L(\mathbf{\theta}; \mathbf{y})}{\partial \theta_1 \delta \theta_2} & \frac{\partial^2 \ln L(\mathbf{\theta}; \mathbf{y})}{\partial \theta_2^2} \end{bmatrix}$$

Calculate its inverse

$$\widehat{\mathbb{V}}(\mathbf{\theta};\mathbf{y})=\mathbf{H}^{-1}$$

Extract element and take square root

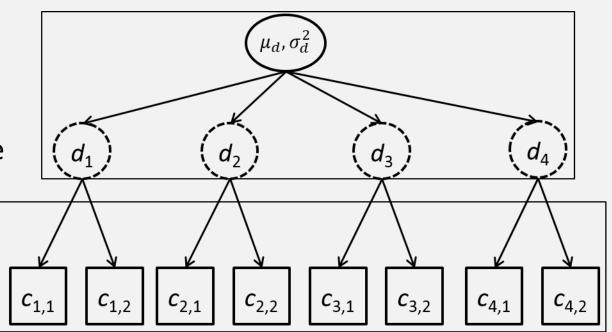
$$\widehat{SE}(\theta_i; \mathbf{y}) = \sqrt{\widehat{\mathbb{V}}(\mathbf{\theta}; \mathbf{y})_{i,i}}$$

#### Example – Hierarchical count samples

$$\log(d_j) \sim Normal(\mu_d, \sigma_d^2)$$
$$c_{i,j} \sim Poisson(d_j)$$

#### Counts

- 4 sites
- 2 observations/site
- 3 fixed effects
- 4 random effects



### Example – Hierarchical count samples

$$\log(d_j) \sim Normal(\mu_d, \sigma_d^2)$$
$$c_{i,j} \sim Poisson(d_i)$$

#### **Questions:**

- 1. What is the mean of  $d_j$  across all sites j?
- 2. What is the variance of  $d_i$  across all sites?
- 3. What is the mean of  $c_{i,j}$  across all sites j and samples i?
- 4. What is the variance of  $c_{i,j}$  across all sites and samples?

- Simulating data
  - [See R code]

### Fit using R

- Using *lme4* package
- formula: way to specify model
- 1. Linear model *lm(formula= ... )* 
  - Count ~ 0 + factor(Site)
  - "Count" response variable
  - "0" Don't include intercept
  - "factor(Site)" Include a fixed effect for each site
- 2. Linear mixed model *lm(formula = ... | ... )* 
  - Count ~ (1 | factor(Site))
  - "(1 | factor(Site))" Include a random effect for each site

### Fit using R

– [See R code]

### Fit using TMB

### Steps during optimization

1. Write joint log-likelihood  $Pr(y, \varepsilon | \theta)$  in CPP file

$$f(\theta, \varepsilon) = \log(\Pr(y|\theta_1, \varepsilon) \Pr(\varepsilon|\theta_2))$$

- 2. Choose initial values for fixed  $\theta_0$  and random  $\varepsilon_0$
- 3. "Inner optimization" Optimize random effects with  $\theta_0$  held constant

$$\hat{\varepsilon} = \operatorname{argmax}_{\varepsilon} (f(\theta_0, \varepsilon))$$

4. Calculate Laplace approx. for marginal likelihood of fixed effects

$$\ln L(\theta_0; y) \cong f(\theta_0, \hat{\varepsilon}) - \frac{1}{2} \log(|\mathbf{H}|)$$

- TMB also provides the gradient of the penalized likelihood with respect to fixed effects
- 5. "Outer optimization" Repeat steps 2-3
  - Outer optimization is done in R using the function value and gradient provided by TMB

### Fit using TMB

[See R code]

- Benefits of using linear mixed models
  - Separate estimate of measurement and between-site variability
  - Include covariates for either one
  - Improved precision
  - "Shrinkage"

- Draw-backs
  - Biased if random effects aren't "exchangeable"

#### Restricted maximum likelihood models (REML)

- Maximum likelihood (ML) estimates of variance parameters are biased
  - ML estimate  $\hat{\sigma}_{ML}^{2}$

$$\hat{\sigma}_{ML}^{2} = \frac{1}{n_i} \sum_{i=1}^{n_i} (y - \hat{\mu}_i)^2$$

Expectation

$$\hat{\sigma}_{unbiased}^2 = \frac{1}{n_i - 1} \sum_{i=1}^{n_i} (y - \hat{\mu}_i)^2$$

- Same problem arises for variance estimates of random effects
- REML gives unbiased estimates of random-effect variances
  - Also sometimes helps convergence
  - Important when log-likelihood function is correlated with respect to random and fixed effects

#### Confidence interval:

- Parameter estimates are normally distributed
- Computation

$$CI_{\chi\%}(\hat{\theta}) = \hat{\theta} \pm \widehat{SE}(\hat{\theta}) \times \Phi^{-1}(\frac{x}{2})$$

- Where  $CI_{x\%}$  contains the true value x% of the time if the model is correct
- $-\Phi^{-1}$  is the inverse cumulative distribution for a normal distribution
- $\hat{\theta}$  is the estimate for parameter  $\theta$
- $-\widehat{SE}(\widehat{\theta})$  is the estimated standard error for parameter  $\theta$

#### Confidence interval coverage

Coverage – the expected proportion of times that an estimated x% confidence interval contains the true value given an estimation model and true "data-generating process"

#### **Estimation:**

- 1. Simulate data with a known value for parameter  $\theta$
- 2. Record true parameter values
- 3. Apply estimator
- 4. Record confidence interval  $CI_{x\%}(\widehat{\theta})$  for parameter  $\theta$
- 5. Repeat steps 1-4 hundreds of times
- 6. Compute the proportion of times where  $CI_{x\%}(\widehat{\theta})$  contains the true value for parameter  $\theta$

#### **Separability**

 What if different components of the model are statistically independent?

$$\Pr(y|\theta_1, \varepsilon) \Pr(\varepsilon|\theta_2) = \prod_{i=1}^{N} \Pr(y|\theta_1, \varepsilon_i) \Pr(\varepsilon_i|\theta_2)$$

- Examples:
  - Overdispersed samples

$$C_i \sim Poisson(\lambda_i)$$
  
 $\log(\lambda_i) \sim Normal(\mu, \sigma^2)$ 

– Each  $\lambda_i$  is independent conditional on  $\mu$ ,  $\sigma^2$ 

$$\Pr(C|\lambda)\Pr(\lambda|\mu,\sigma^2) = \prod_{i=1}^{N} \Pr(C_i|\lambda_i)\Pr(\lambda_i|\mu,\sigma^2)$$

### **Separability**

Then we can factor the integral

$$\int \Pr(y|\theta_1,\varepsilon) \Pr(\varepsilon|\theta_2) d\varepsilon = \prod_{i=1}^N \int \Pr(y|\theta_1,\varepsilon_i) \Pr(\varepsilon_i|\theta_2) d\varepsilon_i$$

 Where we replace a N-dimensional integral with N 1dimensional integrals

#### Uses

- 1. Meta-analysis: species are often independent
- 2. Time series: years are often "conditionally" independent

[Explore "map" argument to TMB]